

# AN INTRODUCTION TO MISSING DATA ANALYSES FOR EDUCATION RESEARCH

Craig Enders, UCLA

Brian Keller, University of Missouri

Remus Mitchell, UCLA



# COURSE MATERIAL DOWNLOAD

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WWW.APPLIEDMISSINGDATA.COM/BLIMP-PAPERS

## APPLIED MISSING DATA

[home](#)[analysis examples](#)[blimp](#)[blimp papers](#)[videos](#)[centerstat workshop](#)[quantitude podcast](#)

### Workshops and Training

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Enders, C., Keller, B., & Mitchell, R. (2025, April). *An introduction to missing data analyses for educational research*. Professional development workshop presented at the annual meeting of the American Educational Research Association. Denver, CO.


[DOWNLOAD WEBINAR MATERIALS](#)

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# COURSE MATERIALS

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## ✓ AERA Analysis Examples


 Blimp Studio Analysis 1 Script.imp


 Blimp Studio Analysis 2 Script.imp

 Blimp Studio Analysis 3 Script.imp


 Blimp Studio Analysis 4 Script.imp

 rBlimp Analysis Scripts.R


 reading.dat

 AERA Missing Data Workshop.pdf

## ✓ IES Missing Data Toolkit

>  Dealing With Missing Data Analysis Scripts

 Dealing With Missing Data in Educational Research - Software Tutorials.pdf

 Dealing With Missing Data in Educational Research.pdf



## DEALING WITH MISSING DATA IN EDUCATIONAL RESEARCH

METHODOLOGICAL INNOVATIONS AND  
CONTEMPORARY RECOMMENDATIONS

CRAIG K. ENDERS, PHD

## DEALING WITH MISSING DATA IN EDUCATIONAL RESEARCH

SOFTWARE TUTORIALS

CRAIG K. ENDERS  
REMUS MITCHELL  
MICHAEL P. WOLLER



# BLIMP



[WWW.APPLIEDMISSINGDATA.COM/VIDEOS](http://WWW.APPLIEDMISSINGDATA.COM/VIDEOS)

## BLIMP VIDEO SERIES

The Blimp video series and corresponding YouTube channel provide researchers with training for using the Blimp software. Each video provides a short, step-by-step tutorial that walks viewers through a particular aspect of a missing data analysis. Check back for updates, as new videos are continually added.

# INSTALLING BLIMP

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[WWW.APPLIEDMISSINGDATA.COM/BLIMP](http://WWW.APPLIEDMISSINGDATA.COM/BLIMP)

## BLIMP 3.0

Blimp 3 offers powerful latent variable modeling and imputation for incomplete data sets with up to three levels. Blimp's unique Bayesian computational architecture allows easy specification of complex analyses that are difficult or impossible to fit in other software packages.

[Download Now](#)

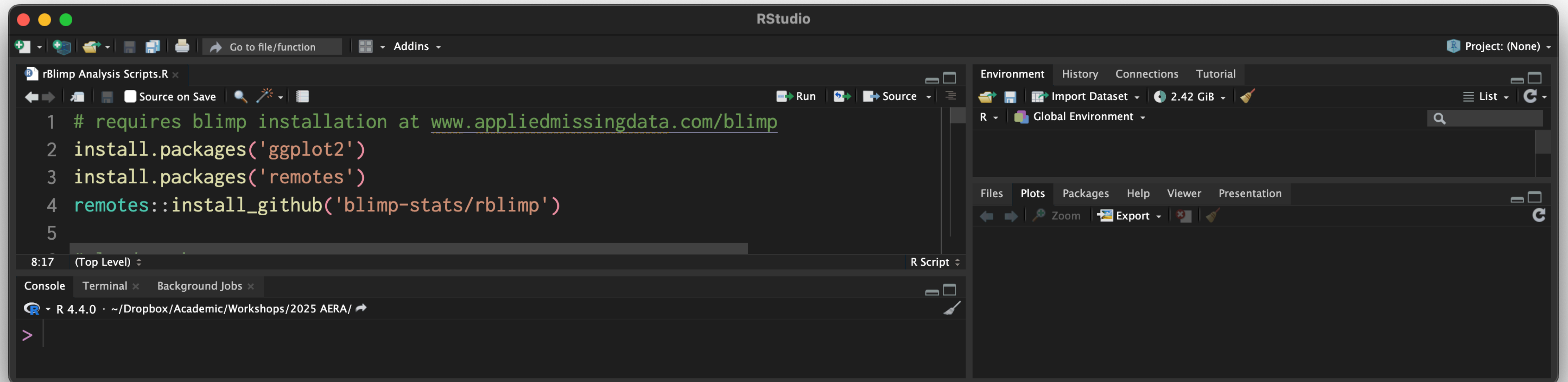
[User's Guide](#)





# INSTALLING RBLIMP (OPTIONAL)

- Package installation lines are at the top of the rBlimp analysis scripts file in the course materials



[WWW.APPLIEDMISSINGDATA.COM/BLIMP](http://WWW.APPLIEDMISSINGDATA.COM/BLIMP)

## BLIMP 3



## User's Guide

BRIAN T. KELLER & CRAIG K. ENDERS

## USER'S GUIDE

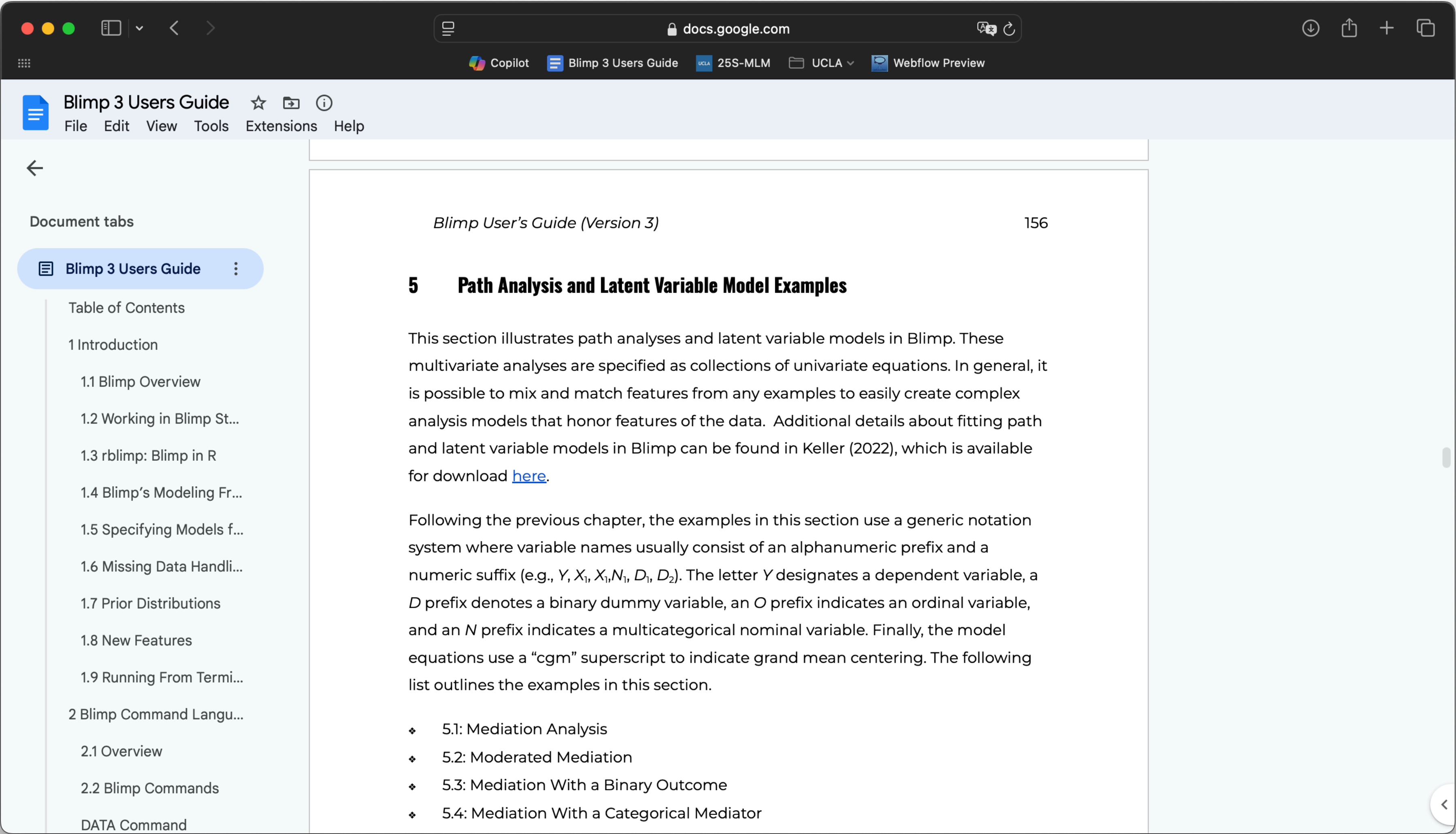
User Guide

Discussion Board  
(Coming Soon)

The User Guide provides an accessible overview of Blimp's simple scripting language, including dozens of new analysis examples covering regression models, path and latent variable models, psychometric models, multilevel models, and missing not at random models. All user examples are accessible and executable from Blimp Studio pull-down menus.



# BLIMP USER GUIDE



# OUTLINE

1

Modern Missing Data Methods

2

Missing Data Mechanisms

3

MCMC Estimation and Missing Data Imputation

4

Fitting Regression Models in Blimp

5

Incomplete Categorical Variables

6

Interaction Effects

7

Reporting Results From a Missing Data Analysis



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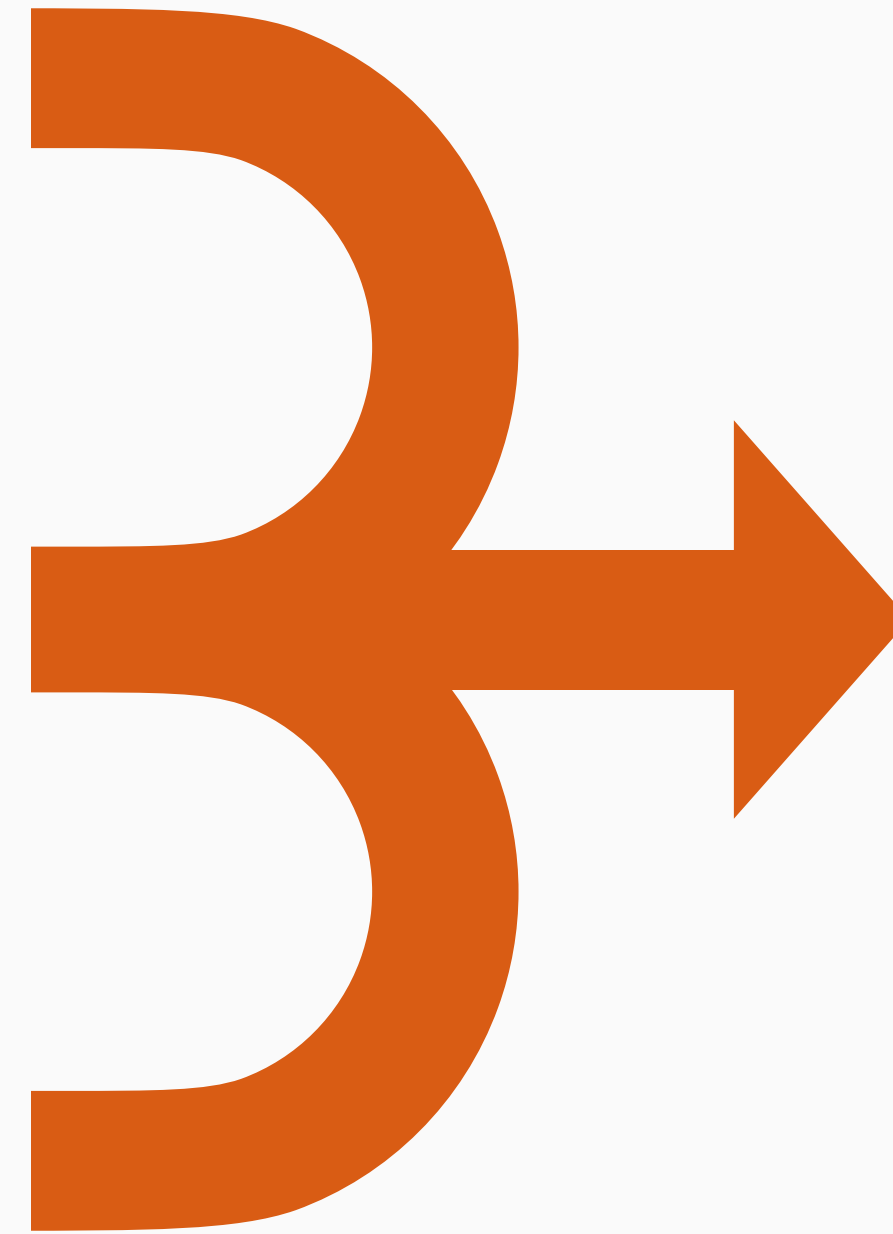
# MODERN MISSING DATA METHODS

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Maximum likelihood

Bayesian MCMC estimation

Multiple imputation



**the  
Big  
Three**



# KEY ADVANTAGES OF BIG THREE

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- Achieve unbiasedness with a more realistic assumption about the missing data process
- Allow for alternate assumptions about nonresponse process
- Maximize power
- Use all available data, no wasted resources

# CHOOSING A MISSING DATA METHOD

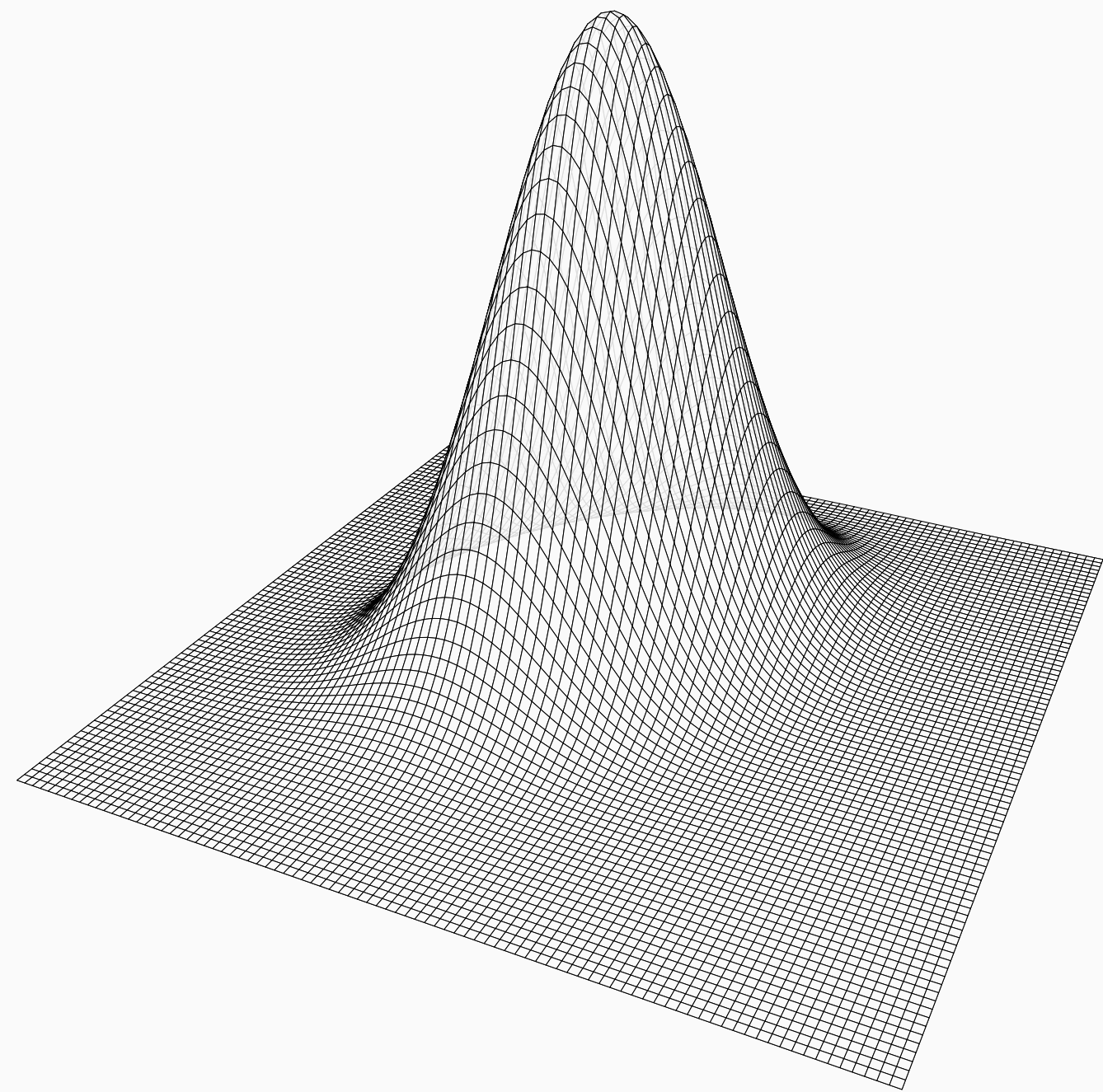
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- All things being equal—same data, same variables, same assumptions—the Big Three rarely produce different results
- Missing data analyses require distributional assumptions
- How we represent those distributions—multivariate versus factored specifications—is what matters

# MODELING FRAMEWORKS

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## Multivariate modeling



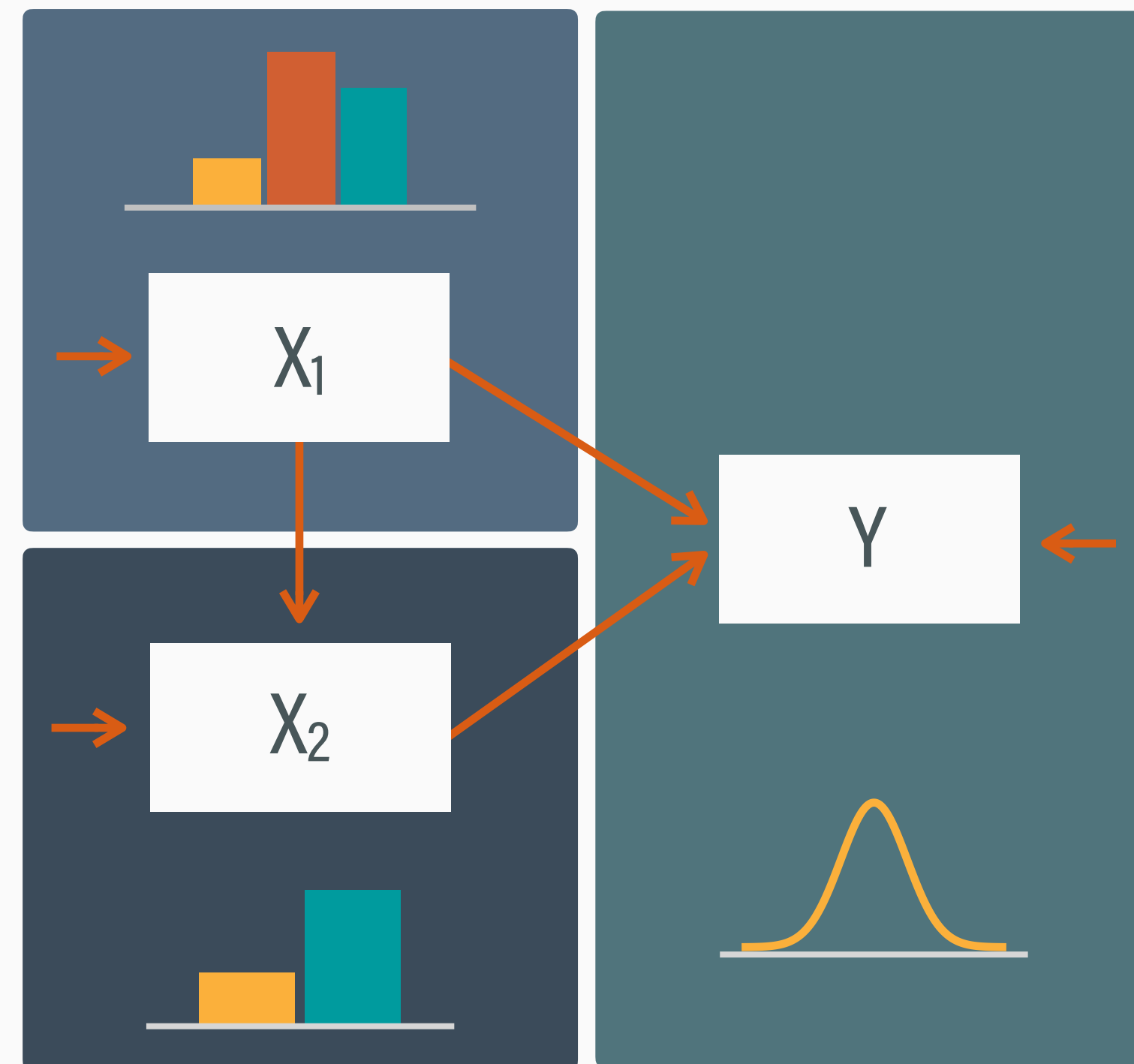
- Classic approaches often assume multivariate normality
- Most applications of maximum likelihood and multiple imputation



# MODELING FRAMEWORKS

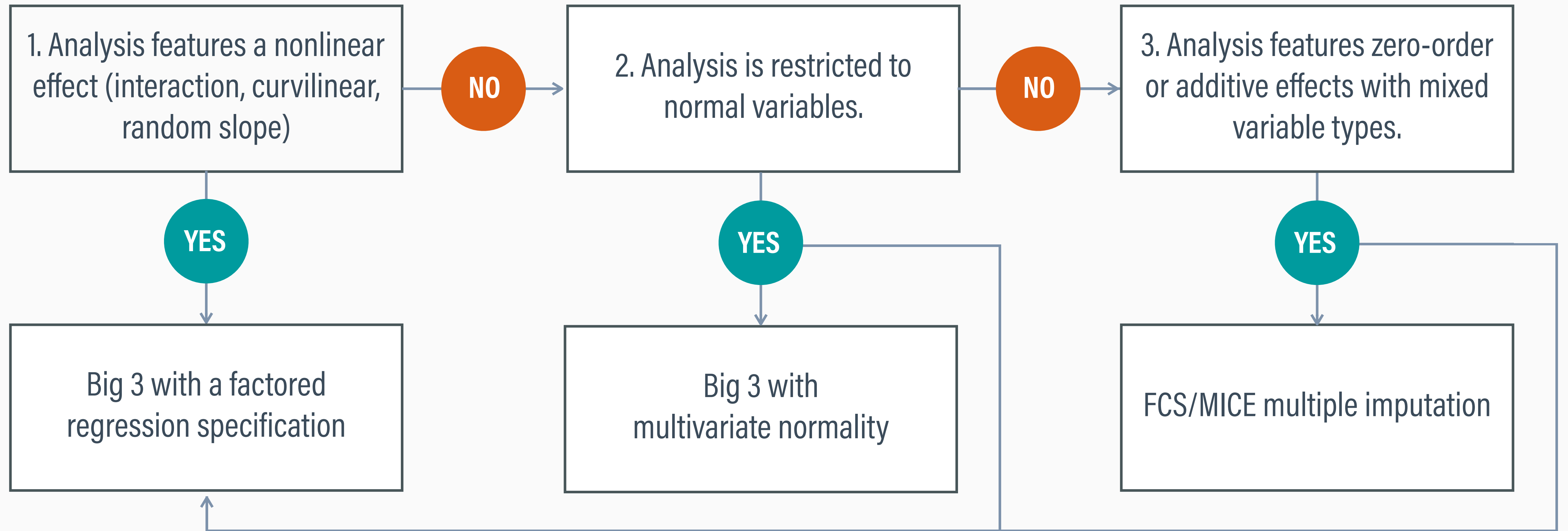
## Multivariate modeling

## Factored regression specification



- Factored regression invokes a unique model and distribution for each variable
- Each model can include terms that are at odds with multivariate normality (e.g., categorical variables, interactions, random slopes)

# MISSING DATA DECISION TREE



# WHY CHOOSE MCMC?

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- MCMC readily handles complex missing data problems, including:
  - Mixed metrics (normal, ordinal, nominal, skewed, count, latent)
  - Nonlinear effects (interactions, curvilinear effects)
  - Multilevel data (random coefficients, interactions)
  - Latent variable modeling (interactions)
- FIML estimators with factored specifications are limited



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# HOW MUCH MISSING DATA IS TOO MUCH?

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- The Big Three can tolerate substantial amounts of missing data
- The Big Three are increasingly better than ad hoc methods (e.g., deleting incomplete cases) as missingness increases
- The amount of missing data is less important than why the data are missing (the missingness process or mechanism)

# RUBIN'S MISSING DATA MECHANISMS

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- ⦿ Missing data mechanisms (processes) describe different ways in which the data relate to nonresponse
- ⦿ Missingness may be completely random or systematically related to different parts of the data
- ⦿ Mechanisms function as statistical assumptions



# PARTITIONING THE DATA

Complete			=	Observed			+	Missing			Indicators		
Y <sub>1</sub>	Y <sub>2</sub>	Y <sub>3</sub>		Y <sub>1</sub>	Y <sub>2</sub>	Y <sub>3</sub>		Y <sub>1</sub>	Y <sub>2</sub>	Y <sub>3</sub>	M <sub>1</sub>	M <sub>2</sub>	M <sub>3</sub>
4	4	3		4	4	3					0	0	0
3	3	5		3	NA	5			3		0	1	0
7	1	6		7	1	6					0	0	0
2	1	6		NA	1	6		2			1	0	0
5	9	3	=	5	9	3	+				0	0	0
3	2	2		3	NA	NA			2	2	0	1	1
1	6	7		1	6	7					0	0	0
9	4	9		9	4	9					0	0	0
2	5	6		2	NA	6			5		0	1	0

# MISSING COMPLETELY AT RANDOM

- The probability of missing values is completely unrelated to the data

$$f(M = 1 \mid \text{data}_{\text{obs}}, \text{data}_{\text{mis}}) = f(M = 1)$$

- MCAR is purely random missingness
- We don't care about this process or testing for it (e.g., Little's MCAR test)

## Missingness

### Indicators

M <sub>1</sub>	M <sub>2</sub>	M <sub>3</sub>
0	0	0
0	1	0
0	0	0
1	0	0
0	0	0
0	1	1
0	0	0
0	0	0
0	1	0

## Predictors of nonresponse

### Observed

Y <sub>1</sub>	Y <sub>2</sub>	Y <sub>3</sub>
4	4	3
3	NA	5
7	1	6
NA		6
5	9	3
3	NA	NA
1	6	7
9	4	9
2	NA	6

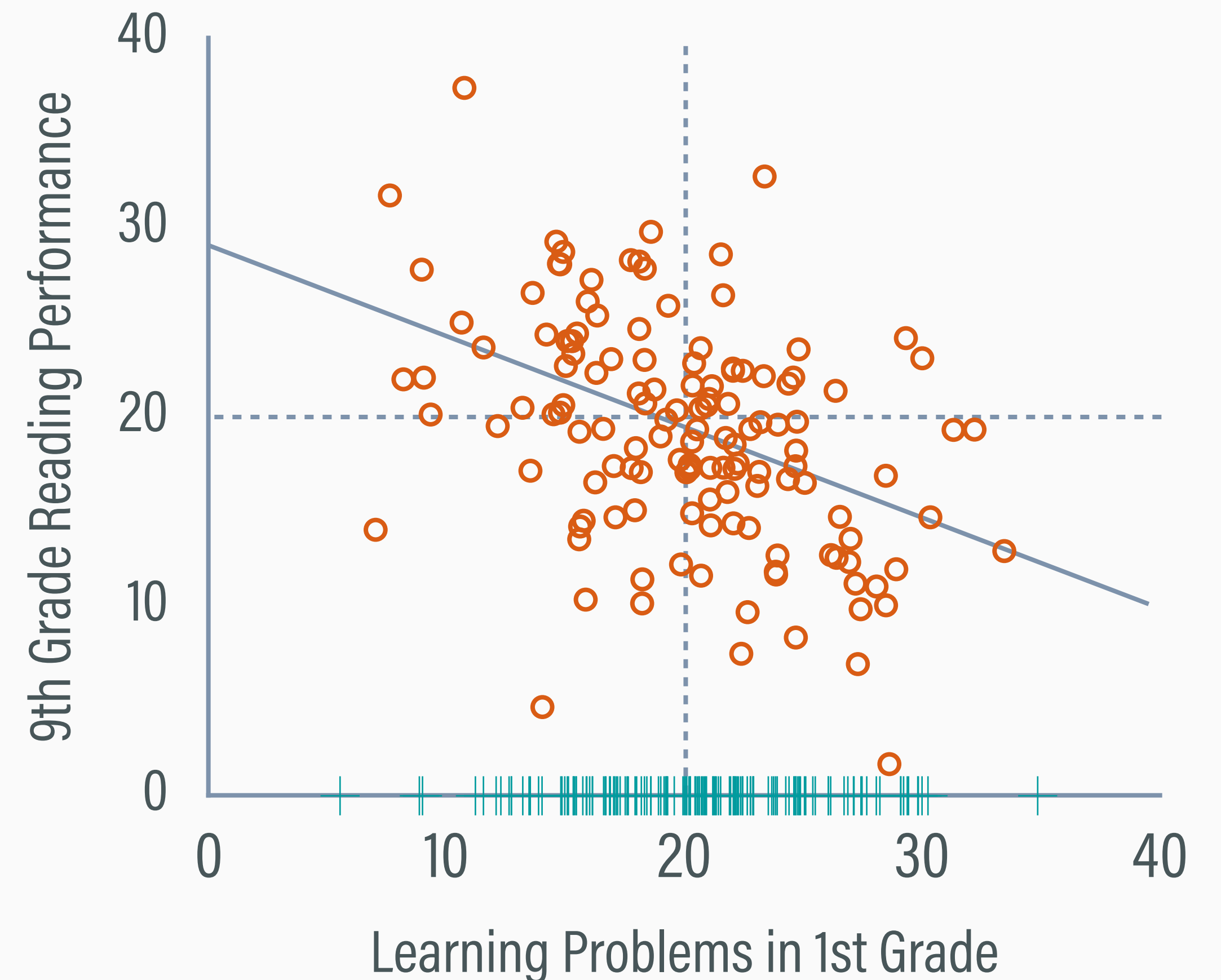
### Missing

Y <sub>1</sub>	Y <sub>2</sub>	Y <sub>3</sub>
	3	
2		
	2	2
	5	

# RESEARCH SCENARIO

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- Study investigating association between learning problems in 1st grade and reading performance in 9th grade
- Learning problems ratings are complete and reading scores are incomplete





# MCAR EXAMPLE

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- Missingness is **unrelated** to the observed learning problems measure and **unrelated** to the unseen reading scores
- Planned missing data design where 9th grade reading scores are collected from a random subset of the original sample in order to reduce data collection costs
- Unplanned missingness is unrelated to the data (e.g., scheduling conflicts, administrative errors, family relocation)

# (CONDITIONALLY) MISSING AT RANDOM

- Systematic missingness related to the observed data but unrelated to the unseen latent data


$$f(M = 1 \mid \text{data}_{\text{obs}}, \text{data}_{\text{mis}}) = f(M = 1 \mid \text{data}_{\text{obs}})$$

- Most Big Three applications assume CMAR

## Missingness

Indicators		
M <sub>1</sub>	M <sub>2</sub>	M <sub>3</sub>
0	0	0
0	1	0
0	0	0
1	0	0
0	0	0
0	1	1
0	0	0
0	0	0
0	1	0

## Predictors of nonresponse

Observed			Missing		
Y <sub>1</sub>	Y <sub>2</sub>	Y <sub>3</sub>	Y <sub>1</sub>	Y <sub>2</sub>	Y <sub>3</sub>
4	4	3			
3	NA	5		3	
7	1	6			
NA	1	6	2		
5	9	3			
3	NA	NA		2	2
1	6	7			
9	4	9			
2	NA	6		5	

# CONDITIONALLY MAR EXAMPLE

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- Missingness is **related** to the observed learning problems measure but **unrelated** to the unseen reading scores
- Students with high levels of learning problems are more likely to have missing data due to increased dropout risk, disciplinary actions, or family or situational instability
- The Big Three assume a CMAR process by default



# MISSING NOT AT RANDOM

- Systematic missingness related to the observed data and the unseen latent data

$$f(M = 1 \mid \text{data}_{\text{obs}}, \text{data}_{\text{mis}})$$

- The Big Three also allow MNAR processes (selection and pattern mixture models)

## Missingness

### Indicators

M <sub>1</sub>	M <sub>2</sub>	M <sub>3</sub>
0	0	0
0	1	0
0	0	0
1	0	0
0	0	0
0	1	1
0	0	0
0	0	0
0	1	0

## Predictors of nonresponse

### Observed      Missing

Y <sub>1</sub>	Y <sub>2</sub>	Y <sub>3</sub>	Y <sub>1</sub>	Y <sub>2</sub>	Y <sub>3</sub>
4	4	3			
3	NA	5		3	
7	1	6			
NA	1	6	2		
5	9	3			
3	NA	NA		2	2
1	6	7			
9	4	9			
2	NA	6		5	

# MNAR EXAMPLE

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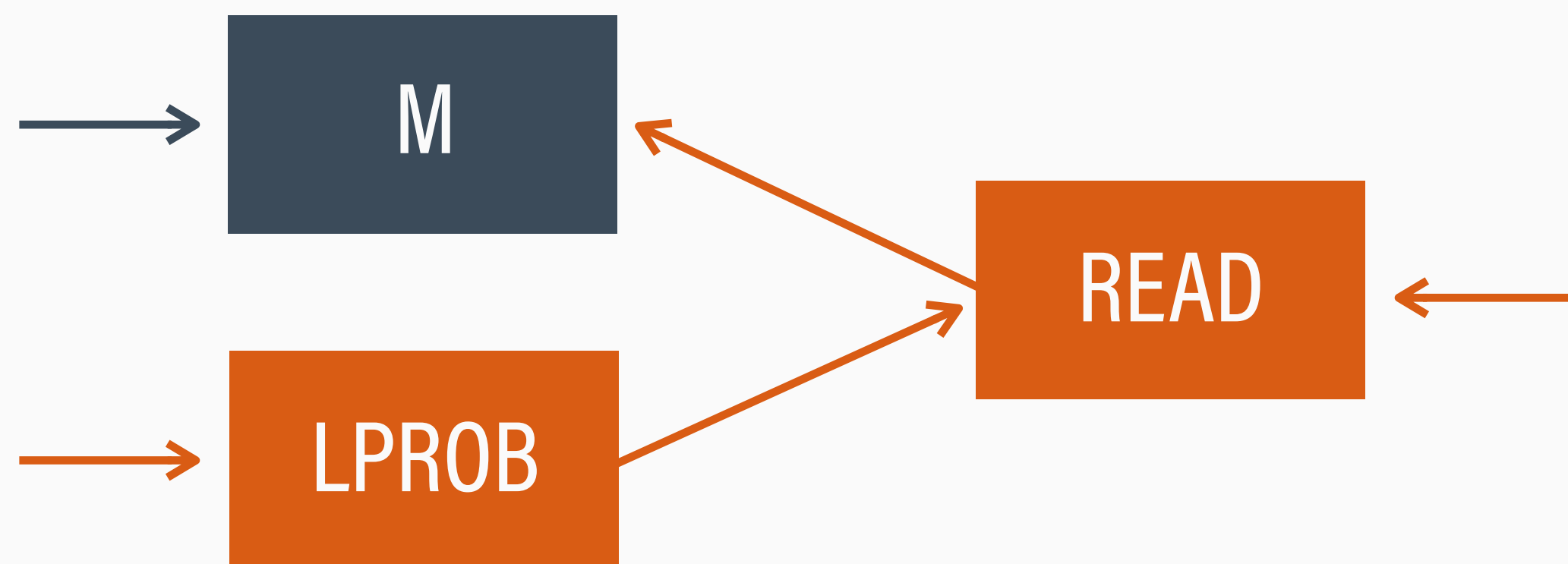
- Missingness is **related** to the observed learning problems measure and also **related** to the unseen reading scores
- Individuals with the low reading levels opt out because they feel discouraged or anxious about testing or because they were moved to specialized programs or alternative educational settings where standardized testing protocols differ

# MNAR MODELING

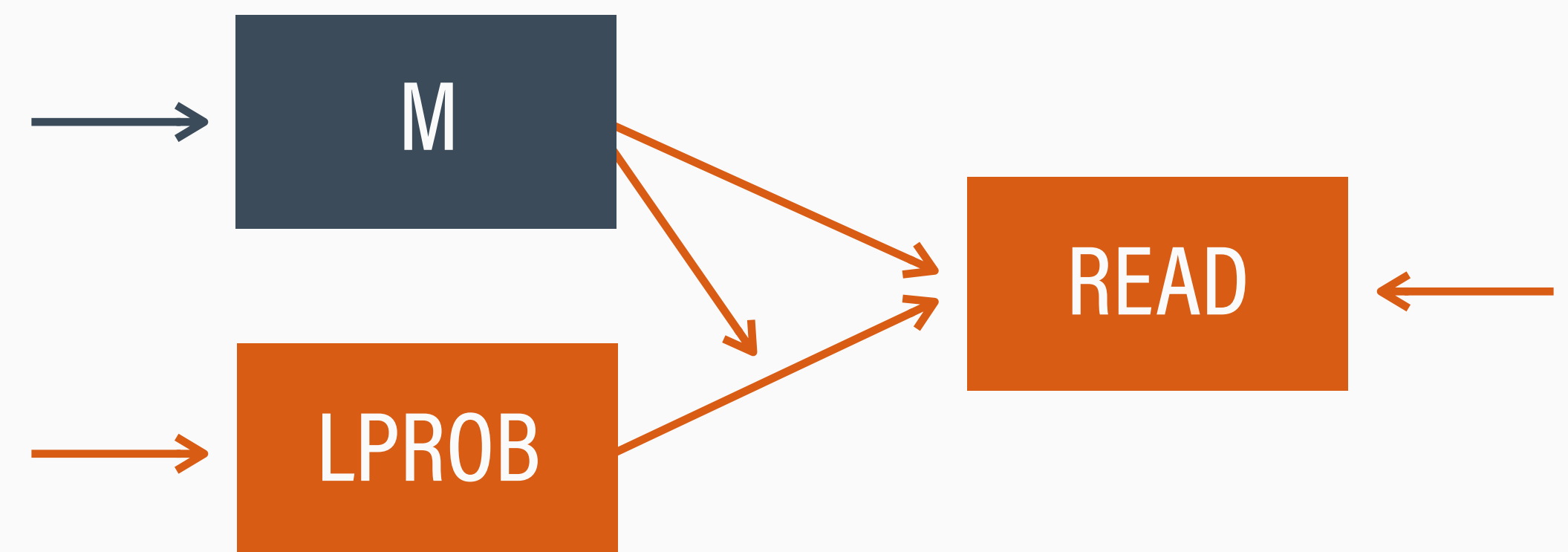
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- Missing not at random processes require an explicit model that incorporates the missing data indicator (M)

Selection Model



Pattern Mixture Model



# TESTING THE CMAR ASSUMPTION

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- The CMAR assumption is untestable because it stipulates no relation between missingness and the unseen scores
- We must rely on logical arguments about why the unseen scores should not be related to missingness
- When in doubt, conduct sensitivity analyses that compare the estimates from CMAR and MNAR assumptions



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# FREQUENTIST VS. BAYESIAN PARADIGMS

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## Frequentist

- The parameter is a fixed quantity, estimates vary across different samples
- Statements about probability, precision, and confidence refer to estimates
- Probability = long run frequency of outcomes across many samples

## Bayesian

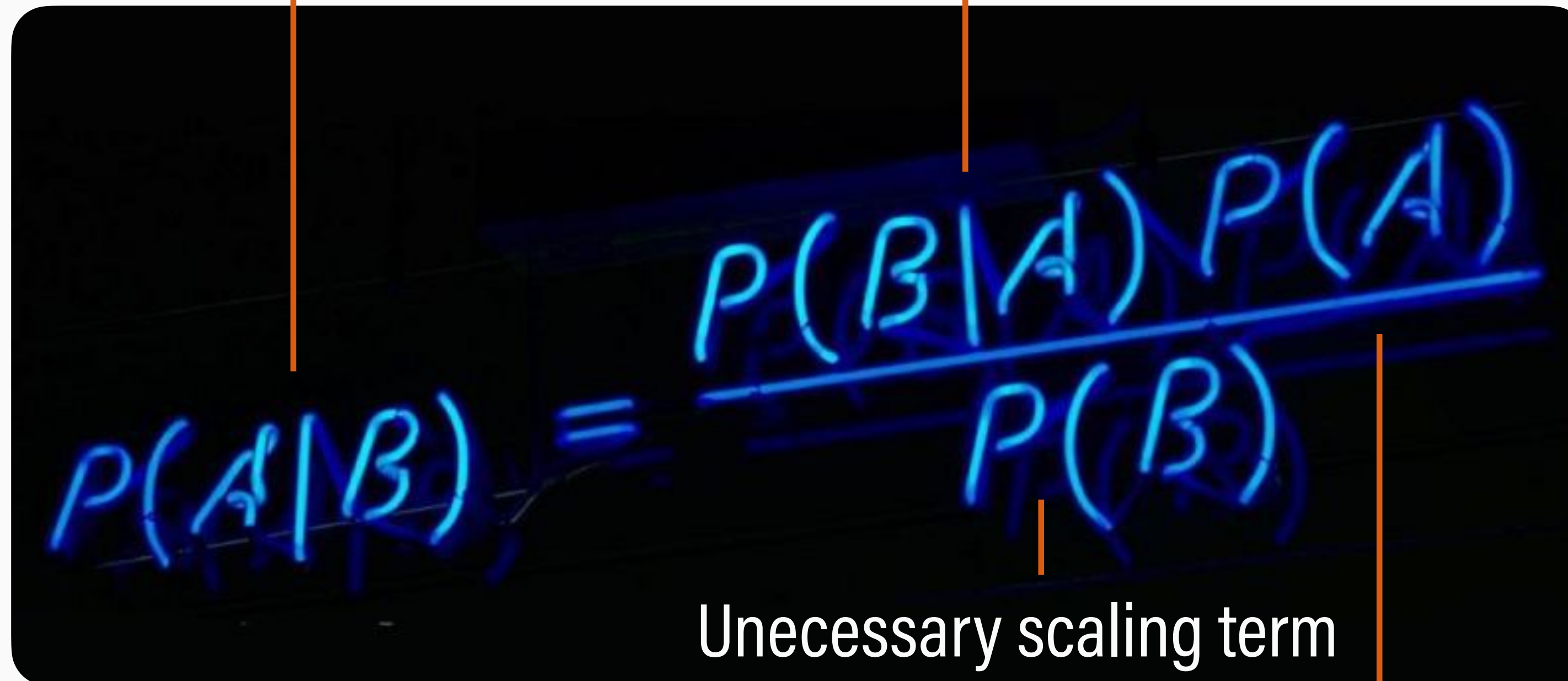
- Parameters are random variables with a distribution of plausible realizations
- Statements about probability, precision, and intervals refer to the parameter
- Probability = our degree of certainty about a parameter after analyzing data

# BAYES' THEOREM

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Posterior = parameters (A) given the data (B)

Frequentist likelihood = data (B) given the parameters (A)



The image shows a handwritten version of Bayes' Theorem on a chalkboard. The formula is  $P(A|B) = \frac{P(B|A)P(A)}{P(B)}$ . Three orange vertical lines point from text labels to parts of the formula: one points to  $P(A|B)$ , another points to  $P(B|A)$ , and a third points to  $P(A)$ . The text 'Unnecessary scaling term' is written below the denominator  $P(B)$ .

