

A Predictive Test of Voters' Economic Benchmarking: The 2013 German Bundestag Election*

Mark Andreas Kayser [†] Arndt Leininger [‡]

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Abstract

Do voters judge their national economy relative to economic performance abroad? In 2013 we took advantage of the German Bundestag election to test this hypothesis predictively. Nearly two months prior to the election, we published an election forecast relying on a theory-driven empirical model of election outcomes that draws on previous election outcomes; characteristics of the government and of voters; and, most originally, the relative economic performance of Germany (“benchmarked” growth) in comparison to the three other most important economies in Europe, France, the UK and Italy. Our forecast put the outgoing coalition government of CDU/CSU and FDP at 47.05% of the popular vote deviating from the actual outcome of 46.3 by 0.75 points. This makes our forecast one of the most accurate in this election cycle. Despite one and a half months of lead time, our forecast performed on par or slightly better than the last poll results issued only two days before the election.

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[†]Hertie School of Governance, Berlin, e-mail: kayser@hertie-school.org

[‡]Hertie School of Governance, Berlin, e-mail: a.leininger@phd.hertie-school.org

*“Forecasts are difficult – especially if they concern the future.”*¹
– Karl Valentin

1 Introduction

Predictive validity is a cornerstone of science, more difficult to achieve than explanation (Hempel, 1963) and correspondingly all too often neglected in the social sciences (Rosenberg, 1992; Schrodt, 2014; Friedman, 1953). In the run-up to the German federal election of 2013, forecasts and prognostications about who would win attracted considerable attention, both within Germany and beyond. We entered the fray, using it as an opportunity to test a recent hypothesis about how voters form economic evaluations about government performance.

The economy perennially serves as one of the most important determinants of electoral outcomes (Nadeau, Lewis-Beck and Bélanger, 2013) but precisely how voters form their economic assessments is less clear. Recently, Kayser and Peress (2012) claimed, and Hansen, Olsen and Beck (2013) experimentally confirmed – in the case of Denmark and Sweden – that voters judge national economic performance in comparison to that abroad. Such effects emerge in observational data because the media report more positively on the domestic economy when it outdistances its neighbors. Both the observational and experimental data of these respective studies, however, sought explanation rather than prediction. Here we offer a different test, an election forecast estimated two months prior to the 2013 Bundestag election based on a spartan model of voting behavior that incorporates only benchmarked (i.e., comparative) economic growth, previous vote share, terms in office and partisan identification. To set our results in the context of what constitutes an accurate prediction, we contrast our model’s performance with other forecasts and polls,² also “post-predicting” past elections out of sample and calculating average

¹“Prognosen sind schwierig – vor allem, wenn sie die Zukunft betreffen” quote ascribed to German cabaret artist Karl Valentin (1882-1948)

²As we address below, polls are snapshots in time (Jennings and Wlezien, Forthcoming 2015) rather

prediction error.

There are many ways to forecast an election. Leaving casual punditry aside, prognosticators can turn to structural models, survey polling – asking respondents whom they will vote for or who they think will win – political betting markets, poll averaging techniques and more recently, synthetic models (Erikson and Wlezien, 2014; Lewis-Beck and Dassonneville, 2015) that mix structural forecasts and polls. The 2013 Bundestag election saw a welcome expansion in the number of election forecasts in a country that has historically had few.³ In contrast to many of the other forecasts, we constructed a structural forecasting model that made no attempt at improving methods (cf., for example, Selb and Munzert, 2013; Graefe et al., 2014; Jérôme, Jérôme-Speziari and Lewis-Beck, 2013); rather we employed simple linear regression with some of the most conventional vote-choice predictors identified in electoral research. We added only a single innovation intended to test whether voters use cross-national comparisons to assess the economy: in place of economic growth, we employed the deviation in German real GDP growth from that of its largest European benchmarks, the UK, France and Italy.

More specifically, our forecast, released in early August 2013 (FOCUS Online, 2013), relies on a theory-driven forecasting model of election outcomes that draws on previous election outcomes; characteristics of the government and of voters; and, most originally, the relative economic performance of Germany (“benchmarked” growth). It put the outgoing coalition government of CDU/CSU and FDP at roughly 47.05% of the popular vote,⁴ overshooting the actual result by just .75 percentage points. This compares rather favorably to the last poll results issued only two days before the election as well as to other forecasting models. When we consider the average out-of-sample prediction error of the nine most recent Bundestag elections – single forecasts can, of course, be misleading

than predictions of outcomes on election day. We account for this by evaluating our forecast in comparison to polls taken shortly before (usually two days prior to) the election

³The notable exception are Gschwend and Norpoth (2001) and Jérôme, Jérôme-Speziari and Lewis-Beck (2009) whose Chancellor and Political Economy models were the only the only serious forecasts for several elections.

⁴As originally reported in our early August forecast (FOCUS Online, 2013).

– our benchmarking model also outperforms other models, including an identical model using non-benchmarked growth.

Forecasting models have already made important theoretical contributions to electoral research. Most notably, the volatility of political polls over a campaign contrasts sharply with the ability of fundamentals-based structural models to predict outcomes months ahead of an election. This has highlighted the often ephemeral effects of events in election campaigns and led to insights into when and how voters receive information and form voting preferences (Jennings and Wlezien, Forthcoming 2015). As elections near, they both pay more attention and weight “fundamentals” such as the economy more heavily (Gelman and King, 1993). As Mayer (2014) points out, election forecasting has contributed greatly to our understanding of the retrospective nature of the economic vote. We hope that forecasts can similarly contribute to fundamental insights on the effect of cross-national economic benchmarking on vote choice.

2 Benchmarking and the Economic Vote

The effect of the economy on the vote is one of the most researched relationships in the study of politics. Scholars, using a broad variety of objective economic variables as well as surveys of perceived economic performance have assembled strong evidence across national and institutional contexts demonstrating that voters punish the party in power when the economy sours but reward it when the economy expands (Nadeau, Lewis-Beck and Bélanger, 2013; Duch and Stevenson, 2008). Careful estimates associate a unit change in economic growth with between a .8 and 1.4 percentage point increase in the incumbent party’s vote share (Becher and Donnelly, 2013).

Yet, for all of the scholarly confidence about the economic vote, considerable uncertainties remain about how it arises and manifests itself (Anderson, 2007; Kayser, 2014). How do voters attribute responsibility to specific governing parties (Fortunato and Stevenson,

2013; [Duch and Stevenson, 2013](#))? Do they punish governments more for downturns than they reward them for expansions ([Soroka, 2006](#); [Stanig, 2013](#))? How do institutions and political arrangements that clarify or obfuscate party responsibility for outcomes influence attribution and voting ([Anderson, 2006](#); [Hellwig, 2012](#); [Hobolt, Tilley and Banducci, 2012](#); [Powell and Whitten, 1993](#))? Do voters' abilities ([Gomez and Wilson, 2006](#)) and partisanship ([Kayser and Wlezien, 2011](#); [Eggers, 2015](#)) influence their response to economic change? Do voters hold governing parties accountable for developments, economic and other, beyond the government's control ([Healy, Malhotra and Mo, 2010](#); [Hellwig, 2001](#))?

One fundamental issue is how voters evaluate the economy in the first place. A growth rate of two percent would be welcomed as strong growth in much of Europe today but looks anemic in the context of historic growth rates of earlier decades and would constitute an economic crisis in contemporary China. [Palmer and Whitten \(1999\)](#) addressed this issue eloquently with respect to comparisons over time but it is only recently that scholars have considered how the performance of foreign economies affects voters assessment of their own. Cross-jurisdiction benchmarking has been an established empirical regularity in the study of economic policies, especially tax rates, at least since the seminal work on "yardstick competition" by [Besley and Case \(1995\)](#). Only recently, however, has it gained attention in empirical voting studies when [Kayser and Peress \(2012\)](#) demonstrated that benchmarked growth – the deviation in the growth rate from that of comparison countries – predicts the vote for lead governing parties more strongly than non-decomposed growth.

