

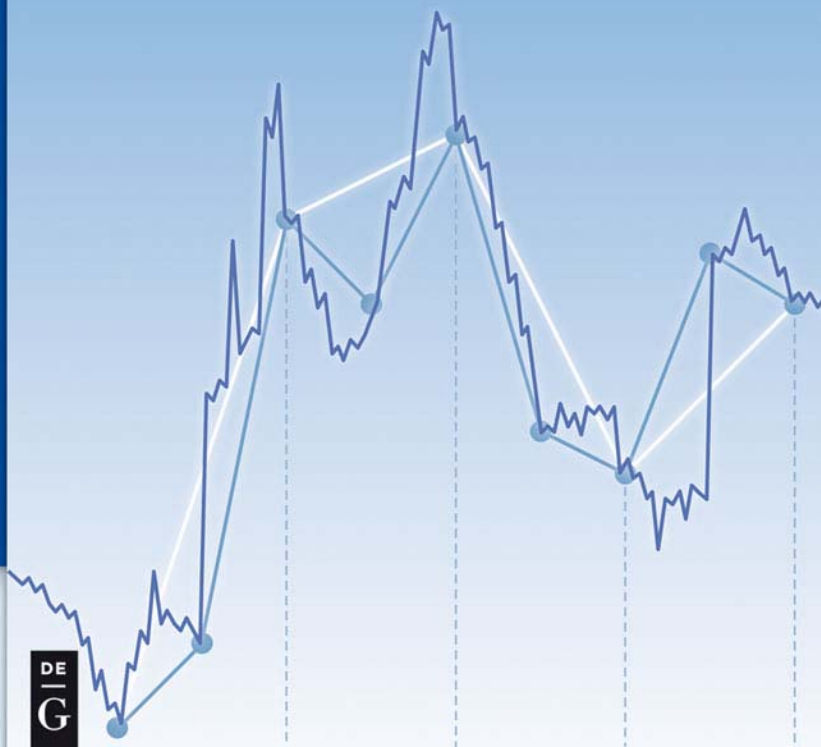
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René L. Schilling, Lothar Partzsch

BROWNIAN MOTION

AN INTRODUCTION TO STOCHASTIC PROCESSES



De Gruyter Graduate

Schilling / Partzsch · Brownian Motion

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Brownian Motion

An Introduction to Stochastic Processes

With a Chapter on Simulation by Björn Böttcher

De Gruyter

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Preface

Brownian motion is arguably the single most important stochastic process. Historically it was the first stochastic process in continuous time and with a continuous state space, and thus it influenced the study of Gaussian processes, martingales, Markov processes, diffusions and random fractals. Its central position within mathematics is matched by numerous applications in science, engineering and mathematical finance.

The present book grew out of several courses which we taught at the University of Marburg and TU Dresden, and it draws on the lecture notes [141] by one of us. Many students are interested in applications of probability theory and it is important to teach Brownian motion and stochastic calculus at an early stage of the curriculum. Such a course is very likely the first encounter with stochastic processes in continuous time, following directly on an introductory course on rigorous (i. e. measure-theoretic) probability theory. Typically, students would be familiar with the classical limit theorems of probability theory and basic discrete-time martingales, as it is treated, for example, by Jacod & Protter *Probability Essentials* [88], Williams *Probability with Martingales* [189], or in the more voluminous textbooks by Billingsley [11] and Durrett [50].

General textbooks on probability theory cover however, if at all, Brownian motion only briefly. On the other hand, there is a quite substantial gap to more specialized texts on Brownian motion which is not so easy to overcome for the novice. Our aim was to write a book which can be used in the classroom as an introduction to Brownian motion and stochastic calculus, and as a first course in continuous-time and continuous-state Markov processes. We also wanted to have a text which would be both a readily accessible mathematical back-up for contemporary applications (such as mathematical finance) and a foundation to get easy access to advanced monographs, e. g. Karatzas & Shreve [99], Revuz & Yor [156] or Rogers & Williams [161] (for stochastic calculus), Marcus & Rosen [129] (for Gaussian processes), Peres & Mörters [133] (for random fractals), Chung [23] or Port & Stone [149] (for potential theory) or Blumenthal & Gettoor [13] (for Markov processes) to name but a few.


Things the readers are expected to know: Our presentation is basically self-contained, starting from ‘scratch’ with continuous-time stochastic processes. We do, however, assume some basic measure theory (as in [169]) and a first course on probability theory and discrete-time martingales (as in [88] or [189]). Some ‘remedial’ material is collected in the appendix, but this is really intended as a back-up.

How to read this book: Of course, nothing prevents you from reading it linearly. But there is more material here than one could cover in a one-semester course. De-

pending on your needs and likings, there are at least three possible selections: *BM and Itô calculus*, *BM and its sample paths* and *BM as a Markov process*. The diagram on page xi will give you some ideas how things depend on each other and how to construct your own ‘Brownian sample path’ through this book.



Ex. N.N.

Whenever special attention is needed and to point out traps & pitfalls, we have used the  sign in the margin. Also in the margin, there are cross-references to exercises at the end of each chapter which we think fit (and are sometimes needed) at that point.¹ They are not just drill problems but contain variants, excursions from and extensions of the material presented in the text. The proofs of the core material do not seriously depend on any of the problems.

Writing an introductory text also meant that we had to omit many beautiful topics. Often we had to stop at a point where we, hopefully, got you really interested... Therefore, we close every chapter with a brief outlook on possible texts for further reading.

Many people contributed towards the completion of this project: First of all the students who attended our courses and helped – often unwittingly – to shape the presentation of the material. We profited a lot from comments by Niels Jacob (Swansea) and Panki Kim (Seoul National University) who used an early draft of the manuscript in one of his courses. Special thanks go to our colleagues and students Björn Böttcher, Katharina Fischer, Julian Hollender, Felix Lindner and Michael Schwarzenberger who read substantial parts of the text, often several times and at various stages. They found countless misprints, inconsistencies and errors which we would never have spotted. Björn helped out with many illustrations and, more importantly, contributed Chapter 20 on simulation. Finally we thank our colleagues and friends at TU Dresden and our families who contributed to this work in many uncredited ways. We hope that they approve of the result.

Dresden, February 2012

René L. Schilling
Lothar Partzsch

¹ For the readers’ convenience there is a web page where additional material and solutions are available. The URL is http://www.motapa.de/brownian_motion/index.html

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Dependence chart

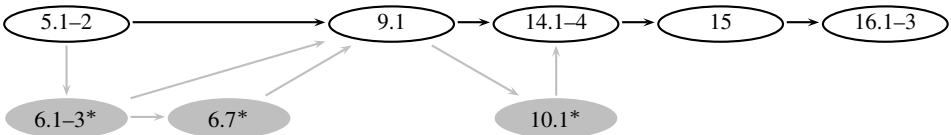
As we have already mentioned in the preface, there are at least three paths through this book which highlight different aspects of Brownian motion: Brownian motion and Itô calculus, Brownian motion as a Markov process, and Brownian motion and its sample paths. Below we suggest some fast tracks “C”, “M” and “S” for each route, and we indicate how the other topics covered in this book depend on these fast tracks. This should help you to find your own personal sample path. Starred sections (in the grey ovals) contain results which can be used without proof and without compromising too much on rigour.

Getting started

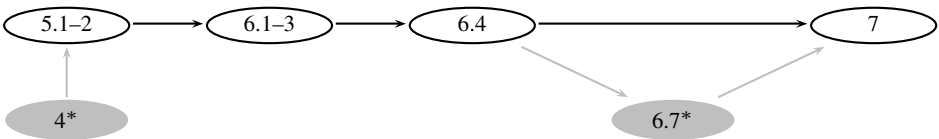
For all three fast tracks you need to read Chapters 1 and 2 first. If you are not too much in a hurry, you should choose *one* construction of Brownian motion from Chapter 3. For the beginner we recommend either 3.1 or 3.2.



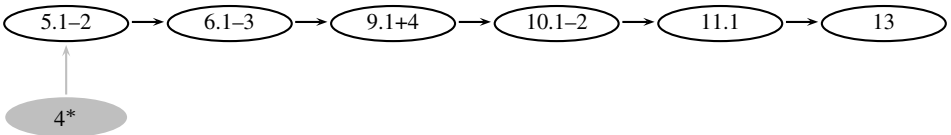
Basic stochastic calculus (C)



Basic Markov processes (M)



Basic sample path properties (S)



Dependence to the sections 5.1–19.2

The following list shows which prerequisites are needed for each section. A star as in 4* or 6.7* indicates that some result(s) from Chapter 4 or Section 6.7 are used which may be used without proof and without compromising too much on rigour. Starred sections are mentioned only where they are actually needed, while other prerequisites are repeated to indicate the full line of dependence. For example,

6.6: *M or S or C*, 6.1–3

indicates that the prerequisites for Section 6.6 are covered by either “M” or “S” or “C if you add 6.1, 6.2, 6.3”. Since we do not refer to later sections with higher numbers, you will only need those sections in “M”, “S”, or “C and 6.1, 6.2, 6.3” with section numbers below 6.6. Likewise,

17.1: *C*, 16.4–5, 14.6*

means that 17.1 requires “C” plus the Sections 16.4 and 16.5. Some results from 14.6 are used, but they can be quoted without proof.

5.1:	<i>C or M or S</i>	9.2:	<i>S or C or M</i> , 9.1	16.3:	<i>C</i>
5.2:	<i>C or M or S</i>	9.3:	<i>S or C or M</i> , 9.1	16.4:	<i>C</i>
5.3:	<i>C or M or S</i>	9.4:	<i>S or C or M</i> , 9.1	16.5:	<i>C</i> , 16.4
6.1:	<i>M or S or C</i>	10.1:	<i>S or C or M</i>	16.6:	<i>C</i>
6.2:	<i>M or S or C</i> , 6.1	10.2:	<i>S or C or M</i>	17.1:	<i>C</i> , 16.4–5, 14.6*
6.3:	<i>M or S or C</i> , 6.1–2	10.3:	<i>S or C or M</i>	17.2:	<i>C</i> , 16.4–5
6.4:	<i>M or S or C</i> , 6.1–3	11.1:	<i>S</i> , 10.3* <i>or C</i> , 10.3* <i>or M</i> , 10.3*	17.3:	<i>C</i> , 16.4–5, 17.1
6.5:	<i>M or S or C</i> , 6.1–3			17.4:	<i>C</i> , 16.4–5
6.6:	<i>M or S or C</i> , 6.1–3	11.2:	<i>S or C</i> , 11.1 <i>or M</i> , 11.1	17.5:	<i>C</i> , 14.5–6, 16.4–5, 17.2
6.7:	<i>M or S or C</i> , 6.1–3	12.1:	<i>S</i>	17.6:	<i>C</i> , 16.4–5, 17.2
7.1:	<i>M</i> , 4.2* <i>or C</i> , 6.1, 4.2*	12.2:	<i>S</i> , 12.1	17.7:	<i>C</i> , 16.4–5
7.2:	<i>M or C</i> , 6.1, 7.1	12.3:	<i>S</i> , 12.1–2, 4*	18.1:	<i>C</i>
7.3:	<i>M or C</i> , 6.1, 7.1–2	13:	<i>S or C or M</i>	18.2:	<i>C</i> , 18.1
7.4:	<i>M or C</i> , 6.1, 7.1–3	14.1:	<i>C or M</i>	18.3:	<i>C</i> , 16.4–5, 18.1–2
7.5:	<i>M or C</i> , 6.1, 7.1–4	14.2:	<i>C</i> , 6.7* <i>or M</i> , 14.1, 6.7*	18.4:	<i>C</i> , 6.1, 16.4–5, 18.1–3
8.1:	<i>M or C</i> , 6.1, 7.1–3	14.3:	<i>C or M</i> , 14.1–2	18.5:	<i>C</i> , 16.4–5, 18.1–3
8.2:	<i>M</i> , 8.1 <i>or C</i> , 6.1, 7.1–3, 8.1	14.4:	<i>C or M</i> , 14.1–3, 9.1*	18.6:	<i>C</i> , 16.4–5, 18.1–3, 10.1*, 17.7*
8.3:	<i>M</i> , 8.1–2 <i>or C</i> , 6.1, 7.1–4, 8.1–2	14.5:	<i>C</i> , 6.7*	19.1:	<i>M or C</i> , 6.1, 7
8.4:	<i>M</i> , 6.7*, 8.1–3* <i>or C</i> , 6.1–4, 7, 6.9*, 8.1–3*	14.6:	<i>C</i> , 14.5	19.2:	<i>C</i> , 6.1, 7, 16.4–5, 18, 19.1
9.1:	<i>S or C or M</i>	15:	<i>C or M</i> , 14.1–4		
		16.1:	<i>C</i>		
		16.2:	<i>C</i>		

Index of notation

This index is intended to aid cross-referencing, so notation that is specific to a single section is generally not listed. Some symbols are used locally, without ambiguity, in senses other than those given below; numbers following an entry are page numbers.

Unless otherwise stated, functions are real-valued and binary operations between functions such as $f \pm g$, $f \cdot g$, $f \wedge g$, $f \vee g$, comparisons $f \leq g$, $f < g$ or limiting relations $f_j \xrightarrow{j \rightarrow \infty} f$, $\lim_j f_j$, $\underline{\lim}_j f_j$, $\overline{\lim}_j f_j$, $\sup_j f_j$ or $\inf_j f_j$ are understood pointwise. ‘Positive’ and ‘negative’ always means ‘ ≥ 0 ’ and ‘ ≤ 0 ’.

General notation: analysis

$\inf \emptyset$	$\inf \emptyset = +\infty$
$a \vee b$	maximum of a and b
$a \wedge b$	minimum of a and b
a^+	$a \vee 0$
a^-	$-(a \wedge 0)$
$\lfloor x \rfloor$	largest integer $n \leq x$
$ x $	Euclidean norm in \mathbb{R}^d , $ x ^2 = x_1^2 + \cdots + x_d^2$, $d \geq 1$
$\langle x, y \rangle$	scalar product in \mathbb{R}^d , $\sum_{j=1}^d x_j y_j$, $d \geq 2$
I_d	unit matrix in $\mathbb{R}^{d \times d}$
$\mathbb{1}_A$	$\mathbb{1}_A(x) = \begin{cases} 1, & x \in A \\ 0, & x \notin A \end{cases}$
$\langle f, g \rangle_{L^2(\mu)}$	scalar product $\int f g \, d\mu$
Leb	Lebesgue measure
λ_T	Leb. measure on $[0, T]$
δ_x	point mass at x
$\mathfrak{D}, \mathfrak{R}$	domain/range
Δ	Laplace operator
$\mathfrak{D}(\Delta)$	domain of Brownian generator, 92–94, 102
∂_j	partial derivative $\frac{\partial}{\partial x_j}$

General notation: probability

\sim	“is distributed as”
$\overset{s}{\sim}$	“is sample of”, 312
$\perp\!\!\!\perp$	“is stochastically independent”
\xrightarrow{d}	convergence in law
$\xrightarrow{\mathbb{P}}$	convergence in probab.
$\xrightarrow{L^p}$	convergence in $L^p(\mathbb{P})$
a. s.	almost surely (w. r. t. \mathbb{P})
iid	independent and identically distributed
LIL	law of iterated logarithm
\mathbb{P}, \mathbb{E}	probability, expectation
\mathbb{V}, Cov	variance, covariance
$N(\mu, \sigma^2)$	normal law in \mathbb{R} , mean μ , variance σ^2
$N(m, \Sigma)$	normal law in \mathbb{R}^d , mean $m \in \mathbb{R}^d$, cov. $\Sigma \in \mathbb{R}^{d \times d}$
BM	Brownian motion, 4
BM^1, BM^d	1-, d -dim. BM, 4
(B0)–(B4)	4
(B3')	6

Sets and σ -algebras

A^c	complement of the set A
\overline{A}	closure of the set A
$\mathbb{B}(x, r)$	open ball, centre x , radius r
$\overline{\mathbb{B}}(x, r)$	closed ball, centre x , radius r
$\text{supp } f$	support, $\overline{\{f \neq 0\}}$
$\mathcal{B}(E)$	Borel sets of E
\mathcal{F}_t^X	$\sigma(X_s : s \leq t)$
\mathcal{F}_{t+}	$\bigcap_{u>t} \mathcal{F}_u$
$\overline{\mathcal{F}}_t$	completion of \mathcal{F}_t with all subsets of \mathbb{P} null sets
\mathcal{F}_∞	$\sigma\left(\bigcup_{t \geq 0} \mathcal{F}_t\right)$
$\mathcal{F}_\tau, \mathcal{F}_{\tau+}$	55, 342
\mathcal{P}	progressive σ -algebra, 219

Processes

$(X_t, \mathcal{F}_t)_{t \geq 0}$	adapted process, 48
$\mathbb{P}^x, \mathbb{E}^x$	law of BM, starting at x , 63 law of Feller process, starting at x , 89–91
σ, τ	stopping times: $\{\sigma \leq t\} \in \mathcal{F}_t$, $t \geq 0$
τ_D, τ_D°	first hitting/entry time, 53
X_t^τ	stopped process $X_{t \wedge \tau}$
$\langle X \rangle_t$	quadratic variation, 204, 212, 354
$\langle X, Y \rangle_t$	quadratic covariation, 206
$\text{var}_p(f; t)$	p -variation on $[0, t]$, 137

\mathcal{E}_T	simple processes, 207
\mathcal{L}_T^2	closure of \mathcal{E}_T , 212
$\mathcal{L}_{T, \text{loc}}^2$	227
$L_{\mathcal{P}}^2$	$f \in L^2$ with \mathcal{P} mble. representative, 219
$L_{\mathcal{P}}^2 = \mathcal{L}_T^2$	221
$\mathcal{M}^2, \mathcal{M}_T^2$	L^2 martingales, 203, 207
$\mathcal{M}_T^{2,c}$	continuous L^2 martingales, 207

Spaces of functions

$\mathcal{B}(E)$	Borel functions on E
$\mathcal{B}_b(E)$	– – , bounded
$\mathcal{C}(E)$	continuous functions on E
$\mathcal{C}_b(E)$	– – , bounded
$\mathcal{C}_\infty(E)$	– – , $\lim_{ x \rightarrow \infty} f(x) = 0$
$\mathcal{C}_c(E)$	– – , compact support
$\mathcal{C}_{(0)}(E)$	– – , $f(0) = 0$,
$\mathcal{C}^k(E)$	k times continuously diff'ble functions on E
$\mathcal{C}_b^k(E)$	– – , bounded (with all derivatives)
$\mathcal{C}_\infty^k(E)$	– – , 0 at infinity (with all derivatives)
$\mathcal{C}_c^k(E)$	– – , compact support
$\mathcal{C}^{1,2}(I \times E)$	$f(\cdot, x) \in \mathcal{C}^1(I)$ and $f(t, \cdot) \in \mathcal{C}^2(E)$
\mathcal{H}^1	Cameron–Martin space, 175
$L^p(E, \mu), L^p(\mu), L^p(E)$	L^p space w. r. t. the measure space (E, \mathcal{A}, μ)

Chapter 1

Robert Brown's new thing¹

If you observe plant pollen in a drop of water through a microscope, you will see an incessant, irregular movement of the particles. The Scottish Botanist Robert Brown was not the first to describe this phenomenon – he refers to W. F. Gleichen-Rußwurm as *the discoverer of the motions of the Particles of the Pollen* [16, p. 164] – but his 1828 and 1829 papers [15, 16] are the first scientific publications investigating ‘Brownian motion’. Brown points out that

- the motion is very irregular, composed of translations and rotations;
- the particles appear to move independently of each other;
- the motion is more active the smaller the particles;
- the composition and density of the particles have no effect;
- the motion is more active the less viscous the fluid;
- the motion never ceases;
- the motion is not caused by flows in the liquid or by evaporation;
- the particles are not animated.

Let us briefly sketch how the story of Brownian motion evolved.

Brownian motion and physics. Following Brown's observations, several theories emerged, but it was Einstein's 1905 paper [57] which gave the correct explanation: The atoms of the fluid perform a temperature-dependent movement and bombard the (in this scale) macroscopic particles suspended in the fluid. These collisions happen frequently and they do not depend on position nor time. In the introduction Einstein remarks: *It is possible, that the movements to be discussed here are identical with the so-called “Brownian molecular motion”; [...] If the movement discussed here can actually be observed (together with the laws relating to it that one would expect to find), then classical thermodynamics can no longer be looked upon as applicable with precision to bodies even of dimensions distinguishable in a microscope: An exact determination of actual atomic dimensions is then possible.*² [58, pp. 1–2]. And between

¹ ‘I have some sea-mice – five specimens – in spirits. And I will throw in Robert Brown's new thing – “Microscopic Observations on the Pollen of Plants” – if you don't happen to have it already.’ in: George Eliot, *Middlemarch*, [59, book II, chapter xvii].

² Es ist möglich, daß die hier zu behandelnden Bewegungen mit der sogenannten “Brownschen Molekularbewegung” identisch sind; [...] Wenn sich die hier zu behandelnde Bewegung samt den

the lines: This would settle the then ongoing discussion on the existence of atoms. It was Jean Perrin who combined in 1909 Einstein's theory and experimental observations of Brownian motion to prove the existence and determine the size of atoms, cf. [144, 145]. Independently of Einstein, M. von Smoluchowski arrived at an equivalent interpretation of Brownian motion, cf. [173].

Brownian motion and mathematics. As a mathematical object, Brownian motion can be traced back to the not completely rigorous definition of Bachelier [4] who makes no connection to Brown or Brownian motion. Bachelier's work was only rediscovered by economists in the 1960s, cf. [31]. The first rigorous mathematical construction of Brownian motion is due to Wiener [185] who introduces the *Wiener measure* on the space $\mathcal{C}[0, 1]$ (which he calls *differential-space*) building on Einstein's and von Smoluchowski's work. Further constructions of Brownian motion were subsequently given by Wiener [186] (Fourier-Wiener series), Kolmogorov [105, 106] (giving a rigorous justification of Bachelier [4]), Lévy [120, 121, pp. 492–494, 17–20] (interpolation argument), Ciesielski [26] (Haar representation) and Donsker [39] (limit of random walks, invariance principle), see Chapter 3.

Let us start with Brown's observations to build a mathematical model of Brownian motion. To keep things simple, we consider a one-dimensional setting where each particle performs a random walk. We assume that each particle

- starts at the origin $x = 0$,
- changes its position only at discrete times $k\Delta t$ where $\Delta t > 0$ is fixed and for all $k = 1, 2, \dots$;
- moves Δx units to the left or to the right with equal probability;

and that

- Δx does not depend on any past positions nor the current position x nor on time $t = k\Delta t$;

Letting $\Delta t \rightarrow 0$ and $\Delta x \rightarrow 0$ in an appropriate way should give a random motion which is continuous in time and space.

Let us denote by X_t the random position of the particle at time $t \in [0, T]$. During the time $[0, T]$, the particle has changed its position $N = \lfloor T/\Delta t \rfloor$ times. Since the decision to move left or right is random, we will model it by independent, identically distributed Bernoulli random variables, $\epsilon_k, k \geq 1$, where

$$\mathbb{P}(\epsilon_1 = 1) = \mathbb{P}(\epsilon_1 = 0) = \frac{1}{2}$$

für sie zu erwartenden Gesetzmäßigkeiten wirklich beobachten läßt, so ist die klassische Thermodynamik schon für mikroskopisch unterscheidbare Räume nicht mehr als genau gültig anzusehen und es ist dann eine exakte Bestimmung der wahren Atomgröße möglich. [57, p. 549]

so that

$$S_N = \epsilon_1 + \cdots + \epsilon_N \quad \text{and} \quad N - S_N$$

denote the number of right and left moves, respectively. Thus

$$X_T = S_N \Delta x - (N - S_N) \Delta x = (2S_N - N) \Delta x = \sum_{k=1}^N (2\epsilon_k - 1) \Delta x$$

is the position of the particle at time $T = N\Delta t$. Since $X_0 = 0$ we find for any two times $t = n\Delta t$ and $T = N\Delta t$ that

$$X_T = (X_T - X_t) + (X_t - X_0) = \sum_{k=n+1}^N (2\epsilon_k - 1) \Delta x + \sum_{k=1}^n (2\epsilon_k - 1) \Delta x.$$

Since the ϵ_k are iid random variables, the two increments $X_T - X_t$ and $X_t - X_0$ are independent and

$$X_T - X_t \sim X_{T-t} - X_0$$

(‘ \sim ’ indicates that the random variables have the same probability distribution). We write $\sigma^2(t) := \mathbb{V} X_t$. By Bienaymé's identity we get

$$\mathbb{V} X_T = \mathbb{V}(X_T - X_t) + \mathbb{V}(X_t - X_0) = \sigma^2(T - t) + \sigma^2(t)$$

which means that $t \mapsto \sigma^2(t)$ is linear:

$$\mathbb{V} X_T = \sigma^2(T) = \sigma^2 T,$$

where $\sigma > 0$ is the so-called *diffusion coefficient*. On the other hand, since $\mathbb{E} \epsilon_1 = \frac{1}{2}$ and $\mathbb{V} \epsilon_1 = \frac{1}{4}$ we get by a direct calculation that

$$\mathbb{V} X_T = N(\Delta x)^2 = \frac{T}{\Delta t} (\Delta x)^2$$

which reveals that

$$\frac{(\Delta x)^2}{\Delta t} = \sigma^2 = \text{const.}$$

The particle's position X_T at time $T = N\Delta t$ is the sum of N iid random variables,

$$X_T = \sum_{k=1}^N (2\epsilon_k - 1) \Delta x = (2S_N - N) \Delta x = S_N^* \sqrt{T} \sigma,$$

where

$$S_N^* = \frac{2S_N - N}{\sqrt{N}} = \frac{S_N - \mathbb{E} S_N}{\sqrt{\mathbb{V} S_N}}$$

is the normalization – i. e. mean 0, variance 1 – of the random variable S_N . A simple application of the central limit theorem now shows that in distribution

$$X_T = \sqrt{T}\sigma S_N^* \xrightarrow[N \rightarrow \infty]{(\text{i.e. } \Delta x, \Delta t \rightarrow 0)} \sqrt{T}\sigma G$$

where $G \sim \mathbf{N}(0, 1)$ is a standard normal distributed random variable. This means that, in the limit, the particle's position $B_T = \lim_{\Delta x, \Delta t \rightarrow 0} X_T$ is normally distributed with law $\mathbf{N}(0, T\sigma^2)$.

This approximation procedure yields for each $t \in [0, T]$ some random variable $B_t \sim \mathbf{N}(0, t\sigma^2)$. More generally

1.1 Definition. Let $(\Omega, \mathcal{A}, \mathbb{P})$ be a probability space. A d -dimensional stochastic process indexed by $I \subset [0, \infty)$ is a family of random variables $X_t : \Omega \rightarrow \mathbb{R}^d, t \in I$. We write $X = (X_t)_{t \in I}$. I is called the *index set* and \mathbb{R}^d the *state space*.

The only requirement of Definition 1.1 is that the $X_t, t \in I$, are $\mathcal{A}/\mathcal{B}(\mathbb{R}^d)$ measurable. This definition is, however, too general to be mathematically useful; more information is needed on $(t, \omega) \mapsto X_t(\omega)$ as a function of two variables. Although the family $(B_t)_{t \in [0, T]}$ satisfies the condition of Definition 1.1, a realistic model of Brownian motion should have at least *continuous trajectories*: For all ω the *sample path* $[0, T] \ni t \mapsto B_t(\omega)$ should be a continuous function.

1.2 Definition. A d -dimensional Brownian motion $B = (B_t)_{t \geq 0}$ is a stochastic process indexed by $[0, \infty)$ taking values in \mathbb{R}^d such that

$$B_0(\omega) = 0 \quad \text{for almost all } \omega; \tag{B0}$$

$$B_{t_n} - B_{t_{n-1}}, \dots, B_{t_1} - B_{t_0} \text{ are independent} \tag{B1}$$

for all $n \geq 0, 0 = t_0 \leq t_1 < t_2 < \dots < t_n < \infty$;

$$B_t - B_s \sim B_{t+h} - B_{s+h} \quad \text{for all } 0 \leq s < t, h \geq -s; \tag{B2}$$

$$B_t - B_s \sim \mathbf{N}(0, t-s)^{\otimes d}, \quad \mathbf{N}(0, t)(dx) = \frac{1}{\sqrt{2\pi t}} \exp\left(-\frac{x^2}{2t}\right) dx; \tag{B3}$$

$$t \mapsto B_t(\omega) \quad \text{is continuous for all } \omega. \tag{B4}$$

We use BM^d as shorthand for d -dimensional Brownian motion.

We will also speak of a Brownian motion if the index set is an interval of the form $[0, T]$ or $[0, \infty)$. We say that $(B_t + x)_{t \geq 0}, x \in \mathbb{R}^d$, is a d -dimensional Brownian motion *started at* x . Frequently we write $B(t, \omega)$ and $B(t)$ instead of $B_t(\omega)$ and B_t ; this should cause no confusion.

By definition, Brownian motion is an \mathbb{R}^d -valued stochastic process starting at the origin (B0) with *independent increments* (B1), *stationary increments* (B2) and continuous paths (B4).

We will see later that this definition is redundant: (B0)–(B3) entail (B4) at least for almost all ω . On the other hand, (B0)–(B2) and (B4) automatically imply that the increment $B(t) - B(s)$ has a (possibly degenerate) Gaussian distribution (this is a consequence of the central limit theorem).

Before we discuss such details, we should settle the question, if there exists a process satisfying the requirements of Definition 1.2.

1.3 Further reading. A good general survey on the history of continuous-time stochastic processes is the paper [29]. The role of Brownian motion in mathematical finance is explained in [6] and [31]. Good references for Brownian motion in physics are [130] and [135], for applications in modelling and engineering [125].



[6] Bachelier, Davis (ed.), Etheridge (ed.): *Louis Bachelier's Theory of Speculation: The Origins of Modern Finance*.

[29] Cohen: The history of noise.

[31] Cootner (ed.): *The Random Character of Stock Market Prices*.

[125] MacDonald: *Noise and Fluctuations*.

[130] Mazo: *Brownian Motion*.

[135] Nelson: *Dynamical Theories of Brownian Motion*.

Problems

Recall that a sequence of random variables $X_n : \Omega \rightarrow \mathbb{R}^d$ converges *weakly* (also: *in distribution* or *in law*) to a random variable X , $X_n \xrightarrow{d} X$ if, and only if for all bounded and continuous functions $f \in \mathcal{C}_b(\mathbb{R}^d)$ $\lim_{n \rightarrow \infty} \mathbb{E} f(X_n) = \mathbb{E} f(X)$. This is equivalent to the convergence of the characteristic functions $\lim_{n \rightarrow \infty} \mathbb{E} e^{i \langle \xi, X_n \rangle} = \mathbb{E} e^{i \langle \xi, X \rangle}$ for all $\xi \in \mathbb{R}^d$.

1. Let $X, Y, X_n, Y_n : \Omega \rightarrow \mathbb{R}$, $n \geq 1$, be random variables.

(a) If, for all $n \geq 1$, $X_n \perp\!\!\!\perp Y_n$ and if $(X_n, Y_n) \xrightarrow{d} (X, Y)$, then $X \perp\!\!\!\perp Y$.

(b) Let $X \perp\!\!\!\perp Y$ such that $X, Y \sim \beta_{1/2} := \frac{1}{2}(\delta_0 + \delta_1)$ are Bernoulli random variables. We set $X_n := X + \frac{1}{n}$ and $Y_n := 1 - X_n$. Then $X_n \xrightarrow{d} X$, $Y_n \xrightarrow{d} Y$, $X_n + Y_n \xrightarrow{d} 1$ but (X_n, Y_n) does not converge weakly to (X, Y) .

(c) Assume that $X_n \xrightarrow{d} X$ and $Y_n \xrightarrow{d} Y$. Is it true that $X_n + Y_n \xrightarrow{d} X + Y$?

2. (Slutsky's Theorem) Let $X_n, Y_n : \Omega \rightarrow \mathbb{R}^d$, $n \geq 1$, be two sequences of random variables such that $X_n \xrightarrow{d} X$ and $X_n - Y_n \xrightarrow{\mathbb{P}} 0$. Then $Y_n \xrightarrow{d} X$.

3. (Slutsky's Theorem 2) Let $X_n, Y_n : \Omega \rightarrow \mathbb{R}, n \geq 1$, be two sequences of random variables.

(a) If $X_n \xrightarrow{d} X$ and $Y_n \xrightarrow{\mathbb{P}} c$, then $X_n Y_n \xrightarrow{d} cX$. Is this still true if $Y_n \xrightarrow{d} c$?

(b) If $X_n \xrightarrow{d} X$ and $Y_n \xrightarrow{\mathbb{P}} 0$, then $X_n + Y_n \xrightarrow{d} X$. Is this still true if $Y_n \xrightarrow{d} 0$?

4. Let $X_n, X, Y : \Omega \rightarrow \mathbb{R}, n \geq 1$, be random variables. If for all $f \in \mathcal{C}_b(\mathbb{R})$ and $g \in \mathcal{B}_b(\mathbb{R})$

$$\lim_{n \rightarrow \infty} \mathbb{E}(f(X_n)g(Y)) = \mathbb{E}(f(X)g(Y))$$

holds, then $(X_n, Y) \xrightarrow{d} (X, Y)$. If $X = \phi(Y)$ for some $\phi \in \mathcal{B}(\mathbb{R})$, then $X_n \xrightarrow{\mathbb{P}} X$.

5. Let $\delta_j, j \geq 1$, be iid Bernoulli random variables with $\mathbb{P}(\delta_j = \pm 1) = 1/2$. We set

$$S_0 := 0, \quad S_n := \delta_1 + \cdots + \delta_n \quad \text{and} \quad X_t^n := \frac{1}{\sqrt{n}} S_{[nt]}.$$

A one-dimensional random variable G is Gaussian if it has the characteristic function $\mathbb{E} e^{i\xi G} = \exp(im\xi - \frac{1}{2}\sigma^2\xi^2)$ with $m \in \mathbb{R}$ and $\sigma \geq 0$. Prove that

(a) $X_t^n \xrightarrow{d} G_t$ where $t > 0$ and G_t is a Gaussian random variable.

(b) $X_t^n - X_s^n \xrightarrow{d} G_{t-s}$ where $t \geq s \geq 0$ and G_u is a Gaussian random variable. Do we have $G_{t-s} = G_t - G_s$?

(c) Let $0 \leq t_1 \leq \cdots \leq t_m, m \geq 1$. Determine the limit as $n \rightarrow \infty$ of the random vector $(X_{t_m}^n - X_{t_{m-1}}^n, \dots, X_{t_2}^n - X_{t_1}^n, X_{t_1}^n)$.

6. Consider the condition

$$\text{for all } s < t \text{ the random variables } \frac{B(t) - B(s)}{\sqrt{t-s}} \text{ are} \quad (\text{B3'})$$

identically distributed, centered and square integrable.

Show that (B0), (B1), (B2), (B3) and (B0), (B1), (B2), (B3') are equivalent.

Hint: If $X \sim Y$, $X \perp\!\!\!\perp Y$ and $X \sim \frac{1}{\sqrt{2}}(X + Y)$, then $X \sim N(0, 1)$, cf. Rényi [153, VI.5 Theorem 2].

Chapter 2

Brownian motion as a Gaussian process

Recall that a one-dimensional random variable Γ is *Gaussian* if it has the characteristic function

$$\mathbb{E} e^{i\xi\Gamma} = e^{im\xi - \frac{1}{2}\sigma^2\xi^2} \quad (2.1)$$

for some real numbers $m \in \mathbb{R}$ and $\sigma \geq 0$. If we differentiate (2.1) two times with respect to ξ and set $\xi = 0$, we see that

$$m = \mathbb{E} \Gamma \quad \text{and} \quad \sigma^2 = \mathbb{V} \Gamma. \quad (2.2)$$

A random vector $\Gamma = (\Gamma_1, \dots, \Gamma_n) \in \mathbb{R}^n$ is *Gaussian*, if $\langle \ell, \Gamma \rangle$ is for every $\ell \in \mathbb{R}^n$ a one-dimensional Gaussian random variable. This is the same as to say that

$$\mathbb{E} e^{i\langle \xi, \Gamma \rangle} = e^{i\mathbb{E}\langle \xi, \Gamma \rangle - \frac{1}{2}\mathbb{V}\langle \xi, \Gamma \rangle}. \quad (2.3)$$

Setting $m = (m_1, \dots, m_n) \in \mathbb{R}^n$ and $\Sigma = (\sigma_{jk})_{j,k=1,\dots,n} \in \mathbb{R}^{n \times n}$ where

$$m_j := \mathbb{E} \Gamma_j \quad \text{and} \quad \sigma_{jk} := \mathbb{E}(\Gamma_j - m_j)(\Gamma_k - m_k) = \text{Cov}(\Gamma_j, \Gamma_k),$$

we can rewrite (2.3) in the following form

$$\mathbb{E} e^{i\langle \xi, \Gamma \rangle} = e^{i\langle \xi, m \rangle - \frac{1}{2}\langle \xi, \Sigma \xi \rangle}. \quad (2.4)$$

We call m the *mean vector* and Σ the *covariance matrix* of Γ .

2.1 The finite dimensional distributions

Let us quickly establish some first consequences of the definition of Brownian motion. To keep things simple, we assume throughout this section that $(B_t)_{t \geq 0}$ is a one-dimensional Brownian motion.

2.1 Proposition. *Let $(B_t)_{t \geq 0}$ be a one-dimensional Brownian motion. Then $B_t, t \geq 0$, are Gaussian random variables with mean 0 and variance t :* Ex. 2.1

$$\mathbb{E} e^{i\xi B_t} = e^{-t\xi^2/2} \quad \text{for all } t \geq 0, \xi \in \mathbb{R}. \quad (2.5)$$

Proof. Set $\phi_t(\xi) = \mathbb{E} e^{i\xi B_t}$. If we differentiate ϕ_t with respect to ξ , and use integration by parts we get

$$\begin{aligned}\phi_t'(\xi) &= \mathbb{E} \left(i B_t e^{i\xi B_t} \right) \stackrel{(B3)}{=} \frac{1}{\sqrt{2\pi t}} \int_{\mathbb{R}} e^{ix\xi} (ix) e^{-x^2/(2t)} dx \\ &= \frac{1}{\sqrt{2\pi t}} \int_{\mathbb{R}} e^{ix\xi} (-it) \frac{d}{dx} e^{-x^2/(2t)} dx \\ &\stackrel{\text{parts}}{=} -t\xi \frac{1}{\sqrt{2\pi t}} \int_{\mathbb{R}} e^{ix\xi} e^{-x^2/(2t)} dx \\ &= -t\xi \phi_t(\xi).\end{aligned}$$

Since $\phi_t(0) = 1$, (2.5) is the unique solution of the differential equation

$$\frac{\phi_t'(\xi)}{\phi_t(\xi)} = -t\xi \quad \square$$

From the elementary inequality $1 \leq \exp([\frac{y}{2} - c]^2)$ we see that $e^{cy} \leq e^{c^2} e^{y^2/4}$ for all $c, y \in \mathbb{R}$. Therefore, $e^{cy} e^{-y^2/2} \leq e^{c^2} e^{-y^2/4}$ is integrable. Considering real and imaginary parts separately, it follows that the integrals in (2.5) converge for all $\xi \in \mathbb{C}$ and define an analytic function.

2.2 Corollary. *A one-dimensional Brownian motion $(B_t)_{t \geq 0}$ has exponential moments of all orders, i. e.*

$$\mathbb{E} e^{\zeta B_t} = e^{t \zeta^2/2} \text{ for all } \zeta \in \mathbb{C}. \quad (2.6)$$

2.3 Moments. Note that for $k = 0, 1, 2, \dots$

$$\mathbb{E}(B_t^{2k+1}) = \frac{1}{\sqrt{2\pi t}} \int_{\mathbb{R}} x^{2k+1} e^{-x^2/(2t)} dx = 0 \quad (2.7)$$

and

$$\begin{aligned}\mathbb{E}(B_t^{2k}) &= \frac{1}{\sqrt{2\pi t}} \int_{\mathbb{R}} x^{2k} e^{-x^2/(2t)} dx \\ &\stackrel{x=\sqrt{2ty}}{=} \frac{2}{\sqrt{2\pi t}} \int_0^\infty (2ty)^k e^{-y} \frac{2t dy}{2\sqrt{2ty}} \\ &= \frac{2^k t^k}{\sqrt{\pi}} \int_0^\infty y^{k-1/2} e^{-y} dy \\ &= t^k \frac{2^k \Gamma(k + 1/2)}{\sqrt{\pi}}\end{aligned} \quad (2.8)$$

where $\Gamma(\cdot)$ denotes Euler's Gamma function. In particular,

$$\mathbb{E} B_t = \mathbb{E} B_t^3 = 0, \quad \forall B_t = \mathbb{E} B_t^2 = t \quad \text{and} \quad \mathbb{E} B_t^4 = 3t^2.$$

2.4 Covariance. For $s, t \geq 0$ we have

$$\text{Cov}(B_s, B_t) = \mathbb{E} B_s B_t = s \wedge t.$$

Indeed, if $s \leq t$,

$$\mathbb{E} B_s B_t = \mathbb{E} (B_s(B_t - B_s)) + \mathbb{E} (B_s^2) \stackrel{(B1)}{=} s = s \wedge t.$$

2.5 Definition. A one-dimensional stochastic process $(X_t)_{t \geq 0}$ is called a *Gaussian process* if all vectors $\Gamma = (X_{t_1}, \dots, X_{t_n})$, $n \geq 1$, $0 \leq t_1 < t_2 < \dots < t_n$ are (possibly degenerate) Gaussian random vectors.

Let us show that a Brownian motion is a Gaussian process.

2.6 Theorem. A one-dimensional Brownian motion $(B_t)_{t \geq 0}$ is a Gaussian process.

For $t_0 := 0 < t_1 < \dots < t_n$, $n \geq 1$, the vector $\Gamma := (B_{t_1}, \dots, B_{t_n})^\top$ is a Gaussian random variable with a strictly positive definite, symmetric covariance matrix $C = (t_j \wedge t_k)_{j,k=1,\dots,n}$ and mean vector $m = 0 \in \mathbb{R}^n$: Ex. 2.2

$$\mathbb{E} e^{i\langle \xi, \Gamma \rangle} = e^{-\frac{1}{2}\langle \xi, C\xi \rangle}. \quad (2.9)$$

Moreover, the probability distribution of Γ is given by

$$\mathbb{P}(\Gamma \in dx) = \frac{1}{(2\pi)^{n/2}} \frac{1}{\sqrt{\det C}} \exp\left(-\frac{1}{2}\langle x, C^{-1}x \rangle\right) dx \quad (2.10a)$$

$$= \frac{1}{(2\pi)^{n/2} \sqrt{\prod_{j=1}^n (t_j - t_{j-1})}} \exp\left(-\frac{1}{2} \sum_{j=1}^n \frac{(x_j - x_{j-1})^2}{t_j - t_{j-1}}\right) dx. \quad (2.10b)$$

Proof. Set $\Delta := (B_{t_1} - B_{t_0}, B_{t_2} - B_{t_1}, \dots, B_{t_n} - B_{t_{n-1}})^\top$ and observe that we can write $B(t_k) - B(t_0) = \sum_{j=1}^k (B_{t_j} - B_{t_{j-1}})$. Thus,

$$\Gamma = \begin{pmatrix} 1 & 0 & \dots & 0 \\ 1 & \ddots & \ddots & \vdots \\ \vdots & \ddots & \ddots & 0 \\ 1 & \dots & 1 & 1 \end{pmatrix} \Delta = M \Delta$$

where $M \in \mathbb{R}^{n \times n}$ is a lower triangular matrix with entries 1 on and below the diagonal. Ex. 2.3

Therefore,

$$\begin{aligned}
 \mathbb{E} \left(\exp \left[i \langle \xi, \Gamma \rangle \right] \right) &= \mathbb{E} \left(\exp \left[i \langle M^\top \xi, \Delta \rangle \right] \right) \\
 &\stackrel{(B1)}{=} \prod_{j=1}^n \mathbb{E} \left(\exp \left[i (B_{t_j} - B_{t_{j-1}}) (\xi_j + \cdots + \xi_n) \right] \right) \\
 &\stackrel{(B2)}{=} \prod_{j=1}^n \exp \left[-\frac{1}{2} (t_j - t_{j-1}) (\xi_j + \cdots + \xi_n)^2 \right].
 \end{aligned} \tag{2.11}$$

Observe that

$$\begin{aligned}
 &\sum_{j=1}^n t_j (\xi_j + \cdots + \xi_n)^2 - \sum_{j=1}^n t_{j-1} (\xi_j + \cdots + \xi_n)^2 \\
 &= t_n \xi_n^2 + \sum_{j=1}^{n-1} t_j ((\xi_j + \cdots + \xi_n)^2 - (\xi_{j+1} + \cdots + \xi_n)^2) \\
 &= t_n \xi_n^2 + \sum_{j=1}^{n-1} t_j \xi_j (\xi_j + 2\xi_{j+1} + \cdots + 2\xi_n) \\
 &= \sum_{j=1}^n \sum_{k=1}^n (t_j \wedge t_k) \xi_j \xi_k.
 \end{aligned} \tag{2.12}$$

This proves (2.9). Since C is strictly positive definite and symmetric, the inverse C^{-1} exists and is again positive definite; both C and C^{-1} have unique positive definite, symmetric square roots. Using the uniqueness of the Fourier transform, the following calculation proves (2.10a):

$$\begin{aligned}
 &\left(\frac{1}{2\pi} \right)^{n/2} \frac{1}{\sqrt{\det C}} \int_{\mathbb{R}^n} e^{i \langle x, \xi \rangle} e^{-\frac{1}{2} \langle x, C^{-1} x \rangle} dx \\
 &\stackrel{y=C^{-1/2}x}{=} \left(\frac{1}{2\pi} \right)^{n/2} \int_{\mathbb{R}^n} e^{i \langle (C^{1/2}y), \xi \rangle} e^{-\frac{1}{2} |y|^2} dy \\
 &= \left(\frac{1}{2\pi} \right)^{n/2} \int_{\mathbb{R}^n} e^{i \langle y, C^{1/2} \xi \rangle} e^{-\frac{1}{2} |y|^2} dy \\
 &= e^{-\frac{1}{2} |C^{1/2} \xi|^2} = e^{-\frac{1}{2} \langle \xi, C \xi \rangle}.
 \end{aligned}$$

Let us finally determine $\langle x, C^{-1} x \rangle$ and $\det C$. Since the entries of Δ are independent $N(0, t_j - t_{j-1})$ distributed random variables we get

$$\mathbb{E} e^{i \langle \xi, \Delta \rangle} \stackrel{(2.5)}{=} \exp \left(-\frac{1}{2} \sum_{j=1}^n (t_j - t_{j-1}) \xi_j^2 \right) = \exp \left(-\frac{1}{2} \langle \xi, D \xi \rangle \right)$$

where $D \in \mathbb{R}^{n \times n}$ is a diagonal matrix with entries $(t_1 - t_0, \dots, t_n - t_{n-1})$. On the other hand, we have

$$e^{-\frac{1}{2}\langle \xi, C\xi \rangle} = \mathbb{E} e^{i\langle \xi, \Gamma \rangle} = \mathbb{E} e^{i\langle \xi, M\Delta \rangle} = \mathbb{E} e^{i\langle M^\top \xi, \Delta \rangle} = e^{-\frac{1}{2}\langle M^\top \xi, DM^\top \xi \rangle}.$$

Thus, $C = MDM^\top$ and, therefore $C^{-1} = (M^\top)^{-1}D^{-1}M^{-1}$. Since M^{-1} is a two-band matrix with entries 1 on the diagonal and -1 on the first sub-diagonal below the diagonal, we see

Ex. 2.4

$$\langle x, C^{-1}x \rangle = \langle M^{-1}x, D^{-1}M^{-1}x \rangle = \sum_{j=1}^n \frac{(x_j - x_{j-1})^2}{t_j - t_{j-1}}$$

as well as $\det C = \det(MDM^\top) = \det D = \prod_{j=1}^n (t_j - t_{j-1})$. This shows (2.10b). \square

The proof of Theorem 2.6 actually characterizes Brownian motion among all Gaussian processes.

2.7 Corollary. *Let $(X_t)_{t \geq 0}$ be a one-dimensional Gaussian process such that the vector $\Gamma = (X_{t_1}, \dots, X_{t_n})^\top$ is a Gaussian random variable with mean 0 and covariance matrix $C = (t_j \wedge t_k)_{j,k=1,\dots,n}$. If $(X_t)_{t \geq 0}$ has continuous sample paths, then $(X_t)_{t \geq 0}$ is a one-dimensional Brownian motion.*

Proof. The properties (B4) and (B0) follow directly from the assumptions; note that $X_0 \sim N(0, 0) = \delta_0$. Set $\Delta = (X_{t_1} - X_{t_0}, \dots, X_{t_n} - X_{t_{n-1}})^\top$ and let $M \in \mathbb{R}^{n \times n}$ be the lower triangular matrix with entries 1 on and below the diagonal. Then, as in Theorem 2.6, $\Gamma = M\Delta$ or $\Delta = M^{-1}\Gamma$ where M^{-1} is a two-band matrix with entries 1 on the diagonal and -1 on the first sub-diagonal below the diagonal. Since Γ is Gaussian, we see

Ex. 2.3

$$\begin{aligned} \mathbb{E} e^{i\langle \xi, \Delta \rangle} &= \mathbb{E} e^{i\langle \xi, M^{-1}\Gamma \rangle} = \mathbb{E} e^{i\langle (M^{-1})^\top \xi, \Gamma \rangle} \stackrel{(2.9)}{=} e^{-\frac{1}{2}\langle (M^{-1})^\top \xi, C(M^{-1})^\top \xi \rangle} \\ &= e^{-\frac{1}{2}\langle \xi, M^{-1}C(M^{-1})^\top \xi \rangle}. \end{aligned}$$

A straightforward calculation shows that $M^{-1}C(M^{-1})^\top$ is just

$$\begin{pmatrix} 1 & & & \\ -1 & \ddots & & \\ & \ddots & \ddots & \\ & & -1 & 1 \end{pmatrix} \begin{pmatrix} t_1 & t_1 & \cdots & t_1 \\ t_1 & t_2 & \cdots & t_2 \\ \vdots & \vdots & \ddots & \vdots \\ t_1 & t_2 & \cdots & t_n \end{pmatrix} \begin{pmatrix} 1 & -1 & & \\ & \ddots & \ddots & \\ & & \ddots & -1 \\ & & & 1 \end{pmatrix} = \begin{pmatrix} t_1 - t_0 & & & \\ & t_2 - t_1 & & \\ & & \ddots & \\ & & & t_n - t_{n-1} \end{pmatrix}.$$

Thus, Δ is a Gaussian random vector with uncorrelated, hence independent, components which are $N(0, t_j - t_{j-1})$ distributed. This proves (B1), (B3) and (B2). \square

2.2 Invariance properties of Brownian motion

The fact that a stochastic process is a Brownian motion is preserved under various operations at the level of the sample paths. Throughout this section $(B_t)_{t \geq 0}$ denotes a d -dimensional Brownian motion.

Ex. 2.10 2.8 Reflection. If $(B_t)_{t \geq 0}$ is a BM^d , so is $(-B_t)_{t \geq 0}$.

2.9 Renewal. Let $(B(t))_{t \geq 0}$ be a Brownian motion and fix some time $a > 0$. Then $(W(t))_{t \geq 0}$, $W(t) := B(t + a) - B(a)$, is again a BM^d . The properties (B0) and (B4) are obvious for $W(t)$. For all $s \leq t$

$$\begin{aligned} W(t) - W(s) &= B(t + a) - B(a) - (B(s + a) - B(a)) \\ &= B(t + a) - B(s + a) \\ &\stackrel{\text{(B3)}}{\sim} \text{N}(0, t - s) \end{aligned}$$

which proves (B3) and (B2) for the process W . Finally, if $t_0 = 0 < t_1 < \dots < t_n$, then

$$W(t_j) - W(t_{j-1}) = B(t_j + a) - B(t_{j-1} + a) \quad \text{for all } j = 1, \dots, n$$

i. e. the independence of the W -increments follows from (B1) for B at the times $t_j + a$, $j = 1, \dots, d$.