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

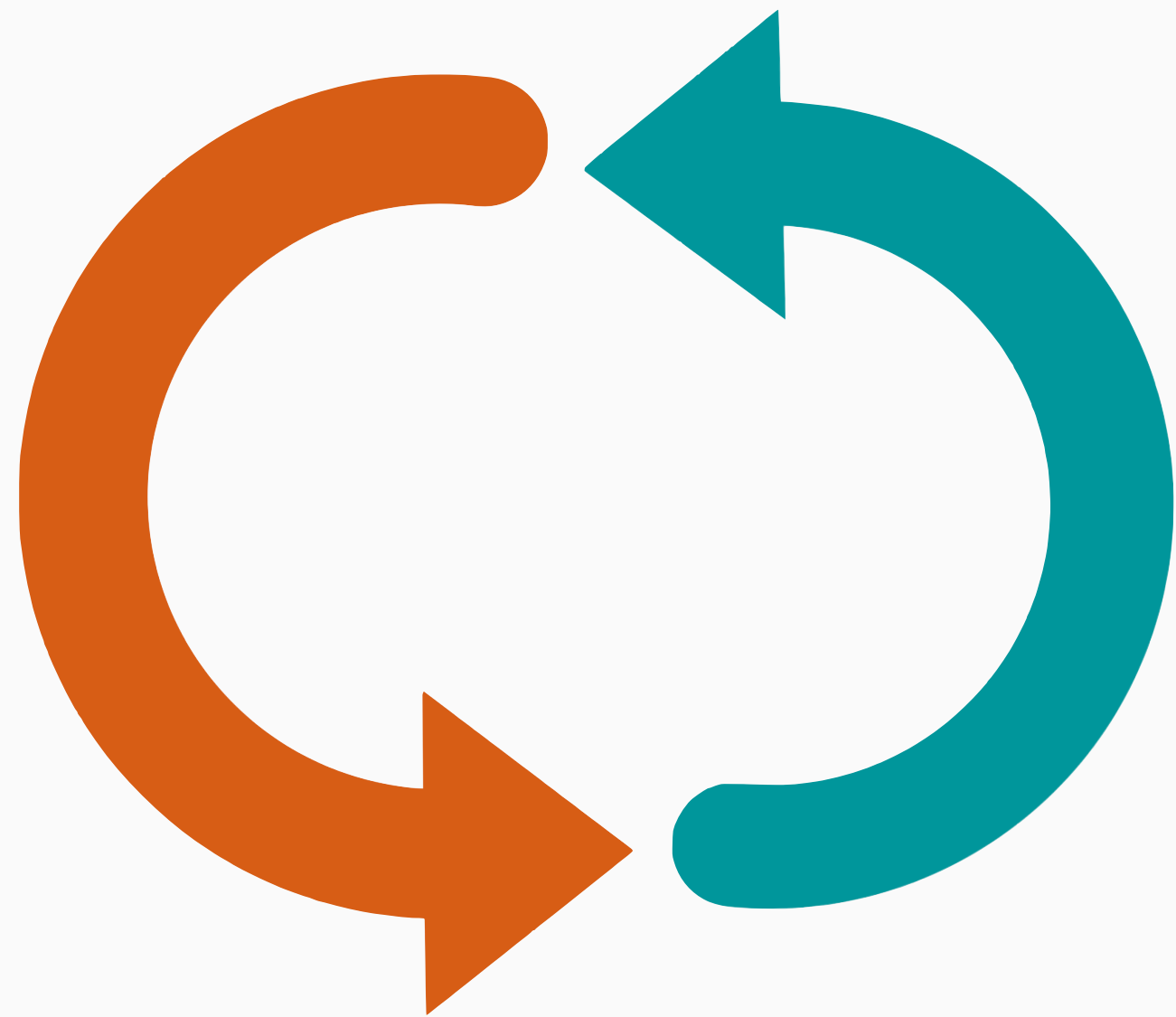
Unnecessary scaling term

Prior = a priori belief about parameters (A)

# MCMC ESTIMATION

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Estimate regression models



Impute missing values

Do for  $t = 1$  to 10,000 iterations

- » Estimate model parameters, conditional on the filled-in data
- » Impute missing values, conditional on the model parameters

Repeat

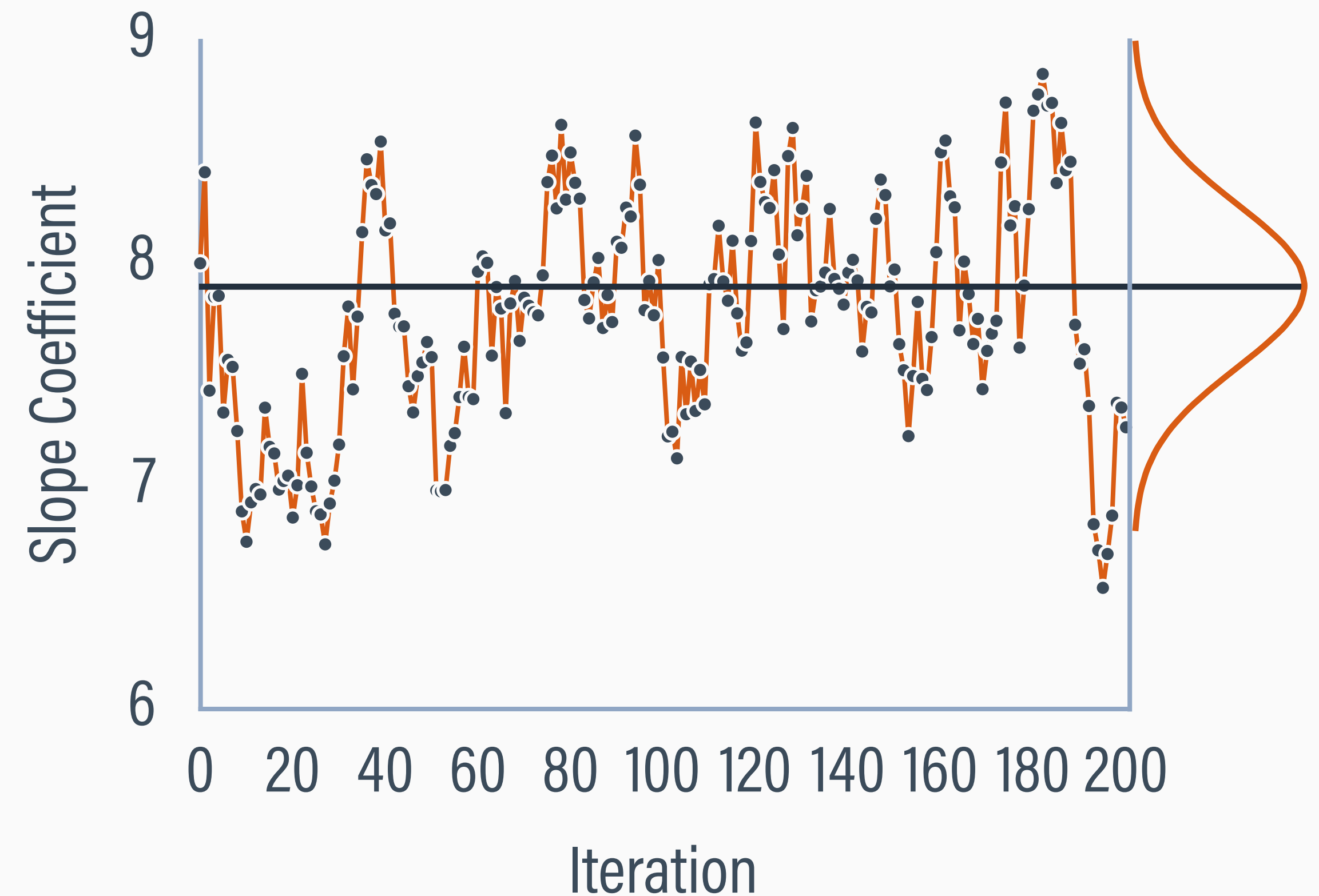
Summarize model parameters



# MEANING OF ESTIMATION

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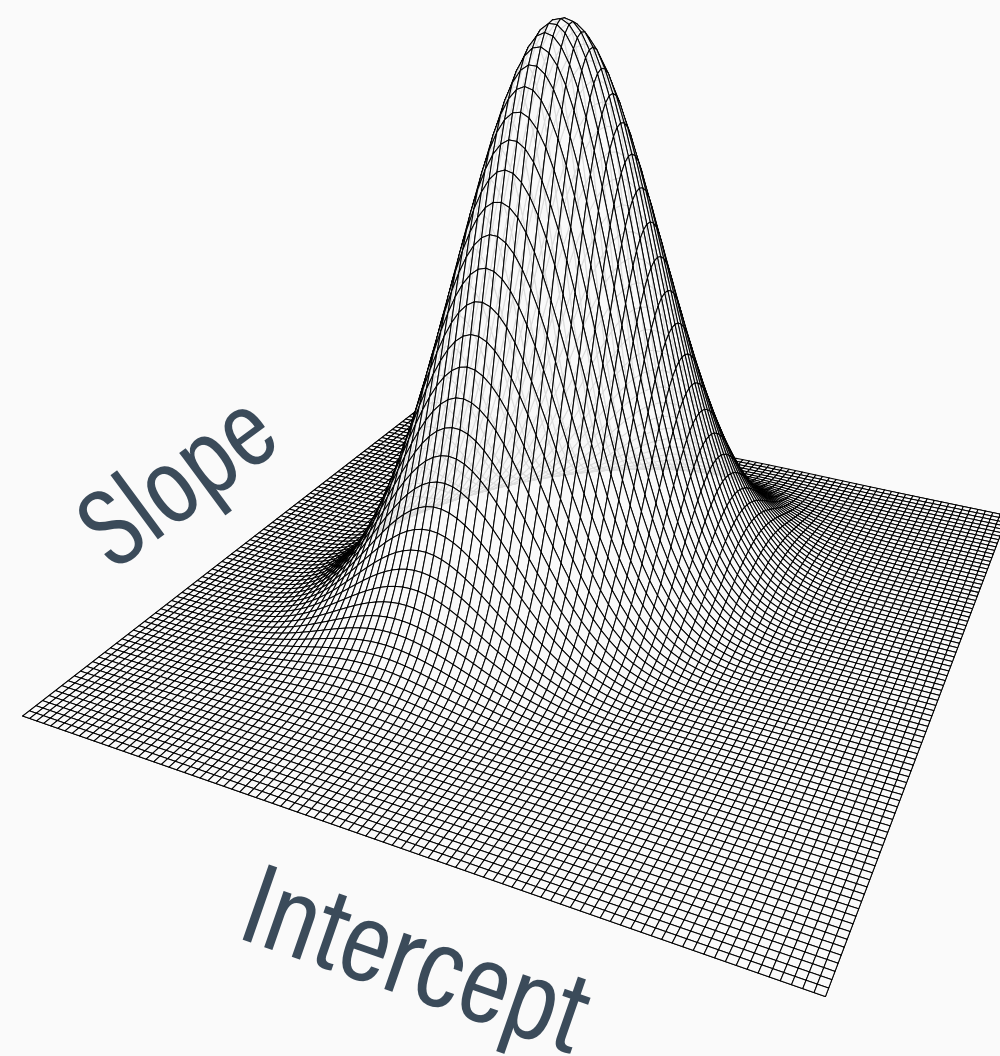
- MCMC uses computer simulation to “sample” parameters from a distribution
- Estimates continually vary across iterations in a random pattern
- Each iteration gives plausible parameter values that could have produced **our data**



# PARAMETER-GENERATING DISTRIBUTIONS

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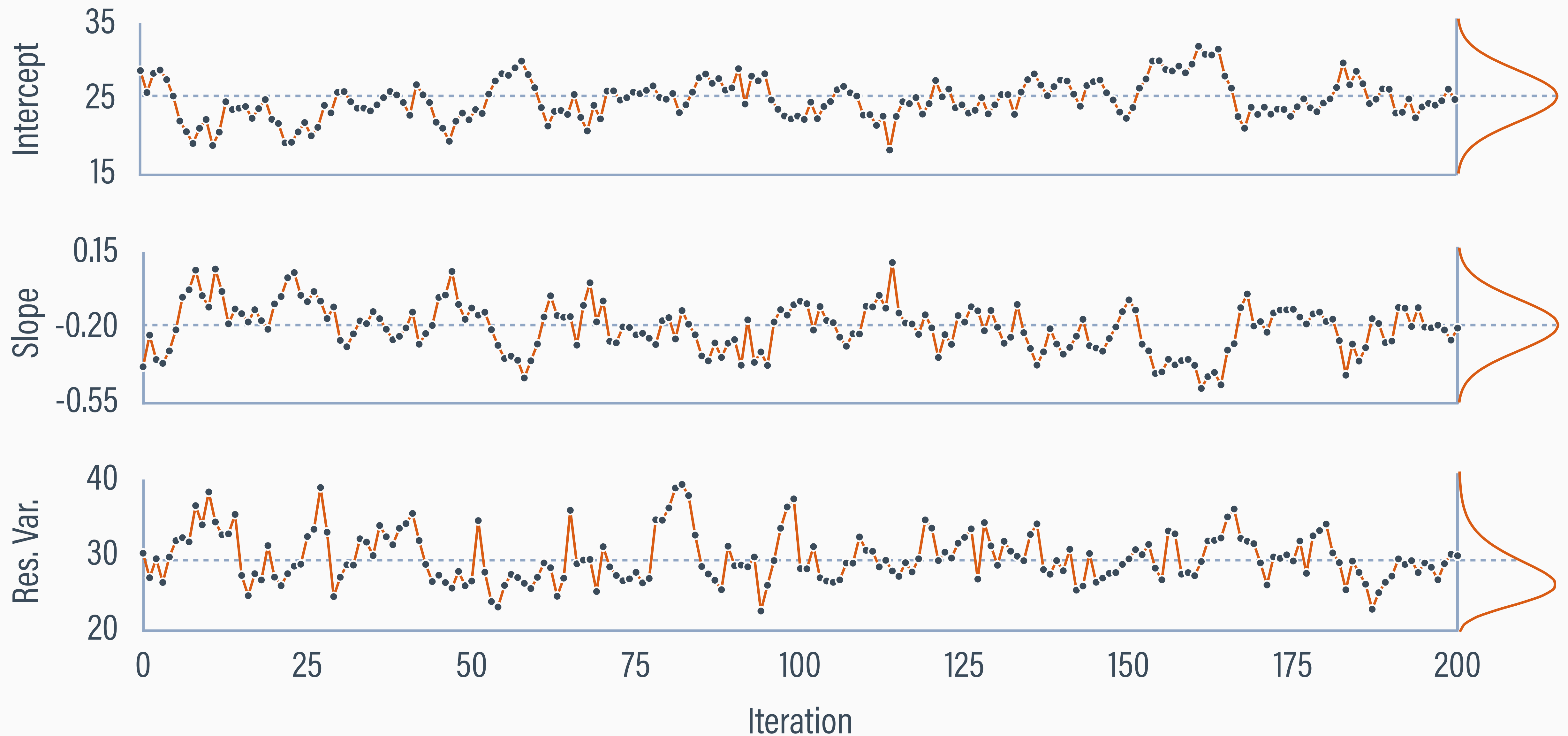
- MCMC draws coefficients from a multivariate normal distribution, with least-squares estimates defining shape



- MCMC draws variances from an inverse gamma distribution with its shape determined by the df and residual SS



# PARAMETERS FROM 200 MCMC CYCLES



# PRIOR DISTRIBUTIONS

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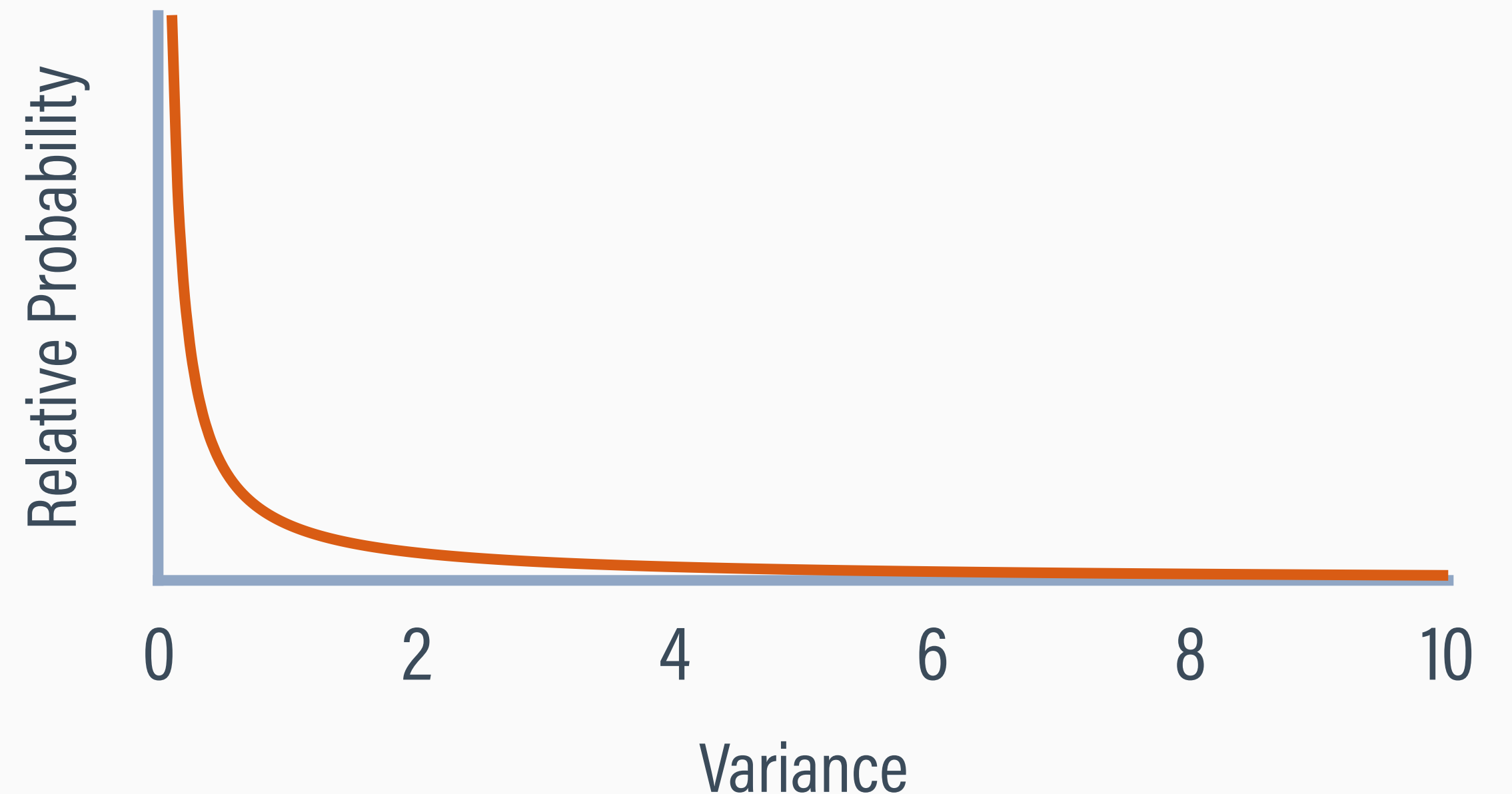
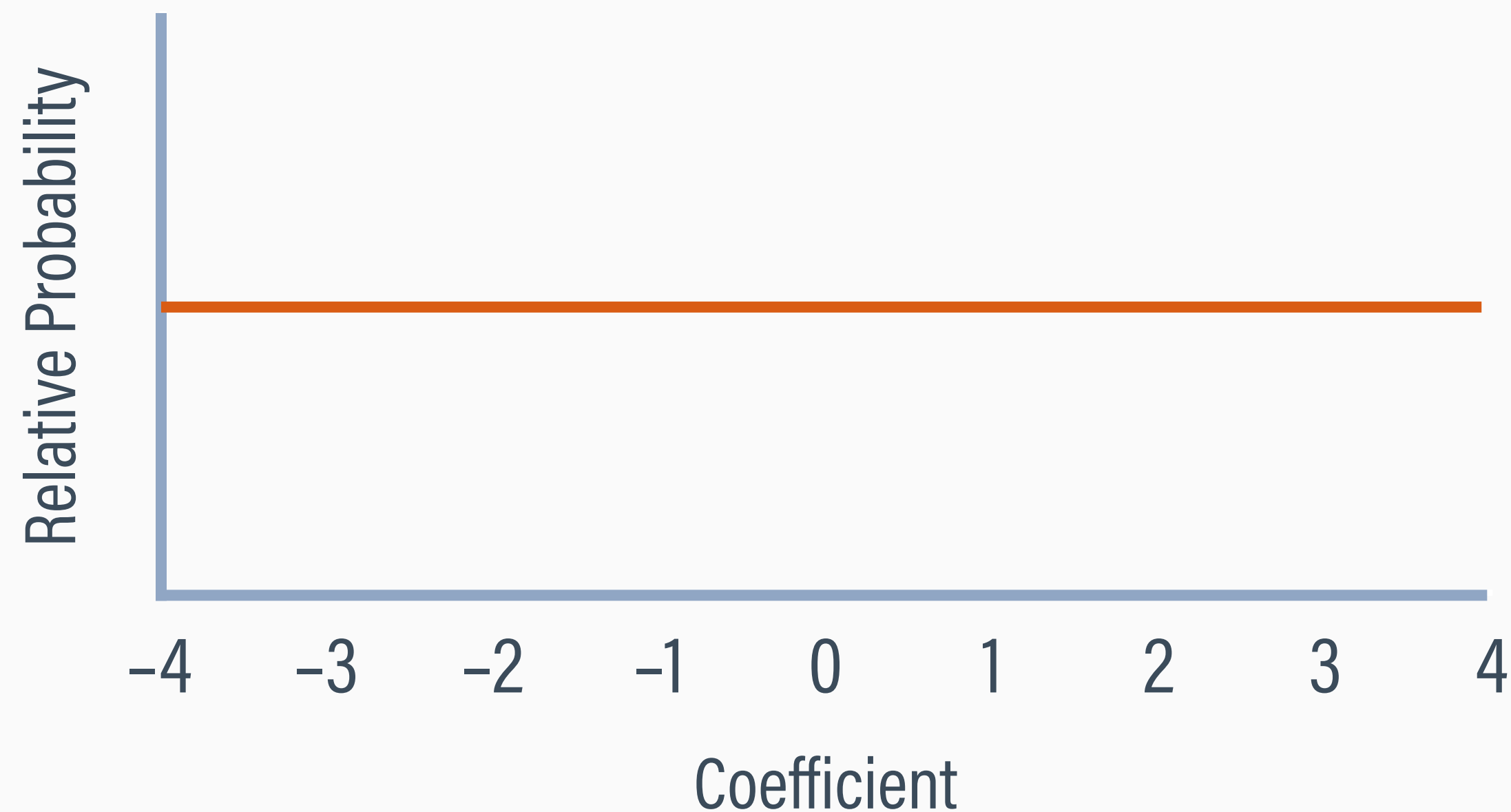
- Bayesian analyses require prior distributions that encode our beliefs about the parameter values prior to analyzing the data
- Blimp adopts non-informative (diffuse) priors that impart as little information as possible (i.e., let the data do the talking)
- Prior distributions for variances can influence estimates when the  $N$  is small, so sensitivity analyses may be warranted (FIML variance estimates are also biased in the same scenarios)



# PRIOR DISTRIBUTIONS

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- A diffuse prior for means and coefficients conveys that all possible parameter values are equally likely a priori
- Diffuse priors for variances are slightly informative, and different options function like df adjustments in regression



# SUMMARIZING MCMC ESTIMATES

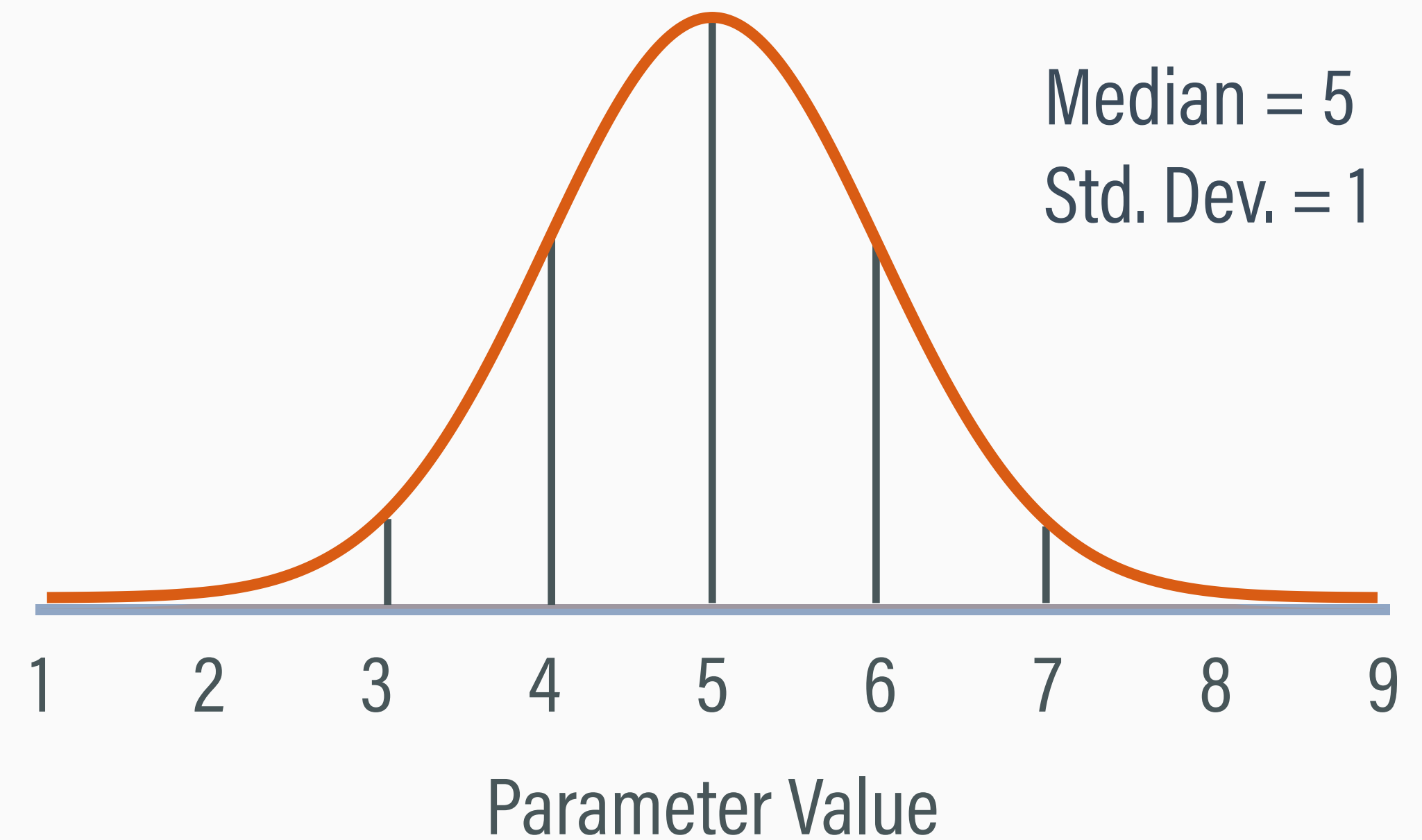
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- MCMC iterates for thousands of cycles, and each cycle produces estimates based on one filed-in data set
- Bayesian estimation yields a distribution of parameters—called a posterior—that averages over thousands of imputations
- The posterior is a distribution of plausible parameter values that could have produced our particular data

# POSTERIOR MEDIAN AND STD. DEV.

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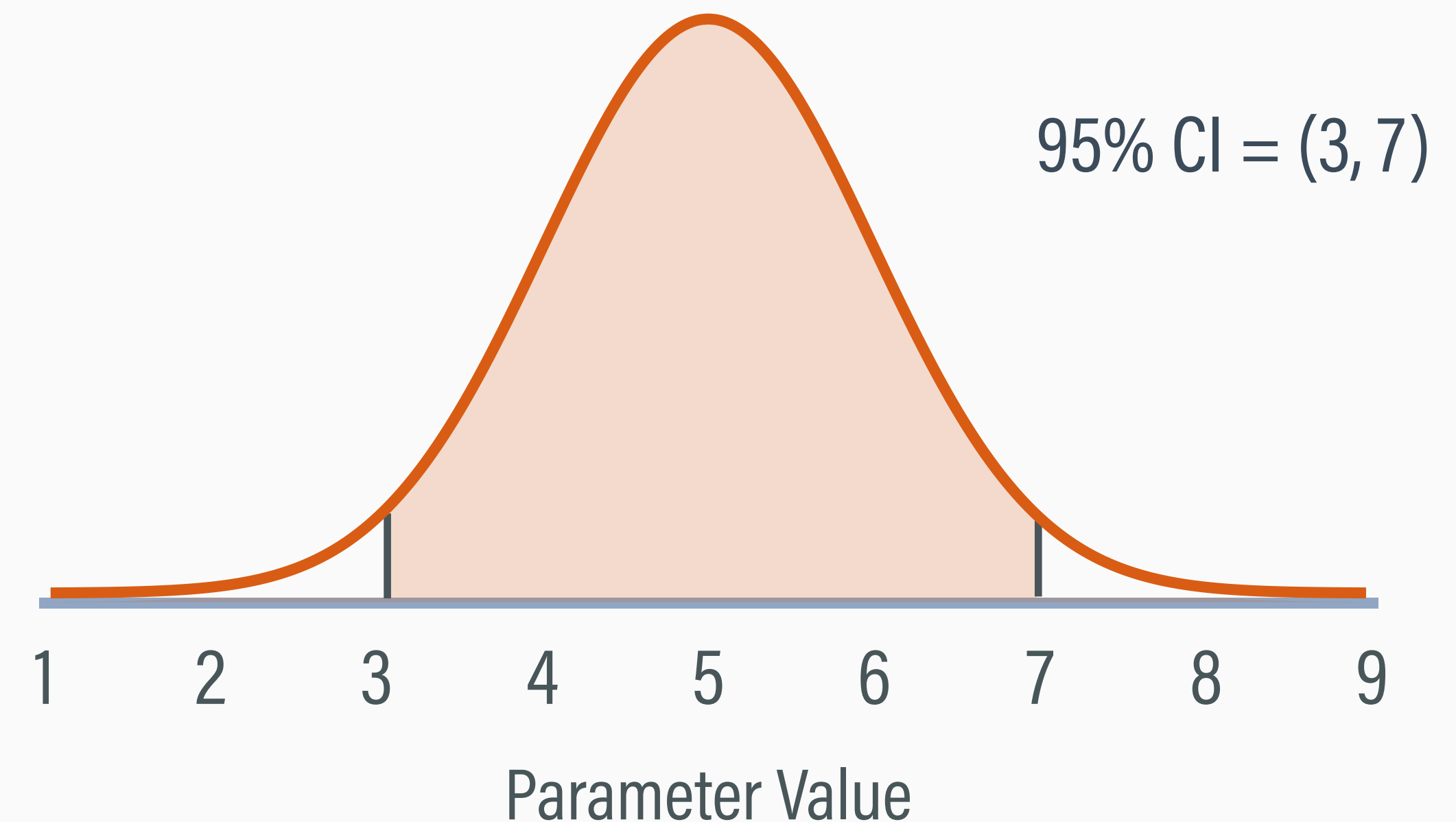
- The posterior median and standard deviation quantify the most likely parameter value and uncertainty
- Analogous to a point estimate and standard error but no repeated sampling



# 95% CREDIBLE INTERVALS

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- The 95% credible interval gives limits spanning 95% of the parameter's range
- Akin to a confidence interval, but references a range of highly plausible parameter values





# SIMPLE REGRESSION ILLUSTRATION

- Study that seeks to determine whether reading levels in 1st grade predict 9th grade reading achievement in middle school

$$\text{read}_9 = \beta_0 + \beta_1(\text{lrnprob}_1) + \varepsilon$$

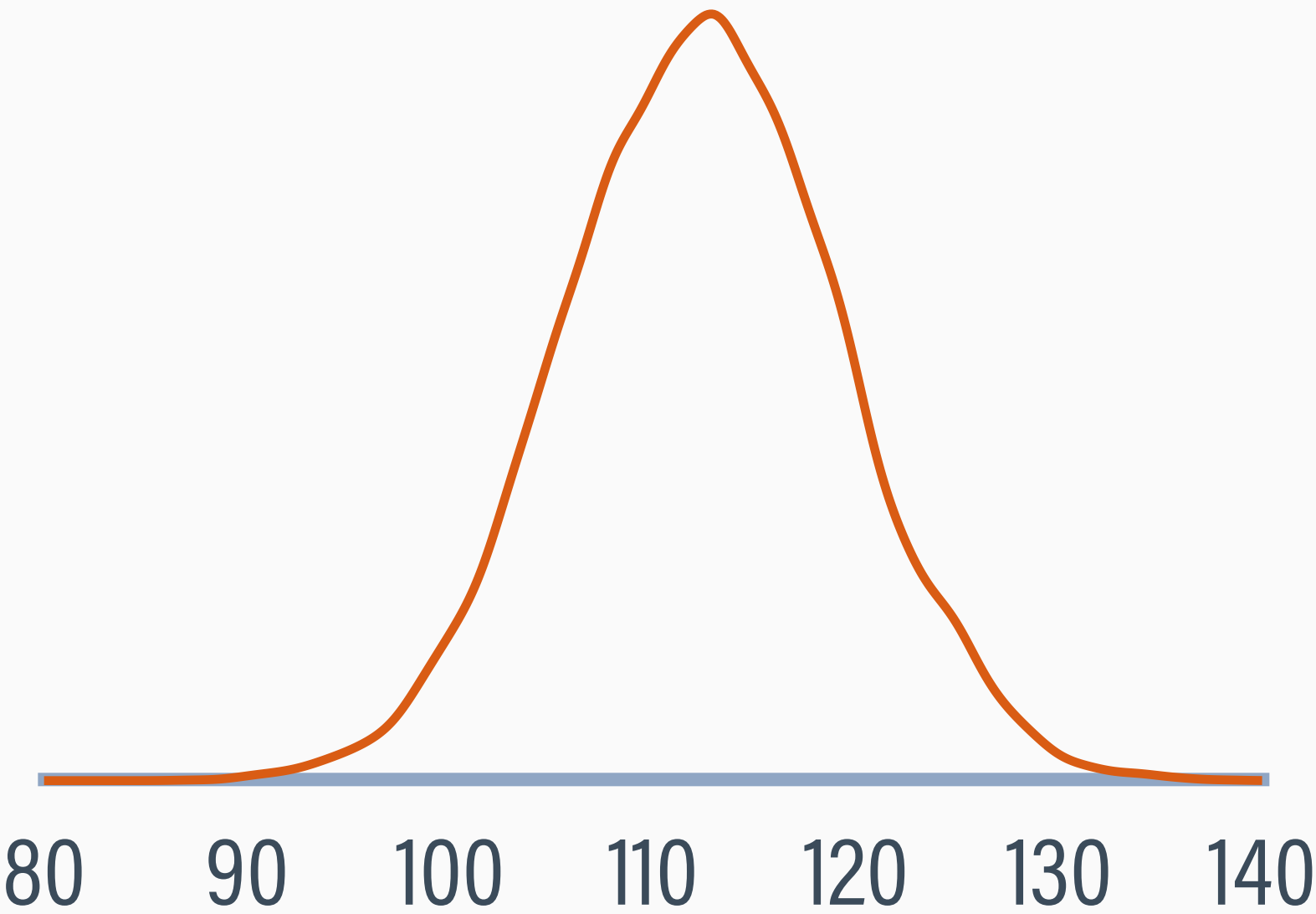
Variable	Definition	Missing %	Scale
atrisk	Emotional/behavioral risk code	2.2	0 = Low, 1 = Medium/high
lrnprob1	1st grade learning problems	2.2	Numeric (31 to 88)
read1	1st grade broad reading composite	6.5	Numeric (39 to 153)
read9	9th grade broad reading composite	17.4	Numeric (41 to 123)

# POSTERIOR DISTRIBUTIONS

Median = 112.48

Std. Dev. = 6.81

95% CI = (99.09, 125.87)

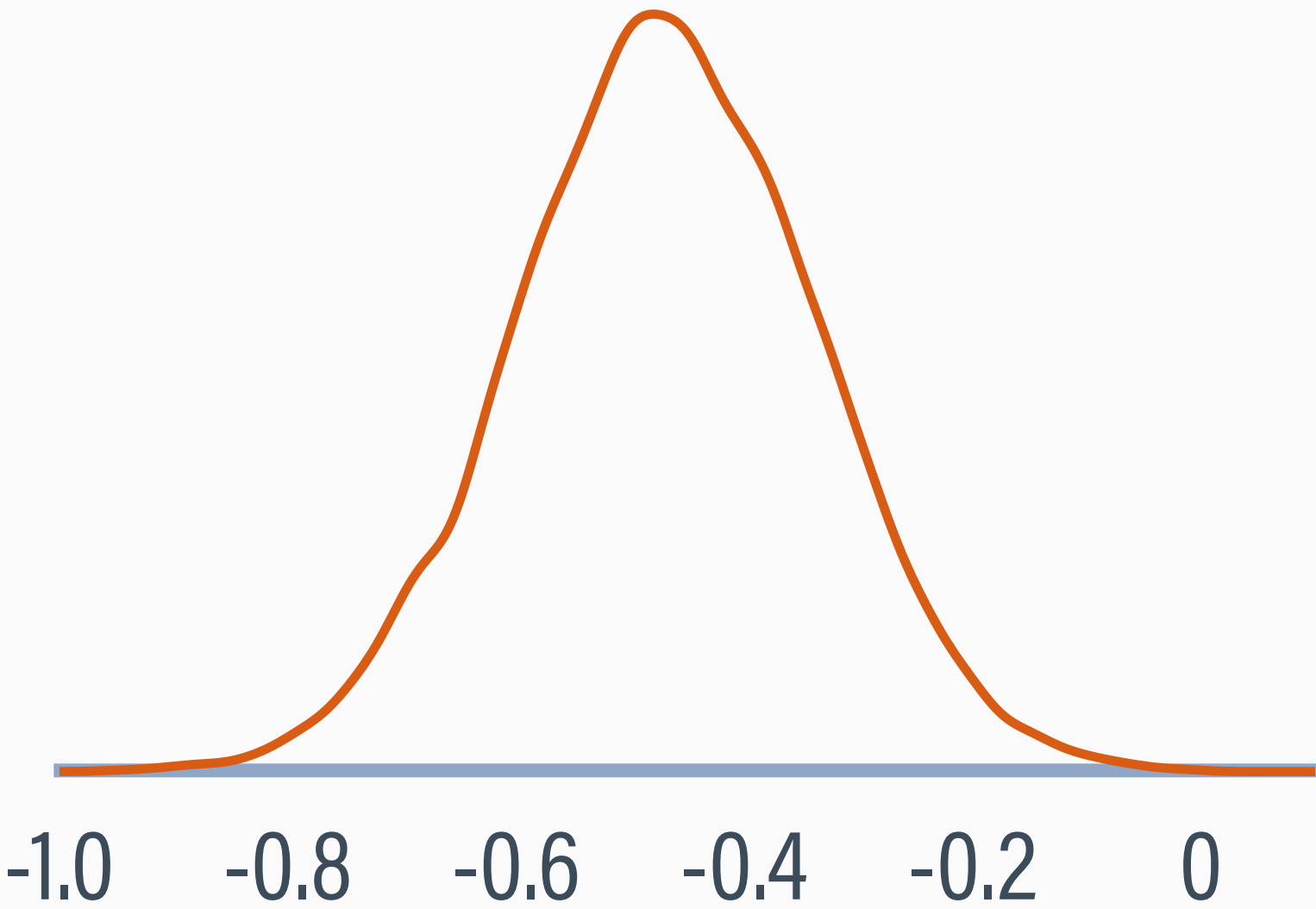


Intercept

Median = -0.47

Std. Dev. = 0.13

95% CI = (-0.73, -0.21)

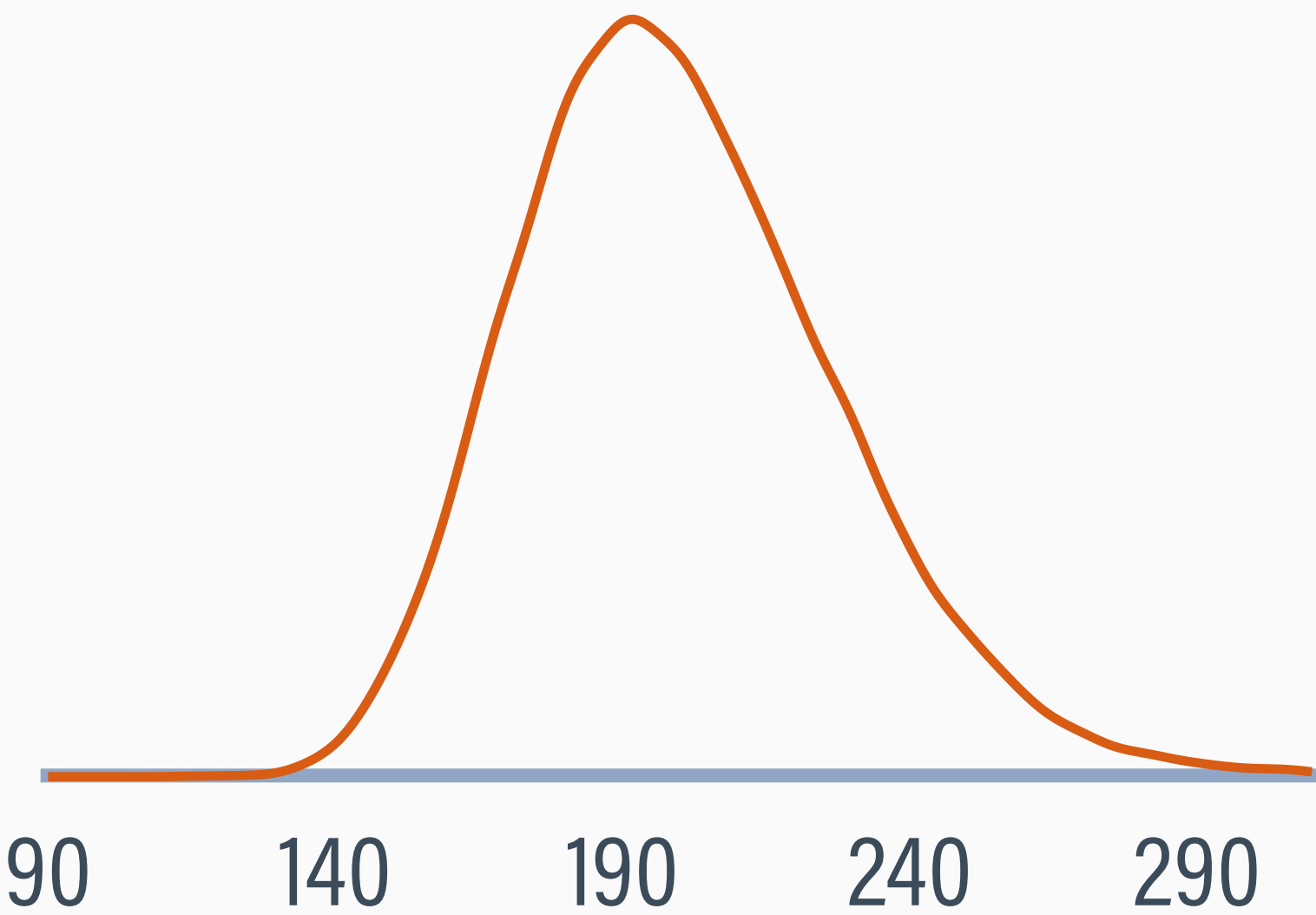


Grade 1 Learning Problems Slope

Median = 197.63

Std. Dev. = 27.62

95% CI = (153.26, 260.29)



Residual Var.