[Duch and Stevenson \(2010\)](#) had already argued that voters are sufficiently informed to compare the variance in economic outcomes across countries when assessing governmental performance and, indeed, [Hansen, Olsen and Beck \(2013\)](#) have shown in an experimental setting that when voters are informed of a comparison country's performance, they use it as a benchmark for holding their own governing parties accountable. [Kayser and Peress \(2012\)](#), however, are more skeptical of voter information levels and show, using data

from the Times of London that the press reports more positively on the economy when it outperforms that of other countries, raising the possibility of “pre-benchmarking”. Voters need not know anything about the performance of other economies so long as media reports on the economy influence their vote. Using a dataset of 32 newspapers from 16 countries in six languages [Kayser and Peress \(2015\)](#) indeed show that the effect of economic growth (but not of inflation or unemployment) on the vote is heavily mediated by media reporting.

Such results are encouraging but the social science are littered with false positives ([Choudoin, Hays and Hicks, 2015](#)) and the danger of confirmation bias ([Nickerson, 1998](#)) lurks for scholars in all areas. Prediction offers an alternative means to test for cross-national benchmarking in the economic vote and the run-up to the 2013 Bundestag election offered promising conditions. The rate of real GDP growth in Germany in the months prior to the election (.8% in the second and .3% in the third quarter) ([Federal Statistical Office, 2015](#)) was low compared to prior periods nevertheless higher than the average of other large European economies still suffering from the aftermath of the Financial Crisis, the Euro Crisis and/or fiscal austerity. If German voters judged the economy only in absolute terms or in comparison to past growth, one would expect cross-nationally benchmarked growth to emerge as a significantly weaker predictor than simple real GDP growth and the forecasting model to predict the outcomes poorly. If they indeed respond to relative growth, then benchmarked growth should predict the vote more strongly than non-benchmarked growth and the model should predict the election accurately. Of course, what constitutes accurate prediction must itself be benchmarked against other predictions. We therefore survey other forecasting models next.

3 Other forecasting models

Electoral forecasting of the sort presented here, just like polling, originated in the context of US presidential elections. While a small cottage industry has developed a variety of models that perform rather well in the United States (see, for example, the 13 forecasts in the October 2012 issue of *PS: Political Science and Politics*) election forecasting in Germany is still in its infancy. Nevertheless, it is striking that to the best of our knowledge our “benchmarking” model was only the third model to be proposed for forecasting the 2013 German federal election – the others being the “Chancellor model” first proposed by [Gschwend and Norpoth \(2001\)](#) who also provide a prediction for the upcoming election ([Norpoth and Gschwend, 2013](#)) and a “Political Economy Model” by [Jérôme, Jérôme-Speziari and Lewis-Beck \(2013\)](#). [Selb and Munzert \(2013\)](#) released a fourth forecast a bit later than ours.

Election forecasting in Germany was pioneered by [Gschwend and Norpoth \(2001\)](#).⁵ Their forecasts provide a three-months lead and they have a very good track record predicting the 2002, 2005 and 2009 elections. In 2002 they forecast the outgoing red-green government vote share exactly right. Their model includes chancellor support (percentage of people favoring the chancellor over his or her challenger), long-term partisanship (average of past three election results of coalition parties) and the log of the number of terms the chancellor’s party has been in government. Estimating their model over all past 16 elections and fitting it to up-to-date data they provided not one but three slightly differing forecasts for the 2013 election. These forecasts were, in chronological order, 51.7% vote share for the outgoing government, published in July (but with data from April) in *PS: Political Science and Politics* ([Norpoth and Gschwend, 2013](#)), 49.7%, published on 26 July 2013 in a blog by the German weekly *Die Zeit* ([Zeit Blog, 2013a](#)), and 51.2% in another blog post on 24 August 2013 ([Zeit Blog, 2013b](#)). All forecasts were produced using

⁵[Jérôme, Jérôme-Speziari and Lewis-Beck \(2009\)](#) have also developed a forecasting model for German elections that they have used to forecast elections since 1998. They have published their forecasts in the French press and have received little attention in Germany.

the same model, forecasts differ slightly because each time they use updated polling data on chancellor approval. Technically, their model is similar to ours but also their forecast is the farthest off of all models we review. They attribute this to the unforeseen success of the Alternative für Deutschland, a eurosceptic party founded just months before the election, and the exceptional popularity of chancellor Merkel, which extends well beyond her own party, diminishing the predictive quality of chancellor support ([Norpoth and Gschwend, 2014](#)).

[Jérôme, Jérôme-Speziari and Lewis-Beck \(2013\)](#) have produced many forecasts of different national elections. For 2013 they presented a new ‘Political Economy Model.’ They use a seemingly unrelated regression (SUR) model to estimate the vote shares of all parties represented in parliament. Their variable choice is also informed by electoral research and includes an economic measure like unemployment but also political variables like chancellor approval as well as vote intention polls. Their forecast of the coalition’s vote share is on par with ours but their estimates for the opposition are slightly off. Their economic predictors are measured one to two quarters before the election whereas some of their other political predictors are in principle measurable even farther in advance, for instance whether the FDP was an outgoing party with the CDU. They calculated their forecast in April 2013 using vote intention polls available at the time (used to forecast vote shares of the Green Party, Die Linke and the combined vote share of all other parties) which gives their model a lead of about five months.

Selb and Munzert’s ([2013](#)) approach is similar to the above mentioned ones in using data from prior elections to arrive at a forecast well ahead of the election. They differ in that they exclusively rely on polling data. However by estimating the relationship between pre-election polling and actual results they hope to eliminate the biases inherent in pre-election polling which might be party- or polling-company specific and therefore arrive at a more accurate forecast than the polls they use. Interestingly they find, based on out-of-sample predictions, that they obtain the most accurate forecasts with polls taken

six months before the election. Accuracy decreases when using polls taken closer to the election.

3.1 Polling

Polls differ markedly in technique and theory from forecasts; in fact they are atheoretical and ahistorical. What they provide is a snapshot in time, not more and not less. Thus, a poll, say, two months prior to an election, is not technically a forecast of the election, although the media and voters often treat it as such (Hillygus, 2011). “Election-eve” polls, however, are much closer to predictions and, consequently, serve as a reasonable comparison for the accuracy of our forecast.

Polling companies this year provided surprisingly accurate forecasts and substantially improved their performance compared to previous elections. Forschungsgruppe Wahlen for instance, the non-profit company behind the ZDF’s “Politbarometer”, provided estimates of the coalition vote share that hovered between 45 and 47 percent since May 2013 (see Figure 4). 2013 was an exceptionally good year for polling in Germany; in prior years even final pre-election polls have deviated considerably from the actual result. Schnell and Noack (2014) recently evaluated 232 pre-election polls from 1957 to 2013 and found that conventional 95% confidence intervals only contained the actual election result 69% of the time. In the prior election in 2009, for instance, when the outgoing government was a grand coalition between the CDU/CSU and the SPD the final pre-election “Politbarometer” overestimated both the vote share of the CDU/CSU and the SPD leading to an error of 4.2 percentage points. This occurred partly because they underestimated the extent of “vote lending” from CDU to FDP – CDU supporters giving their second vote (“*Zweitstimme*”) to the FDP.⁶ They made no such mistake this time providing reasonable estimates for both CDU/CSU and FDP. This might be due to improved methodology, however, the reform of the apportionment system which elimi-

⁶These might be better described as ‘rental votes’ (Meffert and Gschwend, 2010).

nates the so-called “Überhangmandate” and therewith any electoral advantage of ticket splitting may have been more important in this regard. All point predictions from polls suggested an FDP vote share of 5% or more, while the FDP, in fact, fell just below the 5% hurdle. The actual result for the FDP, however, is within the poll’s margin of error.