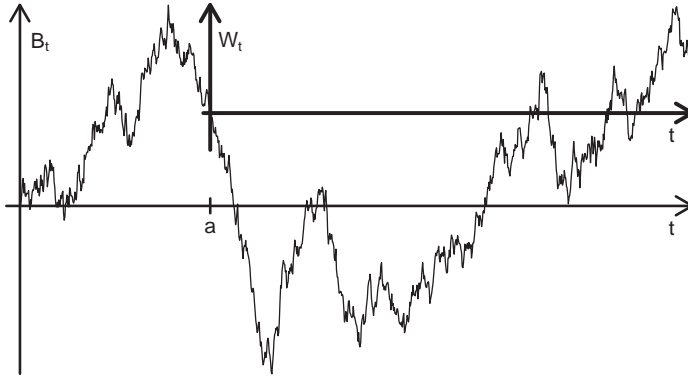


Figure 2.1. Renewal at time a .

A consequence of the independent increments property is that a Brownian motion has no memory. This is the essence of the next lemma.

2.10 Lemma (Markov property of BM). *Let $(B(t))_{t \geq 0}$ be a BM^d and denote by $W(t) := B(t+a) - B(a)$ the shifted Brownian motion constructed in Paragraph 2.9. Then $(B(t))_{0 \leq t \leq a}$ and $(W(t))_{t \geq 0}$ are independent, i.e. the σ -algebras generated by these processes are independent:* Ex. 2.9

$$\sigma(B(t) : 0 \leq t \leq a) = \mathcal{F}_a^B \perp \mathcal{F}_\infty^W := \sigma(W(t) : 0 \leq t < \infty). \quad (2.13)$$

In particular, $B(t) - B(s) \perp \mathcal{F}_s^B$ for all $0 \leq s < t$.

Proof. Let X_0, X_1, \dots, X_n be d -dimensional random variables. Then

$$\sigma(X_j : j = 0, \dots, n) = \sigma(X_0, X_j - X_{j-1} : j = 1, \dots, n). \quad (2.14)$$

Since X_0 and $X_j - X_{j-1}$ are $\sigma(X_j : j = 0, \dots, n)$ measurable, we see the inclusion ‘ \supset ’. For the converse we observe that $X_k = \sum_{j=1}^k (X_j - X_{j-1}) + X_0$, $k = 0, \dots, n$.

Let $0 = s_0 < s_1 < \dots < s_m = a = t_0 < t_1 < \dots < t_n$. By (B1) the random variables

$$B(s_1) - B(s_0), \dots, B(s_m) - B(s_{m-1}), B(t_1) - B(t_0), \dots, B(t_n) - B(t_{n-1})$$

are independent, thus

$$\sigma(B(s_j) - B(s_{j-1}) : j = 1, \dots, m) \perp \sigma(B(t_k) - B(t_{k-1}) : k = 1, \dots, n).$$

Using $W(t_k - t_0) - W(t_{k-1} - t_0) = B(t_k) - B(t_{k-1})$ and $B(0) = W(0) = 0$, we can apply (2.14) to get

$$\sigma(B(s_j) : j = 1, \dots, m) \perp \sigma(W(t_k - t_0) : k = 1, \dots, n)$$

and

$$\bigcup_{\substack{0 < s_1 < \dots < s_m \leq a \\ m \geq 1}} \sigma(B(s_j) : j = 1, \dots, m) \perp \bigcup_{\substack{0 < u_1 < \dots < u_n \\ n \geq 1}} \sigma(W(u_k) : k = 1, \dots, n).$$

The families on the left and right-hand side are \cap -stable generators of \mathcal{F}_a^B and \mathcal{F}_∞^W , respectively, thus $\mathcal{F}_a^B \perp \mathcal{F}_\infty^W$.

Finally, taking $a = s$, we see that $B(t) - B(s) = W(t-s)$ which is \mathcal{F}_∞^W measurable and therefore independent of \mathcal{F}_s^B . □

2.11 Time inversion. Let $(B_t)_{t \geq 0}$ be a Brownian motion and fix some time $a > 0$. Ex. 2.14
Then $W_t := B_{a-t} - B_a$, $t \in [0, a]$, is again a BM^d . This follows as in 2.9 (see Fig. 2.2).

2.12 Scaling. For all $c > 0$ and $t > 0$ we have $B_{ct} \sim c^{1/2} B_t$. In particular, Ex. 2.10
 $(c^{-1/2} B_{ct})_{t \geq 0}$ is again a BM^d .

Denote by $\mathbf{N}(0, t)$ the normal law with mean 0 and variance t . The first assertion follows easily from (B0) and (B3) as $\mathbf{N}(0, ct) = c^{1/2} \mathbf{N}(0, t)$, i.e.

$$B_{ct} \sim \mathbf{N}(0, ct)^{\otimes d} = c^{1/2} \mathbf{N}(0, t)^{\otimes d} \sim c^{1/2} B_t.$$

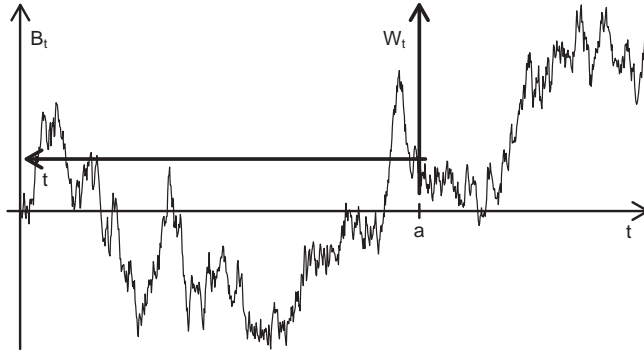


Figure 2.2. Time inversion.

The second claim is now obvious since scaling does not change the independence of the increments or the continuity of the sample paths.

Ex. 2.15 **2.13 Projective reflection at $t = \infty$.** Let $(B_t)_{t \geq 0}$ be a BM^d . Then

Ex. 2.12

$$W(t) := \begin{cases} t B(\frac{1}{t}), & t > 0; \\ 0, & t = 0 \end{cases}$$

is again a BM^d .

It is clear that $(W(t_1), \dots, W(t_n))$ is for $0 < t_1 < \dots < t_n$ a Gaussian random vector. The mean is 0, and the covariance is given by

$$\text{Cov}(W(t_j), W(t_k)) = \text{Cov}\left(t_j B\left(\frac{1}{t_j}\right), t_k B\left(\frac{1}{t_k}\right)\right) = t_j t_k \left(\frac{1}{t_j} \wedge \frac{1}{t_k}\right) = t_j \wedge t_k.$$

As $t \mapsto W_t$, $t > 0$, is continuous, Corollary 2.7 shows that $(W_t)_{t > 0}$ satisfies (B1)–(B4) on $(0, \infty)$. All that remains to be shown is $\lim_{t \downarrow 0} W(t) = W(0) = 0$, i.e. the continuity of the sample paths at $t = 0$.

Note that the limit $\lim_{t \rightarrow 0} W(t, \omega) = 0$ if, and only if,

$$\forall n \geq 1 \quad \exists m \geq 1 \quad \forall r \in \mathbb{Q} \cap (0, \frac{1}{m}] : |r B(\frac{1}{r})| \leq \frac{1}{n}.$$

Thus,

$$\Omega^W := \left\{ \lim_{t \rightarrow 0} W(t) = 0 \right\} = \bigcap_{n \geq 1} \bigcup_{m \geq 1} \bigcap_{r \in \mathbb{Q} \cap (0, 1/m]} \left\{ |W(r)| \leq \frac{1}{n} \right\}.$$

We know already that $(W_t)_{t > 0}$ and $(B_t)_{t > 0}$ have the same finite dimensional distributions. Since Ω^W and the analogously defined set Ω^B are determined by countably many sets of the form $\{|W(r)| \leq \frac{1}{n}\}$ and $\{|B(r)| \leq \frac{1}{n}\}$, we conclude that $\mathbb{P}(\Omega^W) = \mathbb{P}(\Omega^B)$.

Consequently,

$$\mathbb{P}(\Omega^W) = \mathbb{P}(\Omega^B) \stackrel{(B4)}{=} \mathbb{P}(\Omega) = 1.$$

This shows that $(W_t)_{t \geq 0}$ is, on the smaller probability space $(\Omega^W, \mathbb{P}, \Omega^W \cap \mathcal{A})$ equipped with the trace σ -algebra $\Omega^W \cap \mathcal{A}$, a Brownian motion.

2.3 Brownian Motion in \mathbb{R}^d

We will now show that $B_t = (B_t^1, \dots, B_t^d)$ is a BM^d if, and only if, its coordinate processes B_t^j are independent one-dimensional Brownian motions. We call two stochastic processes $(X_t)_{t \geq 0}$ and $(Y_t)_{t \geq 0}$ (defined on the same probability space) *independent*, Ex. 2.18 if the σ -algebras generated by these processes are independent:

$$\mathcal{F}_\infty^X \perp\!\!\!\perp \mathcal{F}_\infty^Y \quad (2.15)$$

where

$$\mathcal{F}_\infty^X := \sigma \left(\bigcup_{n \geq 1} \bigcup_{0 \leq t_1 < \dots < t_n < \infty} \sigma(X(t_1), \dots, X(t_n)) \right). \quad (2.16)$$

Note that the family of sets $\bigcup_n \bigcup_{t_1, \dots, t_n} \sigma(X(t_1), \dots, X(t_n))$ is stable under finite intersections. Therefore, (2.15) follows already if

Ex. 2.16

Ex. 2.17

$$(X(s_1), \dots, X(s_n)) \perp\!\!\!\perp (Y(t_1), \dots, Y(t_m))$$

for all $m, n \geq 1$, $s_1 < \dots < s_m$ and $t_1 < \dots < t_n$. Without loss of generality we can even assume that $m = n$ and $s_j = t_j$ for all j . This follows easily if we take the common refinement of the s_j and t_j .

The following simple characterization of d -dimensional Brownian motion will be very useful for our purposes.

2.14 Lemma. Let $(X_t)_{t \geq 0}$ be a d -dimensional stochastic process. X satisfies (B0)– Ex. 2.19 (B3) if, and only if, for all $n \geq 0$, $0 = t_0 < t_1 < \dots < t_n$, and $\xi_0, \dots, \xi_n \in \mathbb{R}^d$

$$\mathbb{E} \left[\exp \left(i \sum_{j=1}^n \langle \xi_j, X_{t_j} - X_{t_{j-1}} \rangle + i \langle \xi_0, X_{t_0} \rangle \right) \right] = \exp \left(-\frac{1}{2} \sum_{j=1}^n |\xi_j|^2 (t_j - t_{j-1}) \right) \quad (2.17)$$

holds. If X has continuous sample paths, it is a BM^d .

Proof. Assume that X satisfies (B0)–(B3). Since the characteristic function of a Gaussian $\mathbf{N}(0, (t-s)I_d)$ -random variable is $\exp(-\frac{1}{2}(t-s)|\xi|^2)$, we get

$$\begin{aligned} \mathbb{E} \left[\exp \left(i \sum_{j=1}^n \langle \xi_j, X_{t_j} - X_{t_{j-1}} \rangle + i \langle \xi_0, X_{t_0} \rangle \right) \right] &\stackrel{(B1)}{=} \prod_{j=1}^n \mathbb{E}[\exp(i \langle \xi_j, X_{t_j} - X_{t_{j-1}} \rangle)] \\ &\stackrel{(B2)}{=} \prod_{j=1}^n \mathbb{E}[\exp(i \langle \xi_j, X_{t_j - t_{j-1}} \rangle)] \\ &\stackrel{(B3)}{=} \prod_{j=1}^n \exp \left(-\frac{1}{2} (t_j - t_{j-1}) |\xi_j|^2 \right). \end{aligned}$$

Conversely, assume that (2.17) holds. Fix $k \in \{0, 1, \dots, n\}$ and pick $\xi_j = 0$ for all $j \neq k$. Then

$$\mathbb{E} [\exp(i \langle \xi_k, X_{t_k} - X_{t_{k-1}} \rangle)] = e^{-\frac{1}{2} |\xi_k|^2 (t_k - t_{k-1})} = \mathbb{E} [\exp(i \langle \xi_k, X_{t_k - t_{k-1}} \rangle)].$$

This proves (B2), (B3) and, if we take $n = k = 0$, also (B0). Since $X_{t_j} - X_{t_{j-1}} \sim \mathbf{N}(0, (t_j - t_{j-1})I_d)$, (2.17) shows that the increments $X_{t_1} - X_{t_0}, \dots, X_{t_n} - X_{t_{n-1}}$ are independent, i. e. (B1). \square

2.15 Corollary. *Let B be a d -dimensional Brownian motion. Then the coordinate processes B^j , $j = 1, \dots, d$, are independent BM¹.*

Proof. Fix $n \geq 1$ and $t_0 = 0 < t_1 < \dots < t_n < \infty$. Since B is a BM ^{d} , it satisfies (2.17) for all $\xi_1, \dots, \xi_n \in \mathbb{R}^d$. If we take $\xi_j = z_j e_k$ where e_k is the k th unit vector of \mathbb{R}^d , we see that B^k satisfies (2.17) for all $z_1, \dots, z_n \in \mathbb{R}$. Since B^k inherits the continuity of its sample paths from B , Lemma 2.14 shows that B^k is a one-dimensional Brownian motion.

In order to see the independence of the coordinate processes, we have to show that the σ -algebras $(\mathcal{F}_t^{B^1})_{t \geq 0}, \dots, (\mathcal{F}_t^{B^d})_{t \geq 0}$ are independent. As we have seen at the beginning of the section – for two processes, but the argument stays the same for finitely many processes – it is enough to verify that the random vectors

$$(B_{t_1}^k, \dots, B_{t_n}^k), \quad k = 1, 2, \dots, d,$$

are independent for all choices of $n \geq 1$ and $t_0 = 0 \leq t_1 < \dots < t_n$. Since each $B_{t_\ell}^k$, $\ell = 1, \dots, n$, can be written as $B_{t_\ell}^k = \sum_{j=1}^\ell (B_{t_j}^k - B_{t_{j-1}}^k)$, it is enough to show that all increments $(B_{t_j}^k - B_{t_{j-1}}^k)$, $j = 1, \dots, n$ and $k = 1, \dots, d$ are independent. This

follows again with (2.17), since for all $\xi_1, \dots, \xi_n \in \mathbb{R}^d$

$$\begin{aligned}
 \mathbb{E} \left[\exp \left(i \sum_{j=1}^n \sum_{k=1}^d \xi_j^k (B_{t_j}^k - B_{t_{j-1}}^k) \right) \right] &= \mathbb{E} \left[\exp \left(i \sum_{j=1}^n \langle \xi_j, B_{t_j} - B_{t_{j-1}} \rangle \right) \right] \\
 &= \exp \left(-\frac{1}{2} \sum_{j=1}^n |\xi_j|^2 (t_j - t_{j-1}) \right) \\
 &= \exp \left(-\frac{1}{2} \sum_{j=1}^n \sum_{k=1}^d (\xi_j^k)^2 (t_j - t_{j-1}) \right) \\
 &= \prod_{j=1}^n \prod_{k=1}^d \mathbb{E} \exp(i \xi_j^k (B_{t_j}^k - B_{t_{j-1}}^k)). \quad \square
 \end{aligned}$$

The converse of Corollary 2.15 is also true:

2.16 Theorem. B is a BM^d if, and only if, the coordinate processes B^1, \dots, B^d are independent BM^1 .

Proof. Because of Corollary 2.15 it is enough to check that (B_t^1, \dots, B_t^d) is a BM^d provided the B^j are independent one-dimensional Brownian motions. Fix $n \geq 1$, $t_0 = 0 < t_1 < \dots < t_n < \infty$ and $\xi_j^k \in \mathbb{R}$ where $j = 1, \dots, n$ and $k = 1, \dots, d$. By assumption, each B^j satisfies (2.17), i. e. for all $k = 1, \dots, d$

$$\mathbb{E} \left[\exp \left(i \sum_{j=1}^n \xi_j^k (B_{t_j}^k - B_{t_{j-1}}^k) \right) \right] = \exp \left(-\frac{1}{2} \sum_{j=1}^n (\xi_j^k)^2 (t_j - t_{j-1}) \right).$$

Multiply the d resulting equalities; since B^1, \dots, B^d are independent, we get

$$\mathbb{E} \left[\prod_{k=1}^d \exp \left(i \sum_{j=1}^n \xi_j^k (B_{t_j}^k - B_{t_{j-1}}^k) \right) \right] = \exp \left(-\frac{1}{2} \sum_{k=1}^d \sum_{j=1}^n (\xi_j^k)^2 (t_j - t_{j-1}) \right)$$

which is just (2.17) for $B = (B^1, \dots, B^d)$ and $\xi_j = (\xi_j^1, \dots, \xi_j^d) \in \mathbb{R}^d$ (note that we can neglect the t_0 -term as $B(t_0) = 0$). The claim follows from Lemma 2.14. \square

Since a d -dimensional Brownian motion is a vector of independent one-dimensional Brownian motions, many properties of BM^1 also hold for BM^d , and often we need only consider the one-dimensional setting. Things are slightly different, if the coordinate processes are mixed:

Ex. 2.24 **2.17 Definition** (Q -Brownian motion). Let $Q \in \mathbb{R}^{d \times d}$ be a symmetric, positive semi-definite $d \times d$ matrix. A Q -Brownian motion is a d -dimensional process $(X_t)_{t \geq 0}$ satisfying (B0)–(B2), (B4) and

$$X_t - X_s \sim N(0, (t-s)Q) \quad \text{for all } s < t. \quad (\text{QB3})$$

Clearly, BM^d is an I_d -BM. If $(B_t)_{t \geq 0}$ is a BM^d and if $\Sigma \in \mathbb{R}^{d \times d}$, then $X_t := \Sigma B_t$ is a Q -BM with $Q = \Sigma \Sigma^\top$. This follows immediately from

$$\mathbb{E} e^{i\langle \xi, X_t \rangle} = \mathbb{E} e^{i\langle \Sigma^\top \xi, B_t \rangle} = e^{-\frac{1}{2} \langle \Sigma^\top \xi \rangle^2} = e^{-\frac{1}{2} \langle \xi, \Sigma \Sigma^\top \xi \rangle},$$

since the map $x \mapsto \Sigma x$ does not destroy the properties (B0)–(B2) and (B4).

The same calculation shows that any Q -BM $(X_t)_{t \geq 0}$ with a non-degenerate (i. e. strictly positive definite) Q is of the form ΣB_t where Σ is the unique positive definite square root of Q and $(B_t)_{t \geq 0}$ is some BM^d .

Since for Gaussian random vectors ‘independent’ and ‘not correlated’ coincide, it is easy to see that a Q -BM has independent coordinates if, and only if, Q is a diagonal matrix.



2.18 Further reading. The literature on Gaussian processes is quite specialized and technical. Among the most accessible monographs are [129] (with a focus on the Dynkin isomorphism theorem and local times) and [123]. The seminal, and still highly readable, paper [44] is one of the most influential original contributions in the field.

[44] Dudley, R. M.: Sample functions of the Gaussian process.

[123] Lifshits, M. A.: *Gaussian Random Functions*.

[129] Marcus, Rosen: *Markov Processes, Gaussian Processes, and Local Times*.

Problems



1. Show that there exist a random vector (U, V) such that U and V are one-dimensional Gaussian random variables but (U, V) is not Gaussian.

Hint: Try $f(u, v) = g(u)g(v)(1 - \sin u \sin v)$ where $g(u) = (2\pi)^{-1/2} e^{-u^2/2}$.

2. Show that the covariance matrix $C = (t_j \wedge t_k)_{j,k=1,\dots,n}$ appearing in Theorem 2.6 is positive definite.
3. Verify that the matrix M in the proof of Theorem 2.6 and Corollary 2.7 is a lower triangular matrix with entries 1 on and below the diagonal. Show that the inverse matrix M^{-1} is a lower triangular matrix with entries 1 on the diagonal and -1 directly below the diagonal.

4. Let $A, B \in \mathbb{R}^{n \times n}$ be symmetric matrices. If $\langle Ax, x \rangle = \langle Bx, x \rangle$ for all $x \in \mathbb{R}^n$, then $A = B$.
5. Let $(B_t)_{t \geq 0}$ be a BM^1 . Decide which of the following processes are Brownian motions:
 - (a) $X_t := 2B_{t/4}$; (b) $Y_t := B_{2t} - B_t$; (c) $Z_t := \sqrt{t} B_1$. ($t \geq 0$)
6. Let $(B(t))_{t \geq 0}$ be a BM^1 .
 - (a) Find the density of the random vector $(B(s), B(t))$ where $0 < s < t < \infty$.
 - (b) (Brownian bridge) Find the conditional density of the vector $(B(s), B(t))$, $0 < s < t < 1$ under the condition $B(1) = 0$, and use this to determine $\mathbb{E}(B(s)B(t) | B(1) = 0)$.
 - (c) Let $0 < t_1 < t_2 < t_3 < t_4 < \infty$. Determine the conditional density of the bivariate random variable $(B(t_2), B(t_3))$ given that $B(t_1) = B(t_4) = 0$. What is the conditional correlation of $B(t_2)$ and $B(t_3)$?
7. Find the covariance function $C(s, t) := \mathbb{E}(X_s X_t)$, $s, t \geq 0$, of the stochastic process $X_t := B_t^2 - t$, $t \geq 0$, where $(B_t)_{t \geq 0}$ is a BM^1 .
8. (Ornstein–Uhlenbeck process) Let $(B_t)_{t \geq 0}$ be a BM^1 , $\alpha > 0$ and consider the stochastic process $X_t := e^{-\alpha t/2} B_{e^{\alpha t}}$, $t \geq 0$.
 - (a) Determine the mean value and covariance functions $m(t) = \mathbb{E} X_t$ and $C(s, t) = \mathbb{E}(X_s X_t)$, $s, t \geq 0$.
 - (b) Find the probability density of $(X_{t_1}, \dots, X_{t_n})$ where $0 \leq t_1 < \dots < t_n < \infty$.
9. Show that (B1) is equivalent to $B_t - B_s \perp \mathcal{F}_s^B$ for all $0 \leq s < t$.
10. Let $(B_t)_{t \geq 0}$ be a BM^1 and set $\tau_a = \inf\{s \geq 0 : B_s = a\}$ where $a \in \mathbb{R}$. Show that $\tau_a \sim \tau_{-a}$ and $\tau_a \sim a^2 \tau_1$.
11. A one-dimensional stochastic process $(X_t)_{t \geq 0}$ such that $\mathbb{E}(X_t^2) < \infty$ is called *stationary (in the wide sense)* if $m(t) = \mathbb{E} X_t \equiv \text{const.}$ and $C(s, t) = \mathbb{E}(X_s X_t) = g(t - s)$, $0 \leq s \leq t < \infty$ for some even function $g : \mathbb{R} \rightarrow \mathbb{R}$. Which of the following processes is stationary?
 - (a) $W_t = B_t^2 - t$; (b) $X_t = e^{-\alpha t/2} B_{e^{\alpha t}}$;
 - (c) $Y_t = B_{t+h} - B_t$; (d) $Z_t = B_{e^t}$?
12. Let $(B_t)_{t \in [0,1]}$ and $(\beta_t)_{t \in [0,1]}$ be independent one-dimensional Brownian motions. Show that the following process is again a Brownian motion:

$$W_t := \begin{cases} B_t, & \text{if } t \in [0, 1], \\ B_1 + t\beta_{1/t} - \beta_1, & \text{if } t \in (1, \infty). \end{cases}$$

13. Find out whether the processes

$$X(t) := B(e^t) \quad \text{and} \quad X(t) := e^{-t/2} B(e^t) \quad (t \geq 0)$$

have the *no-memory property*, i. e. $\sigma(X(t) : t \leq a) \perp\!\!\!\perp \sigma(X(t+a) - X(a) : t \geq 0)$ for $a > 0$.

14. Prove the time inversion property from Paragraph 2.11.

15. (Strong law of large numbers) Let $(B_t)_{t \geq 0}$ be a BM^1 . Use Paragraph 2.13 to show that $\lim_{t \rightarrow \infty} B_t/t = 0$ a. s. and in mean square sense.

16. Let $(B_t)_{t \geq 0}$ be a BM^1 . Show that $\mathcal{F}_\infty^B = \bigcup_{\substack{J \subset [0, \infty) \\ J \text{ countable}}} \sigma(B(t_j) : t \in J)$.

17. Let $X(t), Y(t)$ be any two stochastic processes. Show that

$$(X(u_1), \dots, X(u_p)) \perp\!\!\!\perp (Y(u_1), \dots, Y(u_p)) \quad \text{for all } u_1 < \dots < u_p, \quad p \geq 1$$

implies

$$(X(s_1), \dots, X(s_n)) \perp\!\!\!\perp (Y(t_1), \dots, Y(t_m)) \\ \text{for all } s_1 < \dots < s_m, \quad t_1 < \dots < t_n, \quad m, n \geq 1.$$

18. Let $(\mathcal{F}_t)_{t \geq 0}$ and $(\mathcal{G}_t)_{t \geq 0}$ be any two filtrations and define $\mathcal{F}_\infty = \sigma(\bigcup_{t \geq 0} \mathcal{F}_t)$ and $\mathcal{G}_\infty = \sigma(\bigcup_{t \geq 0} \mathcal{G}_t)$. Show that $\mathcal{F}_t \perp\!\!\!\perp \mathcal{G}_t$ ($\forall t \geq 0$) if, and only if, $\mathcal{F}_\infty \perp\!\!\!\perp \mathcal{G}_\infty$.

19. Use Lemma 2.14 to show that (all finite dimensional distributions of) a BM^d is invariant under rotations.

20. Let $B_t = (b_t, \beta_t), t \geq 0$, be a two-dimensional Brownian motion.

(a) Show that $W_t := \frac{1}{\sqrt{2}}(b_t + \beta_t)$ is a BM^1 .

(b) Are $X_t := (W_t, \beta_t)$ and $Y_t := \frac{1}{\sqrt{2}}(b_t + \beta_t, b_t - \beta_t), t \geq 0$, two-dimensional Brownian motions?

21. Let $B_t = (b_t, \beta_t), t \geq 0$, be a two-dimensional Brownian motion. For which values of $\lambda, \mu \in \mathbb{R}$ is the process $X_t := \lambda b_t + \mu \beta_t$ a BM^1 ?

22. Let $B_t = (b_t, \beta_t), t \geq 0$, be a two-dimensional Brownian motion. Decide whether for $s > 0$ the process $X_t := (b_t, \beta_{s-t} - \beta_t), 0 \leq t \leq s$ is a two-dimensional Brownian motion.

23. Let $B_t = (b_t, \beta_t), t \geq 0$, be a two-dimensional Brownian motion and $\alpha \in [0, 2\pi)$. Show that $W_t = (b_t \cdot \cos \alpha + \beta_t \cdot \sin \alpha, \beta_t \cdot \cos \alpha - b_t \cdot \sin \alpha)^\top$ is a two-dimensional Brownian motion. Find a suitable d -dimensional generalization of this observation.

24. Let X be a Q -BM. Determine $\text{Cov}(X_t^j, X_s^k)$, the characteristic function of $X(t)$ and the transition probability (density) of $X(t)$.

Chapter 3

Constructions of Brownian motion

There are several different ways to construct Brownian motion and we will present a few of them here. Since a d -dimensional Brownian motion can be obtained from d independent one-dimensional Brownian motions, see Section 2.3, we restrict our attention to the one-dimensional setting. If this is your first encounter with Brownian motion, you should restrict your attention to the Sections 3.1 and 3.2 which contain the most intuitive constructions.

3.1 The Lévy–Ciesielski construction

This approach goes back to Lévy [120, pp. 492–494] but it got its definitive form in the hands of Ciesielski, cf. [26, 27]. The idea is to write the paths $[0, 1] \ni t \mapsto B_t(\omega)$ for (almost) every ω as a random series with respect to a complete orthonormal system (ONS) in the Hilbert space $L^2(dt) = L^2([0, 1], dt)$ with canonical scalar product $\langle f, g \rangle_{L^2} = \int_0^1 f(t)g(t) dt$. Assume that $(\phi_n)_{n \geq 0}$ is any complete ONS and let $(G_n)_{n \geq 0}$ be a sequence of real-valued iid Gaussian $N(0, 1)$ -random variables on the probability space $(\Omega, \mathcal{A}, \mathbb{P})$. Set

$$\begin{aligned} W_N(t) &:= \sum_{n=0}^{N-1} G_n \langle \mathbb{1}_{[0,t)}, \phi_n \rangle_{L^2} \\ &= \sum_{n=0}^{N-1} G_n \int_0^t \phi_n(s) ds. \end{aligned}$$

We want to show that $\lim_{N \rightarrow \infty} W_N(t)$ defines a Brownian motion on $[0, 1]$.

3.1 Lemma. *The limit $W(t) := \lim_{N \rightarrow \infty} W_N(t)$ exists for every $t \in [0, 1]$ in $L^2(\mathbb{P})$ and the process $W(t)$ satisfies (B0)–(B3).* Ex. 3.1

Proof. Using the independence of the $G_n \sim \mathbf{N}(0, 1)$ and Parseval's identity we get for every $t \in [0, 1]$

$$\begin{aligned}
 \mathbb{E}(W_N(t)^2) &= \mathbb{E} \left[\sum_{m,n=0}^{N-1} G_n G_m \langle \mathbb{1}_{[0,t]}, \phi_m \rangle_{L^2} \langle \mathbb{1}_{[0,t]}, \phi_n \rangle_{L^2} \right] \\
 &= \sum_{m,n=1}^{N-1} \underbrace{\mathbb{E}(G_n G_m)}_{=0 \text{ } (n \neq m), \text{ or } =1 \text{ } (n=m)} \langle \mathbb{1}_{[0,t]}, \phi_m \rangle_{L^2} \langle \mathbb{1}_{[0,t]}, \phi_n \rangle_{L^2} \\
 &= \sum_{n=1}^{N-1} \langle \mathbb{1}_{[0,t]}, \phi_n \rangle_{L^2}^2 \xrightarrow[N \rightarrow \infty]{} \langle \mathbb{1}_{[0,t]}, \mathbb{1}_{[0,t]} \rangle_{L^2} = t.
 \end{aligned}$$

This shows that $W(t) = L^2\text{-}\lim_{N \rightarrow \infty} W_N(t)$ exists. An analogous calculation yields for $s < t$ and $u < v$

$$\begin{aligned}
 &\mathbb{E}(W(t) - W(s))(W(v) - W(u)) \\
 &= \sum_{n=0}^{\infty} \langle \mathbb{1}_{[0,t]} - \mathbb{1}_{[0,s]}, \phi_n \rangle_{L^2} \langle \mathbb{1}_{[0,v]} - \mathbb{1}_{[0,u]}, \phi_n \rangle_{L^2} \\
 &= \langle \mathbb{1}_{[s,t]}, \mathbb{1}_{[u,v]} \rangle_{L^2} = \begin{cases} t - s, & [s, t] = [u, v]; \\ 0, & [s, t] \cap [u, v] = \emptyset; \\ (v \wedge t - u \vee s)^+, & \text{in general.} \end{cases}
 \end{aligned}$$

With this calculation we find for all $0 \leq s < t \leq u < v$ and $\xi, \eta \in \mathbb{R}$

$$\begin{aligned}
 &\mathbb{E} \left[\exp(i\xi(W(t) - W(s)) + i\eta(W(v) - W(u))) \right] \\
 &= \lim_{N \rightarrow \infty} \mathbb{E} \left[\exp \left(i \sum_{n=0}^{N-1} (\xi \langle \mathbb{1}_{[s,t]}, \phi_n \rangle + \eta \langle \mathbb{1}_{[u,v]}, \phi_n \rangle) G_n \right) \right] \\
 &\stackrel{\text{iid}}{=} \lim_{N \rightarrow \infty} \prod_{n=0}^{N-1} \mathbb{E} \left[\exp \left((i\xi \langle \mathbb{1}_{[s,t]}, \phi_n \rangle + i\eta \langle \mathbb{1}_{[u,v]}, \phi_n \rangle) G_n \right) \right] \\
 &= \lim_{N \rightarrow \infty} \prod_{n=0}^{N-1} \exp \left(-\frac{1}{2} |\xi \langle \mathbb{1}_{[s,t]}, \phi_n \rangle + \eta \langle \mathbb{1}_{[u,v]}, \phi_n \rangle|^2 \right)
 \end{aligned}$$

(since $\mathbb{E} e^{i\xi G_n} = \exp(-\frac{1}{2} |\xi|^2)$)

$$\begin{aligned}
 &= \lim_{N \rightarrow \infty} \exp \left(-\frac{1}{2} \sum_{n=0}^{N-1} \left[\xi^2 \langle \mathbb{1}_{[s,t]}, \phi_n \rangle^2 + \eta^2 \langle \mathbb{1}_{[u,v]}, \phi_n \rangle^2 \right] \right) \\
 &\quad \times \lim_{N \rightarrow \infty} \exp \left(-\underbrace{\sum_{n=0}^{N-1} \xi \eta \langle \mathbb{1}_{[s,t]}, \phi_n \rangle \langle \mathbb{1}_{[u,v]}, \phi_n \rangle}_{\rightarrow 0} \right) \\
 &= \exp \left(-\frac{1}{2} \xi^2 (t-s) \right) \exp \left(-\frac{1}{2} \eta^2 (v-u) \right).
 \end{aligned}$$

This calculation shows

- $W(t) - W(s) \sim \mathbf{N}(0, t-s)$, if we take $\eta = 0$;
- $W(t-s) \sim \mathbf{N}(0, t-s)$, if we take $\eta = 0, s = 0$ and replace t by $t-s$;
- $W(t) - W(s) \perp\!\!\!\perp W(v) - W(u)$, since ξ, η are arbitrary.

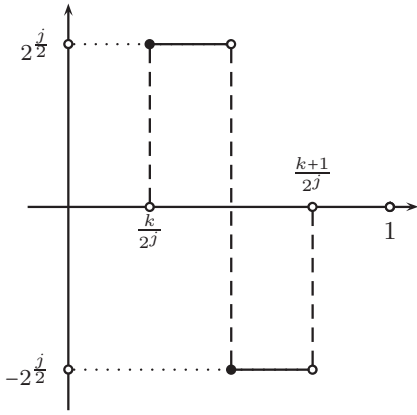
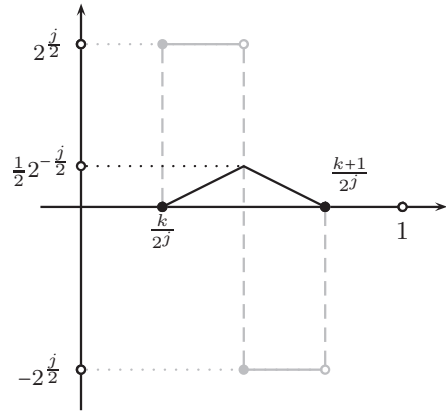
The independence of finitely many increments can be seen in the same way. Since $W(0) = 0$ is obvious, we are done. \square

Since $W(t)$ is an $L^2(\mathbb{P})$ -convergent, hence stochastically convergent, series of independent random variables, we know from classical probability theory that $\lim_{N \rightarrow \infty} W_N(t) = W(t)$ almost surely. But this is not enough to ensure that the path $t \mapsto W(t, \omega)$ is (\mathbb{P} almost surely) continuous. The general theorem due to Itô–Nisio [86] would do the trick, but we prefer to give a direct proof using a special complete orthonormal system.

3.2 The Haar and Schauder systems. We will now give an explicit construction of Brownian motion. For this we need the families of Haar H_{2^j+k} and Schauder S_{2^j+k} functions. For $n = 0$ and $n = 2^j + k, j \geq 0, k = 0, 1, \dots, 2^j - 1$, they are defined as

$$\begin{aligned}
 H_0(t) &= 1, & S_0(t) &= t \\
 H_{2^j+k}(t) &= \begin{cases} +2^{\frac{j}{2}}, & \text{on } \left[\frac{k}{2^j}, \frac{2k+1}{2^{j+1}} \right) \\ -2^{\frac{j}{2}}, & \text{on } \left[\frac{2k+1}{2^{j+1}}, \frac{k+1}{2^j} \right) \\ 0, & \text{otherwise} \end{cases}, & S_{2^j+k}(t) &= \langle \mathbb{1}_{[0,t]}, H_{2^j+k} \rangle_{L^2} \\
 & & &= \int_0^t H_{2^j+k}(s) ds, \\
 & & \text{supp } S_n &= \text{supp } H_n.
 \end{aligned}$$

For $n = 2^j + k$ the graphs of the Haar and Schauder functions are shown in Figures 3.1 and 3.2.

**Figure 3.1.** The Haar functions H_{2^j+k} .**Figure 3.2.** The Schauder functions S_{2^j+k} .

Obviously, $\int_0^1 H_n(t) dt = 0$ for all $n \geq 1$ and

$$H_{2^j+k} H_{2^j+\ell} = S_{2^j+k} S_{2^j+\ell} = 0 \quad \text{for all } j \geq 0, k \neq \ell.$$

The Schauder functions S_{2^j+k} are tent-functions with support $[k2^{-j}, (k+1)2^{-j}]$ and maximal value $\frac{1}{2} 2^{-j/2}$ at the tip of the tent. The Haar functions are also an orthonormal system in $L^2 = L^2([0, 1], dt)$, i. e.

$$\langle H_n, H_m \rangle_{L^2} = \int_0^1 H_n(t) H_m(t) dt = \begin{cases} 1, & m = n \\ 0, & m \neq n, \end{cases} \quad \text{for all } n, m \geq 0,$$

and they are a basis of $L^2([0, 1], ds)$, i. e. a complete orthonormal system, see Theorem A.44 in the appendix.

Ex. 3.2 3.3 Theorem (Lévy 1940; Ciesielski 1959). *There is a probability space $(\Omega, \mathcal{A}, \mathbb{P})$ and a sequence of iid standard normal random variables $(G_n)_{n \geq 0}$ such that*

$$W(t, \omega) = \sum_{n=0}^{\infty} G_n(\omega) \langle \mathbb{1}_{[0,t]}, H_n \rangle_{L^2}, \quad t \in [0, 1], \quad (3.1)$$

is a Brownian motion.

Proof. Let $W(t)$ be as in Lemma 3.1 where we take the Haar functions H_n as orthonormal system in $L^2(dt)$. As probability space $(\Omega, \mathcal{A}, \mathbb{P})$ we use the probability space which supports the (countably many) independent Gaussian random variables $(G_n)_{n \geq 0}$ from the construction of $W(t)$. It is enough to prove that the sample paths

$t \mapsto W(t, \omega)$ are continuous. By definition, the partial sums

$$t \mapsto W_N(t, \omega) = \sum_{n=0}^{N-1} G_n(\omega) \langle \mathbb{1}_{[0,t)}, H_n \rangle_{L^2} = \sum_{n=0}^{N-1} G_n(\omega) S_n(t). \quad (3.2)$$

are continuous for all $N \geq 1$, and it is sufficient to prove that (a subsequence of) $(W_N(t))_{N \geq 0}$ converges uniformly to $W(t)$.

The next step of the proof is similar to the proof of the Riesz–Fischer theorem on the completeness of L^p spaces, see e. g. [169, Theorem 12.7]. Consider the random variable

$$\Delta_j(t) := W_{2^{j+1}}(t) - W_{2^j}(t) = \sum_{k=0}^{2^j-1} G_{2^j+k} S_{2^j+k}(t).$$

If $k \neq \ell$, then $S_{2^j+k} S_{2^j+\ell} = 0$, and we see

$$\begin{aligned} |\Delta_j(t)|^4 &= \sum_{k,\ell,p,q=0}^{2^j-1} G_{2^j+k} G_{2^j+\ell} G_{2^j+p} G_{2^j+q} S_{2^j+k}(t) S_{2^j+\ell}(t) S_{2^j+p}(t) S_{2^j+q}(t) \\ &= \sum_{k=0}^{2^j-1} G_{2^j+k}^4 |S_{2^j+k}(t)|^4 \leq \sum_{k=0}^{2^j-1} G_{2^j+k}^4 2^{-2j}. \end{aligned}$$

Since the right-hand side does not depend on $t \in [0, 1]$ and since $\mathbb{E} G_n^4 = 3$ (use 2.3 for $G_n \sim \mathcal{N}(0, 1)$) we get

$$\mathbb{E} \left[\sup_{t \in [0,1]} |\Delta_j(t)|^4 \right] \leq 3 \sum_{k=0}^{2^j-1} 2^{-2j} = 3 \cdot 2^{-j} \quad \text{for all } j \geq 1.$$

For $n < N$ we find using Minkowski's inequality in $L^4(\mathbb{P})$

$$\begin{aligned} \left\| \sup_{t \in [0,1]} |W_{2^N}(t) - W_{2^n}(t)| \right\|_{L^4} &\leq \left\| \sum_{j=n+1}^N \sup_{t \in [0,1]} \underbrace{|W_{2^j}(t) - W_{2^{j-1}}(t)|}_{=|\Delta_j(t)|} \right\|_{L^4} \\ &\leq \sum_{j=n+1}^N \left\| \sup_{t \in [0,1]} |\Delta_j(t)| \right\|_{L^4} \\ &\leq 3^{1/4} \sum_{j=n+1}^N 2^{-j/4} \xrightarrow{n, N \rightarrow \infty} 0. \end{aligned}$$

Fatou's lemma gives

$$\left\| \lim_{n, N \rightarrow \infty} \sup_{t \in [0,1]} |W_{2^N}(t) - W_{2^n}(t)| \right\|_{L^4} \leq \lim_{n, N \rightarrow \infty} \left\| \sup_{t \in [0,1]} |W_{2^N}(t) - W_{2^n}(t)| \right\|_{L^4} = 0.$$

Ex. 3.3 This shows that there exists a subset $\Omega_0 \subset \Omega$ with $\mathbb{P}(\Omega_0) = 1$ such that



$$\lim_{n, N \rightarrow \infty} \sup_{t \in [0, 1]} |W_{2^N}(t, \omega) - W_{2^n}(t, \omega)| = 0 \quad \text{for all } \omega \in \Omega_0.$$

By the completeness of the space of continuous functions, there is a subsequence of $(W_{2^j}(t, \omega))_{j \geq 1}$ which converges uniformly in $t \in [0, 1]$. Since we know that $W(t, \omega)$ is the limiting function, we conclude that $W(t, \omega)$ inherits the continuity of the $W_{2^j}(t, \omega)$ for all $\omega \in \Omega_0$.

It remains to define Brownian motion everywhere on Ω . We set

$$\tilde{W}(t, \omega) := \begin{cases} W(t, \omega), & \omega \in \Omega_0 \\ 0, & \omega \notin \Omega_0. \end{cases}$$

As $\mathbb{P}(\Omega \setminus \Omega_0) = 0$, Lemma 3.1 remains valid for \tilde{W} which means that $\tilde{W}(t, \omega)$ is a one-dimensional Brownian motion indexed by $[0, 1]$. \square

Strictly speaking, W is not a Brownian motion but has a restriction \tilde{W} which is a Brownian motion.

3.4 Definition. Let $(X_t)_{t \in I}$ and $(Y_t)_{t \in I}$ be two \mathbb{R}^d -valued stochastic processes with the same index set. We call X and Y

Ex. 3.5 *indistinguishable* if they are defined on the same probability space and if

$$\mathbb{P}(X_t = Y_t \quad \forall t \in I) = 1;$$

Ex. 3.4 *modifications* if they are defined on the same probability space and

$$\mathbb{P}(X_t = Y_t) = 1 \quad \text{for all } t \in I;$$

equivalent if they have the same *finite dimensional distributions*, i. e.

$$(X_{t_1}, \dots, X_{t_n}) \sim (Y_{t_1}, \dots, Y_{t_n})$$

for all $t_1, \dots, t_n \in I$, $n \geq 1$ (but X, Y need not be defined on the same probability space).

Now it is easy to construct a real-valued Brownian motion for all $t \geq 0$.

3.5 Corollary. *Let $(W^k(t))_{t \in [0,1]}$, $k \geq 0$, be independent real-valued Brownian motions on the same probability space. Then*

$$B(t) := \begin{cases} W^0(t), & t \in [0, 1); \\ W^0(1) + W^1(t - 1), & t \in [1, 2); \\ \sum_{j=0}^{k-1} W^j(1) + W^k(t - k), & t \in [k, k + 1), k \geq 2. \end{cases} \quad (3.3)$$

is a BM^1 indexed by $t \in [0, \infty)$.

Proof. Let B^k be a copy of the process W from Theorem 3.3 and denote the corresponding probability space by $(\Omega_k, \mathcal{A}_k, \mathbb{P}_k)$. Define, on the product space $(\Omega, \mathcal{A}, \mathbb{P}) = \bigotimes_{k \geq 1} (\Omega_k, \mathcal{A}_k, \mathbb{P}_k)$, $\omega = (\omega_1, \omega_2, \dots)$, processes $W^k(\omega) := B^k(\omega_k)$; by construction, these are independent Brownian motions on $(\Omega, \mathcal{A}, \mathbb{P})$.

Let us check (B0)–(B4) for the process $(B(t))_{t \geq 0}$ defined by (3.3). The properties (B0) and (B4) are obvious.

Let $s < t$ and assume that $s \in [\ell, \ell + 1), t \in [m, m + 1)$ where $\ell \leq m$. Then

$$\begin{aligned} B(t) - B(s) &= \sum_{j=0}^{m-1} W^j(1) + W^m(t - m) - \sum_{j=0}^{\ell-1} W^j(1) - W^\ell(s - \ell) \\ &= \sum_{j=\ell}^{m-1} W^j(1) + W^m(t - m) - W^\ell(s - \ell) \\ &= \begin{cases} W^m(t - m) - W^m(s - m) \sim \text{N}(0, t - s), & \ell = m \\ \underbrace{(W^\ell(1) - W^\ell(s - \ell)) + \sum_{j=\ell+1}^{m-1} W^j(1) + W^m(t - m)}_{\substack{\text{indep} \\ \sim \text{N}(0, 1-s+\ell) \star \text{N}(0, 1) \star m-\ell-1 \star \text{N}(0, t-m) = \text{N}(0, t-s)}}}, & \ell < m. \end{cases} \end{aligned}$$

This proves (B2). We show (B1) only for two increments $B(u) - B(t)$ and $B(t) - B(s)$ where $s < t < u$. As before, $s \in [\ell, \ell + 1), t \in [m, m + 1)$ and $u \in [n, n + 1)$ where $\ell \leq m \leq n$. Then

$$B(u) - B(t) = \begin{cases} W^m(u - m) - W^m(t - m), & m = n \\ (W^m(1) - W^m(t - m)) + \sum_{j=m+1}^{n-1} W^j(1) + W^n(u - n), & m < n. \end{cases}$$

By assumption, the random variables appearing in the representation of the increments

$B(u) - B(t)$ and $B(t) - B(s)$ are independent which means that the increments are independent. The case of finitely many, not necessarily adjacent, increments is similar. \square

3.2 Lévy's original argument

Lévy's original argument is based on interpolation. Let $t_0 < t < t_1$ and assume that $(W_t)_{t \in [0,1]}$ is a Brownian motion. Then

$$G' := W(t) - W(t_0) \quad \text{and} \quad G'' := W(t_1) - W(t)$$

are independent Gaussian random variables with mean 0 and variances $t - t_0$ and $t_1 - t$, respectively. In order to simulate the positions $W(t_0)$, $W(t)$, $W(t_1)$ we could either determine $W(t_0)$ and then G' and G'' independently or, equivalently, simulate first $W(t_0)$ and $W(t_1) - W(t_0)$, and obtain $W(t)$ by interpolation. Let Γ be a further standard normal random variable which is independent of $W(t_0)$ and $W(t_1) - W(t_0)$. Then

$$\begin{aligned} & \frac{(t_1 - t)W(t_0) + (t - t_0)W(t_1)}{t_1 - t_0} + \sqrt{\frac{(t - t_0)(t_1 - t)}{t_1 - t_0}} \Gamma \\ &= W(t_0) + \frac{t - t_0}{t_1 - t_0} (W(t_1) - W(t_0)) + \sqrt{\frac{(t - t_0)(t_1 - t)}{t_1 - t_0}} \Gamma \end{aligned}$$

is – because of the independence of $W(t_0)$, $W(t_1) - W(t_0)$ and Γ – a Gaussian random variable with mean zero and variance t . Thus, we get a random variable with the same distribution as $W(t)$ and so

Ex. 3.7

$$\begin{aligned} \mathbb{P}(W(t) \in \cdot \mid W(t_0) = x_0, W(t_1) = x_1) &= \mathbf{N}(m_t, \sigma_t^2) \\ m_t &= \frac{(t_1 - t)x_0 + (t - t_0)x_1}{t_1 - t_0} \quad \text{and} \quad \sigma_t^2 = \frac{(t - t_0)(t_1 - t)}{t_1 - t_0}. \end{aligned} \tag{3.4}$$

If t is the midpoint of $[t_0, t_1]$, Lévy's method gives the following prescription: Assume we know already $W(t_0)$ and $W(t_1)$. Interpolate these two values linearly, take the midpoint of the line and add to this value (the outcome of the simulation of) an independent Gaussian random variable with mean zero and variance $\sigma^2 = (t - t_0)(t_1 - t)/(t_1 - t_0) = \frac{1}{4}(t_1 - t_0)$. If we start with $t_0 = 0$ and $t_1 = 1$, this allows us to simulate the values

$$W(k2^{-j}) \quad \text{for all } j \geq 0 \quad \text{and} \quad k = 0, 1, \dots, 2^j - 1,$$

and the following picture emerges:

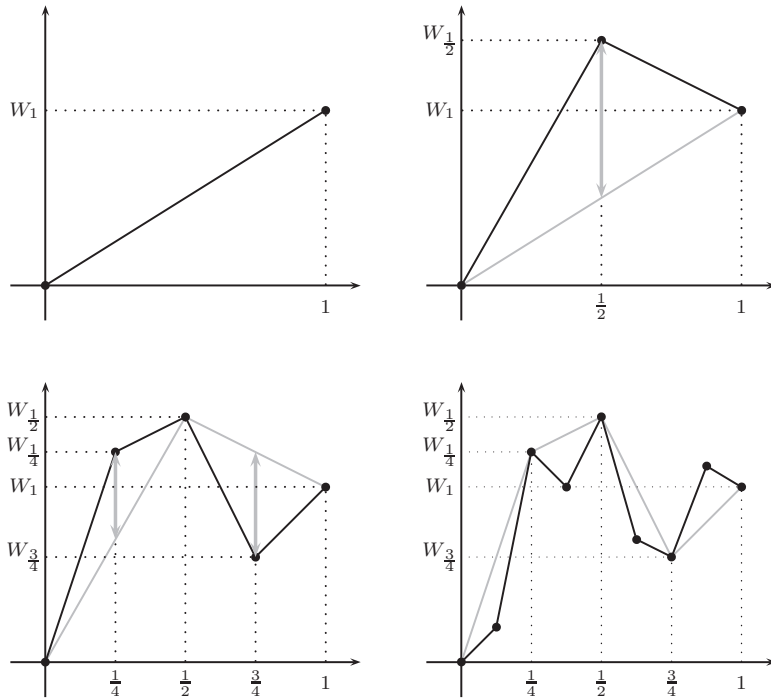


Figure 3.3. The first 4 interpolation steps.