# ESTIMATOR COMPARISON

The two estimators are effectively numerically equivalent even with a small N!!!

Parameter	MCMC			FIML		
	Median	SD	95% CI	Est.	SE	95% CI
Intercept	112.48	6.81	(99.09, 125.87)	112.43	6.58	(99.55, 125.32)
Learning Problems	−0.47	0.13	(−0.73, −0.21)	−0.47	0.13	(−0.72, −0.22)
Residual variance	197.63	27.62	(153.26, 260.29)	189.94	25.24	(140.46, 239.41)
R <sup>2</sup>	.11	.06	(.03, .24)	.12	.06	—

# MCMC AS COMPUTATIONAL FREQUENTISM

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- Researchers adopting a computational frequentism view can use MCMC results as surrogates for ML estimates and frequentist inference (Levy & McNeish, 2021)
- MCMC is simply a flexible way to estimate frequentist quantities in cases where FIML solutions are unavailable (e.g., missing data)



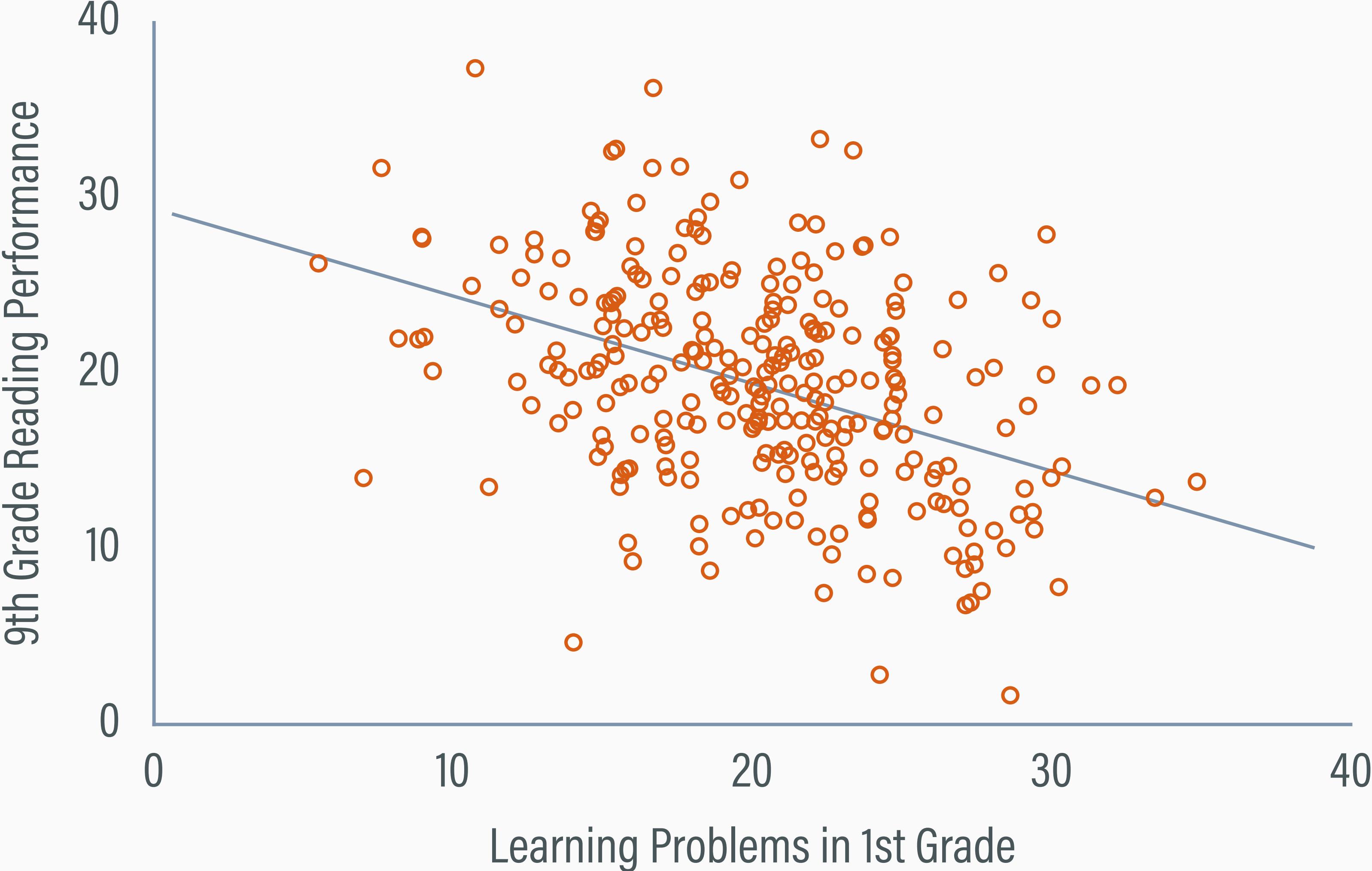
# MISSING DATA IMPUTATION STEP

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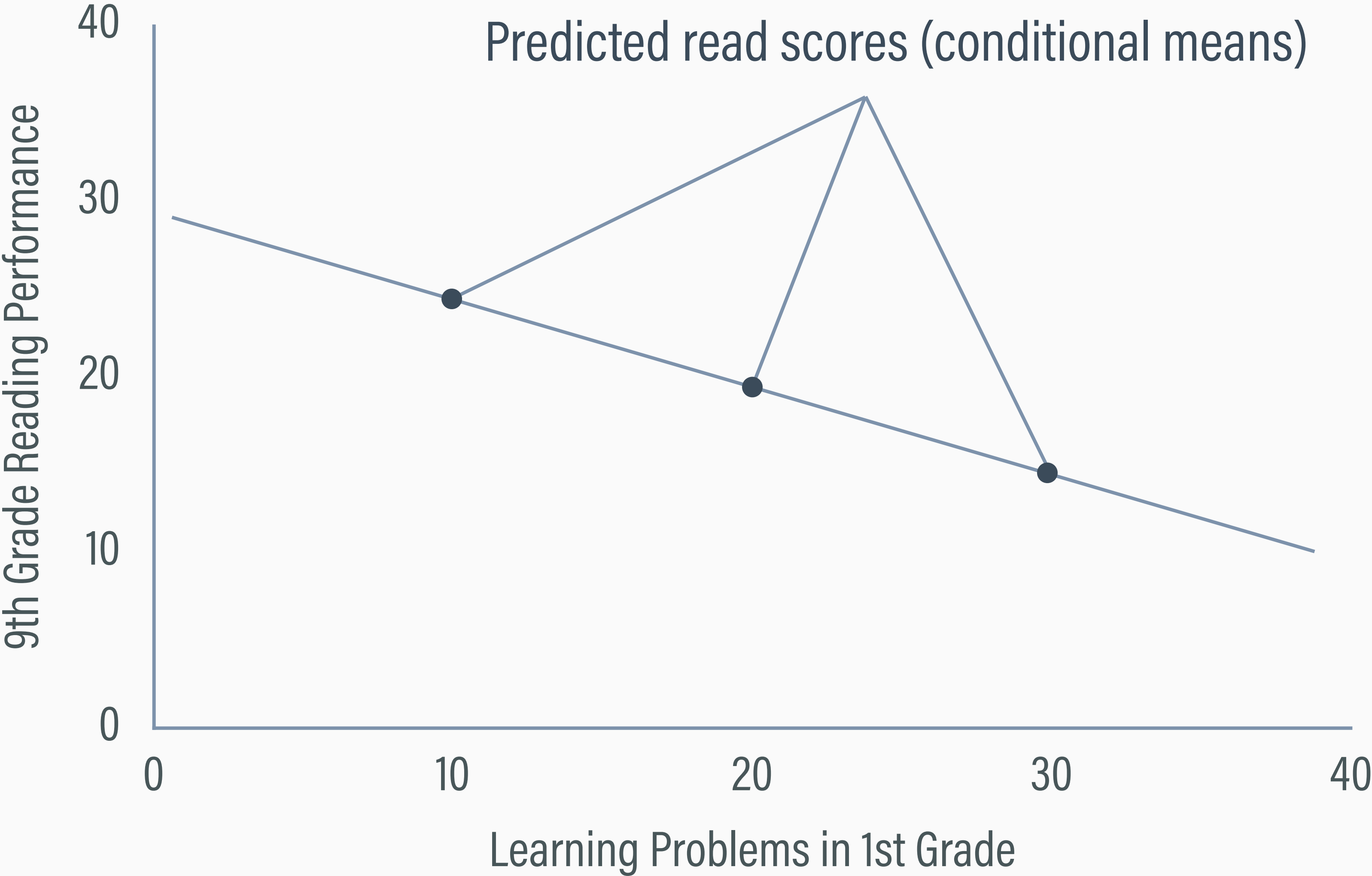
- Missing scores are imputed by drawing replacement scores at random from a distribution of plausible values
- The model parameters combine to define the center and spread of the missing data imputations
- Each iteration yields unique model parameters and unique imputations based on those parameters



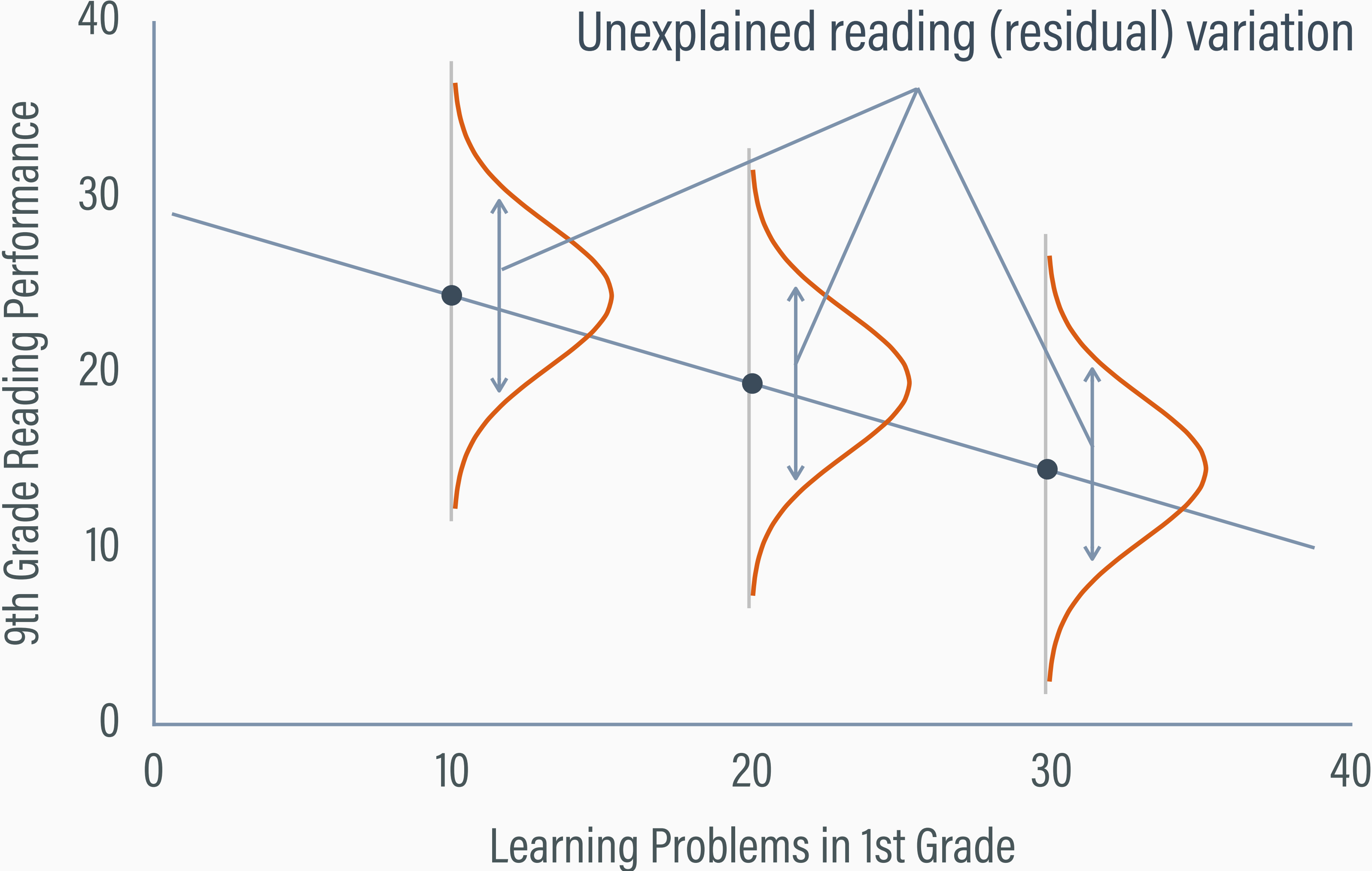
# REGRESSION FROM ONE FILLED-IN DATA SET



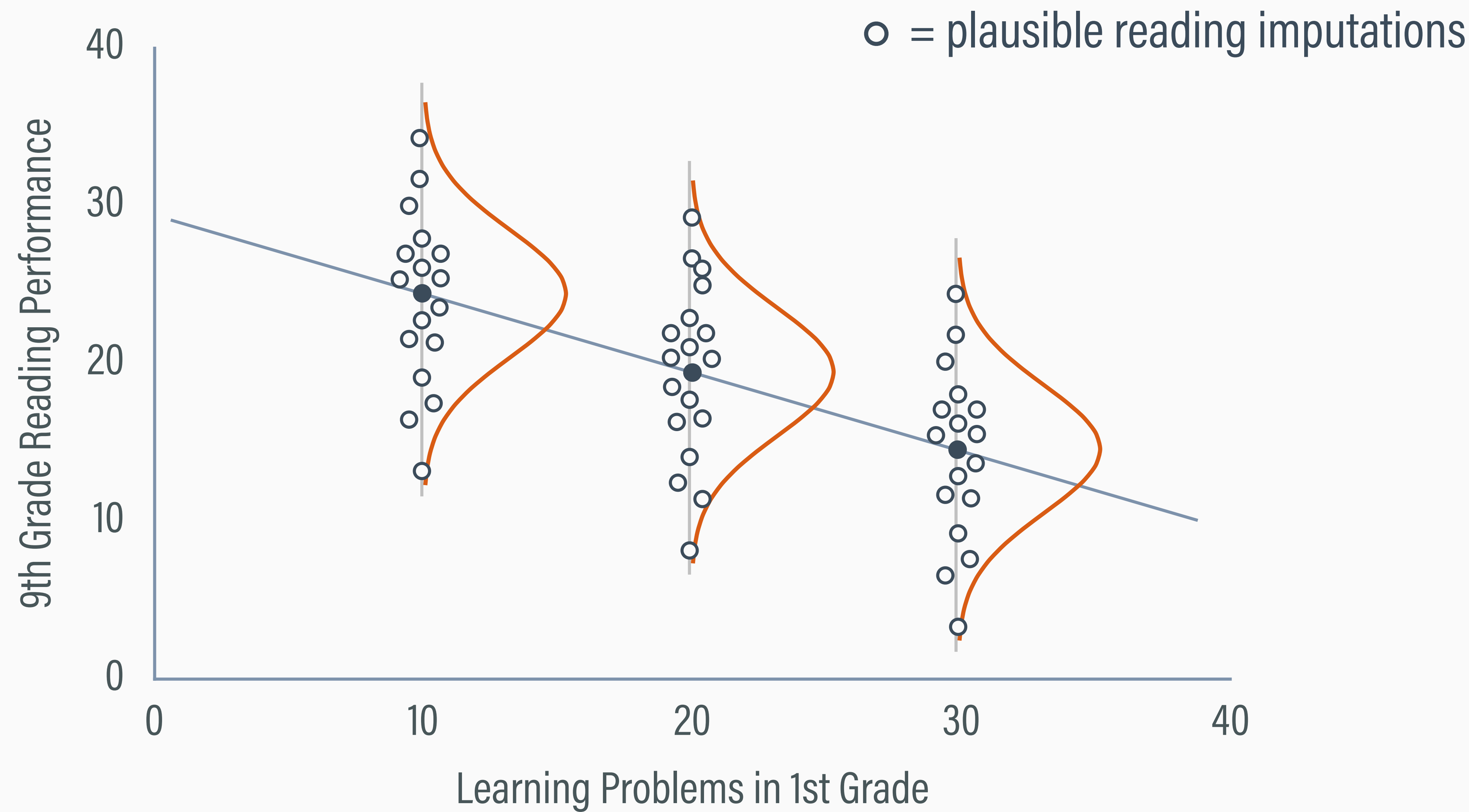
# PREDICTED VALUES



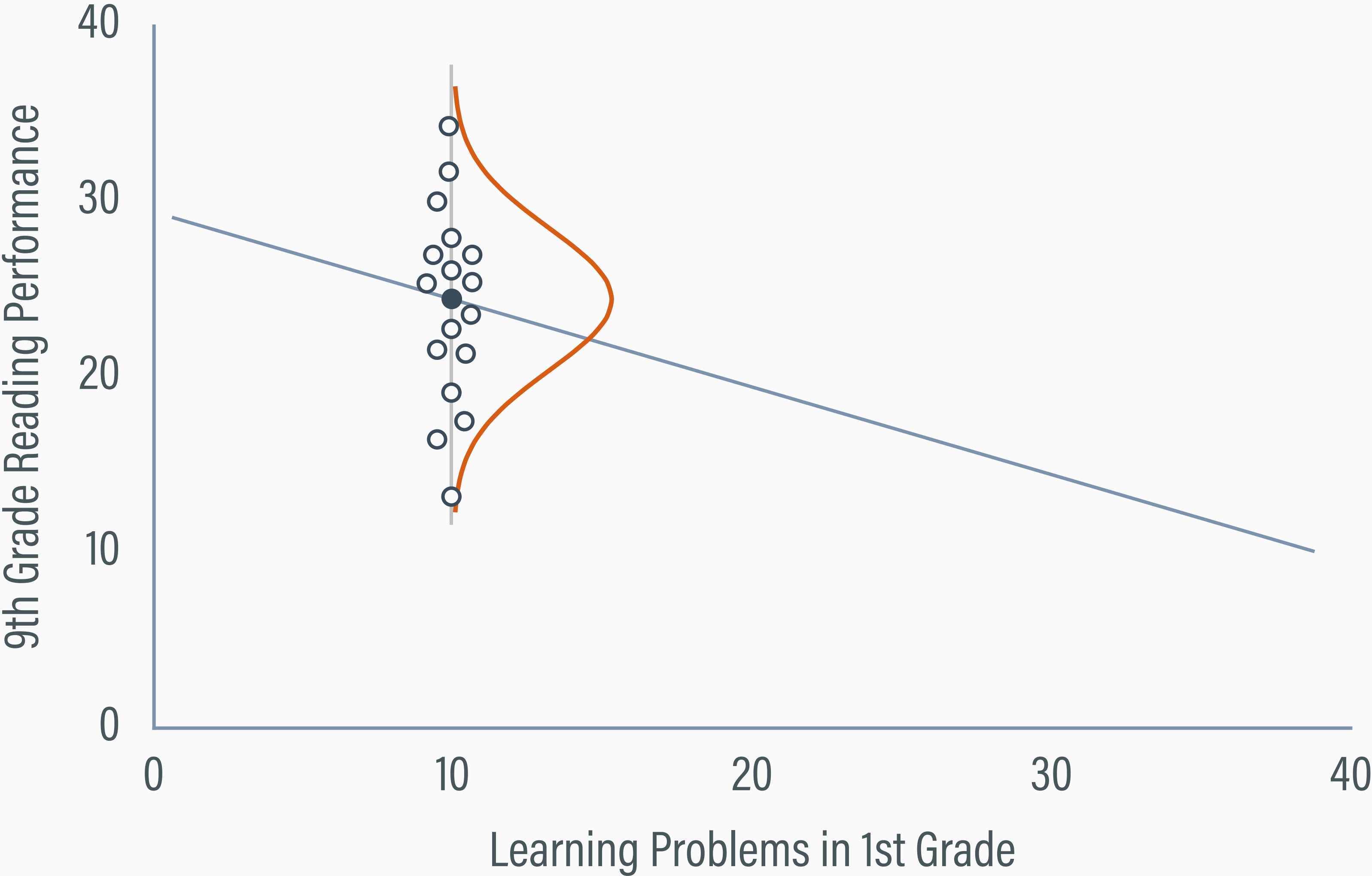
# RESIDUAL VARIATION



# DISTRIBUTIONS OF IMPUTATIONS

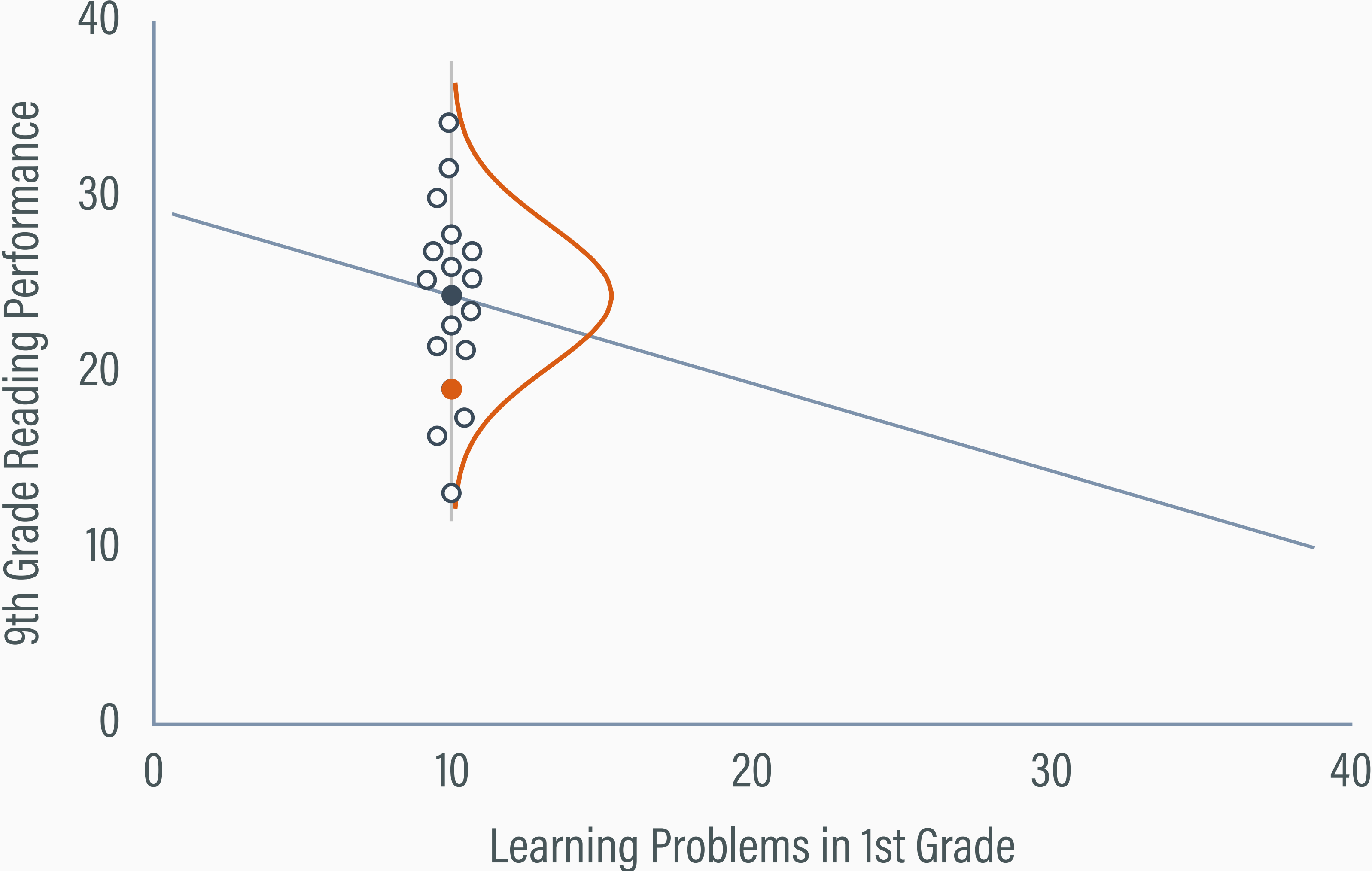


# IMPUTATION FOR LOW LEARNING PROBLEMS



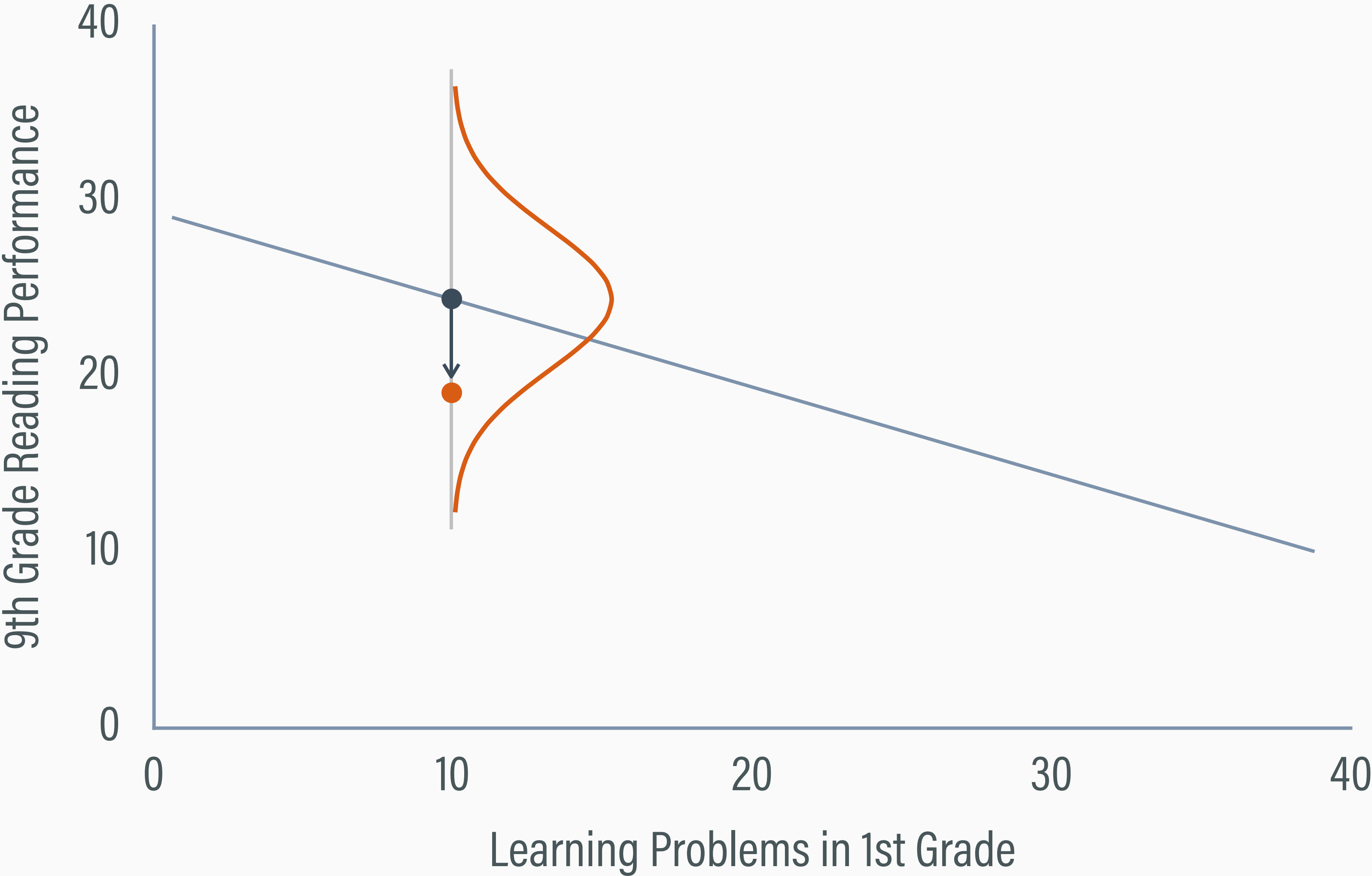


# DRAW AN IMPUTATION AT RANDOM

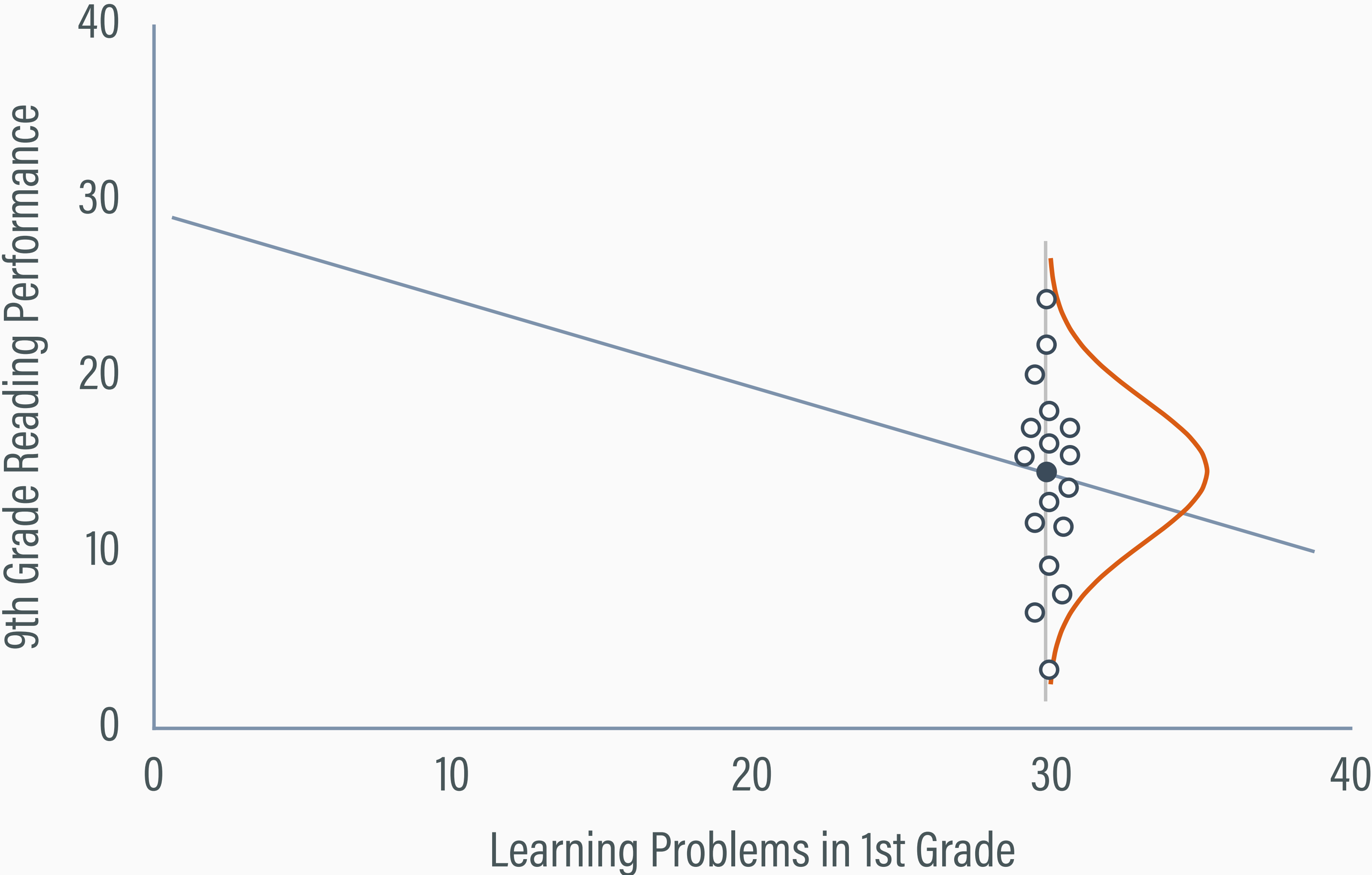


# IMPUTATION = PREDICTION + NOISE

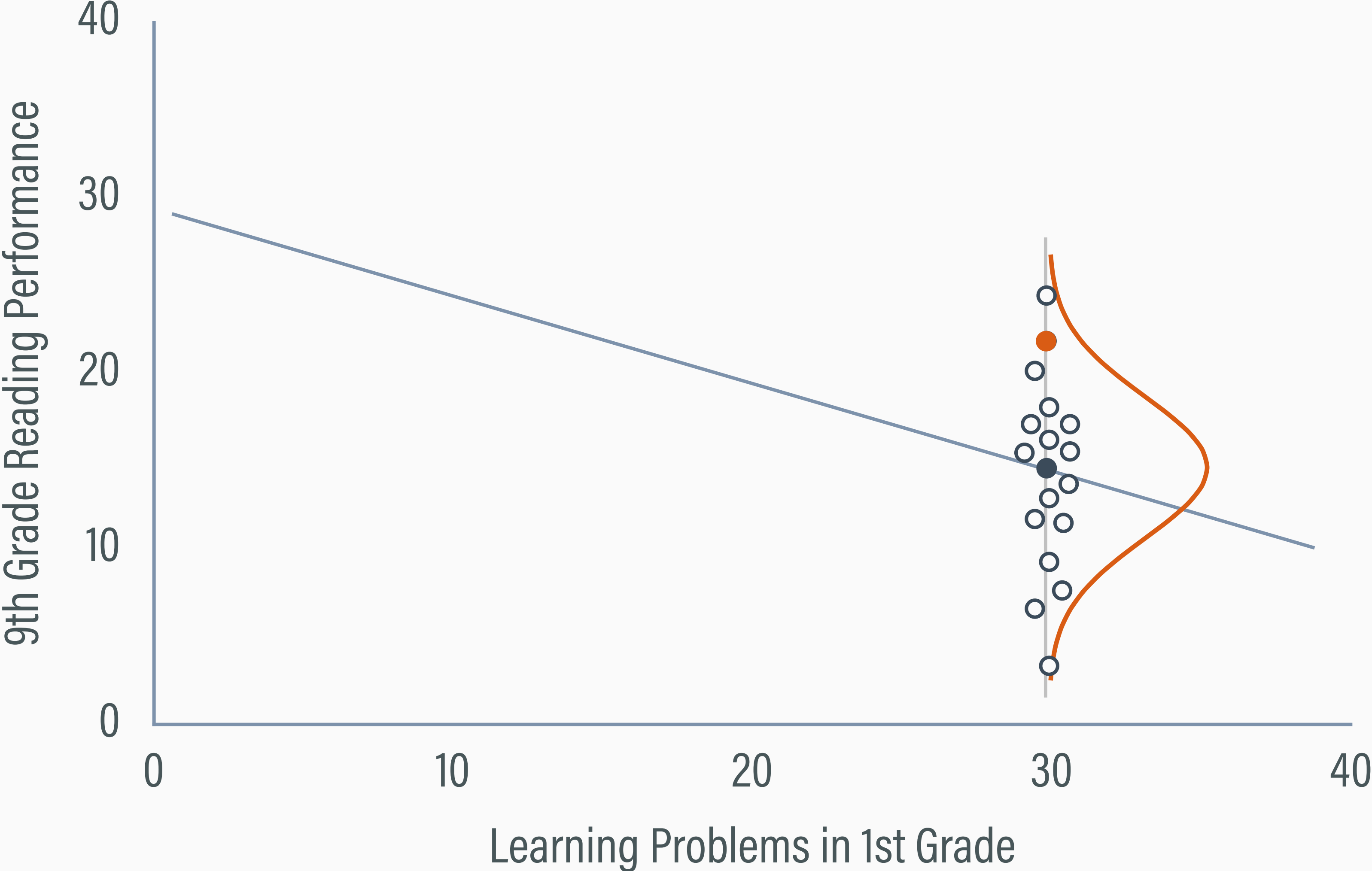
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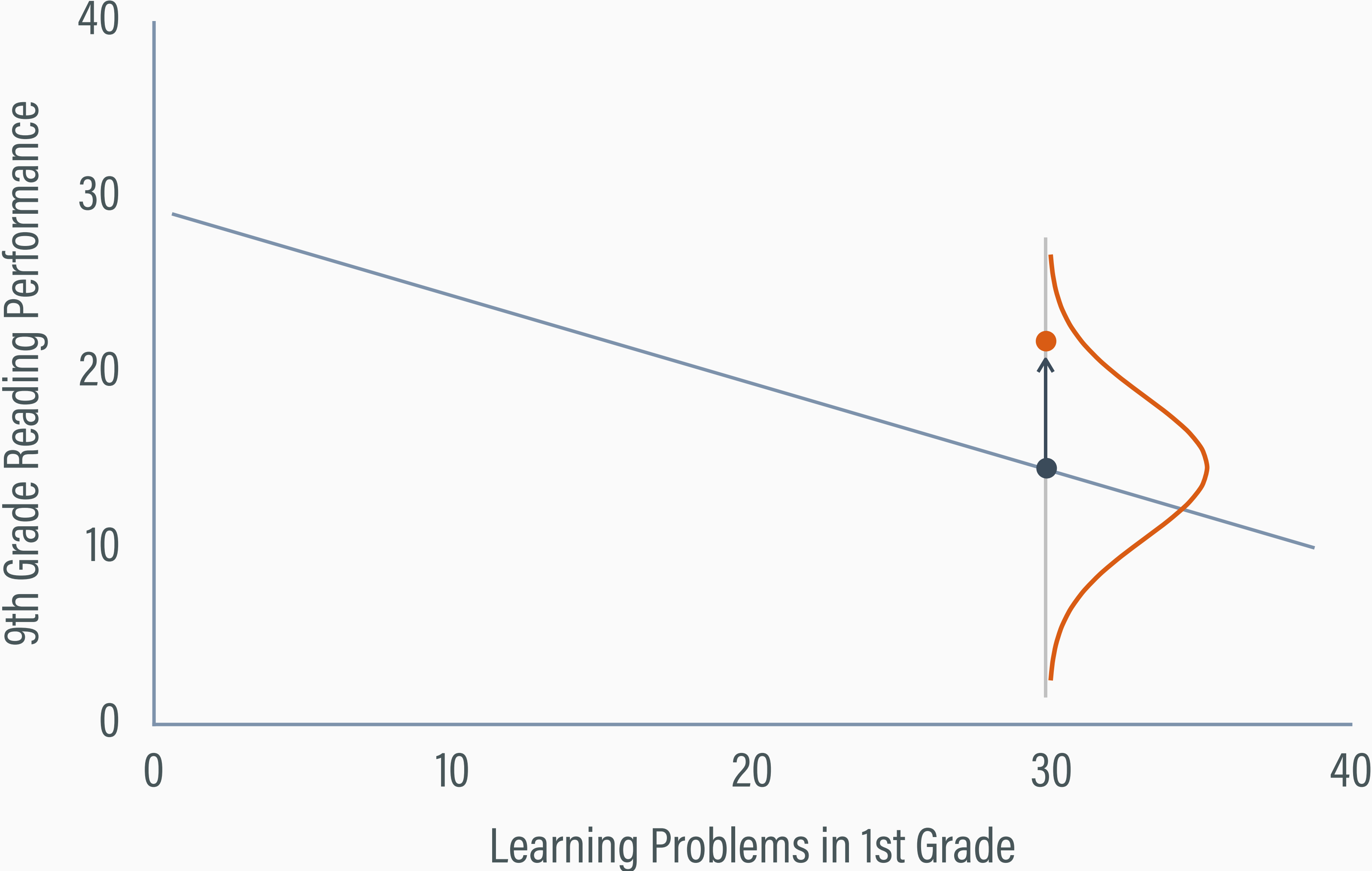
# IMPUTATION FOR HIGH LEARNING PROBLEMS