4 Our model

Election forecasting in multi-party systems with proportional representation entails unique challenges, these include multiple parties, coalition governments and strategic and mechanic effects of electoral thresholds. Most people are interested in forecasts because they want to know who will win the election. However, in a multi-party system like Germany’s the identity of the winner also depends on the outcome of post-election coalition bargaining. Since it is difficult to predict winners or losers, it makes sense to focus rather on parties’ vote shares, a determinant of the set of feasible governing coalitions. Specifically, we focus on the vote share (to be) obtained by the outgoing governing coalition. The coalition vote share is the percentage of the popular vote received by the parties forming the governing coalition. This is usually the sum of the vote shares of two or more parties – only once (in 1961) has Germany seen an outgoing single party government and even then only when one counts the CDU and CSU as one party. We focus on the coalition’s vote share, rather than on that of individual parties, because it is of the greatest substantive interest. With this forecast, we are able to predict whether the same government will continue in power.

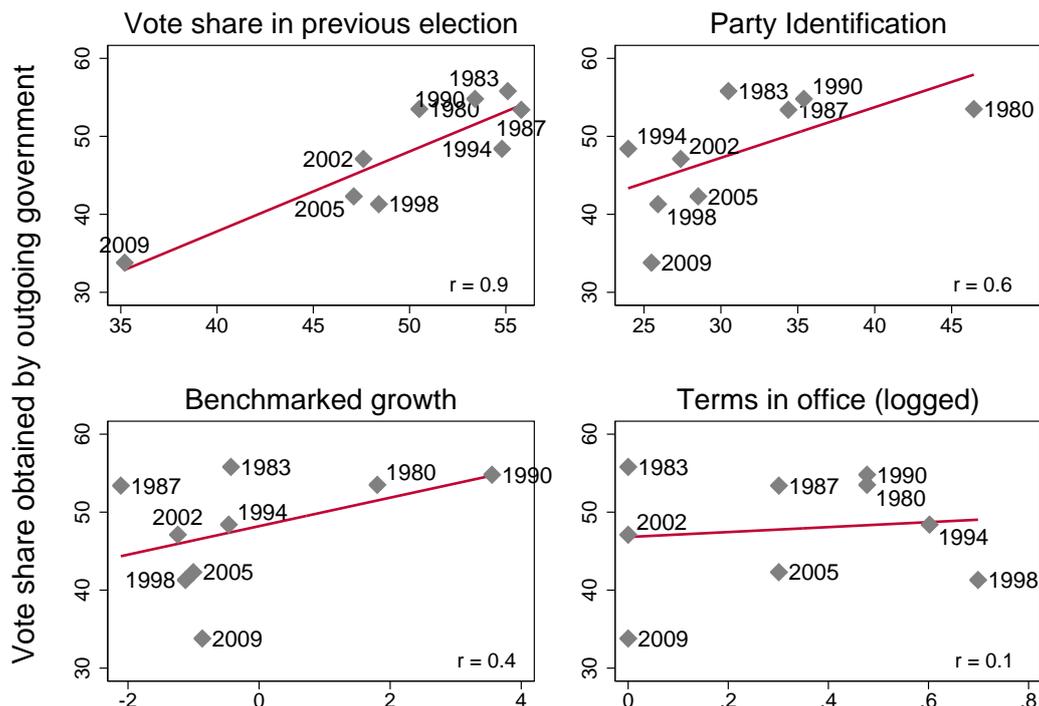
We thus make the assumption that a governing coalition wants to and will stay in power if it secures the necessary votes to re-capture a parliamentary majority. Grand coalitions are the exception here. They are mere coalitions of necessity (“*Staatsräson*”) when no other options involving a larger and one or more smaller parties seem feasible. Usually neither coalition partner has an interest in continuing the coalition beyond the next election. Furthermore, voters have no credible alternative government in opposition. Our

modeling choice – to use only the larger party’s vote share – was therefore motivated by theory. The larger party in both of Germany’s grand coalitions (1969 and 2009) was the CDU/CSU. For reasons we discuss further below our time-series is limited to 1980-2013 and therefore only includes the latter. When faced with a coalition government, voters attribute responsibility primary to the party of the prime minister (Duch and Stevenson, 2008, 2013), which in Germany has always been is the larger coalition partner. Norpoth and Gschwend in their forecasting model also rely on the larger party’s vote share for the same reasons.

Common predictors of a government’s electoral fortune work well for a time-series composed of ‘regular’ coalitions and the larger party of a grand coalition but not if the full vote share of a grand coalition is used. An alternative would be to use the grand coalition vote share but include a grand coalition dummy variable. However, as our time-series only contains one grand coalition, the dummy simply captures the difference in expected outcomes between grand-coalition and regular governments. In effect removing the grand coalition from our time-series seems rather arbitrary to us. It also generates a fundamental problem for forecasting models. If we were to forecast an election result given an outgoing grand coalition – which is in fact what the situation will be in 2017 – our estimate would also be based on a model containing a coefficient that is based on solely the deviation of one single election.

Four explanatory variables appear in our model: (1) the vote share received by the current governing parties in the previous election, (2) the proportion of people identifying with one of the governing parties, (3) the difference between Germany’s growth rate and the “benchmark”, i.e., the average of the growth rates in France, the UK and Italy, and (4) the (log of the) number of terms a government has been in power. See Figure 1 for plots of our dependent variable against our predictors. With the exception of the economic benchmarking variable, an innovation for an election forecasting model, we chose these variables for their broad acceptance as predictors of the vote in the voting literature.

Figure 1: Predictors of the vote



We include the *vote share of the current incumbent parties in the previous election* (even if they were not in government then) to form a baseline prediction. Past outcomes are a strong predictor of future outcomes, so controlling for previous vote share effectively focuses the other predictors on changes from the previous vote share. The combined vote share of the parties making up the outgoing government correlates strongly with their results in the previous election ($r = 0.88$). This is also because many people exhibit strong partisanship, leading them to vote for the same party in successive elections.

Our second variable, *party identification*, captures the proportion of respondents expressing an attachment with one of the governing parties. Our data come from the Politbarometer.⁷ The question put to respondents reads (our own translation), “Many people in Germany lean toward a certain party, although they might vote for a different party from time to time. What about you? Do you – generally speaking – lean toward a certain

⁷We thank the *Forschungsgruppe Wahlen* for providing us with their 2013 aggregate party identification data.

party?”⁸ Party identification does not imply a formal attachment but, rather, simply feeling close to a party. However, since the number of people with a party identification still changes over the medium and long term, the proportion of people identifying with one of the governing parties correlates significantly with the vote ($r = 0.61$). As party identification has declined over the years, by about .6 percentage points per year (Arzheimer, 2012), so has the vote share obtained by the outgoing government. In the US setting, partisan identification is often the strongest predictor of vote choice (Campbell et al., 1960). In Europe, scholars debate its theoretical and empirical grounding in the context of multi-party systems but empirical tests have demonstrated its stability in individuals over time and its sound performance as a predictor in voting models (Ohr and Quandt, 2012, section 2.2). As party identification aggregates vary from month to month by about 3%-points due to sampling error and idiosyncratic events (Arzheimer, 2012) we take the average partisan identification for the governing parties in the six months before the given election – for an election in September this would be the months of February through July. Since monthly party identification data is only available for the elections from 1980 onward, we can only use elections since then in our analysis (for an overview see Table 1).