At the dyadic points $t = k2^{-j}$ we get the ‘true’ value of $W(t, \omega)$. This observation shows that the piecewise linear functions are successive approximations of the random function $t \mapsto W(t, \omega)$. This polygonal arc can be expressed by Schauder functions. From (3.2) and the proof of Theorem 3.3 we know that

$$W(t, \omega) = G_0(\omega)S_0(t) + \sum_{j=0}^{\infty} \sum_{k=0}^{2^j-1} G_{2^j+k}(\omega) S_{2^j+k}(t),$$

where $S_n(t)$ are the Schauder functions. For dyadic $t = k/2^j$, the expansion of W_t is finite. Assume that we have already constructed $W(k2^{-j})$ for some fixed $j \geq 1$ and all $k = 0, 1, \dots, 2^j$. Then

$$W(\ell 2^{-j-1}) = \begin{cases} W(k2^{-j}), & \ell = 2k, \\ \frac{1}{2}(W(k2^{-j}) + W((k+1)2^{-j})) + \Gamma_{2^j+k}, & \ell = 2k+1, \end{cases}$$

where $\Gamma_{2^j+k} = G_{2^j+k}S_{2^j+k}((2k+1)2^{-j-1})$ is a Gaussian random variable with mean zero and variance $2^{-j}/4$. This means that each *new* node is constructed by adding

an independent, suitably scaled normal random variable to the linear interpolation of the nodes immediately to the left and right of the new node – and this is exactly Lévy's construction. The first few steps are

$$\begin{aligned}
 1^\circ \quad & W(1, \omega) = G_0(\omega)S_0(1) \\
 2^\circ \quad & W\left(\frac{1}{2}, \omega\right) = \underbrace{G_0(\omega)S_0\left(\frac{1}{2}\right)}_{\text{interpolation}} + \underbrace{G_1(\omega)S_1\left(\frac{1}{2}\right)}_{=\Gamma_1 \sim N(0, 1/4)} \\
 3^\circ \quad & W\left(\frac{1}{4}, \omega\right) = \underbrace{G_0(\omega)S_0\left(\frac{1}{4}\right) + G_1(\omega)S_1\left(\frac{1}{4}\right)}_{\text{interpolation}} + \underbrace{G_2(\omega)S_2\left(\frac{1}{4}\right)}_{=\Gamma_2 \sim N(0, 1/8)} = \underbrace{G_3(\omega)S_3\left(\frac{3}{4}\right)}_{=\Gamma_3 \sim N(0, 1/8)} \\
 & W\left(\frac{3}{4}, \omega\right) = \underbrace{G_0(\omega)S_0\left(\frac{3}{4}\right) + G_1(\omega)S_1\left(\frac{3}{4}\right)}_{\text{interpolation}} + \underbrace{G_2(\omega)S_2\left(\frac{3}{4}\right)}_{=0} + \underbrace{G_3(\omega)S_3\left(\frac{3}{4}\right)}_{=\Gamma_3 \sim N(0, 1/8)} \\
 4^\circ \quad & \dots
 \end{aligned}$$

The 2^j th partial sum, $W_{2^j}(t, \omega)$, is therefore a piecewise linear interpolation of the Brownian path $W_t(\omega)$ at the points $(k2^{-j}, W(k2^{-j}, \omega))$, $k = 0, 1, \dots, 2^j$.

3.6 Theorem (Lévy 1940). *The series*

$$W(t, \omega) := \sum_{n=0}^{\infty} (W_{2^{n+1}}(t, \omega) - W_{2^n}(t, \omega)) + W_1(t, \omega), \quad t \in [0, 1],$$

converges a. s. uniformly. In particular $(W(t))_{t \in [0, 1]}$ is a BM¹.

Proof. Set $\Delta_n(t, \omega) := W_{2^{n+1}}(t, \omega) - W_{2^n}(t, \omega)$. By construction,

$$\Delta_n((2k-1)2^{-n-1}, \omega) = \Gamma_{2^n+(k-1)}(\omega), \quad k = 1, 2, \dots, 2^n,$$

are iid $N(0, 2^{-(n+2)})$ distributed random variables. Therefore,

$$\mathbb{P}\left(\max_{1 \leq k \leq 2^n} |\Delta_n((2k-1)2^{-n-1})| > \frac{x_n}{\sqrt{2^{n+2}}}\right) \leq 2^n \mathbb{P}\left(|\sqrt{2^{n+2}} \Delta_n(2^{-n-1})| > x_n\right)$$

and the right-hand side equals

$$\frac{2 \cdot 2^n}{\sqrt{2\pi}} \int_{x_n}^{\infty} e^{-r^2/2} dr \leq \frac{2^{n+1}}{\sqrt{2\pi}} \int_{x_n}^{\infty} \frac{r}{x_n} e^{-r^2/2} dr = \frac{2^{n+1}}{x_n \sqrt{2\pi}} e^{-x_n^2/2}.$$

Choose $c > 1$ and $x_n := c\sqrt{2n \log 2}$. Then

$$\begin{aligned} \sum_{n=1}^{\infty} \mathbb{P} \left(\max_{1 \leq k \leq 2^n} |\Delta_n((2k-1)2^{-n-1})| > \frac{x_n}{\sqrt{2^{n+2}}} \right) &\leq \sum_{n=1}^{\infty} \frac{2^{n+1}}{c\sqrt{2\pi}} e^{-c^2 \log 2^n} \\ &= \frac{2}{c\sqrt{2\pi}} \sum_{n=1}^{\infty} 2^{-(c^2-1)n} < \infty. \end{aligned}$$

Using the Borel–Cantelli lemma we find a set $\Omega_0 \subset \Omega$ with $\mathbb{P}(\Omega_0) = 1$ such that for every $\omega \in \Omega_0$ there is some $N(\omega) \geq 1$ with

$$\max_{1 \leq k \leq 2^n} |\Delta_n((2k-1)2^{-n-1})| \leq c \sqrt{\frac{n \log 2}{2^{n+1}}} \quad \text{for all } n \geq N(\omega).$$

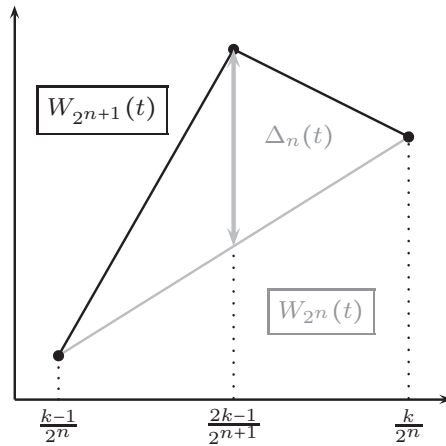


Figure 3.4. The n th interpolation step.

By definition, $\Delta_n(t)$ is the distance between the polygonal arcs $W_{2^{n+1}}(t)$ and $W_{2^n}(t)$; the maximum is attained at one of the midpoints of the intervals $[(k-1)2^{-n}, k2^{-n}]$, $k = 1, \dots, 2^n$, see Figure 3.4. Thus

$$\begin{aligned} \sup_{0 \leq t \leq 1} |W_{2^{n+1}}(t, \omega) - W_{2^n}(t, \omega)| &\leq \max_{1 \leq k \leq 2^n} |\Delta_n((2k-1)2^{-n-1}, \omega)| \\ &\leq c \sqrt{\frac{n \log 2}{2^{n+1}}} \quad \text{for all } n \geq N(\omega), \end{aligned}$$

which shows that the limit

$$W(t, \omega) := \lim_{N \rightarrow \infty} W_{2^N}(t, \omega) = \sum_{n=0}^{\infty} (W_{2^{n+1}}(t, \omega) - W_{2^n}(t, \omega)) + W_1(t, \omega)$$

exists for all $\omega \in \Omega_0$ uniformly in $t \in [0, 1]$. Therefore, $t \mapsto W(t, \omega)$, $\omega \in \Omega_0$, inherits the continuity of the polygonal arcs $t \mapsto W_{2^n}(t, \omega)$. Set

$$\tilde{W}(t, \omega) := \begin{cases} W(t, \omega), & \omega \in \Omega_0, \\ 0, & \omega \notin \Omega_0. \end{cases}$$

By construction, we find for all $0 \leq j \leq k \leq 2^n$

$$\begin{aligned} \tilde{W}(k2^{-n}) - \tilde{W}(j2^{-n}) &= W_{2^n}(k2^{-n}) - W_{2^n}(j2^{-n}) \\ &= \sum_{\ell=j+1}^k (W_{2^n}(\ell2^{-n}) - W_{2^n}((\ell-1)2^{-n})) \\ &\stackrel{\text{iid}}{\sim} \mathbf{N}(0, (k-j)2^{-n}). \end{aligned}$$

Since $t \mapsto \tilde{W}(t)$ is continuous and since the dyadic numbers are dense in $[0, t]$, we conclude that the increments $\tilde{W}(t_j) - \tilde{W}(t_{j-1})$, $0 = t_0 < t_1 < \dots < t_N \leq 1$ are independent $\mathbf{N}(0, t_j - t_{j-1})$ distributed random variables. This shows that $(\tilde{W}(t))_{t \in [0, 1]}$ is a Brownian motion. \square

Ex. 3.7 3.7 Rigorous proof of (3.4). We close this section with a rigorous proof of Lévy's formula (3.4). Observe that for all $0 \leq s < t$ and $\xi \in \mathbb{R}$

$$\begin{aligned} \mathbb{E}(e^{i\xi W(t)} \mid W(s)) &= e^{i\xi W(s)} \mathbb{E}(e^{i\xi(W(t)-W(s))} \mid W(s)) \\ &\stackrel{\text{(B1)}}{=} e^{i\xi W(s)} \mathbb{E}(e^{i\xi(W(t)-W(s))}) \\ &\stackrel{\text{(B2)}}{\stackrel{(2.5)}}{=} e^{-\frac{1}{2}(t-s)\xi^2} e^{i\xi W(s)}. \end{aligned}$$

If we apply this equality to the projectively reflected (cf. 2.13) Brownian motion $(tW(1/t))_{t>0}$, we get

$$\mathbb{E}(e^{i\xi tW(1/t)} \mid W(1/s)) = \mathbb{E}(e^{i\xi tW(1/t)} \mid sW(1/s)) = e^{-\frac{1}{2}(t-s)\xi^2} e^{i\xi sW(1/s)}.$$

Now set $b := \frac{1}{s} > \frac{1}{t} =: a$ and $\eta := \xi/a = \xi t$. Then

$$\begin{aligned} \mathbb{E}(e^{i\eta W(a)} \mid W(b)) &= \exp\left[-\frac{1}{2}a^2\left(\frac{1}{a} - \frac{1}{b}\right)\eta^2\right] \exp\left[i\eta \frac{a}{b}W(b)\right] \\ &= \exp\left[-\frac{1}{2}\frac{a}{b}(b-a)\eta^2\right] \exp\left[i\eta \frac{a}{b}W(b)\right]. \end{aligned} \tag{3.5}$$

This is (3.4) for the case where $(t_0, t, t_1) = (0, a, b)$ and $x_0 = 0$.

Now we fix $0 \leq t_0 < t < t_1$ and observe that $W(s + t_0) - W(t_0)$, $s \geq 0$, is again a Brownian motion. Applying (3.5) yields

$$\mathbb{E} \left(e^{i\eta(W(t)-W(t_0))} \mid W(t_1) - W(t_0) \right) = e^{-\frac{1}{2} \frac{t-t_0}{t_1-t_0} (t_1-t) \eta^2} e^{i\eta \frac{t-t_0}{t_1-t_0} (W(t_1)-W(t_0))}.$$

On the other hand, $W(t_0) \perp\!\!\!\perp (W(t) - W(t_0), W(t_1) - W(t_0))$ and

$$\begin{aligned} & \mathbb{E} \left(e^{i\eta(W(t)-W(t_0))} \mid W(t_1) - W(t_0) \right) \\ &= \mathbb{E} \left(e^{i\eta(W(t)-W(t_0))} \mid W(t_1) - W(t_0), W(t_0) \right) \\ &= \mathbb{E} \left(e^{i\eta(W(t)-W(t_0))} \mid W(t_1), W(t_0) \right) \\ &= e^{-i\eta W(t_0)} \mathbb{E} \left(e^{i\eta W(t)} \mid W(t_1), W(t_0) \right). \end{aligned}$$

This proves

$$\mathbb{E} \left(e^{i\eta W(t)} \mid W(t_1), W(t_0) \right) = e^{-\frac{1}{2} \frac{t-t_0}{t_1-t_0} (t_1-t) \eta^2} e^{i\eta \frac{t-t_0}{t_1-t_0} (W(t_1)-W(t_0))} e^{i\eta W(t_0)},$$

and (3.4) follows.

3.3 Wiener's construction

This is also a series approach, but Wiener used the trigonometric functions $(e^{in\pi t})_{n \in \mathbb{Z}}$ as orthonormal basis for $L^2[0, 1]$. In this case we obtain Brownian motion on $[0, 1]$ as a Wiener-Fourier series

$$W(t, \omega) := \sum_{n=1}^{\infty} \frac{\sin(n\pi t)}{n} G_n(\omega), \quad (3.6)$$

where $(G_n)_{n \geq 0}$ are iid standard normal random variables. Lemma 3.1 remains valid for (3.6) and shows that the series converges in L^2 and that the limit satisfies (B0)–(B3); only the proof that the limiting process is continuous, Theorem 3.3, needs some changes.

Proof of the continuity of (3.6). Let

$$W_N(t, \omega) := \sum_{n=1}^N \frac{\sin(n\pi t)}{n} G_n(\omega).$$

It is enough to show that $(W_{2^j})_{j \geq 1}$ is a Cauchy sequence in $L^2(\mathbb{P})$ uniformly for all $t \in [0, 1]$. Set

$$\Delta_j(t) := W_{2^{j+1}}(t) - W_{2^j}(t).$$

Using $|\operatorname{Im} z| \leq |z|$ for $z \in \mathbb{C}$, we see

$$|\Delta_j(t)|^2 = \left(\sum_{k=2^j+1}^{2^{j+1}} \frac{\sin(k\pi t)}{k} G_k \right)^2 \leq \left| \sum_{k=2^j+1}^{2^{j+1}} \frac{e^{ik\pi t}}{k} G_k \right|^2,$$

and since $|z|^2 = z\bar{z}$ we get

$$\begin{aligned} |\Delta_j(t)|^2 &\leq \sum_{k=2^j+1}^{2^{j+1}} \sum_{\ell=2^j+1}^{2^{j+1}} \frac{e^{ik\pi t} e^{-i\ell\pi t}}{k\ell} G_k G_\ell \\ &= \sum_{k=2^j+1}^{2^{j+1}} \frac{G_k^2}{k^2} + 2 \sum_{k=2^j+1}^{2^{j+1}} \sum_{\ell=2^j+1}^{k-1} \frac{e^{ik\pi t} e^{-i\ell\pi t}}{k\ell} G_k G_\ell \\ &\stackrel{m=k-\ell}{=} \sum_{k=2^j+1}^{2^{j+1}} \frac{G_k^2}{k^2} + 2 \sum_{m=1}^{2^j-1} \sum_{\ell=2^j+1}^{2^{j+1}-m} \frac{e^{im\pi t}}{\ell(\ell+m)} G_\ell G_{\ell+m} \\ &\leq \sum_{k=2^j+1}^{2^{j+1}} \frac{G_k^2}{k^2} + 2 \sum_{m=1}^{2^j-1} \left| \sum_{\ell=2^j+1}^{2^{j+1}-m} \frac{G_\ell G_{\ell+m}}{\ell(\ell+m)} \right|. \end{aligned}$$

Note that the right-hand side is independent of t . On both sides we can therefore take the supremum over all $t \in [0, 1]$ and then the mean value $\mathbb{E}(\dots)$. From Jensen's inequality, $\mathbb{E}|Z| \leq \sqrt{\mathbb{E}|Z|^2}$, we conclude

$$\begin{aligned} \mathbb{E} \left(\sup_{t \in [0,1]} |\Delta_j(t)|^2 \right) &\leq \sum_{k=2^j+1}^{2^{j+1}} \frac{\mathbb{E} G_k^2}{k^2} + 2 \sum_{m=1}^{2^j-1} \mathbb{E} \left| \sum_{\ell=2^j+1}^{2^{j+1}-m} \frac{G_\ell G_{\ell+m}}{\ell(\ell+m)} \right| \\ &\leq \sum_{k=2^j+1}^{2^{j+1}} \frac{\mathbb{E} G_k^2}{k^2} + 2 \sum_{m=1}^{2^j-1} \sqrt{\mathbb{E} \left(\sum_{\ell=2^j+1}^{2^{j+1}-m} \frac{G_\ell G_{\ell+m}}{\ell(\ell+m)} \right)^2}. \end{aligned}$$

If we expand the square we get expressions of the form $\mathbb{E}(G_\ell G_{\ell+m} G_{\ell'} G_{\ell'+m})$; these expectations are zero whenever $\ell \neq \ell'$. Thus,

$$\begin{aligned} \mathbb{E} \left(\sup_{t \in [0,1]} |\Delta_j(t)|^2 \right) &= \sum_{k=2^j+1}^{2^{j+1}} \frac{1}{k^2} + 2 \sum_{m=1}^{2^j-1} \sqrt{\sum_{\ell=2^j+1}^{2^{j+1}-m} \mathbb{E} \left(\frac{G_\ell G_{\ell+m}}{\ell(\ell+m)} \right)^2} \\ &= \sum_{k=2^j+1}^{2^{j+1}} \frac{1}{k^2} + 2 \sum_{m=1}^{2^j-1} \sqrt{\sum_{\ell=2^j+1}^{2^{j+1}-m} \frac{1}{\ell^2(\ell+m)^2}} \end{aligned}$$

$$\begin{aligned}
&\leq \sum_{k=2^j+1}^{2^{j+1}} \frac{1}{(2^j)^2} + 2 \sum_{m=1}^{2^j-1} \sqrt{\sum_{\ell=2^j+1}^{2^{j+1}-m} \frac{1}{(2^j)^2(2^j)^2}} \\
&\leq 2^j \cdot 2^{-2j} + 2 \cdot 2^j \cdot \sqrt{2^j \cdot 2^{-4j}} \\
&\leq 3 \cdot 2^{-j/2}.
\end{aligned}$$

From this point onwards we can follow the proof of Theorem 3.3: For $n < N$ we find using Minkowski's inequality in $L^2(\mathbb{P})$

$$\begin{aligned}
\left\| \sup_{t \in [0,1]} |W_{2^N}(t) - W_{2^n}(t)| \right\|_{L^2} &\leq \left\| \sum_{j=n+1}^N \sup_{t \in [0,1]} \underbrace{|W_{2^j}(t) - W_{2^{j-1}}(t)|}_{=|\Delta_j(t)|} \right\|_{L^2} \\
&\leq \sum_{j=n+1}^N \left\| \sup_{t \in [0,1]} |\Delta_j(t)| \right\|_{L^2} \\
&\leq 3^{1/2} \sum_{j=n+1}^N 2^{-j/4} \xrightarrow{n, N \rightarrow \infty} 0.
\end{aligned}$$

Fatou's lemma gives

$$\left\| \lim_{n, N \rightarrow \infty} \sup_{t \in [0,1]} |W_{2^N}(t) - W_{2^n}(t)| \right\|_{L^2} \leq \lim_{n, N \rightarrow \infty} \left\| \sup_{t \in [0,1]} |W_{2^N}(t) - W_{2^n}(t)| \right\|_{L^2} = 0.$$

This shows that there exists a subset $\Omega_0 \subset \Omega$ with $\mathbb{P}(\Omega_0) = 1$ such that

$$\lim_{n, N \rightarrow \infty} \sup_{t \in [0,1]} |W_{2^N}(t, \omega) - W_{2^n}(t, \omega)| = 0 \quad \text{for all } \omega \in \Omega_0.$$

Ex. 3.3



By the completeness of the space of continuous functions, there is a subsequence of $(W_{2^j}(t, \omega))_{j \geq 1}$ which converges uniformly in $t \in [0, 1]$. Since we know that $W(t, \omega)$ is the limiting function, we conclude that $W(t, \omega)$ inherits the continuity of the $W_{2^j}(t, \omega)$ for all $\omega \in \Omega_0$.

It remains to define the Brownian motion everywhere on Ω . We set

$$\tilde{W}(t, \omega) := \begin{cases} W(t, \omega), & \omega \in \Omega_0 \\ 0, & \omega \notin \Omega_0. \end{cases}$$

Since $\mathbb{P}(\Omega \setminus \Omega_0) = 0$, Lemma 3.1 remains valid for \tilde{W} which means that $\tilde{W}(t, \omega)$ is a one-dimensional Brownian motion indexed by $[0, 1]$. \square

3.4 Donsker's construction

Donsker's invariance theorem shows that Brownian motion is a limit of linearly interpolated random walks – pretty much in the way we have started the discussion in Chapter 1. As before, the difficult point is to prove the sample continuity of the limiting process.

Let, on a probability space $(\Omega, \mathcal{A}, \mathbb{P})$, $\epsilon_n, n \geq 1$, be iid Bernoulli random variables such that $\mathbb{P}(\epsilon_1 = 1) = \mathbb{P}(\epsilon_1 = -1) = \frac{1}{2}$. Then

$$S_n := \epsilon_1 + \cdots + \epsilon_n$$

is a simple random walk. Interpolate linearly and apply Gaussian scaling

$$S^n(t) := \frac{1}{\sqrt{n}} (S_{[nt]} - (nt - [nt])\epsilon_{[nt]+1}), \quad t \in [0, 1].$$

In particular, $S^n(\frac{j}{n}) = \frac{1}{\sqrt{n}} S_j$. If $j = j(n)$ and $j/n = s = \text{const.}$, the central limit theorem shows that $S^n(\frac{j}{n}) = \sqrt{s} S_j / \sqrt{j} \xrightarrow{d} \sqrt{s} G$ as $n \rightarrow \infty$ where G is a standard normal random variable. Moreover, with $s = j/n$ and $t = k/n$, the increment $S^n(t) - S^n(s) = (S_k - S_j)/\sqrt{n}$ is independent of $\epsilon_1, \dots, \epsilon_j$, and therefore of all earlier increments of the same form. Moreover,

$$\mathbb{E}(S^n(t) - S^n(s)) = 0 \quad \text{and} \quad \mathbb{V}(S^n(t) - S^n(s)) = \frac{k-j}{n} = t - s;$$

in the limit we get a Gaussian increment with mean zero and variance $t - s$. Since independence and stationarity of the increments are distributional properties, they are inherited by the limiting process – which we will denote by $(B_t)_{t \in [0,1]}$. We have seen that $(B_q)_{q \in [0,1] \cap \mathbb{Q}}$ would have the properties (B0)–(B3) and it qualifies as a candidate for Brownian motion. If it had continuous sample paths, (B0)–(B3) would hold not only for rational times but for all $t \geq 0$. That the limit exists and is uniform in t is the essence of Donsker's invariance principle.

3.8 Theorem (Donsker 1951). *Let $(B(t))_{t \in [0,1]}$ be a one-dimensional Brownian motion, $(S^n(t))_{t \in [0,1]}$, $n \geq 1$, be as above and $\Phi : \mathcal{C}([0, 1], \mathbb{R}) \rightarrow \mathbb{R}$ a uniformly continuous bounded functional. Then*

$$\lim_{n \rightarrow \infty} \mathbb{E} \Phi(S^n(\cdot)) = \mathbb{E} \Phi(B(\cdot)),$$

i.e. $S^n(\cdot)$ converges weakly to $B(\cdot)$.

We will give a proof of Donsker's theorem in Chapter 13, Theorem 13.5. Since our proof relies on the existence of a Brownian motion, we cannot use it to construct $(B_t)_{t \geq 0}$. Nevertheless, it is an important result for the intuitive understanding of a

Brownian motion as limit of random walks as well as for (simple) proofs of various limit theorems.

3.5 The Bachelier–Kolmogorov point of view

The starting point of this construction is the observation that the finite dimensional marginal distributions of a Brownian motion are Gaussian random variables. More precisely, for any number of times $t_0 = 0 < t_1 < \dots < t_n$, $t_j \in I$, and all Borel sets $A_1, \dots, A_n \in \mathcal{B}(\mathbb{R})$ the finite dimensional distributions

$$p_{t_1, \dots, t_n}(A_1 \times \dots \times A_n) = \mathbb{P}(B_{t_1} \in A_1, \dots, B_{t_n} \in A_n)$$

are mean-zero normal laws with covariance matrix $C = (t_j \wedge t_k)_{j,k=1, \dots, n}$. From Theorem 2.6 we know that they are given by

$$\begin{aligned} p_{t_1, \dots, t_n}(A_1 \times \dots \times A_n) &= \frac{1}{(2\pi)^{n/2}} \frac{1}{\sqrt{\det C}} \int_{A_1 \times \dots \times A_n} \exp\left(-\frac{1}{2}\langle x, C^{-1}x \rangle\right) dx \\ &= \frac{1}{(2\pi)^{n/2} \sqrt{\prod_{j=1}^n (t_j - t_{j-1})}} \int_{A_1 \times \dots \times A_n} \exp\left(-\frac{1}{2} \sum_{j=1}^n \frac{(x_j - x_{j-1})^2}{t_j - t_{j-1}}\right) dx. \end{aligned}$$

We will characterize a stochastic process in terms of its finite dimensional distributions – and we will discuss this approach in Chapter 4 below. The idea is to identify a stochastic process with an infinite dimensional measure \mathbb{P} on the space of all sample paths Ω such that the finite dimensional projections of \mathbb{P} are exactly the finite dimensional distributions.

3.9 Further reading. Yet another construction of Brownian motion, using interpolation arguments and Bernstein polynomials, can be found in the well-written paper [108]. Donsker’s theorem, without assuming the existence of BM, is e. g. proved in the classic monograph [10]; in a more general context it is contained in [45], the presentation in [99] is probably the easiest to read. Ciesielski’s construction became popular through his very influential Aarhus lecture notes [27], one of the first books on Brownian motion after Wiener [187] and Lévy [121]. Good sources on random Fourier series and wavelet expansions are [95] and [146]. Constructions of Brownian motion in infinite dimensions, e. g. cylindrical Brownian motions etc., can be found in [150] and [34], but this needs a good knowledge of functional analysis.



- [10] Billingsley: *Convergence of Probability Measures*.
- [27] Ciesielski: *Lectures on Brownian Motion, Heat Conduction and Potential Theory*.
- [34] DaPrato, Zabczyk: *Stochastic Equations in Infinite Dimensions*.
- [45] Dudley: *Uniform Central Limit Theorems*.
- [95] Kahane: *Some Random Series of Functions*.
- [99] Karatzas, Shreve: *Brownian Motion and Stochastic Calculus*.
- [108] Kowalski: Bernstein Polynomials and Brownian Motion.
- [146] Pinsky: *Introduction to Fourier Analysis and Wavelets*.
- [150] Prévôt, Röckner: *A Concise Course on Stochastic Partial Differential Equations*.

Problems

1. Use the Lévy–Ciesielski representation $B(t) = \sum_{n=0}^{\infty} G_n S_n(t)$, $t \in [0, 1]$, to obtain a series representation for the random variable $X := \int_0^1 B(t) dt$ and find the distribution of X .
2. Let $\phi_n = H_n$, $n = 0, 1, 2, \dots$, be the Haar functions. The following steps show that the full sequence $W_N(t, \omega) := \sum_{n=0}^{N-1} G_n(\omega) S_n(t)$, $N = 1, 2, \dots$, converges uniformly for \mathbb{P} -almost all $\omega \in \Omega$.
 - (a) Let $(a_n)_{n \geq 0} \subset \mathbb{R}$ be a sequence satisfying $a_n = O(n^\epsilon)$ for some $\epsilon \in (0, 1/2)$. Show that $\sum_{n=0}^{\infty} a_n S_n(t)$ converges absolutely and uniformly in $t \in [0, 1]$.
 - (b) Let $(G_n)_{n \geq 0}$ be a sequence of iid $N(0, 1)$ random variables. Show that almost surely $|G_n| = O(\sqrt{\log n})$.
Hint: Use the estimate $\mathbb{P}(|G_1| > x) \leq 2(2\pi)^{-1/2} x^{-1} e^{-x^2/2}$, $x > 0$, and the Borel–Cantelli lemma.
3. Let (S, d) be a complete metric space equipped with the σ -algebra $\mathcal{B}(S)$ of its Borel sets. Assume that $(X_n)_{n \geq 1}$ is a sequence of S -valued random variables such that

$$\lim_{n \rightarrow \infty} \sup_{m \geq n} \mathbb{E} (d(X_n, X_m)^p) = 0$$

for some $p \in [1, \infty)$. Show that there is a subsequence $(n_k)_{k \geq 1}$ and a random variable X such that $\lim_{k \rightarrow \infty} X_{n_k} = X$ almost surely.

4. Let $(X_t)_{t \geq 0}$ and $(Y_t)_{t \geq 0}$ be two stochastic processes which are modifications of each other. Show that they have the same finite dimensional distributions.
5. Let $(X_t)_{t \in I}$ and $(Y_t)_{t \in I}$ be two processes with the same index set $I \subset [0, \infty)$ and state space. Show that

$$X, Y \text{ indistinguishable} \implies X, Y \text{ modifications} \implies X, Y \text{ equivalent}.$$

Assume that the processes are defined on the same probability space and $t \mapsto X_t$ and $t \mapsto Y_t$ are right-continuous (or that I is countable). In this case, the reverse implications hold, too.

6. Let $(B_t)_{t \geq 0}$ be a real-valued stochastic process with exclusively continuous sample paths. Assume that $(B_q)_{q \in \mathbb{Q}}$ satisfies (B0)–(B3). Show that $(B_t)_{t \geq 0}$ is a BM.
7. Give a direct proof of the formula (3.4) using the joint probability distribution $(W(t_0), W(t), W(t_1))$ of the Brownian motion $W(t)$.

Chapter 4

The canonical model

- Ex. 4.1 Often we encounter the statement ‘Let X be a random variable with law μ ’ where μ is an *a priori* given probability distribution. There is no reference to the underlying probability space $(\Omega, \mathcal{A}, \mathbb{P})$, and actually the nature of this space is not important: While X is defined by the probability distribution, there is considerable freedom in the choice of our model $(\Omega, \mathcal{A}, \mathbb{P})$. The same situation is true for stochastic processes: Brownian motion is defined by distributional properties and our construction of BM, see e. g. Theorem 3.3, not only furnished us with a process but also a suitable probability space.
- Ex. 4.2 By definition, a d -dimensional stochastic process $(X(t))_{t \in I}$ is a family of \mathbb{R}^d -valued random variables on the space $(\Omega, \mathcal{A}, \mathbb{P})$. Alternatively, we may understand a process as a map $(t, \omega) \mapsto X(t, \omega)$ from $I \times \Omega$ to \mathbb{R}^d or as a map $\omega \mapsto \{t \mapsto X(t, \omega)\}$ from Ω into the space $(\mathbb{R}^d)^I = \{w, w : I \rightarrow \mathbb{R}^d\}$. If we go one step further and identify Ω with (a subset of) $(\mathbb{R}^d)^I$, we get the so-called *canonical model*. Of course, it is a major task to identify the correct σ -algebra and measure in $(\mathbb{R}^d)^I$. For Brownian motion it is enough to consider the subspace $\mathcal{C}_{(0)}[0, \infty) \subset (\mathbb{R}^d)^{[0, \infty)}$,

$$\mathcal{C}_{(0)} := \mathcal{C}_{(0)}[0, \infty) := \{w : [0, \infty) \rightarrow \mathbb{R}^d : w \text{ is continuous and } w(0) = 0\}.$$

In the first part of this chapter we will see how we can obtain a canonical model for BM defined on the space of continuous functions; in the second part we discuss Kolmogorov’s construction of stochastic processes with given finite dimensional distributions.

4.1 Wiener measure

Let $I = [0, \infty)$ and denote by $\pi_t : (\mathbb{R}^d)^I \rightarrow \mathbb{R}^d$, $w \mapsto w(t)$, the canonical projection onto the t th coordinate. The natural σ -algebra on the infinite product $(\mathbb{R}^d)^I$ is the product σ -algebra

$$\mathcal{B}^I(\mathbb{R}^d) = \sigma \left\{ \pi_t^{-1}(B) : B \in \mathcal{B}(\mathbb{R}^d), t \in I \right\} = \sigma \{ \pi_t : t \in I \};$$

this shows that $\mathcal{B}^I(\mathbb{R}^d)$ is the smallest σ -algebra which makes all projections π_t measurable. Since Brownian motion has exclusively continuous sample paths, it is natural

to replace $(\mathbb{R}^d)^I$ by $\mathcal{C}_{(o)}$. Unfortunately, cf. Corollary 4.6, $\mathcal{C}_{(o)}$ is not contained in $\mathcal{B}^I(\mathbb{R}^d)$; therefore, we have to consider the trace σ -algebra

$$\mathcal{C}_{(o)} \cap \mathcal{B}^I(\mathbb{R}^d) = \sigma(\pi_t|_{\mathcal{C}_{(o)}} : t \in I).$$

If we equip $\mathcal{C}_{(o)}$ with the metric of locally uniform convergence,

$$\rho(w, v) = \sum_{n=1}^{\infty} \left(1 \wedge \sup_{0 \leq t \leq n} |w(t) - v(t)| \right) 2^{-n},$$

$\mathcal{C}_{(o)}$ becomes a complete separable metric space. Denote by \mathcal{O}_ρ the topology induced by ρ and consider the Borel σ -algebra $\mathcal{B}(\mathcal{C}_{(o)}) := \sigma(\mathcal{O}_\rho)$ on $\mathcal{C}_{(o)}$.

4.1 Lemma. *We have $\mathcal{C}_{(o)} \cap \mathcal{B}^I(\mathbb{R}^d) = \mathcal{B}(\mathcal{C}_{(o)})$.*

Proof. Under the metric ρ the map $\mathcal{C}_{(o)} \ni w \mapsto \pi_t(w)$ is (Lipschitz) continuous for every t , hence measurable with respect to the topological σ -algebra $\mathcal{B}(\mathcal{C}_{(o)})$. This means that

$$\sigma(\pi_t|_{\mathcal{C}_{(o)}}) \subset \mathcal{B}(\mathcal{C}_{(o)}) \quad \forall t \in I, \quad \text{and so,} \quad \sigma(\pi_t|_{\mathcal{C}_{(o)}} : t \in I) \subset \mathcal{B}(\mathcal{C}_{(o)}).$$

Conversely, we have for $v, w \in \mathcal{C}_{(o)}$

$$\begin{aligned} \rho(v, w) &= \sum_{n=1}^{\infty} \left(1 \wedge \sup_{t \in [0, n] \cap \mathbb{Q}} |w(t) - v(t)| \right) 2^{-n} \\ &= \sum_{n=1}^{\infty} \left(1 \wedge \sup_{t \in [0, n] \cap \mathbb{Q}} |\pi_t(w) - \pi_t(v)| \right) 2^{-n}. \end{aligned} \tag{4.1}$$

Thus, $(v, w) \mapsto \rho(v, w)$ is measurable with respect to $\sigma(\pi_t|_{\mathcal{C}_{(o)}} : t \in \mathbb{Q}^+)$ and, consequently, with respect to $\mathcal{C}_{(o)} \cap \mathcal{B}^I(\mathbb{R}^d)$. Since $(\mathcal{C}_{(o)}, \rho)$ is separable, there exists a countable dense set $\mathcal{D} \subset \mathcal{C}_{(o)}$ and it is easy to see that every $U \in \mathcal{O}_\rho$ can be written as a countable union of the form

$$U = \bigcup_{\substack{\mathbb{B}^{(\rho)}(w, r) \subset U \\ w \in \mathcal{D}, r \in \mathbb{Q}^+}} \mathbb{B}^{(\rho)}(w, r)$$

where $\mathbb{B}^{(\rho)}(w, r) = \{v \in \mathcal{C}_{(o)} : \rho(v, w) < r\}$ is an open ball in the metric ρ . Since $\mathbb{B}^{(\rho)}(w, r) \in \mathcal{C}_{(o)} \cap \mathcal{B}^I(\mathbb{R}^d)$, we get

$$\mathcal{O}_\rho \subset \mathcal{C}_{(o)} \cap \mathcal{B}^I(\mathbb{R}^d) \quad \text{and therefore} \quad \mathcal{B}(\mathcal{C}_{(o)}) = \sigma(\mathcal{O}_\rho) \subset \mathcal{C}_{(o)} \cap \mathcal{B}^I(\mathbb{R}^d). \quad \square$$

Let $(B_t)_{t \geq 0}$ be a Brownian motion on some probability space $(\Omega, \mathcal{A}, \mathbb{P})$ and consider the map

$$\Psi : \Omega \rightarrow \mathcal{C}_{(o)}, \quad \omega \mapsto w := \Psi(\omega) := B(\cdot, \omega). \tag{4.2}$$

The metric balls $\mathbb{B}^{(\rho)}(v, r) \subset \mathcal{C}_{(0)}$ generate $\mathcal{B}(\mathcal{C}_{(0)})$. From (4.1) we conclude that the map $\omega \mapsto \rho(B(\cdot, \omega), w)$ is \mathcal{A} measurable; therefore

$$\{\omega : \Psi(\omega) \in \mathbb{B}^{(\rho)}(v, r)\} = \{\omega : \rho(B(\cdot, \omega), v) < r\}$$

shows that Ψ is $\mathcal{A}/\mathcal{B}(\mathcal{C}_{(0)})$ measurable. The image measure

$$\mu(\Gamma) := \mathbb{P}(\Psi^{-1}(\Gamma)) = \mathbb{P}(\Psi \in \Gamma), \quad \Gamma \in \mathcal{B}(\mathcal{C}_{(0)}), \quad (4.3)$$

is a measure on $\mathcal{C}_{(0)}$. Assume that Γ is a *cylinder set*, i. e. a set of the form

$$\begin{aligned} \Gamma &= \{w \in \mathcal{C}_{(0)} : w(t_1) \in C_1, \dots, w(t_n) \in C_n\} \\ &= \{w \in \mathcal{C}_{(0)} : \pi_{t_1}(w) \in C_1, \dots, \pi_{t_n}(w) \in C_n\} \end{aligned}$$

where $n \geq 1$, $t_1 < t_2 < \dots < t_n$ and $C_1, \dots, C_n \in \mathcal{B}(\mathbb{R}^d)$. Because of $w = \Psi(\omega)$ and $\pi_t(w) = w(t) = B(t, \omega)$ we see

$$\mu(\Gamma) = \mu(\pi_{t_1} \in C_1, \dots, \pi_{t_n} \in C_n) = \mathbb{P}(B(t_1) \in C_1, \dots, B(t_n) \in C_n). \quad (4.4)$$

Since the family of all cylinder sets is a \cap -stable generator of the trace σ -algebra $\mathcal{C}_{(0)} \cap \mathcal{B}^I(\mathbb{R}^d)$, we see that the finite dimensional distributions (4.4) uniquely determine the measure μ (and \mathbb{P}). This proves the following theorem.

4.2 Theorem. *On the probability space $(\mathcal{C}_{(0)}, \mathcal{B}(\mathcal{C}_{(0)}), \mu)$ the family of projections $(\pi_t)_{t \geq 0}$ satisfies (B0)–(B4), i. e. $(\pi_t)_{t \geq 0}$ is a Brownian motion.*

Theorem 4.2 says that every Brownian motion $(\Omega, \mathcal{A}, \mathbb{P}, B_t, t \geq 0)$ admits an equivalent Brownian motion $(W_t)_{t \geq 0}$ on the probability space $(\mathcal{C}_{(0)}, \mathcal{B}(\mathcal{C}_{(0)}), \mu)$. Since $W(t)$ is just the projection π_t , we can identify $W(t, w)$ with its sample path $w(t)$.

4.3 Definition. Let $(\mathcal{C}_{(0)}, \mathcal{B}(\mathcal{C}_{(0)}), \mu, \pi_t)$ be as in Theorem 4.2. The measure μ is called the *Wiener measure*, $(\mathcal{C}_{(0)}, \mathcal{B}(\mathcal{C}_{(0)}), \mu)$ is called *Wiener space* or *path space*, and $(\pi_t)_{t \geq 0}$ is the *canonical (model of the) Brownian motion process*.

We will write $(\mathcal{C}_{(0)}(I), \mathcal{B}(\mathcal{C}_{(0)}(I)), \mu)$ if we consider Brownian motion on an interval I other than $[0, \infty)$.

4.4 Remark. We have seen in §§ 2.8–2.13 that a Brownian motion is invariant under reflection, scaling, renewal, time inversion and projective reflection at $t = \infty$. Since these operations leave the finite dimensional distributions unchanged, and since the Wiener measure μ is uniquely determined by the finite dimensional distributions, μ will also be invariant. This observation justifies arguments of the type

$$\mathbb{P}(B(\cdot) \in A) = \mathbb{P}(W(\cdot) \in A)$$

where $A \in \mathcal{A}$ and W is the (path-by-path) transformation of B under any of the above operations. Below is an overview on the transformations both at the level of the sample paths and the canonical space $\mathcal{C}_{(o)}$. Note that the maps $S : \mathcal{C}_{(o)} \rightarrow \mathcal{C}_{(o)}$ in Table 4.1 are bijective.

Table 4.1. Transformations of Brownian motion.

Operation	time set I	Transformation at the level of	
		paths $W(t, \omega)$	$w \in \mathcal{C}_{(o)}(I), w \mapsto Sw(\cdot)$
Reflection	$[0, \infty)$	$-B(t, \omega)$	$-w(t)$
Renewal	$[0, \infty)$	$B(t + a, \omega) - B(a, \omega)$	$w(t + a) - w(a)$
Time inversion	$[0, a]$	$B(a - t, \omega) - B(a, \omega)$	$w(a - t) - w(a)$
Scaling	$[0, \infty)$	$\sqrt{c} B(\frac{t}{c}, \omega)$	$\sqrt{c} w(\frac{t}{c})$
Proj. reflection	$[0, \infty)$	$t B(\frac{1}{t}, \omega)$	$t w(\frac{1}{t})$

Let us finally show that the product σ -algebra $\mathcal{B}^I(\mathbb{R}^d)$ is fairly small in the sense that the set $\mathcal{C}_{(o)}$ is not measurable. We begin with an auxiliary result which shows that $\Gamma \in \mathcal{B}^I(\mathbb{R}^d)$ is uniquely determined by *countably many* indices.

4.5 Lemma. *Let $I = [0, \infty)$. For every $\Gamma \in \mathcal{B}^I(\mathbb{R}^d)$ there exists some countable set $S = S_\Gamma \subset I$ such that*

$$f \in (\mathbb{R}^d)^I, w \in \Gamma : f|_S = w|_S \implies f \in \Gamma. \quad (4.5)$$

Proof. Set $\Sigma := \{\Gamma \subset (\mathbb{R}^d)^I : (4.5) \text{ holds for some countable set } S = S_\Gamma\}$. We claim that Σ is a σ -algebra.

- Clearly, $\emptyset \in \Sigma$;
- Let $\Gamma \in \Sigma$ with $S = S_\Gamma$. Then we find for Γ^c and S that

$$f \in (\mathbb{R}^d)^I, w \in \Gamma^c, f|_S = w|_S \implies f \notin \Gamma.$$

(Otherwise, we would have $f \in \Gamma$ and then (4.5) would imply that $w \in \Gamma$.) This means that (4.5) holds for Γ^c with $S = S_\Gamma = S_{\Gamma^c}$, i. e. $\Gamma^c \in \Sigma$.

- Let $\Gamma_1, \Gamma_2, \dots \in \Sigma$ and set $S := \bigcup_{j \geq 1} S_{\Gamma_j}$. Obviously, S is countable and

$$\begin{aligned} f \in (\mathbb{R}^d)^I, w \in \bigcup_{j \geq 1} \Gamma_j, f|_S = w|_S &\implies \exists j_0 : w \in \Gamma_{j_0}, f|_{S_{\Gamma_{j_0}}} = w|_{S_{\Gamma_{j_0}}} \\ &\stackrel{(4.5)}{\implies} f \in \Gamma_{j_0} \text{ hence } f \in \bigcup_{j \geq 1} \Gamma_j. \end{aligned}$$

This proves $\bigcup_{j \geq 1} \Gamma_j \in \Sigma$.

Since all sets of the form $\{\pi_t \in C\}$, $t \in I$, $C \in \mathcal{B}(\mathbb{R}^d)$ are contained in Σ , we find

$$\mathcal{B}^I(\mathbb{R}^d) = \sigma(\pi_t : t \in I) \subset \sigma(\Sigma) = \Sigma. \quad \square$$



4.6 Corollary. *If $I = [0, \infty)$ then $\mathcal{C}_{(0)} \notin \mathcal{B}^I(\mathbb{R}^d)$.*

Proof. Assume that $\mathcal{C}_{(0)} \in \mathcal{B}^I(\mathbb{R}^d)$. By Lemma 4.5 there exists a countable set $S \subset I$ such that

$$f \in (\mathbb{R}^d)^I, w \in \mathcal{C}_{(0)} : f|_S = w|_S \implies f \in \mathcal{C}_{(0)}.$$

This relation remains valid if we enlarge S ; therefore we can assume that S is a countable dense subset of I . Fix $w \in \mathcal{C}_{(0)}$, pick $t_0 \notin S$, $\gamma \in \mathbb{R}^d \setminus \{w(t_0)\}$ and define a new function by

$$f(t) := \begin{cases} w(t), & t \neq t_0 \\ \gamma, & t = t_0 \end{cases}.$$

Then $f|_S = w|_S$, but $f \notin \mathcal{C}_{(0)}$, contradicting our assumption. \square

4.2 Kolmogorov's construction

We have seen in the previous section that the finite dimensional distributions of a d -dimensional Brownian motion $(B_t)_{t \geq 0}$ uniquely determine a measure μ on the space of all sample paths, such that the projections $(\pi_t)_{t \geq 0}$ are, under μ , again a Brownian motion. This measure μ is uniquely determined by the (finite dimensional) projections

$$\pi_{t_1, \dots, t_n}(w) := (w(t_1), \dots, w(t_n))$$

and the corresponding finite dimensional distributions

$$p_{t_1, \dots, t_n}(C_1 \times \dots \times C_n) = \mu(\pi_{t_1, \dots, t_n} \in C_1 \times \dots \times C_n)$$

where $t_1, \dots, t_n \in I$ and $C_1, \dots, C_n \in \mathcal{B}(\mathbb{R}^d)$. From (4.4) it is obvious that the conditions

$$\left. \begin{aligned} p_{t_1, \dots, t_n}(C_1 \times \dots \times C_n) &= p_{t_{\sigma(1)}, \dots, t_{\sigma(n)}}(C_{\sigma(1)} \times \dots \times C_{\sigma(n)}) \\ &\text{for all permutations } \sigma : \{1, 2, \dots, n\} \rightarrow \{1, 2, \dots, n\} \\ p_{t_1, \dots, t_{n-1}, t_n}(C_1 \times \dots \times C_{n-1} \times \mathbb{R}^d) &= p_{t_1, \dots, t_{n-1}}(C_1 \times \dots \times C_{n-1}) \end{aligned} \right\} \quad (4.6)$$

are necessary for a family p_{t_1, \dots, t_n} , $n \geq 1$, $t_1, \dots, t_n \in I$ to be finite dimensional distributions of a stochastic process.

4.7 Definition. Assume that for every $n \geq 1$ and all $t_1, \dots, t_n \in I$, p_{t_1, \dots, t_n} is a probability measure on $((\mathbb{R}^d)^n, \mathcal{B}((\mathbb{R}^d)^n))$. If these measures satisfy the *consistency conditions* (4.6), we call them *consistent* or *projective*.

In fact, the consistency conditions (4.6) are even sufficient for p_{t_1, \dots, t_n} to be finite dimensional distributions of some stochastic process. The following deep theorem is due to Kolmogorov. A proof is given in Theorem A.2 in the appendix.

4.8 Theorem (Kolmogorov 1933). *Let $I \subset [0, \infty)$ and p_{t_1, \dots, t_n} be probability measures defined on $((\mathbb{R}^d)^n, \mathcal{B}((\mathbb{R}^d)^n))$ for all $t_1, \dots, t_n \in I$, $n \geq 1$. If the family of measures is consistent, then there exists a probability measure μ on $((\mathbb{R}^d)^I, \mathcal{B}^I(\mathbb{R}^d))$ such that*

$$p_{t_1, \dots, t_n}(C) = \mu(\pi_{t_1, \dots, t_n}^{-1}(C)) \text{ for all } C \in \mathcal{B}((\mathbb{R}^d)^n).$$

Using Theorem 4.8 we can construct a stochastic process for any family of consistent finite dimensional probability distributions.

4.9 Corollary (Canonical process). *Let $I \subset [0, \infty)$ and p_{t_1, \dots, t_n} be probability measures defined on $((\mathbb{R}^d)^n, \mathcal{B}((\mathbb{R}^d)^n))$ for all $t_1, \dots, t_n \in I$, $n \geq 1$. If the family of measures is consistent, there exist a probability space $(\Omega, \mathcal{A}, \mathbb{P})$ and a d -dimensional stochastic process $(X_t)_{t \geq 0}$ such that*

Ex. 4.2

$$\mathbb{P}(X(t_1) \in C_1, \dots, X(t_n) \in C_n) = p_{t_1, \dots, t_n}(C_1 \times \dots \times C_n).$$

for all $C_1, \dots, C_n \in \mathcal{B}(\mathbb{R}^d)$.

Proof. We take $\Omega = (\mathbb{R}^d)^I$, the product σ -algebra $\mathcal{A} = \mathcal{B}^I(\mathbb{R}^d)$ and $\mathbb{P} = \mu$, where μ is the measure constructed in Theorem 4.8. Then $X_t := \pi_t$ defines a stochastic process with finite dimensional distributions $\mu \circ (\pi_{t_1}, \dots, \pi_{t_n})^{-1} = p_{t_1, \dots, t_n}$. \square

4.10 Kolmogorov's construction of BM. For $d = 1$ and $0 < t_1 < \dots < t_n$ the n -variate Gaussian measures

$$p_{t_1, \dots, t_n}(A_1 \times \dots \times A_n) = \left(\frac{1}{2\pi}\right)^{n/2} \frac{1}{\sqrt{\det C}} \int_{A_1 \times \dots \times A_n} e^{-\frac{1}{2}(x, C^{-1}x)} dx \quad (4.7)$$

with $C = (t_j \wedge t_k)_{j,k=1, \dots, n}$, are probability measures. If $t_1 = 0$, we use $\delta_0 \otimes p_{t_2, \dots, t_n}$ instead. It is not difficult to check that the family (4.7) is consistent. Therefore, Corollary 4.9 proves that there exists an \mathbb{R} -valued stochastic process $(B_t)_{t \geq 0}$. By construction, (B0) and (B3) are satisfied. From (B3) we get immediately (B2); (B1) follows if we read (2.11) backwards: Let $C = (t_j \wedge t_k)_{j,k=1, \dots, n}$ where $0 < t_1 < \dots < t_n < \infty$;

then we have for all $\xi = (\xi_1, \dots, \xi_n)^\top \in \mathbb{R}^n$

$$\begin{aligned} \mathbb{E} \left[\exp \left(i \sum_{j=1}^n \xi_j B_{t_j} \right) \right] &= \exp \left(-\frac{1}{2} \langle \xi, C \xi \rangle \right) \\ &\stackrel{(2.12)}{=} \prod_{j=1}^n \exp \left(-\frac{1}{2} (t_j - t_{j-1}) (\xi_j + \dots + \xi_n)^2 \right) \\ &\stackrel{(B2)}{\stackrel{(2.5)}{=}} \prod_{j=1}^n \mathbb{E} \left[\exp \left(i (B_{t_j} - B_{t_{j-1}}) (\xi_j + \dots + \xi_n) \right) \right]. \end{aligned}$$

On the other hand, we can easily rewrite the expression on the left-hand side in the following form

$$\mathbb{E} \left[\exp \left(i \sum_{j=1}^n \xi_j B_{t_j} \right) \right] = \mathbb{E} \left[\exp \left(i \sum_{j=1}^n (B_{t_j} - B_{t_{j-1}}) (\xi_j + \dots + \xi_n) \right) \right].$$

Using the substitution $\eta_j := \xi_j + \dots + \xi_n$ we see that

$$\mathbb{E} \left[\exp \left(i \sum_{j=1}^n \eta_j (B_{t_j} - B_{t_{j-1}}) \right) \right] = \prod_{j=1}^n \mathbb{E} \left[\exp(i \eta_j (B_{t_j} - B_{t_{j-1}})) \right]$$

for all $\eta_1, \dots, \eta_n \in \mathbb{R}$ which finally proves (B1).