# DRAW AN IMPUTATION AT RANDOM



# DRAW AN IMPUTATION AT RANDOM





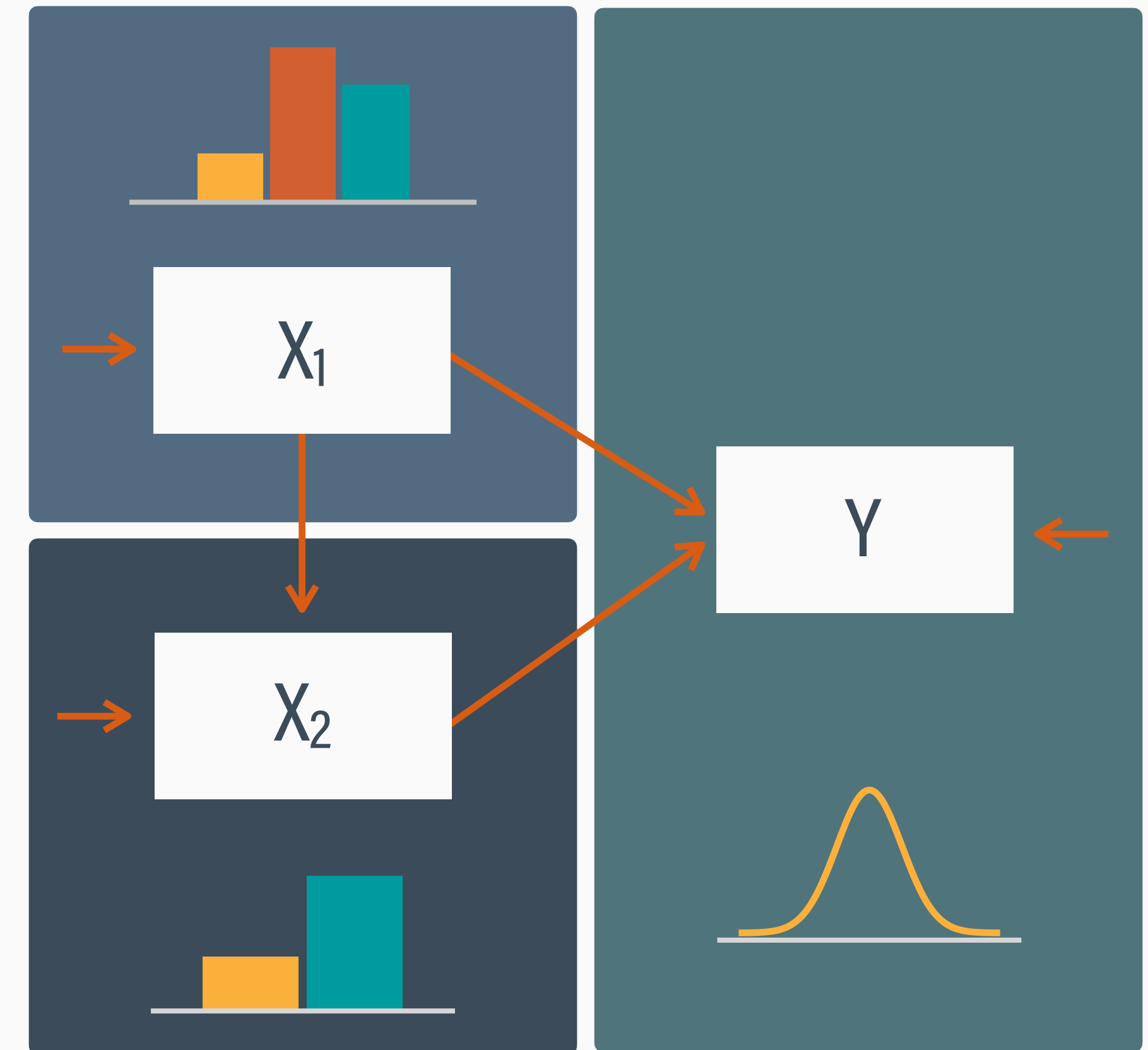
# INCOMPLETE PREDICTORS

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- Incomplete predictors require their own model and distributional assumptions
- Multivariate normal methods (e.g., FIML) can mis-specify the data distributions in a way that introduces bias
- Factored regression uses a modular specification where a sequence of models replaces a general multivariate model

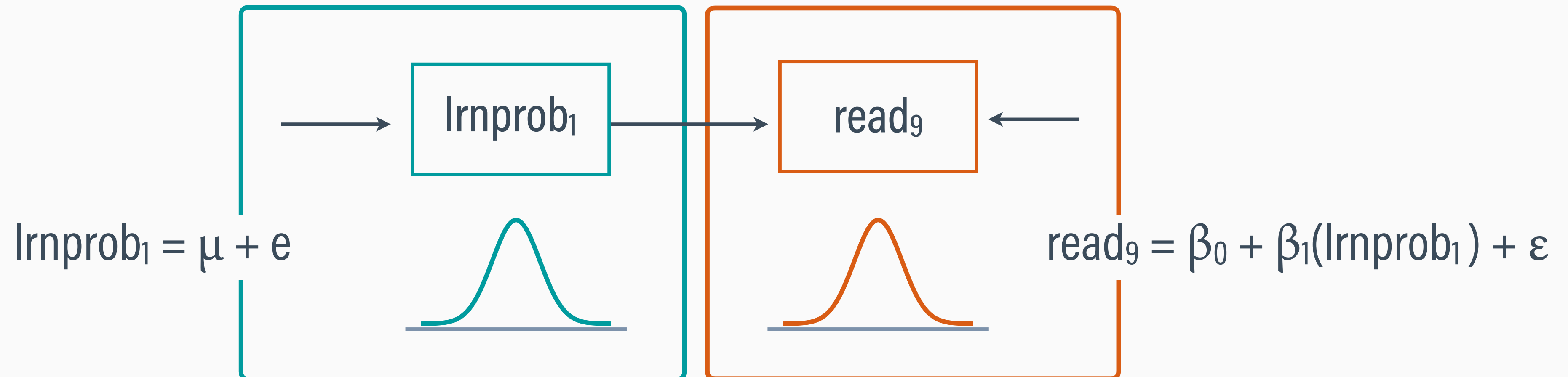
# FACTORED REGRESSION SPECIFICATIONS

- Factored regression specifications invoke a unique distribution for each variable
- The analysis consists of a collection of univariate regression models
- Each model can include terms that are at odds with multivariate normality



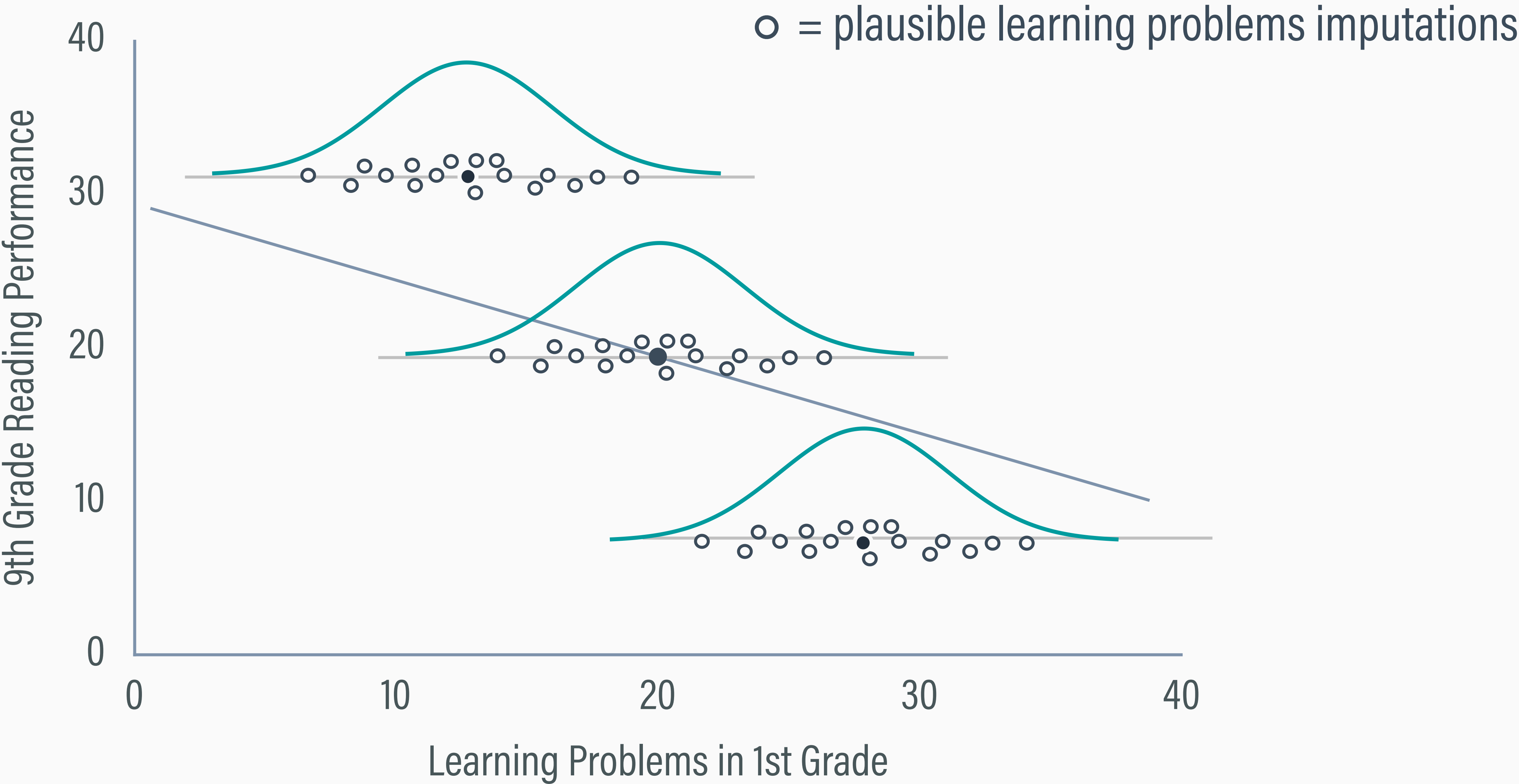
# INCOMPLETE PREDICTORS

- Learning problems is the regressor in the focal model and an outcome in its own empty model



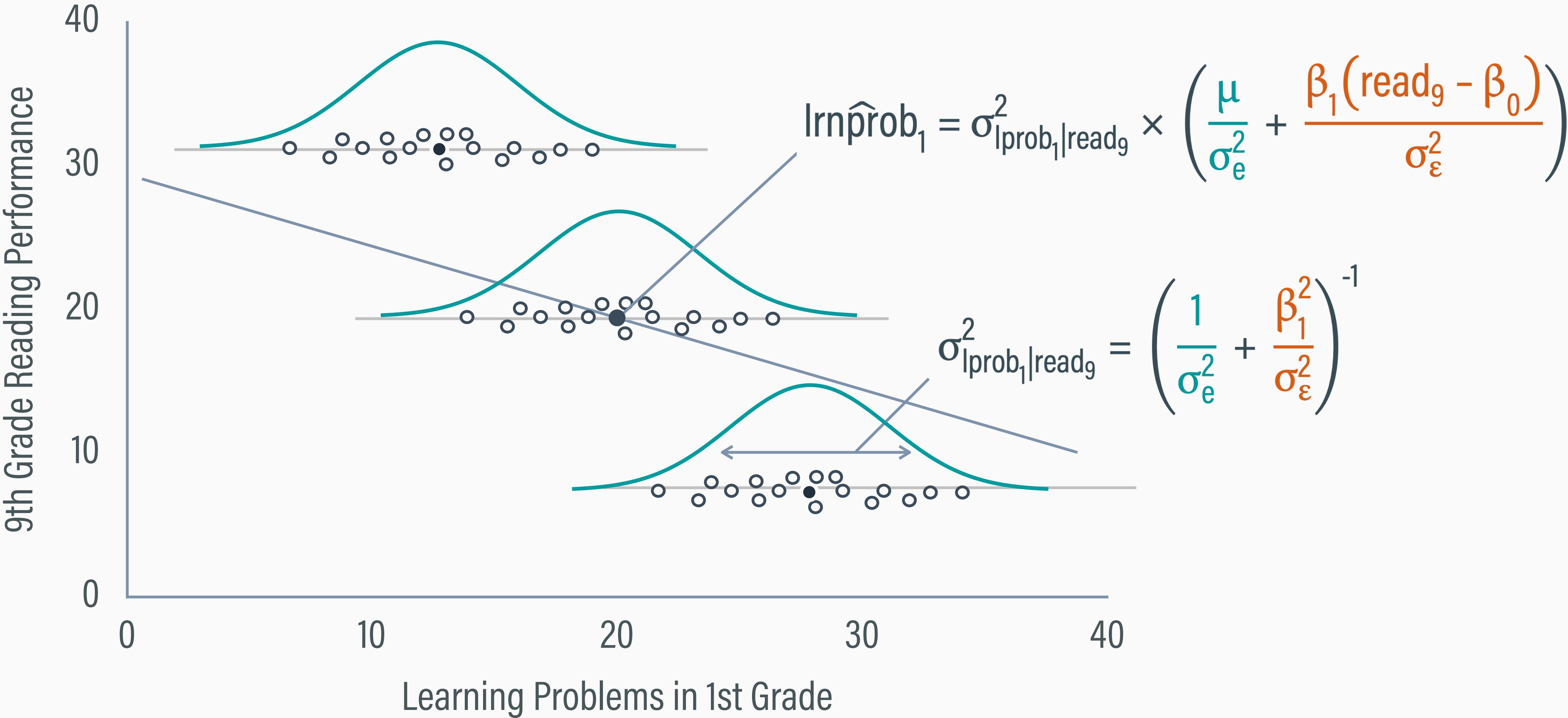
- Both sets of parameter estimates inform the distribution of predictor imputations

# DISTRIBUTIONS OF IMPUTATIONS



# PREDICTED VALUES AND VARIATION

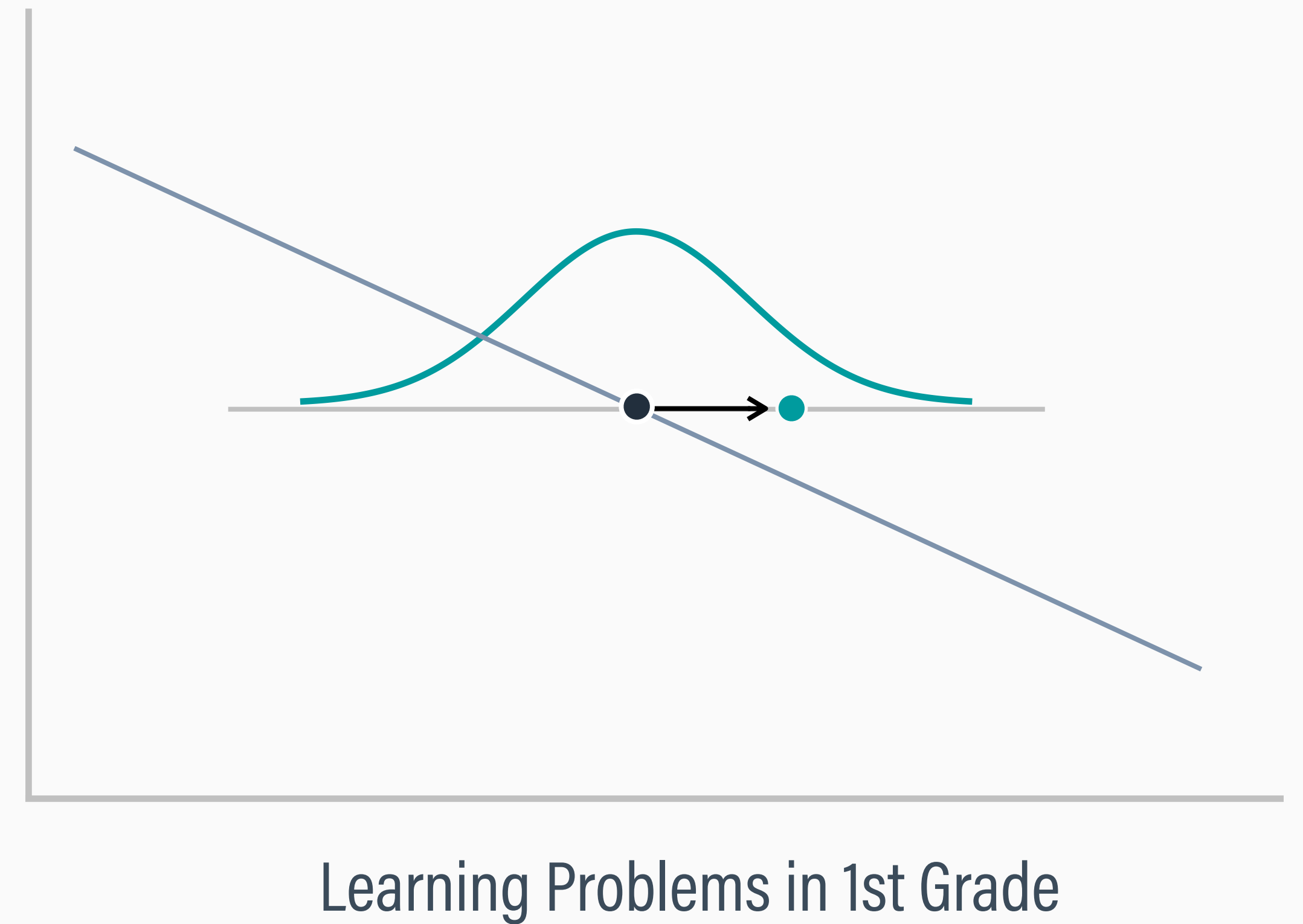
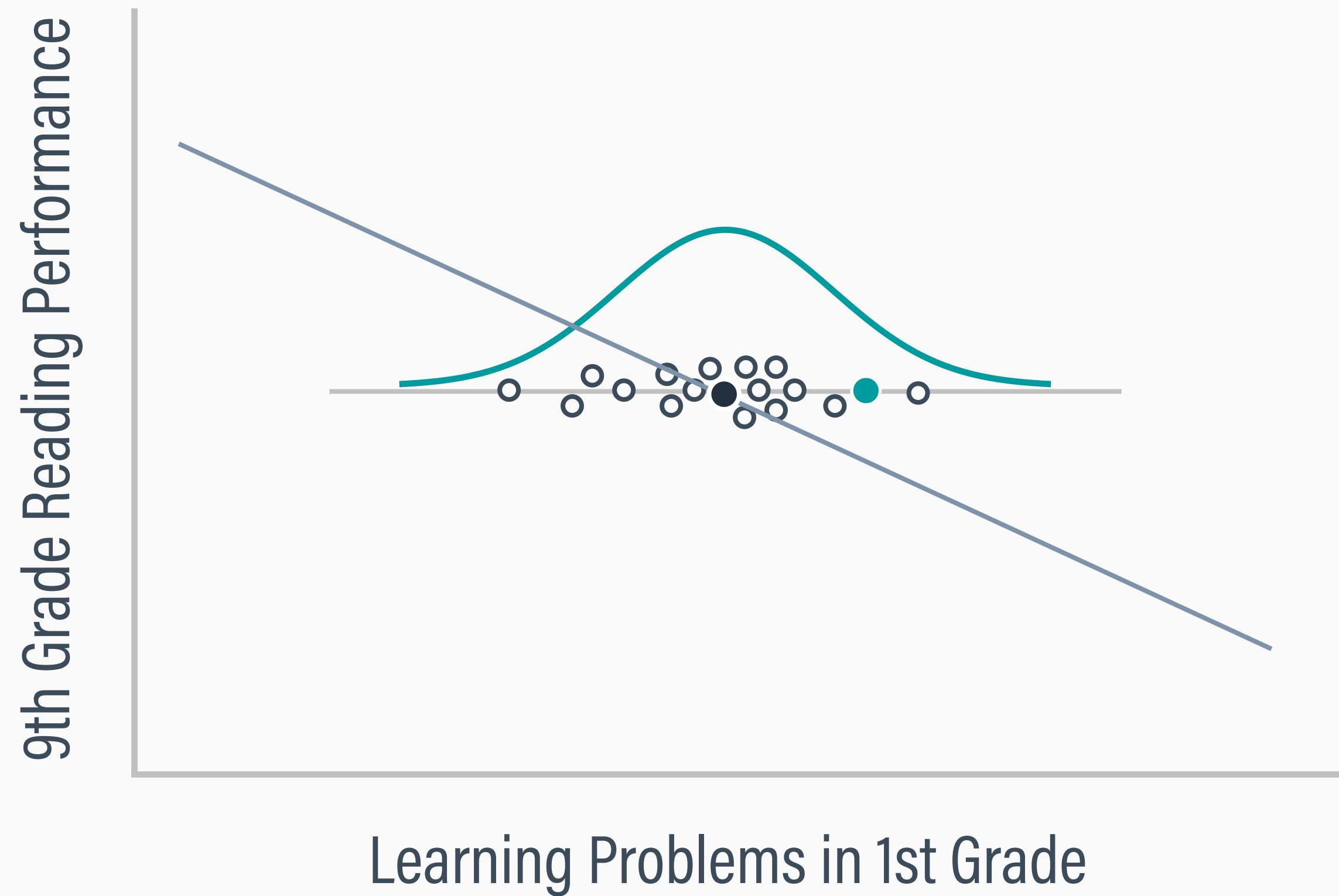
Multiple sets of model parameters define the mean and spread of the imputations





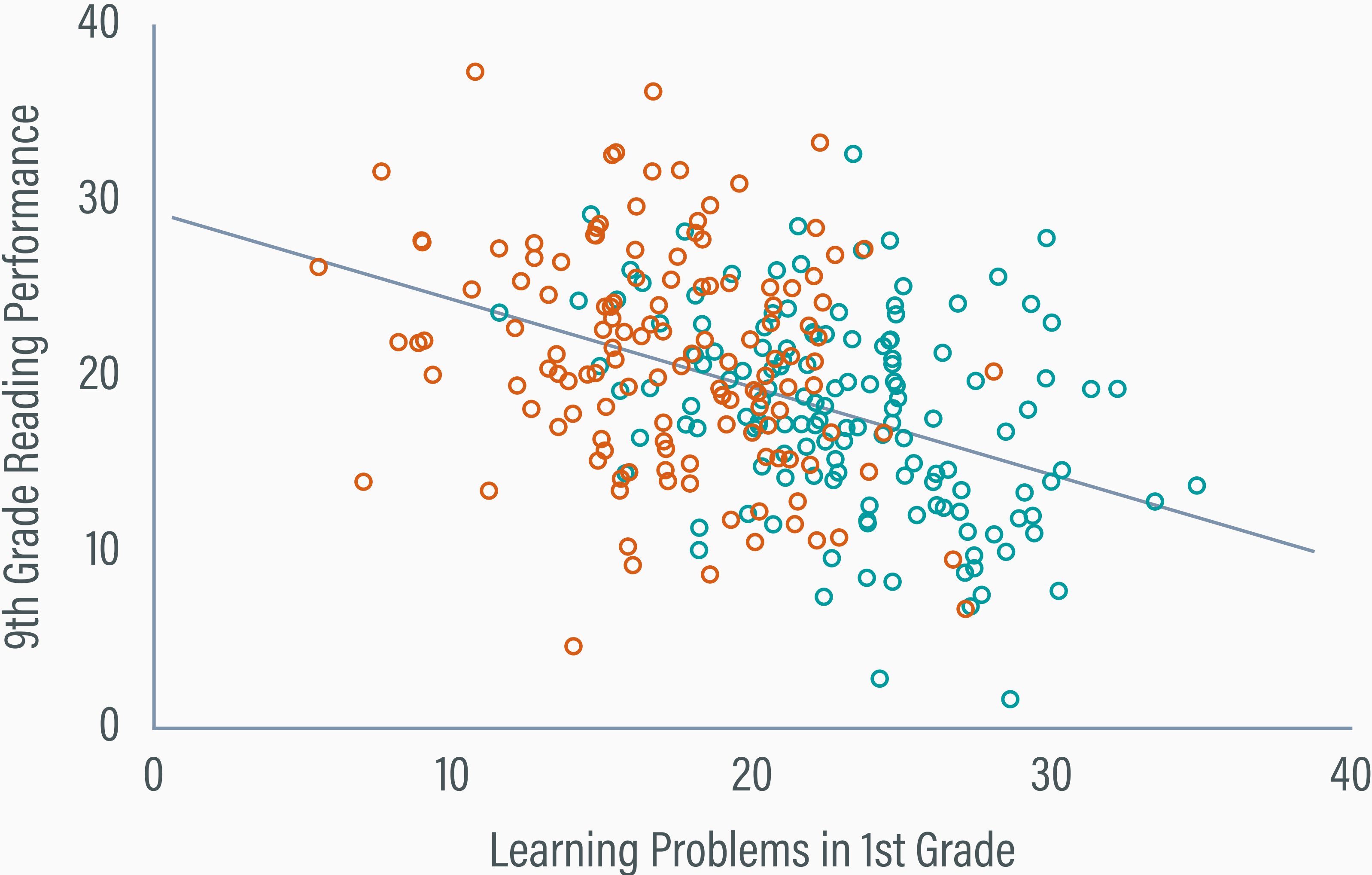
# IMPUTATION EXAMPLE

Imputation = predicted value + random normal noise



# FILLED-IN DATA FROM ONE ITERATION

- Cases with imputed scores
- Cases with complete data



# OUTLINE

1

Modern Missing Data Methods

2

Missing Data Mechanisms

3

MCMC Estimation and Missing Data Imputation

4

Fitting Regression Models in Blimp

5

Incomplete Categorical Variables

6

Interaction Effects

7

Reporting Results From a Missing Data Analysis

# ANALYSIS 1 REGRESSION MODEL

- Academic variables from 1st grade predicting 9th grade reading achievement in middle school

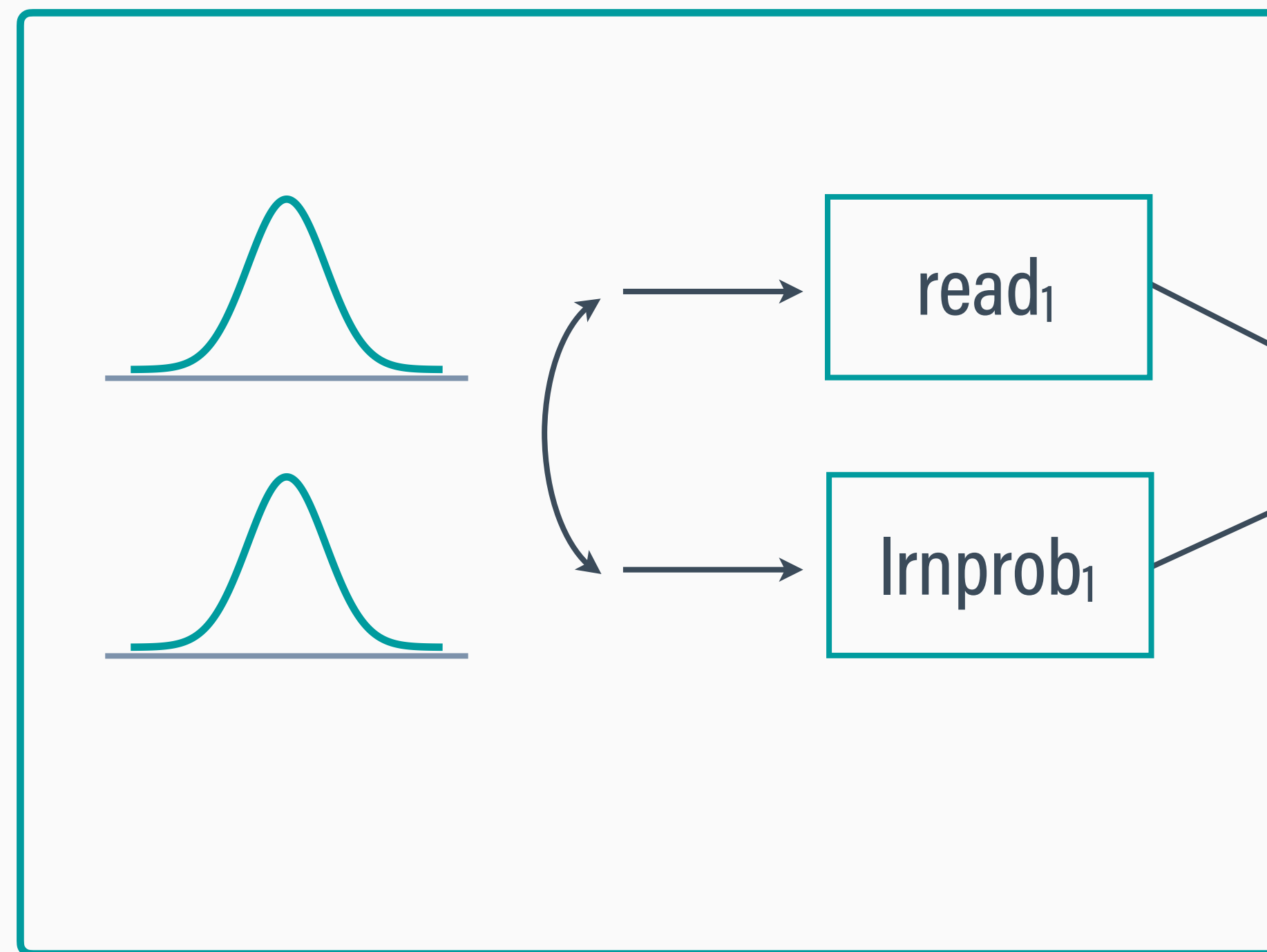
$$\text{read}_9 = \beta_0 + \beta_1(\text{read}_1) + \beta_2(\text{lrnprob}_1) + \varepsilon$$

Variable	Definition	Missing %	Scale
atrisk	Emotional/behavioral risk code	2.2	0 = Low, 1 = Medium/high
lrnprob1	1st grade learning problems	2.2	Numeric (31 to 88)
read1	1st grade broad reading composite	6.5	Numeric (39 to 153)
read9	9th grade broad reading composite	17.4	Numeric (41 to 123)

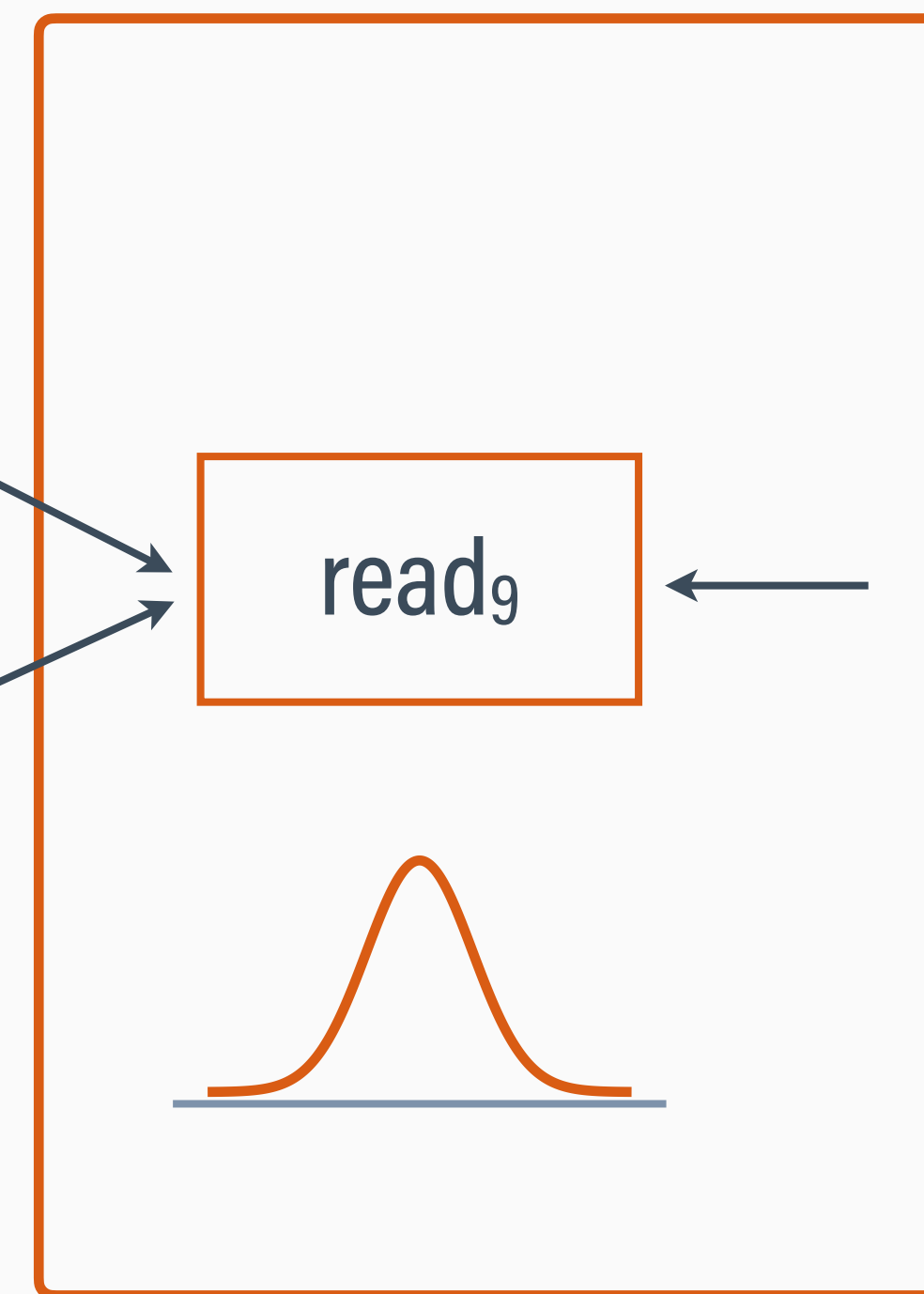
# TWO-PART FACTORED SPECIFICATION

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Incomplete Predictor Model



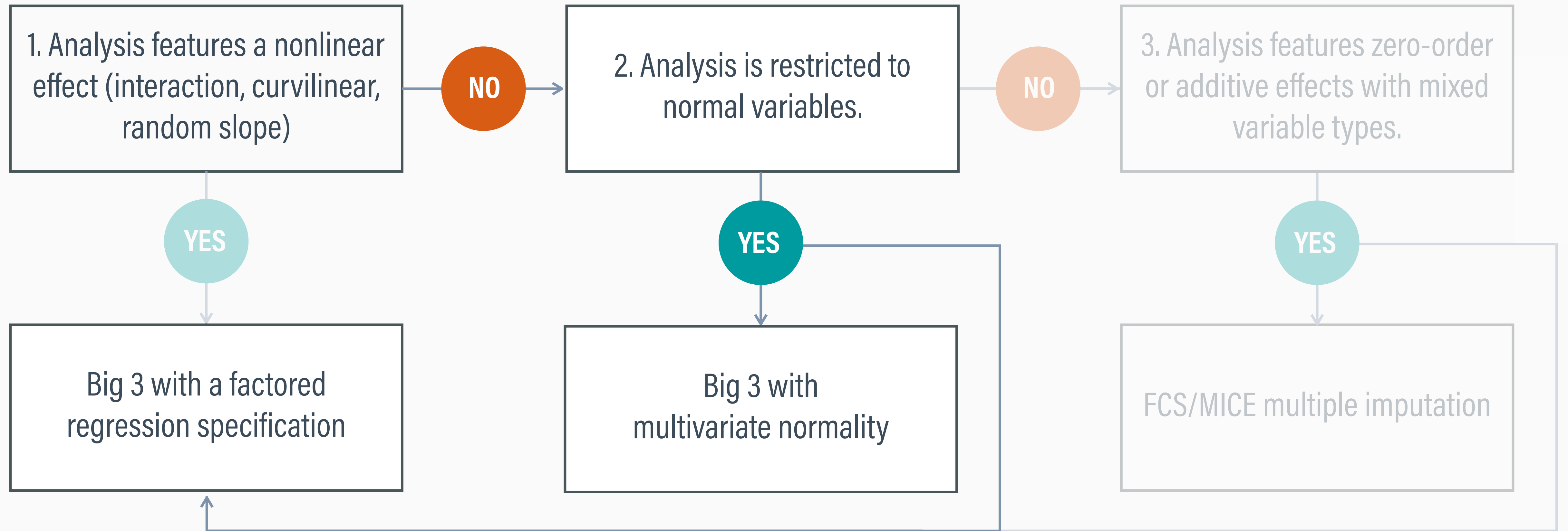
Outcome Model





# MISSING DATA DECISION TREE

This is an analysis where we have multiple equivalent options for missing data handling!