The third variable is the most novel and also the namesake of our “benchmarking model”. We calculate Germany’s growth, as benchmarked against the three next most important economies in Europe: the UK, France and Italy. Data are from the World Bank’s World Economic Outlook which goes back to 1980 and includes a forecast for 2013.⁹ Implicitly by using the deviation of German growth rates from the average of British, French and

⁸“In Deutschland neigen viele Leute längere Zeit einer bestimmten politischen Partei zu, obwohl sie auch ab und zu eine andere Partei wählen. Wie ist das bei Ihnen: Neigen Sie – ganz allgemein gesprochen – einer bestimmten Partei zu?” If respondents answer in the affirmative they are asked for the party they lean toward and the strength of that orientation toward the named party. We use two of the three questions used to measure party identification in the Politbarometer survey. We do not use the third question asking about the intensity of a citizen’s affinity to the party, that is we do not weigh respondents by their answers to this question.

⁹Kayser and Leininger (2015) show that real-time economic data yield smaller forecast errors but in short time-series such as ours little difference between forecasts based on vintage and real-time economic data should emerge.

Table 1: Elections in model

Year	Outgoing Government Coalition	Outgoing government vote share (%)
1980	SPD and FDP	53.5
1983	CDU/CSU and FDP	55.8
1987	CDU/CSU and FDP	53.4
1990	CDU/CSU and FDP	54.8
1994	CDU/CSU and FDP	48.4
1998	CDU/CSU and FDP	41.3
2002	SPD and Greens	47.1
2005	SPD and Greens	42.3
2009	CDU/CSU and SPD	33.8
2013	CDU/CSU and FDP	46.3 (forecast: 47.05)

Italian growth, we presume that voters judge, albeit unwittingly, the state of the economy relative to that of other countries. Evidence of such “benchmarking across borders” comes from [Kayser and Peress \(2012\)](#) who explain the phenomenon with evidence that the media report more positively on the economy when it is outperforming that of comparison countries. This measure of *relative economic performance* is of special importance to the forecast for the 2013 election since German growth is sluggish but looks better when compared to that of other European states suffering from the aftermath of the financial and Euro crises. Indeed, of the 16 national governments up for re-election within an 18 months period from July 2008 to the end of 2009 only nine of them lost vote share.¹⁰ Yet, their economic performance was similarly dismal to that experienced by the OECD in general, with every country experiencing at least two quarters of negative growth in this period and 13 of 16 experiencing negative growth in the quarter of or before the election. Lastly, we rely on the empirical regularity that governments, on average, lose support the longer they remain in office ([Paldam, 1986](#); [Green and Jennings, in progress](#), Ch.5). The major governing parties (CDU/CSU or SPD) in Germany have on average lost 3.2 percentage points per term. We capture this with the \log_{10} of the *number of terms* that

¹⁰The sample is defined as all countries that held a general election or, in the case of the United States a presidential election, between July 2008 and December 2008, which was recorded in the election appendix of the European Journal of Political Research. They are Austria, Bulgaria, Canada, Germany, Greece, Iceland, Israel, Japan, Luxembourg, Lithuania, New Zealand, Norway, Portugal, Romania, Slovenia, and the United States.

a government has been in office.

4.1 Model estimates

Putting all variables in one linear regression model estimated over the elections 1980-2009 we can explain 99% of the variance in the vote share of the outgoing government in the past 9 elections (see Table 2). An R^2 of this magnitude raises worries about possibly fitting noise but we were hesitant to change our specification due to too good a fit. All coefficients are statistically significant and have the expected sign. Our first predictor is the ‘Previous vote share’ – the vote share obtained by the governing parties in the prior election. Controlling for previous vote share effectively focuses the other predictors on changes from the previous vote share: the coefficient on previous vote share is 1.025 implying that, fixing all other predictors at zero, a government should obtain roughly the same vote share as in the previous election minus 8%-points (the constant). Of course we cannot set all other predictors to zero. An outgoing government has to spent at least one term in office. We include the common log of the terms in office to capture the empirical regularity that governments, on average, lose support the longer they remain in office (Paldam, 1986; Green and Jennings, in progress) As the vote loss per term increases non-monotonically we log the variable. The coefficient on ‘Log Terms’ is -8.881. Consequently, the expected decrease in vote share after a secon term in office, holding all other predictors constant, would be 2.67%-points.

Party identification, as expected, has a positive sign. Party identification is a strong predictor of the vote, so that at the aggregate the more people identify with a party the greater a vote share it obtains - even when controlling for previous vote share. The coefficient on party identification is positive and significant despite the fact that we include past vote share. This indicates that party identification at the aggregate does change over the medium term, in-between elections. Every percentage-point increase (decrease) in party ID is associated with an increase (decrease) in vote share of roughly a fifth of a

percentage point, holding all other predictors constant.

Benchmarked growth lastly provides predictive validity for the existence of economic voting in Germany, more specifically the benchmarking mechanism outlined by (Kayser and Peress, 2012). A one percentage-points difference in the growth of Germany and the average of growth rates in France, Italy and the UK is associated with an increase in vote share of roughly one percentage-point.

Table 2: A regression model of past elections to predict future elections

Predictors	Coefficient	(S.E.)
Previous Vote Share	1.025***	(0.0544)
Party ID	0.276**	(0.0563)
Benchmarked Growth	0.930*	(0.224)
Log Terms	-8.881**	(1.295)
Constant	-8.708*	(2.724)
Observations	9	
R-squared	0.993	
Adj. R ²	0.986	
RMSE*	1.424	
Durbin-Watson d	2.561	

Standard errors in parentheses

*** p<0.001, ** p<0.01, * p<0.05

* RMSE calculated from 9 out-of-sample predictions for elections 1980 to 2009.

Note: Model estimated on elections 1980-2009

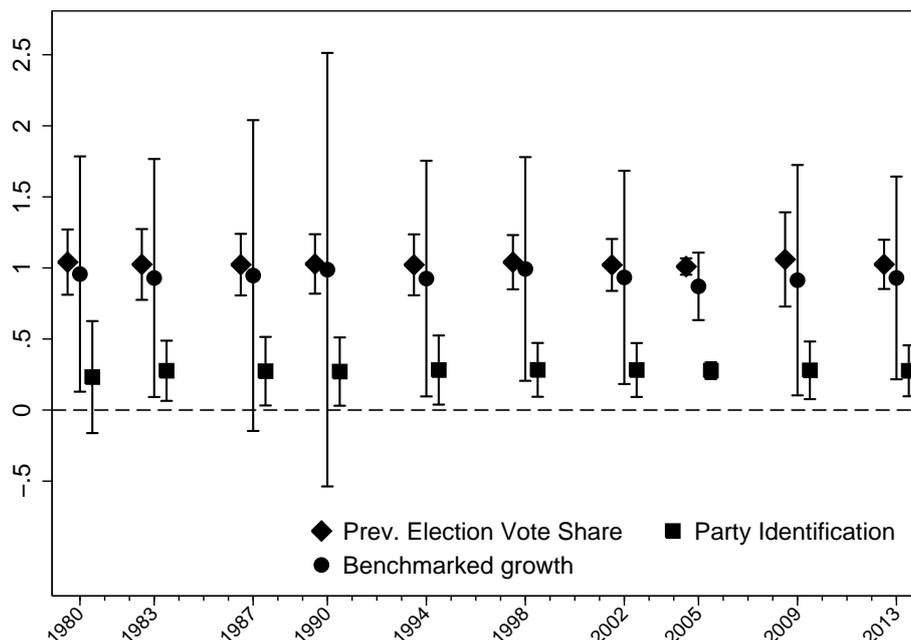
Only once we predict elections outside our sample, however, do we get an idea of our model’s predictive validity. The 2013 election provided an essential test in that regard. In hindsight we know that it fared quite well, but when we constructed our model in early summer 2013 we had to rely on other means to test our model’s ability to make out-of-sample predictions. We did so with the help of out-of-sample predictions. By omitting one election, re-estimating the model on the remaining elections and calculating a prediction using the values of the omitted observation we ‘post-predicted’ historical elections to test our model.¹¹ We did this for all 9 previous elections, compared our

¹¹Due to the size of our dataset (n=9) we were not able to rely on consecutive out-of-sample forecasts

prediction to the actual outcome, squared the differences, averaged them and, finally, took the square root to obtain the root mean square error (RMSE).