This gives a new way to prove the existence of a Brownian motion. The problem is, as in all other constructions of Brownian motion, to check the continuity of the sample paths (B4). This follows from yet another theorem of Kolmogorov.

Ex. 4.1 4.11 Theorem (Kolmogorov 1934; Slutsky 1937; Chentsov 1956). *Denote by $(X_t)_{t \geq 0}$ a stochastic process on $(\Omega, \mathcal{A}, \mathbb{P})$ taking values in \mathbb{R}^d . If*

$$\mathbb{E} (|X(t) - X(s)|^\alpha) \leq c |t - s|^{1+\beta} \quad \text{for all } s, t \geq 0 \quad (4.8)$$

holds for some constants $c > 0$ and $\alpha, \beta > 0$, then $(X_t)_{t \geq 0}$ has a modification $(X'_t)_{t \geq 0}$ which has only continuous sample paths.

We will give the proof in a different context in Chapter 10 below. For a process satisfying (B0)–(B3), we can take $\alpha = 4$ and $\beta = 1$ since, cf. (2.8),

$$\mathbb{E} (|B(t) - B(s)|^4) = \mathbb{E} (|B(t - s)|^4) = 3 |t - s|^2.$$



4.12 Further reading. The concept of Wiener space has been generalized in many directions. Here we only mention Gross' concept of abstract Wiener space [114], [191], [127] and the Malliavin calculus, i. e. stochastic calculus of variations on an (abstract) Wiener space [128]. The results of Cameron and Martin are nicely summed up in [187].

- [114] Kuo: *Gaussian Measures in Banach Spaces*.
- [127] Malliavin: *Integration and Probability*.
- [128] Malliavin: *Stochastic Analysis*.
- [187] Wiener et al.: *Differential Space, Quantum Systems, and Prediction*.
- [191] Yamasaki: *Measures on infinite dimensional spaces*.

Problems

1. Let $F : \mathbb{R} \rightarrow [0, 1]$ be a distribution function.
 - (a) Show that there exists a probability space $(\Omega, \mathcal{A}, \mathbb{P})$ and a random variable X such that $F(x) = \mathbb{P}(X \leq x)$.
 - (b) Show that there exists a probability space $(\Omega, \mathcal{A}, \mathbb{P})$ and an iid sequence of random variables X_n such that $F(x) = \mathbb{P}(X_n \leq x)$.
 - (c) State and prove the corresponding assertions for the d -dimensional case.
2. Show that there exists a stochastic process $(X_t)_{t \geq 0}$ such that the random variables X_t are independent $N(0, t)$ random variables. This process also satisfies $\mathbb{P}(\lim_{s \rightarrow t} X_s \text{ exists}) = 0$ for every $t > 0$.

Chapter 5

Brownian motion as a martingale

Let I be some subset of $[0, \infty]$ and $(\mathcal{F}_t)_{t \in I}$ a filtration, i. e. an increasing family of sub- σ -algebras of \mathcal{A} . Recall that a *martingale* $(X_t, \mathcal{F}_t)_{t \in I}$ is a real or complex stochastic process $X_t : \Omega \rightarrow \mathbb{R}^d$ or $X_t : \Omega \rightarrow \mathbb{C}$ satisfying

- a) $\mathbb{E}|X_t| < \infty$ for all $t \in I$;
- b) X_t is \mathcal{F}_t measurable for every $t \in I$;
- c) $\mathbb{E}(X_t | \mathcal{F}_s) = X_s$ for all $s, t \in I, s \leq t$.

If X_t is real-valued and if we have in c) ‘ \geq ’ or ‘ \leq ’ instead of ‘ $=$ ’, we call $(X_t, \mathcal{F}_t)_{t \in I}$ a *sub-* or *supermartingale*, respectively.

We call a stochastic process $(X_t)_{t \in I}$ *adapted* to the filtration $(\mathcal{F}_t)_{t \in I}$ if X_t is for each $t \in I$ an \mathcal{F}_t measurable random variable. We write briefly $(X_t, \mathcal{F}_t)_{t \geq 0}$ for an adapted process. Clearly, X is \mathcal{F}_t adapted if, and only if, $\mathcal{F}_t^X \subset \mathcal{F}_t$ for all $t \in I$ where $\mathcal{F}_t^X = \sigma(X_s : s \in I, s \leq t)$ is the natural filtration of X .

5.1 Some ‘Brownian’ martingales

Brownian motion is a martingale with respect to its *natural filtration*, i. e. the family of σ -algebras $\mathcal{F}_t^B := \sigma(B_s : s \leq t)$. Recall that, by Lemma 2.10,

$$B_t - B_s \perp\!\!\!\perp \mathcal{F}_s^B \quad \text{for all } 0 \leq s \leq t. \quad (5.1)$$

It is often necessary to enlarge the canonical filtration $(\mathcal{F}_t^B)_{t \geq 0}$ of a Brownian motion $(B_t)_{t \geq 0}$. The property (5.1) is equivalent to the independent increments property (B1) of $(B_t)_{t \geq 0}$, see Lemma 2.10 and Problem 2.9, and it is this property which preserves the Brownian character of a filtration.

Ex. 5.1 **5.1 Definition.** Let $(B_t)_{t \geq 0}$ be a d -dimensional Brownian motion. A filtration $(\mathcal{F}_t)_{t \geq 0}$
Ex. 5.2 is called *admissible*, if

- a) $\mathcal{F}_t^B \subset \mathcal{F}_t$ for all $t \geq 0$;
- b) $B_t - B_s \perp\!\!\!\perp \mathcal{F}_s$ for all $0 \leq s \leq t$.

If \mathcal{F}_0 contains all subsets of \mathbb{P} null sets, $(\mathcal{F}_t)_{t \geq 0}$ is an *admissible complete* filtration.

The natural filtration $(\mathcal{F}_t^B)_{t \geq 0}$ is always admissible. We will discuss further examples of admissible filtrations in Lemma 6.20.

5.2 Example. Let $(B_t)_{t \geq 0}$, $B_t = (B_t^1, \dots, B_t^d)$, be a d -dimensional Brownian motion and $(\mathcal{F}_t)_{t \geq 0}$ an admissible filtration.

a) $(B_t)_{t \geq 0}$ is a martingale with respect to \mathcal{F}_t .

Indeed: Let $0 \leq s \leq t$. Using the conditions (5.1) a) and b) we get

$$\mathbb{E}(B_t | \mathcal{F}_s) = \mathbb{E}(B_t - B_s | \mathcal{F}_s) + \mathbb{E}(B_s | \mathcal{F}_s) = \mathbb{E}(B_t - B_s) + B_s = B_s.$$

b) $M(t) := |B_t|^2$ is a positive submartingale with respect to \mathcal{F}_t .

Indeed: for all $0 \leq s < t$ we can use (the conditional form of) Jensen’s inequality to get

$$\mathbb{E}(M_t | \mathcal{F}_s) = \sum_{j=1}^d \mathbb{E}((B_t^j)^2 | \mathcal{F}_s) \geq \sum_{j=1}^d \mathbb{E}(B_t^j | \mathcal{F}_s)^2 = \sum_{j=1}^d (B_s^j)^2 = M_s.$$

c) $M_t := |B_t|^2 - d \cdot t$ is a martingale with respect to \mathcal{F}_t .

Indeed: since $|B_t|^2 - d \cdot t = \sum_{j=1}^d ((B_t^j)^2 - t)$ it is enough to consider $d = 1$.

For $s < t$ we see

$$\begin{aligned} \mathbb{E}[B_t^2 - t | \mathcal{F}_s] &= \mathbb{E}[(B_t - B_s) + B_s]^2 - t | \mathcal{F}_s] \\ &= \mathbb{E}[(B_t - B_s)^2 + 2B_s(B_t - B_s) + B_s^2 - t | \mathcal{F}_s] \\ &\stackrel{B_t - B_s \perp \mathcal{F}_s}{=} \underbrace{\mathbb{E}[(B_t - B_s)^2]}_{=t-s} + 2B_s \underbrace{\mathbb{E}[B_t - B_s]}_{=0} + B_s^2 - t \\ &= B_s^2 - s. \end{aligned}$$

d) $M^\xi(t) := e^{i\langle \xi, B(t) \rangle + \frac{t}{2} |\xi|^2}$ is for all $\xi \in \mathbb{R}^d$ a complex-valued martingale with respect to \mathcal{F}_t . Ex. 5.5

Indeed: for $0 \leq s < t$ we have, cf. Example a),

$$\begin{aligned} \mathbb{E}(M^\xi(t) | \mathcal{F}_s) &= e^{\frac{t}{2} |\xi|^2} \mathbb{E}(e^{i\langle \xi, B(t) - B(s) \rangle} | \mathcal{F}_s) e^{i\langle \xi, B(s) \rangle} \\ &\stackrel{B_t - B_s \perp \mathcal{F}_s}{=} e^{\frac{t}{2} |\xi|^2} \mathbb{E}(e^{i\langle \xi, B(t-s) \rangle}) e^{i\langle \xi, B(s) \rangle} \\ &\stackrel{B_t - B_s \sim B_{t-s}}{=} e^{\frac{t}{2} |\xi|^2} e^{-\frac{t-s}{2} |\xi|^2} e^{i\langle \xi, B(s) \rangle} \\ &= M^\xi(s). \end{aligned}$$

e) $M^\zeta(t) := e^{\langle \zeta, B(t) \rangle - \frac{t}{2} |\zeta|^2}$ is for all $\zeta \in \mathbb{C}^d$ a complex-valued martingale with respect to \mathcal{F}_t . Ex. 5.6

Indeed: since $\mathbb{E} e^{|\zeta| |B(t)|} < \infty$ for all $\zeta \in \mathbb{C}^d$, cf. (2.6), we can use the argument of Example d).

Since Brownian motion is a martingale we have all martingale tools at our disposal. The following maximal estimate will be particularly useful. It is a direct consequence of Doob's maximal inequality A.10 and the observation that for a process with continuous sample paths $\sup_{s \in D \cap [0, t]} |B_s| = \sup_{s \in [0, t]} |B_s|$ holds for any dense subset $D \subset [0, \infty)$.

5.3 Lemma. *Let $(B_t)_{t \geq 0}$ be a d -dimensional Brownian motion. Then we have for all $t \geq 0$ and $p > 1$*

$$\mathbb{E} \left[\sup_{s \leq t} |B_s|^p \right] \leq \left(\frac{p}{p-1} \right)^p \mathbb{E}[|B_t|^p].$$

The following characterization of a Brownian motion through conditional characteristic functions is quite useful.

5.4 Lemma. *Let $(X_t)_{t \geq 0}$ be a continuous \mathbb{R}^d -valued stochastic process with $X_0 = 0$ and which is adapted to the filtration $(\mathcal{F}_t)_{t \geq 0}$. Then $(X_t, \mathcal{F}_t)_{t \geq 0}$ is a BM^d if, and only if, for all $0 \leq s \leq t$ and $\xi \in \mathbb{R}^d$*

$$\mathbb{E}[e^{i\langle \xi, X_t - X_s \rangle} \mid \mathcal{F}_s] = e^{-\frac{1}{2} |\xi|^2 (t-s)}. \quad (5.2)$$

Proof. We verify (2.17), i. e. for all $n \geq 0$, $0 = t_0 < t_1 < \dots < t_n$, $\xi_0, \dots, \xi_n \in \mathbb{R}^d$ and $X_{t_{-1}} := 0$:

$$\mathbb{E} \left[\exp \left(i \sum_{j=0}^n \langle \xi_j, X_{t_j} - X_{t_{j-1}} \rangle \right) \right] = \exp \left[-\frac{1}{2} \sum_{j=1}^n |\xi_j|^2 (t_j - t_{j-1}) \right].$$

Using the tower property we find

$$\begin{aligned} & \mathbb{E} \left[\exp \left(i \sum_{j=0}^n \langle \xi_j, X_{t_j} - X_{t_{j-1}} \rangle \right) \right] \\ &= \mathbb{E} \left[\mathbb{E} \left[\exp \left(i \langle \xi_n, X_{t_n} - X_{t_{n-1}} \rangle \right) \mid \mathcal{F}_{t_{n-1}} \right] \exp \left(i \sum_{j=0}^{n-1} \langle \xi_j, X_{t_j} - X_{t_{j-1}} \rangle \right) \right] \\ &= \exp \left(-\frac{1}{2} |\xi_n|^2 (t_n - t_{n-1}) \right) \mathbb{E} \left[\exp \left(i \sum_{j=0}^{n-1} \langle \xi_j, X_{t_j} - X_{t_{j-1}} \rangle \right) \right]. \end{aligned}$$

Iterating this step we get (2.17), and the claim follows from Lemma 2.14. \square

All examples in 5.2 were of the form ' $f(B_t)$ '. We can get many more examples of this type. For this we begin with a real-variable lemma.

5.5 Lemma. Let $p(t, x) := (2\pi t)^{-d/2} \exp(-|x|^2/2t)$ be the transition density of a d -dimensional Brownian motion. Then Ex. 5.7

$$\frac{\partial}{\partial t} p(t, x) = \frac{1}{2} \Delta_x p(t, x) \quad \text{for all } x \in \mathbb{R}^d, t > 0, \quad (5.3)$$

where $\Delta = \Delta_x = \sum_{j=1}^d \frac{\partial^2}{\partial x_j^2}$ is the d -dimensional Laplace operator. Moreover,

$$\begin{aligned} \int p(t, x) \frac{1}{2} \Delta_x f(t, x) dx &= \int f(t, x) \frac{1}{2} \Delta_x p(t, x) dx \\ &= \int f(t, x) \frac{\partial}{\partial t} p(t, x) dx \end{aligned} \quad (5.4)$$

for all functions $f \in \mathcal{C}^{1,2}((0, \infty) \times \mathbb{R}^d) \cap \mathcal{C}([0, \infty) \times \mathbb{R}^d)$ satisfying

$$|f(t, x)| + \left| \frac{\partial f(t, x)}{\partial t} \right| + \sum_{j=1}^d \left| \frac{\partial f(t, x)}{\partial x_j} \right| + \sum_{j,k=1}^d \left| \frac{\partial^2 f(t, x)}{\partial x_j \partial x_k} \right| \leq c(t) e^{C|x|} \quad (5.5)$$

for all $t > 0$, $x \in \mathbb{R}^d$, some constant $C > 0$ and with a locally bounded function $c : (0, \infty) \rightarrow [0, \infty)$.

Proof. The identity (5.3) can be directly verified. We leave this (lengthy but simple) computation to the reader. The first equality in (5.4) follows if we integrate by parts (twice); note that the condition (5.5) imposed on $f(t, x)$ guarantees that the marginal terms vanish and that all integrals exist. Finally, the second equality in (5.4) follows if we plug (5.3) into the second integral. \square

5.6 Theorem. Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a d -dimensional Brownian motion and assume that $f \in \mathcal{C}^{1,2}((0, \infty) \times \mathbb{R}^d, \mathbb{R}) \cap \mathcal{C}([0, \infty) \times \mathbb{R}^d, \mathbb{R})$ satisfies (5.5). Then Ex. 5.8
Ex. 16.4

$$M_t^f := f(t, B_t) - f(0, B_0) - \int_0^t Lf(r, B_r) dr, \quad t \geq 0, \quad (5.6)$$

is an \mathcal{F}_t martingale. Here $Lf(t, x) = \frac{\partial}{\partial t} f(t, x) + \frac{1}{2} \Delta_x f(t, x)$.

Proof. Observe that for $0 \leq s < t$

$$M_t^f - M_s^f = f(t, B_t) - f(s, B_s) - \int_s^t Lf(r, B_r) dr;$$

we have to show that $\mathbb{E}(M_t^f - M_s^f | \mathcal{F}_s) = 0$. Using $B_t - B_s \perp \mathcal{F}_s$ we see from

Lemma A.3 that

$$\begin{aligned}
& \mathbb{E} (M_t^f - M_s^f \mid \mathcal{F}_s) \\
&= \mathbb{E} \left(f(t, B_t) - f(s, B_s) - \int_s^t Lf(r, B_r) dr \mid \mathcal{F}_s \right) \\
&= \mathbb{E} \left(f(t, (B_t - B_s) + B_s) - f(s, B_s) - \int_s^t Lf(r, (B_r - B_s) + B_s) dr \mid \mathcal{F}_s \right) \\
&\stackrel{A.3}{=} \mathbb{E} \left(f(t, (B_t - B_s) + z) - f(s, z) - \int_s^t Lf(r, (B_r - B_s) + z) dr \right) \Big|_{z=B_s}.
\end{aligned}$$

Set $\phi(t, x) = \phi_{s,z}(t, x) := f(t + s, x + z)$, $s > 0$, and observe that $B_t - B_s \sim B_{t-s}$. Since $f \in \mathcal{C}^{1,2}((0, \infty) \times \mathbb{R}^d)$, the shifted function ϕ satisfies (5.5) on $[0, T] \times \mathbb{R}^d$ where we can take $C = c(t) \equiv \gamma$ and the constant γ depends only on s, T and z .

We have to show that for all $z \in \mathbb{R}$

$$\mathbb{E} (M_{t-s}^\phi - M_0^\phi) = \mathbb{E} \left(\phi(t-s, B_{t-s}) - \phi(0, B_0) - \int_0^{t-s} L\phi(r, B_r) dr \right) = 0.$$

Let $0 < \epsilon < u$. Then

$$\begin{aligned}
& \mathbb{E}(M_u^\phi - M_\epsilon^\phi) \\
&= \mathbb{E} \left(\phi(u, B_u) - \phi(\epsilon, B_\epsilon) - \int_\epsilon^u L\phi(r, B_r) dr \right) \\
&\stackrel{(*)}{=} \int (p(u, x)\phi(u, x) - p(\epsilon, x)\phi(\epsilon, x)) dx \\
&\quad - \int_\epsilon^u \int p(r, x) \left(\frac{\partial \phi(r, x)}{\partial r} + \frac{1}{2} \Delta_x \phi(r, x) \right) dx dr \\
&\stackrel{5.5}{=} \int (p(u, x)\phi(u, x) - p(\epsilon, x)\phi(\epsilon, x)) dx \\
&\quad - \int_\epsilon^u \int \left(p(r, x) \frac{\partial \phi(r, x)}{\partial r} + \phi(r, x) \frac{\partial p(r, x)}{\partial r} \right) dx dr \\
&= \int (p(u, x)\phi(u, x) - p(\epsilon, x)\phi(\epsilon, x)) dx - \int_\epsilon^u \int \frac{\partial}{\partial r} (p(r, x)\phi(r, x)) dx dr \\
&\stackrel{(*)}{=} \int (p(u, x)\phi(u, x) - p(\epsilon, x)\phi(\epsilon, x)) dx - \int \int_\epsilon^u \frac{\partial}{\partial r} (p(r, x)\phi(r, x)) dr dx
\end{aligned}$$

and this equals zero. In the lines marked with an asterisk (*) we used Fubini's theorem which is indeed applicable because of (5.5) and $\epsilon > 0$. Moreover,

$$|M_u^\phi - M_\epsilon^\phi| \leq \gamma e^{\gamma|B_u|} + \gamma e^{\gamma|B_\epsilon|} + (u - \epsilon) \sup_{\epsilon \leq r \leq u} \gamma e^{\gamma|B_r|} \leq (2 + u) \gamma \sup_{0 \leq r \leq u} e^{\gamma|B_r|}.$$

By Doob's maximal inequality for $p = \gamma$ – without loss of generality we can assume that $\gamma > 1$ – and for the submartingale $e^{|B_r|}$ we see that

$$\mathbb{E} \left(\sup_{0 \leq r \leq u} e^{\gamma |B_r|} \right) \leq \left(\frac{\gamma}{\gamma - 1} \right)^\gamma \mathbb{E} \left(e^{\gamma |B_u|} \right) < \infty.$$

Now we can use dominated convergence to get

$$\mathbb{E} (M_u^\phi - M_0^\phi) = \lim_{\epsilon \downarrow 0} \mathbb{E} (M_u^\phi - M_\epsilon^\phi) = 0. \quad \square$$

5.2 Stopping and sampling

If we want to know when a Brownian motion $(B_t)_{t \geq 0}$

- leaves or enters a set for the first time,
- hits its running maximum,
- returns to zero,

we have to look at random times. A random time $\tau : \Omega \rightarrow [0, \infty]$ is a *stopping time* (with respect to $(\mathcal{F}_t)_{t \geq 0}$) if

$$\{\tau \leq t\} \in \mathcal{F}_t \quad \text{for all } t \geq 0. \quad (5.7)$$

Typical examples of stopping times are entry and hitting times of a process $(X_t)_{t \geq 0}$ into a set $A \in \mathcal{B}(\mathbb{R}^d)$:

(first) entry time into A : $\tau_A^\circ := \inf\{t \geq 0 : X_t \in A\}$,

(first) hitting time of A : $\tau_A := \inf\{t > 0 : X_t \in A\}$,

($\inf \emptyset = \infty$); sometimes τ_{A^c} is called the (first) exit time from A . Note that $\tau_A^\circ \leq \tau_A$. If $t \mapsto X_t$ and $(\mathcal{F}_t)_{t \geq 0}$ are sufficiently regular, τ_A°, τ_A are stopping times for every Borel set $A \in \mathcal{B}(\mathbb{R}^d)$. In this generality, the proof is very hard, cf. [13, Chapter I.10]. Ex. 5.11

For our purposes it is enough to consider closed and open sets A . The natural filtration $\mathcal{F}_t^X := \sigma(X_s : s \leq t)$ of a stochastic process $(X_t)_{t \geq 0}$ is relatively small. For many interesting stopping times we have to consider the slightly larger filtration

$$\mathcal{F}_{t+}^X := \bigcap_{u > t} \mathcal{F}_u^X = \bigcap_{n \geq 1} \mathcal{F}_{t+\frac{1}{n}}^X. \quad (5.8)$$

5.7 Lemma. Let $(X_t)_{t \geq 0}$ be a d -dimensional stochastic process with right-continuous sample paths and $U \subset \mathbb{R}^d$ an open set. The first hitting time τ_U satisfies

$$\{\tau_U < t\} \in \mathcal{F}_t^X \quad \text{and} \quad \{\tau_U \leq t\} \in \mathcal{F}_{t+}^X,$$

i. e. it is a stopping time with respect to \mathcal{F}_{t+}^X .

Ex. 5.12 It is not difficult to see that, for open sets U , the stopping times τ_U° and τ_U coincide.

Proof. The second assertion follows immediately from the first as

$$\{\tau_U \leq t\} = \bigcap_{n \geq 1} \{\tau_U < t + \frac{1}{n}\} \in \bigcap_{n \geq 1} \mathcal{F}_{t+\frac{1}{n}}^X = \mathcal{F}_{t+}^X. \quad (5.9)$$

For the first assertion we claim that for all $t > 0$

$$\{\tau_U < t\} = \bigcup_{\mathbb{Q}^+ \ni r \leq t} \underbrace{\{X(r) \in U\}}_{\in \mathcal{F}_r^X} \in \mathcal{F}_t^X.$$

Indeed: ‘ \subset ’ if $\tau_U(\omega) < t$, then there is some $s < t$ with $X(s, \omega) \in U$. Since the paths are continuous from the right and U is open, we find some $s < r < t$, $r \in \mathbb{Q}$, with $X(r, \omega) \in U$. Hence $\omega \in \bigcup_{\mathbb{Q}^+ \ni r < t} \{X(r) \in U\}$.

‘ \supset ’ Conversely, let $\omega \in \{X(r) \in U\}$ for some $r \in (0, t] \cap \mathbb{Q}$. Since U is open and $t \mapsto X(t)$ right-continuous, we have $\tau_U(\omega) < r \leq t$. \square

5.8 Lemma. Let $(X_t)_{t \geq 0}$ be a d -dimensional stochastic process with continuous sample paths and $F \subset \mathbb{R}^d$ a closed set. Then τ_F° is an \mathcal{F}_t^X stopping time whereas τ_F is an \mathcal{F}_{t+}^X stopping time.

Ex. 5.13 *Proof.* Denote by $d(x, F) := \inf_{y \in F} |x - y|$ the distance of x and F . If we can show that

$$\{\tau_F^\circ \leq t\} = \left\{ \omega \in \Omega : \inf_{r \in \mathbb{Q} \cap [0, t]} d(X_r(\omega), F) = 0 \right\} =: \Omega_t,$$

then τ_F° is a stopping time since, obviously, $\Omega_t \in \mathcal{F}_t$.

Indeed: ‘ \supset ’ assume that $\omega \in \Omega_t$. Then there is a sequence $(r_j)_{j \geq 1} \subset \mathbb{Q}^+ \cap [0, t]$ and some $s \leq t$ such that

$$r_j \xrightarrow{j \rightarrow \infty} s, \quad d(X(r_j, \omega), F) \xrightarrow{j \rightarrow \infty} 0 \quad \text{and} \quad X(r_j, \omega) \xrightarrow{j \rightarrow \infty} X(s, \omega).$$

Since $d(x, F)$ is continuous, we get $d(X(s, \omega), F) = 0$, i.e. $X(s, \omega) \in F$, as F is closed. Thus, $\tau_F^\circ(\omega) \leq s \leq t$.

‘ \subset ’ Now we assume that $\omega \notin \Omega_t$. Then

$$\inf_{\mathbb{Q}^+ \ni r \leq t} d(X(r, \omega), F) \geq \delta > 0 \xrightarrow[\text{paths}]{\text{cts.}} X(s, \omega) \notin F, \quad \forall s \leq t \implies \tau_F^\circ(\omega) > t.$$

For τ_F we set $\tau_n := \inf \{s \geq 0 : X_{1/n+s} \in F\}$. Then

$$\tau_F = \inf_{n \geq 1} \inf \left\{ t \geq \frac{1}{n} : X_t \in F \right\} = \inf_{n \geq 1} \left(\frac{1}{n} + \tau_n \right)$$

and we have $\{\tau_F < t\} = \bigcup_{n \geq 1} \{\tau_n < t - \frac{1}{n}\} \in \mathcal{F}_t^X$. As in (5.9) we conclude that $\{\tau_F \leq t\} \in \mathcal{F}_{t+}^X$. \square

5.9 Example (First passage time). For a BM¹ $(B_t)_{t \geq 0}$ the entry time $\tau_b := \tau_{\{b\}}^\circ$ into the closed set $\{b\}$ is often called *first passage time*. Observe that

$$\sup_{t \geq 0} B_t \geq \sup_{n \geq 1} B_n = \sup_{n \geq 1} (\xi_1 + \xi_2 + \cdots + \xi_n)$$

where the random variables $\xi_j = B_j - B_{j-1}$ are iid standard normal random variables. Then

$$S := \sup_{n \geq 1} (\xi_1 + \xi_2 + \cdots + \xi_n) = \xi_1 + \sup_{n \geq 2} (\xi_2 + \cdots + \xi_n) =: \xi_1 + S'.$$

By the iid property of the ξ_j we get that $S \sim S'$ and $\xi_1 \perp\!\!\!\perp S'$. Since $\xi_1 \sim \mathbf{N}(0, 1)$ is not trivial, we conclude that $S = \infty$ a.s. The same argument applies to the infimum $\inf_{t \geq 0} B_t$ and we get

$$\mathbb{P} \left(\sup_{t \geq 0} B_t = +\infty, \quad \inf_{t \geq 0} B_t = -\infty \right) = 1, \quad (5.10)$$

hence

$$\mathbb{P}(\tau_b < \infty) = 1 \quad \text{for all } b \in \mathbb{R}. \quad (5.11)$$

For an alternative martingale proof of this fact, see Theorem 5.13 below.

Because of the continuity of the sample paths, (5.10) immediately implies that the random set $\{t \geq 0 : B_t(\omega) = b\}$ is a.s. unbounded. In particular, a one-dimensional Brownian motion is point recurrent, i.e. it returns to each level $b \in \mathbb{R}$ time and again.

As usual, we can associate σ -algebras with an \mathcal{F}_t stopping time τ :

$$\mathcal{F}_{\tau(+)} := \left\{ A \in \mathcal{F}_\infty := \sigma \left(\bigcup_{t \geq 0} \mathcal{F}_t \right) : A \cap \{\tau \leq t\} \in \mathcal{F}_{t(+)} \quad \forall t \geq 0 \right\}.$$

Note that, despite the slightly misleading notation, \mathcal{F}_τ is a family of sets and \mathcal{F}_τ does not depend on ω !

If we apply the optional stopping theorem, see A.18, we obtain rather surprising results.

5.10 Theorem (Wald's identities. Wald 1944). *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM¹ and assume that τ is an \mathcal{F}_t stopping time. If $\mathbb{E} \tau < \infty$, then $B_\tau \in L^2(\mathbb{P})$ and we have*

$$\mathbb{E} B_\tau = 0 \quad \text{and} \quad \mathbb{E} B_\tau^2 = \mathbb{E} \tau.$$

Proof. Since $(\tau \wedge t)_{t \geq 0}$ is a family of bounded stopping times, the optional stopping theorem (Theorem A.18) applies and shows that $(B_{\tau \wedge t}, \mathcal{F}_{\tau \wedge t})_{t \geq 0}$ is a martingale. In particular, $\mathbb{E} B_{\tau \wedge t} = \mathbb{E} B_0 = 0$.

Using Lemma 5.3 for $p = 2$ we see that

$$\mathbb{E} [B_{\tau \wedge t}^2] \leq \mathbb{E} \left[\sup_{s \leq t} B_s^2 \right] \leq 4 \mathbb{E} [B_t^2] = 4t,$$

Ex. 5.14
Ex. 5.15



Ex. 5.18

i. e. $B_{\tau \wedge t} \in L^2(\mathbb{P})$. Again by optional stopping we conclude with Example 5.2 c) that $(M_t^2 := B_{\tau \wedge t}^2 - \tau \wedge t, \mathcal{F}_{\tau \wedge t})_{t \geq 0}$ is a martingale, hence

$$\mathbb{E}[M_t^2] = \mathbb{E}[M_0^2] = 0 \quad \text{or} \quad \mathbb{E}[B_{\tau \wedge t}^2] = \mathbb{E}(\tau \wedge t).$$

By the martingale property we see $\mathbb{E}[B_{\tau \wedge t} B_{\tau \wedge s}] = \mathbb{E}[B_{\tau \wedge s}^2]$ for all $s \leq t$, i. e.

$$\mathbb{E}[(B_{\tau \wedge t} - B_{\tau \wedge s})^2] = \mathbb{E}[B_{\tau \wedge t}^2 - B_{\tau \wedge s}^2] = \mathbb{E}(\tau \wedge t - \tau \wedge s) \xrightarrow[s, t \rightarrow \infty]{\text{dom. conv.}} 0.$$

Thus, $(B_{\tau \wedge t})_{t \geq 0}$ is an L^2 -Cauchy sequence; the fact that $\tau < \infty$ a. s. and the continuity of Brownian motion yield L^2 - $\lim_{t \rightarrow \infty} B_{\tau \wedge t} = B_\tau$. In particular, we have $B_\tau \in L^2(\mathbb{P})$ and, by monotone convergence,

$$\mathbb{E}[B_\tau^2] = \lim_{t \rightarrow \infty} \mathbb{E}[B_{\tau \wedge t}^2] = \lim_{t \rightarrow \infty} \mathbb{E}[\tau \wedge t] = \mathbb{E} \tau.$$

In particular, we see that $\mathbb{E} B_\tau = \lim_{t \rightarrow \infty} \mathbb{E} B_{\tau \wedge t} = 0$, as L^2 -convergence implies L^1 -convergence. \square

The following corollaries are typical applications of Wald's identities.

Ex. 5.16 5.11 Corollary. *Let $(B_t)_{t \geq 0}$ be a BM¹ and $\tau = \inf\{t \geq 0 : B_t \notin (-a, b)\}$ be the first entry time into the set $(-a, b)^c$, $a, b \geq 0$. Then*

$$\mathbb{P}(B_\tau = -a) = \frac{b}{a+b}, \quad \mathbb{P}(B_\tau = b) = \frac{a}{a+b} \quad \text{and} \quad \mathbb{E} \tau = ab. \quad (5.12)$$

Proof. Note that $t \wedge \tau$ is a bounded stopping time. Since $B_{t \wedge \tau} \in [-a, b]$, we have $|B_{t \wedge \tau}| \leq a \vee b$ and so

$$\mathbb{E}(t \wedge \tau) \stackrel{5.10}{=} \mathbb{E}(B_{t \wedge \tau}^2) \leq a^2 \vee b^2.$$

By monotone convergence $\mathbb{E} \tau \leq a^2 \vee b^2 < \infty$. We can apply Theorem 5.10 and get

$$-a \mathbb{P}(B_\tau = -a) + b \mathbb{P}(B_\tau = b) = \mathbb{E} B_\tau = 0.$$

Since $\tau < \infty$ a. s., we also know

$$\mathbb{P}(B_\tau = -a) + \mathbb{P}(B_\tau = b) = 1.$$

Solving this system of equations yields

$$\mathbb{P}(B_\tau = -a) = \frac{b}{a+b} \quad \text{and} \quad \mathbb{P}(B_\tau = b) = \frac{a}{a+b}.$$

The second identity from Theorem 5.10 finally gives

$$\mathbb{E} \tau = \mathbb{E}(B_\tau^2) = a^2 \mathbb{P}(B_\tau = -a) + b^2 \mathbb{P}(B_\tau = b) = ab. \quad \square$$

The next result is quite unexpected: A one-dimensional Brownian motion hits 1 infinitely often with probability one – but the average time the Brownian motion needs to get there is infinite.



5.12 Corollary. *Let $(B_t)_{t \geq 0}$ be a BM^1 and $\tau_1 = \inf\{t \geq 0 : B_t = 1\}$ be the first passage time of the level 1. Then $\mathbb{E} \tau_1 = \infty$.*

Proof. We know from Example 5.9 that $\tau_1 < \infty$ almost surely. Then, $B_{\tau_1} = 1$ and $\mathbb{E} B_{\tau_1} = \mathbb{E} B_{\tau_1}^2 = 1$. In view of Wald's identities, Theorem 5.10, this is only possible if $\mathbb{E} \tau_1 = \infty$. \square

5.3 The exponential Wald identity

Let $(B_t)_{t \geq 0}$ be a BM^1 and $\tau_b = \tau_{\{b\}}^\circ$ the first passage time of the level b . Recall from Example 5.2 d) that $M^\xi(t) := e^{\xi B(t) - \frac{1}{2} \xi^2 t}$, $t \geq 0$ and $\xi > 0$, is a martingale. Applying the optional stopping theorem (Theorem A.18) to the bounded stopping times $t \wedge \tau_b$ we see that

$$1 = \mathbb{E} M^\xi(0) = \mathbb{E} M^\xi(t \wedge \tau_b) = \mathbb{E}[e^{\xi B(t \wedge \tau_b) - \frac{1}{2} \xi^2 (t \wedge \tau_b)}].$$

Since $B(t \wedge \tau_b) \leq b$ for $b > 0$, we have $0 \leq e^{\xi B(t \wedge \tau_b) - \frac{1}{2} \xi^2 (t \wedge \tau_b)} \leq e^{\xi b}$. Using the fact that $e^{-\infty} = 0$ we get

$$\lim_{t \rightarrow \infty} e^{\xi B(t \wedge \tau_b) - \frac{1}{2} \xi^2 (t \wedge \tau_b)} = \begin{cases} e^{\xi B(\tau_b) - \frac{1}{2} \xi^2 \tau_b}, & \text{if } \tau_b < \infty, \\ 0, & \text{if } \tau_b = \infty. \end{cases}$$

Thus,

$$1 = \mathbb{E} \left[e^{\xi B(t \wedge \tau_b) - \frac{1}{2} \xi^2 (t \wedge \tau_b)} \right] \xrightarrow[t \rightarrow \infty]{\text{dom. conv.}} e^{\xi b} \mathbb{E} \left[\mathbb{1}_{\{\tau_b < \infty\}} e^{-\frac{1}{2} \xi^2 \tau_b} \right]$$

which shows

$$\mathbb{E} \left[\mathbb{1}_{\{\tau_b < \infty\}} e^{-\frac{1}{2} \xi^2 \tau_b} \right] = e^{-\xi b}.$$

By monotone convergence we get

$$\mathbb{P}(\tau_b < \infty) = \lim_{\xi \downarrow 0} \mathbb{E} \left[\mathbb{1}_{\{\tau_b < \infty\}} e^{-\frac{1}{2} \xi^2 \tau_b} \right] = 1.$$

Inserting this into the previous equality is (almost) the proof of

5.13 Theorem. *Let $(B_t)_{t \geq 0}$ be a BM^1 . Then the first passage time $\tau_b = \tau_{\{b\}}^\circ$, $b \in \mathbb{R}$, is a.s. finite and its Laplace transform is given by*

$$\mathbb{E} e^{-\zeta \tau_b} = e^{-\sqrt{2\zeta} |b|}, \quad \zeta \geq 0.$$

Proof. For $b \geq 0$ there is nothing left to show. If $b < 0$ we apply the calculation preceding Theorem 5.13 with $-\xi$ instead of ξ . \square

In the proof of Theorem 5.13 we use $\mathbb{E} \exp(\xi B(\tau_b) - \frac{1}{2} \xi^2 \tau_b) = 1$ for $\xi > 0$. A similar stopping result holds for stopping times τ which satisfy $\mathbb{E} e^{\tau/2} < \infty$. This is a very special case of the Novikov condition, cf. Theorem 17.4 below or [99, Chapter 3.5.D] and [156, pp. 332-3].

5.14 Theorem (Exponential Wald identity). *Let $(B_t)_{t \geq 0}$ be a BM¹ and let τ be a \mathcal{F}_t^B stopping time such that $\mathbb{E} e^{\tau/2} < \infty$. Then*

$$\mathbb{E} e^{B(\tau) - \frac{1}{2} \tau} = 1.$$

Proof. Fix some $c \in (0, 1)$ and pick $p = p(c) > 1$ such that $p < \frac{1}{c} \wedge \frac{1}{c(2-c)}$. From the optional stopping theorem we know that $M_t := e^{cB_{t \wedge \tau} - \frac{1}{2} c^2 (t \wedge \tau)}$, $t \geq 0$, is a martingale. Using the Hölder inequality for the conjugate exponents $1/pc$ and $1/(1-pc)$ we find

$$\begin{aligned} \mathbb{E} \left[e^{pcB_{t \wedge \tau} - \frac{1}{2} pc^2 (t \wedge \tau)} \right] &= \mathbb{E} \left[e^{pc(B_{t \wedge \tau} - \frac{1}{2} (t \wedge \tau))} e^{\frac{1}{2} pc(1-c)(t \wedge \tau)} \right] \\ &\leq \underbrace{\left[\mathbb{E} e^{B_{t \wedge \tau} - \frac{1}{2} (t \wedge \tau)} \right]^{pc}}_{= [\mathbb{E} M_t]^{pc} = [\mathbb{E} M_0]^{pc} = 1} \left[\mathbb{E} e^{\frac{1}{2} \frac{pc(1-c)}{1-pc} \tau} \right]^{1-pc} \leq (\mathbb{E} e^{\tau/2})^{1-pc}. \end{aligned}$$

In the last step we used that $pc(1-c)/(1-pc) \leq 1$. This shows that the p th moment of the martingale M_t is uniformly bounded for all $t > 0$, i.e. $(M_t)_{t \geq 0}$ is uniformly integrable.

Therefore we can let $t \rightarrow \infty$ and find, using uniform integrability and Hölder's inequality for the exponents $1/c$ and $1/(1-c)$,

$$1 = \mathbb{E} e^{cB_\tau - \frac{1}{2} c^2 \tau} = \mathbb{E} \left[e^{cB_\tau - \frac{1}{2} c \tau} e^{\frac{1}{2} c(1-c)\tau} \right] \leq \left[\mathbb{E} e^{B_\tau - \frac{1}{2} \tau} \right]^c \left[\mathbb{E} e^{\frac{1}{2} c \tau} \right]^{1-c}.$$

Since the last factor is bounded by $(\mathbb{E} e^{\tau/2})^{1-c} \xrightarrow{c \rightarrow 1} 1$, we find as $c \rightarrow 1$

$$1 \leq \mathbb{E} e^{B_\tau - \frac{1}{2} \tau} = \mathbb{E} \left[\lim_{t \rightarrow \infty} e^{B_{t \wedge \tau} - \frac{1}{2} (t \wedge \tau)} \right] \leq \lim_{t \rightarrow \infty} \mathbb{E} \left[e^{B_{t \wedge \tau} - \frac{1}{2} (t \wedge \tau)} \right] = 1,$$

and the assertion follows. \square

5.15 Remark. A close inspection of the proof of Theorem 5.14 shows that we have actually shown $\mathbb{E} e^{cB(\tau) - \frac{1}{2} c^2 \tau} = 1$ for all $0 \leq c \leq 1$ and all stopping times τ satisfying $\mathbb{E} e^{\tau/2} < \infty$. If we replace in this calculation B by $-B$, we get

$$\mathbb{E} e^{\xi B(\tau) - \frac{1}{2} \xi^2 \tau} = 1 \quad \text{for all } -1 \leq \xi \leq 1.$$



5.16 Further reading. The monograph [156] is one of the best and most comprehensive sources on martingales in connection with Brownian motion. A completely ‘martingale’ view of the world and everything else can be found in (the second volume of) [161]. Applications of Brownian motion and martingales to analysis are in [49].

[49] Durrett: *Brownian Motion and Martingales in Analysis*.

[156] Revuz, Yor: *Continuous Martingales and Brownian Motion*

[161] Rogers, Williams: *Diffusions, Markov Processes and Martingales*.

Problems

- Let $(B_t)_{t \geq 0}$ be a BM^d and assume that X is a d -dimensional random variable which is independent of \mathcal{F}_∞^B .
 - Show that $\tilde{\mathcal{F}}_t := \sigma(X, B_s : s \leq t)$ defines an admissible filtration for $(B_t)_{t \geq 0}$.
 - The completion $\bar{\mathcal{F}}_t^B$ of \mathcal{F}_t^B is the smallest σ -algebra which contains \mathcal{F}_t^B and all subsets of \mathbb{P} null sets. Show that $(\bar{\mathcal{F}}_t^B)_{t \geq 0}$ is admissible.
- Let $(\mathcal{F}_t)_{t \geq 0}$ be an admissible filtration for the Brownian motion $(B_t)_{t \geq 0}$. Mimic the proof of Lemma 2.10 and show that $\mathcal{F}_t \perp\!\!\!\perp \mathcal{F}_{[t, \infty)}^W := \sigma(B_u - B_t : u \geq t)$ for each $t > 0$.
- Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a martingale and denote by \mathcal{F}_t^* be the completion of \mathcal{F}_t (completion means to add all subsets of \mathbb{P} -null sets).
 - Show that $(X_t, \mathcal{F}_t^*)_{t \geq 0}$ is a martingale.
 - Let $(\tilde{X}_t)_{t \geq 0}$ be a modification of $(X_t)_{t \geq 0}$. Show that $(\tilde{X}_t, \mathcal{F}_t^*)_{t \geq 0}$ is a martingale.
- Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a sub-martingale with continuous paths and $\mathcal{F}_{t+} = \bigcap_{u>t} \mathcal{F}_u$. Show that $(X_t, \mathcal{F}_{t+})_{t \geq 0}$ is again a sub-martingale.
- Let $(B_t)_{t \geq 0}$ be a BM^1 . Find a polynomial $\pi(t, x)$ in x and t which is of order 4 in the variable x such that $\pi(t, B_t)$ is a martingale.
Hint: Use the exponential Wald identity $\mathbb{E} \exp(\xi B_\tau - \frac{1}{2} \xi^2 \tau) = 1$, $-1 \leq \xi \leq 1$ for a suitable stopping time τ and use a power-series expansion of the left-hand side in ξ .
- Let $(B_t)_{t \geq 0}$ be a BM^d . Find all $c \in \mathbb{R}$ such that $\mathbb{E} e^{c|B_t|}$ and $\mathbb{E} e^{c|B_t|^2}$ are finite.
- Let $p(t, x) = (2\pi t)^{-d/2} \exp(-|x|^2/(2t))$, $x \in \mathbb{R}^d$, $t > 0$, be the transition density of a d -dimensional Brownian motion.

(a) Show that $p(t, x)$ is a solution for the heat equation, i. e.

$$\frac{\partial}{\partial t} p(t, x) = \frac{1}{2} \Delta_x p(t, x) = \frac{1}{2} \sum_{j=1}^d \frac{\partial^2}{\partial x_j^2} p(t, x), \quad x \in \mathbb{R}^d, \quad t > 0.$$

(b) Let $f \in \mathcal{C}^{1,2}((0, \infty) \times \mathbb{R}^d) \cap \mathcal{C}([0, \infty) \times \mathbb{R}^d)$ be such that for some constants $c, c_t > 0$ and all $x \in \mathbb{R}^d, t > 0$

$$|f(t, x)| + \left| \frac{\partial f(t, x)}{\partial t} \right| + \sum_{j=1}^d \left| \frac{\partial f(t, x)}{\partial x_j} \right| + \sum_{j,k=1}^d \left| \frac{\partial^2 f(t, x)}{\partial x_j \partial x_k} \right| \leq c_t e^{c|x|}.$$

Show that

$$\int p(t, x) \frac{1}{2} \Delta_x f(t, x) dx = \int f(t, x) \frac{1}{2} \Delta_x p(t, x) dx.$$

8. Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a one-dimensional Brownian motion. Which of the following processes are martingales?

- | | |
|---|---|
| (a) $U_t = e^{cB_t}, \quad c \in \mathbb{R};$ | (b) $V_t = tB_t - \int_0^t B_s ds;$ |
| (c) $W_t = B_t^3 - tB_t;$ | (d) $X_t = B_t^3 - 3 \int_0^t B_s ds;$ |
| (e) $Y_t = \frac{1}{3} B_t^3 - tB_t;$ | (f) $Z_t = e^{B_t - ct}, \quad c \in \mathbb{R}.$ |

9. Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM¹. Show that $X_t = \exp(aB_t + bt), t \geq 0$, is a martingale if, and only if, $a^2/2 + b > 0$.

10. Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM^d. Show that $X_t = \frac{1}{d} |B_t|^2 - t, t \geq 0$, is a martingale.

11. Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a d -dimensional stochastic process and $A, A_n, C \in \mathcal{B}(\mathbb{R}^d), n \geq 1$. Then

- $A \subset C$ implies $\tau_A^\circ \geq \tau_C^\circ$ and $\tau_A \geq \tau_C$;
- $\tau_{A \cup C}^\circ = \min\{\tau_A^\circ, \tau_C^\circ\}$ and $\tau_{A \cup C} = \min\{\tau_A, \tau_C\}$;
- $\tau_{A \cap C}^\circ = \max\{\tau_A^\circ, \tau_C^\circ\}$ and $\tau_{A \cap C} = \max\{\tau_A, \tau_C\}$;
- $A = \bigcup_{n \geq 1} A_n$, then $\tau_A^\circ = \inf_{n \geq 1} \tau_{A_n}^\circ$ and $\tau_A = \inf_{n \geq 1} \tau_{A_n}$;
- $\tau_A^\circ = \inf_{n \geq 1} (\frac{1}{n} + \tau_n^\circ)$ where $\tau_n^\circ = \inf\{s \geq 0 : X_{s+1/n} \in A\}$;
- find an example of a set A such that $\tau_A^\circ < \tau_A$.

12. Let $U \in \mathbb{R}^d$ be an open set and assume that $(X_t)_{t \geq 0}$ is a stochastic process with continuous paths. Show that $\tau_U = \tau_U^\circ$.

13. Show that the function $d(x, A) := \inf_{y \in A} |x - y|, A \subset \mathbb{R}^d$, is continuous.

14. Let τ be a stopping time. Check that \mathcal{F}_τ and $\mathcal{F}_{\tau+}$ are σ -algebras.
15. Let τ be a stopping time for the filtration $(\mathcal{F}_t)_{t \geq 0}$. Show that
- (a) $F \in \mathcal{F}_{\tau+} \iff \forall t \geq 0 : F \cap \{\tau < t\} \in \mathcal{F}_t$;
 - (b) $\{\tau \leq t\} \in \mathcal{F}_{\tau \wedge t}$ for all $t \geq 0$.
16. Let $\tau := \tau_{(-a,b)^c}^\circ$ be the first entry time of a BM^1 into the set $(-a, b)^c$.
- (a) Show that τ has finite moments $\mathbb{E} \tau^n$ of any order $n \geq 1$.
Hint: Use Example 5.2 d) and show that $\mathbb{E} e^{c\tau} < \infty$ for some $c > 0$.
 - (b) Evaluate $\mathbb{E} \int_0^\tau B_s ds$.
Hint: Find a martingale which contains $\int_0^t B_s ds$.
17. Let $(B_t)_{t \geq 0}$ be a BM^d . Find $\mathbb{E} \tau_R$ where $\tau_R = \inf \{t \geq 0 : |B_t| = R\}$.
18. Let $(B_t)_{t \geq 0}$ be a BM^1 and σ, τ be two stopping times such that $\mathbb{E} \tau, \mathbb{E} \sigma < \infty$.
- (a) Show that $\sigma \wedge \tau$ is a stopping time.
 - (b) Show that $\{\sigma \leq \tau\} \in \mathcal{F}_{\sigma \wedge \tau}$.
 - (c) Show that $\mathbb{E}(B_\tau B_\sigma) = \mathbb{E} \tau \wedge \sigma$
Hint: Consider $\mathbb{E}(B_\sigma B_\tau \mathbb{1}_{\{\sigma \leq \tau\}}) = \mathbb{E}(B_{\sigma \wedge \tau} B_\tau \mathbb{1}_{\{\sigma \leq \tau\}})$ and use optional stopping.
 - (d) Deduce from (c) that $\mathbb{E}(|B_\tau - B_\sigma|^2) = \mathbb{E} |\tau - \sigma|$.

Chapter 6

Brownian motion as a Markov process

We have seen in 2.9 that for a d -dimensional Brownian motion $(B_t)_{t \geq 0}$ and any $s > 0$ the shifted process $W_t := B_{t+s} - B_s, t \geq 0$, is again a BM^d which is independent of $(B_t)_{0 \leq t \leq s}$. Since $B_{t+s} = W_t + B_s$, we can interpret this as a renewal property: Rather than going from $0 = B_0$ straight away to $x = B_{t+s}$ in $(t + s)$ units of time, we stop after time s at $x' = B_s$ and move, using a further Brownian motion W_t for t units of time, from x' to $x = B_{t+s}$. This situation is shown in Figure 6.1:

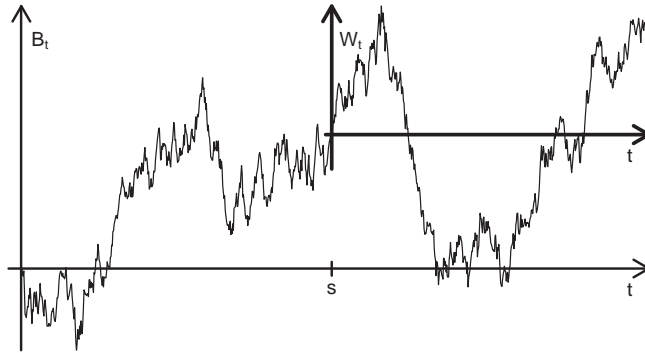


Figure 6.1. $W_t := B_{t+s} - B_s, t \geq 0$, is a Brownian motion in the new coordinate system with origin at (s, B_s) .