# BLIMP STUDIO SCRIPT 1

---

**DATA:** reading.dat;

**VARIABLES:** id male hispanic riskgrp atrisk behsymp1 lnprob1 read1 read2 read3 read9  
read9grp stanread7 math1 math2 math3 math9 math9grp stanmath7;

**MISSING:** 999;

**MODEL:** read9 ~ read1 lnprob1;

**BURN:** 10000;

**ITER:** 10000;

**SEED:** 90291;

# DATA AND VARIABLES

---

```
DATA: reading.dat;                # data in same directory as the script
VARIABLES: id male hispanic riskgrp atrisk behsymp1 lnprob1 read1 read2 read3 read9
               read9grp stanread7 math1 math2 math3 math9 math9grp stanmath7; # name data columns
MISSING: 999;                   # missing value code
MODEL: read9 ~ read1 lnprob1;
BURN: 10000;
ITER: 10000;
SEED: 90291;
```

# MODEL DETAILS

---

**DATA:** reading.dat;

**VARIABLES:** id male hispanic riskgrp atrisk behsymp1 lnprob1 read1 read2 read3 read9  
read9grp stanread7 math1 math2 math3 math9 math9grp stanmath7;

**MISSING:** 999;

**MODEL:** read9 ~ read1 lnprob1;      # regression model

**BURN:** 10000;

**ITER:** 10000;

**SEED:** 90291;

# COMPUTATIONAL DETAILS

---

**DATA:** reading.dat;

**VARIABLES:** id male hispanic riskgrp atrisk behsymp1 lnprob1 read1 read2 read3 read9  
read9grp stanread7 math1 math2 math3 math9 math9grp stanmath7;

**MISSING:** 999;

**MODEL:** read9  $\sim$  read1 lnprob1;

**BURN:** 10000;

# number of warm-up iterations

**ITER:** 10000;

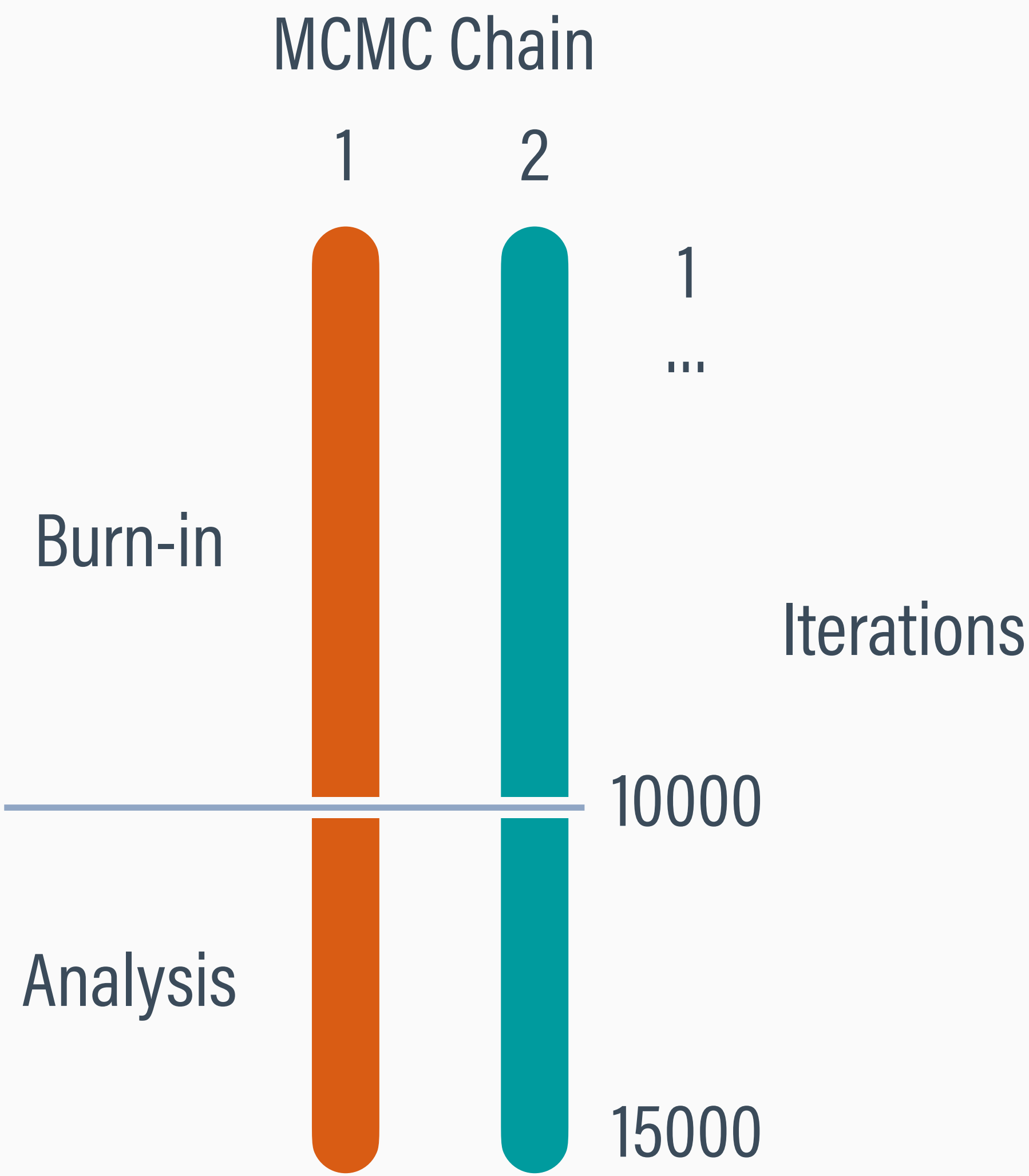
# number of parameter values for the analysis

**SEED:** 90291;

# integer seed for Monte Carlo simulation

# COMPUTATIONAL DETAILS

**BURN:** 10000; # warm-up iterations (per chain)  
**ITER:** 10000; # estimates to summarize (both chains)  
**SEED:** 90291; # random number seed





# RBLIMP SCRIPT (MODEL 1)

---

# fit model

model1 <- rblimp(

  data = reading,

  model = 'read9 ~ read1 llnprob1',

  seed = 90291,

  burn = 10000,

  iter = 10000)

# summarize results

output(model1)

posterior\_plot(model1, 'read9')

# R data frame

# regression model

# integer seed for Monte Carlo simulation

# number of warm-up iterations

# number of parameter values for the analysis

# print output

# plot parameter distributions

# MISSING DATA INFORMATION

- Per-variable missingness rates appear in the rows, and missing data patterns appear in the columns (M = missing)

DATA INFORMATION:

Sample Size: 138

Missing Data Info:

	miss %	1	2	3	4	5
read9 = 17.4		-	M	-	-	M
lrrnprob1 = 2.2		-	-	-	M	-
read1 = 6.5		-	-	M	-	M
	%	74.6	16.7	5.8	2.2	0.7

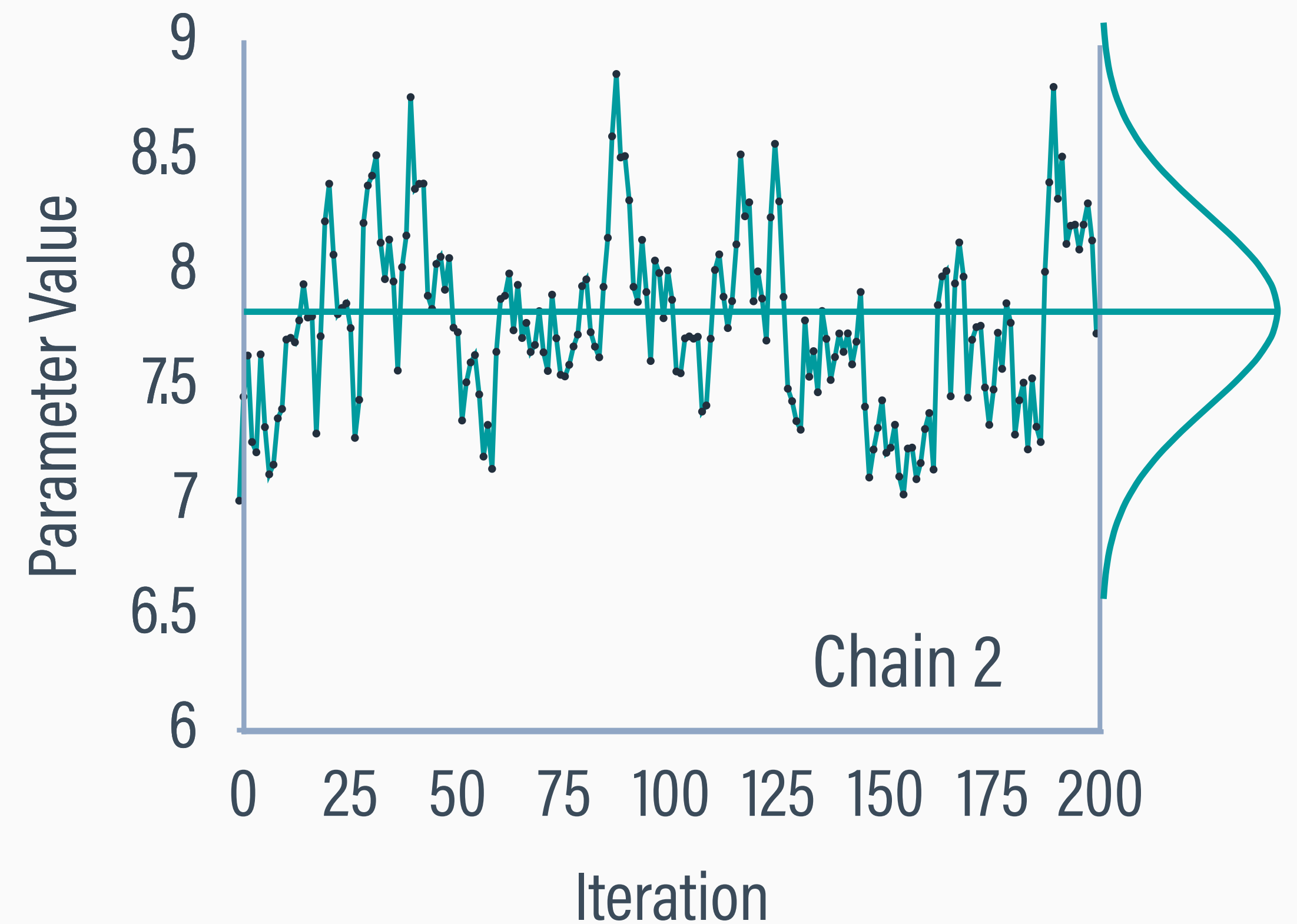
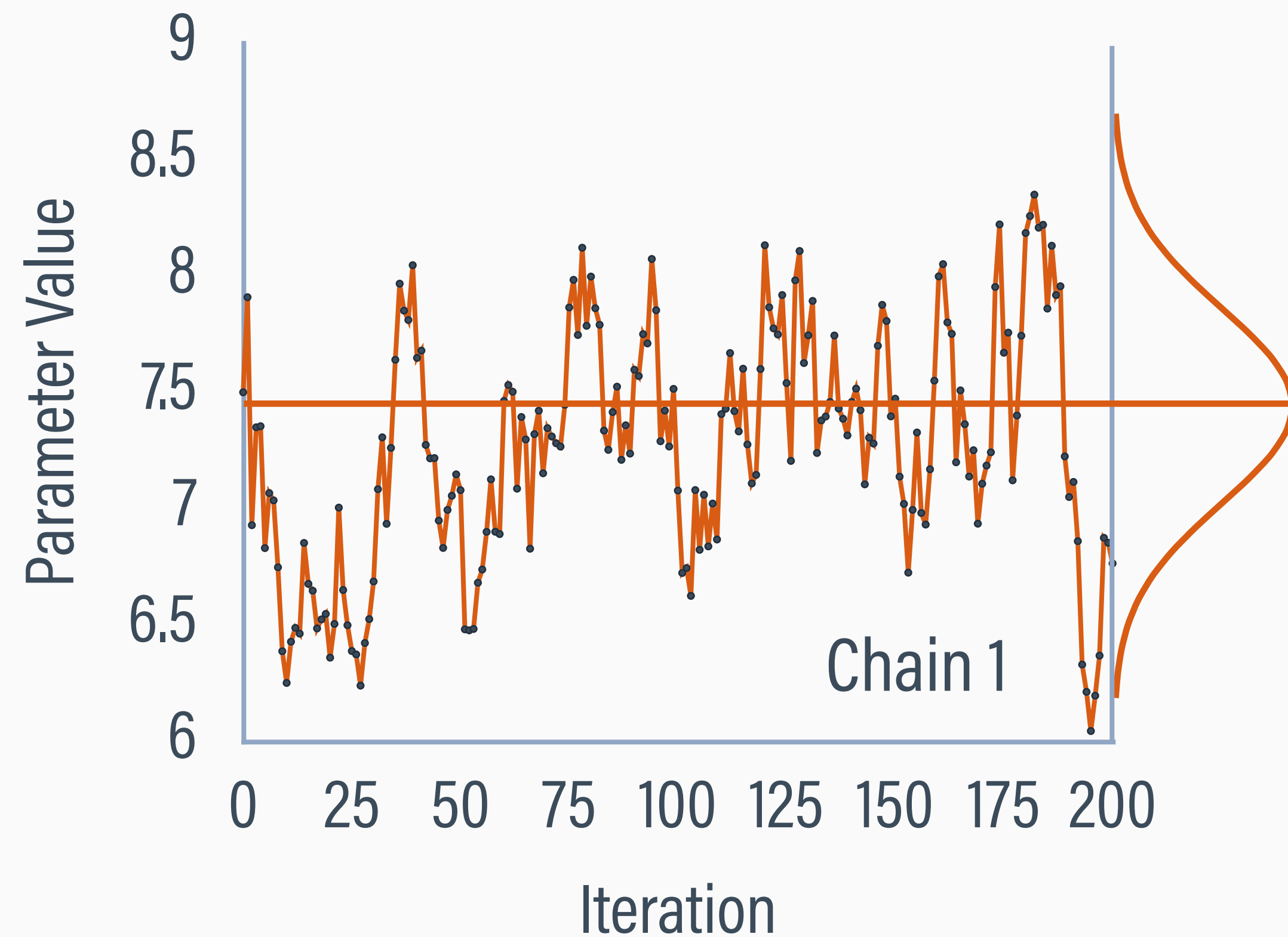
# MCMC CONVERGENCE

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- MCMC parameter estimates continually vary across iterations
- MCMC converges when parameter estimates oscillate around a stable mean, and variation doesn't change with more iterations
- The potential scale reduction factor (PSRF) compares the similarity of parameters generated from two MCMC processes

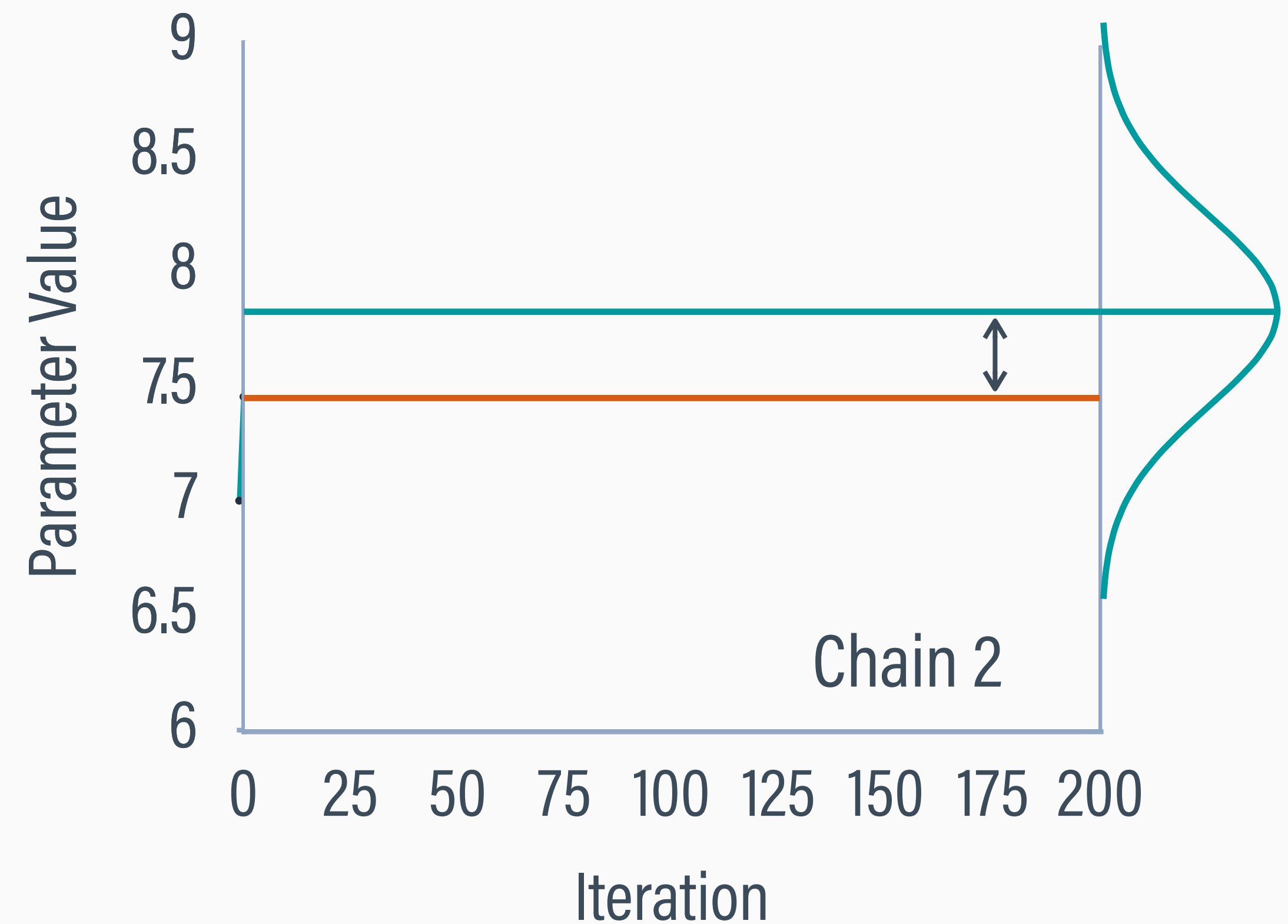
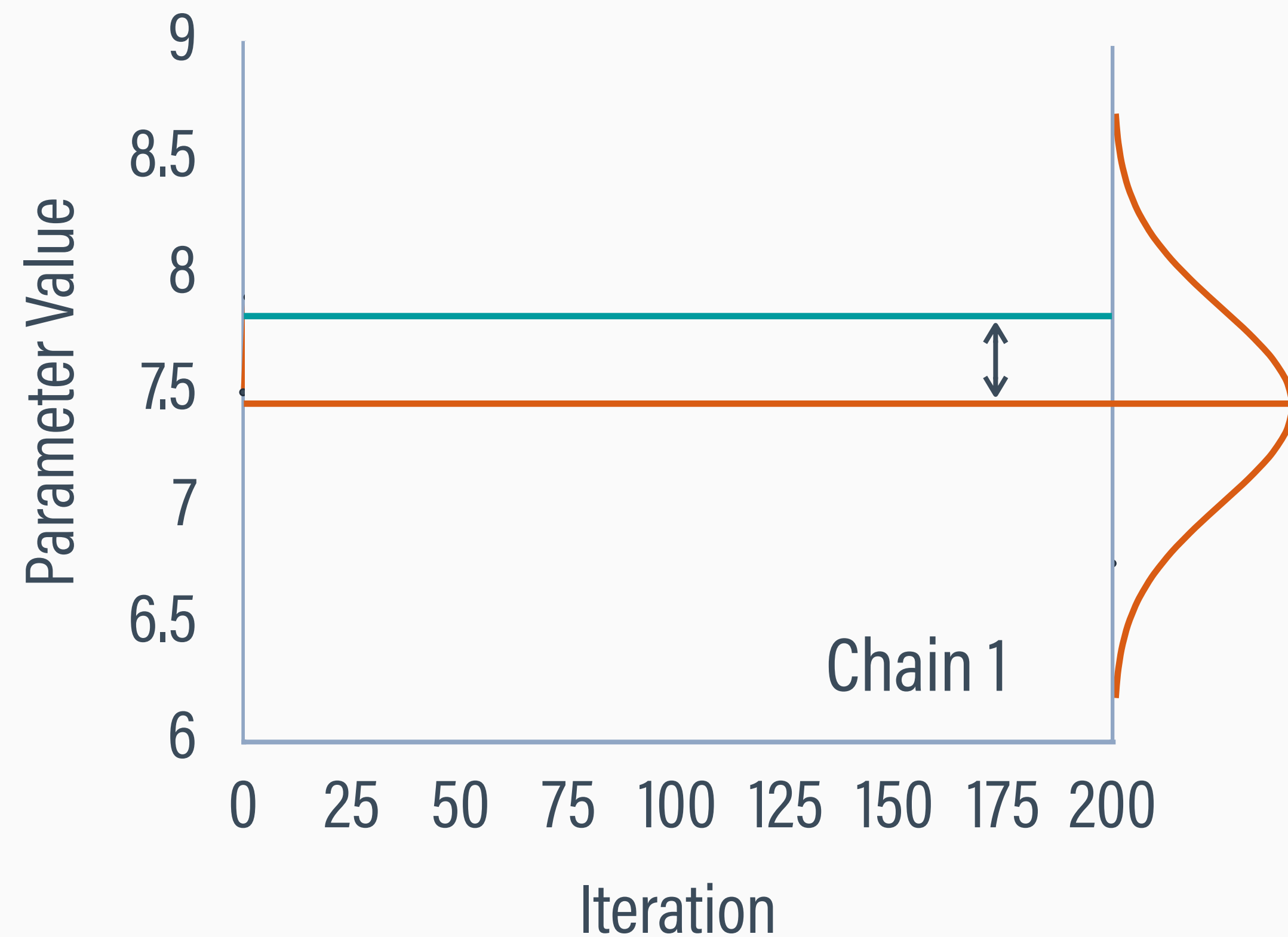
# POTENTIAL SCALE REDUCTION FACTOR

$$\text{PSRF} = \sqrt{\frac{\text{mean difference between chains} + \text{within-chain variation}}{\text{within-chain variation}}}$$



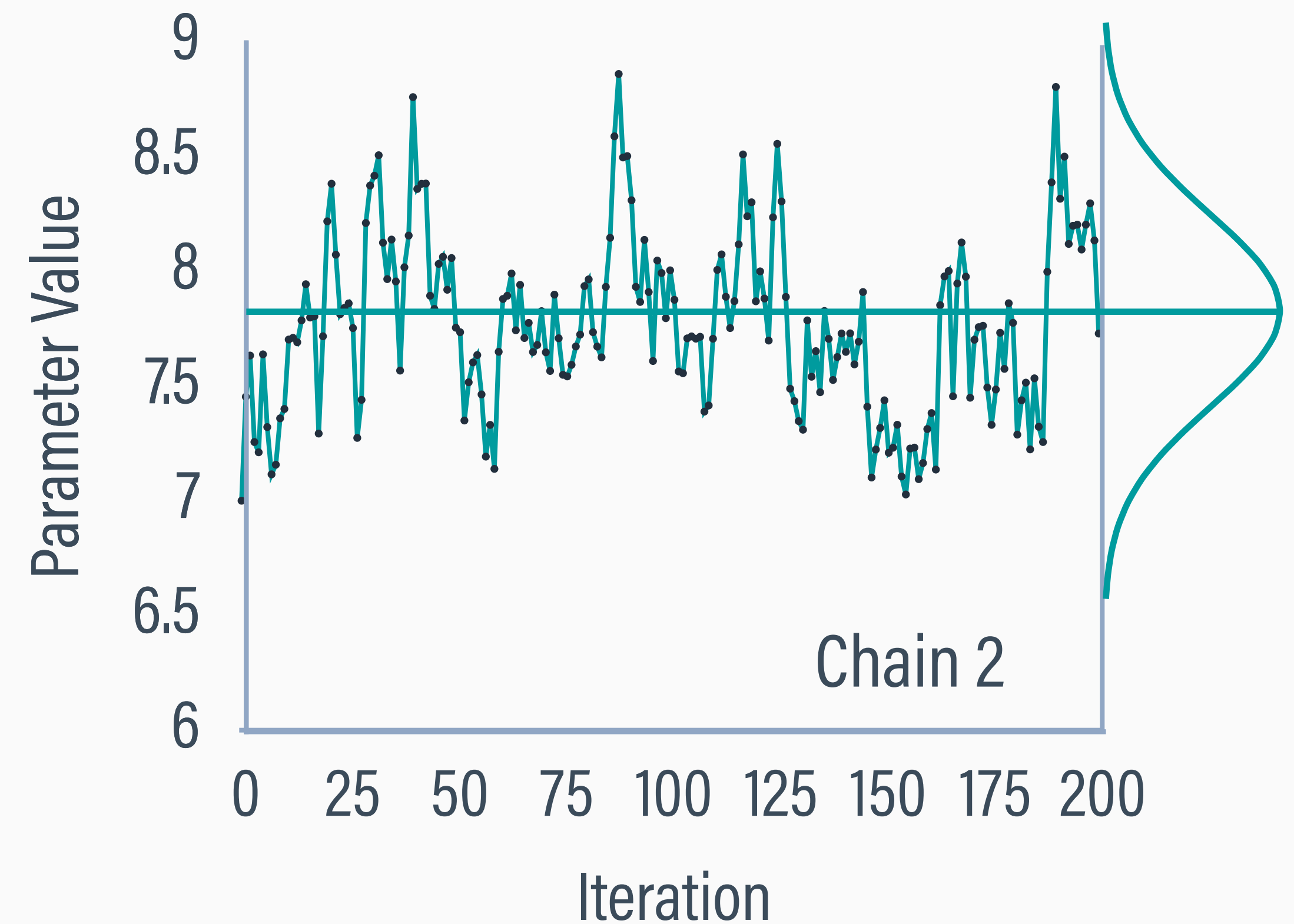
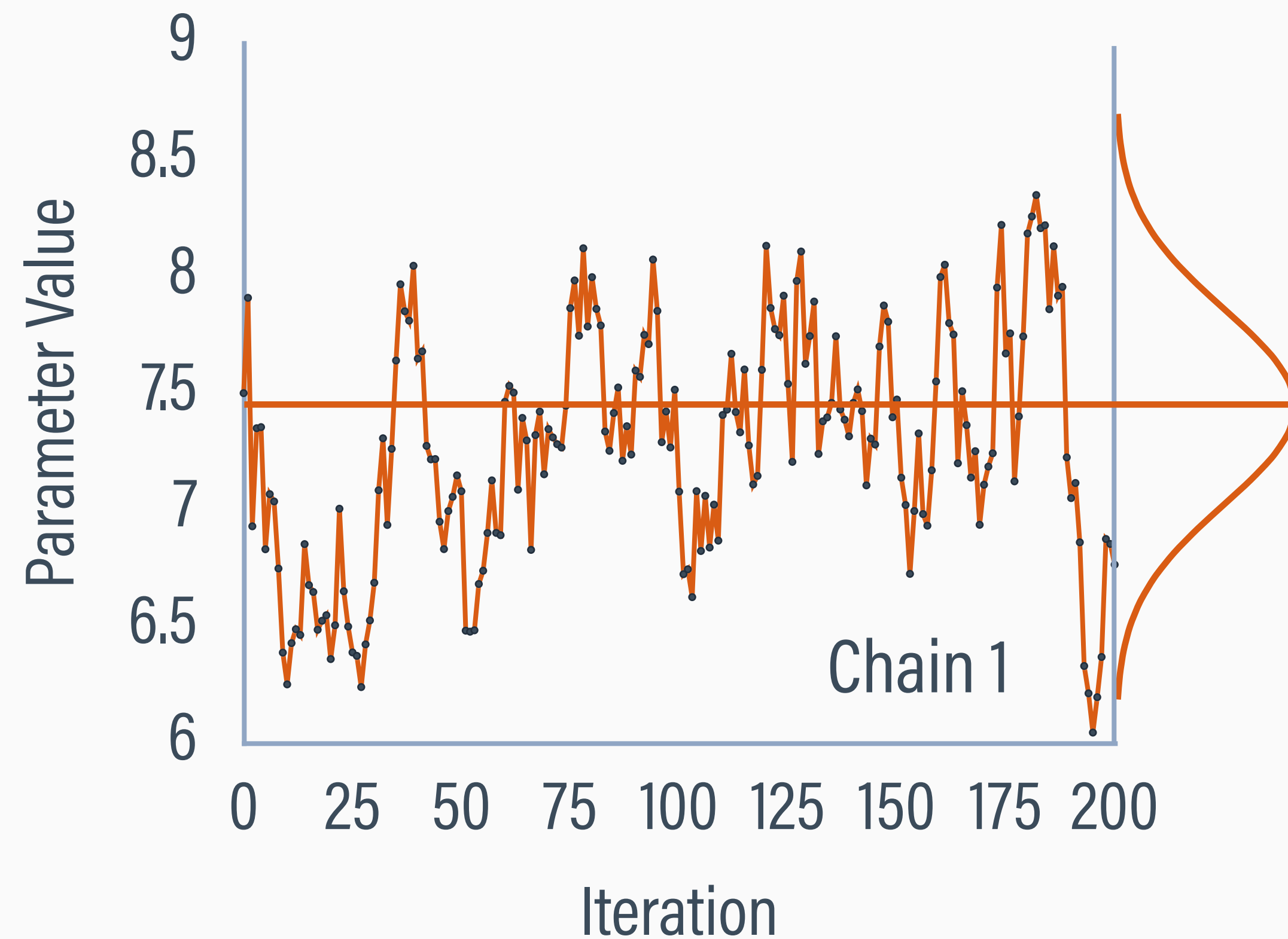
# BETWEEN-CHAIN MEAN DIFFERENCE

$$\text{PSRF} = \sqrt{\frac{\text{mean difference between chains} + \text{within-chain variation}}{\text{within-chain variation}}}$$



# WITHIN-CHAIN VARIATION

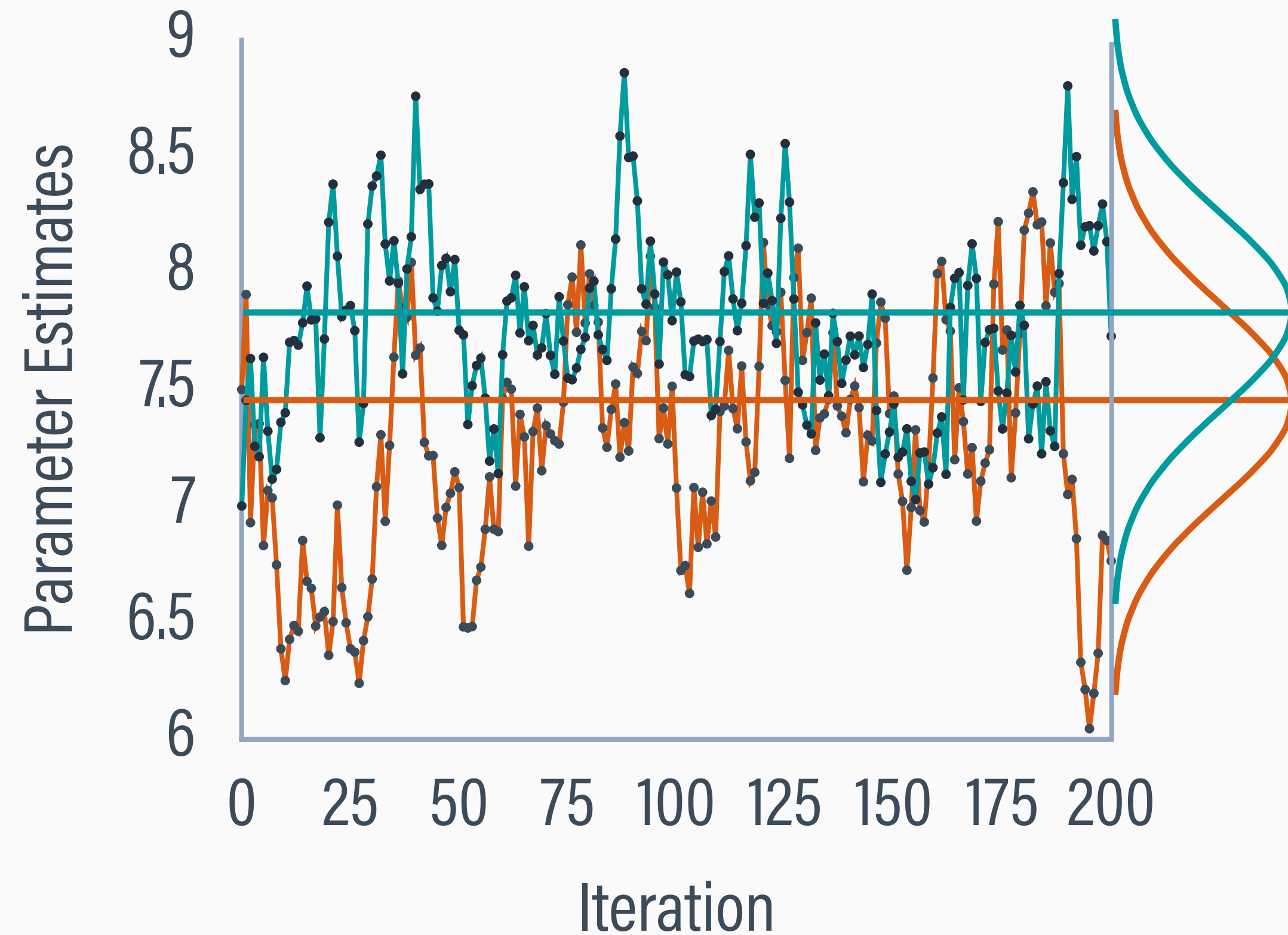
$$\text{PSRF} = \sqrt{\frac{\text{mean difference between chains} + \text{within-chain variation}}{\text{within-chain variation}}}$$



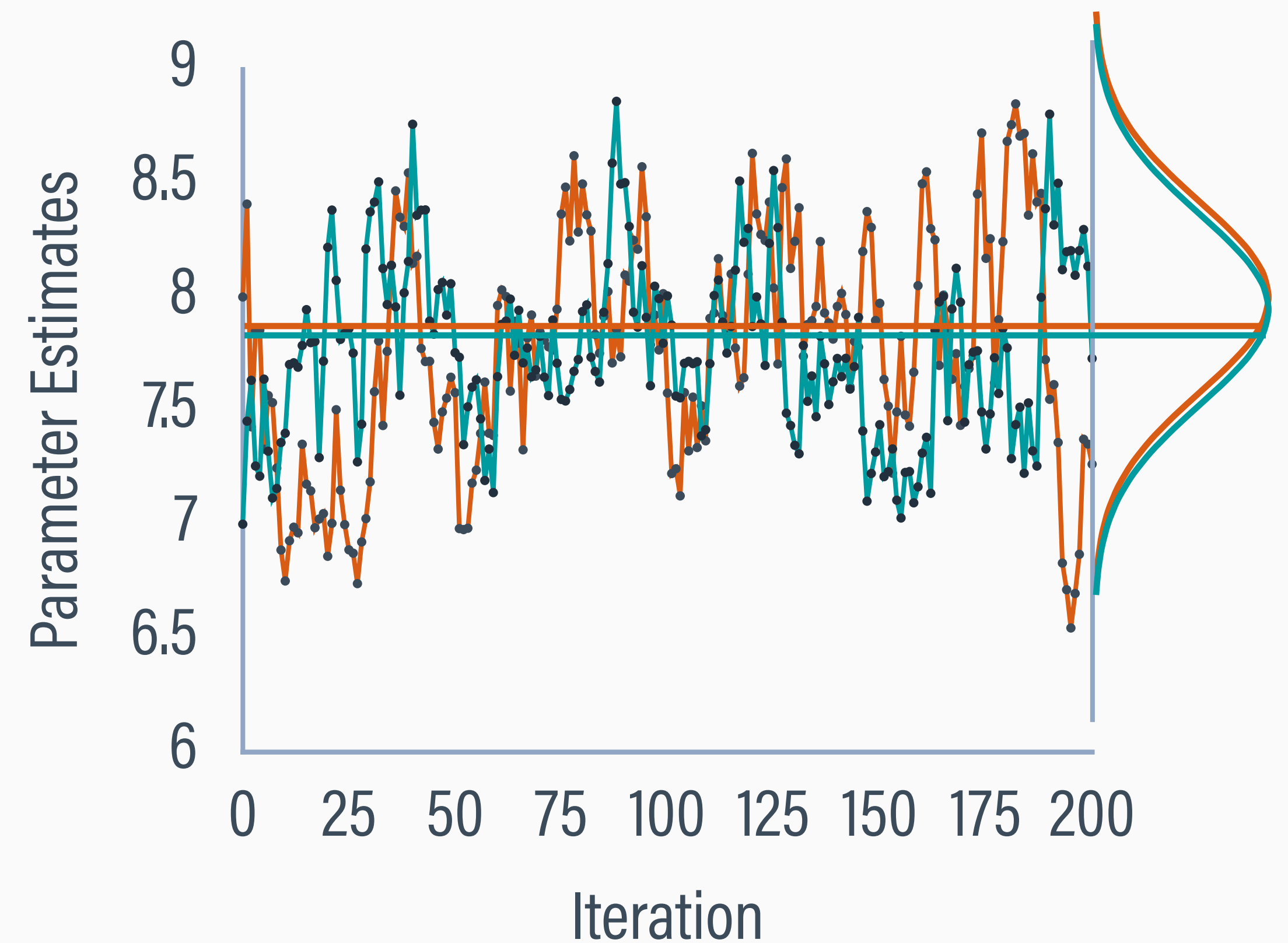


# CONVERGENCE

MCMC has not converged because between-chain mean difference is large ( $\text{PSR} > 1.05$ )



MCMC has converged because between-chain mean difference is very small ( $\text{PSR} < 1.05$ )



# PSR DIAGNOSTIC OUTPUT

The number of burn-in iterations is sufficient because the highest PSRF across all parameters is  $< 1.05$  at the end of the burn-in period

## BURN-IN POTENTIAL SCALE REDUCTION (PSR) OUTPUT:

NOTE: Split chain PSR is being used. This splits each chain's iterations to create twice as many chains.

Comparing iterations across 2 chains	Highest PSR	Parameter #
251 to 500	1.029	6
501 to 1000	1.012	1
751 to 1500	1.006	8
1001 to 2000	1.003	1
...	...	...
4001 to 8000	1.001	6
4251 to 8500	1.001	4
4501 to 9000	1.001	14
4751 to 9500	1.000	14
5001 to 10000	>> 1.001	6

# SUMMARY TABLE: VARIANCES AND COVARIANCES

## OUTCOME MODEL ESTIMATES:

Summaries based on 10000 iterations using 2 chains.  
NOTE: Estimate column based on posterior median.

Outcome Variable: read9

Parameters	Estimate	StdDev	2.5%	97.5%	ChiSq	PValue	N_Eff
Variances:							
Residual Var.	94.449	13.550	72.536	125.090	---	---	5358.708
Coefficients:							
Intercept	65.176	6.748	51.876	78.518	93.364	0.000	3601.650
read1	0.505	0.048	0.409	0.597	112.098	0.000	5608.777
lnnprob1	-0.404	0.098	-0.599	-0.209	16.811	0.000	2401.917
Standardized Coefficients:							
read1	0.680	0.045	0.577	0.753	226.950	0.000	4685.157
lnnprob1	-0.297	0.069	-0.426	-0.156	18.342	0.000	2399.433
Proportion Variance Explained							
by Coefficients	0.562	0.055	0.438	0.654	---	---	5115.914
by Residual Variation	0.438	0.055	0.346	0.562	---	---	5115.914

## SUMMARY TABLE: REGRESSION COEFFICIENTS

## OUTCOME MODEL ESTIMATES:

Summaries based on 10000 iterations using 2 chains.

NOTE: Estimate column based on posterior median.

Outcome Variable: **read9**

Parameters	Estimate	StdDev	2.5%	97.5%	ChiSq	PValue	N_Eff
<hr/>							
Variances:							
Residual Var.	94.449	13.550	72.536	125.090	---	---	5358.708
<hr/>							
Coefficients:							
Intercept	65.176	6.748	51.876	78.518	93.364	0.000	3601.650
read1	0.505	0.048	0.409	0.597	112.098	0.000	5608.777
lrrnprob1	-0.404	0.098	-0.599	-0.209	16.811	0.000	2401.917
<hr/>							
Standardized Coefficients:							
read1	0.680	0.045	0.577	0.753	226.950	0.000	4685.157
lrrnprob1	-0.297	0.069	-0.426	-0.156	18.342	0.000	2399.433
<hr/>							
Proportion Variance Explained							
by Coefficients	0.562	0.055	0.438	0.654	---	---	5115.914
by Residual Variation	0.438	0.055	0.346	0.562	---	---	5115.914
<hr/>							

## SUMMARY TABLE: STANDARDIZED COEFFICIENTS

## OUTCOME MODEL ESTIMATES:

Summaries based on 10000 iterations using 2 chains.