The RMSE gives us an estimate of the ‘average’ error of our model in out-of-sample forecasting. Figure 3 plots the actual vote shares received by the outgoing government against our prediction. The farthest we are off is 1.8 percentage points in 2005; in 1983 we get within one tenth of a percentage point of the actual result. The RMSE for our model is 1.4 percentage points which, considering that the model only rests on nine previous elections, compares rather favorably to standard errors of the regression in surveys involving many more observations. Our coefficients remain stable across all out-of-sample estimations showing that our results are not unduly influenced by outliers and that the effects of our explanatory variables are, as we expected, stable over time (Figure 2).

Figure 2: Stability of coefficient estimates



predicting one observation ahead. Instead, we predicted the omitted election using also elections following the election to be predicted. Consequently, only our 2009 forecast is a true out-of-sample forecast into the future. All other out-of-sample forecast predict a value within or at the beginning of the time-series used to fit the model.

4.2 Our forecast for 2013

When we regress the vote share of the outgoing government on our four explanatory variables we obtain coefficient estimates that allow us to calculate the 2013 vote share of the outgoing government by using up-to-date values for our explanatory variables. Inserting 2013 values for our explanatory variables into the below equation, we obtain a point prediction of 47.05% for the combined vote share of the CDU/CSU and FDP in the 2013 Bundestag election.

$$\widehat{\text{Outgoing government vote share}} = -8.708 + 1.025 * \text{vote share in previous election} + 0.276 * \text{party identification} + 0.930 * \text{benchmarked growth} - 8.881(\log_{10}(\text{terms}))$$

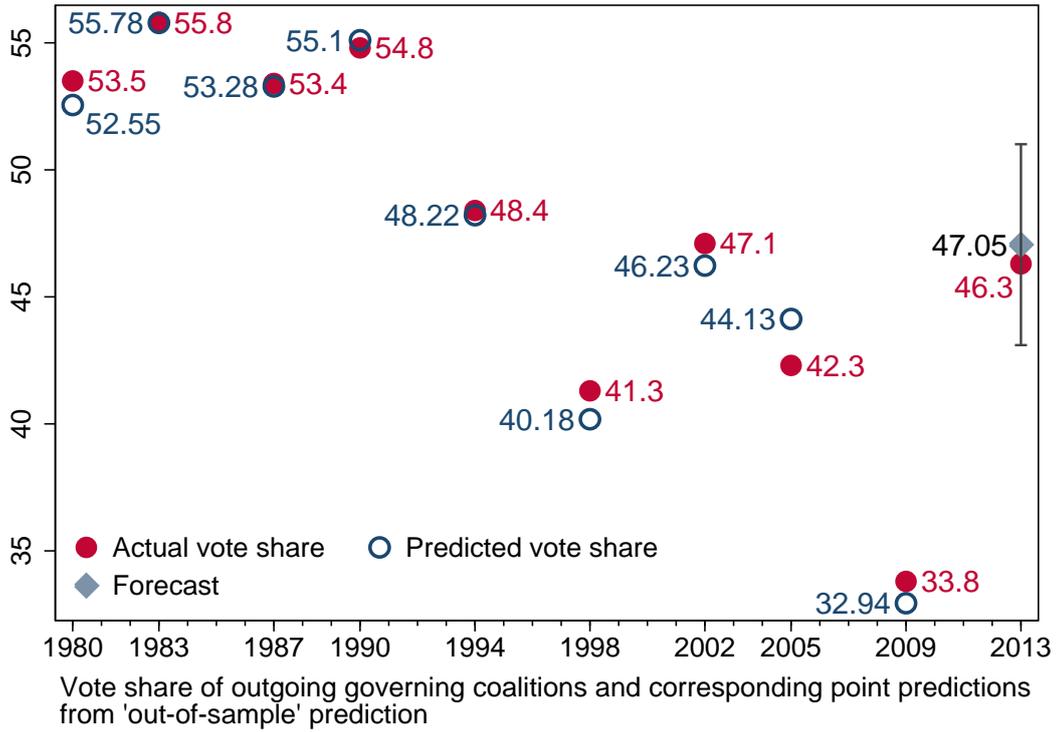
Using the RMSE calculated from our “out-of-sample” predictions we can calculate the probability of the coalition obtaining a majority of seats in parliament necessary to continue in office. As this year’s change to the electoral system in Germany largely eliminates distortions of the vote-seat relationship arising from so called “*Überhangmandate*” we only have to worry about the votes obtained by parties that will not be represented in parliament. Polling prior to the election suggested that about 8%¹² to 9.5%¹³ of votes would go to the Pirates, Alternative für Deutschland and other fringe parties that would likely not exceed the 5% threshold. We assumed that the FDP and Die Linke would clear the threshold. Die Linke was consistently well above 5-percent while the FDP would, so we thought, again profit from “rental votes” from CDU/CSU-voters which polls have difficulty predicting. Of course, as we now know, the FDP stumbled on the 5% threshold. We discuss this issue below.

Considering the substantial error in poll-based projections 2 months ahead of the elections this gave us a range of 6 to 12 percentage points. If the polling for these minor parties was

¹²Forschungsgruppe Wahlen 24.07.2013, Infratest Dimap 25.07.2013

¹³Allensbach 12.07.2013

Figure 3: Actual and predicted vote share



correct, however, about 45.5% of the vote should suffice for Mrs. Merkel to continue the coalition of CDU/CSU and FDP. Our point prediction already was above this threshold; taking into account the statistical error inherent in the forecast, we predicted a 83% probability that the outgoing coalition would have a majority in the next parliament. In hindsight, the complex assumptions necessary to estimate this probability and the fact that the FDP did not, in fact, obtain the minimal 5% vote share necessary to enter parliament, may make our probability estimate seem overly naive as we discuss in section 6. However, the presentation of the election prediction as a probability is a useful exercise because it emphasizes the uncertainty of the point prediction which might be missed if only the point prediction is presented.

5 A comparative assessment

At first blush, our model seems to have fared well. Our prediction of a coalition vote share of 47.05% deviated from the actual vote share (46.3%) by only 0.75 percentage points. Whether this actually constitutes a success, however, depends on the performance of the comparison models. We assess the performance of both our model and of the benchmarked growth variable in three ways. Comparing the performance of our model to a variant replacing benchmarked with non-benchmarked growth (i.e., German GDP growth) enables a direct comparison of benchmarked and non-benchmarked growth, both as a predictor and in forecast errors (section 5.1). Of course, a model that surpassed its non-benchmarked alternative specification but did a poor job of predicting the election would not necessarily be of much use. We therefore place our benchmarking model's performance in context by comparing it to other forecasts of the 2013 election and the final election polls, two days before the election (section 5.2).

5.1 A forecast with non-benchmarked growth

To assess whether benchmarked growth is a superior predictor to real growth in GDP – probably the most common economic measure in structural election forecasting models – we compare two forms of our model, one with benchmarked growth and one with non-benchmarked growth in German real GDP. We are wary about over-drawing conclusions based on the small sample size, the results are nevertheless informative, given the constraints of the data.