6.1 The Markov property

In fact, we know from Lemma 2.10 that the paths up to time s and thereafter are stochastically independent, i. e. $\mathcal{F}_s^B \perp \mathcal{F}_\infty^W := \sigma(B_{t+s} - B_s : t \geq 0)$. If we use any admissible filtration $(\mathcal{F}_t)_{t \geq 0}$ instead of the natural filtration $(\mathcal{F}_t^B)_{t \geq 0}$, the argument of Lemma 2.10

Ex. 5.2 remains valid, and we get

- Ex. 6.1 **6.1 Theorem** (Markov property of BM). *Let $(B_t)_{t \geq 0}$ be a BM^d and $(\mathcal{F}_t)_{t \geq 0}$ some*
 Ex. 6.2 *admissible filtration. For every $s > 0$, the process $W_t := B_{t+s} - B_s, t \geq 0$, is also a BM^d and $(W_t)_{t \geq 0}$ is independent of \mathcal{F}_s , i. e. $\mathcal{F}_\infty^W = \sigma(W_t, t \geq 0) \perp \mathcal{F}_s$.*

Theorem 6.1 justifies our intuition that we can split a Brownian path into two independent pieces

$$B(t + s) = B(t + s) - B(s) + B(s) = W(t) + y|_{y=B(s)}.$$

Observe that $W(t) + y$ is a Brownian motion started at $y \in \mathbb{R}^d$. Let us introduce the following notation

$$\mathbb{P}^x(B_{t_1} \in A_1, \dots, B_{t_n} \in A_n) := \mathbb{P}(B_{t_1} + x \in A_1, \dots, B_{t_n} + x \in A_n) \quad (6.1)$$

where $0 \leq t_1 < \dots < t_n$ and $A_1, \dots, A_n \in \mathcal{B}(\mathbb{R}^d)$. We will write \mathbb{E}^x for the corresponding mathematical expectation. Clearly, $\mathbb{P}^0 = \mathbb{P}$ and $\mathbb{E}^0 = \mathbb{E}$. This means that $\mathbb{P}^x(B_s \in A)$ denotes the probability that a Brownian particle starts at time $t = 0$ at the point x and travels in s units of time into the set A .

Since the finite dimensional distributions (6.1) determine the measure \mathbb{P}^x uniquely, \mathbb{P}^x is a well-defined measure on $(\Omega, \mathcal{F}_\infty)$.¹

Moreover, $x \mapsto \mathbb{E}^x u(B_t) = \mathbb{E}^0 u(B_t + x)$ is for all $u \in \mathcal{B}_b(\mathbb{R}^d)$ a measurable function. Ex. 6.3

We will need the following formulae for conditional expectations, a proof of which can be found in the appendix, Lemma A.3. Assume that $\mathcal{X} \subset \mathcal{A}$ and $\mathcal{Y} \subset \mathcal{A}$ are independent σ -algebras. If X is an $\mathcal{X}/\mathcal{B}(\mathbb{R}^d)$ and Y a $\mathcal{Y}/\mathcal{B}(\mathbb{R}^d)$ measurable random variable, then

$$\mathbb{E}[\Phi(X, Y) | \mathcal{X}] = \mathbb{E}[\Phi(x, Y)]_{x=X} = \mathbb{E}[\Phi(X, Y) | X] \quad (6.2)$$

for all bounded Borel measurable functions $\Phi : \mathbb{R}^d \times \mathbb{R}^d \rightarrow \mathbb{R}$.

If $\Psi : \mathbb{R}^d \times \Omega \rightarrow \mathbb{R}$ is bounded and $\mathcal{B}(\mathbb{R}^d) \otimes \mathcal{Y}/\mathcal{B}(\mathbb{R})$ measurable, then

$$\mathbb{E}[\Psi(X(\cdot), \cdot) | \mathcal{X}] = \mathbb{E}[\Psi(x, \cdot)]_{x=X} = \mathbb{E}[\Psi(X(\cdot), \cdot) | X]. \quad (6.3)$$

6.2 Theorem (Markov property). *Let $(B_t)_{t \geq 0}$ be a BM^d with admissible filtration $(\mathcal{F}_t)_{t \geq 0}$. Then*

$$\mathbb{E}[u(B_{t+s}) | \mathcal{F}_s] = \mathbb{E}[u(B_t + x)]_{x=B_s} = \mathbb{E}^{B_s}[u(B_t)] \quad (6.4)$$

holds for all bounded measurable functions $u : \mathbb{R}^d \rightarrow \mathbb{R}$ and

$$\mathbb{E}[\Psi(B_{\cdot+s}) | \mathcal{F}_s] = \mathbb{E}[\Psi(B_{\cdot} + x)]_{x=B_s} = \mathbb{E}^{B_s}[\Psi(B_{\cdot})] \quad (6.5)$$

holds for all bounded $\mathcal{B}(\mathbb{C})/\mathcal{B}(\mathbb{R})$ measurable functionals $\Psi : \mathcal{C}[0, \infty) \rightarrow \mathbb{R}$ which may depend on the whole Brownian path.

¹ The corresponding canonical Wiener measure μ^x from Theorem 4.2 is, consequently, concentrated on the continuous functions $w : [0, \infty) \rightarrow \mathbb{R}^d$ satisfying $w(0) = x$.

Proof. We write $u(B_{t+s}) = u(B_s + B_{t+s} - B_s) = u(B_s + (B_{t+s} - B_s))$. Then (6.4) follows from (6.2) and (B1) with $\Phi(x, y) = u(x + y)$, $X = B_s$, $Y = B_{t+s} - B_s$ and $\mathcal{X} = \mathcal{F}_s$.

In the same way we can derive (6.5) from (6.3) and 2.10 if we take $\mathcal{Y} = \mathcal{F}_\infty^W$ where $W_t := B_{t+s} - B_s$, $\mathcal{X} = \mathcal{F}_s$ and $X = B_s$. \square

6.3 Remark.

- a) Theorem 6.1 really is a result for processes with stationary and independent increments, whereas Theorem 6.2 is not necessarily restricted to such processes. In general any d -dimensional, \mathcal{F}_t adapted stochastic process $(X_t)_{t \geq 0}$ with (right-) continuous paths is called a *Markov process* if it satisfies

$$\mathbb{E}[u(X_{t+s}) | \mathcal{F}_s] = \mathbb{E}[u(X_{t+s}) | X_s] \quad \text{for all } s, t \geq 0, u \in \mathcal{B}_b(\mathbb{R}^d). \quad (6.4a)$$

Since we can express $\mathbb{E}[u(X_{t+s}) | X_s]$ as an (essentially unique) function of X_s , say $g_{u,s,t+s}(X_s)$, we have

$$\mathbb{E}[u(X_{t+s}) | \mathcal{F}_s] = \mathbb{E}^{s, X_s}[u(X_{t+s})] \quad \text{with} \quad \mathbb{E}^{s, x}[u(X_{t+s})] := g_{u,s,t+s}(x). \quad (6.4b)$$

Note that we can always select a Borel measurable version of $g_{u,s,t+s}(\cdot)$. If $g_{u,s,t+s}$ depends only on the difference $(t+s) - s = t$, we call the Markov process *homogeneous* and we write

$$\mathbb{E}[u(X_{t+s}) | \mathcal{F}_s] = \mathbb{E}^{X_s}[u(X_t)] \quad \text{with} \quad \mathbb{E}^x[u(X_t)] := g_{u,t}(x). \quad (6.4c)$$

Obviously (6.4) implies both (6.4a) and (6.4c), i.e. Brownian motion is also a Markov process in this more general sense. For homogeneous Markov processes we have

$$\mathbb{E}[\Psi(X_{t+s}) | \mathcal{F}_s] = \mathbb{E}[\Psi(X_{t+s}) | X_s] = \mathbb{E}^{X_s}[\Psi(X_t)] \quad (6.5a)$$

for bounded measurable functionals $\Psi : \mathcal{C}[0, \infty) \rightarrow \mathbb{R}$ of the sample paths. Although this relation seems to be more general than (6.4a), (6.4b), it is actually possible to derive (6.5a) directly from (6.4a), (6.4b) by standard measure-theoretic techniques.



- b) It is useful to write (6.4) in integral form and with all ω :

$$\begin{aligned} \mathbb{E}[u(B_{t+s}) | \mathcal{F}_s](\omega) &= \int_{\Omega} u(B_t(\omega') + B_s(\omega)) \mathbb{P}(d\omega') \\ &= \int_{\Omega} u(B_t(\omega')) \mathbb{P}^{B_s(\omega)}(d\omega') = \mathbb{E}^{B_s(\omega)}[u(B_t)]. \end{aligned}$$

Markov processes are often used in the literature. In fact, the Markov property (6.4a), (6.4b) can be seen as a means to obtain the finite dimensional distributions from one-dimensional distributions. We will show this for a Brownian motion $(B_t)_{t \geq 0}$

where we only use the property (6.4). The guiding idea is as follows: If a particle starts at x and is by time t in a set A and by time u in C , we can realize this as

$$x \xrightarrow[\text{Brownian motion}]{\text{time } t} A \quad \left\{ \begin{array}{l} \text{stop there at } B_t = a, \\ \text{then restart from } a \in A \end{array} \right\} \xrightarrow[\text{Brownian motion}]{\text{time } u-t} C.$$

Let us become a bit more rigorous. Let $0 \leq s < t < u$ and let $\phi, \psi, \chi : \mathbb{R}^d \rightarrow \mathbb{R}$ bounded measurable functions. Then

$$\begin{aligned} \mathbb{E} [\phi(B_s) \psi(B_t) \chi(B_u)] &= \mathbb{E} [\phi(B_s) \psi(B_t) \mathbb{E} (\chi(B_u) | \mathcal{F}_t)] \\ &\stackrel{(6.4)}{=} \mathbb{E} [\phi(B_s) \psi(B_t) \mathbb{E}^{B_t} \chi(B_{u-t})] \\ &\stackrel{(6.4)}{=} \mathbb{E} [\phi(B_s) \underbrace{\mathbb{E}^{B_s} (\psi(B_{t-s}) \mathbb{E}^{B_{t-s}} \chi(B_{u-t}))}_{= \int \psi(a) \int \chi(b) \mathbb{P}^a(B_{u-t} \in db) \mathbb{P}^{B_s}(B_{t-s} \in da)}] \\ &= \int \psi(a) \int \chi(b) \mathbb{P}^a(B_{u-t} \in db) \mathbb{P}^{B_s}(B_{t-s} \in da) \mathbb{P}(B_s \in dx). \end{aligned}$$

If we rewrite all expectations as integrals we get

$$\begin{aligned} &\iiint \phi(x) \psi(y) \chi(z) \mathbb{P}(B_s \in dx, B_t \in dy, B_u \in dz) \\ &= \int \left[\phi(c) \int \psi(a) \int \chi(b) \mathbb{P}^a(B_{u-t} \in db) \mathbb{P}^c(B_{t-s} \in da) \right] \mathbb{P}(B_s \in dc). \end{aligned}$$

This shows that we find the joint distribution of the vector (B_s, B_t, B_u) if we know the one-dimensional distributions of each B_t . For $s = 0$ and $\phi(0) = 1$ these relations are also known as *Chapman-Kolmogorov equations*. By iteration we get

6.4 Theorem. *Let $(B_t)_{t \geq 0}$ be a BM^d (or a general Markov process), $x_0 = 0$ and $t_0 = 0 < t_1 < \dots < t_n$. Then*

$$\mathbb{P}^0(B_{t_1} \in dx_1, \dots, B_{t_n} \in dx_n) = \prod_{j=1}^n \mathbb{P}^{x_{j-1}}(B_{t_j - t_{j-1}} \in dx_j). \quad (6.6)$$

6.2 The strong Markov property

Let us now show that (6.4) remains valid if we replace s by a stopping time $\sigma(\omega)$. Since stopping times can attain the value $+\infty$, we have to take care that B_σ is defined even in this case. If $(X_t)_{t \geq 0}$ is a stochastic process and σ a stopping time, we set

$$X_\sigma(\omega) := \begin{cases} X_{\sigma(\omega)}(\omega) & \text{if } \sigma(\omega) < \infty, \\ \lim_{t \rightarrow \infty} X_t(\omega) & \text{if } \sigma(\omega) = \infty \text{ and if the limit exists,} \\ 0 & \text{in all other cases.} \end{cases}$$

6.5 Theorem (Strong Markov property). *Let $(B_t)_{t \geq 0}$ be a BM^d with admissible filtration $(\mathcal{F}_t)_{t \geq 0}$ and $\sigma < \infty$ some a.s. finite stopping time. Then $(W_t)_{t \geq 0}$, where $W_t := B_{\sigma+t} - B_\sigma$, is again a BM^d , which is independent of $\mathcal{F}_{\sigma+}$.*

Proof. We know, cf. Lemma A.16, that $\sigma_j := (\lfloor 2^j \sigma \rfloor + 1)/2^j$ is a decreasing sequence of stopping times such that $\inf_{j \geq 1} \sigma_j = \sigma$. For all $0 \leq s < t$, $\xi \in \mathbb{R}^d$ and all $F \in \mathcal{F}_{\sigma+}$ we find by the continuity of the sample paths

$$\begin{aligned}
& \mathbb{E} \left[e^{i \langle \xi, B_{\sigma+t} - B_{\sigma+s} \rangle} \mathbb{1}_F \right] \\
&= \lim_{j \rightarrow \infty} \mathbb{E} \left[e^{i \langle \xi, B_{\sigma_j+t} - B_{\sigma_j+s} \rangle} \mathbb{1}_F \right] \\
&= \lim_{j \rightarrow \infty} \sum_{k=1}^{\infty} \mathbb{E} \left[e^{i \langle \xi, B_{k2^{-j}+t} - B_{k2^{-j}+s} \rangle} \mathbb{1}_{\{\sigma_j = k2^{-j}\}} \cdot \mathbb{1}_F \right] \\
&= \lim_{j \rightarrow \infty} \sum_{k=1}^{\infty} \mathbb{E} \left[\overbrace{e^{i \langle \xi, B_{k2^{-j}+t} - B_{k2^{-j}+s} \rangle}}^{\perp \mathcal{F}_{k2^{-j}}, \sim B_{t-s} \text{ by (B1)/(5.1), (B2)}} \underbrace{\mathbb{1}_{\{(k-1)2^{-j} \leq \sigma < k2^{-j}\}}}_{\in \mathcal{F}_{k2^{-j}} \text{ as } F \in \mathcal{F}_{\sigma+}} \mathbb{1}_F \right] \\
&= \lim_{j \rightarrow \infty} \sum_{k=1}^{\infty} \mathbb{E} \left[e^{i \langle \xi, B_{t-s} \rangle} \right] \mathbb{P} \left(\{(k-1)2^{-j} \leq \sigma < k2^{-j}\} \cap F \right) \\
&= \mathbb{E} \left[e^{i \langle \xi, B_{t-s} \rangle} \right] \mathbb{P}(F).
\end{aligned}$$

In the last equality we used $\bigcup_{k=1}^{\infty} \{(k-1)2^{-j} \leq \sigma < k2^{-j}\} = \{\sigma < \infty\}$ for all $j \geq 1$.

The same calculation applies to finitely many increments. Let $t_0 = 0 < t_1 < \dots < t_n$ and $\xi_1, \dots, \xi_n \in \mathbb{R}^d$ and $F \in \mathcal{F}_{\sigma+}$. Then

$$\mathbb{E} \left[e^{i \sum_{j=1}^n \langle \xi_j, B_{\sigma+t_j} - B_{\sigma+t_{j-1}} \rangle} \mathbb{1}_F \right] = \prod_{j=1}^n \mathbb{E} \left[e^{i \langle \xi_j, B_{t_j} - B_{t_{j-1}} \rangle} \right] \mathbb{P}(F).$$

This shows that the increments $B_{\sigma+t_j} - B_{\sigma+t_{j-1}}$ are independent and distributed like $B_{t_j-t_{j-1}}$. Moreover, all increments are independent of $F \in \mathcal{F}_{\sigma+}$. Using the same argument as in the proof of Theorem 6.1 shows that all random vectors of the form $(B_{\sigma+t_1} - B_\sigma, \dots, B_{\sigma+t_n} - B_{\sigma+t_{n-1}})$ and, consequently, $\mathcal{F}_\infty^W := \sigma(B_{\sigma+t} - B_\sigma, t \geq 0)$ are independent of $\mathcal{F}_{\sigma+}$. \square

We can now use (6.2), (6.3) (or Lemma A.3) to deduce from Theorem 6.5 the following consequences of the strong Markov property.

6.6 Theorem. *Let $(B_t)_{t \geq 0}$ be a BM^d and σ be a stopping time. Then we have for all $t \geq 0$, $u \in \mathcal{B}_b(\mathbb{R}^d)$ and \mathbb{P} almost all $\omega \in \{\sigma < \infty\}$* Ex. 6.4

$$\mathbb{E}[u(B_{t+\sigma}) | \mathcal{F}_{\sigma+}](\omega) = \mathbb{E}[u(B_t + x)]|_{x=B_\sigma(\omega)} = \mathbb{E}^{B_\sigma(\omega)} u(B_t). \quad (6.7)$$

For all bounded $\mathcal{B}(\mathbb{C})/\mathcal{B}(\mathbb{R})$ measurable functionals $\Psi : \mathbb{C}[0, \infty) \rightarrow \mathbb{R}$ which may depend on a whole Brownian path and \mathbb{P} almost all $\omega \in \{\sigma < \infty\}$ this becomes

$$\mathbb{E}[\Psi(B_{\cdot+\sigma}) | \mathcal{F}_{\sigma+}] = \mathbb{E}[\Psi(B_{\cdot} + x)]|_{x=B_\sigma} = \mathbb{E}^{B_\sigma}[\Psi(B_{\cdot})]. \quad (6.8)$$

Let $\tau = \tau(B_{\cdot})$ be a stopping time which can be expressed as a functional of a Brownian path, e. g. a first hitting time, and assume that σ is a further stopping time such that $\sigma \leq \tau$ a. s. Denote by $\tau' = \tau(B_{\cdot+\sigma})$ the stopping time τ for the shifted process $B_{\cdot+\sigma}$, which is the remaining time, counting from σ , until the event described by τ happens. Set $W_{\cdot} := B_{\cdot+\sigma} - B_\sigma$; then $\tau' = \tau(W_{\cdot} + B_\sigma)$, and the functionals $u(B_{\tau})$ and $u(W_{\tau'} + B_\sigma)$ have the same distribution. Thus, (6.8) implies the following result.

6.7 Corollary. *Let $(B_t)_{t \geq 0}$ be a BM^d , $\tau = \tau(B_{\cdot})$ a first hitting time and σ a stopping time such that $\sigma \leq \tau$ a. s. Set $\tau' = \tau(B_{\cdot+\sigma})$ and $W_{\cdot} := B_{\cdot+\sigma} - B_\sigma$. Then*

$$\mathbb{E}[u(B_{\tau}) | \mathcal{F}_{\sigma+}](\omega) = \mathbb{E}[u(W_{\tau'} + x)]|_{x=B_\sigma(\omega)} = \mathbb{E}^{B_\sigma(\omega)} u(W_{\tau'}) \quad (6.9)$$

holds for all $u \in \mathcal{B}_b(\mathbb{R}^d)$ and \mathbb{P} almost all $\omega \in \{\tau < \infty\}$.

Often one identifies the Brownian motions B and W appearing in (6.9), and writes only B .

6.8 Remark. Let $(X_t)_{t \geq 0}$ be an \mathcal{F}_t adapted, d -dimensional stochastic process with (right-)continuous paths. If $(X_t)_{t \geq 0}$ satisfies for all $t \geq 0$ and $u \in \mathcal{B}_b(\mathbb{R}^d)$

$$\mathbb{E}[u(X_{\sigma+t}) | \mathcal{F}_{\sigma+}] = \mathbb{E}[u(X_{\sigma+t}) | X_\sigma] \quad \text{on } \{\sigma < \infty\}, \quad (6.10)$$

it is called a *strong Markov process*. For a Brownian motion this follows from (6.7). Just as for the simple Markov property, the (general analogue of) property (6.8) seems only stronger than (6.7); as a matter of fact one can derive it directly from (6.7) using standard techniques.

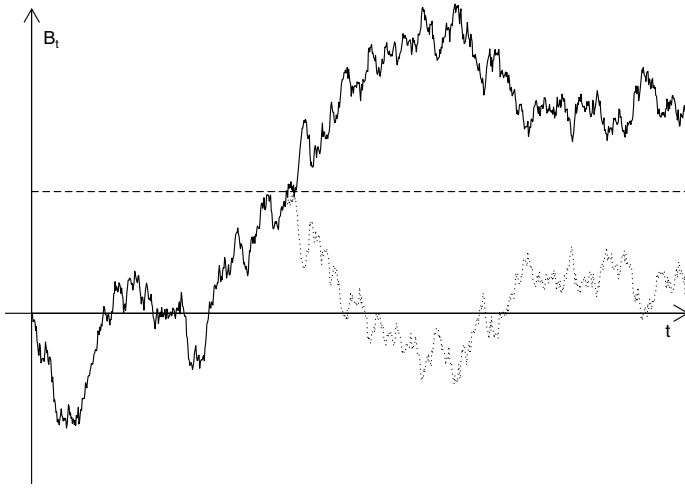


Figure 6.2. Reflection upon reaching the level b for the first time.

6.3 Desiré André's reflection principle

Let us briefly discuss one of the most famous applications of the strong Markov property of a Brownian motion. Let $(B_t)_{t \geq 0}$ be a one-dimensional Brownian motion and $\tau_b = \inf\{t \geq 0 : B_t = b\}$ the first passage time for b . Assume that B has reached the level b for the first time and that $\tau_b < t$. Stop at τ_b and start anew from $B_{\tau_b} = b$. Since $W_s + b = B_{\tau_b+s} - B_{\tau_b} + b = B_{\tau_b+s}$ is, by Theorem 6.5, again a Brownian motion (started at b), the probability to be at time t above or below b is the same, cf. Figure 6.2. Since $B_{\tau_b} = b$ and since τ_b is $\mathcal{F}_{\tau_b}^B$ measurable, cf. Lemma A.15, we have

$$\mathbb{P}(\tau_b \leq t, B_t < b) = \mathbb{P}(\underbrace{\{\tau_b \leq t\}}_{\in \mathcal{F}_{\tau_b}^B} \cap \underbrace{\{B_{\tau_b+(t-\tau_b)} - B_{\tau_b} < 0\}}_{\substack{\in \mathcal{F}_{\infty}^W \perp \mathcal{F}_{\tau_b}^B}}) = \frac{1}{2} \mathbb{P}(\tau_b \leq t)$$

and so

$$\begin{aligned} \mathbb{P}(\tau_b \leq t) &= \mathbb{P}(\tau_b \leq t, B_t \geq b) + \mathbb{P}(\tau_b \leq t, B_t < b) \\ &= \mathbb{P}(\tau_b \leq t, B_t \geq b) + \frac{1}{2} \mathbb{P}(\tau_b \leq t) \\ &= \mathbb{P}(B_t \geq b) + \frac{1}{2} \mathbb{P}(\tau_b \leq t). \end{aligned}$$

This shows that $\mathbb{P}(\tau_b \leq t) = 2 \mathbb{P}(B_t \geq b)$. Observe that $\{\tau_b \leq t\} = \{M_t \geq b\}$ where $M_t := \sup_{s \leq t} B_s$. Therefore,

$$\mathbb{P}(M_t \geq b) = \mathbb{P}(\tau_b \leq t) = 2 \mathbb{P}(B_t \geq b) \quad \text{for all } b > 0. \quad (6.11)$$

From this we can calculate the probability distributions of various functionals of a Brownian motion.

6.9 Theorem (Lévy 1939). *Let $(B_t)_{t \geq 0}$ be a BM^1 , $b \in \mathbb{R}$, and set*

Ex. 6.5

$$\tau_b = \inf\{t \geq 0 : B_t = b\}, \quad M_t := \sup_{s \leq t} B_s, \quad \text{and} \quad m_t := \inf_{s \leq t} B_s.$$

Ex. 6.6

Ex. 6.7

Then, for all $x \geq 0$,

$$M_t \sim |B_t| \sim -m_t \sim M_t - B_t \sim B_t - m_t \sim \sqrt{\frac{2}{\pi t}} e^{-x^2/(2t)} dx \quad (6.12)$$

and for $t > 0$

$$\tau_b \sim \frac{|b|}{\sqrt{2\pi t^3}} e^{-b^2/(2t)} dt. \quad (6.13)$$

Proof. The equalities (6.11) immediately show that $M_t \sim |B_t|$. Using the symmetry 2.8 of a Brownian motion we get $m_t = \inf_{s \leq t} B_s = -\sup_{s \leq t} (-B_s) \sim -M_t$. The invariance of BM^1 under time reversals, cf. 2.11, gives

$$M_t - B_t = \sup_{s \leq t} (B_s - B_t) = \sup_{s \leq t} (B_{t-s} - B_t) \sim \sup_{s \leq t} B_s = M_t,$$

and if we combine this again with the symmetry 2.8 we get $M_t - B_t \sim B_t - m_t$.

The formula (6.13) follows for $b > 0$ from the fact that $\{\tau_b \geq t\} = \{M_t \leq b\}$ if we differentiate the relation

$$\mathbb{P}(\tau_b \geq t) = \mathbb{P}(M_t \leq b) = \sqrt{\frac{2}{\pi t}} \int_0^b e^{-x^2/(2t)} dx.$$

For $b < 0$ we use the symmetry of a Brownian motion. □

6.10 Remark. (6.12) tells us that for each single instant of time the random variables, e.g. M_t and $|B_t|$, have the same distribution. This is no longer true if we consider the finite dimensional distributions, i.e. the laws of processes $(M_t)_{t \geq 0}$ and $(|B_t|)_{t \geq 0}$ are different. In our example this is obvious since $t \mapsto M_t$ is a.s. increasing while $t \mapsto |B_t|$ is positive but not necessarily monotone. On the other hand, the processes $-m$ and M are equivalent, i.e. they do have the same law.



In order to derive (6.11) we used the Markov property of a Brownian motion, i.e. the fact that \mathcal{F}_s^B and \mathcal{F}_∞^W , $W_t = B_{t+s} - B_s$ are independent. Often the proof is based

on (6.7) but this leads to the following difficulty:

$$\begin{aligned}
 \mathbb{P}(\tau_b \leq t, B_t < b) &= \mathbb{E}(\mathbb{1}_{\{\tau_b \leq t\}} \mathbb{E}[\mathbb{1}_{(-\infty, 0)}(B_t - b) \mid \mathcal{F}_{\tau_b}^B]) \\
 &\stackrel{?!}{=} \mathbb{E}(\mathbb{1}_{\{\tau_b \leq t\}} \mathbb{E}^{B_{\tau_b}}[\mathbb{1}_{(-\infty, 0)}(B_{t-\tau_b} - b)]) \\
 &= \mathbb{E}(\mathbb{1}_{\{\tau_b \leq t\}} \mathbb{E}^b[\mathbb{1}_{(-\infty, 0)}(B_{t-\tau_b} - b)]) \\
 &= \mathbb{E}(\mathbb{1}_{\{\tau_b \leq t\}} \underbrace{\mathbb{E}[\mathbb{1}_{(-\infty, 0)}(B_{t-\tau_b})]}_{=1/2}).
 \end{aligned}$$



Notice that in the step marked by ‘?!’ there appears the expression $t - \tau_b$. As it is written, we do not know whether $t - \tau_b$ is positive – and if it were, our calculation would certainly not be covered by (6.7).

In order to make such arguments work we need the following stronger version of the strong Markov property. Mind the dependence of the random times on the fixed path ω in the expression below!

6.11 Theorem. *Let $(B_t)_{t \geq 0}$ be a BM^d , τ an \mathcal{F}_t^B stopping time and $\eta \geq \tau$ where η is an $\mathcal{F}_{\tau+}^B$ measurable random time. Then we have for all $\omega \in \{\eta < \infty\}$ and all $u \in \mathcal{B}_b(\mathbb{R}^d)$*

$$\begin{aligned}
 \mathbb{E}[u(B_\eta) \mid \mathcal{F}_{\tau+}^B](\omega) &= \mathbb{E}^{B_\tau(\omega)}[u(B_{\eta(\omega)-\tau(\omega)}(\cdot))] \\
 &= \int u(B_{\eta(\omega)-\tau(\omega)}(\omega')) \mathbb{P}^{B_\tau(\omega)}(d\omega').
 \end{aligned} \tag{6.14}$$

The requirement that τ is a stopping time and η is a $\mathcal{F}_{\tau+}$ measurable random time with $\eta \geq \tau$ looks artificial. Typical examples of (τ, η) pairs are $(\tau, \tau + t)$, $(\tau, \tau \vee t)$ or $(\sigma \wedge t, t)$ where σ is a further stopping time.

Proof of Theorem 6.11. Replacing $u(B_\eta)$ by $u(B_\eta)\mathbb{1}_{\{\eta < \infty\}}$ we may assume $\eta < \infty$. Set $\sigma_j := (\lfloor 2^j(\eta - \tau) \rfloor + 1)/2^j$. Then $\inf_{j \geq 1} \sigma_j = \eta - \tau$ and σ_j is clearly $\mathcal{F}_{\tau+}^B$ measurable, since η and τ are $\mathcal{F}_{\tau+}^B$ measurable.

For all $F \in \mathcal{F}_{\tau+}^B$ we know that $F \cap \{\sigma_j = k/2^j\} \in \mathcal{F}_{\tau+}^B$, i. e.

$$\begin{aligned} \int_F u(B_{\tau+\sigma_j}) d\mathbb{P} &= \sum_{k=1}^{\infty} \int_{F \cap \{\sigma_j = k/2^j\}} u(B_{\tau+k/2^j}) d\mathbb{P} \\ &= \sum_{k=1}^{\infty} \int_{F \cap \{\sigma_j = k/2^j\}} \mathbb{E}[u(B_{\tau+k/2^j}) \mid \mathcal{F}_{\tau+}^B] d\mathbb{P} \\ &\stackrel{(6.7)}{=} \sum_{k=1}^{\infty} \int_{F \cap \{\sigma_j = k/2^j\}} \mathbb{E}^{B_{\tau}(\omega)} u(B_{k/2^j}) \mathbb{P}(d\omega) \\ &= \int_F \mathbb{E}^{B_{\tau}(\omega)} u(B_{\sigma_j(\omega)}) \mathbb{P}(d\omega). \end{aligned}$$

Since $t \mapsto B_t$ is continuous, we have $\lim_{j \rightarrow \infty} B_{\sigma_j} = B_{\eta-\tau}$, and we get (6.14) for $u \in \mathcal{C}_b(\mathbb{R}^d)$ using dominated convergence.

For every compact set $K \subset \mathbb{R}^d$ we can approximate $\mathbb{1}_K$ from above by a sequence of continuous functions. Therefore, Ex. 6.10

$$\int_F \mathbb{1}_K(B_{\eta}) d\mathbb{P} = \int_F \mathbb{E}^{B_{\tau}(\omega)} \mathbb{1}_K(B_{\eta(\omega)-\tau(\omega)}) \mathbb{P}(d\omega).$$

As the compact sets generate the Borel σ -algebra, we can use the uniqueness theorem for measures to extend the above equality to all functions of the form $\mathbb{1}_K$ where $K \in \mathcal{B}(\mathbb{R}^d)$; then we get (6.14) for all measurable step functions and a Beppo–Levi argument finally shows (6.14) for $\mathcal{B}_b(\mathbb{R}^d)$. \square

The proof of the reflection principle uses only the strong Markov property and the symmetry of a Brownian motion. Therefore, we can rephrase the reflection principle in the following more general version. Let $(B(t), \mathcal{F}_t)_{t \geq 0}$ be a Brownian motion and τ an a. s. finite stopping time. We consider the process

$$W(t, \omega) := \begin{cases} B(t, \omega), & \text{if } 0 \leq t < \tau(\omega) \leq +\infty, \\ 2B(\tau(\omega), \omega) - B(t, \omega), & \text{if } \tau(\omega) \leq t < \infty, \end{cases} \quad (6.15)$$

see Figure 6.3.

Observe that the trajectories of B and W coincide in the interval $[0, \tau(\omega))$. On $[\tau(\omega), \infty)$ we have $W(t, \omega) = B(\tau(\omega), \omega) - (B(t, \omega) - B(\tau(\omega), \omega))$, i. e. every trajectory of W is the reflection of the original trajectory with respect to the axis $y = B(\tau(\omega), \omega)$.

From Theorem 6.5 we know that $(B(\tau) - B(\tau + t))_{t \geq 0}$ is a Brownian motion which is independent of the stopped process $(B(t \wedge \tau))_{t \geq 0}$. Therefore it is plausible that the process W , which is a concatenation of these two processes, is again a Brownian motion.

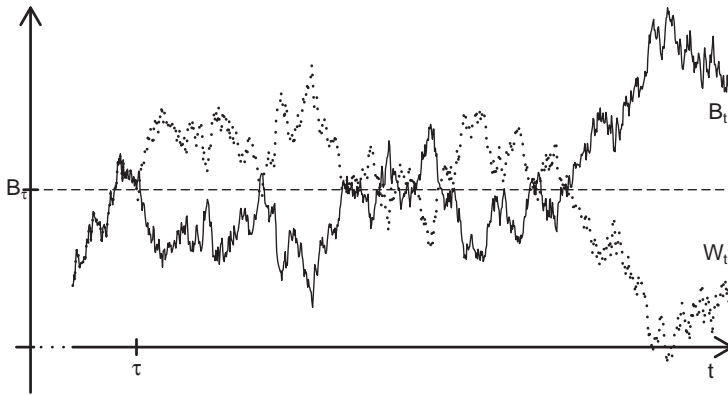


Figure 6.3. Both B_t and the reflection W_t are Brownian motions.

6.12 Theorem. Let $(B(t), \mathcal{F}_t)_{t \geq 0}$ be a BM^d and let $(W(t))_{t \geq 0}$ the process given by (6.15). Then $(W(t))_{t \geq 0}$ is again a Brownian motion.

Proof. Without restriction we may assume that $\tau(\omega) < \infty$ for all $\omega \in \Omega$. The mapping $\Phi : \mathcal{C}_{(0)} \times [0, \infty) \times \mathcal{C}_{(0)} \rightarrow \mathcal{C}_{(0)}$, where $\mathcal{C}_{(0)} = \{f \in \mathcal{C}[0, \infty) : f(0) = 0\}$, defined by

$$\Phi(f, t, g) := \begin{cases} f(s), & \text{if } 0 \leq s < t, \\ f(t) + g(s - t), & \text{if } t \leq s < \infty, \end{cases}$$

is continuous if we equip $\mathcal{C}_{(0)} \times [0, \infty) \times \mathcal{C}_{(0)}$ with the topology of locally uniform convergence (in $\mathcal{C}_{(0)}$) and pointwise convergence (in $[0, \infty)$); in particular, the mapping Φ is $\mathcal{B}(\mathcal{C}_{(0)}) \otimes \mathcal{B}[0, \infty) \otimes \mathcal{B}(\mathcal{C}_{(0)}) / \mathcal{B}(\mathcal{C}_{(0)})$ measurable.

By Lemma 6.13, $\omega \mapsto (B(\cdot \wedge \tau(\omega)), \tau(\omega))$ is $\mathcal{F}_{\tau+} / \mathcal{B}(\mathcal{C}_{(0)}) \otimes \mathcal{B}[0, \infty)$ measurable. Moreover, $\omega \mapsto B(\cdot \wedge \tau(\omega), \omega) - B(\tau(\omega), \omega)$ is independent of $\mathcal{F}_{\tau+}$ and has the same probability distribution as $\omega \mapsto B(\cdot, \omega)$. Because of the symmetry of the Wiener measure, $B(\cdot \wedge \tau) - B(\tau) \sim -B(\cdot \wedge \tau) - B(\tau)$, and we see that

$$W = \Phi(B(\cdot \wedge \tau), \tau, B(\tau) - B(\cdot + \tau)) \quad \text{and} \quad B = \Phi(B(\cdot \wedge \tau), \tau, B(\cdot + \tau) - B(\tau))$$

have the same law. Since W has, by construction, continuous paths, the claim follows. \square

Ex. 6.4 6.13 Lemma (cf. Corollary 6.25). Let $(B(t), \mathcal{F}_t)_{t \geq 0}$ be a BM^1 and $\tau < \infty$ an a. s. finite stopping time. Then $B(t \wedge \tau)$ is $\mathcal{F}_{\tau+}$ measurable for every $t \geq 0$.

Proof. We have to show that $\{B(t \wedge \tau) \in U\} \cap \{\tau < s\} \in \mathcal{F}_s$ for all Borel sets $U \subset \mathbb{R}$ and $s \geq 0$, cf. Definition A.14. It is clearly enough to consider only open sets U and dyadic numbers $s = k2^{-n}$. From Lemma A.16 we know that $\tau_j := (\lfloor 2^j \tau \rfloor + 1)/2^j$ is a decreasing sequence of stopping times such that $\inf_{j \geq 1} \tau_j = \tau$. Since $t \mapsto B(t, \omega)$ is continuous and U open, we see for all $n \geq 1$

$$\{B(t \wedge \tau) \in U\} = \bigcup_{m \geq n} \bigcap_{l \geq m} \{B(t \wedge \tau_l) \in U\}.$$

By definition, $\{\tau < k2^{-n}\} = \{\tau_n \leq k2^{-n}\} = \{\tau_m \leq k2^{-n}\}$ for all $m \geq n$. Therefore,

$$\{B(t \wedge \tau) \in U\} \cap \{\tau < k2^{-n}\} = \bigcup_{m \geq n} \bigcap_{l \geq m} \underbrace{\left(\{B(t \wedge \tau_l) \in U\} \cap \{\tau_l \leq k2^{-n}\} \right)}_{=A(l,k,n)}.$$

As $A(l, k, n) = \bigcup_{j: j2^{-l} \leq k2^{-n}} \{B(t \wedge j2^{-l}) \in U\} \cap \{\tau_l = j2^{-l}\} \in \mathcal{F}_{k2^{-n}}$, we see $\{B(t \wedge \tau) \in U\} \cap \{\tau < k2^{-n}\} \in \mathcal{F}_{k2^{-n}}$. \square

6.4 Transience and recurrence

In Section 5.2 we used Wald's identities to obtain the distribution of a BM^1 when exiting from the interval $(-a, b)$, cf. Corollary 5.11. Without much effort we can extend this to continuous martingales.

6.14 Corollary. *Let $(M_t^x, \mathcal{F}_t)_{t \geq 0}$ be a real-valued martingale with continuous paths and $M_0 = x \in \mathbb{R}$ a.s. If the first exit time $\tau = \tau_{(r,R)^c}$ from the interval (r, R) , $r < R$, is a.s. finite, then we have for all $x \in (r, R)$*

$$\mathbb{P}(M_\tau^x = r) = \frac{R - x}{R - r} \quad \text{and} \quad \mathbb{P}(M_\tau^x = R) = \frac{x - r}{R - r}.$$

If we combine Corollary 6.14 with Theorem 5.6 for a suitable function f we obtain the exit probabilities of a BM^d from an annulus $\mathbb{B}(0, R) \setminus \mathbb{B}(0, r)$.

6.15 Theorem. *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a Brownian motion with right-continuous filtration, $r < R$ and denote by $\tau_{\mathbb{B}(0,r)}$ and $\tau_{\mathbb{B}^c(0,R)}$ the first hitting times of the sets $\mathbb{B}(0, r)$ and*

$\mathbb{B}^c(0, R)$. Then, for all $r < |x| < R$,

$$\mathbb{P}^x(\tau_{\mathbb{B}(0,r)} < \tau_{\mathbb{B}^c(0,R)}) = \begin{cases} \frac{R - |x|}{R - r}, & \text{if } d = 1, \\ \frac{\log R - \log |x|}{\log R - \log r}, & \text{if } d = 2, \\ \frac{R^{2-d} - |x|^{2-d}}{R^{2-d} - r^{2-d}}, & \text{if } d \geq 3. \end{cases} \quad (6.16)$$

Proof. Note that $\tau := \tau_{\mathbb{B}(0,r)} \wedge \tau_{\mathbb{B}^c(0,R)}$ is the first exit time from $\mathbb{B}(0, R) \setminus \mathbb{B}(0, r)$. Therefore, (6.16) follows, in dimension $d = 1$, directly from Corollary 6.14 where we set $M_t = B_t + x$.

If $d = 2$, we use Theorem 5.6 with $f(|x|) = -\log |x|$ and $0 < r \leq |x| \leq R$ (the attentive reader might first want to multiply f by a cut-off function $\chi_{r,R} \in \mathcal{C}_c(\mathbb{R}^2)$ such that $\mathbb{1}_{\mathbb{B}(0,r)} \leq \chi_{r,R} \leq \mathbb{1}_{\mathbb{B}(0,R)}$). Since

$$\frac{\partial f(|x|)}{\partial x_j} = \frac{-x_j}{|x|^2} \quad \text{and} \quad \frac{\partial^2 f(|x|)}{\partial x_j^2} = \frac{2x_j^2 - |x|^2}{|x|^4}, \quad j = 1, 2,$$

we find $\Delta f(x) \equiv 0$ on $\mathbb{B}(0, R) \setminus \mathbb{B}(0, r)$. Therefore, $M_t = f(|B_t + x|)$ is a martingale with continuous paths. Moreover, $|B_t|^2 = (B_t^1)^2 + (B_t^2)^2 \geq (B_t^1)^2$ where the coordinate processes $(B_t^j)_{t \geq 0}$ are one-dimensional Brownian motions; since any one-dimensional Brownian motion leaves every interval $[-R, R]$ with probability one, cf. Corollary 5.11, we see that $\tau < \infty$ a.s. We can now apply Corollary 6.14 and conclude that

$$\mathbb{P}^x(\tau_{\mathbb{B}(0,r)} < \tau_{\mathbb{B}^c(0,R)}) = \frac{f(R) - f(x)}{f(R) - f(r)} = \frac{\log R - \log |x|}{\log R - \log r}.$$

If $d = 3$, we can use the same argument with $f(|x|) = |x|^{2-d}$. Now $\Delta f(|x|) = 0$ since for $j = 1, 2, \dots, d$

$$\frac{\partial f(|x|)}{\partial x_j} = \frac{(2-d)x_j}{|x|^d} \quad \text{and} \quad \frac{\partial^2 f(|x|)}{\partial x_j^2} = \frac{(2-d)|x|^2 - d(2-d)x_j^2}{|x|^{d+2}}. \quad \square$$

The results of Theorem 6.15 become especially interesting for degenerate annuli, i.e. for $r \rightarrow 0$ and $R \rightarrow \infty$. Letting first $r \rightarrow 0$ and then $R \rightarrow \infty$ yields

$$\mathbb{P}^x(\tau_{\mathbb{B}(0,r)} < \tau_{\mathbb{B}^c(0,R)}) \xrightarrow{r \rightarrow 0} \mathbb{P}^x(\tau_{\{0\}} \leq \tau_{\mathbb{B}^c(0,R)}) \xrightarrow{R \rightarrow \infty} \mathbb{P}^x(\tau_{\{0\}} < \infty).$$

If we consider only $R \rightarrow \infty$ we get

$$\mathbb{P}^x(\tau_{\mathbb{B}(0,r)} < \tau_{\mathbb{B}^c(0,R)}) \xrightarrow{R \rightarrow \infty} \mathbb{P}^x(\tau_{\mathbb{B}(0,r)} < \infty).$$

Since we know from Theorem 6.15 the values for the respective probabilities, we arrive at

6.16 Corollary. *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM^d with a right-continuous filtration. Then we have for all $|x| \geq r > 0$*

$$\mathbb{P}^x (\tau_{\{0\}} < \infty) = \begin{cases} 1, & \text{if } d = 1, \\ 0, & \text{if } d \geq 2, \end{cases} \quad (6.17)$$

$$\mathbb{P}^x (\tau_{\mathbb{B}(0,r)} < \infty) = \begin{cases} 1, & \text{if } d = 1, 2, \\ \left(\frac{|x|}{r}\right)^{2-d}, & \text{if } d \geq 3. \end{cases} \quad (6.18)$$

If $x = 0$, $\tau_{\{0\}}$ is the first return time to 0, and (6.17) is still true:

$$\{\tau_{\{0\}} < \infty\} = \{\exists t > 0 : B_t = 0\} = \bigcup_{n \geq 1} \{\exists t > 1/n : B_t = 0\}.$$

By the Markov property, and the fact that $\mathbb{P}^0(B_{1/n} = 0) = 0$, we have

$$\mathbb{P}^0(\exists t > 1/n : B_t = 0) = \mathbb{E}^0 \mathbb{P}^{B_{1/n}}(\exists t > 0 : B_t = 0) = \mathbb{E}^0 \underbrace{\mathbb{P}^{B_{1/n}}(\tau_{\{0\}} < \infty)}_{=0 \text{ or } 1 \text{ by (6.14)}}$$

which is 0 or 1 according to $d = 1$ or $d \geq 2$.

By the (strong) Markov property, a re-started Brownian motion is again a Brownian motion. This is the key to the following result.

6.17 Corollary. *Let $(B(t))_{t \geq 0}$ be a BM^d and $x \in \mathbb{R}^d$.*

a) *If $d = 1$, Brownian motion is point recurrent, i. e.*

$$\mathbb{P}^x (B(t) = 0 \text{ infinitely often}) = 1.$$

b) *If $d = 2$, Brownian motion is neighbourhood recurrent, i. e.*

$$\mathbb{P}^x (B(t) \in \mathbb{B}(0, r) \text{ infinitely often}) = 1 \text{ for every } r > 0.$$

c) *If $d \geq 3$, Brownian motion is transient, i. e. $\mathbb{P}^x (\lim_{t \rightarrow \infty} |B(t)| = \infty) = 1$.*

Proof. Let $d = 1, 2$, $r > 0$ and denote by $\tau_{\mathbb{B}(0,r)}$ the first hitting time of $\mathbb{B}(0, r)$ for $(B(t))_{t \geq 0}$. We know from Corollary 6.16 that $\tau_{\mathbb{B}(0,r)}$ is a. s. finite. By the strong Markov property, Theorem 6.6, we find

$$\mathbb{P}^0 (\exists t > 0 : B(t + \tau_{\mathbb{B}^c(0,r)}) \in \mathbb{B}(0, r)) = \mathbb{E}^0 \left(\mathbb{P}^{B(\tau_{\mathbb{B}^c(0,r)})} (\tau_{\mathbb{B}(0,r)} < \infty) \right) \stackrel{(6.18)}{=} 1.$$

This means that $(B(t))_{t \geq 0}$ visits every neighbourhood $\mathbb{B}(0, r)$ time and again. Since $B(t) - y$, $y \in \mathbb{R}^d$, is also a Brownian motion, the same argument shows that $(B(t))_{t \geq 0}$ visits any ball $\mathbb{B}(y, r)$ infinitely often.

In one dimension this is, because of the continuity of the paths of a Brownian motion, only possible if every *point* is visited infinitely often.

Let $d \geq 3$. By the strong Markov property, Theorem 6.6, and Corollary 6.16 we find for all $r < R$

$$\mathbb{P}^0(\exists t \geq \tau_{\mathbb{B}^c(0, R)} : B(t) \in \mathbb{B}(0, r)) = \mathbb{E}^0(\mathbb{P}^{B(\tau_{\mathbb{B}^c(0, R)})}(\tau_{\mathbb{B}(0, r)} < \infty)) = \left(\frac{r}{R}\right)^{d-2}.$$

Take $r = \sqrt{R}$ and let $R \rightarrow \infty$. This shows that the return probability of $(B(t))_{t \geq 0}$ to any compact neighbourhood of 0 is zero, i.e. $\lim_{t \rightarrow \infty} |B(t)| = \infty$. \square

6.5 Lévy's triple law

In this section we show how we can apply the reflection principle repeatedly to obtain for a BM^1 , $(B_t)_{t \geq 0}$, the joint distribution of B_t , the running minimum $m_t = \inf_{s \leq t} B_s$ and maximum $M_t = \sup_{s \leq t} B_s$. This is known as P. Lévy's *loi à trois variables*, [121, VI.42.6, 4°, p. 213]. In a different context this formula appears already in Bachelier [5, Nos. 413 and 504].

Ex. 6.8 6.18 Theorem (Lévy 1948). *Let $(B_t)_{t \geq 0}$ be a BM^1 and denote by m_t and M_t its running minimum and maximum respectively. Then*

$$\begin{aligned} & \mathbb{P}(m_t > a, M_t < b, B_t \in dx) \\ &= \frac{dx}{\sqrt{2\pi t}} \sum_{n=-\infty}^{\infty} \left[e^{-\frac{(x+2n(b-a))^2}{2t}} - e^{-\frac{(x-2a-2n(b-a))^2}{2t}} \right] \end{aligned} \quad (6.19)$$

for all $t > 0$ and $a < 0 < b$.

Proof. Let $a < 0 < b$, denote by $\tau = \inf\{t \geq 0 : B_t \notin (a, b)\}$ the first exit time from the interval (a, b) , and pick $I = [c, d] \subset (a, b)$. We know from Corollary 5.11 that τ is an a. s. finite stopping time.

Since $\{\tau > t\} = \{m_t > a, M_t < b\}$ we have

$$\begin{aligned} & \mathbb{P}(m_t > a, M_t < b, B_t \in I) \\ &= \mathbb{P}(\tau > t, B_t \in I) \\ &= \mathbb{P}(B_t \in I) - \mathbb{P}(\tau \leq t, B_t \in I) \\ &= \mathbb{P}(B_t \in I) - \mathbb{P}(B_\tau = a, \tau \leq t, B_t \in I) - \mathbb{P}(B_\tau = b, \tau \leq t, B_t \in I). \end{aligned}$$

For the last step we used the fact that $\{B_\tau = a\}$ and $\{B_\tau = b\}$ are, up to the null set $\{\tau = \infty\}$, a disjoint partition of Ω .

We want to use the reflection principle repeatedly. Denote by $r_a x := 2a - x$ and $r_b x := 2b - x$ the reflection in the (x, t) -plane with respect to the line $x = a$ and $x = b$, respectively, cf. Figure 6.4. Using Theorem 6.11 we find

$$\mathbb{P}(B_\tau = b, \tau \leq t, B_t \in I \mid \mathcal{F}_\tau)(\omega) = \mathbb{1}_{\{B_\tau = b\} \cap \{\tau \leq t\}}(\omega) \mathbb{P}^{B_\tau(\omega)}(B_{t-\tau(\omega)} \in I).$$

Because of the symmetry of Brownian motion,

$$\begin{aligned} \mathbb{P}^b(B_{t-\tau(\omega)} \in I) &= \mathbb{P}(B_{t-\tau(\omega)} \in I - b) \\ &= \mathbb{P}(B_{t-\tau(\omega)} \in b - I) \\ &= \mathbb{P}^b(B_{t-\tau(\omega)} \in r_b I), \end{aligned}$$

and, therefore,

$$\begin{aligned} \mathbb{P}(B_\tau = b, \tau \leq t, B_t \in I \mid \mathcal{F}_\tau)(\omega) &= \mathbb{1}_{\{B_\tau = b\} \cap \{\tau \leq t\}}(\omega) \mathbb{P}^{B_\tau(\omega)}(B_{t-\tau(\omega)} \in r_b I) \\ &= \mathbb{P}(B_\tau = b, \tau \leq t, B_t \in r_b I \mid \mathcal{F}_\tau)(\omega). \end{aligned}$$

This shows that

$$\mathbb{P}(B_\tau = b, \tau \leq t, B_t \in I) = \mathbb{P}(B_\tau = b, \tau \leq t, B_t \in r_b I) = \mathbb{P}(B_\tau = b, B_t \in r_b I).$$

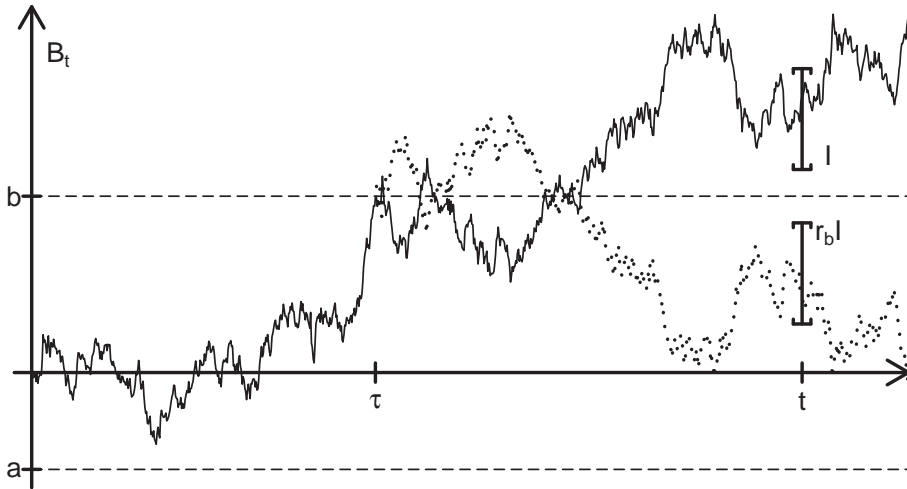


Figure 6.4. A [reflected] Brownian path visiting the [reflected] interval at time t .