NOTE: Estimate column based on posterior median.

Outcome Variable: **read9**

Parameters	Estimate	StdDev	2.5%	97.5%	ChiSq	PValue	N_Eff
Variances:							
Residual Var.	94.449	13.550	72.536	125.090	---	---	5358.708
Coefficients:							
Intercept	65.176	6.748	51.876	78.518	93.364	0.000	3601.650
read1	0.505	0.048	0.409	0.597	112.098	0.000	5608.777
lnnprob1	-0.404	0.098	-0.599	-0.209	16.811	0.000	2401.917
Standardized Coefficients:							
read1	0.680	0.045	0.577	0.753	226.950	0.000	4685.157
lnnprob1	-0.297	0.069	-0.426	-0.156	18.342	0.000	2399.433
Proportion Variance Explained							
by Coefficients	0.562	0.055	0.438	0.654	---	---	5115.914
by Residual Variation	0.438	0.055	0.346	0.562	---	---	5115.914

## SUMMARY TABLE: R<sup>2</sup> EFFECT SIZES

## OUTCOME MODEL ESTIMATES:

Summaries based on 10000 iterations using 2 chains.

NOTE: Estimate column based on posterior median.

Outcome Variable: **read9**

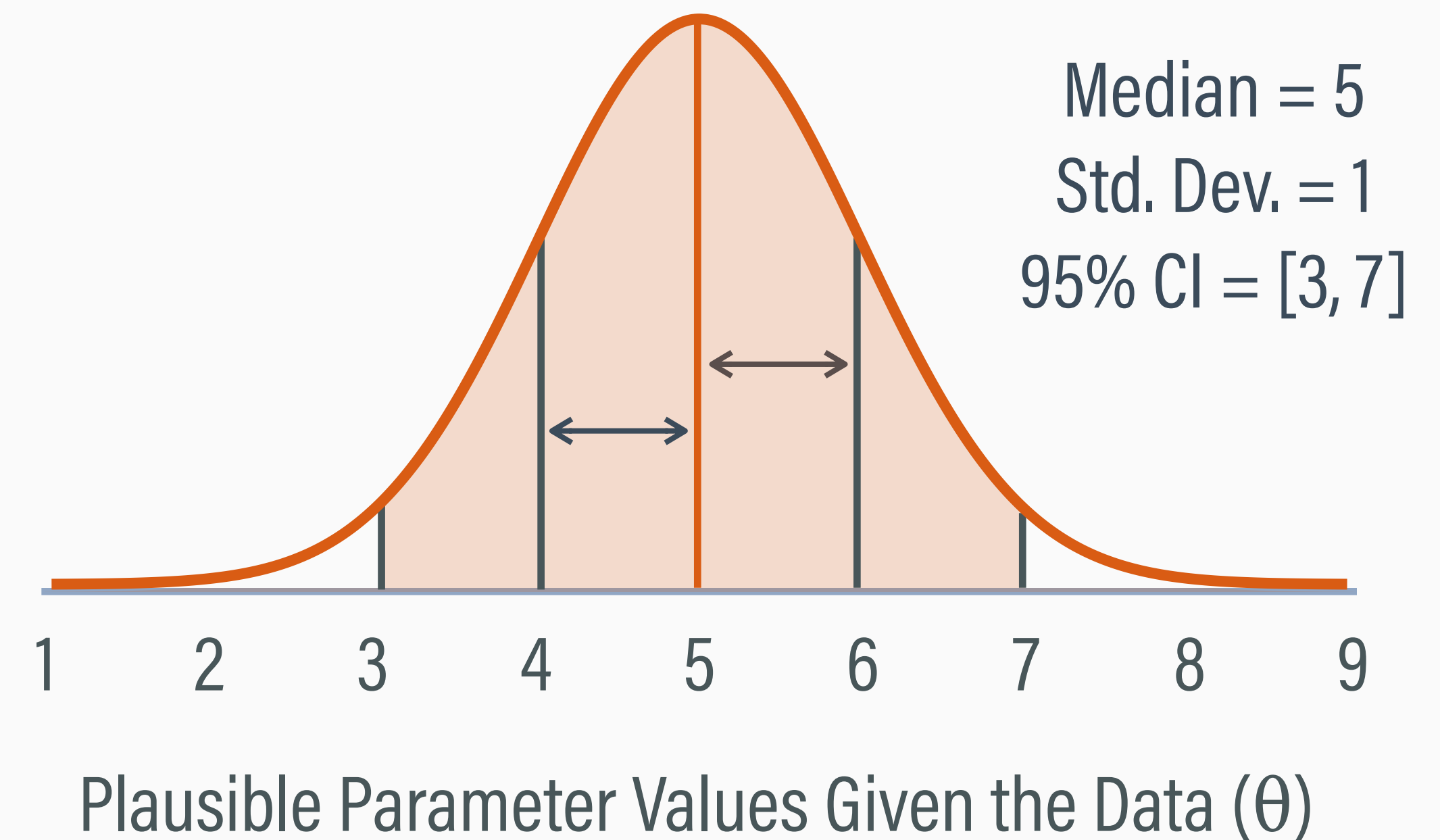
Parameters	Estimate	StdDev	2.5%	97.5%	ChiSq	PValue	N_Eff
-----							
Variances:							
Residual Var.	94.449	13.550	72.536	125.090	---	---	5358.708
Coefficients:							
Intercept	65.176	6.748	51.876	78.518	93.364	0.000	3601.650
read1	0.505	0.048	0.409	0.597	112.098	0.000	5608.777
lnnprob1	-0.404	0.098	-0.599	-0.209	16.811	0.000	2401.917
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-----							
Proportion Variance Explained							
by Coefficients	0.562	0.055	0.438	0.654	---	---	5115.914
by Residual Variation	0.438	0.055	0.346	0.562	---	---	5115.914



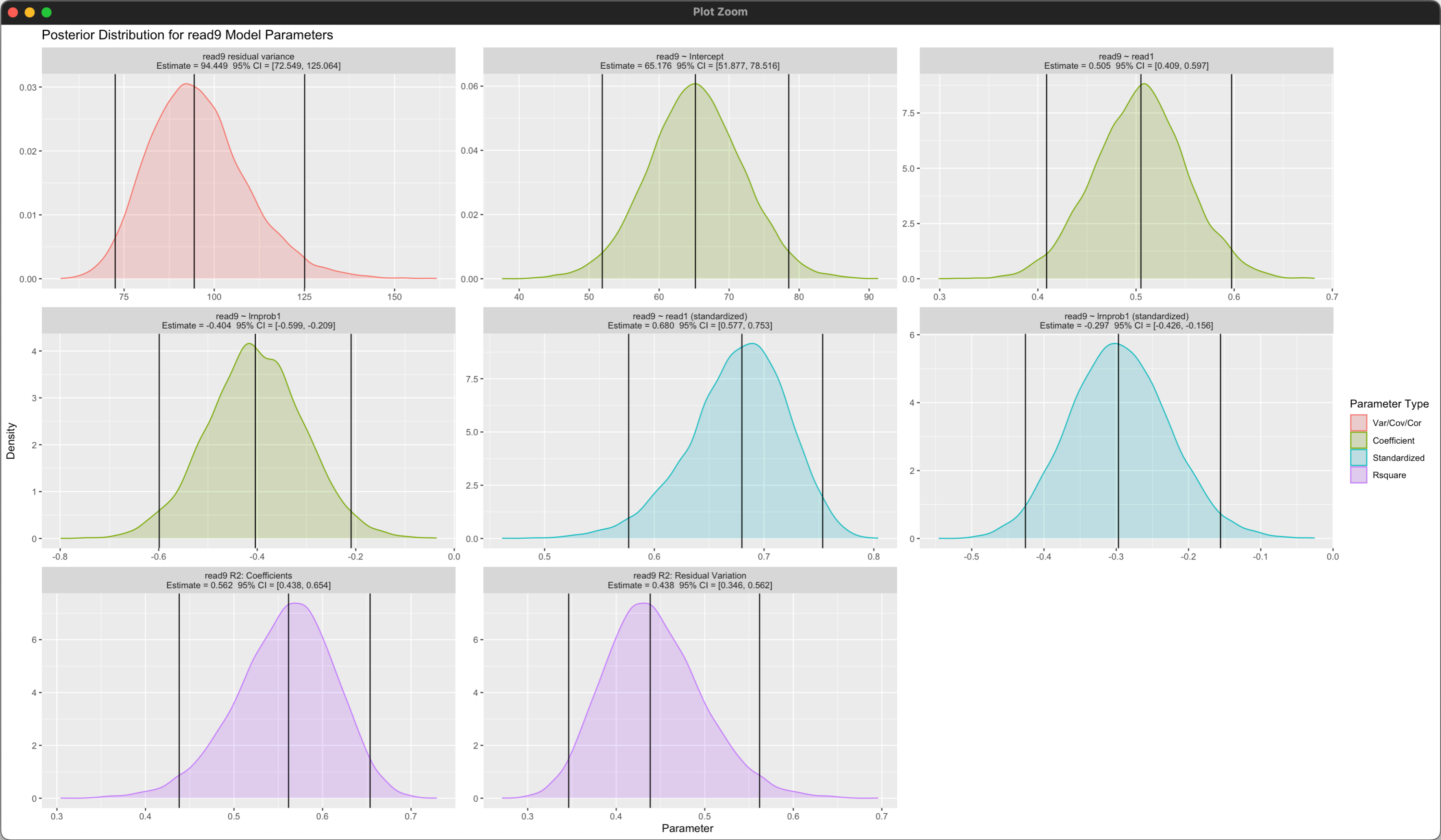
# DISTRIBUTION SUMMARIES

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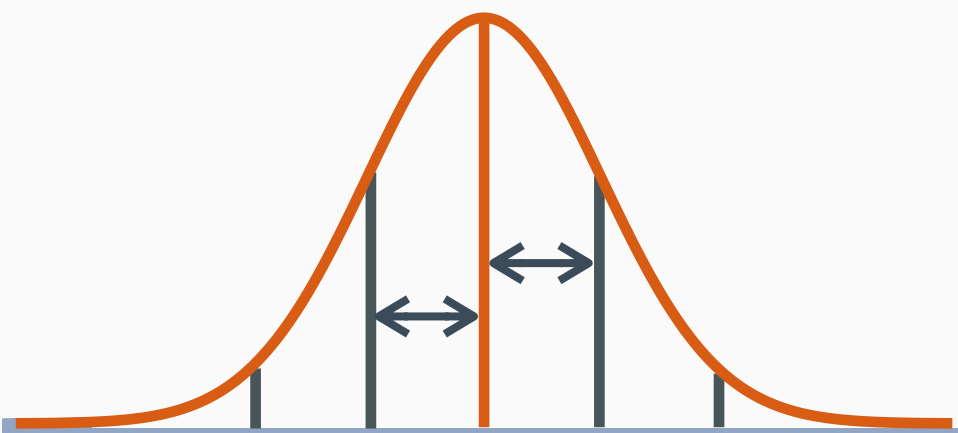
- ◉ The median (or mean) quantifies the most likely parameter value
- ◉ The standard deviation quantifies spread of the parameter's distribution
- ◉ 95% intervals define plausible parameter values that could have produced the data



# DISTRIBUTION PLOTS (RBLIMP ONLY)



# ESTIMATES AND “BAYESIAN STANDARD ERRORS”



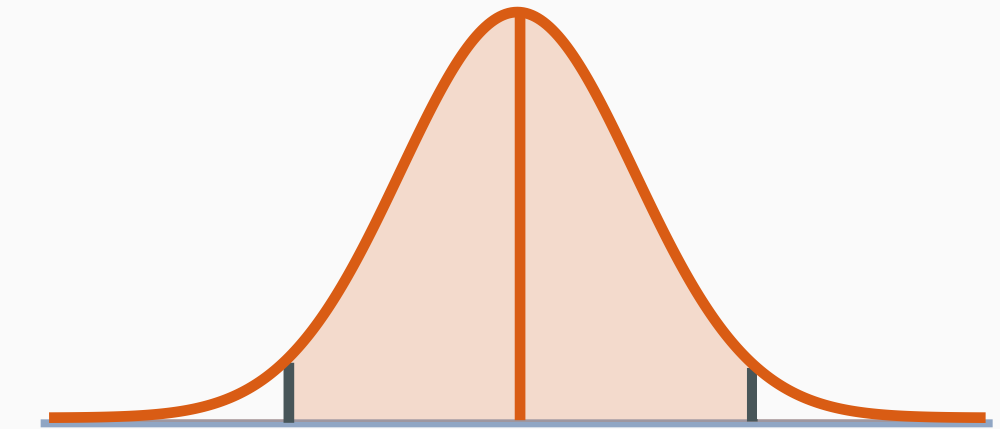
## OUTCOME MODEL ESTIMATES:

Summaries based on 10000 iterations using 2 chains.  
NOTE: Estimate column based on posterior median.

Outcome Variable: read9

Parameters	Estimate	StdDev	2.5%	97.5%	ChiSq	PValue	N_Eff
Variances:							
Residual Var.	94.449	13.550	72.536	125.090	---	---	5358.708
Coefficients:							
Intercept	65.176	6.748	51.876	78.518	93.364	0.000	3601.650
read1	0.505	0.048	0.409	0.597	112.098	0.000	5608.777
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Proportion Variance Explained							
by Coefficients	0.562	0.055	0.438	0.654	---	---	5115.914
by Residual Variation	0.438	0.055	0.346	0.562	---	---	5115.914

## 95% CREDIBLE INTERVALS



## OUTCOME MODEL ESTIMATES:

Summaries based on 10000 iterations using 2 chains.

NOTE: Estimate column based on posterior median.

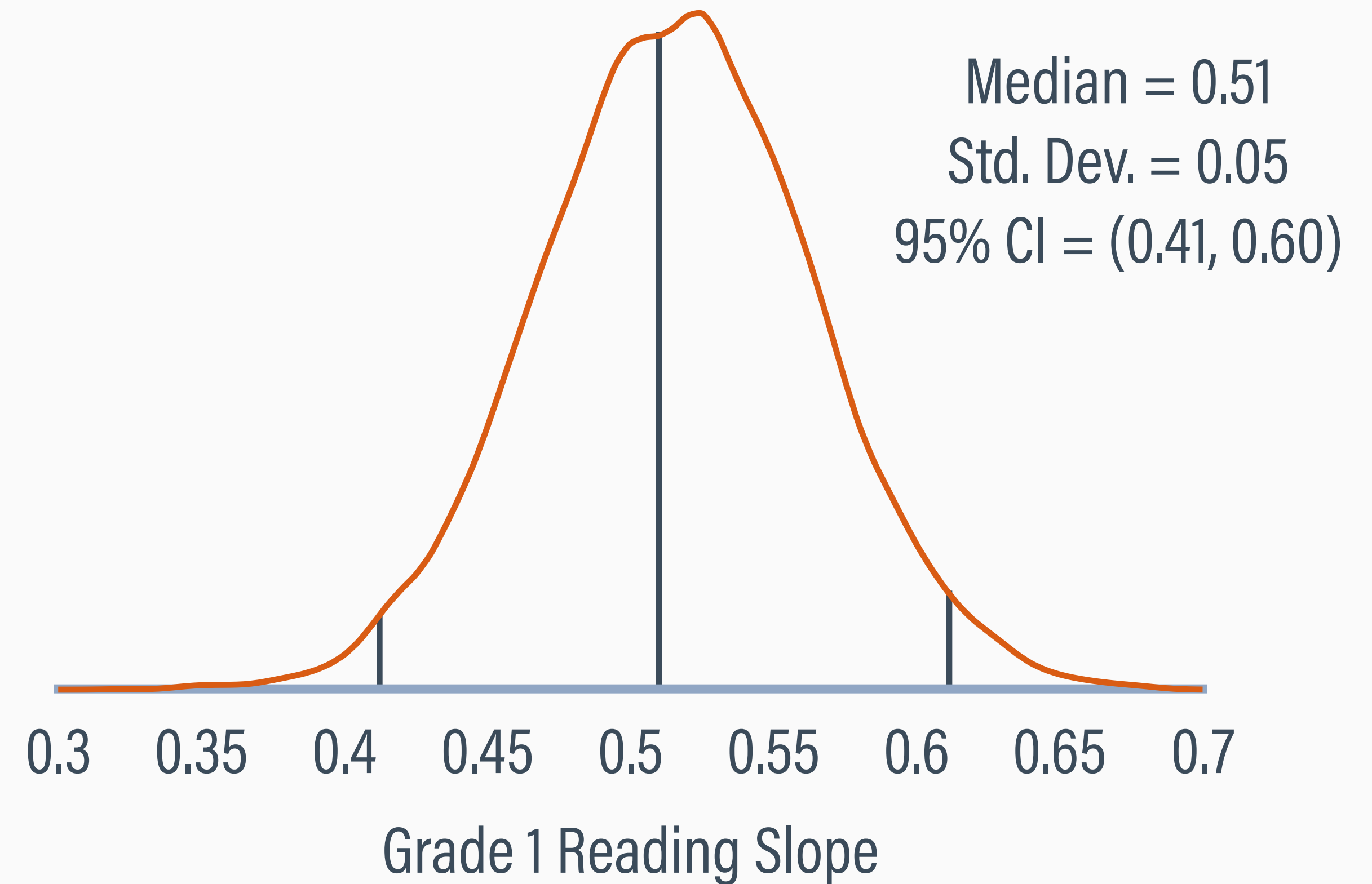
Outcome Variable: **read9**

Parameters	Estimate	StdDev	2.5%	97.5%	ChiSq	PValue	N_Eff
<hr/>							
Variances:							
Residual Var.	94.449	13.550	72.536	125.090	---	---	5358.708
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by Residual Variation	0.438	0.055	0.346	0.562	---	---	5115.914

# INTERPRETATIONS

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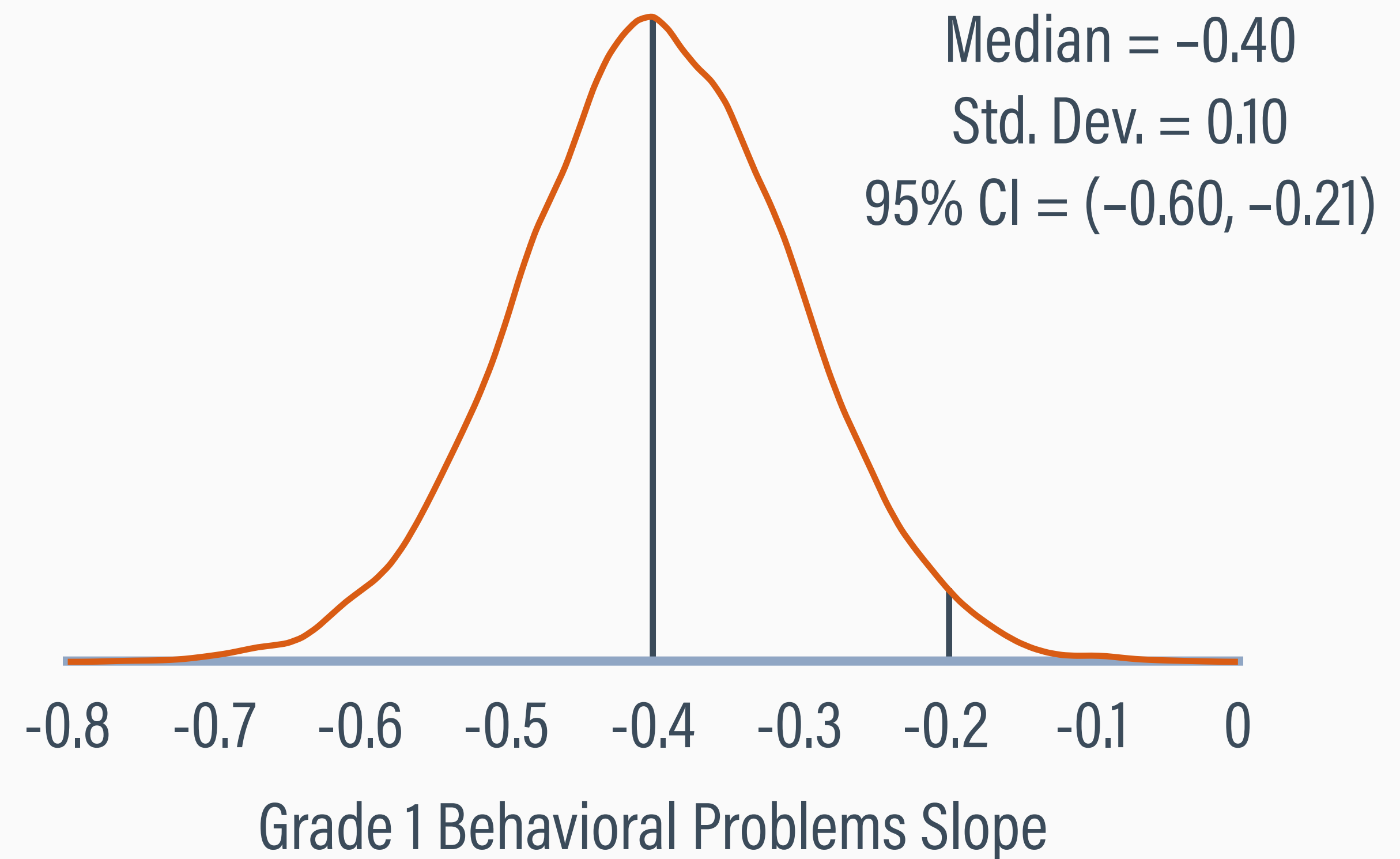
- For two students with same first grade learning problems rating, scoring one point higher on the first grade reading test predicts a 0.51 increase in grade 9 reading
- The parameter's standard deviation is 0.05
- The range from 0.41 to 0.60 captures 95% of the plausible parameter values that could have produced these data



# INTERPRETATIONS, CONTINUED

---

- For two students with same first grade reading, scoring one point higher on the first grade learning problems measure predicts a  $-0.44$  decrease in grade 9 reading
- The parameter's standard deviation is 0.10
- The range from  $-0.60$  to  $-0.21$  captures 95% of the plausible parameter values that could have produced these data



# COMPUTATIONAL FREQUENTISM

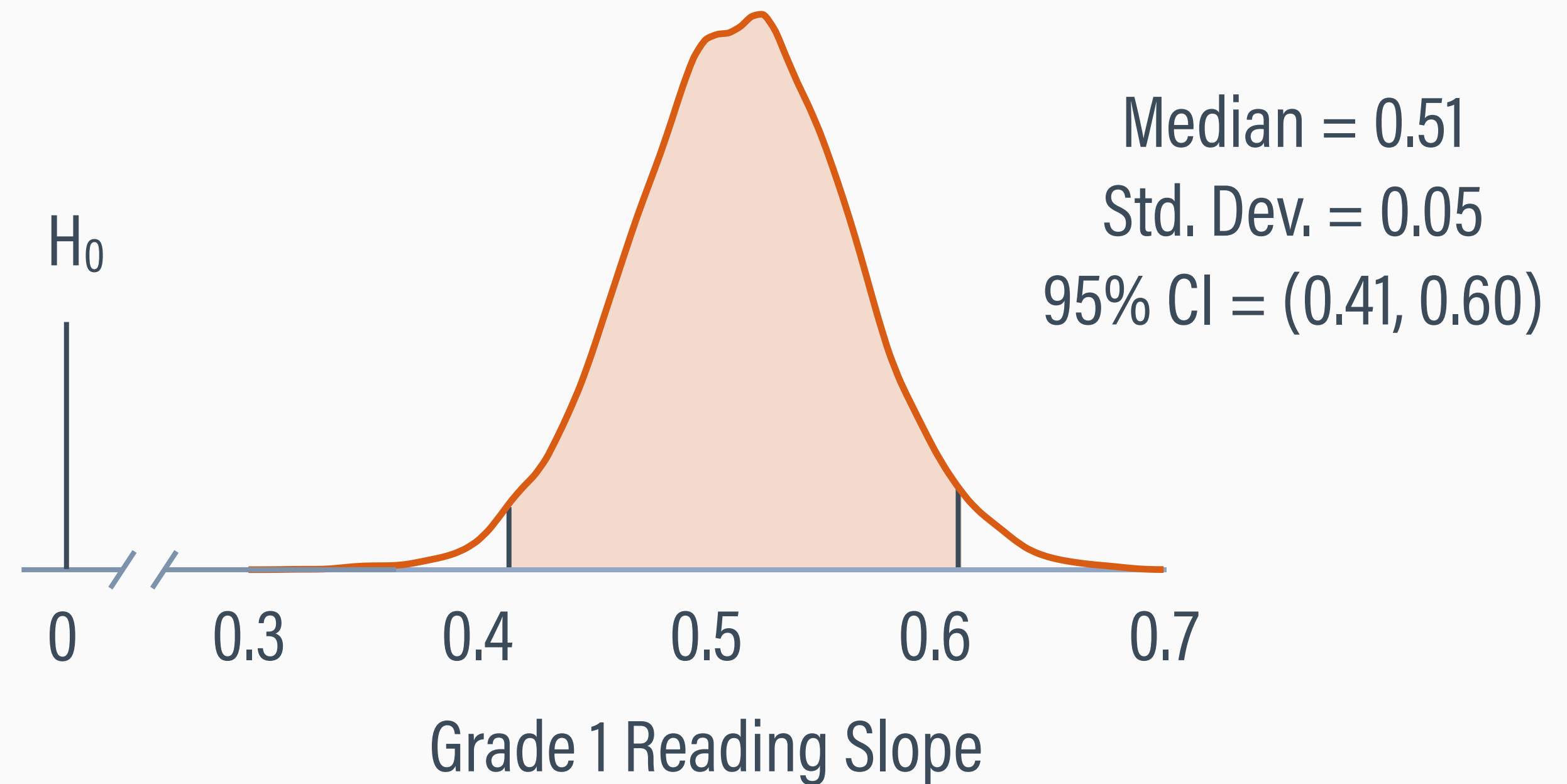
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- Many common missing data problems are challenging or impossible with maximum likelihood
- A computational frequentist perspective views MCMC parameter summaries as replacements for unobtainable maximum likelihood estimates
- Use MCMC-generated quantities for frequentist inference



# SIGNIFICANCE TESTING VIA INTERVALS

- If the 95% credible interval does not include zero, so then we refute the null hypothesis ( $p < .05$ )
- A population slope equal to zero is unlikely to have produced these data



# FREQUENTIST WALD TEST

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- ◉ The Wald chi-square is an alternate test statistic that equals the square of the z-statistic (or t-test)

$$\chi_{\text{ML}}^2 = \frac{(\hat{\theta} - \theta_0)^2}{\text{SE}^2} = \frac{(\text{estimate} - \text{null})^2}{(\text{standard error})^2} = z^2$$

- ◉ Multivariate versions of the Wald test can evaluate multiple parameters simultaneously

# MCMC WALD TEST

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- The Bayesian or MCMC Wald test statistic (Asparouhov & Muthén, 2021) replaces the point estimate and standard error with the posterior mean and standard deviation

$$\chi^2_{\text{MCMC}} = \frac{(\theta - \theta_0)^2}{\text{SD}^2} = \frac{(\text{posterior mean} - \text{null})^2}{(\text{posterior standard deviation})^2} = z^2$$

- MCMC-generated test statistic and p-value for frequentist inference (computational frequentism)

# WALD CHI-SQUARE TESTS

OUTCOME MODEL ESTIMATES:

Summaries based on 10000 iterations using 2 chains.  
NOTE: Estimate column based on posterior median.

Outcome Variable: read9

Parameters	Estimate	StdDev	2.5%	97.5%	ChiSq	PValue	N_Eff
Variances:							
Residual Var.	94.449	13.550	72.536	125.090	---	---	5358.708
Coefficients:							
Intercept	65.176	6.748	51.876	78.518	93.364	0.000	3601.650
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lnnprob1	-0.404	0.098	-0.599	-0.209	16.811	0.000	2401.917
Standardized Coefficients:							
read1	0.680	0.045	0.577	0.753	226.950	0.000	4685.157
lnnprob1	-0.297	0.069	-0.426	-0.156	18.342	0.000	2399.433
Proportion Variance Explained							
by Coefficients	0.562	0.055	0.438	0.654	---	---	5115.914
by Residual Variation	0.438	0.055	0.346	0.562	---	---	5115.914

# ESTIMATOR COMPARISON

The two estimators are effectively numerically equivalent!!!

Parameter	MCMC						FIML					
	Median	SD	2.5%	97.5%	Chi-Sq.	p	Est.	SE	2.5%	97.5%	z	p
Intercept	65.18	6.75	51.88	78.52	93.36	< .001	65.13	6.50	52.39	77.87	10.02	< .001
1st Grade Reading	0.51	0.05	0.41	0.60	112.10	< .001	0.51	0.05	0.42	0.59	11.07	< .001
Learning Problems	−0.40	0.10	−0.60	−0.21	16.81	< .001	−0.40	0.10	−0.59	−0.22	−4.24	< .001
Residual variance	94.45	13.55	72.54	125.09	—	—	89.35	12.33	65.19	113.51	—	—
R <sup>2</sup>	.56	.06	.44	.65	—	—	0.57	0.06	—	—	—	—

## QUALITY CONTROL CHECK: EFFECTIVE SAMPLE SIZE

---

- The effective sample size ( $N_{\text{Eff}}$ ) diagnostic quantifies the number of independent MCMC estimates contributing to the parameter summaries after removing autocorrelation
- An acceptable value ( $N_{\text{Eff}} > 100$ ) implies that the number of iterations after the burn-in period is sufficient, whereas low values suggest to increase the number of iterations
- Low values often indicate that the data lack support for certain model parameters (e.g., due to overfitting)

# DIAGNOSTIC OUTPUT

All N\_Eff values > 100, the number of iterations for the summary is sufficient!

## OUTCOME MODEL ESTIMATES:

Summaries based on 10000 iterations using 2 chains.

NOTE: Estimate column based on posterior median.

Outcome Variable: **read9**

Parameters	Estimate	StdDev	2.5%	97.5%	ChiSq	PValue	N_Eff
<hr/>							
Variances:							
Residual Var.	94.449	13.550	72.536	125.090	---	---	5358.708
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lnnprob1	-0.404	0.098	-0.599	-0.209	16.811	0.000	2401.917
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Proportion Variance Explained							
by Coefficients	0.562	0.055	0.438	0.654	---	---	5115.914
by Residual Variation	0.438	0.055	0.346	0.562	---	---	5115.914



# OUTLINE

1

Modern Missing Data Methods

2

Missing Data Mechanisms

3

MCMC Estimation and Missing Data Imputation

4

Fitting Regression Models in Blimp

5

Incomplete Categorical Variables

6

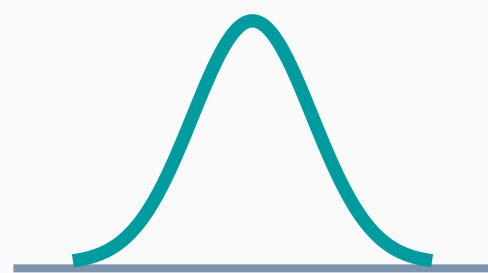
Interaction Effects

7

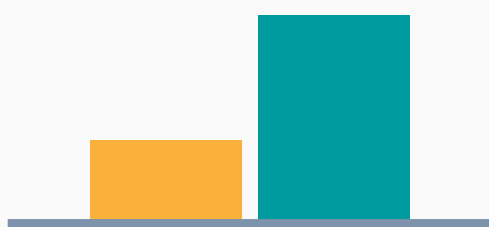
Reporting Results From a Missing Data Analysis

# BLIMP VARIABLE TYPES

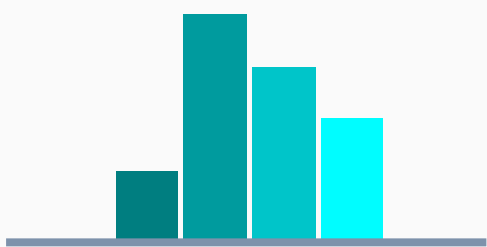
## Exogenous Predictors



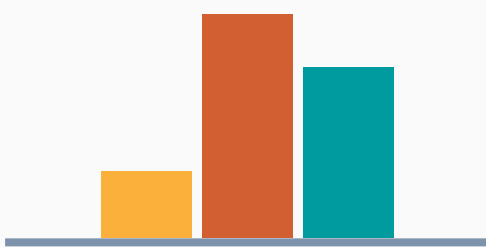
Normal  
(Manifest)



Binary

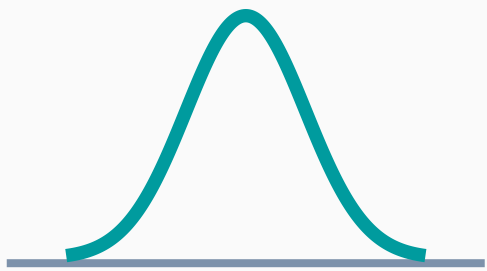


Ordinal

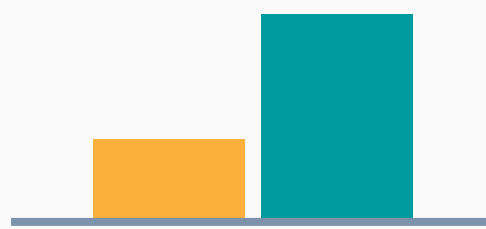


Nominal

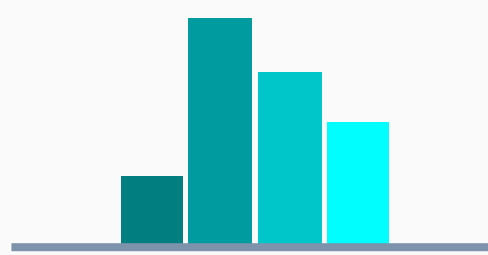
## Univariate Outcomes



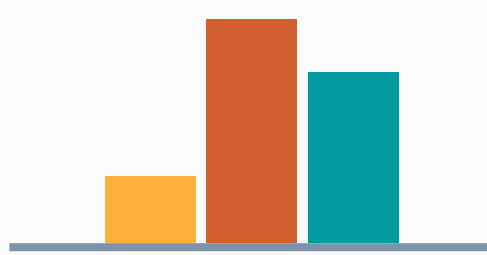
Normal  
(Manifest or Latent)



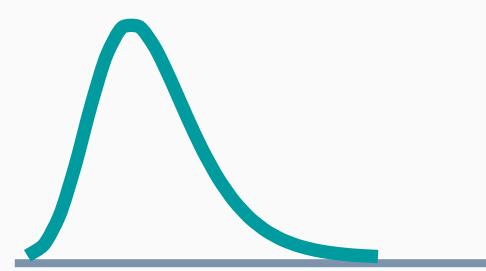
Binary



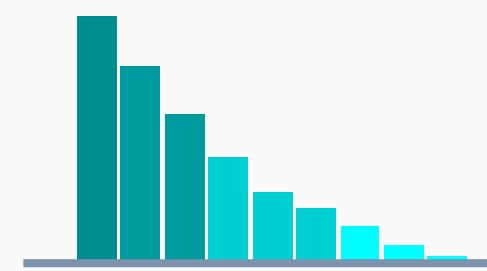
Ordinal



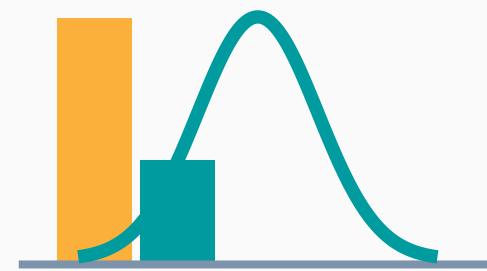
Nominal



Skewed  
(Manifest or Latent)

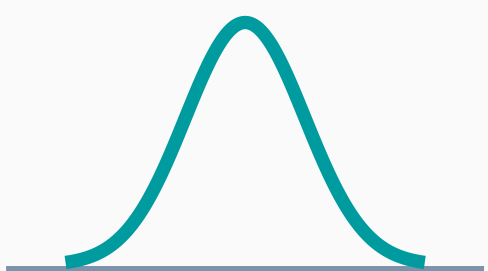


Count



Two-Part  
(Floor Effects)

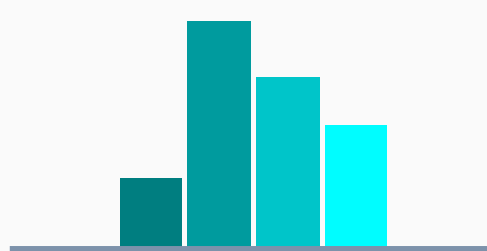
## Multivariate Outcomes



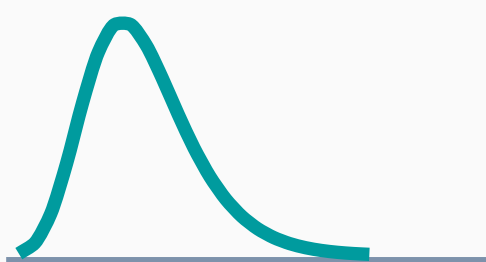
Normal  
(Manifest or Latent)



Binary



Ordinal

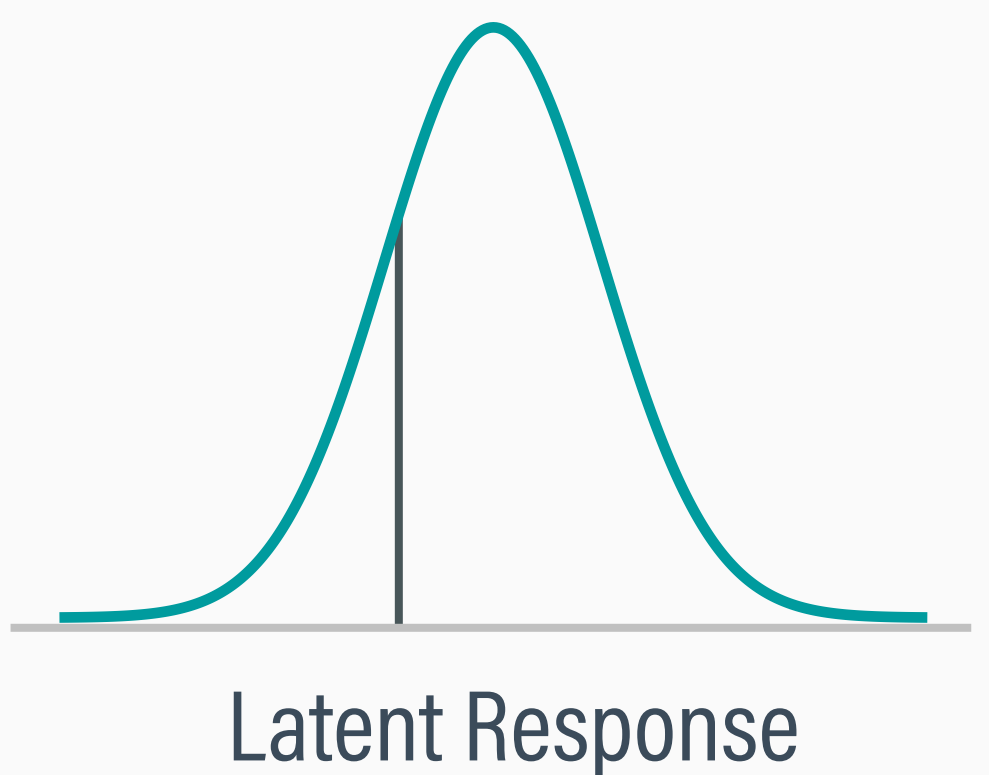


Skewed  
(Manifest or Latent)