Benchmarked growth reveals a stronger effect than its non-benchmarked growth, as shown in Table 5.1, associating a unit change with correspondingly larger increase (.93 v. .69) in predicted vote share for the coalition. Benchmarked growth is also more precisely estimated than its non-benchmarked counterpart. But how did the models fare in predicting out-of-sample elections?

	benchmarked	not benchmarked
Previous Vote Share	1.025*** (0.0544)	0.789** (0.163)
Party ID	0.276** (0.0563)	0.395** (0.0834)
Benchmarked Growth	0.930* (0.224)	
Growth Germany		0.693 (0.426)
Log Terms	-8.881** (1.295)	-9.616* (2.618)
Constant	-8.708* (2.724)	-1.335 (8.194)
MAE	1.01	2.09
RMSE	1.42	2.47
2013 Forecast	47.05	45.79
N	9	9
R^2	0.993	0.978
adj. R^2	0.986	0.956

Table 3: *A comparison of benchmarked and non-benchmarked growth. Standard errors in parentheses. * $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$*

When all elections are predicted, one-by-one, out of sample, benchmarked growth outperforms its non-benchmarked alternative: When growth is benchmarked, the mean average error (MAE) and root mean squared error (RMSE) of the model were 48% and 57% as large, respectively, as when it was not. In other words, the forecast error from the model with the non-benchmarked growth was over twice as large (2.09 v. 1.01) when measured as MAE and 1.74 times as large measured as RMSE.

For the 2013 election, the point forecasts from both models deviated from from the final outcome by less than a percentage point. While the forecast error from the non-benchmarked model was even slightly smaller than that from the benchmarked model they were not statistically distinguishable from each other at normal significance levels. We further note that the non-benchmarked growth model predicts a smaller vote share

than the original model and actually underestimates the actual result. This suggests that in contexts such as the 2013 German election, when Germany had sluggish growth compared to prior periods but relatively high growth rates when compared to other Eurozone economies, using non-benchmarked growth may underestimate the popularity of the government.

In sum, while non-benchmarked growth actually produced a slightly more accurate prediction for 2013, benchmarked growth produced more accurate out-of-sample predictions in most of the other eight elections, yielding an average prediction error roughly half the size of that from the non-benchmarked model.

5.2 Other forecasts and polls

Let us also benchmark our benchmarking model. It is informative to see how a simple structural forecasting model with benchmarked growth performs relative to one with non-benchmarked growth. But this comparison would be of less interest if both models predicted poorly relative to other approaches. How did our benchmarking model fare relative to other forecasts estimated well before the election and polls taken nearly on its eve?

Our forecast put the outgoing coalition government of CDU/CSU and FDP at 47.05% of the popular vote – quite close to the 46.3% it actually obtained.¹⁴ This makes our forecast one of the most accurate among the forecasts that were released prior to the election. Only one other forecast ([Jérôme, Jérôme-Speziari and Lewis-Beck, 2013](#)) provided similar accuracy. Our forecast was also on par or slightly more accurate than the last poll results collected only two days before the election (for a quick comparison of forecasts, final polls and results see [Table 4](#)). It also bested averages of pre-election polls, as well as average predictions of expert assessments and prediction markets. These have been calculated by [Graefe \(2015a\)](#) for his ‘PollyVote’ forecast. It is produced through calculation of simple

¹⁴Taking the additional step of assuming FDP inclusion in the next parliament and estimating the coalition reelection probability appears, in retrospect, to have been less well advised.

averages of the forecasts in four categories of forecasts: polls, prediction markets, expert judgment and quantitative models. Graefe also calculates the overall average across the categories.

He explicitly calculates simple averages because as he argues there are no advantages in terms of accuracy in using more complicated averaging procedures. The idea behind averaging over different forecasting methods is that no individual method that the PollyVote model relies on is consistently better than the other. In fact, in an analysis of six elections from 1992 to 2012 [Graefe et al. \(2014\)](#) found that “methods that provided the most accurate forecasts in one election were often among the least accurate in another election.” For 2013 the PollyVote aggregate was more accurate than any of the four component methods, however it was not the single most accurate model. Thus, overall, our benchmarking model fared well but we need to acknowledge that there is no guarantee that it will remain as accurate in the future.

5.3 Other assessment criteria

Accuracy, of course, is not the only criterion in assessing which forecast fared best. Forecast assessment is multifaceted and many other criteria demand consideration ([Lewis-Beck, 2005](#)). Parsimony, forecasting with a small number of predictors, is valuable where observations are scarce. Lead time – how far ahead of the election the forecast is estimated – is another quite important criterion that we consider in our comparative assessment. After all, what is a forecast worth that is dead accurate but issued only a day before the election? A forecast that is, say, two percentage points off but published three months ahead of the election can arguably be considered a better forecast. There is a tradeoff between lead time and accuracy and people will disagree what constitutes the optimal balance between the two. For this reason, we refrain from presenting a ranking of forecasts. Rather we summarize how the different forecasts performed on all the relevant criteria (see [Table 5](#)).

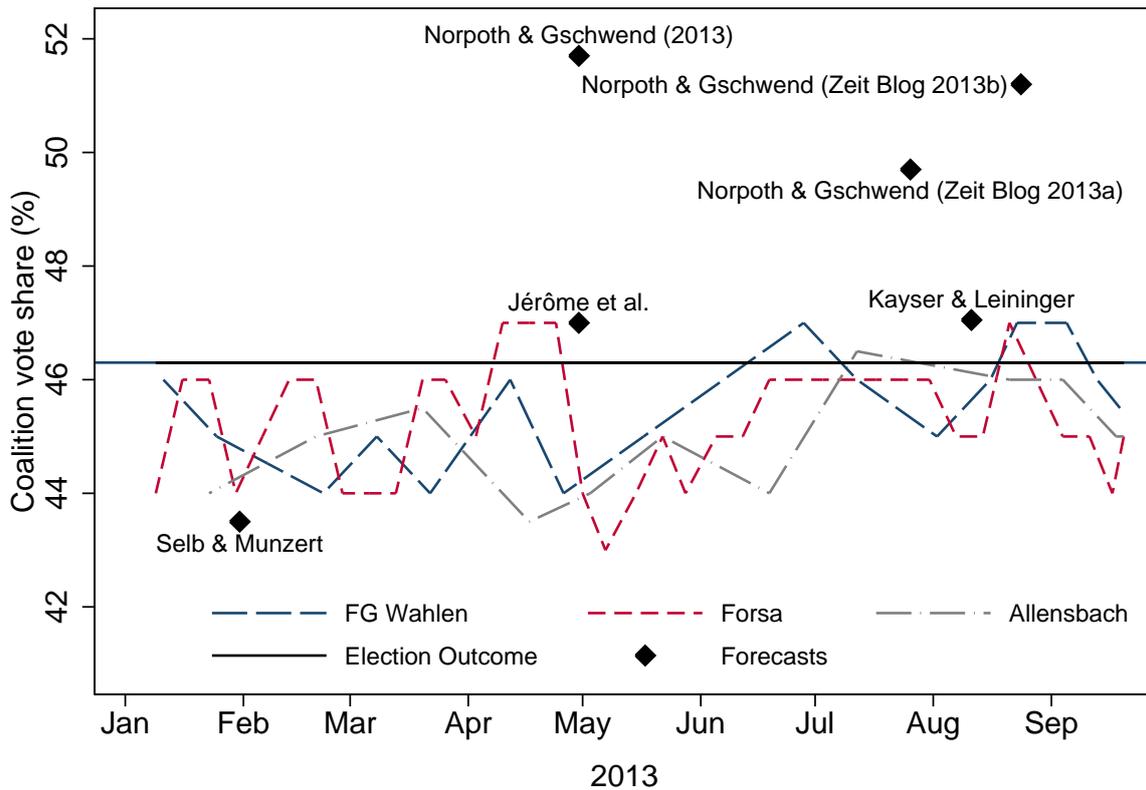
Table 4: Election result, forecasts and final pre-election polls

	Official Re- sults	Kayser & Leininger	Norpoth & Gschwend	Jérôme et al	Selb & Munzert	Forschungs-Forsa, gruppe Wahlen, 09/19	Forsa, 09/20	Allensbach, 09/20
Coalition	46.3	47.05 (.75)	51.2 (4.9)	47 (.7)	43.5 (-2.8)	45.5 (-.8)	45 (-1.3)	45 (-1.3)
CDU/CSU	41.5			41 (-.5)	38.1 (3.4)	40 (-1.5)	40 (-1.5)	39.5 (-2)
FDP	4.8			6 (1.2)	5.4 (.6)	5.5 (-.7)	5 (.2)	5.5 (.7)
SPD	25.7			28 (2.3)	28.2 (2.5)	27 (1.3)	26 (.3)	27 (1.3)
Bündnis 90 / Die Grünen	8.4			10 (1.6)	13.5 (5.1)	9 (.6)	10 (1.6)	9 (.6)
Die Linke	8.6			9 (.4)	7.7 (-.9)	8.5 (.1)	9 (.4)	9 (.4)
Others	10.9			6 (-4.9)	6.5 (-4.4)	10 (-.9)	10 (-.9)	10 (-.9)
MAE	-	.75	4.9	1.82	2.82	.84	.89	.89

Numbers in parentheses indicate difference to official results.