Further applications of the reflection principle yield

$$\begin{aligned}
 \mathbb{P}(B_\tau = b, B_t \in r_b I) &= \mathbb{P}(B_t \in r_b I) - \mathbb{P}(B_\tau = a, B_t \in r_b I) \\
 &= \mathbb{P}(B_t \in r_b I) - \mathbb{P}(B_\tau = a, B_t \in r_a r_b I) \\
 &= \mathbb{P}(B_t \in r_b I) - \mathbb{P}(B_t \in r_a r_b I) + \mathbb{P}(B_\tau = b, B_t \in r_a r_b I)
 \end{aligned}$$

and, finally,

$$\begin{aligned}
 \mathbb{P}(B_\tau = b, \tau \leq t, B_t \in I) \\
 = \mathbb{P}(B_t \in r_b I) - \mathbb{P}(B_t \in r_a r_b I) + \mathbb{P}(B_t \in r_b r_a r_b I) - + \dots
 \end{aligned}$$

Since the quantities a and b play exchangeable roles in all calculations up to this point, we get

$$\begin{aligned}
 \mathbb{P}(m_t > a, M_t < b, B_t \in I) \\
 = \mathbb{P}(B_t \in I) - \mathbb{P}(B_t \in r_b I) + \mathbb{P}(B_t \in r_a r_b I) - \mathbb{P}(B_t \in r_b r_a r_b I) + \dots \\
 - \mathbb{P}(B_t \in r_a I) + \mathbb{P}(B_t \in r_b r_a I) - \mathbb{P}(B_t \in r_a r_b r_a I) + \dots \\
 = \mathbb{P}(B_t \in I) - \mathbb{P}(r_a B_t \in r_a r_b I) + \mathbb{P}(B_t \in r_a r_b I) - \mathbb{P}(r_a B_t \in (r_a r_b)^2 I) + \dots \\
 - \mathbb{P}(r_a B_t \in I) + \mathbb{P}(B_t \in r_b r_a I) - \mathbb{P}(r_a B_t \in r_b r_a I) + \dots
 \end{aligned}$$

In the last step we used that $r_a r_a = \text{id}$. All probabilities in these identities are probabilities of mutually disjoint events, i. e. the series converges absolutely (with a sum < 1); therefore we may rearrange the terms in the series. Since $r_a I = 2a - I$ and $r_b r_a I = 2b - (2a - I) = 2(b - a) + I$, we get for all $n \geq 0$

$$(r_b r_a)^n I = 2n(b - a) + I \quad \text{and} \quad (r_a r_b)^n I = 2n(a - b) + I = 2(-n)(b - a) + I.$$

Therefore,

$$\begin{aligned}
 \mathbb{P}(m_t > a, M_t < b, B_t \in I) \\
 = \sum_{n=-\infty}^{\infty} \mathbb{P}(B_t \in 2n(a - b) + I) - \sum_{n=-\infty}^{\infty} \mathbb{P}(2a - B_t \in 2n(a - b) + I) \\
 = \sum_{n=-\infty}^{\infty} \int_I \left(e^{-\frac{(x-2n(a-b))^2}{2t}} - e^{-\frac{(2a-x-2n(a-b))^2}{2t}} \right) \frac{dx}{\sqrt{2\pi t}}
 \end{aligned}$$

which is the same as (6.19). □

6.6 An arc-sine law

There are quite a few functionals of a Brownian path which follow an arc-sine distribution. Most of them are more or less connected with the zeros of $(B_t)_{t \geq 0}$. Consider, for example, the largest zero of B_s in the interval $[0, t]$:

$$\xi_t := \sup\{s \leq t : B_s = 0\}.$$

Obviously, ξ_t is not a stopping time since $\{\xi_t \leq r\}$, $r < t$, cannot be contained in \mathcal{F}_r^B ; otherwise we could forecast on the basis of $[0, r] \ni s \mapsto B_s$ whether $B_s = 0$ for some future time $s \in (r, t)$ or not. Nevertheless we can determine the probability distribution of ξ_t .

6.19 Theorem. *Let $(B_t)_{t \geq 0}$ be a BM¹ and write ξ_t for the largest zero of B_s in the interval $[0, t]$. Then*

Ex. 6.11

Ex. 6.12

Ex. 6.13

$$\mathbb{P}(\xi_t < s) = \frac{2}{\pi} \arcsin \sqrt{\frac{s}{t}} \quad \text{for all } 0 \leq s \leq t. \quad (6.20)$$

Proof. Set $h(s) := \mathbb{P}(\xi_t < s)$. By the Markov property we get

$$\begin{aligned} h(s) &= \mathbb{P}(B_u \neq 0 \quad \text{for all } u \in [s, t]) \\ &= \int \mathbb{P}^{B_s(\omega)}(B_{u-s} \neq 0 \quad \text{for all } u \in [s, t]) \mathbb{P}(d\omega) \\ &= \int \mathbb{P}^b(B_{u-s} \neq 0 \quad \text{for all } u \in [s, t]) \mathbb{P}(B_s \in db) \\ &= \int \mathbb{P}^0(B_v \neq -b \quad \text{for all } v \in [0, t-s]) \mathbb{P}(B_s \in db) \\ &= \int \mathbb{P}^0(\tau_{-b} \geq t-s) \mathbb{P}(B_s \in db), \end{aligned}$$

where τ_{-b} is the first passage time for $-b$. Using $\mathbb{P}(B_s \in db) = \mathbb{P}(B_s \in -db)$ and $\mathbb{P}(\tau_b \geq t-s) = \sqrt{\frac{2}{\pi(t-s)}} \int_0^b e^{-x^2/2(t-s)} dx$, cf. Theorem 6.9, we get

$$\begin{aligned} h(s) &= 2 \int_0^\infty \sqrt{\frac{2}{\pi(t-s)}} \int_0^b e^{-\frac{x^2}{2(t-s)}} dx \frac{1}{\sqrt{2\pi s}} e^{-\frac{b^2}{2s}} db \\ &= \frac{2}{\pi} \int_0^\infty \int_0^{\beta \sqrt{\frac{s}{t-s}}} e^{-\frac{\xi^2}{2}} d\xi e^{-\frac{\beta^2}{2}} d\beta. \end{aligned}$$

If we differentiate the last expression with respect to s we find

$$h'(s) = \frac{1}{\pi} \int_0^\infty e^{-\frac{\beta^2}{2}} e^{-\frac{\beta^2}{2} \frac{s}{t-s}} \frac{t}{t-s} \frac{\beta}{\sqrt{s(t-s)}} d\beta = \frac{1}{\pi} \frac{1}{\sqrt{s(t-s)}}, 0 < s < t.$$

Since $\frac{d}{ds} \arcsin \sqrt{\frac{s}{t}} = \frac{1}{2} \frac{1}{\sqrt{s(t-s)}} = \frac{\pi}{2} h'(s)$ and $0 = \arcsin 0 = h(0)$, we are done. \square

6.7 Some measurability issues

Let $(B_t)_{t \geq 0}$ be a BM^1 . Recall from Definition 5.1 that a filtration $(\mathcal{G}_t)_{t \geq 0}$ is admissible if

$$\mathcal{F}_t^B \subset \mathcal{G}_t \quad \text{and} \quad B_t - B_s \perp \mathcal{G}_s \quad \text{for all } s < t.$$

Typically, one considers the following types of enlargements:

$$\mathcal{G}_t := \sigma(\mathcal{F}_t^B, X) \quad \text{where } X \perp \mathcal{F}_\infty^B \text{ is a further random variable,}$$

$$\mathcal{G}_t := \mathcal{F}_{t+}^B := \bigcap_{u>t} \mathcal{F}_u^B,$$

$$\mathcal{G}_t := \overline{\mathcal{F}}_t^B := \sigma(\mathcal{F}_t^B, \mathcal{N})$$

where $\mathcal{N} = \{M \subset \Omega : \exists N \in \mathcal{A}, M \subset N, \mathbb{P}(N) = 0\}$ is the system of all subsets of measurable \mathbb{P} null sets; $\overline{\mathcal{F}}_t^B$ is the *completion* of \mathcal{F}_t^B .

Ex. 6.1 6.20 Lemma. *Let $(B_t)_{t \geq 0}$ be a BM^d . Then \mathcal{F}_{t+}^B and $\overline{\mathcal{F}}_t^B$ are admissible in the sense of Definition 5.1.*

Proof. **1°** \mathcal{F}_{t+}^B is admissible: Let $0 \leq t \leq u$, $F \in \mathcal{F}_{t+}$, and $f \in \mathcal{C}_b(\mathbb{R}^d)$. Since $F \in \mathcal{F}_{t+\epsilon}$ for all $\epsilon > 0$, we find with (5.1)

$$\mathbb{E} [\mathbb{1}_F \cdot f(B(u+\epsilon) - B(t+\epsilon))] = \mathbb{E} [\mathbb{1}_F] \cdot \mathbb{E} [f(B(u+\epsilon) - B(t+\epsilon))].$$

Letting $\epsilon \rightarrow 0$ proves 5.1 b); condition 5.1 a) is trivially satisfied.

2° $\overline{\mathcal{F}}_t^B$ is admissible: Let $0 \leq t \leq u$. By the very definition of the completion we can find for every $\tilde{F} \in \overline{\mathcal{F}}_t^B$ some $F \in \mathcal{F}_t^B$ and $N \in \mathcal{N}$ such that the symmetric difference $(\tilde{F} \setminus F) \cup (F \setminus \tilde{F})$ is in \mathcal{N} . Therefore,

$$\begin{aligned} \mathbb{E}(\mathbb{1}_{\tilde{F}} \cdot (B_u - B_t)) &= \mathbb{E}(\mathbb{1}_F \cdot (B_u - B_t)) \stackrel{(5.1)}{=} \mathbb{P}(F) \mathbb{E}(B_u - B_t) \\ &= \mathbb{P}(\tilde{F}) \mathbb{E}(B_u - B_t), \end{aligned}$$

which proves 5.1 b); condition 5.1 a) is clear. \square

6.21 Theorem. Let $(B_t)_{t \geq 0}$ be a BM^d . Then $(\overline{\mathcal{F}}_t^B)_{t \geq 0}$ is a right-continuous filtration, i. e. $\overline{\mathcal{F}}_{t+}^B = \overline{\mathcal{F}}_t^B$ holds for all $t \geq 0$.

Proof. Let $F \in \overline{\mathcal{F}}_{t+}^B$. It is enough to show that $\mathbb{1}_F = \mathbb{E}[\mathbb{1}_F | \overline{\mathcal{F}}_t^B]$ almost surely. If this is true, $\mathbb{1}_F$ is, up to a set of measure zero, $\overline{\mathcal{F}}_t^B$ measurable. Since $\mathcal{N} \subset \overline{\mathcal{F}}_t^B \subset \overline{\mathcal{F}}_{t+}^B$, we get $\overline{\mathcal{F}}_{t+}^B \subset \overline{\mathcal{F}}_t^B$.

Fix $t < u$ and pick $0 \leq s_1 < \dots < s_m \leq t < u = t_0 \leq t_1 < \dots < t_n$ and $\eta_1, \dots, \eta_m, \xi_1, \dots, \xi_n \in \mathbb{R}^d, m, n \geq 1$. Then we have

$$\mathbb{E}[e^{i \sum_{j=1}^m \langle \eta_j, B(s_j) \rangle} | \overline{\mathcal{F}}_t^B] = e^{i \sum_{j=1}^m \langle \eta_j, B(s_j) \rangle} = \mathbb{E}[e^{i \sum_{j=1}^m \langle \eta_j, B(s_j) \rangle} | \overline{\mathcal{F}}_{t+}^B]. \quad (6.21)$$

As in the proof of Theorem 6.1 we get

$$\sum_{j=1}^n \langle \xi_j, B(t_j) \rangle = \sum_{k=1}^n \langle \xi_k, B(t_k) - B(t_{k-1}) \rangle + \langle \xi_1, B(u) \rangle, \quad \xi_k = \xi_k + \dots + \xi_n.$$

Therefore,

$$\begin{aligned} \mathbb{E}[e^{i \sum_{j=1}^n \langle \xi_j, B(t_j) \rangle} | \overline{\mathcal{F}}_u^B] &= \mathbb{E}[e^{i \sum_{k=1}^n \langle \xi_k, B(t_k) - B(t_{k-1}) \rangle} | \overline{\mathcal{F}}_u^B] e^{i \langle \xi_1, B(u) \rangle} \\ &\stackrel{(6.20)}{=} \prod_{k=1}^n \mathbb{E}[e^{i \langle \xi_k, B(t_k) - B(t_{k-1}) \rangle} | \overline{\mathcal{F}}_u^B] e^{i \langle \xi_1, B(u) \rangle} \\ &\stackrel{(2.17)}{=} \prod_{k=1}^n e^{-\frac{1}{2}(t_k - t_{k-1})|\xi_k|^2} e^{i \langle \xi_1, B(u) \rangle} \end{aligned}$$

Exactly the same calculation with u replaced by t and $t_0 = t$ tells us

$$\mathbb{E}[e^{i \sum_{j=1}^n \langle \xi_j, B(t_j) \rangle} | \overline{\mathcal{F}}_t^B] = \prod_{k=1}^n e^{-\frac{1}{2}(t_k - t_{k-1})|\xi_k|^2} e^{i \langle \xi_1, B(t) \rangle}.$$

Now we can use the backwards martingale convergence Theorem A.8 and let $u \downarrow t$ (along a countable sequence) in the first calculation. This gives

$$\begin{aligned} \mathbb{E}[e^{i \sum_{j=1}^n \langle \xi_j, B(t_j) \rangle} | \overline{\mathcal{F}}_{t+}^B] &= \lim_{u \downarrow t} \mathbb{E}[e^{i \sum_{j=1}^n \langle \xi_j, B(t_j) \rangle} | \overline{\mathcal{F}}_u^B] \\ &= \lim_{u \downarrow t} \prod_{k=1}^n e^{-\frac{1}{2}(t_k - t_{k-1})|\xi_k|^2} e^{i \langle \xi_1, B(u) \rangle} \\ &= \prod_{k=1}^n e^{-\frac{1}{2}(t_k - t_{k-1})|\xi_k|^2} e^{i \langle \xi_1, B(t) \rangle} \\ &= \mathbb{E}[e^{i \sum_{j=1}^n \langle \xi_j, B(t_j) \rangle} | \overline{\mathcal{F}}_t^B]. \end{aligned} \quad (6.22)$$

Given $0 \leq u_1 < u_2 < \dots < u_N$, $N \geq 1$, and $t \geq 0$, then t splits the u_j into two groups, $s_1 < \dots < s_n \leq t < t_1 < \dots < t_m$, where $n + m = N$ and $\{u_1, \dots, u_N\} = \{s_1, \dots, s_n\} \cup \{t_1, \dots, t_m\}$. If we combine the equalities (6.21) and (6.22), we see for $\Gamma := (B(u_1), \dots, B(u_N)) \in \mathbb{R}^{Nd}$ and all $\theta \in \mathbb{R}^{Nd}$

$$\mathbb{E}[e^{i\langle \theta, \Gamma \rangle} \mid \overline{\mathcal{F}}_{t+}^B] = \mathbb{E}[e^{i\langle \theta, \Gamma \rangle} \mid \overline{\mathcal{F}}_t^B].$$

Using the definition and then the L^2 symmetry of the conditional expectation we get for all $F \in \overline{\mathcal{F}}_{t+}^B$

$$\int \mathbb{1}_F \cdot e^{i\langle \theta, \Gamma \rangle} d\mathbb{P} = \int \mathbb{1}_F \cdot \mathbb{E}[e^{i\langle \theta, \Gamma \rangle} \mid \overline{\mathcal{F}}_t^B] d\mathbb{P} = \int \mathbb{E}[\mathbb{1}_F \mid \overline{\mathcal{F}}_t^B] \cdot e^{i\langle \theta, \Gamma \rangle} d\mathbb{P}.$$

This shows that the image measures $(\mathbb{1}_F \mathbb{P}) \circ \Gamma^{-1}$ and $(\mathbb{E}[\mathbb{1}_F \mid \overline{\mathcal{F}}_t^B] \mathbb{P}) \circ \Gamma^{-1}$ have the same characteristic functions, i. e.

$$\int_A \mathbb{1}_F d\mathbb{P} = \int_A \mathbb{E}[\mathbb{1}_F \mid \overline{\mathcal{F}}_t^B] d\mathbb{P} \quad (6.23)$$

holds for all $A \in \sigma(\Gamma) = \sigma(B(u_1), \dots, B(u_N))$ and all $0 \leq u_1 < \dots < u_N$ and $N \geq 1$. Since $\bigcup_{u_1 < \dots < u_N, N \geq 1} \sigma(B(u_1), \dots, B(u_N))$ is a \cap -stable generator of \mathcal{F}_∞^B , the uniqueness theorem for measures shows that (6.23) holds for all $A \in \mathcal{F}_\infty^B$. Therefore, $\mathbb{1}_F = \mathbb{E}[\mathbb{1}_F \mid \overline{\mathcal{F}}_t^B]$ almost surely. \square

6.22 Corollary (Blumenthal's 0-1-law. Blumenthal 1957). *Let $(B(t))_{t \geq 0}$ be a BM^d . Then \mathcal{F}_{0+}^B is trivial, i. e. if $F \in \mathcal{F}_{0+}^B$ then $\mathbb{P}(F) = 0$ or 1 .*

Proof. By Theorem 6.21 we know that $\overline{\mathcal{F}}_0^B = \overline{\mathcal{F}}_{0+}^B \supset \mathcal{F}_{0+}^B$. On the other hand we have $\overline{\mathcal{F}}_0^B = \sigma(\mathcal{N}, \mathcal{F}_0^B) = \sigma(\mathcal{N})$ since $\mathcal{F}_0^B = \{\emptyset, \Omega\}$. This proves our claim.

Let us indicate the *classical proof* which does not rely on Theorem 6.21. Since \mathcal{F}_{t+}^B is admissible, cf. Lemma 6.20, we find that \mathcal{F}_{0+}^B is independent of $B_t - B_0 = B_t$ for all $t \geq 0$. This means that

$$\mathcal{F}_{0+}^B \perp\!\!\!\perp \bigcap_{t>0} \underbrace{\sigma\left(\bigcup_{0 \leq t_1 < \dots < t_n \leq t, n \geq 1} \sigma(B_{t_1}, \dots, B_{t_n})\right)}_{\cap\text{-stable generator}} = \mathcal{F}_{0+}^B,$$

i. e. $\mathbb{P}(F) = \mathbb{P}(F \cap F) = \mathbb{P}(F) \cdot \mathbb{P}(F)$ so that $\mathbb{P}(F) = 0$ or 1 . \square

6.23 Remark (Usual conditions). Many authors assume that the underlying filtered probability space $(\Omega, \mathcal{A}, \mathbb{P}, \mathcal{F}_t, t \geq 0)$ satisfies the *usual conditions* or *usual hypotheses*, i. e.

a) \mathcal{F}_0 contains all subsets of \mathbb{P} null sets of \mathcal{A} ;

b) $(\mathcal{F}_t)_{t \geq 0}$ is right-continuous, i. e. $\mathcal{F}_t = \mathcal{F}_{t+}$ holds for all $t \geq 0$.

Theorem 6.21 ensures that a Brownian motion together with the completion of its canonical filtration satisfies the usual conditions.

We close this section with a result that exhibits the interaction of a Brownian motion with the underlying measurability structure.

6.24 Lemma (Progressive measurability). *Let $(B_t)_{t \geq 0}$ be a BM^d with an admissible filtration $(\mathcal{F}_t)_{t \geq 0}$. Then $(B_t)_{t \geq 0}$ is progressively measurable i. e.*

$$B(\cdot, \cdot) : ([0, t] \times \Omega, \mathcal{B}[0, t] \otimes \mathcal{F}_t) \rightarrow (\mathbb{R}^d, \mathcal{B}(\mathbb{R}^d))$$

is, for every $t \geq 0$, $\mathcal{B}[0, t] \otimes \mathcal{F}_t / \mathcal{B}(\mathbb{R}^d)$ measurable.

Proof. Set

$$B^j(s, \omega) := \begin{cases} B(\frac{k}{2^j}t, \omega), & k2^{-j}t < s \leq (k+1)2^{-j}t, \quad k = 0, 1, \dots, 2^j - 1, \\ B(0, \omega), & s = 0. \end{cases}$$

Then we find for all $A \in \mathcal{B}(\mathbb{R}^d)$

$$\begin{aligned} & \{(s, \omega) \in [0, t] \times \Omega : B^j(s, \omega) \in A\} \\ &= \{0\} \times \{B(0, \cdot) \in A\} \cup \bigcup_{k=0}^{2^j-1} \left(\frac{k}{2^j}t, \frac{k+1}{2^j}t \right] \times \{B(\frac{k}{2^j}t, \cdot) \in A\}. \end{aligned}$$

This set is contained in $\mathcal{B}[0, t] \otimes \mathcal{F}_t$. Therefore, the processes $(B^j(t))_{t \geq 0}$, $j \geq 0$, are progressively measurable. Because of the (left-)continuity of the sample paths, we get

$$B(s, \omega) = \lim_{j \rightarrow \infty} B^j(s, \omega) \quad \text{for all } s \in [0, t], \omega \in \Omega,$$

and this shows that $(B(t))_{t \geq 0}$ is also progressively measurable. \square

6.25 Corollary. *Let $(B_t)_{t \geq 0}$ be a BM^1 and τ be an \mathcal{F}_t stopping time. Then $B_\tau \mathbb{1}_{\{\tau < \infty\}}$ is \mathcal{F}_τ measurable.* Ex. 6.4

Proof. Consider the maps

$$(\{\tau \leq t\}, \mathcal{F}_t) \xrightarrow{\omega \mapsto (\tau(\omega), \omega)} ([0, t] \times \Omega, \mathcal{B}[0, t] \otimes \mathcal{F}_t) \xrightarrow{(s, \omega) \mapsto B(s, \omega)} (\mathbb{R}^d, \mathcal{B}(\mathbb{R}^d)).$$

The second map is measurable, since $(B_t)_{t \geq 0}$ is progressively measurable. For the first map we observe that for all $r \leq t$ and $\Gamma \in \mathcal{F}_t$

$$\underbrace{\{\tau \leq r\}}_{\in \mathcal{F}_r \subset \mathcal{F}_t} \cap \Gamma \in \mathcal{F}_t.$$

Thus, $\{B(\tau) \in A\} \cap \{\tau \leq t\} = \{\omega \in \{\tau \leq t\} : B(\tau(\omega), \omega) \in A\} \in \mathcal{F}_t$, which proves that $\{B(\tau) \in A\} \cap \{\tau < \infty\} \in \mathcal{F}_\tau$, i. e. $B(\tau)\mathbb{1}_{\{\tau < \infty\}}$ is \mathcal{F}_τ measurable. \square



6.26 Further reading. The (strong) Markov property is a central theme in most books on Brownian motion. The account in [23] excels in its clarity, elegance and precision. Shift operators and the Markov property on the path space are explored in [99]. For the general theory of Markov processes we refer to [61] and the classic book [13].

[13] Blumenthal, Gettoor: *Markov Processes and Potential Theory*.

[23] Chung: *Lectures from Markov Processes to Brownian Motion*

[61] Ethier, Kurtz: *Markov Processes: Characterization and Convergence*.

[99] Karatzas, Shreve: *Brownian Motion and Stochastic Calculus*.

Problems

1. Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM¹.
 - (a) Show that $X_t := |B_t|$, $t \geq 0$, is also a Markov process for the filtration $(\mathcal{F}_t)_{t \geq 0}$, i. e. for all $s, t \geq 0$ and $u \in \mathcal{B}_b(\mathbb{R})$

$$\mathbb{E}(u(X_{t+s}) \mid \mathcal{F}_s) = \mathbb{E}(u(X_{t+s}) \mid X_s).$$

- (b) Find in a) the function $g_{u,s,t+s}(x)$ with $\mathbb{E}(u(X_{t+s}) \mid X_s) = g_{u,s,t+s}(X_s)$.
 - (c) Solve a) and b) for the process $Y_t := \sup_{s \leq t} B_s - B_t$, $t \geq 0$.
 - (d) Compare the transition functions from b) and c) for X and Y , respectively.
2. Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM¹ and set $M_t := \sup_{s \leq t} B_s$ and $I_t = \int_0^t B_s ds$.
 - (a) Show that the two-dimensional process $(B_t, M_t)_{t \geq 0}$ is a Markov process for $(\mathcal{F}_t)_{t \geq 0}$.
 - (b) Show that the two-dimensional process $(B_t, I_t)_{t \geq 0}$ is a Markov process for $(\mathcal{F}_t)_{t \geq 0}$.
 - (c) Are $(M_t)_{t \geq 0}$ and $(I_t)_{t \geq 0}$ Markov processes for $(\mathcal{F}_t)_{t \geq 0}$?

3. Let $(B(t))_{t \geq 0}$ be a BM^d and let Z be a bounded \mathcal{F}_∞^B measurable random variable. Then $x \mapsto \mathbb{E}^x Z$ is in $\mathcal{B}_b(\mathbb{R}^d)$.
Hint: \mathcal{F}_∞^B is generated by sets of the form $G := \bigcap_{j=1}^n \{B(t_j) \in A_j\}$. Then, $\mathbb{E}^x \mathbb{1}_G \in \mathcal{B}_b(\mathbb{R}^d)$. Consider the family $\Sigma := \{A : \mathbb{E}^x \mathbb{1}_A \in \mathcal{B}_b(\mathbb{R}^d)\}$; this is a σ -algebra containing all sets G , hence \mathcal{F}_∞^B .
4. Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM^1 and τ a stopping time for the filtration $(\mathcal{F}_t)_{t \geq 0}$. Show that B_τ is measurable with respect to $\mathcal{F}_{\tau+}$.
Hint: Use in Theorem 6.6 $u(x) = u_n(x) \uparrow x$, $u \in \mathcal{C}_b(\mathbb{R}^d)$; see also Corollary 6.25.
5. Let $(B_t)_{t \geq 0}$ be a BM^1 . Show that $\mathbb{P}(\sup_{s \leq t} |B_s| \geq x) \leq 2 \mathbb{P}(|B_t| \geq x)$, $x, t \geq 0$.
6. Let $(B_t)_{t \geq 0}$ be a BM^1 and denote by $\tau_b = \inf\{s \geq 0 : B_s = b\}$ the first time when B_t reaches $b \in \mathbb{R}$. Show that
 - (a) $\tau_b \sim \tau_{-b}$;
 - (b) $\tau_{cb} \sim c^2 \tau_b$, $c \in \mathbb{R}$;
 - (c) if $0 < a < b$, then $\tau_b - \tau_a \perp\!\!\!\perp \{\tau_\alpha : \alpha \in [0, a]\}$.
7. Let $\tau = \tau_{(a,b)^c}^\circ$ be the first exit time of a Brownian motion from the interval (a, b) .
 - (a) Find $\mathbb{E}^x e^{-\lambda \tau}$ for all $x \in (a, b)$ and $\lambda > 0$.
 - (b) Find $\mathbb{E}^x (e^{-\lambda \tau} \mathbb{1}_{\{B_\tau = a\}})$ for all $x \in (a, b)$ and $\lambda > 0$.
8. Let $(B_t)_{t \geq 0}$ be a BM^1 and set $M_t := \sup_{s \leq t} B_s$. Find the distribution of (M_t, B_t) .
9. Let $(B_t)_{t \geq 0}$ be a BM^1 and let τ_0 be the first hitting time of 0. Find the ‘density’ of $\mathbb{P}^x(B_t \in dz, \tau_0 > t)$, i. e. find the function $f_{t,x}(z)$ such that

$$\mathbb{P}^x(B_t \in A, \tau_0 > t) = \int_A f_{t,x}(z) dz \quad \text{for all } A \in \mathcal{B}(\mathbb{R}).$$

10. Let $K \subset \mathbb{R}^d$ be a compact set. Show that there is a decreasing sequence of continuous functions $\phi_n(x)$ such that $\mathbb{1}_K = \inf_n \phi_n$.
Hint: Let $U \supset K$ be an open set and $\phi(x) := d(x, U^c)/(d(x, K) + d(x, U^c))$.
11. Let $(B_t)_{t \geq 0}$ be a BM^1 . Find the distribution of $\widetilde{\xi}_t := \inf\{s \geq t : B_s = 0\}$.
12. Let $(B_t)_{t \geq 0}$ be a BM^1 and $0 < u < v < w < \infty$. Find the following probabilities:
 - (a) $\mathbb{P}(B_t = 0 \text{ for some } t \in (u, v))$;
 - (b) $\mathbb{P}(B_t \neq 0 \forall t \in (u, w) \mid B_t \neq 0 \forall t \in (u, v))$;
 - (c) $\mathbb{P}(B_t \neq 0 \forall t \in (0, w) \mid B_t \neq 0 \forall t \in (0, v))$.
13. Let $(B_t)_{t \geq 0}$ be a BM^1 and set $M_t = \sup_{s \leq t} B_s$. Denote by ξ_t the largest zero of B_s before time t and by η_t the largest zero of $Y_s = M_s - B_s$ before time t . Show that $\xi_t \sim \eta_t$.

Chapter 7

Brownian motion and transition semigroups

The Markov property allows us to study Brownian motion through an analytic approach. For this we associate with a Brownian motion $(B_t)_{t \geq 0}$ two families of linear operators, the *(transition) semigroup*

$$P_t u(x) = \mathbb{E}^x u(B_t), \quad u \in \mathcal{B}_b(\mathbb{R}^d), \quad x \in \mathbb{R}^d, \quad t \geq 0, \quad (7.1)$$

and the *resolvent*

$$U_\alpha u(x) = \mathbb{E}^x \left[\int_0^\infty u(B_t) e^{-\alpha t} dt \right], \quad u \in \mathcal{B}_b(\mathbb{R}^d), \quad x \in \mathbb{R}^d, \quad \alpha > 0. \quad (7.2)$$

Throughout this chapter we assume that $(B_t)_{t \geq 0}$ is a d -dimensional Brownian motion which is adapted to the filtration $(\mathcal{F}_t)_{t \geq 0}$.¹ We will assume that $(\mathcal{F}_t)_{t \geq 0}$ is right-continuous, i. e. $\mathcal{F}_t = \mathcal{F}_{t+}$ for all $t \geq 0$, cf. Theorem 6.21 for a sufficient condition.

7.1 The semigroup

A semigroup $(P_t)_{t \geq 0}$ on a Banach space $(\mathfrak{B}, \|\cdot\|)$ is a family of linear operators $P_t : \mathfrak{B} \rightarrow \mathfrak{B}$, $t \geq 0$, which satisfies

$$P_t P_s = P_{t+s} \quad \text{and} \quad P_0 = \text{id}. \quad (7.3)$$

In this chapter we will consider the following Banach spaces:

$\mathcal{B}_b(\mathbb{R}^d)$ the family of all bounded Borel measurable functions $f : \mathbb{R}^d \rightarrow \mathbb{R}$ equipped with the uniform norm $\|\cdot\|_\infty$;

Ex. 7.1 $\mathcal{C}_\infty(\mathbb{R}^d)$ the family of all continuous functions $f : \mathbb{R}^d \rightarrow \mathbb{R}$ vanishing at infinity, i. e. $\lim_{|x| \rightarrow \infty} u(x) = 0$, equipped with the uniform norm $\|\cdot\|_\infty$.

Unless it is ambiguous, we write \mathcal{B}_b and \mathcal{C}_∞ instead of $\mathcal{B}_b(\mathbb{R}^d)$ and $\mathcal{C}_\infty(\mathbb{R}^d)$.

¹ Many results of this chapter will be true, with only few changes in the proofs, for more general Markov processes.

7.1 Lemma. Let $(B_t)_{t \geq 0}$ be a d -dimensional Brownian motion with filtration $(\mathcal{F}_t)_{t \geq 0}$. Then (7.1) defines a semigroup of operators on $\mathcal{B}_b(\mathbb{R}^d)$.

Proof. It is obvious that P_t is a linear operator. Since $x \mapsto e^{-|x-y|^2/2t}$ is measurable, we see that

$$P_t u(x) = \int_{\mathbb{R}^d} u(y) \mathbb{P}^x(B_t \in dy) = \frac{1}{(2\pi t)^{d/2}} \int_{\mathbb{R}^d} u(y) e^{-|x-y|^2/2t} dy$$

is again a Borel measurable function. (For a general Markov process we must use (6.4c) with a Borel measurable $g_{u,t}(\cdot)$.) The semigroup property (7.3) is now a direct consequence of the Markov property (6.4): For $s, t \geq 0$ and $u \in \mathcal{B}_b(\mathbb{R}^d)$



$$\begin{aligned} P_{t+s} u(x) &= \mathbb{E}^x u(B_{t+s}) \stackrel{\text{tower}}{=} \mathbb{E}^x [\mathbb{E}^x (u(B_{t+s}) | \mathcal{F}_s)] \\ &\stackrel{(6.4)}{=} \mathbb{E}^x [\mathbb{E}^{B_s} (u(B_t))] \\ &= \mathbb{E}^x [P_t u(B_s)] \\ &= P_s P_t u(x) \end{aligned} \quad \square$$

For many properties of the semigroup the Banach space $\mathcal{B}_b(\mathbb{R}^d)$ is too big, and we have to consider the smaller space $\mathcal{C}_\infty(\mathbb{R}^d)$.

7.2 Lemma. A d -dimensional Brownian motion $(B_t)_{t \geq 0}$ is uniformly stochastically continuous, i. e.

$$\lim_{t \rightarrow 0} \sup_{x \in \mathbb{R}^d} \mathbb{P}^x(|B_t - x| > \delta) = 0 \quad \text{for all } \delta > 0. \quad (7.4)$$

Proof. This follows immediately from the translation invariance of a Brownian motion and Chebyshev's inequality

$$\mathbb{P}^x(|B_t - x| > \delta) = \mathbb{P}(|B_t| > \delta) \leq \frac{\mathbb{E}|B_t|^2}{\delta^2} = \frac{td}{\delta^2}. \quad \square$$

After this preparation we can discuss the properties of the semigroup $(P_t)_{t \geq 0}$.

7.3 Proposition. Let $(P_t)_{t \geq 0}$ be the transition semigroup of a d -dimensional Brownian motion $(B_t)_{t \geq 0}$ and $\mathcal{C}_\infty = \mathcal{C}_\infty(\mathbb{R}^d)$ and $\mathcal{B}_b = \mathcal{B}_b(\mathbb{R}^d)$. For all $t \geq 0$

Ex. 7.2

Ex. 7.3

Ex. 7.4

Ex. 7.5

- a) P_t is conservative: $P_t 1 = 1$.
- b) P_t is a contraction on \mathcal{B}_b : $\|P_t u\|_\infty \leq \|u\|_\infty$ for all $u \in \mathcal{B}_b$.
- c) P_t is positivity preserving: For $u \in \mathcal{B}_b$ we have $u \geq 0 \Rightarrow P_t u \geq 0$;
- d) P_t is sub-Markovian: For $u \in \mathcal{B}_b$ we have $0 \leq u \leq 1 \Rightarrow 0 \leq P_t u \leq 1$;
- e) P_t has the Feller property: If $u \in \mathcal{C}_\infty$ then $P_t u \in \mathcal{C}_\infty$;
- f) P_t is strongly continuous on \mathcal{C}_∞ : $\lim_{t \rightarrow 0} \|P_t u - u\|_\infty = 0$ for all $u \in \mathcal{C}_\infty$;

g) P_t has the strong Feller property: If $u \in \mathcal{B}_b$, then $P_t u \in \mathcal{C}_b$.

7.4 Remark. A semigroup of operators $(P_t)_{t \geq 0}$ on $\mathcal{B}_b(\mathbb{R}^d)$ which satisfies

- a)–d) is called a *Markov semigroup*;
- b)–d) is called a *sub-Markov semigroup*;
- b)–f) is called a *Feller semigroup*;
- b)–d), g) is called a *strong Feller semigroup*.

Proof of Proposition 7.3. a) We have $P_t 1(x) = \mathbb{E}^x \mathbb{1}_{\mathbb{R}^d}(B_t) = \mathbb{P}^x(X_t \in \mathbb{R}^d) = 1$ for all $x \in \mathbb{R}^d$.

b) We have $P_t u(x) = \mathbb{E}^x u(B_t) = \int u(B_t) d\mathbb{P}^x \leq \|u\|_\infty$ for all $x \in \mathbb{R}^d$.

c) Since $u \geq 0$ we see that $P_t u(x) = \mathbb{E}^x u(B_t) \geq 0$.

d) Positivity follows from c). The proof of b) shows that $P_t u \leq 1$ if $u \leq 1$.

e) Note that $P_t u(x) = \mathbb{E}^x u(B_t) = \mathbb{E} u(B_t + x)$. Since $|u(B_t + x)| \leq \|u\|_\infty$ is integrable, the dominated convergence theorem, can be used to show

$$\lim_{x \rightarrow y} u(B_t + x) = u(B_t + y) \quad \text{and} \quad \lim_{|x| \rightarrow \infty} u(B_t + x) = 0.$$

f) Fix $\epsilon > 0$. Since $u \in \mathcal{C}_\infty(\mathbb{R}^d)$ is uniformly continuous, there is some $\delta > 0$ such that

$$|u(x) - u(y)| < \epsilon \quad \text{for all } |x - y| < \delta.$$

Thus,

$$\begin{aligned} \|P_t u - u\|_\infty &= \sup_{x \in \mathbb{R}^d} |\mathbb{E}^x u(B_t) - u(x)| \\ &\leq \sup_{x \in \mathbb{R}^d} \mathbb{E}^x |u(B_t) - u(x)| \\ &\leq \sup_{x \in \mathbb{R}^d} \left(\int_{|B_t - x| < \delta} |u(B_t) - u(x)| d\mathbb{P}^x + \int_{|B_t - x| \geq \delta} |u(B_t) - u(x)| d\mathbb{P}^x \right) \\ &\leq \epsilon \sup_{x \in \mathbb{R}^d} \mathbb{P}^x(|B_t - x| < \delta) + 2\|u\|_\infty \sup_{x \in \mathbb{R}^d} \mathbb{P}^x(|B_t - x| \geq \delta) \\ &\leq \epsilon + 2\|u\|_\infty \mathbb{P}(|B_t| \geq \delta). \end{aligned}$$

Since $(B_t)_{t \geq 0}$ is stochastically continuous, we get $\overline{\lim}_{t \rightarrow 0} \|P_t u - u\|_\infty \leq \epsilon$, cf. (7.4), and the claim follows as $\epsilon \rightarrow 0$.

g) Since continuity is a local property, it is enough to consider $z \rightarrow x$ for all points $z, x \in \mathbb{B}(0, R)$ and any $R > 0$. Let $u \in \mathcal{B}_b(\mathbb{R}^d)$. Then $P_t u$ is given by

$$P_t u(z) = p_t \star u(z) = \frac{1}{(2\pi t)^{d/2}} \int u(y) e^{-|z-y|^2/2t} dy,$$

and so

$$P_t u(z) = \frac{1}{(2\pi t)^{d/2}} \left(\int_{|y| < 2R} u(y) e^{-|z-y|^2/2t} dy + \int_{|y| \geq 2R} u(y) e^{-|z-y|^2/2t} dy \right).$$

If $|y| \geq 2$ we get $|z-y|^2 \geq (|y| - |z|)^2 \geq \frac{1}{4}|y|^2$ which shows that the integrand is bounded by

$$|u(y)| e^{-|z-y|^2/2t} \leq \|u\|_\infty \left(\mathbb{1}_{\mathbb{B}(0,2R)}(y) + e^{-|y|^2/8t} \mathbb{1}_{\mathbb{B}^c(0,2R)}(y) \right)$$

and the right-hand side is integrable. Therefore we can use dominated convergence to get $\lim_{z \rightarrow x} P_t u(z) = P_t u(x)$. This proves that $P_t u \in \mathcal{C}_b$. \square

The semigroup notation offers a simple way to express the finite dimensional distributions of a Markov process. Let $(B_t)_{t \geq 0}$ be a d -dimensional Brownian motion, $s < t$ and $f, g \in \mathcal{B}_b(\mathbb{R}^d)$. Then

$$\begin{aligned} \mathbb{E}^x [f(B_s)g(B_t)] &= \mathbb{E}^x [f(B_s) \mathbb{E}^{B_s} g(B_{t-s})] \\ &= \mathbb{E}^x [f(B_s) P_{t-s} g(B_s)] = P_s [f P_{t-s} g](x). \end{aligned}$$

If we iterate this, we get the following counterpart of Theorem 6.4.

7.5 Theorem. *Let $(B_t)_{t \geq 0}$ be a BM^d , $x \in \mathbb{R}^d$, $t_0 = 0 < t_1 < \dots < t_n$ and $C_1, \dots, C_n \in \mathcal{B}(\mathbb{R}^d)$. If $(P_t)_{t \geq 0}$ is the transition semigroup of $(B_t)_{t \geq 0}$, then*

$$\mathbb{P}^x(B_{t_1} \in C_1, \dots, B_{t_n} \in C_n) = P_{t_1} [\mathbb{1}_{C_1} P_{t_2-t_1} [\mathbb{1}_{C_2} \dots P_{t_n-t_{n-1}} \mathbb{1}_{C_n}] \dots](x).$$

7.6 Remark (From semigroups to processes). So far we have considered semigroups given by a Brownian motion. It is an interesting question whether we can construct a Markov process starting from a semigroup of operators. If $(T_t)_{t \geq 0}$ is a Markov semigroup where $p_t(x, C) := T_t \mathbb{1}_C(x)$ is an *everywhere defined* kernel such that

$$\begin{aligned} C &\mapsto p_t(x, C) \quad \text{is a probability measure for all } x \in \mathbb{R}^d, \\ (t, x) &\mapsto p_t(x, C) \quad \text{is measurable for all } C \in \mathcal{B}(\mathbb{R}^d), \end{aligned}$$

this is indeed possible. We can use the formula from Theorem 7.5 to define for every $x \in \mathbb{R}^d$

$$p_{t_1, \dots, t_n}^x(C_1 \times \dots \times C_n) := T_{t_1} [\mathbb{1}_{C_1} T_{t_2-t_1} [\mathbb{1}_{C_2} \dots T_{t_n-t_{n-1}} \mathbb{1}_{C_n}] \dots](x)$$

where $0 \leq t_1 < t_2 < \dots < t_n$ and $C_1, \dots, C_n \in \mathcal{B}(\mathbb{R}^d)$. For fixed $x \in \mathbb{R}^d$, this

Ex. 7.6



Ex. 7.7 defines a family of finite dimensional measures on $(\mathbb{R}^{d \cdot n}, \mathcal{B}^n(\mathbb{R}^d))$ which satisfy the Kolmogorov consistency conditions (4.6). The permutation property is obvious, while the projectivity condition follows from the semigroup property:

$$\begin{aligned}
 & p_{t_1, \dots, t_k, \dots, t_n}^x (C_1 \times \dots \times \mathbb{R}^d \times \dots \times C_n) \\
 &= T_{t_1} \left[\mathbb{1}_{C_1} T_{t_2-t_1} [\mathbb{1}_{C_2} \dots \mathbb{1}_{C_{k-1}} \underbrace{T_{t_k-t_{k-1}} \mathbb{1}_{\mathbb{R}^d} T_{t_{k+1}-t_k}}_{= T_{t_k-t_{k-1}} T_{t_{k+1}-t_k} = T_{t_{k+1}-t_{k-1}}} \mathbb{1}_{C_{k+1}} \dots T_{t_n-t_{n-1}} \mathbb{1}_{C_n}] \dots \right] (x) \\
 &= T_{t_1} \left[\mathbb{1}_{C_1} T_{t_2-t_1} [\mathbb{1}_{C_2} \dots \mathbb{1}_{C_{k-1}} T_{t_{k+1}-t_{k-1}} \mathbb{1}_{C_{k+1}} \dots T_{t_n-t_{n-1}} \mathbb{1}_{C_n}] \dots \right] (x) \\
 &= p_{t_1, \dots, t_{k-1}, t_{k+1}, \dots, t_n}^x (C_1 \times \dots \times C_{k-1} \times C_{k+1} \times \dots \times C_n).
 \end{aligned}$$

Now we can use Kolmogorov's construction, cf. Theorem 4.8 and Corollary 4.9, to get measures \mathbb{P}^x and a canonical process $(X_t)_{t \geq 0}$ such that $\mathbb{P}^x(X_0 = x) = 1$ for every $x \in \mathbb{R}^d$.

We still have to check that the process $X = (X_t)_{t \geq 0}$ is a Markov process. Let $\mathcal{F}_t := \sigma(X_s : s \leq t)$ denote the canonical filtration of X . We will show that

$$p_{t-s}(X_s, C) = \mathbb{P}^x(X_t \in C \mid \mathcal{F}_s) \quad \text{for all } s < t \text{ and } C \in \mathcal{B}(\mathbb{R}^d). \quad (7.5)$$

A further application of the tower property gives (6.4a). Since \mathcal{F}_s is generated by an \cap -stable generator \mathcal{G} consisting of sets of the form

$$G = \bigcap_{j=0}^m \{X_{s_j} \in C_j\}, \quad m \geq 1, \quad s_0 = 0 < s_1 < \dots < s_m = s, \quad C_1, \dots, C_m \in \mathcal{B}(\mathbb{R}^d),$$

it is enough to show that for every $G \in \mathcal{G}$

$$\int_G p_{t-s}(X_s, C) d\mathbb{P}^x = \int_G \mathbb{1}_C(X_t) d\mathbb{P}^x.$$

Because of the form of G and since $s = s_m$, this is the same as

$$\int \prod_{j=1}^m \mathbb{1}_{C_j}(X_{s_j}) p_{t-s_m}(X_{s_m}, C) d\mathbb{P}^x = \int \prod_{j=1}^m \mathbb{1}_{C_j}(X_{s_j}) \mathbb{1}_C(X_t) d\mathbb{P}^x.$$

By the definition of the finite dimensional distributions we see

$$\begin{aligned}
 & \int \prod_{j=1}^m \mathbb{1}_{C_j}(X_{s_j}) p_{t-s_m}(X_{s_m}, C) d\mathbb{P}^x \\
 &= T_{s_1} \left[\mathbb{1}_{C_1} T_{s_2-s_1} \left[\mathbb{1}_{C_2} \dots \mathbb{1}_{C_{m-1}} T_{s_m-s_{m-1}} \left[\mathbb{1}_{C_m} p_{t-s_m}(\cdot, C) \right] \dots \right] \right] (x) \\
 &= T_{s_1} \left[\mathbb{1}_{C_1} T_{s_2-s_1} \left[\mathbb{1}_{C_2} \dots \mathbb{1}_{C_{m-1}} T_{s_m-s_{m-1}} \left[\mathbb{1}_{C_m} T_{t-s_m} \mathbb{1}_C \right] \dots \right] \right] (x) \\
 &= p_{s_1, \dots, s_m, t}^x(C_1 \times \dots \times C_{m-1} \times C_m \times C) \\
 &= \int \prod_{j=1}^m \mathbb{1}_{C_j}(X_{s_j}) \mathbb{1}_C(X_t) d\mathbb{P}^x. \quad \square
 \end{aligned}$$

7.7 Remark. Feller processes. Denote by $(T_t)_{t \geq 0}$ a Feller semigroup on $\mathcal{C}_\infty(\mathbb{R}^d)$. Remark 7.6 allows us to construct a d -dimensional Markov process $(X_t)_{t \geq 0}$ whose transition semigroup is $(T_t)_{t \geq 0}$. The process $(X_t)_{t \geq 0}$ is called a *Feller process*. Let us briefly outline this construction: By the Riesz representation theorem, cf. e. g. Rudin [164, Theorem 2.14], there exists a unique (compact inner regular) kernel $p_t(x, \cdot)$ such that

$$T_t u(x) = \int u(y) p_t(x, dy), \quad u \in \mathcal{C}_\infty(\mathbb{R}^d). \quad (7.6)$$

Obviously, (7.6) extends T_t to $\mathcal{B}_b(\mathbb{R}^d)$, and $(T_t)_{t \geq 0}$ becomes a Feller semigroup in the sense of Remark 7.4.

Let $K \subset \mathbb{R}^d$ be compact. Then $U_n := \{x + y : x \in K, y \in \mathbb{B}(0, 1/n)\}$ is open and Ex. 7.8

$$u_n(x) := \frac{d(x, U_n^c)}{d(x, K) + d(x, U_n^c)}, \quad \text{where } d(x, A) := \inf_{a \in A} |x - a|, \quad (7.7)$$

is a sequence of continuous functions such that $\inf_{n \geq 1} u_n = \mathbb{1}_K$. By monotone convergence, $p_t(x, K) = \inf_{n \geq 1} T_t u_n(x)$, and $(t, x) \mapsto p_t(x, K)$ is measurable. Since the kernel is inner regular, we have $p_t(x, C) = \sup\{p_t(x, K) : K \subset C, K \text{ compact}\}$ for all Borel sets $C \subset \mathbb{R}^d$ and, therefore, $(t, x) \mapsto p_t(x, C)$ is measurable.

Obviously, $p_t(x, \mathbb{R}^d) \leq 1$. If $p_t(x, \mathbb{R}^d) < 1$, the following trick can be used to make Remark 7.6 work: Consider the one-point compactification of \mathbb{R}^d by adding the point ∂ and define

$$\begin{cases} p_t^\partial(x, \{\partial\}) := 1 - p_t(x, \mathbb{R}^d) & \text{for } x \in \mathbb{R}^d, t \geq 0, \\ p_t^\partial(\partial, C) := \delta_\partial(C) & \text{for } C \in \mathcal{B}(\mathbb{R}^d \cup \{\partial\}), t \geq 0 \\ p_t^\partial(x, C) := p_t(x, C) & \text{for } x \in \mathbb{R}^d, C \in \mathcal{B}(\mathbb{R}^d), t \geq 0. \end{cases} \quad (7.8)$$

With some effort we can always get a version of $(X_t)_{t \geq 0}$ with right-continuous sample paths. Moreover, a Feller process is a strong Markov process, cf. Section A.5 in the appendix.