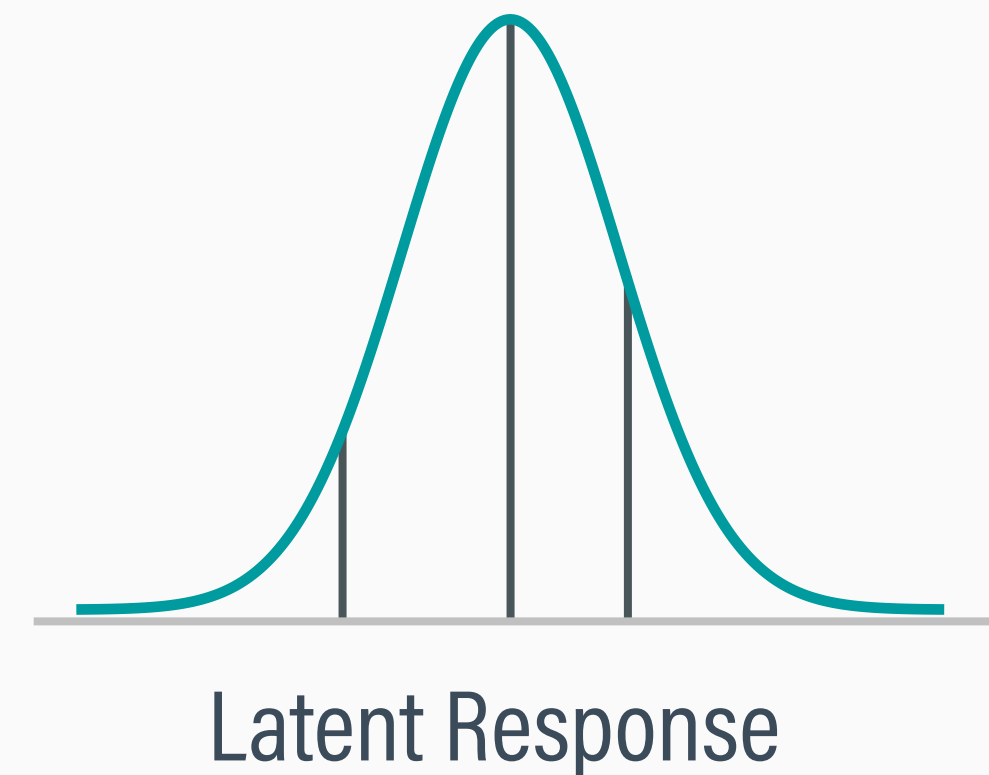
# LATENT RESPONSE FORMULATION

---

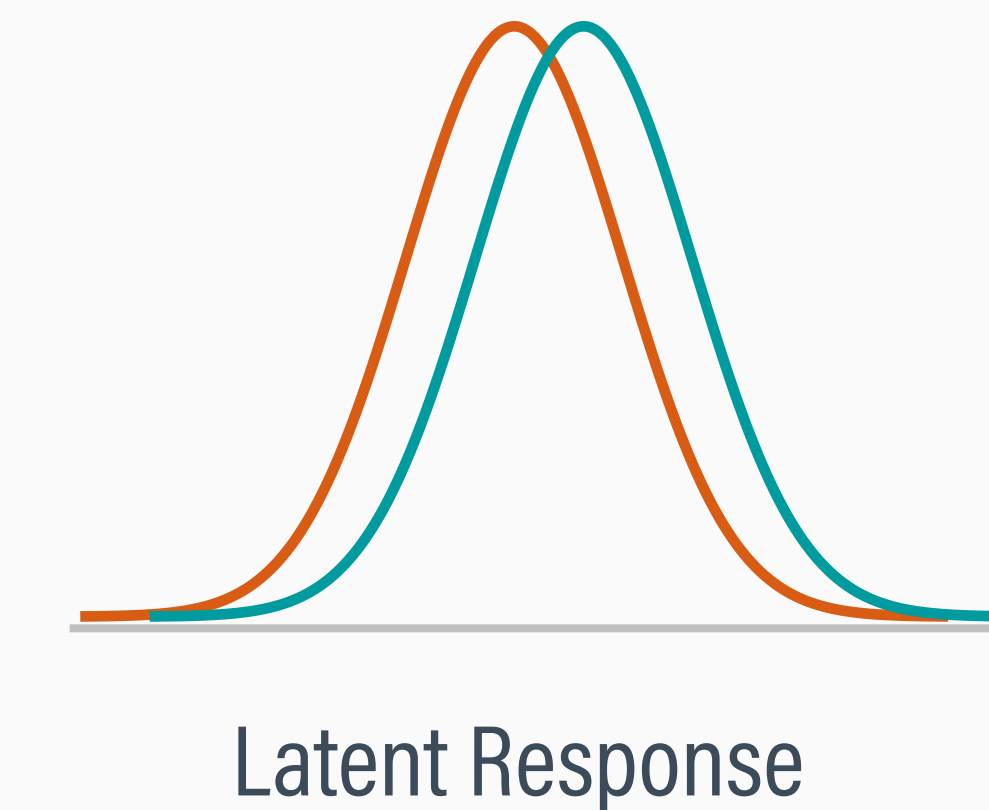
Binary



Ordinal



Multicategorical



# READING DATA

---

- In first grade, students are classified according to their risk of developing emotional or behavioral problems
- 65% are classified as being at medium or high risk

Variable	Definition	Missing %	Scale
atrisk	Emotional/behavioral risk code	2.2	0 = Low, 1 = Medium/high
lrnprob1	1st grade learning problems	2.2	Numeric (31 to 88)
read1	1st grade broad reading composite	6.5	Numeric (39 to 153)
read9	9th grade broad reading composite	17.4	Numeric (41 to 123)

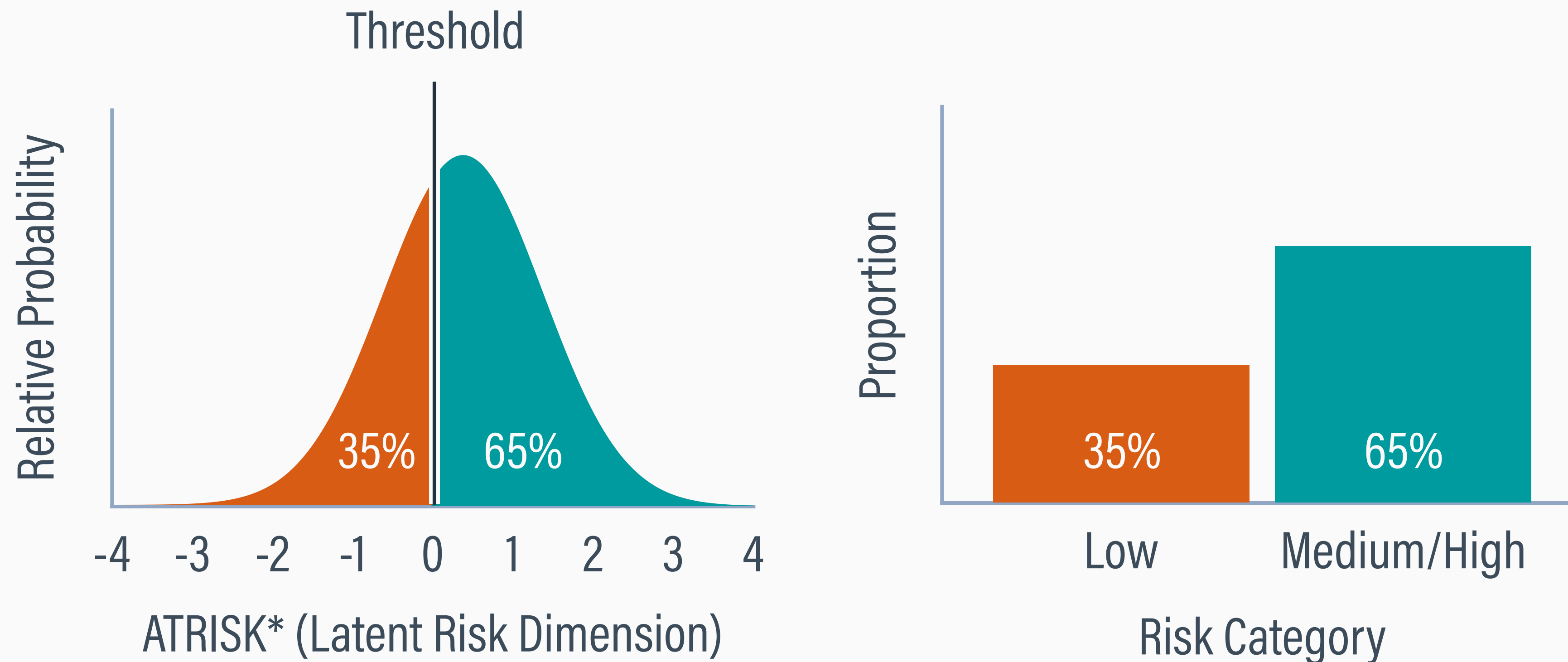
# INCOMPLETE BINARY VARIABLES

---

- ◉ Probit regression envisions binary and ordinal variables arising from an underlying normal latent response variable
- ◉ Applied to the at risk indicator, the latent variable represents an unobserved, continuous propensity for emotional disorders
- ◉ A threshold carves the latent distribution into segments

# LATENT AND DISCRETE DISTRIBUTIONS

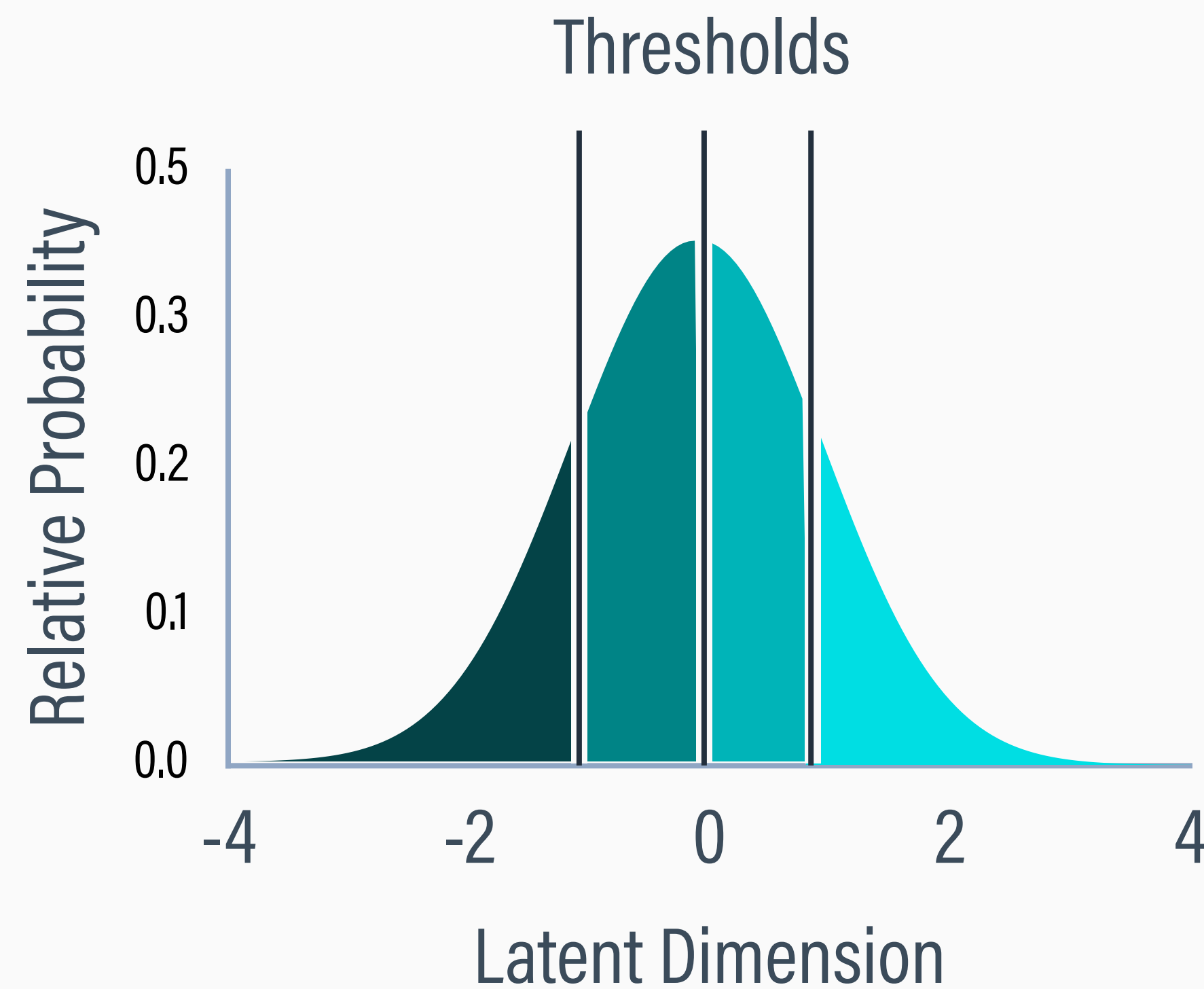
- The threshold parameter divides the latent distribution into segments, with areas under the curve matching the bar plot



# ORDINAL VARIABLES

---

- Multiple threshold parameters divide the latent distribution into segments, with areas under the curve matching the bar plot





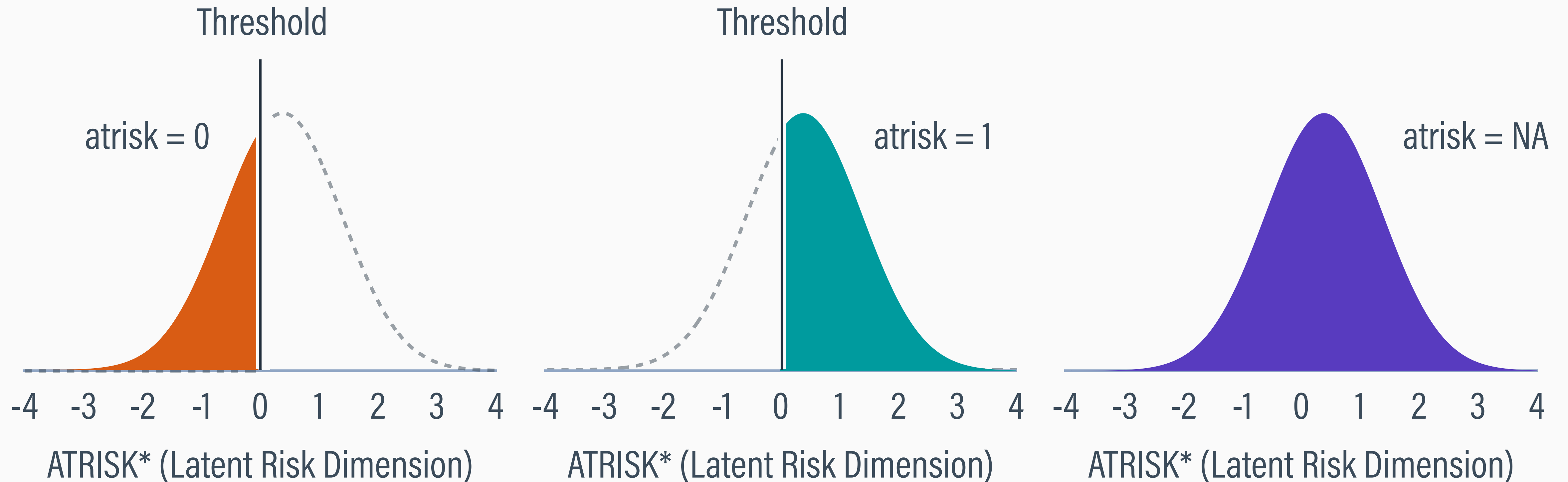
# IMPUTING LATENT RESPONSE SCORES

---

- Latent response scores are missing data to be imputed
- MCMC uses computer simulation to “sample” latent response scores distributions, just like any other incomplete variable
- Blimp uses the latent response scores to link the categorical predictor to other continuous predictors, but the binary dummy code is the regressor in the focal model

# LATENT AND DISCRETE DISTRIBUTIONS

- Latent imputations must fall below or above threshold if the binary variable is observed, and they are unconstrained if missing



# ANALYSIS 2 REGRESSION MODEL

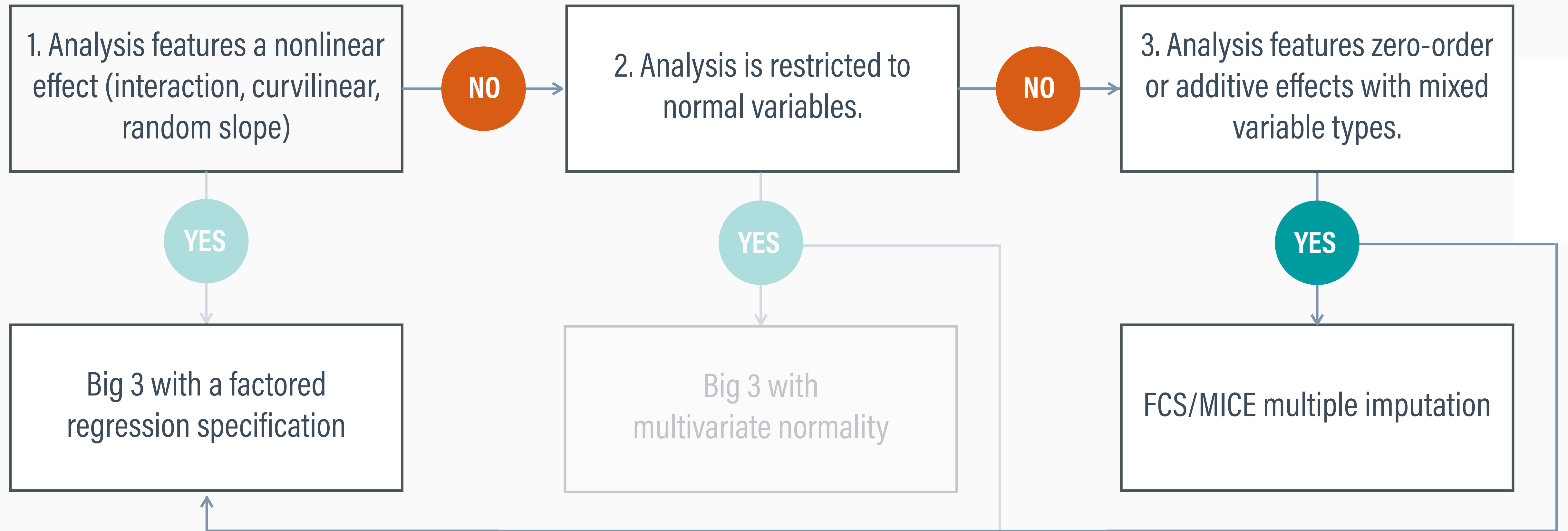
- Academic variables from 1st grade predicting 9th grade reading achievement in middle school

$$\text{read}_9 = \beta_0 + \beta_1(\text{read}_1) + \beta_2(\text{lrnprob}_1) + \beta_3(\text{atrisk}) + \varepsilon$$

Variable	Definition	Missing %	Scale
atrisk	Emotional/behavioral risk code	2.2	0 = Low, 1 = Medium/high
lrnprob1	1st grade learning problems	2.2	Numeric (31 to 88)
read1	1st grade broad reading composite	6.5	Numeric (39 to 153)
read9	9th grade broad reading composite	17.4	Numeric (41 to 123)

# MISSING DATA DECISION TREE

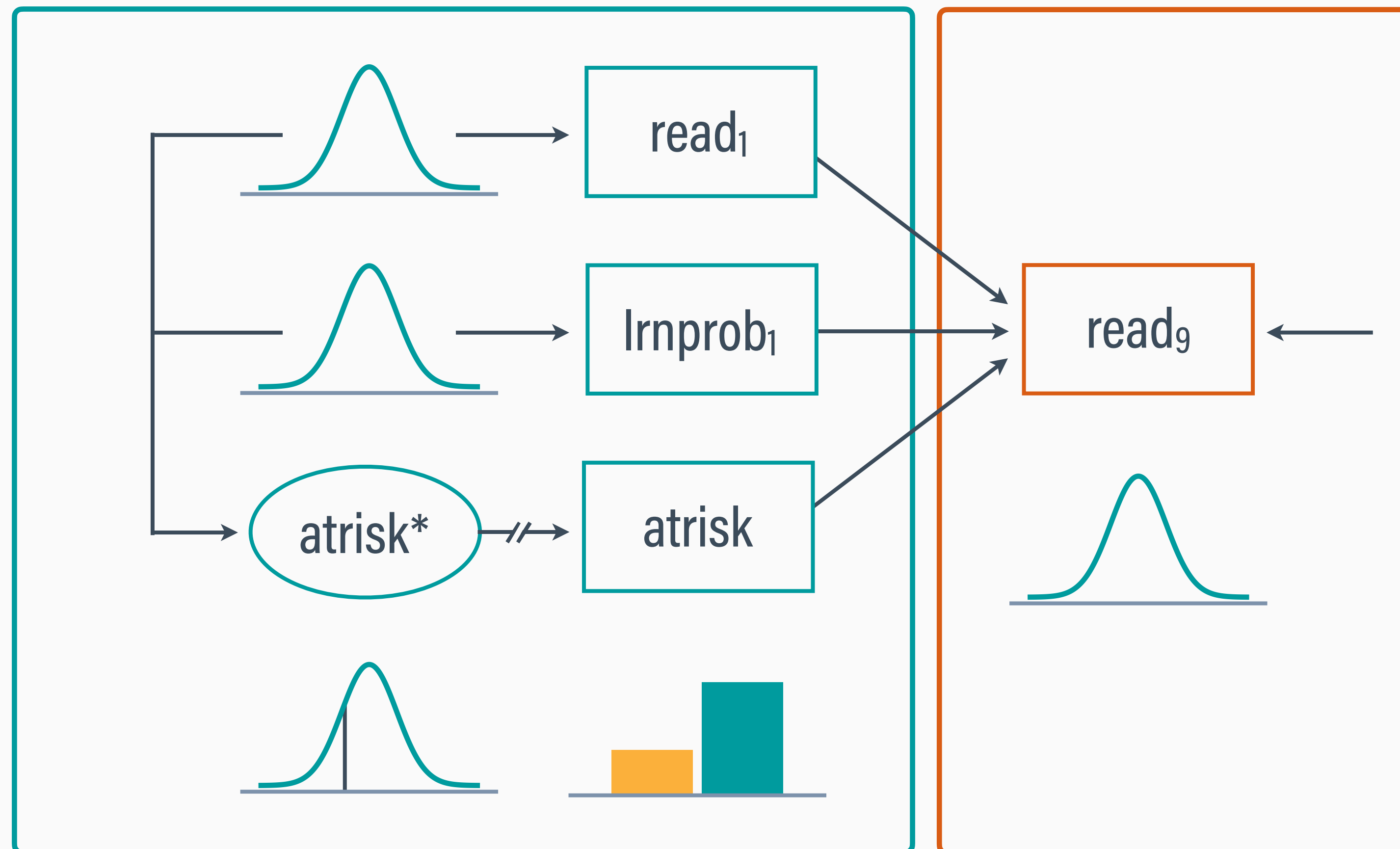
Mixtures of numeric and categorical variables reduce the number of viable options.



# TWO-PART FACTORED SPECIFICATION

Incomplete Predictor Model

Outcome Model



# BLIMP STUDIO SCRIPT 2

---

**DATA:** reading.dat;

**VARIABLES:** id male hispanic riskgrp atrisk behsymp1 lnprob1 read1 read2 read3 read9  
read9grp stanread7 math1 math2 math3 math9 math9grp stanmath7;

**MISSING:** 999;

**NOMINAL:** atrisk;      # automatic dummy coding with lowest score as the reference

**MODEL:** read9 ~ read1 lnprob1 atrisk;

**BURN:** 10000;

**ITER:** 10000;

**SEED:** 90291;

# RBLIMP SCRIPT (MODEL 2)

---

```
# fit model
```

```
model2 <- rblimp(
```

```
  data = reading,
```

```
  nominal = 'atrisk',      # automatic dummy coding with lowest score as the reference
```

```
  model = 'read9 ~ read1 lnp1 atrisk',
```

```
  seed = 90291,
```

```
  burn = 10000,
```

```
  iter = 10000)
```

```
# summarize results
```

```
output(model2)
```

```
posterior_plot(model2, 'read9')
```



# PSR DIAGNOSTIC OUTPUT

The number of burn-in iterations is sufficient because the highest PSRF across all parameters is  $< 1.05$  at the end of the burn-in period

## BURN-IN POTENTIAL SCALE REDUCTION (PSR) OUTPUT:

NOTE: Split chain PSR is being used. This splits each chain's iterations to create twice as many chains.

Comparing iterations across 2 chains	Highest PSR	Parameter #
251 to 500	1.034	21
501 to 1000	1.010	11
751 to 1500	1.010	9
1001 to 2000	1.006	13
... ..	...	...
4001 to 8000	1.001	18
4251 to 8500	1.001	21
4501 to 9000	1.001	8
4751 to 9500	1.001	8
5001 to 10000	>> 1.001	8

# MISSING DATA INFORMATION

- Per-variable missingness rates appear in the rows, and missing data patterns appear in the columns (M = missing)

DATA INFORMATION:

Sample Size: 138  
Nominal Dummy Codes:

atrisk = atrisk.1

Missing Data Info:

	miss %	1	2	3	4	5	6
-----							
read9 =	17.4	-	M	-	-	-	M
atrisk =	2.2	-	-	-	M	-	-
lnnprob1 =	2.2	-	-	-	-	M	-
read1 =	6.5	-	-	M	-	-	M
-----							
	%	72.5	16.7	5.8	2.2	2.2	0.7

# REGRESSION SUMMARY TABLE

## OUTCOME MODEL ESTIMATES:

Summaries based on 10000 iterations using 2 chains.

NOTE: Estimate column based on posterior median.

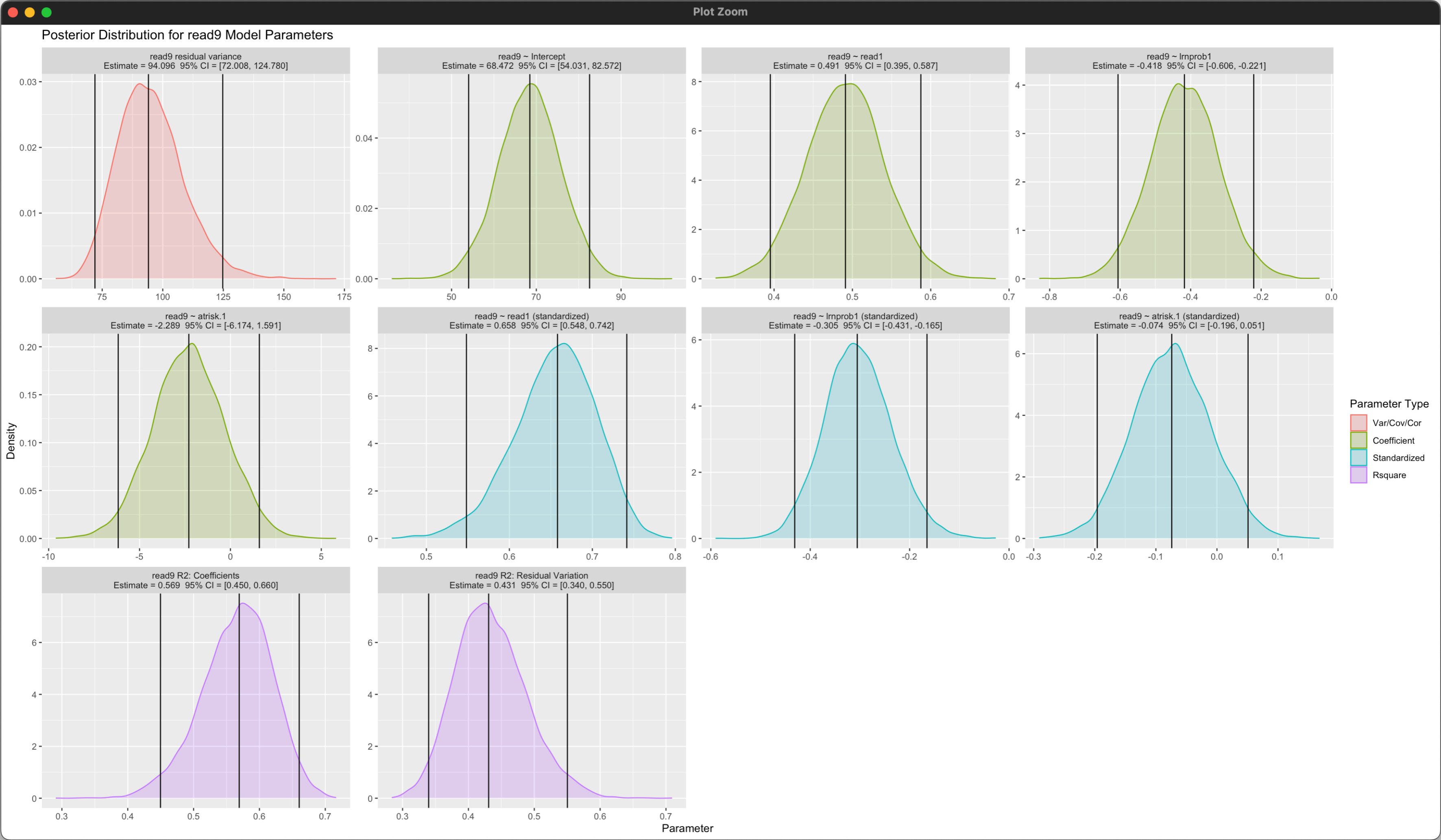
Outcome Variable: **read9**

Parameters	Estimate	StdDev	2.5%	97.5%	ChiSq	PValue	N_Eff
-----							
Variances:							
Residual Var.	94.096	13.521	71.998	124.781	---	---	5093.390
Coefficients:							
Intercept	68.472	7.253	54.030	82.578	89.064	0.000	3563.037
read1	0.491	0.049	0.395	0.587	99.355	0.000	6011.826
lnnprob1	-0.418	0.098	-0.606	-0.221	18.237	0.000	2740.547
atrisk.1	-2.289	1.990	-6.180	1.591	1.320	0.251	5821.121

• • •

Proportion Variance Explained							
by Coefficients	0.569	0.054	0.450	0.660	---	---	5125.545
by Residual Variation	0.431	0.054	0.340	0.550	---	---	5125.545

# DISTRIBUTION PLOTS (RBLIMP ONLY)



# INTERPRETATIONS

---

- Blimp automatically dummy codes nominal variables, treating the lowest numeric code (low risk) as the reference
- For two students with same first grade reading and learning problems scores, being classified as at risk is associated with a -2.29 point decrease in grade 9 reading (parameter SD = 1.99)
- Slopes for numeric predictors are similar to the first analysis

# ANALYSIS 3 REGRESSION MODEL

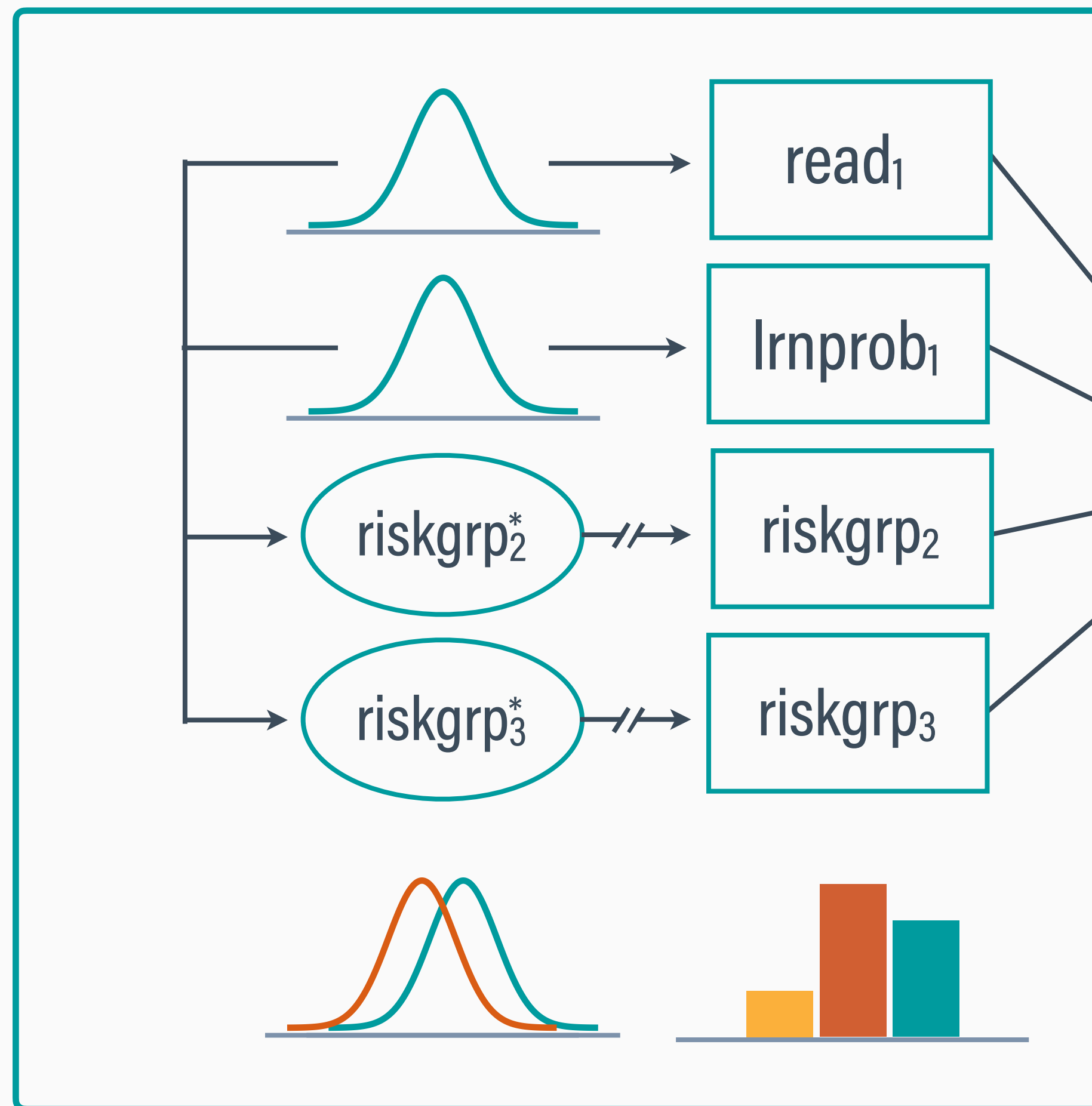
- Risk status is represented by two dummy codes with the low-risk group (riskgrp = 1) serving as the reference group

$$\text{read}_9 = \beta_0 + \beta_1(\text{read}_1) + \beta_2(\text{lrnprob}_1) + \beta_3(\text{riskgrp}_2) + \beta_4(\text{riskgrp}_3) + \varepsilon$$

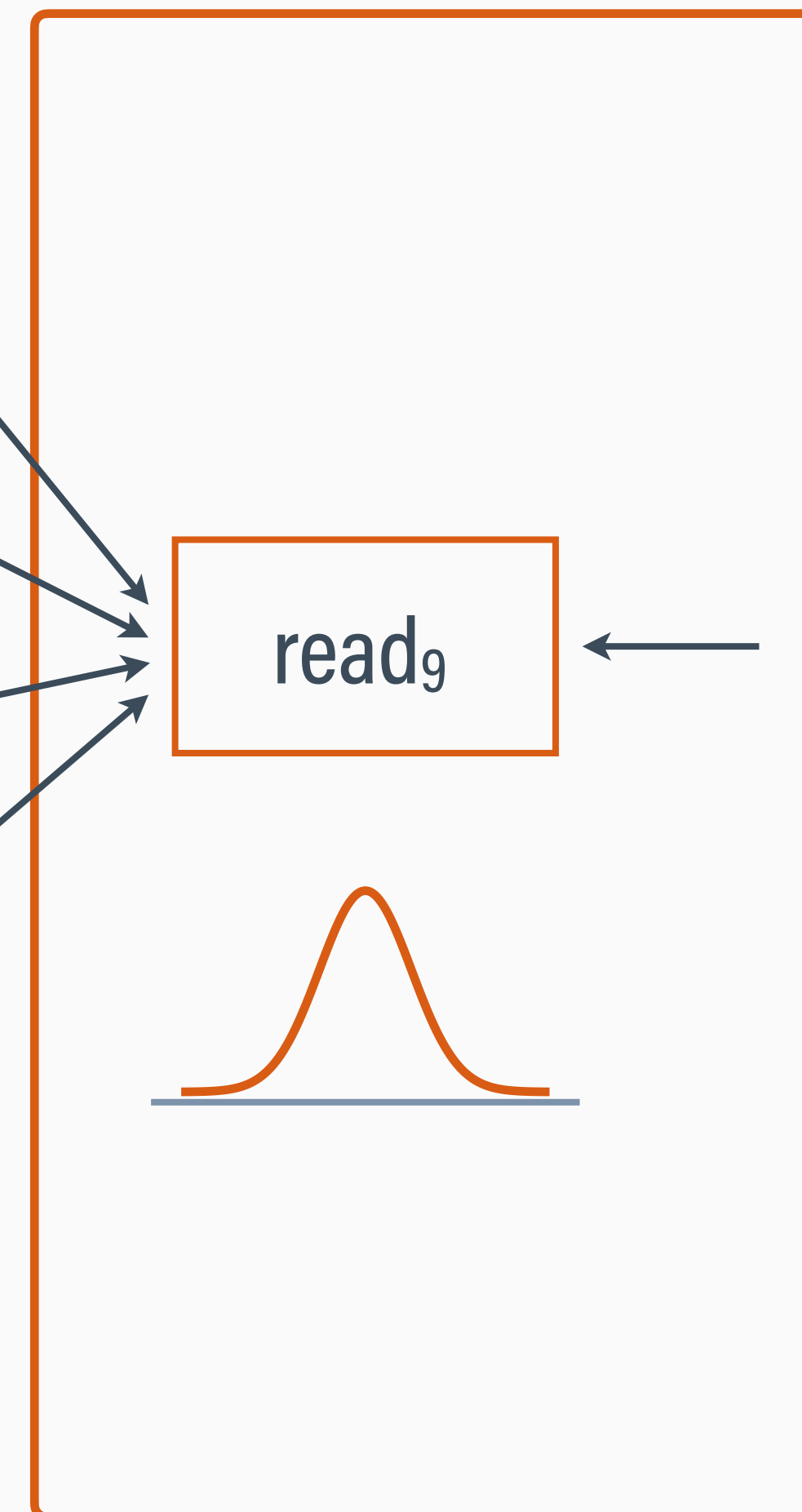
Variable	Definition	Missing %	Scale
riskgrp	Emotional/behavioral risk code	2.2	1 = Low, 2 = Medium, 3 = High
lrnprob1	1st grade learning problems	2.2	Numeric (31 to 88)
read1	1st grade broad reading composite	6.5	Numeric (39 to 153)
read9	9th grade broad reading composite	17.4	Numeric (41 to 123)

# TWO-PART FACTORED SPECIFICATION

Incomplete Predictor Model



Outcome Model





# BLIMP STUDIO SCRIPT 3

---

**DATA:** reading.dat;

**VARIABLES:** id male hispanic riskgrp atrisk behsymp1 lnprob1 read1 read2 read3 read9  
read9grp stanread7 math1 math2 math3 math9 math9grp stanmath7;

**MISSING:** 999;

**NOMINAL:** riskgrp;      # automatic dummy coding with lowest score as the reference

**MODEL:** read9 ~ read1 lnprob1 riskgrp;

**BURN:** 10000;

**ITER:** 10000;

**SEED:** 90291;

# RBLIMP SCRIPT (MODEL 3)

---

```
# fit model
```

```
model3 <- rblimp(
```

```
  data = reading,
```

```
  nominal = 'riskgrp',      # automatic dummy coding with lowest score as the reference
```

```
  model = 'read9 ~ read1 lnp1 riskgrp',
```

```
  seed = 90291,
```

```
  burn = 10000,
```

```
  iter = 10000)
```

```
# summarize results
```

```
output(model3)
```

```
posterior_plot(model3, 'read9')
```

# PSR DIAGNOSTIC OUTPUT

The number of burn-in iterations is sufficient because the highest PSRF across all parameters is < 1.05 at the end of the burn-in period

## BURN-IN POTENTIAL SCALE REDUCTION (PSR) OUTPUT:

NOTE: Split chain PSR is being used. This splits each chain's iterations to create twice as many chains.

Comparing iterations across 2 chains	Highest PSR	Parameter #
251 to 500	1.096	18
501 to 1000	1.037	18
751 to 1500	1.012	14
1001 to 2000	1.018	14
... ..	...	...
4001 to 8000	1.003	25
4251 to 8500	1.003	18
4501 to 9000	1.002	18
4751 to 9500	1.002	18
5001 to 10000	>> 1.002	14

# MISSING DATA INFORMATION

- Per-variable missingness rates appear in the rows, and missing data patterns appear in the columns (M = missing)

DATA INFORMATION:

Sample Size: 138  
Nominal Dummy Codes:

riskgrp = riskgrp.2 riskgrp.3

Missing Data Info:

	miss %	1	2	3	4	5	6
read9 = 17.4		-	M	-	-	-	M
riskgrp = 2.2		-	-	-	M	-	-
lnnprob1 = 2.2		-	-	-	-	M	-
read1 = 6.5		-	-	M	-	-	M
	%	72.5	16.7	5.8	2.2	2.2	0.7

# REGRESSION SUMMARY TABLE

## OUTCOME MODEL ESTIMATES:

Summaries based on 10000 iterations using 2 chains.

NOTE: Estimate column based on posterior median.