Another feature of election forecasting models that we consider is unique to election forecasts for multi-party systems: the number of vote shares forecast. We predicted only the combined vote share of the parties in the outgoing coalition, CDU, CSU and FDP. Therefore, we were not able to predict the FDP's dropping out of the Bundestag. Models such as those by Selb and Munzert and Jérôme, Jérôme-Speziari and Lewis-Beck seem in principle to enjoy an advantage in that regard, although both models predicted the FDP at above 5%. Intuitively, there is more potential for error when predicting multiple vote shares. To provide a fair comparison of the accuracy of the models we compare their mean absolute errors (MAE). The model by [Selb and Munzert \(2013\)](#) has the greatest lead but an MAE of 2.82 percentage points. Our model has the smallest MAE of .75 percentage points but we also have the smallest lead of just one and a half months and we only predict coalition vote share. Considering that the forecasts were made months

Figure 4: Timeline of election forecasts



Note: Forecasts indicate earliest date that forecast could have been made. For example, although [Selb and Munzert \(2013\)](#) published their forecast in August they use polling up to February of the election year. Their forecast could therefore have been made as early as February.

ahead of the election we believe all of the forecasting models compare quite well to the last pre-election polls issued only days before the election. These were quite accurate but they have practically no lead time (see Table 4).

One model can provide different forecasts, so can have multiple different lead times. For instance Norpoth and Gschwend provided three forecasts in 2013 based on updated data on their poll-based predictor variable chancellor popularity. The limiting factor in our model is also a poll-based variable, party identification in our case. As we published our forecast in early August we chose to incorporate the latest available data. We conducted several out-of-sample forecasts, post-hoc varying the lead time, and found that the average forecasting error tends to increase with lead time. For future elections we could provide

Table 5: Model comparison

Model	Predicted Variables	Predictor Variables	Lead (months)	Accuracy (MAE)
Kayser & Leininger	coalition vote share (1)	previous vote share, party identification, benchmarked growth, log(terms)	1 1/2	.75
Norpoth & Gschwend	coalition vote share (1)	chancellor support, long-term partisanship, log(terms)	1 (last forecast) - 5 (first forecast)	4.9
Jérôme et al.	individual party vote shares (6 parties)	governing party vote share, opposition party vote share, unemployment, chancellor preference, FDP's coalition partner, vote intention polls for various parties, grand coalition, various idiosyncratic dummies	5	1.82
Selb & Munzert	individual party vote shares (6)	vote intention polls	8	2.82

monthly updates to our forecast by changing the measure of party identification – by adding the newest Politbarometer month to the average and dropping the last – with the promise of increasing the accuracy of our forecast as the election draws nearer. Structural models also provide the opportunity to calculate scenarios. Instead of waiting for new data to come in one can use a range of likely values to calculate corresponding forecasts.

6 Discussion: Election Forecasting in Germany

The 2013 election exposed two problems with forecasting models that focus on coalition vote share. First, estimating the probability of a government remaining in power is considerably more difficult than providing a point estimate of its vote share. For the former one has to identify a vote threshold that will likely guarantee a parliamentary majority. We estimated this threshold to be 45.5% given the likely vote shares to be obtained by parties not surpassing the 5% threshold. However to calculate this threshold we ultimately had to rely on the polls that we aimed to outperform. Second, as we did not forecast individual party vote shares, we again had to rely on polls to gauge the likelihood of all coalition parties re-entering parliament. Based on our point prediction we calculated a 83% probability that the outgoing coalition will have a majority in the next parliament. We calculated that probability as a one-sided hypothesis test, treating the effective threshold of 45.5% as null hypothesis. We naively put the FDP's probability of making the threshold at 1, which is obviously wrong as we now know. In hindsight we believe we could have done better. Rather than taking the polls at face value we should have, as one always should with polling data, accounted for the random error. We should have calculated the probability of the FDP surpassing the threshold and included that information in calculating the probability of the government continuing. Models such as those provided by [Jérôme, Jérôme-Speziari and Lewis-Beck \(2013\)](#) and [Selb and Munzert \(2013\)](#) have a clear advantage in that regard as they provide predictions for all parties.

This election also posed a unique challenge to all election forecasters due to the change in the apportionment system. A reform mandated by Germany's constitutional court eliminated the so-called *Überhangmandate*, drastically reducing the incentive for strategic voting. Without the reform, forecasts would likely have been more accurate. The reduction of split-ticket voting ([Bawn, 1999](#)) also goes towards explaining the increased accuracy of pre-election polling.

Minor parties pose another challenge to election forecasters and not only in multi-party

systems – think of third party candidates in the US. They matter more, however, in multi-party systems. Forecasts of US presidential elections usually predict the incumbent party’s share of the two-party popular vote (rather than of the total vote) to account for third party candidates taking some of the vote which in the end is politically inconsequential. Forecasters of German Bundestag elections could similarly forecast the in-Bundestag vote share, the incumbent coalition’s vote share relative to the vote share obtained by all parties surpassing the 5%-threshold needed to be represented in the Bundestag. Such forecasts would be hard to communicate to the public as coverage of the horse race focuses on overall vote share. Presenting such forecasts in seat shares – especially since the reform of the electoral systems promises to make vote share more proportional to seat share – might be more viable.

Another complication that also applies to election forecasts beyond Germany arises from small sample size. For a good forecast we want the error around the prediction to be as small as possible – it should be comparable to confidence intervals in survey research. Given the low number of post-war elections, a forecasting model, besides having all coefficients significant, needs a good fit. Most forecasting models haven an R^2 well above .9. However, also due to the low number of observations, one runs the risk of over-fitting the model – fitting some of the random noise inherent in elections. Striking a balance between tightly fitting and over-fitting the model is a challenge that seems to be as much art as science.

These problems decrease only slightly with every additional Bundestag election held. Yet, there are alternatives to waiting four years for the next election to expand one’s dataset by one observation. Several authors estimate forecasting models at the district and state level and aggregate the results up to arrive at a national prediction ([Klarner, 2012](#); [Jerôme and Jérôme-Speziari, 2012](#)). [Munzert \(2015\)](#) presents the first attempt at constituency-level election forecasting for Germany. Such an approach has certain advantages. In the US the winner of the popular vote might be the loser of the Electoral College vote, which is precisely what happened to Al Gore in 2000. Furthermore, not just national but

also regional factors determine national election outcomes and combining district or state time-series yields a much larger dataset allowing the estimation of more complex models. These reasons in part also apply to the German context as voters vote for regional not national party lists and seat shares are allocated at the regional level. Disaggregating to a lower level also allows forecasters to account for regional peculiarities, for instance the CSU in Bavaria or the strong performance of Die Linke in Eastern Germany (or greater number of protest voters in Eastern Germany).