7.2 The generator

Let $\phi : [0, \infty) \rightarrow \mathbb{R}$ be a continuous function which satisfies the functional equation $\phi(t)\phi(s) = \phi(t+s)$ and $\phi(0) = 1$. Then there is a unique $a \in \mathbb{R}$ such that $\phi(t) = e^{at}$. Obviously,

$$a = \left. \frac{d^+}{dt} \phi(t) \right|_{t=0} = \lim_{t \rightarrow 0} \frac{\phi(t) - 1}{t}.$$

Since a strongly continuous semigroup $(P_t)_{t \geq 0}$ on a Banach space $(\mathfrak{B}, \|\cdot\|)$ also satisfies the functional equation $P_t P_s u = P_{t+s} u$ it is not too wild a guess that – in an appropriate sense – $P_t u = e^{tA} u$ for some operator A in \mathfrak{B} . In order to keep things technically simple, we will only consider Feller semigroups.

7.8 Definition. Let $(P_t)_{t \geq 0}$ be a Feller semigroup on $\mathcal{C}_\infty(\mathbb{R}^d)$. Then

$$Au := \lim_{t \rightarrow 0} \frac{P_t u - u}{t} \quad \left(\text{the limit is taken in } (\mathcal{C}_\infty(\mathbb{R}^d), \|\cdot\|_\infty) \right) \quad (7.9a)$$

$$\mathfrak{D}(A) := \left\{ u \in \mathcal{C}_\infty(\mathbb{R}^d) \mid \exists g \in \mathcal{C}_\infty(\mathbb{R}^d) : \lim_{t \rightarrow 0} \left\| \frac{P_t u - u}{t} - g \right\|_\infty = 0 \right\} \quad (7.9b)$$

is the (*infinitesimal*) generator of the semigroup $(P_t)_{t \geq 0}$.

Ex. 7.9 Clearly, (7.9a) defines a linear operator $A : \mathfrak{D}(A) \rightarrow \mathcal{C}_\infty(\mathbb{R}^d)$.

7.9 Example. Let $(B_t)_{t \geq 0}$, $B_t = (B_t^1, \dots, B_t^d)$, be a d -dimensional Brownian motion and denote by $P_t u(x) = \mathbb{E}^x u(B_t)$ the transition semigroup. Then

$$\mathcal{C}_\infty^2(\mathbb{R}^d) := \{ u \in \mathcal{C}_\infty(\mathbb{R}^d) : \partial_j u, \partial_j \partial_k u \in \mathcal{C}_\infty(\mathbb{R}^d), \\ j, k = 1, \dots, d \} \subset \mathfrak{D}(A)$$

where

$$A = \frac{1}{2} \Delta = \frac{1}{2} \sum_{j=1}^d \partial_j^2$$

and

$$\partial_j = \partial / \partial x_j.$$

Let $u \in \mathcal{C}_\infty^2(\mathbb{R}^d)$. By the Taylor formula we get for some $\theta = \theta(\omega) \in [0, 1]$

$$\begin{aligned} & \left| \mathbb{E} \left[\frac{u(B_t + x) - u(x)}{t} - \frac{1}{2} \sum_{j=1}^d \partial_j^2 u(x) \right] \right| \\ &= \left| \mathbb{E} \left[\frac{1}{t} \sum_{j=1}^d \partial_j u(x) B_t^j + \frac{1}{2t} \sum_{j,k=1}^d \partial_j \partial_k u(x + \theta B_t) B_t^j B_t^k - \frac{1}{2} \sum_{j=1}^d \partial_j^2 u(x) \right] \right| \\ &= \frac{1}{2} \left| \mathbb{E} \left[\sum_{j,k=1}^d (\partial_j \partial_k u(x + \theta B_t) - \partial_j \partial_k u(x)) \frac{B_t^j B_t^k}{t} \right] \right| \end{aligned}$$

In the last equality we used that $\mathbb{E} B_t^j = 0$ and $\mathbb{E} B_t^j B_t^k = \delta_{jk} t$. Applying the Cauchy–Schwarz inequality two times, we find

$$\begin{aligned} & \leq \frac{1}{2} \mathbb{E} \left[\left(\sum_{j,k=1}^d |\partial_j \partial_k u(x + \theta B_t) - \partial_j \partial_k u(x)|^2 \right)^{1/2} \underbrace{\left(\sum_{j,k=1}^d \frac{(B_t^j B_t^k)^2}{t^2} \right)^{1/2}}_{= \sum_{j=1}^d (B_t^j)^2 / t = |B_t|^2 / t} \right] \\ & \leq \frac{1}{2} \left(\sum_{j,k=1}^d \mathbb{E} \left(|\partial_j \partial_k u(x + \theta B_t) - \partial_j \partial_k u(x)|^2 \right) \right)^{1/2} \left(\frac{\mathbb{E} (|B_t|^4)}{t^2} \right)^{1/2}. \end{aligned}$$

By the scaling property 2.12, $\mathbb{E}(|B_t|^4) = t^2 \mathbb{E}(|B_1|^4) < \infty$. Since


$$|\partial_j \partial_k u(x + \theta B_t) - \partial_j \partial_k u(x)|^2 \leq 4 \|\partial_j \partial_k u\|_\infty^2 \quad \text{for all } j, k = 1, \dots, d,$$

and since the expression on the left converges (uniformly in x) to 0 as $t \rightarrow 0$, we get by dominated convergence

$$\lim_{t \rightarrow 0} \sup_{x \in \mathbb{R}^d} \left| \mathbb{E} \left[\frac{u(B_t + x) - u(x)}{t} - \frac{1}{2} \sum_{j=1}^d \partial_j^2 u(x) \right] \right| = 0.$$

This shows that $\mathcal{C}_\infty^2(\mathbb{R}^d) \subset \mathfrak{D}(A)$ and $A = \frac{1}{2} \Delta$ on $\mathcal{C}_\infty^2(\mathbb{R}^d)$.

In abuse of notation we write $\mathfrak{D}(\Delta)$ for the domain of $\frac{1}{2} \Delta$. We will see in Example 7.20 below that $\mathfrak{D}(\Delta) = \mathcal{C}_\infty^2(\mathbb{R})$ if $d = 1$.

For $d > 1$ we have $\mathfrak{D}(\Delta) \subsetneq \mathcal{C}_\infty^2(\mathbb{R}^d)$, cf. [85, pp. 234–5]. In general, $\mathfrak{D}(\Delta)$ is the Sobolev-type space $\{u \in L^2(\mathbb{R}^d) : u, \Delta u \in \mathcal{C}_\infty(\mathbb{R}^d)\}$ where Δu is understood in the sense of distributions. 

Since every Feller semigroup $(P_t)_{t \geq 0}$ is given by a family of measurable kernels, cf. Remark 7.7, we can define linear operators of the type $L = \int_0^t P_s ds$ in the following

way: For $u \in \mathcal{B}_b(\mathbb{R}^d)$

$$Lu \text{ is a Borel function given by } x \mapsto Lu(x) := \int_0^t \int u(y) p_s(x, dy) ds.$$

The following lemma describes the fundamental relation between generators and semigroups.

7.10 Lemma. *Let $(P_t)_{t \geq 0}$ be a Feller semigroup with generator $(A, \mathcal{D}(A))$.*

a) *For all $u \in \mathcal{D}(A)$ and $t > 0$ we have $P_t u \in \mathcal{D}(A)$. Moreover,*

$$\frac{d}{dt} P_t u = A P_t u = P_t A u \text{ for all } u \in \mathcal{D}(A), t > 0. \quad (7.10)$$

b) *For all $u \in \mathcal{C}_\infty(\mathbb{R}^d)$ we have $\int_0^t P_s u ds \in \mathcal{D}(A)$ and*

$$P_t u - u = A \int_0^t P_s u ds, \quad u \in \mathcal{C}_\infty(\mathbb{R}^d), t > 0 \quad (7.11a)$$

$$= \int_0^t A P_s u ds, \quad u \in \mathcal{D}(A), t > 0 \quad (7.11b)$$

$$= \int_0^t P_s A u ds, \quad u \in \mathcal{D}(A), t > 0. \quad (7.11c)$$

Proof. a) Let $0 < \epsilon < t$ and $u \in \mathcal{D}(A)$. By the triangle inequality, we obtain the semigroup and the contraction properties

$$\begin{aligned} \left\| \frac{P_\epsilon P_t u - P_t u}{\epsilon} - P_t A u \right\|_\infty &= \left\| \frac{P_{t+\epsilon} u - P_t u}{\epsilon} - P_t A u \right\|_\infty \\ &= \left\| P_t \frac{P_\epsilon u - u}{\epsilon} - P_t A u \right\|_\infty \\ &\leq \left\| \frac{P_\epsilon u - u}{\epsilon} - A u \right\|_\infty \xrightarrow{\epsilon \rightarrow 0} 0. \end{aligned}$$

By the definition of the generator, $P_t u \in \mathcal{D}(A)$ and $\frac{d^+}{dt} P_t u = A P_t u = P_t A u$. Similarly,

$$\begin{aligned} &\left\| \frac{P_t u - P_{t-\epsilon} u}{\epsilon} - P_t A u \right\|_\infty \\ &\leq \left\| P_{t-\epsilon} \frac{P_\epsilon u - u}{\epsilon} - P_{t-\epsilon} A u \right\|_\infty + \| P_{t-\epsilon} A u - P_{t-\epsilon} P_\epsilon A u \|_\infty \\ &\leq \left\| \frac{P_\epsilon u - u}{\epsilon} - A u \right\|_\infty + \| A u - P_\epsilon A u \|_\infty \xrightarrow{\epsilon \rightarrow 0} 0. \end{aligned}$$

where we used the strong continuity in the last step.

This shows that we have $\frac{d^-}{dt} P_t u = A P_t u = P_t A u$.

b) Let $u \in \mathcal{B}_b(\mathbb{R}^n)$ and $t, \epsilon > 0$. By Fubini's theorem

$$\begin{aligned} P_\epsilon \int_0^t P_s u(x) ds &= \int_0^t \int_0^t \int u(y) p_s(z, dy) ds p_\epsilon(x, dz) \\ &= \int_0^t \int \int u(y) p_s(z, dy) p_\epsilon(x, dz) ds = \int_0^t P_\epsilon P_s u(x) ds, \end{aligned}$$

and so,

$$\begin{aligned} \frac{P_\epsilon - \text{id}}{\epsilon} \int_0^t P_s u(x) ds &= \frac{1}{\epsilon} \int_0^t (P_{s+\epsilon} u(x) - P_s u(x)) ds \\ &= \frac{1}{\epsilon} \int_t^{t+\epsilon} P_s u(x) ds - \frac{1}{\epsilon} \int_0^\epsilon P_s u(x) ds. \end{aligned}$$

The contraction property and the strong continuity give for $r > 0$

$$\begin{aligned} \left| \frac{1}{\epsilon} \int_r^{r+\epsilon} P_s u(x) ds - P_r u(x) \right| &\leq \frac{1}{\epsilon} \int_r^{r+\epsilon} |P_s u(x) - P_r u(x)| ds \\ &\leq \sup_{r \leq s \leq r+\epsilon} \|P_s u - P_r u\|_\infty \\ &\leq \sup_{s \leq \epsilon} \|P_s u - u\|_\infty \xrightarrow{\epsilon \rightarrow 0} 0. \end{aligned} \tag{7.12}$$

This shows that $\int_0^t P_s u ds \in \mathfrak{D}(A)$ as well as (7.11a). If $u \in \mathfrak{D}(A)$, then

$$\begin{aligned} \int_0^t P_s A u(x) ds &\stackrel{(7.10)}{=} \int_0^t A P_s u(x) ds \\ &\stackrel{(7.10)}{=} \int_0^t \frac{d}{ds} P_s u(x) ds \\ &= P_t u(x) - u(x) \stackrel{(7.11a)}{=} A \int_0^t P_s u(x) ds. \end{aligned}$$

This proves (7.11c) and (7.11b) follows now from part a). □

7.11 Corollary. Let $(P_t)_{t \geq 0}$ be a Feller semigroup with generator $(A, \mathfrak{D}(A))$.

Ex. 19.3

- a) $\mathfrak{D}(A)$ is a dense subset of $\mathcal{C}_\infty(\mathbb{R}^d)$.
b) $(A, \mathfrak{D}(A))$ is a closed operator, i. e.

$$\left. \begin{aligned} (u_n)_{n \geq 1} &\subset \mathfrak{D}(A), \lim_{n \rightarrow \infty} u_n = u, \\ (A u_n)_{n \geq 1} &\text{ is a Cauchy sequence} \end{aligned} \right\} \implies \left\{ \begin{aligned} u &\in \mathfrak{D}(A) \text{ and} \\ A u &= \lim_{n \rightarrow \infty} A u_n. \end{aligned} \right. \tag{7.13}$$

- c) If $(T_t)_{t \geq 0}$ is a further Feller semigroup with generator $(A, \mathfrak{D}(A))$, then $P_t = T_t$.

Proof. a) Let $u \in \mathcal{C}_\infty(\mathbb{R}^d)$. From (7.11a) we see $u_\epsilon := \epsilon^{-1} \int_0^\epsilon P_s u \, ds \in \mathfrak{D}(A)$ and, by an argument similar to (7.12), we find $\lim_{\epsilon \rightarrow 0} \|u_\epsilon - u\|_\infty = 0$.

b) Let $(u_n)_{n \geq 1} \subset \mathfrak{D}(A)$ be such that $\lim_{n \rightarrow \infty} u_n = u$ and $\lim_{n \rightarrow \infty} Au_n = w$ uniformly. Therefore

$$\begin{aligned} P_t u(x) - u(x) &= \lim_{n \rightarrow \infty} P_t u_n(x) - u_n(x) \\ &\stackrel{(7.11c)}{=} \lim_{n \rightarrow \infty} \int_0^t P_s A u_n(x) \, ds = \int_0^t P_s w(x) \, ds. \end{aligned}$$

As in (7.12) we find

$$\lim_{t \rightarrow 0} \frac{P_t u(x) - u(x)}{t} = \lim_{t \rightarrow 0} \frac{1}{t} \int_0^t P_s w(x) \, ds = w(x)$$

uniformly for all x , thus $u \in \mathfrak{D}(A)$ and $Au = w$.

c) Let $0 < s < t$ and assume that $0 < |h| < \min\{t-s, s\}$. Then we have for $u \in \mathfrak{D}(A)$

$$\frac{P_{t-s} T_s u - P_{t-s-h} T_{s-h} u}{h} = \frac{P_{t-s} - P_{t-s-h}}{h} T_s u + P_{t-s-h} \frac{T_s u - T_{s-h} u}{h}.$$

Ex. 7.10 Letting $h \rightarrow 0$, Lemma 7.10 shows that the function $s \mapsto P_{t-s} T_s u$, $u \in \mathfrak{D}(A)$, is differentiable, and

$$\frac{d}{ds} P_{t-s} T_s u = \left(\frac{d}{ds} P_{t-s} \right) T_s u + P_{t-s} \frac{d}{ds} T_s u = -P_{t-s} A T_s u + P_{t-s} A T_s u = 0.$$

Thus,

$$0 = \int_0^t \frac{d}{ds} P_{t-s} T_s u(x) \, ds = P_{t-s} T_s u(x) \Big|_{s=0}^{s=t} = T_t u(x) - P_t u(x).$$

Therefore, $T_t u = P_t u$ for all $t > 0$ and $u \in \mathfrak{D}(A)$. Since $\mathfrak{D}(A)$ is dense in $\mathcal{C}_\infty(\mathbb{R}^d)$, we can approximate $u \in \mathcal{C}_\infty(\mathbb{R}^d)$ by a sequence $(u_j)_{j \geq 1} \subset \mathfrak{D}(A)$, i. e.

$$T_t u = \lim_{j \rightarrow \infty} T_t u_j = \lim_{j \rightarrow \infty} P_t u_j = P_t u. \quad \square$$

7.3 The resolvent

Let $(P_t)_{t \geq 0}$ be a Feller semigroup. Then the integral

$$U_\alpha u(x) := \int_0^\infty e^{-\alpha t} P_t u(x) \, dt, \quad \alpha > 0, \, x \in \mathbb{R}^d, \, u \in \mathcal{B}_b(\mathbb{R}^d), \quad (7.14)$$

exists and defines a linear operator $U_\alpha : \mathcal{B}_b(\mathbb{R}^d) \rightarrow \mathcal{B}_b(\mathbb{R}^d)$.

7.12 Definition. Let $(P_t)_{t \geq 0}$ be a Feller semigroup and $\alpha > 0$. The operator U_α given by (7.14) is the α -potential operator.

The α -potential operators share many properties of the semigroup. In particular, we have the following counterpart of Proposition 7.3.

7.13 Proposition. *Let $(P_t)_{t \geq 0}$ be a Feller semigroup with potential operators $(U_\alpha)_{\alpha > 0}$ and generator $(A, \mathcal{D}(A))$.*

Ex. 7.12

Ex. 7.14

- a) αU_α is conservative: $\alpha U_\alpha 1 = 1$;
- b) αU_α is a contraction on \mathcal{B}_b : $\|\alpha U_\alpha u\|_\infty \leq \|u\|_\infty$ for all $u \in \mathcal{B}_b$;
- c) αU_α is positivity preserving: For $u \in \mathcal{B}_b$ we have $u \geq 0 \Rightarrow \alpha U_\alpha u \geq 0$;
- d) αU_α has the Feller property: If $u \in \mathcal{C}_\infty$, then $\alpha U_\alpha u \in \mathcal{C}_\infty$;
- e) αU_α is strongly continuous on \mathcal{C}_∞ : $\lim_{\alpha \rightarrow \infty} \|\alpha U_\alpha u - u\|_\infty = 0$;
- f) $(U_\alpha)_{\alpha > 0}$ is the resolvent: For all $\alpha > 0$, $\alpha \text{id} - A$ is invertible with a bounded inverse $U_\alpha = (\alpha \text{id} - A)^{-1}$;
- g) Resolvent equation: $U_\alpha u - U_\beta u = (\beta - \alpha)U_\beta U_\alpha u$ for all $\alpha, \beta > 0$ and $u \in \mathcal{B}_b$. In particular, $\lim_{\beta \rightarrow \alpha} \|U_\alpha u - U_\beta u\|_\infty = 0$ for all $u \in \mathcal{B}_b$.
- h) There is a one-to-one relationship between $(P_t)_{t \geq 0}$ and $(U_\alpha)_{\alpha > 0}$.
- i) The resolvent $(U_\alpha)_{\alpha > 0}$ is sub-Markovian if, and only if, the semigroup $(P_t)_{t \geq 0}$ is sub-Markovian: Let $0 \leq u \leq 1$. Then $0 \leq P_t u \leq 1 \quad \forall t \geq 0$ if, and only if, $0 \leq \alpha U_\alpha u \leq 1 \quad \forall \alpha > 0$.

Proof. a)–c) follow immediately from the integral formula (7.14) for $U_\alpha u(x)$ and the corresponding properties of the Feller semigroup $(P_t)_{t \geq 0}$.

d) Since $P_t u \in \mathcal{C}_\infty$ and $|\alpha e^{-\alpha t} P_t u(x)| \leq \alpha e^{-\alpha t} \|u\|_\infty \in L^1(dt)$, this follows from dominated convergence.

e) By the integral representation (7.14) and the triangle inequality we find

$$\|\alpha U_\alpha u - u\|_\infty \leq \int_0^\infty \alpha e^{-\alpha t} \|P_t u - u\|_\infty dt = \int_0^\infty e^{-s} \|P_{s/\alpha} u - u\|_\infty ds.$$

As $\|P_{s/\alpha} u - u\|_\infty \leq 2\|u\|_\infty$ and $\lim_{\alpha \rightarrow \infty} \|P_{s/\alpha} u - u\|_\infty = 0$, the claim follows with dominated convergence.

f) For every $\epsilon > 0$, $\alpha > 0$ and $u \in \mathcal{C}_\infty(\mathbb{R}^d)$ we have

$$\begin{aligned} \frac{1}{\epsilon} (P_\epsilon(U_\alpha u) - U_\alpha u) &= \frac{1}{\epsilon} \int_0^\infty e^{-\alpha t} (P_{t+\epsilon} u - P_t u) dt \\ &= \frac{1}{\epsilon} \int_\epsilon^\infty e^{-\alpha(s-\epsilon)} P_s u ds - \frac{1}{\epsilon} \int_0^\infty e^{-\alpha s} P_s u ds \\ &= \underbrace{\frac{e^{\alpha\epsilon} - 1}{\epsilon}}_{\rightarrow \alpha, \epsilon \rightarrow 0} \int_0^\infty e^{-\alpha s} P_s u ds - \underbrace{\frac{e^{\alpha\epsilon}}{\epsilon} \int_0^\epsilon e^{-\alpha s} P_s u ds}_{\rightarrow u, \epsilon \rightarrow 0 \text{ as in (7.12)}} \end{aligned}$$

This shows that we have in the sense of uniform convergence

$$\lim_{\epsilon \rightarrow 0} \frac{P_\epsilon U_\alpha u - U_\alpha u}{\epsilon} = \alpha U_\alpha u - u,$$

i. e. $U_\alpha u \in \mathfrak{D}(A)$ and $AU_\alpha u = \alpha U_\alpha u - u$. Hence, $(\alpha \text{id} - A)U_\alpha u = u$ for all $u \in \mathcal{C}_\infty$.

Since for $u \in \mathfrak{D}(A)$

$$U_\alpha Au = \int_0^\infty e^{-\alpha t} P_t Au dt \stackrel{(7.11)}{=} A \int_0^\infty e^{-\alpha t} P_t u dt = AU_\alpha u,$$

it is easy to see that $u = (\alpha \text{id} - A)U_\alpha u = U_\alpha(\alpha \text{id} - A)u$.

This shows that U_α is the left- and right-inverse of $\alpha \text{id} - A$. Since U_α is bounded, $\alpha \text{id} - A$ has a bounded inverse, i. e. $\alpha > 0$ is in the resolvent set of A , and $(U_\alpha)_{\alpha > 0}$ is the resolvent.

g) Because of f) we see for $\alpha, \beta > 0$ and $u \in \mathcal{B}_b$

$$U_\alpha u - U_\beta u = U_\beta((\beta \text{id} - A)U_\alpha - \text{id})u = U_\beta(\beta - \alpha)U_\alpha u = (\beta - \alpha)U_\beta U_\alpha u.$$

Since αU_α and βU_β are contractive, we find

$$\|U_\alpha u - U_\beta u\|_\infty = \frac{|\beta - \alpha|}{\beta\alpha} \|\alpha U_\alpha \beta U_\beta u\|_\infty \leq \left| \frac{1}{\alpha} - \frac{1}{\beta} \right| \|u\|_\infty \xrightarrow{\beta \rightarrow \alpha} 0.$$

h) Formula (7.14) shows that $(P_t)_{t \geq 0}$ defines $(U_\alpha)_{\alpha > 0}$. On the other hand, (7.14) means that for every $x \in \mathbb{R}^d$ the function $v(\alpha) := U_\alpha u(x)$ is the Laplace transform of $\phi(t) := P_t u(x)$. Therefore, v determines $\phi(t)$ for Lebesgue almost all $t \geq 0$, cf. [170, Proposition 1.2] for the uniqueness of the Laplace transform. Since $\phi(t) = P_t u(x)$ is continuous for $u \in \mathcal{C}_\infty$, this shows that $P_t u(x)$ is uniquely determined for all $u \in \mathcal{C}_\infty$, and so is $(P_t)_{t \geq 0}$, cf. Remark 7.7.

i) Assume that $u \in \mathcal{B}_b$ satisfies $0 \leq u \leq 1$. From (7.14) we get that $0 \leq P_t u \leq 1$, $t \geq 0$, implies $0 \leq \alpha U_\alpha u \leq 1$. Conversely, let $0 \leq u \leq 1$ and $0 \leq \alpha U_\alpha u \leq 1$ for all $\alpha > 0$. From the resolvent equation g) we see for all $u \in \mathcal{B}_b$ with $u \geq 0$

$$\frac{d}{d\alpha} U_\alpha u(x) = \lim_{\beta \rightarrow \alpha} \frac{U_\alpha u(x) - U_\beta u(x)}{\alpha - \beta} \stackrel{g)}{=} - \lim_{\beta \rightarrow \alpha} U_\beta U_\alpha u(x) = -U_\alpha^2 u(x) \leq 0.$$

Ex. 7.13 Iterating this shows that $(-1)^n \frac{d^n}{d\alpha^n} U_\alpha u(x) \geq 0$. This means that $v(\alpha) := U_\alpha u(x)$ is completely monotone, hence the Laplace transform of a *positive* measure, cf. Bernstein's theorem [170, Theorem 1.4]. As $U_\alpha u(x) = \int_0^\infty e^{-\alpha t} P_t u(x) dt$, we infer that $P_t u(x) \geq 0$ for Lebesgue almost all $t > 0$.

The formula for the derivative of U_α yields that $\frac{d}{d\alpha} \alpha U_\alpha u(x) = (\text{id} - \alpha U_\alpha)u(x)$. By iteration we get $(-1)^{n+1} \frac{d^n}{d\alpha^n} \alpha U_\alpha u(x) = n!(\text{id} - \alpha U_\alpha)u(x)$. Plug in $u \equiv 1$; then it follows that $\Upsilon(\alpha) := 1 - \alpha U_\alpha 1(x) = \int_0^\infty \alpha e^{-\alpha t} (1 - P_t 1(x)) dt$ is completely

monotone. Thus, $P_t 1(x) \leq 1$ for Lebesgue almost all $t > 0$, and using the positivity of P_t for the positive function $1 - u$, we get $P_t u \leq P_t 1 \leq 1$ for almost all $t > 0$.

Since $t \mapsto P_t u(x)$ is continuous if $u \in \mathcal{C}_\infty$, we see that $0 \leq P_t u(x) \leq 1$ for all $t > 0$ if $u \in \mathcal{C}_\infty$ and $0 \leq u \leq 1$, and we extend this to $u \in \mathcal{B}_b$ with the help of Remark 7.7. \square

7.14 Example. Let $(B_t)_{t \geq 0}$ be a Brownian motion and denote by $(U_\alpha)_{\alpha > 0}$ the resolvent. For all $u \in \mathcal{B}_b(\mathbb{R}^d)$ Ex. 7.14

$$U_\alpha u(x) = \begin{cases} \int \frac{e^{-\sqrt{2\alpha}y}}{\sqrt{2\alpha}} u(x+y) dy, & \text{if } d = 1, \\ \int \frac{1}{\pi^{d/2}} \left(\frac{\alpha}{2y^2} \right)^{\frac{d}{4}-\frac{1}{2}} K_{\frac{d}{2}-1}(\sqrt{2\alpha}y) u(x+y) dy, & \text{if } d \geq 2, \end{cases} \quad (7.15)$$

where $K_\nu(x)$ is a Bessel function of the third kind, cf. [74, p. 959, 8.432.6 and p. 967, 8.469.3].

Indeed: Let $d = 1$ and denote by $(P_t)_{t \geq 0}$ the transition semigroup. Using Fubini's theorem we find for $\alpha > 0$ and $u \in \mathcal{B}_b(\mathbb{R})$

$$\begin{aligned} U_\alpha u(x) &= \int_0^\infty e^{-\alpha t} P_t u(x) dt \\ &= \int_0^\infty e^{-\alpha t} \mathbb{E} u(x + B_t) dt \\ &= \int \int_0^\infty e^{-\alpha t} u(x+y) \frac{1}{\sqrt{2\pi t}} e^{-y^2/2t} dt dy \\ &= \int r_\alpha(y^2) u(x+y) dy \end{aligned}$$

where

$$r_\alpha(\eta) = \int_0^\infty \frac{1}{\sqrt{2\pi t}} e^{-(\alpha t + \eta/2t)} dt = \frac{e^{-\sqrt{2\alpha\eta}}}{\sqrt{2\alpha}} \int_0^\infty \sqrt{\frac{\alpha}{\pi t}} e^{-(\sqrt{\alpha t} - \sqrt{\eta/2t})^2} dt.$$

Changing variables according to $t = \eta/(2\alpha s)$ and $dt = -\eta/(2\alpha s^2) ds$ gives

$$r_\alpha(\eta) = \frac{e^{-\sqrt{2\alpha\eta}}}{\sqrt{2\alpha}} \int_0^\infty \sqrt{\frac{\eta}{2\pi s^3}} e^{-(\sqrt{\eta/2s} - \sqrt{\alpha s})^2} ds.$$

If we add these two equalities and divide by 2, we get

$$r_\alpha(\eta) = \frac{e^{-\sqrt{2\alpha\eta}}}{\sqrt{2\alpha}} \frac{1}{\sqrt{\pi}} \int_0^\infty \frac{1}{2} \left(\sqrt{\frac{\alpha}{s}} + \sqrt{\frac{\eta}{2s^3}} \right) e^{-(\sqrt{\alpha s} - \sqrt{\eta/2s})^2} ds.$$

A further change of variables, $y = \sqrt{\alpha s} - \sqrt{\eta/2s}$ and $dy = \frac{1}{2} \left(\sqrt{\frac{\alpha}{s}} + \sqrt{\frac{\eta}{2s^3}} \right) ds$, reveals that

$$r_\alpha(\eta) = \frac{e^{-\sqrt{2\alpha\eta}}}{\sqrt{2\alpha}} \frac{1}{\sqrt{\pi}} \int_{-\infty}^{\infty} e^{-y^2} dy = \frac{e^{-\sqrt{2\alpha\eta}}}{\sqrt{2\alpha}}.$$

For a d -dimensional Brownian motion, $d \geq 2$, a similar calculation leads to the kernel

$$r_\alpha(\eta) = \int_0^\infty \frac{1}{(2\pi t)^{d/2}} e^{-(\alpha t + \eta/2t)} dt = \frac{1}{\pi^{d/2}} \left(\frac{\alpha}{2\eta} \right)^{\frac{d}{4} - \frac{1}{2}} K_{\frac{d}{2}-1}(\sqrt{2\alpha\eta})$$

which cannot be expressed by elementary functions if $d \geq 1$.

The following result is often helpful if we want to determine the domain of the generator.

7.15 Theorem (Dynkin 1956; Reuter 1957). *Let $(A, \mathfrak{D}(A))$ be the generator of a Feller semigroup and assume that $(\mathfrak{A}, \mathfrak{D}(\mathfrak{A}))$ extends $(A, \mathfrak{D}(A))$, i. e.*

$$\mathfrak{D}(A) \subset \mathfrak{D}(\mathfrak{A}) \quad \text{and} \quad \mathfrak{A}|_{\mathfrak{D}(A)} = A. \quad (7.16)$$

If for any $u \in \mathfrak{D}(A)$

$$\mathfrak{A}u = u \implies u = 0, \quad (7.17)$$

then $(A, \mathfrak{D}(A)) = (\mathfrak{A}, \mathfrak{D}(\mathfrak{A}))$.

Proof. Pick $u \in \mathfrak{D}(\mathfrak{A})$ and set

$$g := u - \mathfrak{A}u \quad \text{and} \quad h := (\text{id} - A)^{-1}g \in \mathfrak{D}(A).$$

Then

$$h - \mathfrak{A}h \stackrel{(7.16)}{=} h - Ah = (\text{id} - A)(\text{id} - A)^{-1}g = g \stackrel{\text{def}}{=} u - \mathfrak{A}u.$$

Thus, $h - u = \mathfrak{A}(h - u)$ and, by (7.17), $h = u$. Since $h \in \mathfrak{D}(A)$ we get that $u \in \mathfrak{D}(A)$, and so $\mathfrak{D}(\mathfrak{A}) \subset \mathfrak{D}(A)$. \square

7.4 The Hille-Yosida theorem and positivity

We will now discuss the structure of generators of strongly continuous contraction semigroups and Feller semigroups.

7.16 Theorem (Hille 1948; Yosida 1948; Lumer–Phillips 1961). *A linear operator $(A, \mathfrak{D}(A))$ on a Banach space $(\mathfrak{B}, \|\cdot\|)$ is the infinitesimal generator of a strongly continuous contraction semigroup if, and only if,*

- a) $\mathfrak{D}(A)$ is a dense subset of \mathfrak{B} ;
- b) A is dissipative, i.e. $\forall \lambda > 0, u \in \mathfrak{D}(A) : \|\lambda u - Au\| \geq \lambda \|u\|$;
- c) $\Re(\lambda \text{id} - A) = \mathfrak{B}$ for some (or all) $\lambda > 0$.

We do not give a proof of this theorem but refer the interested reader to Ethier–Kurtz [61] and Pazy [142]. In fact, Theorem 7.16 is too general for transition semigroups of stochastic processes. Let us now show which role is played by the positivity and sub-Markovian property of such semigroups. To keep things simple, we restrict ourselves to the Banach space $\mathfrak{B} = \mathbb{C}_\infty$ and Feller semigroups.

7.17 Lemma. *Let $(P_t)_{t \geq 0}$ be a Feller semigroup with generator $(A, \mathfrak{D}(A))$. Then A satisfies the positive maximum principle (PMP)*

$$u \in \mathfrak{D}(A), u(x_0) = \sup_{x \in \mathbb{R}^d} u(x) \geq 0 \implies Au(x_0) \leq 0. \quad (7.18)$$

For the Laplace operator (7.18) is quite familiar: At a (global) maximum the second derivative is negative.

Proof. Let $u \in \mathfrak{D}(A)$ and assume that $u(x_0) = \sup u \geq 0$. Since P_t preserves positivity, we get $u^+ - u \geq 0$ and $P_t(u^+ - u) \geq 0$ or $P_t u^+ \geq P_t u$. Since $u(x_0)$ is positive, $u(x_0) = u^+(x_0)$ and so

$$\frac{1}{t}(P_t u(x_0) - u(x_0)) \leq \frac{1}{t}(P_t u^+(x_0) - u^+(x_0)) \leq \frac{1}{t}(\|u^+\|_\infty - u^+(x_0)) = 0.$$

Letting $t \rightarrow 0$ therefore shows $Au(x_0) = \lim_{t \rightarrow 0} t^{-1}(P_t u(x_0) - u(x_0)) \leq 0$. \square

In fact, the converse of Lemma 7.17 is also true.

7.18 Lemma. *Assume that the linear operator $(A, \mathfrak{D}(A))$ satisfies conditions a) and c) of Theorem 7.16 and the PMP (7.18). Then 7.16 b) holds, and both the resolvent $U_\lambda = (\lambda \text{id} - A)^{-1}$ and the semigroup generated by A are positivity preserving.*

Proof. Let $u \in \mathfrak{D}(A)$ and x_0 be such that $|u(x_0)| = \sup_{x \in \mathbb{R}^d} |u(x)|$. We may assume that $u(x_0) \geq 0$, otherwise we could consider $-u$. Then we find for all $\lambda > 0$

$$\|\lambda u - Au\|_\infty \geq \lambda u(x_0) - \underbrace{Au(x_0)}_{\leq 0 \text{ by PMP}} \geq \lambda u(x_0) = \lambda \|u\|_\infty.$$

This is 7.16 b). Because of Proposition 7.13 h) it is enough to show that U_λ preserves positivity. If we use the PMP for $-u$, we get Feller's minimum principle:

$$v \in \mathfrak{D}(A), \quad v(x_0) = \inf_{x \in \mathbb{R}^d} v(x) \leq 0 \implies Av(x_0) \geq 0.$$

Now assume that $v \geq 0$ and let x_0 be the minimum point of $U_\lambda v$. Since $U_\lambda v$ vanishes at infinity, it is clear that $U_\lambda v(x_0) \leq 0$. By Feller's minimum principle

$$\lambda U_\lambda v(x_0) - v(x_0) \stackrel{7.13 f)}{=} AU_\lambda v(x_0) \geq 0$$

and, therefore, $\lambda U_\lambda v(x) \geq \inf_x \lambda U_\lambda v(x) = \lambda U_\lambda v(x_0) \geq v(x_0) \geq 0$. \square

The next result shows that positivity allows us to determine the domain of the generator using pointwise rather than uniform convergence.

7.19 Theorem. *Let $(P_t)_{t \geq 0}$ be a Feller semigroup generated by $(A, \mathfrak{D}(A))$. Then*

$$\mathfrak{D}(A) = \left\{ u \in \mathcal{C}_\infty(\mathbb{R}^d) \mid \exists g \in \mathcal{C}_\infty(\mathbb{R}^d) \forall x : \lim_{t \rightarrow 0} \frac{P_t u(x) - u(x)}{t} = g(x) \right\}. \quad (7.19)$$

Proof. We define the *weak generator* of $(P_t)_{t \geq 0}$ as

$$A_w u(x) := \lim_{t \rightarrow 0} \frac{P_t u(x) - u(x)}{t} \quad \text{for all } u \in \mathfrak{D}(A_w) \text{ and } x \in \mathbb{R}^d,$$

where $\mathfrak{D}(A_w)$ is set on the right-hand side of (7.19). Clearly, $(A_w, \mathfrak{D}(A_w))$ extends $(A, \mathfrak{D}(A))$ and the calculation in the proof of Lemma 7.17 shows that $(A_w, \mathfrak{D}(A_w))$ also satisfies the positive maximum principle. Using the argument of (the first part of the proof of) Lemma 7.18 we infer that $(A_w, \mathfrak{D}(A_w))$ is dissipative. In particular, we find for all $u \in \mathfrak{D}(A_w)$

$$A_w u = u \implies 0 = \|u - A_w u\| \geq \|u\| \implies \|u\| = 0 \implies u = 0.$$

Therefore, Theorem 7.15 shows that $(A_w, \mathfrak{D}(A_w)) = (A, \mathfrak{D}(A))$. \square

7.20 Example. Let $(A, \mathfrak{D}(A))$ be the generator of the transition semigroup $(P_t)_{t \geq 0}$ of a BM^d. We know from Example 7.9 that $\mathcal{C}_\infty^2(\mathbb{R}^d) \subset \mathfrak{D}(A)$.

If $d = 1$, then $\mathfrak{D}(A) = \mathcal{C}_\infty^2(\mathbb{R}) = \{u \in \mathcal{C}_\infty(\mathbb{R}) : u, u', u'' \in \mathcal{C}_\infty(\mathbb{R})\}$.

Ex. 7.11 *Proof.* We know from Proposition 7.13 f) that U_α is the inverse of $\alpha \text{id} - A$. In particular, $U_\alpha(\mathcal{C}_\infty(\mathbb{R})) = \mathfrak{D}(A)$. This means that any $u \in \mathfrak{D}(A)$ can be written as $u = U_\alpha f$ for some $f \in \mathcal{C}_\infty(\mathbb{R}^d)$.

Using the formula (7.15) for U_α from Example 7.14, we see for all $f \in \mathcal{C}_\infty(\mathbb{R})$

$$U_\alpha f(x) = \frac{1}{\sqrt{2\alpha}} \int_{-\infty}^{\infty} e^{-\sqrt{2\alpha}|y-x|} f(y) dy.$$

By the differentiability lemma for parameter-dependent integrals, cf. [169, Theorem 11.5], we find

$$\begin{aligned} \frac{d}{dx} U_\alpha f(x) &= \int_{-\infty}^{\infty} e^{-\sqrt{2\alpha}|y-x|} \operatorname{sgn}(y-x) f(y) dy \\ &= \int_x^{\infty} e^{-\sqrt{2\alpha}(y-x)} f(y) dy - \int_{-\infty}^x e^{-\sqrt{2\alpha}(x-y)} f(y) dy, \\ \frac{d^2}{dx^2} U_\alpha f(x) &= \sqrt{2\alpha} \int_{-\infty}^{\infty} e^{-\sqrt{2\alpha}|y-x|} f(y) dy - 2f(x). \end{aligned}$$

Since $\left| \frac{d}{dx} U_\alpha f(x) \right| \leq 2\sqrt{2\alpha} U_\alpha |f|(x)$ and $\frac{d^2}{dx^2} U_\alpha f(x) = 2\alpha U_\alpha f(x) - 2f(x)$, Proposition 7.13 d) shows that $U_\alpha f \in \mathcal{C}_\infty(\mathbb{R}^d)$, thus $\mathcal{C}_\infty^2(\mathbb{R}^d) \supset \mathfrak{D}(A)$. \square

If $d > 1$, these arguments are no longer valid, we have $\mathcal{C}_\infty^2(\mathbb{R}^d) \subsetneq \mathfrak{D}(\Delta)$ and

$$\mathfrak{D}(\Delta) = \{u \in \mathcal{C}_\infty(\mathbb{R}^d) : (\text{weak derivative}) \Delta u \in \mathcal{C}_\infty(\mathbb{R}^d)\},$$

see e. g. Itô–McKean [85, pp. 234–5].



7.5 Dynkin's characteristic operator

We will now give a probabilistic characterization of the infinitesimal generator. As so often, a simple martingale relation turns out to be extremely helpful. The following theorem should be compared with Theorem 5.6. Recall that a Feller process is a (strong²) Markov process with right-continuous trajectories whose transition semigroup $(P_t)_{t \geq 0}$ is a Feller semigroup. Of course, a Brownian motion is a Feller process.

7.21 Theorem. *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a Feller process on \mathbb{R}^d with transition semigroup $(P_t)_{t \geq 0}$ and generator $(A, \mathfrak{D}(A))$. Then* Ex. 7.15

$$M_t^u := u(X_t) - u(x) - \int_0^t Au(X_r) dr \quad \text{for all } u \in \mathfrak{D}(A)$$

is an \mathcal{F}_t martingale.

² Feller processes have the strong Markov property, see Theorem A.25 in the appendix

Proof. Let $u \in \mathfrak{D}(A)$, $x \in \mathbb{R}^d$ and $s, t > 0$. By the Markov property (6.4c),

$$\begin{aligned}
 & \mathbb{E}^x (M_{s+t}^u \mid \mathcal{F}_t) \\
 &= \mathbb{E}^x \left(u(X_{s+t}) - u(x) - \int_0^{s+t} Au(X_r) dr \mid \mathcal{F}_t \right) \\
 &= \mathbb{E}^{X_t} u(X_s) - u(x) - \int_0^t Au(X_r) dr - \mathbb{E}^x \left(\int_t^{s+t} Au(X_r) dr \mid \mathcal{F}_t \right) \\
 &= P_s u(X_t) - u(x) - \int_0^t Au(X_r) dr - \mathbb{E}^{X_t} \left(\int_0^s Au(X_r) dr \right) \\
 &= P_s u(X_t) - u(x) - \int_0^t Au(X_r) dr - \int_0^s \mathbb{E}^{X_t} (Au(X_r)) dr \\
 &= P_s u(X_t) - u(x) - \int_0^t Au(X_r) dr - \underbrace{\int_0^s P_r Au(X_t) dr}_{= P_s u(X_t) - u(X_t) \text{ by (7.11c)}} \\
 &= u(X_t) - u(x) - \int_0^t Au(X_r) dr = M_t^u. \quad \square
 \end{aligned}$$

Ex. 7.16 The following result, due to Dynkin, could be obtained from Theorem 7.21 by optional stopping, but we prefer to give an independent proof.

7.22 Proposition (Dynkin 1956). *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a Feller process on \mathbb{R}^d with transition semigroup $(P_t)_{t \geq 0}$, resolvent $(U_\alpha)_{\alpha > 0}$ and generator $(A, \mathfrak{D}(A))$. If σ is an a. s. finite \mathcal{F}_t stopping time, then*

$$U_\alpha f(x) = \mathbb{E}^x \left(\int_0^\sigma e^{-\alpha t} f(X_t) dt \right) + \mathbb{E}^x \left(e^{-\alpha \sigma} U_\alpha f(X_\sigma) \right), \quad f \in \mathcal{B}_b(\mathbb{R}^d) \quad (7.20)$$

for all $x \in \mathbb{R}^d$. If $\mathbb{E}^x \sigma < \infty$, then Dynkin's formula holds

$$\mathbb{E}^x u(X_\sigma) - u(x) = \mathbb{E}^x \left(\int_0^\sigma Au(X_t) dt \right), \quad u \in \mathfrak{D}(A). \quad (7.21)$$

Proof. Since $\sigma < \infty$ a. s., we get

$$\begin{aligned}
 U_\alpha f(x) &= \mathbb{E}^x \left(\int_0^\infty e^{-\alpha t} f(X_t) dt \right) \\
 &= \mathbb{E}^x \left(\int_0^\sigma e^{-\alpha t} f(X_t) dt \right) + \mathbb{E}^x \left(\int_\sigma^\infty e^{-\alpha t} f(X_t) dt \right).
 \end{aligned}$$

Using the strong Markov property (6.7), the second integral becomes

$$\begin{aligned}
 \mathbb{E}^x \left(\int_{\sigma}^{\infty} e^{-\alpha t} f(X_t) dt \right) &= \mathbb{E}^x \left(\int_0^{\infty} e^{-\alpha(t+\sigma)} f(X_{t+\sigma}) dt \right) \\
 &= \int_0^{\infty} \mathbb{E}^x \left(e^{-\alpha(t+\sigma)} \mathbb{E}^x [f(X_{t+\sigma}) | \mathcal{F}_{\sigma+}] \right) dt \\
 &= \int_0^{\infty} e^{-\alpha t} \mathbb{E}^x (e^{-\alpha\sigma} \mathbb{E}^{X_{\sigma}} f(X_t)) dt \\
 &= \mathbb{E}^x \left(e^{-\alpha\sigma} \int_0^{\infty} e^{-\alpha t} P_t f(X_{\sigma}) dt \right) \\
 &= \mathbb{E}^x (e^{-\alpha\sigma} U_{\alpha} f(X_{\sigma})).
 \end{aligned}$$

If $u \in \mathcal{D}(A)$, there is some $f \in \mathcal{C}_{\infty}(\mathbb{R}^d)$ such that $u = U_{\alpha} f$, hence $f = \alpha u - Au$. Using the first identity, we find by dominated convergence

$$\begin{aligned}
 u(x) &= \mathbb{E}^x \left(\int_0^{\sigma} e^{-\alpha t} f(X_t) dt \right) + \mathbb{E}^x (e^{-\alpha\sigma} u(X_{\sigma})) \\
 &= \mathbb{E}^x \left(\int_0^{\sigma} e^{-\alpha t} (\alpha u(X_t) - Au(X_t)) dt \right) + \mathbb{E}^x (e^{-\alpha\sigma} u(X_{\sigma})) \\
 &\xrightarrow{\alpha \rightarrow 0} -\mathbb{E}^x \left(\int_0^{\sigma} Au(X_t) dt \right) + \mathbb{E}^x (u(X_{\sigma})). \quad \square
 \end{aligned}$$

Assume that the filtration $(\mathcal{F}_t)_{t \geq 0}$ is right-continuous. Then the first hitting time of the open set $\mathbb{R}^d \setminus \{x\}$, $\tau_x := \inf\{t > 0 : X_t \neq x\}$, is a stopping time, cf. Lemma 5.7. One can show that $\mathbb{P}^x(\tau_x \geq t) = \exp(-\lambda(x)t)$ for some exponent $\lambda(x) \in [0, \infty]$, see Theorem A.26 in the appendix. For a Brownian motion we have $\lambda(x) \equiv \infty$, since for $t > 0$

$$\mathbb{P}^x(\tau_x > t) = \mathbb{P}^x \left(\sup_{s \leq t} |B_s - x| = 0 \right) \leq \mathbb{P}^x(|B_t - x| = 0) = \mathbb{P}(B_t = 0) = 0.$$

Under \mathbb{P}^x the random time τ_x is the waiting time of the process at the starting point. Therefore, $\lambda(x)$ tells us how quickly a process leaves its starting point, and we can use it to give a classification of the starting points.

7.23 Definition. Let $(X_t)_{t \geq 0}$ be a Feller process. A point $x \in \mathbb{R}^d$ is said to be

- a) an *exponential holding point* if $0 < \lambda(x) < \infty$,
- b) an *instantaneous point* if $\lambda(x) = \infty$,
- c) an *absorbing point* if $\lambda(x) = 0$.

7.24 Lemma. Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a Feller process with a right-continuous filtration, transition semigroup $(P_t)_{t \geq 0}$ and generator $(A, \mathcal{D}(A))$. If x_0 is not absorbing, then

- a) *there exists some $u \in \mathfrak{D}(A)$ such that $Au(x_0) \neq 0$;*
b) *there is an open neighbourhood $U = U(x_0)$ of x_0 such that for $V := \mathbb{R}^d \setminus \overline{U}$ the first hitting time $\tau_V := \inf\{t > 0 : X_t \in V\}$ satisfies $\mathbb{E}^{x_0} \tau_V < \infty$.*

Proof. a) We show that $Au(x_0) = 0$ for all $u \in \mathfrak{D}(A)$ implies that x_0 is an absorbing point. Let $u \in \mathfrak{D}(A)$. By Lemma 7.10 a) $w = P_t u \in \mathfrak{D}(A)$ and

$$\frac{d}{dt} P_t u(x_0) = A P_t u(x_0) = 0 \quad \text{for all } t \geq 0.$$

Thus, (7.11b) shows that $P_t u(x_0) = u(x_0)$. Since $\mathfrak{D}(A)$ is dense in $\mathcal{C}_\infty(\mathbb{R}^d)$ we get $\mathbb{E}^{x_0} u(X_t) = P_t u(x_0) = u(x_0)$ for all $u \in \mathcal{C}_\infty(\mathbb{R}^d)$. Using the sequence from (7.7), we can approximate $x \mapsto \mathbb{1}_{\{x_0\}}(x)$ from above by $(u_n)_{n \geq 0} \subset \mathcal{C}_\infty(\mathbb{R}^d)$. By monotone convergence,

$$\mathbb{P}^{x_0}(X_t = x_0) = \lim_{n \rightarrow \infty} \underbrace{\mathbb{E}^{x_0} u_n(X_t)}_{=u_n(x_0)} = \lim_{n \rightarrow \infty} u_n(x_0) = \mathbb{1}_{\{x_0\}}(x_0) = 1.$$

Ex. 7.17 Since $t \mapsto X_t$ is a.s. right-continuous, we find

$$1 = \mathbb{P}^{x_0}(X_q = x_0 \quad \forall q \in \mathbb{Q} \cap [0, \infty)) = \mathbb{P}^{x_0}(X_t = x_0 \quad \forall t \in [0, \infty))$$

which means that x_0 is an absorbing point.

- b) If x_0 is not absorbing, then part a) shows that there is some $u \in \mathfrak{D}(A)$ such that $Au(x_0) > 0$. Since Au is continuous, there exists some $\epsilon > 0$ and an open neighbourhood $U = U(x_0, \epsilon)$ such that $Au|_U \geq \epsilon > 0$. Let τ_V be the first hitting time of the open set $V = \mathbb{R}^d \setminus \overline{U}$. From Dynkin's formula (7.21) with $\sigma = \tau_V \wedge n$, $n \geq 1$, we deduce

$$\epsilon \mathbb{E}^{x_0}(\tau_V \wedge n) \leq \mathbb{E}^{x_0} \left(\int_0^{\tau_V \wedge n} Au(X_s) ds \right) = \mathbb{E}^{x_0} u(X_{\tau_V \wedge n}) - u(x) \leq 2\|u\|_\infty.$$

Finally, monotone convergence shows that $\mathbb{E}^{x_0} \tau_V \leq 2\|u\|_\infty/\epsilon < \infty$. \square

7.25 Definition (Dynkin 1956). Let $(X(t))_{t \geq 0}$ be a Feller process and $\tau_r^x := \tau_{\overline{\mathbb{B}}^c(x, r)}$ the first hitting time of the set $\overline{\mathbb{B}}^c(x, r)$. Dynkin's characteristic operator is the linear operator defined by

$$\mathfrak{A}u(x) := \begin{cases} \lim_{r \rightarrow 0} \frac{\mathbb{E}^x u(X(\tau_r^x)) - u(x)}{\mathbb{E}^x \tau_r^x}, & \text{if } x \text{ is not absorbing,} \\ 0, & \text{if } x \text{ is absorbing,} \end{cases} \quad (7.22)$$

on the set $\mathfrak{D}(\mathfrak{A})$ consisting of all $u \in \mathcal{B}_b(\mathbb{R}^d)$ such that the limit in (7.22) exists for

each non-absorbing point $x \in \mathbb{R}^d$.