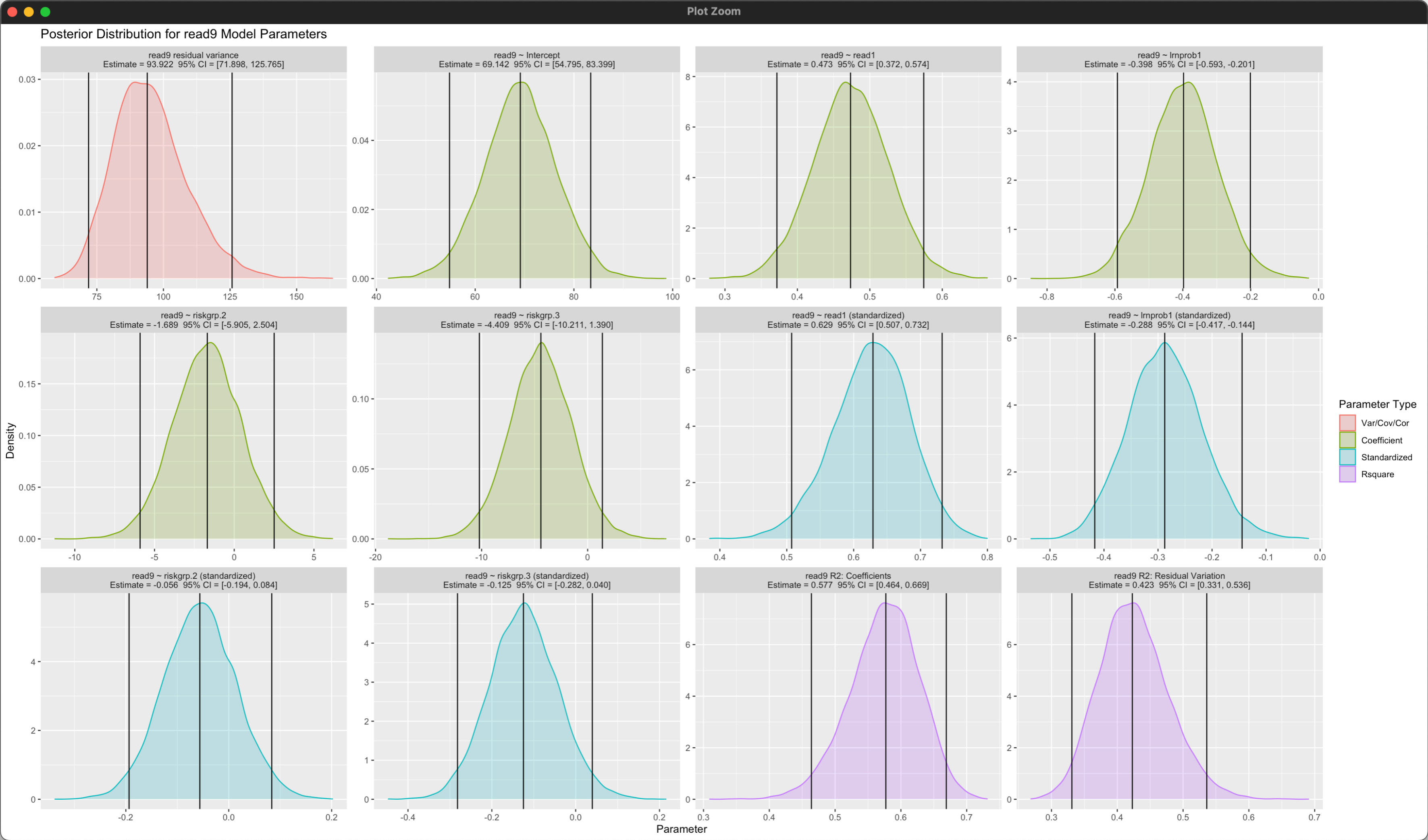
Outcome Variable: **read9**

Parameters	Estimate	StdDev	2.5%	97.5%	ChiSq	PValue	N_Eff
<hr/>							
Variances:							
Residual Var.	93.922	13.762	71.897	125.783	---	---	5004.950
Coefficients:							
Intercept	69.142	7.228	54.792	83.413	91.460	0.000	3685.210
read1	0.473	0.052	0.372	0.574	84.275	0.000	6143.602
lnnprob1	-0.398	0.100	-0.593	-0.201	15.889	0.000	3028.647
riskgrp.2	-1.689	2.132	-5.906	2.505	0.633	0.426	5716.054
riskgrp.3	-4.409	2.936	-10.212	1.391	2.263	0.133	4779.567

• • •

Proportion Variance Explained							
by Coefficients	0.577	0.053	0.464	0.669	---	---	5100.899
by Residual Variation	0.423	0.053	0.331	0.536	---	---	5100.899

# DISTRIBUTION PLOTS (RBLIMP ONLY)



# INTERPRETATIONS

---

- Blimp automatically dummy codes nominal variables, treating the lowest numeric code (low risk) as the reference
- For two students with same first grade reading and learning problems scores, being classified as moderate versus low risk is associated with a  $-1.69$  point decrease in grade 9 reading
- For two students with same first grade reading and learning problems scores, being classified as high versus low risk is associated with a  $-4.41$  point decrease in grade 9 reading



# OUTLINE

1

Modern Missing Data Methods

2

Missing Data Mechanisms

3

MCMC Estimation and Missing Data Imputation

4

Fitting Regression Models in Blimp

5

Incomplete Categorical Variables

6

Interaction Effects

7

Reporting Results From a Missing Data Analysis

# MODERATED REGRESSION

---

- Moderation occurs when a focal predictor's influence on an outcome depends on a third variable called a moderator
- Moderated regression answers the question, for whom does an effect apply?
- Does the diagnostic utility of first-grade reading performance on ninth-grade reading achievement depend on whether a student is experiencing learning problems in first grade?

# ANALYSIS 4 REGRESSION MODEL

---

- An interaction is formed by multiplying two predictors, either of which (or both) could be incomplete

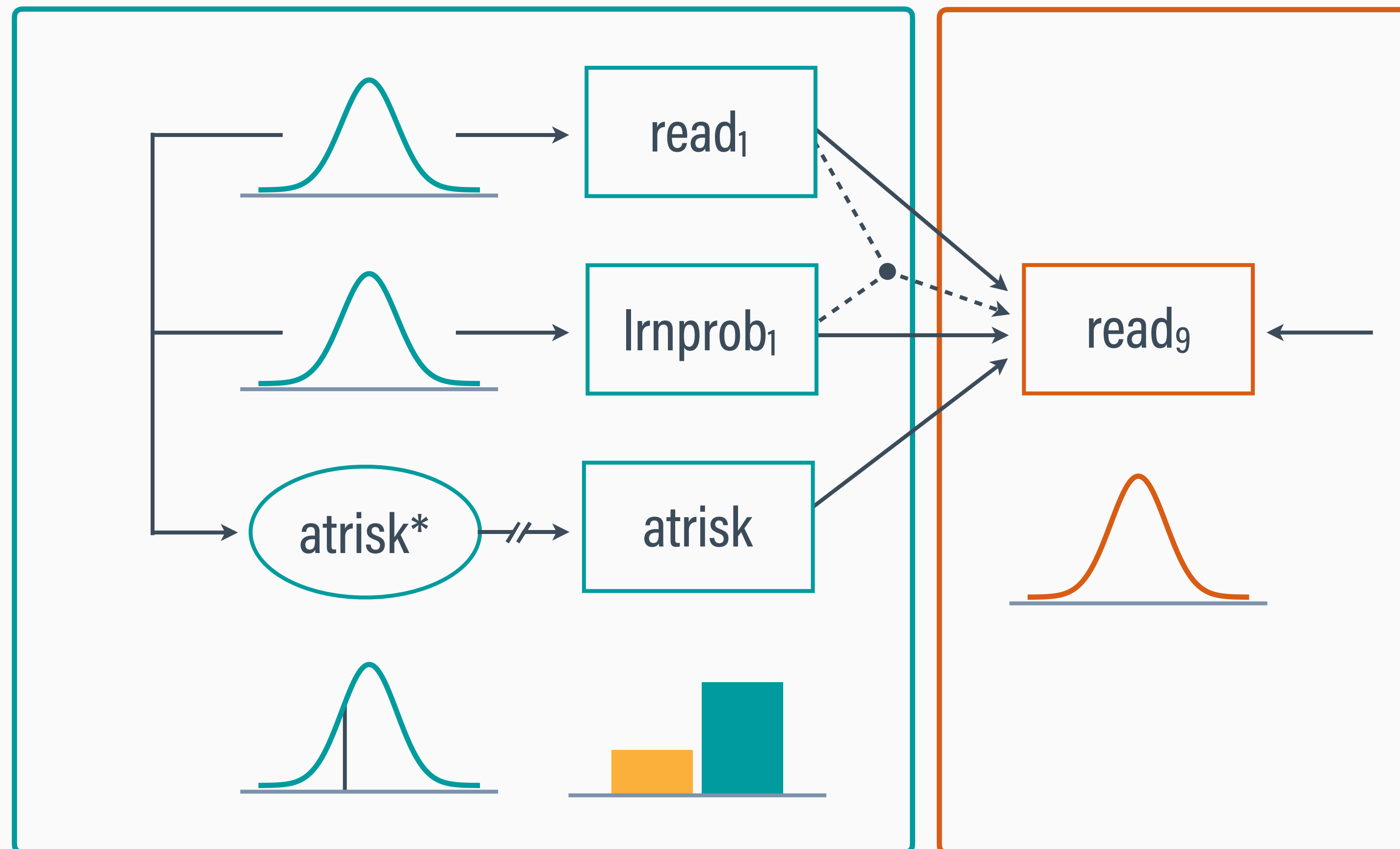
$$\text{read}_9 = \beta_0 + \beta_1(\text{read}_1) + \beta_2(\text{lrnprob}_1) + \beta_3(\text{read}_1)(\text{lrnprob}_1) + \beta_4(\text{atrisk}) + \varepsilon$$

Variable	Definition	Missing %	Scale
riskgrp	Emotional/behavioral risk code	2.2	1 = Low, 2 = Medium, 3 = High
lrnprob1	1st grade learning problems	2.2	Numeric (31 to 88)
read1	1st grade broad reading composite	6.5	Numeric (39 to 153)
read9	9th grade broad reading composite	17.4	Numeric (41 to 123)

# TWO-PART FACTORED SPECIFICATION

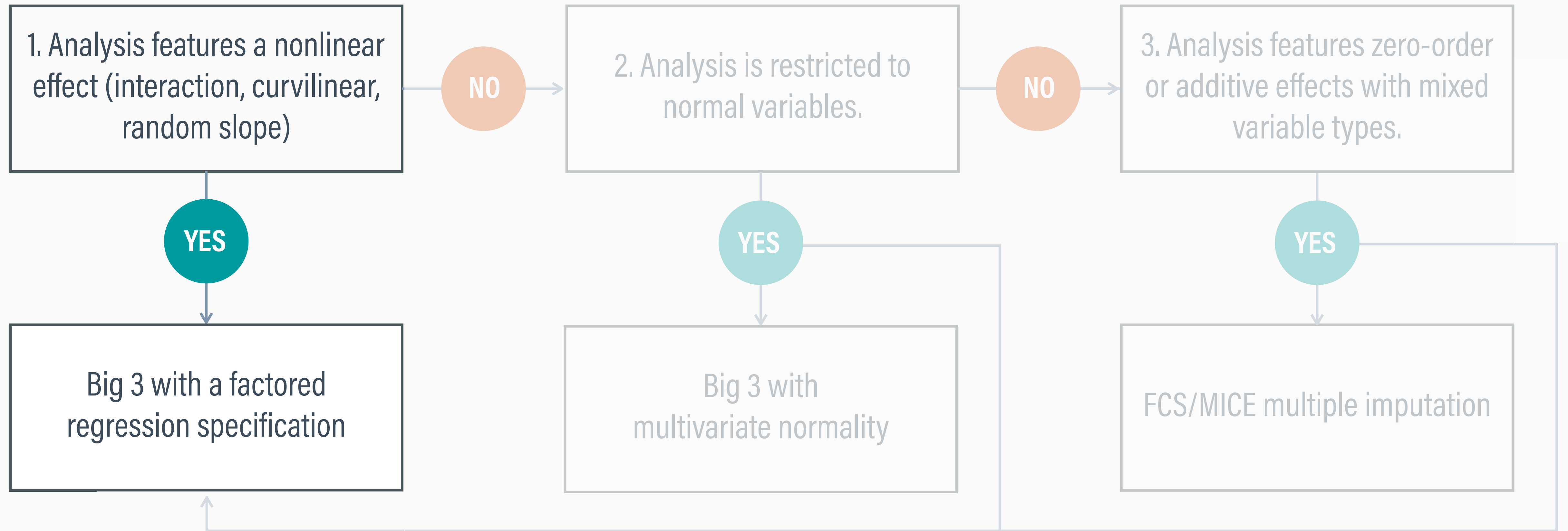
Incomplete Predictor Model

Outcome Model



# MISSING DATA DECISION TREE

Interaction and nonlinear effects require factored specifications with specialized software.



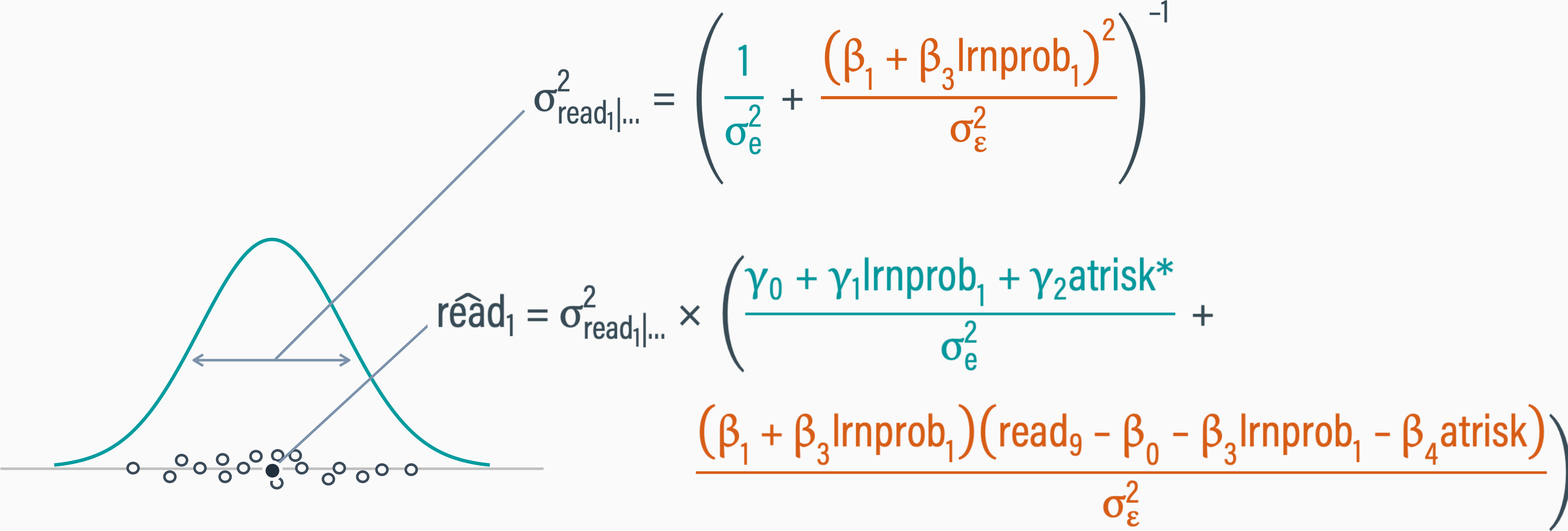
# INCOMPLETE PRODUCT TERMS

---

- Products are deterministic functions of lower-order predictors in the focal model rather than unique variables themselves
- The two-part imputation for incomplete predictors remains the same as before
- When the interaction is non-zero, the two-part specification produces non-normal (heteroscedastic) predictor imputations that accommodate the nonlinear term in the focal model

# PREDICTED VALUES AND VARIATION

Multiple sets of model parameters define the mean and spread of the imputations





# BLIMP STUDIO SCRIPT 4

---

**DATA:** reading.dat;

**VARIABLES:** id male hispanic riskgrp atrisk behsymp1 lnrprob1 read1 read2 read3 read9  
read9grp stanread7 math1 math2 math3 math9 math9grp stanmath7;

**MISSING:** 999;

**NOMINAL:** atrisk;

**CENTER:** read1 lnrprob1;

**MODEL:** read9 ~ read1 lnrprob1 read1\*lnrprob1 atrisk; # product in focal model

**SIMPLE:** read1 | lnrprob1; # conditional effects of read1 at different levels of lnrprob1

**BURN:** 10000;

**ITER:** 10000;

**SEED:** 90291;

# RBLIMP SCRIPT (MODEL 4)

---

# fit model

```
model4 <- rblimp(
```

```
  data = reading,
```

```
  nominal = 'atrisk',
```

```
  center = 'read1 lnrprob1',
```

```
  model = 'read9 ~ read1 lnrprob1 read1*lnrprob1 atrisk',
```

```
  simple = 'read1 | lnrprob1',
```

```
  seed = 90291,
```

```
  burn = 10000,
```

```
  iter = 10000)
```

# product in focal model

# conditional effects of read1 at different levels of lnrprob1

# summarize results

```
output(model4)
```

```
posterior_plot(model4, 'read9')
```

```
simple_plot(read9 ~ read1 | lnrprob1, model4) # plot simple slopes at different values of moderator
```

```
jn_plot(read9 ~ read1 | lnrprob1, model4) # Johnson-Neyman regions of significance
```

# PSR DIAGNOSTIC OUTPUT

The number of burn-in iterations is sufficient because the highest PSRF across all parameters is  $< 1.05$  at the end of the burn-in period

## BURN-IN POTENTIAL SCALE REDUCTION (PSR) OUTPUT:

NOTE: Split chain PSR is being used. This splits each chain's iterations to create twice as many chains.

Comparing iterations across 2 chains	Highest PSR	Parameter #
251 to 500	1.032	15
501 to 1000	1.023	21
751 to 1500	1.054	21
1001 to 2000	1.055	21
...	...	...
4001 to 8000	1.002	21
4251 to 8500	1.005	21
4501 to 9000	1.006	21
4751 to 9500	1.003	21
5001 to 10000	>> 1.007	21

# MISSING DATA INFORMATION

- Per-variable missingness rates appear in the rows, and missing data patterns appear in the columns (M = missing)

DATA INFORMATION:

Sample Size: 138  
Nominal Dummy Codes:

atrisk = atrisk.1

Missing Data Info:

	miss %	1	2	3	4	5	6
read9 = 17.4		-	M	-	-	-	M
atrisk = 2.2		-	-	-	M	-	-
lrrnprob1 = 2.2		-	-	-	-	M	-
read1 = 6.5		-	-	M	-	-	M
	%	72.5	16.7	5.8	2.2	2.2	0.7

# REGRESSION SUMMARY TABLE

## OUTCOME MODEL ESTIMATES:

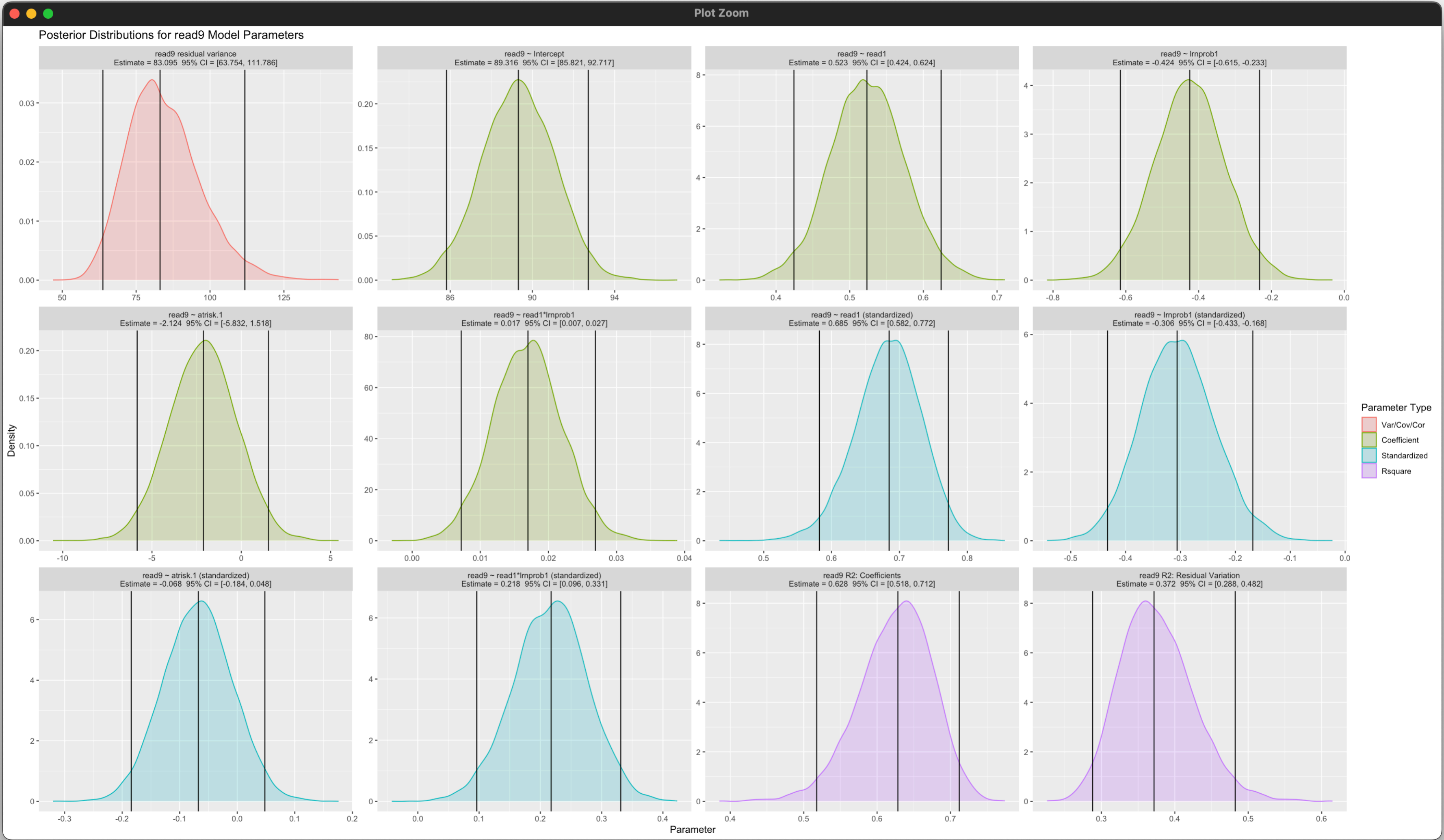
Summaries based on 10000 iterations using 2 chains.

NOTE: Estimate column based on posterior median.

Outcome Variable: **read9**

Parameters	Estimate	StdDev	2.5%	97.5%	ChiSq	PValue	N_Eff
-----							
Variances:							
Residual Var.	83.095	12.307	63.746	111.827	---	---	4424.378
Coefficients:							
Intercept	89.316	1.752	85.821	92.718	2599.229	0.000	1732.732
read1	0.523	0.050	0.424	0.624	108.997	0.000	3908.295
lnnprob1	-0.424	0.097	-0.615	-0.233	19.142	0.000	2155.092
atrisk.1	-2.124	1.886	-5.835	1.520	1.284	0.257	5651.625
read1*lnnprob1	0.017	0.005	0.007	0.027	11.192	0.001	3025.910
...							
Proportion Variance Explained							
by Coefficients	0.628	0.050	0.518	0.712	---	---	4420.718
by Residual Variation	0.372	0.050	0.288	0.482	---	---	4420.718

# DISTRIBUTION PLOTS (RBLIMP ONLY)





# INTERPRETATIONS

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- Lower-order terms are conditional effects that depend on a meaningful zero, which was achieved by centering
- For two students at the mean of the learning problems distribution, scoring one point higher on the first grade reading test is associated with a 0.52 increase in ninth-grade reading
- For two students at the mean of the first grade reading distribution, being rated one point higher on the first grade learning problems measure is associated with a  $-0.42$  decrease in ninth-grade reading

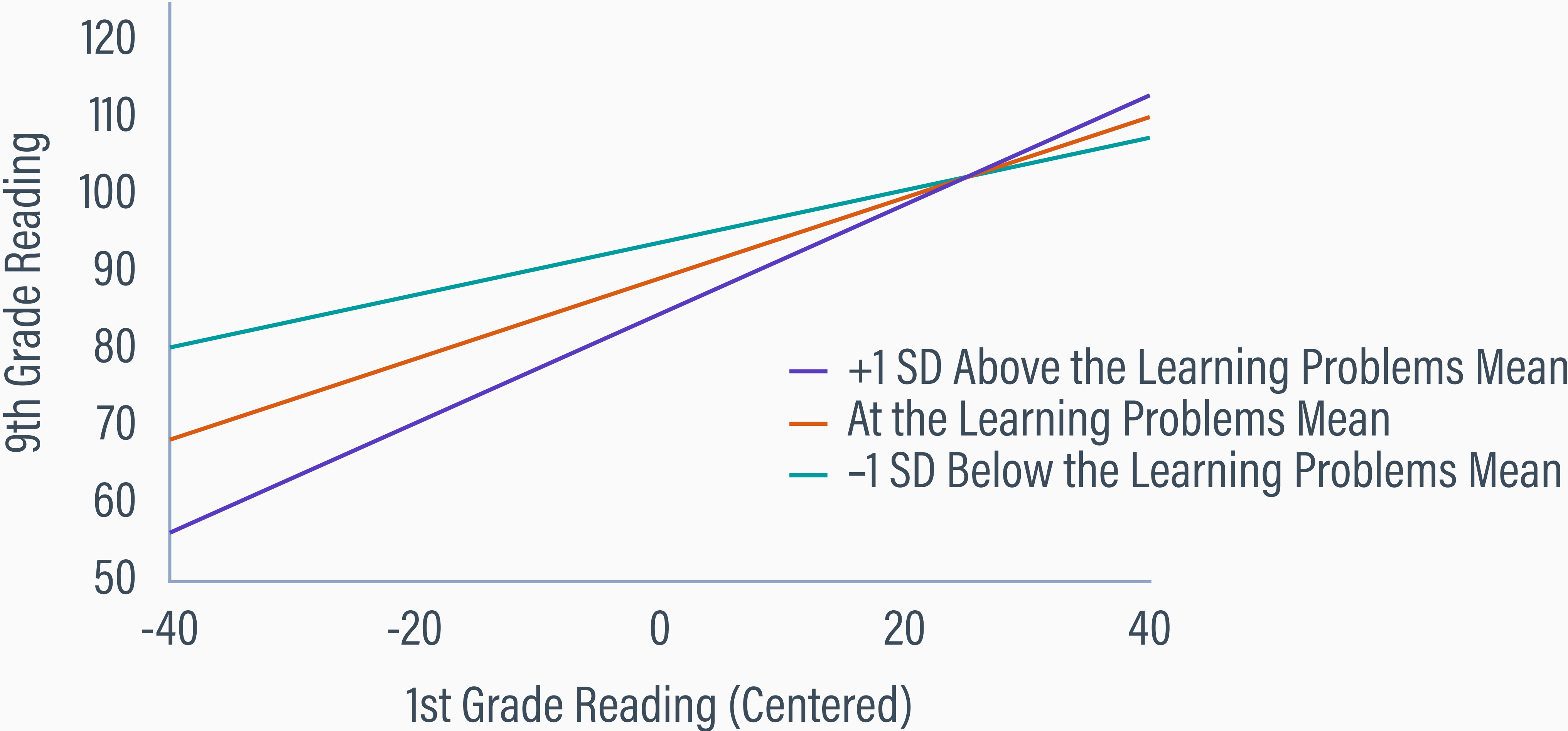


# INTERPRETATIONS, CONTINUED

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- The interaction coefficient,  $= .02$ , is the amount by which the focal slope changes for a one-unit difference on the moderator
- For two students at the mean of the learning problems distribution, scoring one point higher on the first grade reading test is associated with a 0.52 increase in ninth-grade reading
- For two students one point above the learning problems mean, scoring one point higher on the first grade reading test is associated with a  $0.52 + 0.02$  increase in ninth-grade reading

# CONDITIONAL EFFECTS (SIMPLE SLOPES)

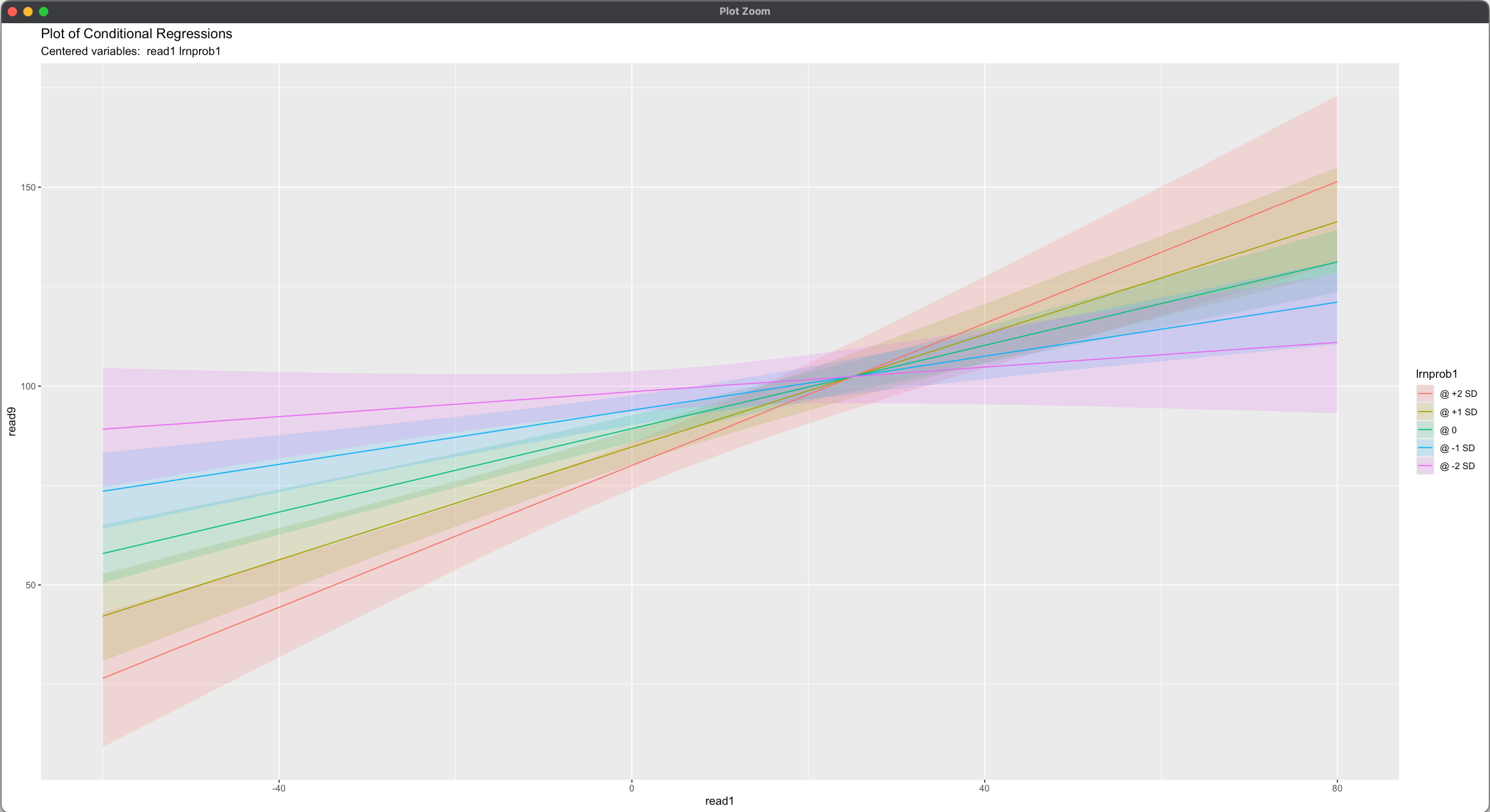


# CONDITIONAL EFFECTS SUMMARY TABLE

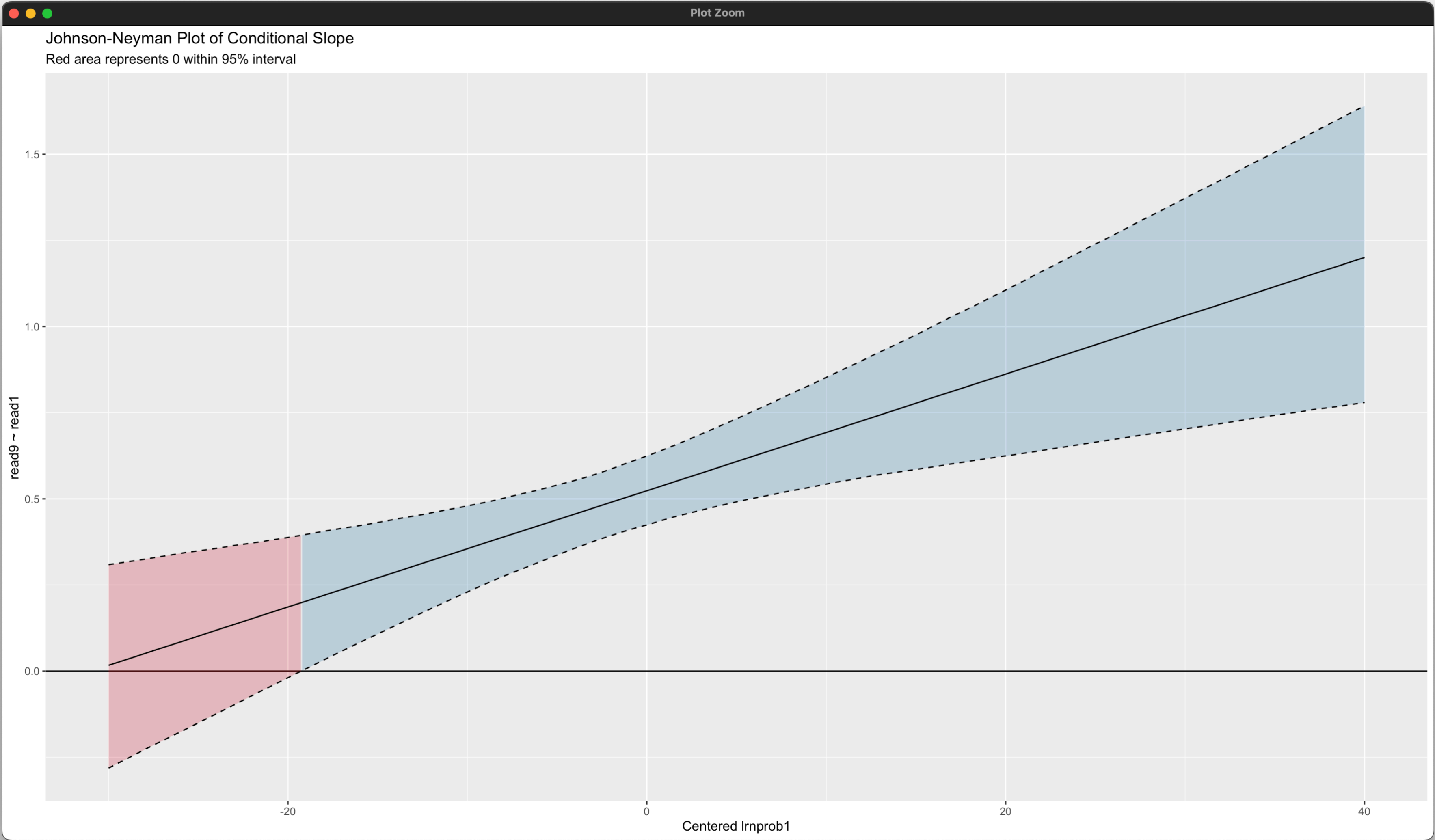
Conditional Effects	Estimate	StdDev	2.5%	97.5%	ChiSq	PValue	N_Eff
<hr/>							
read1   lnrprob1 @ +2 SD							
Intercept	80.052	3.019	74.043	85.852	702.721	0.000	1602.353
Slope	0.893	0.132	0.639	1.158	45.808	0.000	3153.284
read1   lnrprob1 @ +1 SD							
Intercept	84.692	2.211	80.319	89.026	1467.138	0.000	1532.751
Slope	0.709	0.083	0.550	0.876	73.692	0.000	3170.278
read1   lnrprob1 @ 0							
Intercept	89.316	1.752	85.821	92.718	2599.229	0.000	1732.732
Slope	0.523	0.050	0.424	0.624	108.997	0.000	3908.295
read1   lnrprob1 @ -1 SD							
Intercept	93.951	1.913	90.167	97.671	2410.733	0.000	2582.344
Slope	0.340	0.067	0.205	0.470	25.269	0.000	4334.720
read1   lnrprob1 @ -2 SD							
Intercept	98.582	2.582	93.557	103.730	1458.470	0.000	3236.081
Slope	0.156	0.114	-0.073	0.373	1.832	0.176	3688.862
<hr/>							

NOTE: Intercepts are computed by setting all predictors not involved in the conditional effect to zero.

# CONDITIONAL EFFECT PLOTS (RBLIMP ONLY)



# JOHNSON-NEYMAN PLOTS (RBLIMP ONLY)



# OUTLINE

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Modern Missing Data Methods

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Missing Data Mechanisms

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Fitting Regression Models in Blimp

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6

Interaction Effects

7

Reporting Results From a Missing Data Analysis

# REPORTING CHECKLIST

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Missing Data Rates

2

Assumed Missing Data Process

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Missing Data Handling Method

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Algorithmic and Model-Checking Details

6

Interpretation Details



## MISSING DATA RATES AND ASSUMPTIONS

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Missing data rates on predictor variables measured in first grade ranged between 2.2% (risk status and learning problems) and 6.5% (reading performance).

Approximately 17.9% of ninth-grade reading scores were missing. Our analyses are based on the conditionally missing at random assumption where missingness systematically vary as a function of observed scores but not the unseen scores.

## MISSING DATA APPROACH AND SOFTWARE

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We fit our models using Bayesian MCMC estimation in the Blimp 3 software (Keller & Enders, 2021). Given the same assumptions and data, MCMC and likelihood-based missing data handling procedures are numerically equivalent (Enders, 2022). However, MCMC estimation is preferable because classic FIML estimator is known to introduce bias when applied to models interactive effects (the so-called “just another variable” approach; Lüdtke et al., 2020; Zhang & Wang, 2017). MCMC is similar to FIML in the sense that it directly estimates the model of interest, and it is similar to multiple imputation in the sense that it averages over thousands of realizations of the missing values

## ALGORITHMIC DETAILS

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We used Blimp's default diffuse (non-informative) prior distributions, as described in Section 1.7 of the user guide. We used two MCMC chains with random starting values to generate model summaries consisting of 10,000 estimates following an initial burn-in period of 10,000 cycles. The potential scale reduction factor convergence diagnostics (Gelman & Rubin, 1992) indicated that MCMC converged in fewer than 1,000 iterations, so a 10,000-cycle warm-up period was sufficiently conservative. We verified the total number of iterations for the analysis was sufficient by examining the effective number of independent MCMC samples for each parameter, all of which were greater than the recommended value of 100 (Gelman et al., 2014, p. 287).

# REPORTING TEMPLATE CONTINUED

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Table 1 displays the regression summary table from the analysis. From a Bayesian perspective, the posterior medians and standard deviations are analogous to frequentist point estimates and standard errors, and the 95% credible interval limits are akin to confidence intervals. However, these quantities make no reference to repeated samples but instead convey parameter values that are consistent with the observed data. Given the same assumptions and data, Bayesian and likelihood-based missing data handling procedures are numerically equivalent (Enders, 2022). Taking a computational frequentist perspective, these MCMC-generated summaries can also be viewed as surrogates for frequentist point estimates, standard errors, and confidence intervals (Levy & McNeish, 2021)

# SUMMARY TABLE

Table 1  
Regression Model Summary

Parameter	Est.	SD	2.5%	97.5%	Chi-Sq.	p
Intercept	89.32	1.75	85.82	92.72	2599.23	< .001
1st Grade Reading	0.52	0.05	0.42	0.62	109.00	< .001
Learning Problems	−0.42	0.10	−0.62	−0.23	19.14	< .001
Reading by Problems	0.02	0.01	0.01	0.03	11.19	< .001
At Risk Indicator	−2.12	1.89	−5.84	1.52	1.28	.26
Residual variance	83.10	12.31	63.75	111.83	—	—
R <sup>2</sup>	0.63	0.05	0.52	0.71	—	—



# SUMMARY OF RESULTS

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Collectively, the predictors explained approximately 63% of the variation in 9th grade reading scores. At the learning problems mean, first grade reading exhibited a significant positive association with 9th grade reading performance ( $\beta = 0.52$ ,  $SD = 0.05$ ,  $CI = [0.42, 0.62]$ ), controlling for other predictors. The MCMC Wald test (Asparouhov & Muthén, 2021) of the interaction effect was statistically significant,  $\chi^2(1) = 11.19$ ,  $p < .001$ . The positive interaction coefficient indicates that the first grade reading slope increases as learning problems increase ( $\beta = 0.02$ ,  $SD = 0.01$ ,  $CI = [0.01, 0.03]$ ), such that first grade test scores become increasingly predictive of later achievement. Figure 1 displays the simple slopes at three levels of learning problems.



For more information go to

**[WWW.APPLIEDMISSINGDATA.COM](http://WWW.APPLIEDMISSINGDATA.COM)**