6.1 Election Forecasting: what is it good for?

Election forecast are not without criticism. Particularly structural models like ours which are macro-level models of what in the end is an individual-level act – voting – are the object of criticism. They are estimated from small samples and rely on the assumption of time-invariant effects. Therefore, they seem destined to go wrong or require ad-hoc fixes, that may seem arbitrary, if circumstances change..

Such has been the topic of a debate between [Klein \(2005\)](#) and [Gschwend and Norpoth \(2005\)](#) in a 2005 post-election issue of *Politische Vierteljahresschrift*. Klein criticizes Norpoth and Gschwend for a one-time correction to their model, adjusting one variable to account for the emergence of a new electoral coalition between the *WASG* and *PDS*. They were succesful in doing so, missing the actual result by just 0.3%-points. While Norpoth and Gschwend see it as merely ‘a practical adjustment that does not invalidate the logic of the model,’ Klein is more critical. He contends that Norpoth and Gschwend’s 2005 forecast can not be considered a success but rather reveals that their model is not capable of producing universally valid forecasts criticizing their ‘ad-hoc’ adjustment on both theoretical and empirical grounds. We see Klein’s contribution as targeted critique of Norpoth and Gschwend’s model adjustment rather than a general attack on forecasting. Without commenting on the debate, we think that something valuable can be learned. It seems that the popularity of the head of government is less predictive of the vote in a more

fractionalized party system. This hypothesis derived from a forecasting model could be tested in a more standard systematic way. In fact, we do argue that as political scientists we can learn something from forecasting models even when they fail, or particularly when they do. Theoretically informed forecasts provide a baseline (a sort of expected normal vote) against which the actual election can be judged, so even if a forecast is off sometimes it will have explanatory value.

Our model also fares least well in predicting the 2005 election – the forecast is off by about positive 2%-points. We should add that the 2005 election was also considered a debacle for polling as polling companies missed the result by a wide margin, particularly polls overestimated the CDU vote share by more than 5%-points ([Schaffer and Schneider, 2005](#); [Plischke and Rattinger, 2009](#)). One should thus ask what could have made 2005 such an apparently special unpredictable election. Structural forecasting models like ours provide a heuristic to address such questions. We discuss two aspects that are particular about the 2005 election. Firstly, it was an early election called by then chancellor Gerhard Schröder after the SPD lost the state elections in Germany’s most populous state, North Rhine-Westphalia, earlier in the year. Secondly, the 2005 election saw for the first time an electoral alliance of the *PDS* and the newly founded *WASG* formed just months before the election.

We rule out the early election as cause for the relatively large prediction error in our model, although it could pose problems to our measure of term length. The election of 2005 has not been the first early election in Germany. In our time-series the elections of 1983 and 1990 have both been early elections, both taking place roughly three rather than the usual four years after the last election. 2005 is thus not a one time deviation. Also, if at all, the early election should have led us to underpredict the vote share. We use the log count of terms to measure declining support for the government over time, which might actually be a continuous process better measured in say months. If this is the case our model would have predicted too high a term-length penalty because the government did not use its full term in office.

We think the emergence of the electoral alliance of *PDS* and *WASG* is the most likely culprit. [Gschwend and Norpoth \(2005\)](#) adjust their model to account for the new electoral alliance by subtracting the share of vote intentions for the new electoral alliance from the share of people favoring the chancellor. They argue that this correction is necessary to avoid overstating the vote intentions for the SPD based on chancellor popularity. They say that all voters intending to vote for the PDS/WASG should favor Schröder over Merkel but they will nevertheless not vote for the SPD. A similar argument might apply to our party identification measure. We think it likely that in 2005, exclusively because of the relatively recent emergence of the WASG/PDS alliance, a party identification for the SPD was less indicative of a vote for the SPD than it would usually be. We therefore overestimated the coalitions vote share (combined vote share of SPD and Greens) for our 2005 prediction.

A last comment concerns the use of economic data. We use benchmarked growth, as explained above, Jérôme and co-authors use unemployment figures. The use of economic measures in election forecasting models is common practice. It is based on economic voting research that found that voters' evaluations of recent economic developments influence their voting decisions ([Healy and Malhotra, 2013](#)). We do know that the state of the economy matters for voting, however we know less about how voters learn about the economy. We use benchmarked growth because we believe that voters learn about the economy primarily through the media and that the media report more positively on the economy when it is outperforming that of comparison countries as argued by [Kayser and Peress \(2012\)](#). However, how voters learn about the economy is almost never explicitly addressed in the election forecasting literature. As [Kayser and Leininger \(2015\)](#) point out, when fitting models, most forecasters unwittingly assume that the actual state of the economy, a state best estimated by the multiple revisions to official macroeconomic statistics that occur after their initial release, drives voter behavior.

7 Conclusion

The field of election forecasting is changing rapidly. Older “structural” or “political economy” models of forecasting, such as our benchmarking model, are losing ground to poll-averaging techniques and synthetic models that combine polls and fundamentals (Linzer, 2014). Although polls are snapshots in time rather than true forecasts, the public often treat them as prognostications (Hillygus, 2011). In the run-up to the 2012 US presidential election, poll averaging proved impressively accurate with at least four prominent poll aggregators¹⁵ correctly predicting the winner in all or nearly all 50 states accurately, sometimes months in advance of the election (Blumenthal, 2014; Jackman, 2014; Traugott, 2014). Synthetic models that combine elements such as structural forecasts, polls and sometimes even expert predictions have also proliferated and excelled both for US (Linzer, 2013; Erikson and Wlezien, 2014) and European (Graefe, Forthcoming 2015b; Lewis-Beck and Dassonneville, 2015) forecasts. Such developments rightfully beg the question of whether structural forecasts with their obvious shortcomings – limited degrees of freedom and macro-level modeling of micro-level decisions, and a spotty forecasting record (Campbell, 2012; Lauderdale and Linzer, 2014) – serve any role beyond contributing inputs into synthetic models. For example, Nate Silver (2012, ch 2, pp.67-68.), not a stickler for theory, dismisses “fundamentals-based” models as poor predictors compared to synthetic models that combine polls and economic data. For one purpose, however, structural models remain uncontested: testing explicit models of voter behavior. Polls, aggregated or not, do not even attempt to understand why voters behave as they do; they only measure sentiment.

Predictive validity is a hallmark of science and one of the most demanding tests of a model (Friedman, 1953; Rosenberg, 1992). We embraced the challenge of predicting the 2013 German Bundestag election as an opportunity to test the hypothesis that voters benchmark domestic economic performance against that abroad, which influences their

¹⁵Simon Jackman, Drew Linzer, Nate Silver and Sam Wang

support for the government. More specifically, as [Kayser and Peress \(2012\)](#) argue and demonstrate, they unwittingly respond to “pre-benchmarking” in which the media reports more positively on the economy when it outperforms that of neighboring states.

Nearly two months ahead of the election we offered a theory-driven forecasting model predicting the outcome of the upcoming Bundestag election. Our model drew on previous election outcomes, characteristics of the government and of voters and, most originally, the relative economic performance of Germany in comparison to the three other most important economies in Europe, France, the UK and Italy. We predicted that the outgoing coalition of CDU/CSU and FDP would receive a vote share of 47.05% on September 22, 2013. They, in fact, received 46.3% making our forecast one of the most accurate forecasts offered, besting even the final polls released two days before the election. No less importantly, our forecast also proved considerably more accurate in out-of-sample predictions of German Bundestag elections than an otherwise identical forecasting model using non-benchmarked growth.

The great advantage of structural forecasting models is that they explicitly subject theories of voter behavior to one of the most difficult tests in science, prediction. The hypothesis that the economic assessments that motivate vote choice are comparative, i.e., benchmarked against economic performance abroad, has, at least for the 2013 and prior elections, passed this test.

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