Note that, by Lemma 7.24, $\mathbb{E}^x \tau_r^x < \infty$ for sufficiently small $r > 0$.

7.26 Theorem. *Let $(X_t)_{t \geq 0}$ be a Feller process with generator $(A, \mathfrak{D}(A))$ and characteristic operator $(\mathfrak{A}, \mathfrak{D}(\mathfrak{A}))$.*

- a) \mathfrak{A} is an extension of A ;
- b) $(\mathfrak{A}, \mathfrak{D}) = (A, \mathfrak{D}(A))$ if $\mathfrak{D} = \{u \in \mathfrak{D}(\mathfrak{A}) : \mathfrak{A}u \in \mathcal{C}_\infty(\mathbb{R}^d)\}$.

Proof. a) Let $u \in \mathfrak{D}(A)$ and assume that x is not an absorbing point. Since $Au \in \mathcal{C}_\infty(\mathbb{R}^d)$, there exists for every $\epsilon > 0$ an open neighbourhood $U_0 = U_0(\epsilon, x)$ of x such that

$$|Au(y) - Au(x)| < \epsilon \quad \text{for all } y \in U_0.$$

Since x is not absorbing, we find some open neighbourhood $U_1 \subset \overline{U_1} \subset U_0$ of x such that $\mathbb{E}^x \tau_{\overline{U_1}}^x < \infty$. By Dynkin's formula (7.21) for the stopping time $\sigma = \tau_{\mathbb{B}^c(x,r)}^x$ where $\mathbb{B}(x, r) \subset U_1 \subset U_0$ we see

$$\begin{aligned} & \left| \mathbb{E}^x u(X_{\tau_{\mathbb{B}^c(x,r)}^x}) - u(x) - Au(x) \mathbb{E} \tau_{\mathbb{B}^c(x,r)}^x \right| \\ & \leq \mathbb{E}^x \left(\int_{[0, \tau_{\mathbb{B}^c(x,r)}^x)} \underbrace{|Au(X_s) - Au(x)|}_{\leq \epsilon} ds \right) \leq \epsilon \mathbb{E}^x \tau_{\mathbb{B}^c(x,r)}^x. \end{aligned}$$

Thus, $\lim_{r \rightarrow 0} \left(\mathbb{E}^x u(X_{\tau_{\mathbb{B}^c(x,r)}^x}) - u(x) \right) / \mathbb{E}^x \tau_{\mathbb{B}^c(x,r)}^x = Au(x)$.

If x is an absorbing point, $Au(x) = 0$ by Lemma 7.24 a), hence $Au = \mathfrak{A}u$ for all $u \in \mathfrak{D}(A)$.

- b) Since $(\mathfrak{A}, \mathfrak{D})$ satisfies the positive maximum principle (7.18), the claim follows like Theorem 7.19. \square

Let us finally show that processes with continuous sample paths are generated by differential operators.

7.27 Definition. Let $L : \mathfrak{D}(L) \subset \mathcal{B}_b(\mathbb{R}^d) \rightarrow \mathcal{B}_b(\mathbb{R}^d)$ be a linear operator. It is called *local* if for every $x \in \mathbb{R}^d$ and $u, w \in \mathfrak{D}(L)$ satisfying $u|_{\mathbb{B}(x, \epsilon)} \equiv w|_{\mathbb{B}(x, \epsilon)}$ in some open neighbourhood $\mathbb{B}(x, \epsilon)$ of x we have $Lu(x) = Lw(x)$.

Typical local operators are differential operators and multiplication operators, e. g. $u \mapsto \sum_{i,j} a_{ij}(\cdot) \partial_i \partial_j u(\cdot) + \sum_j b_j(\cdot) \partial_j u(\cdot) + c(\cdot) u(\cdot)$. A typical non-local operator is an integral operator, e. g. $u \mapsto \int_{\mathbb{R}^d} k(x, y) u(y) dy$. A deep result by J. Peetre says that local operators are essentially differential operators:

7.28 Theorem (Peetre 1960). *Let $L : \mathcal{C}_c^\infty(\mathbb{R}^d) \rightarrow \mathcal{C}_c^k(\mathbb{R}^d)$ be a linear operator where $k \geq 0$ is fixed. If $\text{supp } Lu \subset \text{supp } u$, then $Lu = \sum_{\alpha \in \mathbb{N}_0^d} a_\alpha(\cdot) \frac{\partial^\alpha}{\partial x^\alpha} u$ with finitely many, uniquely determined distributions $a_\alpha \in \mathcal{D}'(\mathbb{R}^d)$ (i. e. the topological dual of $\mathcal{C}_c^\infty(\mathbb{R}^d)$) which are locally represented by functions of class \mathcal{C}^k .*



Let L be as in Theorem 7.28 and assume that $L = A|_{\mathcal{C}_c^\infty(\mathbb{R}^d)}$ where $(A, \mathcal{D}(A))$ is the generator of a Markov process. Then L has to satisfy the positive maximum principle (7.18). In particular, L is at most a *second order differential operator*. This follows from the fact that we can find test functions $\phi \in \mathcal{C}_c^\infty(\mathbb{R}^d)$ such that x_0 is a global maximum while $\partial_j \partial_k \partial_l \phi(x_0)$ has arbitrary sign.

Every Feller process with continuous sample paths has a local generator.

7.29 Theorem. *Let $(X_t)_{t \geq 0}$ be a Feller process with continuous sample paths. Then the generator $(A, \mathcal{D}(A))$ is a local operator.*

Proof. Because of Theorem 7.26 it is enough to show that $(\mathfrak{A}, \mathcal{D}(\mathfrak{A}))$ is a local operator. If x is absorbing, this is clear. Assume, therefore, that x is not absorbing. If the functions $u, w \in \mathcal{D}(\mathfrak{A})$ coincide in $\mathbb{B}(x, \epsilon)$ for some $\epsilon > 0$, we know for all $r < \epsilon$

$$u(X(\tau_{\overline{\mathbb{B}}^c(x,r)})) = w(X(\tau_{\overline{\mathbb{B}}^c(x,r)}))$$

since, by the continuity of the trajectories, $X(\tau_{\overline{\mathbb{B}}^c(x,r)}) \in \overline{\mathbb{B}}(x, r) \subset \mathbb{B}(x, \epsilon)$. Thus, $\mathfrak{A}u(x) = \mathfrak{A}w(x)$. \square

If the domain $\mathcal{D}(A)$ is sufficiently rich, e. g. if it contains the test functions $\mathcal{C}_c^\infty(\mathbb{R}^d)$, then the local property of A follows from the speed of the process moving away from its starting point.

7.30 Theorem. *Let $(X_t)_{t \geq 0}$ be a Feller process such that the test functions $\mathcal{C}_c^\infty(\mathbb{R}^d)$ are in the domain of the generator $(A, \mathcal{D}(A))$. Then $(A, \mathcal{C}_c^\infty(\mathbb{R}^d))$ is a local operator if, and only if,*

$$\forall \epsilon > 0, r > 0, K \subset \mathbb{R}^d \text{ compact} : \lim_{h \rightarrow 0} \sup_{t \leq h} \sup_{x \in K} \frac{1}{h} \mathbb{P}^x(r > |X_t - x| > \epsilon) = 0. \quad (7.23)$$

Proof. Fix $x \in \mathbb{R}^d$ and assume that $u, w \in \mathcal{C}_c^\infty(\mathbb{R}^d) \subset \mathcal{D}(A)$ coincide on $\mathbb{B}(x, \epsilon)$ for some $\epsilon > 0$. Moreover, pick r in such a way that $\text{supp } u \cup \text{supp } w \subset \mathbb{B}(x, r)$. Then

$$\begin{aligned} |Au(x) - Aw(x)| &= \lim_{t \rightarrow 0} \left| \frac{\mathbb{E}^x u(X_t) - u(x)}{t} - \frac{\mathbb{E}^x w(X_t) - w(x)}{t} \right| \\ &= \lim_{t \rightarrow 0} \frac{1}{t} |\mathbb{E}^x (u(X_t) - w(X_t))| \\ &\leq \lim_{t \rightarrow 0} \frac{1}{t} \int_{\mathbb{R}^d} \underbrace{|u(y) - w(y)|}_{=0 \text{ for } |y-x| < \epsilon \text{ or } |y-x| \geq r} \mathbb{P}^x(X_t \in dy) \\ &\leq \|u - w\|_\infty \lim_{t \rightarrow 0} \frac{1}{t} \mathbb{P}^x(r > |X_t - x| > \epsilon) \stackrel{(7.23)}{=} 0. \end{aligned}$$

Conversely, assume that $(A, \mathcal{C}_c^\infty(\mathbb{R}^d))$ is local. Fix $\epsilon > 0$, $r > 3\epsilon$, pick any compact set K and cover it by finitely many balls $\mathbb{B}(x_j, \epsilon)$, $j = 1, \dots, N$.

Since $\mathcal{C}_c^\infty(\mathbb{R}^d) \subset \mathcal{D}(A)$, there exist functions $u_j \in \mathcal{C}_c^\infty(\mathbb{R}^d)$, $j = 1, \dots, N$, with Ex. 7.8

$$u_j \geq 0, \quad u_j|_{\mathbb{B}(x_j, \epsilon)} \equiv 0 \quad \text{and} \quad u_j|_{\overline{\mathbb{B}(x_j, r+\epsilon)} \setminus \mathbb{B}(x_j, 2\epsilon)} \equiv 1$$

and

$$\lim_{t \rightarrow 0} \left\| \frac{1}{t} (P_t u_j - u_j) - Au_j \right\|_\infty = 0 \quad \text{for all } j = 1, \dots, N.$$

Therefore, there is for every $\delta > 0$ and all $j = 1, \dots, N$ some $h = h_\delta > 0$ such that for all $t < h$ and all $x \in \mathbb{R}^d$

$$\frac{1}{t} |\mathbb{E}^x u_j(X_t) - u_j(x) - t Au_j(x)| \leq \delta.$$

If $x \in \mathbb{B}(x_j, \epsilon)$, then $u_j(x) = 0$ and, by locality, $Au_j(x) = 0$. Thus

$$\frac{1}{t} \mathbb{E}^x u_j(X_t) \leq \delta;$$

for $x \in \mathbb{B}(x_j, \epsilon)$ and $r > |X_t - x| > 3\epsilon$ we have $r + \epsilon > |X_t - x_j| > 2\epsilon$, therefore

$$\frac{1}{t} \mathbb{P}^x(r > |X_t - x| > 3\epsilon) \leq \frac{1}{t} \mathbb{P}^x(r + \epsilon > |X_t - x_j| > 2\epsilon) \leq \frac{1}{t} P_t u_j(x) \leq \delta.$$

This proves

$$\begin{aligned} &\overline{\lim}_{h \rightarrow 0} \frac{1}{h} \sup_{x \in K} \sup_{t \leq h} \mathbb{P}^x(r > |X_t - x| > 3\epsilon) \\ &\leq \overline{\lim}_{h \rightarrow 0} \sup_{x \in K} \sup_{t \leq h} \frac{1}{t} \mathbb{P}^x(r > |X_t - x| > 3\epsilon) \leq \delta. \end{aligned}$$

□



7.31 Further reading. Semigroups in probability theory are nicely treated in the first chapter of [61]; the presentation in (the first volume of) [54] is still up-to-date. Both sources approach things from the analysis side. The survey paper [82] gives a brief self-contained introduction from a probabilistic point of view. The Japanese school has a long tradition approaching stochastic processes through semigroups, see for example [83] and [84]. We did not touch the topic of quadratic forms induced by a generator, so-called Dirichlet forms. The classic text is [70].

[54] Dynkin: *Markov Processes* (volume 1).

[61] Ethier, Kurtz: *Markov Processes. Characterization and Convergence*.

[70] Fukushima, Oshima, Takeda: *Dirichlet Forms and Symmetric Markov Processes*.

[82] Itô: *Semigroups in Probability Theory*.

[83] Itô: *Stochastic Processes*.

[84] Itô: *Essentials of Stochastic Processes*.

Problems

1. Show that $\mathcal{C}_\infty := \{u : \mathbb{R}^d \rightarrow \mathbb{R} : \text{continuous and } \lim_{|x| \rightarrow \infty} u(x) = 0\}$ equipped with the uniform topology is a Banach space. Show that in this topology \mathcal{C}_∞ is the closure of \mathcal{C}_c , i. e. the family of continuous functions with compact support.
2. Let $(B_t)_{t \geq 0}$ be a BM^d (or a Feller process) with transition semigroup $(P_t)_{t \geq 0}$. Show, using arguments similar to those in the proof of Proposition 7.3 f), that $(t, x, u) \mapsto P_t u(x)$ is continuous. As usual, we equip $[0, \infty) \times \mathbb{R}^d \times \mathcal{C}_\infty(\mathbb{R}^d)$ with the norm $|t| + |x| + \|u\|_\infty$.
3. Let $(X_t)_{t \geq 0}$ be a d -dimensional Feller process and let $f, g \in \mathcal{C}_\infty(\mathbb{R}^d)$. Show that the function $x \mapsto \mathbb{E}^x(f(X_t)g(X_{t+s}))$ is also in $\mathcal{C}_\infty(\mathbb{R}^d)$.

Hint: Markov property.

4. Let $(B_t)_{t \geq 0}$ be a BM^d and set $u(t, x) := P_t u(x)$. Adapt the proof of Proposition 7.3 g) and show that for $u \in \mathcal{B}_b(\mathbb{R}^d)$ the function $u(t, \cdot) \in \mathcal{C}^\infty$ and $u \in \mathcal{C}^{1,2}$ (i. e. once continuously differentiable in t , twice continuously differentiable in x).
5. Let $(P_t)_{t \geq 0}$ be the transition semigroup of a BM^d and denote by L^p , $1 \leq p < \infty$, the space of p th power integrable functions with respect to Lebesgue measure on \mathbb{R}^d . Set $u_n := (-n) \wedge u \vee n$ for every $u \in L^p$. Verify that
 - (a) $u_n \in \mathcal{B}_b(\mathbb{R}^d) \cap L^p(\mathbb{R}^d)$ and $L^p\text{-}\lim_{n \rightarrow \infty} u_n = u$.
 - (b) $\tilde{P}_t u := \lim_{n \rightarrow \infty} P_t u_n$ extends P_t and gives a contraction semigroup on L^p .
Hint: Use Young's inequality for convolutions, see e. g. [169, Theorem 14.6].
 - (c) \tilde{P}_t is *sub-Markovian*, i. e. if $u \in L^p$ and $0 \leq u \leq 1$ a. e., then $0 \leq \tilde{P}_t u \leq 1$ a. e.

- (d) \tilde{P}_t is *strongly continuous*, i. e. $\lim_{t \rightarrow 0} \|P_t u - u\|_{L^p} = 0$ for all $u \in L^p$.
Hint: For $u \in L^p$, $y \mapsto \|u(\cdot + y) - u\|_{L^p}$ is continuous, cf. [169, Theorem 14.8].
6. Let $(T_t)_{t \geq 0}$ be a Markov semigroup given by $T_t u(x) = \int_{\mathbb{R}^d} u(y) p_t(x, dy)$ where $p_t(x, C)$ is a kernel in the sense of Remark 7.6. Show that the semigroup property of $(T_t)_{t \geq 0}$ entails the Chapman–Kolmogorov equations
- $$p_{t+s}(x, C) = \int_{\mathbb{R}^d} p_t(y, C) p_s(x, dy) \quad \text{for all } x \in \mathbb{R}^d \text{ and } C \in \mathcal{B}(\mathbb{R}^d).$$
7. Show that the set-functions $p_{t_1, \dots, t_n}^x(C_1 \times \dots \times C_n)$ in Remark 7.6 define measures on $\mathcal{B}(\mathbb{R}^{d \cdot n})$.
8. (a) Let $d(x, A) := \inf_{a \in A} |x - a|$ be the distance between the point $x \in \mathbb{R}^d$ and the set $A \in \mathcal{B}(\mathbb{R}^d)$. Show that $x \mapsto d(x, A)$ is continuous.
 (b) Let u_n be as in (7.7). Show that $u_n \in \mathcal{C}_c(\mathbb{R}^d)$ and $\inf u_n(x) = \mathbb{1}_K$. Draw a picture if $d = 1$ and $K = [0, 1]$.
 (c) Let χ_n be a sequence of type δ , i. e. $\chi_n \in \mathcal{C}_c^\infty(\mathbb{R}^d)$, $\text{supp } \chi_n \subset \mathbb{B}(0, 1/n)$, $\chi_n \geq 0$ and $\int \chi_n(x) dx = 1$. Show that $v_n := \chi_n \star u_n \in \mathcal{C}_c^\infty(\mathbb{R}^d)$ are smooth functions such that $\lim_{n \rightarrow \infty} v_n = \mathbb{1}_K$.
9. Let $A, B \in \mathbb{R}^{d \times d}$ and set $P_t := \exp(tA) := \sum_{j=0}^{\infty} A^j / j!$.
 (a) Show that P_t is a strongly continuous semigroup. Is it contractive?
 (b) Show that $\frac{d}{dt} e^{tA}$ exists and that $\frac{d}{dt} e^{tA} = A e^{tA} = e^{tA} A$.
 (c) Show that $e^{tA} e^{sB} = e^{tB} e^{sA} \forall t \iff AB = BA$.
Hint: Differentiate $e^{tA} e^{sB}$ at $t = 0$ and $s = 0$.
 (d) (Trotter product formula) Show that $e^{A+B} = \lim_{k \rightarrow \infty} (e^{A/k} e^{B/k})^k$.
10. Let $(P_t)_{t \geq 0}$ and $(T_t)_{t \geq 0}$ be two Feller semigroups with generators A and B , respectively.
 (a) Show that $\frac{d}{ds} P_{t-s} T_s = P_{t-s} A T_s - P_{t-s} B T_s$.
 (b) Is it possible that $U_t := T_t P_t$ is again a Feller semigroup?
 (c) (Trotter product formula) Assume that $U_t f := \lim_{n \rightarrow \infty} (T_{t/n} P_{t/n})^n f$ exists strongly and locally uniformly for t . Show that U_t is a Feller semigroup and determine its generator.
11. Complete the following alternative argument for Example 7.20.
 (a) Show that $u_n(x) - u_n(0) - x u'_n(0) = \int_0^x \int_0^y u''_n(z) dz dy \rightarrow \int_0^x \int_0^y 2g(z) dz dy$ uniformly. Conclude that $c' := \lim_{n \rightarrow \infty} u'_n(0)$ exists.
 (b) Show that $u'_n(x) - u'_n(0) \rightarrow \int_0^x g(z) dz$ uniformly. Deduce from this and part a) that $u_n(x)$ converges uniformly to $\int_0^x 2g(z) dt + c'$ and that the constant $c' = \int_{-\infty}^0 g(z) dz$.
 (c) Use a) and b) to show that $u_n(x) - u_n(0) \rightarrow \int_0^x \int_{-\infty}^y 2g(z) dz$ uniformly and argue that $u_n(x)$ has the uniform limit $\int_{-\infty}^x \int_{-\infty}^y 2g(z) dz dy$.

12. Let U_α be the α -potential operator of a BM^d . Give a probabilistic interpretation of $U_\alpha \mathbb{1}_C$ and $\lim_{\alpha \rightarrow 0} U_\alpha \mathbb{1}_C$ if $C \subset \mathbb{R}^d$ is a Borel set.
13. Let $(U_\alpha)_{\alpha > 0}$ be the α -potential operator of a BM^d . Use the resolvent equation to prove the following formulae for $f \in \mathcal{B}_b$ and $x \in \mathbb{R}^d$:

$$\frac{d^n}{d\alpha^n} U_\alpha f(x) = n!(-1)^n U_\alpha^{n+1} f(x)$$

and

$$\frac{d^n}{d\alpha^n} (\alpha U_\alpha) f(x) = n!(-1)^{n+1} (\text{id} - \alpha U_\alpha) f(x).$$

14. Let $(B_t)_{t \geq 0}$ be a BM^d . Then $Uf = \lim_{\alpha \rightarrow 0} U_\alpha f$ is called the zero-potential or potential operator and $Nu := \lim_{\alpha \rightarrow 0} (\alpha \text{id} - A)^{-1} u$ is the zero-resolvent.
- (a) Show that $Uf = Nf = \int_0^\infty P_t f \, dt$.
- (b) Show that $Uf(x) = \int f(y) g(x-y) \, dy$ where $g = \frac{\Gamma(d/2)}{\pi^{d/2}(d-2)} |x|^{2-d}$, $d \geq 3$.
15. Let $(B_t)_{t \geq 0}$ be a BM^1 and consider the two-dimensional process $X_t := (t, B_t)$, $t \geq 0$.
- (a) Show that $(X_t)_{t \geq 0}$ is a Feller process.
- (b) Determine the associated transition semigroup, resolvent and generator.
- (c) State and prove Theorem 7.21 for this process and compare the result with Theorem 5.6
16. Show that Proposition 7.22 follows from Theorem 7.21.
17. Let $t \mapsto X_t$ be a right-continuous stochastic process. Show that for closed sets F

$$\mathbb{P}(X_t \in F \quad \forall t \in \mathbb{R}^+) = \mathbb{P}(X_q \in F \quad \forall q \in \mathbb{Q}^+).$$

Chapter 8

The PDE connection

We want to discuss some relations between partial differential equations (PDEs) and Brownian motion. For many classical PDE problems probability theory yields concrete representation formulae for the solutions in the form of expected values of a Brownian functional. These formulae can be used to get generalized solutions of PDEs (which require less smoothness of the initial/boundary data or the boundary itself) and they are amenable to Monte–Carlo simulations. Purely probabilistic existence proofs for classical PDE problems are, however, rare: Classical solutions require smoothness, which does usually not follow from martingale methods.¹ This explains the role of Proposition 7.3 g) and Proposition 8.10. Let us point out that $\frac{1}{2}\Delta$ has two meanings: On the domain $\mathfrak{D}(\Delta)$, it is the generator of a Brownian motion, but it can also be seen as partial differential operator $L = \sum_{j,k=1}^d \partial_j$ which acts on all \mathcal{C}^2 functions. Of course, on \mathcal{C}_∞^2 both meanings coincide, and it is this observation which makes the method work.

The origin of the probabilistic approach are the pioneering papers by Kakutani [96, 97], Kac [90, 91] and Doob [41]. As an illustration of the method we begin with the elementary (inhomogeneous) heat equation. In this context, classical PDE methods are certainly superior to the clumsy-looking Brownian motion machinery. Its elegance and effectiveness become obvious in the Feynman–Kac formula and the Dirichlet problem which we discuss in Sections 8.3 and 8.4. Moreover, since many second-order differential operators generate diffusion processes, see Chapter 19, only minor changes in our proofs yield similar representation formulae for the corresponding PDE problems.

A martingale relation is the key ingredient. Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM^d . Recall from Theorem 5.6 that for $u \in \mathcal{C}^{1,2}((0, \infty) \times \mathbb{R}^d) \cap \mathcal{C}([0, \infty) \times \mathbb{R}^d)$ satisfying

$$|u(t, x)| + \left| \frac{\partial u(t, x)}{\partial t} \right| + \sum_{j=1}^d \left| \frac{\partial u(t, x)}{\partial x_j} \right| + \sum_{j,k=1}^d \left| \frac{\partial^2 u(t, x)}{\partial x_j \partial x_k} \right| \leq c(t) e^{C|x|} \quad (8.1)$$

for all $t > 0$ and $x \in \mathbb{R}^d$ with some constant $C > 0$ and a locally bounded function $c : (0, \infty) \rightarrow [0, \infty)$, the process

$$M_s^u = u(t-s, B_s) - u(t, B_0) + \int_0^s \left(\frac{\partial}{\partial t} - \frac{1}{2} \Delta_x \right) u(t-r, B_r) dr, \quad s \in [0, t), \quad (8.2)$$

is an \mathcal{F}_s martingale for every measure \mathbb{P}^x , i. e. for every starting point $B_0 = x$ of a Brownian motion.

¹ A notable exception is, of course, Malliavin's calculus, cf. [126, 128]

8.1 The heat equation

In Chapter 7 we have seen that the Laplacian $(\frac{1}{2}\Delta, \mathfrak{D}(\Delta))$ is the infinitesimal generator of a BM^d . In fact, applying Lemma 7.10 to a Brownian motion yields

$$\frac{d}{dt}P_t f(x) = \frac{1}{2}\Delta_x P_t f(x) \quad \text{for all } f \in \mathfrak{D}(\Delta).$$

Setting $u(t, x) := P_t f(x) = \mathbb{E}^x f(B_t)$ this almost proves the following lemma.

8.1 Lemma. *Let $(B_t)_{t \geq 0}$ be a BM^d , $f \in \mathfrak{D}(\Delta)$ and $u(t, x) := \mathbb{E}^x f(B_t)$. Then $u(t, x)$ is the unique bounded solution of the heat equation with initial value f ,*

$$\frac{\partial}{\partial t}u(t, x) - \frac{1}{2}\Delta_x u(t, x) = 0, \tag{8.3a}$$

$$u(0, x) = f(x). \tag{8.3b}$$

Proof. It remains to show uniqueness. Let $u(t, x)$ be a bounded solution of (8.3). Its Laplace transform is given by

$$v_\lambda(x) = \int_0^\infty u(t, x) e^{-\lambda t} dt.$$

Since we may differentiate under the integral sign, we get

$$\begin{aligned} \lambda v_\lambda(x) - \frac{1}{2}\Delta_x v_\lambda(x) &= \lambda v_\lambda(x) - \int_0^\infty \frac{1}{2}\Delta_x u(t, x) e^{-\lambda t} dt \\ &= \lambda v_\lambda(x) - \int_0^\infty \frac{\partial}{\partial t} u(t, x) e^{-\lambda t} dt. \end{aligned}$$

Integration by parts and (8.3b) yield

$$\left(\lambda \text{id} - \frac{1}{2}\Delta_x \right) v_\lambda(x) = \lambda v_\lambda(x) + \int_0^\infty u(t, x) \frac{\partial}{\partial t} e^{-\lambda t} dt - u(t, x) e^{-\lambda t} \Big|_{t=0}^{t=\infty} = f(x).$$

Since the equation $(\lambda \text{id} - \frac{1}{2}\Delta_x)v_\lambda = f$, $f \in \mathfrak{D}(\Delta)$, $\lambda > 0$, has the unique solution $v_\lambda = U_\lambda f$ where U_λ is the resolvent operator, and since the Laplace transform is unique, cf. [170, Proposition 1.2], we conclude that $u(t, x)$ is unique. \square

Ex. 8.1 Lemma 8.1 is a PDE proof. Let us sketch a further, more probabilistic approach, which will be important if we consider more complicated PDEs. It is, admittedly, too complicated for the simple setting considered in Lemma 8.1, but it allows us to remove the restriction that $f \in \mathfrak{D}(\Delta)$, see the discussion in Remark 8.4 below.

8.2 Probabilistic proof of Lemma 8.1. *Uniqueness:* Assume that $u(t, x)$ satisfies (8.1) and $|u(t, x)| \leq c e^{c|x|}$ on $[0, T] \times \mathbb{R}^d$ where $c = c_T$. By (8.2),

$$M_s^u := u(t-s, B_s) - u(t, B_0) + \underbrace{\int_0^s \left(\frac{\partial}{\partial t} - \frac{1}{2} \Delta_x \right) u(t-r, B_r) dr}_{=0 \text{ by (8.3a)}}$$

is a martingale w. r. t. $(\mathcal{F}_s)_{s \in [0, t]}$, and so is $N_s^u := u(t-s, B_s)$, $s \in [0, t]$. By assumption, u is exponentially bounded. Therefore, $(N_s^u)_{s \in [0, t]}$ is dominated by the integrable function $c \sup_{0 \leq s \leq t} \exp(c |B_s|)$, hence uniformly integrable. By the martingale convergence theorem, Theorem A.6, the limit $\lim_{s \rightarrow t} N_s^u$ exists a. s. and in L^1 . Thus,

$$u(t, x) = \mathbb{E}^x u(t, B_0) = \mathbb{E}^x N_0^u = \lim_{s \rightarrow t} \mathbb{E}^x N_s^u = \mathbb{E}^x N_t^u \stackrel{(8.3b)}{=} \mathbb{E}^x f(B_t),$$

which shows that the initial condition f uniquely determines the solution.

Existence: Set $u(t, x) := \mathbb{E}^x f(B_t) = P_t f(x)$ and assume that $f \in \mathfrak{D}(\Delta)$. Lemma 7.10 together with an obvious modification of the proof of Proposition 7.3 g) show that u is in $\mathcal{C}^{1,2}((0, \infty) \times \mathbb{R}^d) \cap \mathcal{C}([0, \infty) \times \mathbb{R}^d)$ and satisfies (8.1). Therefore, (8.2) tells us that

$$M_s^u := u(t-s, B_s) - u(t, B_0) + \int_0^s \left(\frac{\partial}{\partial t} - \frac{1}{2} \Delta_x \right) u(t-r, B_r) dr$$

is a martingale. On the other hand, the Markov property reveals that

$$u(t-s, B_s) = \mathbb{E}^{B_s} f(B_{t-s}) = \mathbb{E}^x (f(B_t) | \mathcal{F}_s), \quad 0 \leq s \leq t,$$

is a martingale, and so is $\int_0^s \left(\frac{\partial}{\partial t} - \frac{1}{2} \Delta_x \right) u(t-r, B_r) dr$, $s \leq t$. Since it has continuous paths of bounded variation, it is identically zero on $[0, t]$, cf. Proposition A.22, and (8.3a) follows. \square

Ex. 8.2

The proof of Proposition 7.4 g) shows that the map $(t, x) \mapsto P_t f(x) = \mathbb{E}^x f(B_t)$, $t > 0, x \in \mathbb{R}^d$, is smooth. This allows us to extend the probabilistic approach using a localization technique described in Remark 8.4 below.

8.3 Theorem. Let $(B_t)_{t \geq 0}$ be a BM^d and $f \in \mathcal{B}(\mathbb{R}^d)$ such that $|f(x)| \leq C e^{C|x|}$. Then $u(t, x) := \mathbb{E}^x f(B_t)$ is the unique solution to the heat equation (8.3) satisfying the growth estimate $|u(x, t)| \leq C e^{C|x|}$.

The growth estimate in the statement of Theorem 8.3 is essential for the uniqueness, cf. John [89, Chapter 7.1].

8.4 Remark (Localization). Let $(B_t)_{t \geq 0}$ be a d -dimensional Brownian motion, $u \in \mathcal{C}^{1,2}((0, \infty) \times \mathbb{R}^d) \cap \mathcal{C}([0, \infty) \times \mathbb{R}^d)$, fix $t > 0$, and consider the following sequence of stopping times

$$\tau_n := \inf \{s > 0 : |B_s| \geq n\} \wedge (t - n^{-1}), \quad n \geq 1/t. \quad (8.4)$$

For every $n \geq 1/t$ we pick some cut-off function $\chi \in \mathcal{C}^{1,2}([0, \infty) \times \mathbb{R}^d)$ with compact support, $\chi_n|_{[0, 1/2n] \times \mathbb{R}^d} \equiv 0$ and $\chi_n|_{[1/n, n] \times \mathbb{B}(0, n)} \equiv 1$. Clearly, for M_s^u as in (8.2),

$$M_{s \wedge \tau_n}^u = M_{s \wedge \tau_n}^{\chi_n u}, \quad \text{for all } s \in [0, t) \quad \text{and} \quad n \geq 1/t.$$

Since $\chi_n u \in \mathcal{C}_b^{1,2}$ satisfies (8.1), $(M_s^{\chi_n u}, \mathcal{F}_s)_{s \in [0, t]}$ is, for each $n \geq 1/t$, a bounded martingale. By optional stopping, Remark A.21,

$$(M_{s \wedge \tau_n}^u, \mathcal{F}_s)_{s \in [0, t]} \quad \text{is a bounded martingale for all } u \in \mathcal{C}^{1,2} \text{ and } n \geq 1/t. \quad (8.5)$$

Assume that $u \in \mathcal{C}^{1,2}((0, \infty) \times \mathbb{R}^d)$. Replacing in the probabilistic proof 8.2 B_s , M_s^u and N_s^u by the stopped versions $B_{s \wedge \tau_n}$, $M_{s \wedge \tau_n}^u$ and $N_{s \wedge \tau_n}^u$, we see that²

$$M_{s \wedge \tau_n}^u = u(t - s \wedge \tau_n, B_{s \wedge \tau_n}) - u(t, B_0) + \underbrace{\int_0^{s \wedge \tau_n} \left(\frac{\partial}{\partial t} - \frac{1}{2} \Delta_x \right) u(t - r, B_r) dr}_{=0 \text{ by (8.3a)}}$$

and $N_{s \wedge \tau_n}^u = u(t - s \wedge \tau_n, B_{s \wedge \tau_n})$, $s \in [0, t]$, are for every $n \geq 1/t$ bounded martingales. In particular, $\mathbb{E}^x N_0^u = \mathbb{E}^x N_{s \wedge \tau_n}^u$ for all $s \in [0, t]$, and by the martingale convergence theorem we get

$$u(t, x) = \mathbb{E}^x u(t, B_0) = \mathbb{E}^x N_0^u = \lim_{s \rightarrow t} \mathbb{E}^x N_{s \wedge \tau_n}^u = \mathbb{E}^x N_{t \wedge \tau_n}^u.$$

As $\lim_{n \rightarrow \infty} \tau_n = t$ a. s., we see

$$u(t, x) = \mathbb{E}^x N_{t \wedge \tau_n}^u = \lim_{k \rightarrow \infty} \mathbb{E}^x u(t - t \wedge \tau_k, B_{t \wedge \tau_k}) = \mathbb{E}^x u(0, B_t) \stackrel{(8.3b)}{=} \mathbb{E}^x f(B_t),$$

which shows that the solution is unique.

To see that $u(t, x) := P_t f(x)$ is a solution, we assume that $f(x) \leq C e^{C|x|}$. We can adapt the proof of Proposition 7.3 g) to see that $u \in \mathcal{C}^{1,2}((0, \infty) \times \mathbb{R}^d)$. Therefore, (8.5) shows that $(M_{s \wedge \tau_n}^u, \mathcal{F}_s)_{s \in [0, t]}$ is, for every $n \geq 1/t$, a bounded martingale. By the Markov property

$$u(t - s, B_s) = \mathbb{E}^{B_s} f(B_{t-s}) = \mathbb{E}^x (f(B_t) | \mathcal{F}_s), \quad 0 \leq s \leq t,$$

and by optional stopping $(u(t - s \wedge \tau_n, B_{s \wedge \tau_n}), \mathcal{F}_s)_{s \in [0, t]}$ is for every $n \geq 1/t$ a martingale. This implies that $\int_0^{s \wedge \tau_n} \left(\frac{\partial}{\partial t} - \frac{1}{2} \Delta_x \right) u(t - r, B_r) dr$ is a martingale with

² Note that $\int_0^{s \wedge \tau_n} \dots dr = \int_{[0, s \wedge \tau_n)} \dots dr$

continuous paths of bounded variation. Hence, it is identically zero on $[0, t \wedge \tau_n)$, cf. Ex. 8.2 Proposition A.22. As $\lim_{n \rightarrow \infty} \tau_n = t$, (8.3a) follows.

8.2 The inhomogeneous initial value problem

If we replace the right-hand side of the heat equation (8.3a) with some function $g(t, x)$, we get the following *inhomogeneous initial value problem*.

$$\frac{\partial}{\partial t} v(t, x) - \frac{1}{2} \Delta_x v(t, x) = g(t, x), \quad (8.6a)$$

$$v(0, x) = f(x). \quad (8.6b)$$

Assume that u solves the corresponding homogeneous problem (8.3). If w solves the inhomogeneous problem (8.6) with zero initial condition $f = 0$, then $v = u + w$ is a solution of (8.6). Therefore, it is customary to consider

$$\frac{\partial}{\partial t} v(t, x) - \frac{1}{2} \Delta_x v(t, x) = g(t, x), \quad (8.7a)$$

$$v(0, x) = 0 \quad (8.7b)$$

instead of (8.6). Let us use the probabilistic approach of Section 8.1 to develop an educated guess about the solution to (8.7).

Let $v(t, x)$ be a solution to (8.7) which is bounded and sufficiently smooth ($\mathcal{C}^{1,2}$ and (8.1) will do). Then, for $s \in [0, t)$,

$$M_s^v := v(t-s, B_s) - v(t, B_0) + \underbrace{\int_0^s \left(\frac{\partial}{\partial t} v(t-r, B_r) - \frac{1}{2} \Delta_x v(t-r, B_r) \right) dr}_{=g(t-r, B_r) \text{ by (8.7a)}}$$

is a martingale, and so is $N_s^v := v(t-s, B_s) + \int_0^s g(t-r, B_r) dr, s \in [0, t)$; in particular, $\mathbb{E}^x N_0^v = \mathbb{E}^x N_s^v$ for all $s \in [0, t)$. If we can use the martingale convergence theorem, e. g. if the martingale is uniformly integrable, we can let $s \rightarrow t$ and get

$$v(t, x) = \mathbb{E}^x N_0^v = \lim_{s \rightarrow t} \mathbb{E}^x N_s^v = \mathbb{E}^x N_t^v \stackrel{(8.7b)}{=} \mathbb{E}^x \left(\int_0^t g(t-r, B_r) dr \right).$$

This consideration already yields uniqueness. For the existence we have to make sure that we can perform all manipulations in the above calculation. With some effort we can justify everything if $g(t, x)$ is bounded or if $g = g(x)$ is from the so-called Kato class. For the sake of clarity, we content ourselves with the following simple version.

Ex. 8.4 8.5 Theorem. Let $(B_t)_{t \geq 0}$ be a BM^d and $g \in \mathcal{D}(\Delta)$. Then the unique solution of the inhomogeneous initial value problem (8.7) satisfying $|v(t, x)| \leq C t$ is given by $v(t, x) = \mathbb{E}^x \int_0^t g(B_s) ds$.

Proof. Since $v(t, x) = \int_0^t \mathbb{E}^x g(B_s) ds = \int_0^t P_s g(x) ds$, Lemma 7.10 and a variation of the proof of Proposition 7.3 g) imply that $v \in \mathcal{C}_b^{1,2}((0, \infty) \times \mathbb{R}^d)$. This means that the heuristic derivation of $v(t, x)$ furnishes a uniqueness proof.

To see that $v(t, x)$ is indeed a solution, we use (8.2) and conclude that

$$\begin{aligned} M_s^v &= v(t-s, B_s) - v(t, B_0) + \int_0^s \left(\frac{\partial}{\partial t} - \frac{1}{2} \Delta_x \right) v(t-r, B_r) dr \\ &= v(t-s, B_s) + \int_0^s g(B_r) dr - v(t, B_0) \\ &\quad + \int_0^s \left[\left(\frac{\partial}{\partial t} - \frac{1}{2} \Delta_x \right) v(t-r, B_r) - g(B_r) \right] dr \end{aligned}$$

is a martingale. On the other hand, the Markov property gives for all $s \leq t$,

$$\begin{aligned} \mathbb{E}^x \left(\int_0^t g(B_r) dr \mid \mathcal{F}_s \right) &= \int_0^s g(B_r) dr + \mathbb{E}^x \left(\int_s^t g(B_r) dr \mid \mathcal{F}_s \right) \\ &= \int_0^s g(B_r) dr + \mathbb{E}^x \left(\int_0^{t-s} g(B_{r+s}) dr \mid \mathcal{F}_s \right) \\ &= \int_0^s g(B_r) dr + \mathbb{E}^{B_s} \left(\int_0^{t-s} g(B_r) dr \right) \\ &= \int_0^s g(B_r) dr + v(t-s, B_s). \end{aligned}$$

This shows that the right-hand side is for $s \leq t$ a martingale.

Consequently, $\int_0^s \left(\frac{\partial}{\partial t} v(t-r, B_r) - \frac{1}{2} \Delta_x v(t-r, B_r) - g(B_r) \right) dr$ is a martingale

Ex. 8.2 with continuous paths of bounded variation; by Proposition A.22 it is zero on $[0, t)$, and (8.7a) follows. \square

If $g = g(x)$ does not depend on time, we can use the localization technique described in Remark 8.4 to show that Theorem 8.5 remains valid if $g \in \mathcal{B}_b(\mathbb{R}^d)$. This requires a modification of the proof of Proposition 7.3 g) to see that the function $v(t, x) := \mathbb{E}^x \left(\int_0^t g(B_r) dr \right)$ is in $\mathcal{C}^{1,2}((0, \infty) \times \mathbb{R}^d)$. Then the proof is similar to the one for the heat equation described in Remark 8.4. As soon as g depends on time, $g = g(t, x)$, things become messy, cf. [49, Chapter 8.2].

8.3 The Feynman–Kac formula

We have seen in Section 8.2 that the inhomogeneous initial value problem (8.7) is solved by the expectation of the random variable $G_t = \int_0^t g(B_r) dr$. It is an interesting question whether we can find out more about the probability distributions of G_t and (B_t, G_t) . A natural approach would be to compute the joint characteristic function $\mathbb{E}^x[\exp(i\xi G_t + i\langle \eta, B_t \rangle)]$. Let us consider the more general problem to calculate

$$w(t, x) := \mathbb{E}^x(f(B_t)e^{C_t}) \quad \text{where} \quad C_t := \int_0^t c(B_r) dr$$

for some functions $f, c : \mathbb{R}^d \rightarrow \mathbb{C}$. Setting $f(x) = e^{i\langle \eta, x \rangle}$ and $c(x) = i\xi g(x)$ brings us back to the original question. We will show that, under suitable conditions on f and c , the function $w(t, x)$ is the unique solution to the following initial value problem

$$\frac{\partial}{\partial t} w(t, x) - \frac{1}{2} \Delta_x w(t, x) - c(x)w(t, x) = 0, \quad (8.8a)$$

$$w(0, x) \text{ is continuous and } w(0, x) = f(x). \quad (8.8b)$$

Considering real and imaginary parts separately, we may assume that f and c are real-valued.

Let us follow the same strategy as in Sections 8.1 and 8.2. Assume that $w \in \mathcal{C}^{1,2}$ is a solution to (8.8) which satisfies (8.1). If $c(x)$ is bounded, we can apply (8.2) to $w(t-s, x) \exp(C_s)$ for $s \in [0, t]$. Since $\frac{d}{dt} \exp(C_t) = c(B_t) \exp(C_t)$, we see that for $s \in [0, t]$

$$M_s^w := w(t-s, B_s) e^{C_s} - w(t, B_0) + \underbrace{\int_0^s \left(\frac{\partial}{\partial t} - \frac{1}{2} \Delta_x - c(B_r) \right) w(t-r, B_r) e^{C_r} dr}_{=0 \text{ by (8.8a)}}$$

is a uniformly integrable martingale with respect to the filtration $(\mathcal{F}_s)_{s \in [0, t]}$. Thus $N_s^w := (w(t-s, B_s) \exp(C_s), \mathcal{F}_s)_{s \in [0, t]}$ is a uniformly integrable martingale. Since $\mathbb{E}^x N_0^w = \mathbb{E}^x N_s^w$ for all $s \in [0, t]$, the martingale convergence theorem, Theorem A.7, gives

$$w(t, x) = \mathbb{E}^x w(t, B_0) = \mathbb{E}^x N_0^w = \lim_{s \rightarrow t} \mathbb{E}^x N_s^w = \mathbb{E}^x N_t^w \stackrel{(8.8b)}{=} \mathbb{E}^x(f(B_t)e^{C_t}).$$

This reveals both the structure of the solution and its uniqueness. To show with martingale methods that $w(t, x) := \mathbb{E}^x(f(B_t)e^{C_t})$ solves (8.8) requires *a priori* knowledge that w is smooth, e. g. $w \in \mathcal{C}^{1,2}$ – but this is quite hard to check, even if $f \in \mathcal{D}(\Delta)$ and $c \in \mathcal{C}_b(\mathbb{R}^d)$.

Therefore, we use semigroup methods to show the next theorem.

8.6 Theorem (Kac 1949). *Let $(B_t)_{t \geq 0}$ be a BM^d , $f \in \mathcal{D}(\Delta)$ and $c \in \mathcal{C}_b(\mathbb{R}^d)$. The unique solution to the initial value problem (8.8) is given by*

$$w(t, x) = \mathbb{E}^x \left[f(B_t) \exp \left(\int_0^t c(B_r) dr \right) \right]. \quad (8.9)$$

This solution is bounded by $e^{\alpha t}$ where $\alpha = \sup_{x \in \mathbb{R}^d} c(x)$.

The representation formula (8.9) is called *Feynman–Kac formula*.

Proof of Theorem 8.6. Denote by P_t the Brownian semigroup. Set $C_t := \int_0^t c(B_r) dr$ and $T_t u(x) := \mathbb{E}^x(u(B_t)e^{C_t})$. If we can show that $(T_t)_{t \geq 0}$ is a Feller semigroup with generator $Au = \frac{1}{2}\Delta u + cu$, we are done: Existence follows from Lemma 7.10, uniqueness from the fact that A has a resolvent which is uniquely determined by T_t , cf. Proposition 7.13.

Since

$$|T_t u(x)| \leq e^{t \sup_{x \in \mathbb{R}^d} c(x)} \mathbb{E}^x |u(B_t)| \leq e^{t\alpha} \|u\|_\infty,$$

the operator T_t is, in general, only a contraction if $\alpha \leq 0$. Let us assume this for the time being and show that $(T_t)_{t \geq 0}$ is a Feller semigroup.

a) *Positivity* is clear from the very definition.

b) *Feller property*: For all $u \in \mathcal{C}_\infty(\mathbb{R}^d)$ we have

$$\begin{aligned} T_t u(x) &= \mathbb{E}^x \left[u(B_t) \exp \left(\int_0^t c(B_r) dr \right) \right] \\ &= \mathbb{E} \left[u(B_t + x) \exp \left(\int_0^t c(B_r + x) dr \right) \right]. \end{aligned}$$

A straightforward application of the dominated convergence theorem yields $\lim_{y \rightarrow x} T_t u(y) = T_t u(x)$ and $\lim_{|x| \rightarrow \infty} T_t u(x) = 0$. Thus, $T_t u \in \mathcal{C}_\infty(\mathbb{R}^d)$.

c) *Strong continuity*: Let $u \in \mathcal{C}_\infty(\mathbb{R}^d)$ and $t \geq 0$. Using the elementary inequality $1 - e^c \leq |c|$ for all $c \leq 0$, we get

$$\begin{aligned} |T_t u(x) - u(x)| &\leq |\mathbb{E}^x (u(B_t)(e^{C_t} - 1))| + |\mathbb{E}^x (u(B_t) - u(x))| \\ &\leq \|u\|_\infty \mathbb{E}^x |C_t| + \|P_t u - u\|_\infty \\ &\leq t \|u\|_\infty \|c\|_\infty + \|P_t u - u\|_\infty. \end{aligned}$$

Since the Brownian semigroup $(P_t)_{t \geq 0}$ is strongly continuous on $\mathcal{C}_\infty(\mathbb{R}^d)$, the strong continuity of $(T_t)_{t \geq 0}$ follows.

d) *Semigroup property*: Let $s, t \geq 0$ and $u \in \mathcal{C}_\infty(\mathbb{R}^d)$. Using the tower property of conditional expectation and the Markov property, we see

$$\begin{aligned}
 T_{t+s}u(x) &= \mathbb{E}^x \left[u(B_{t+s}) e^{\int_0^{t+s} c(B_r) dr} \right] \\
 &= \mathbb{E}^x \left[\mathbb{E}^x \left(u(B_{t+s}) e^{\int_s^{t+s} c(B_r) dr} e^{\int_0^s c(B_r) dr} \mid \mathcal{F}_s \right) \right] \\
 &= \mathbb{E}^x \left[\mathbb{E}^x \left(u(B_{t+s}) e^{\int_0^t c(B_{r+s}) dr} \mid \mathcal{F}_s \right) e^{\int_0^s c(B_r) dr} \right] \\
 &= \mathbb{E}^x \left[\mathbb{E}^{B_s} \left(u(B_t) e^{\int_0^t c(B_r) dr} \right) e^{\int_0^s c(B_r) dr} \right] \\
 &= \mathbb{E}^x \left[T_t f(B_s) e^{\int_0^s c(B_r) dr} \right] \\
 &= T_s T_t f(x).
 \end{aligned}$$

Because of Theorem 7.19, we may calculate the generator A using pointwise convergence: For all $u \in \mathcal{D}(\Delta)$ and $x \in \mathbb{R}^d$

$$\lim_{t \rightarrow 0} \frac{T_t u(x) - u(x)}{t} = \lim_{t \rightarrow 0} \frac{T_t u(x) - P_t u(x)}{t} + \lim_{t \rightarrow 0} \frac{P_t u(x) - u(x)}{t}.$$

The second limit equals $\frac{1}{2} \Delta u(x)$; since $(e^{C_t} - 1)/t \rightarrow c(B_0)$ boundedly – cf. the estimate in the proof of the strong continuity c) – we can use dominated convergence and find

$$\lim_{t \rightarrow 0} \mathbb{E}^x \left(u(B_t) \frac{e^{C_t} - 1}{t} \right) = \mathbb{E}^x (u(B_0) c(B_0)) = u(x) c(x).$$

This completes the proof in the case $\alpha \leq 0$.

If $\alpha > 0$ we can consider the following auxiliary problem: Replace in (8.8) the function $c(x)$ by $c(x) - \alpha$. From the first part we know that $e^{-\alpha t} T_t f = \mathbb{E}^x (f(B_t) e^{C_t - \alpha t})$ is the unique solution of the auxiliary problem

Ex. 8.10

$$\left(\frac{\partial}{\partial t} - \frac{1}{2} \Delta - c(\cdot) + \alpha \right) e^{-\alpha t} T_t f = 0.$$

On the other hand,

$$\frac{d}{dt} e^{-\alpha t} T_t f = -\alpha e^{-\alpha t} T_t f + e^{-\alpha t} \frac{d}{dt} T_t f,$$

and this shows that $T_t f$ is the unique solution of the original problem. \square

8.7 An Application: Lévy's arc-sine law (Lévy 1939, Kac 1949). Denote by $(B_t)_{t \geq 0}$ a BM¹ and by $G_t := \int_0^t \mathbb{1}_{(0, \infty)}(B_s) ds$ the total amount of time a Brownian motion spends in the positive half-axis up to time t . Lévy showed in [119] that G_t has an

arc-sine distribution:

$$\mathbb{P}^0 \left(\int_0^t \mathbb{1}_{(0,\infty)}(B_s) ds \leq x \right) = \frac{2}{\pi} \arcsin \sqrt{\frac{x}{t}}, \quad 0 \leq x \leq t. \quad (8.10)$$

We calculate the Laplace transform $g_\alpha(t) = \mathbb{E}^0 e^{-\alpha G_t}$, $\alpha \geq 0$, to identify the probability distribution. Define for $\lambda > 0$ and $x \in \mathbb{R}$

$$v_\lambda(x) := \int_0^\infty e^{-\lambda t} \mathbb{E}^x e^{-\alpha G_t} dt = \int_0^\infty e^{-\lambda t} \mathbb{E}^x \left(\mathbb{1}_{\mathbb{R}}(B_t) \cdot e^{-\alpha \int_0^t \mathbb{1}_{(0,\infty)}(B_s) ds} \right) dt.$$

Assume for a moment that we *could* apply Theorem 8.6 with $f(x) = \mathbb{1}_{\mathbb{R}}(x)$ and $c(x) = -\alpha \mathbb{1}_{(0,\infty)}(x)$. Then $\lambda \mapsto v_\lambda(x)$ is the Laplace transform of the function $w(t, x) = \mathbb{E}^x(\mathbb{1}_{\mathbb{R}}(B_t) \cdot e^{-\alpha \int_0^t \mathbb{1}_{(0,\infty)}(B_s) ds})$ which, by Theorem 8.6, is the solution of the initial value problem (8.8). A short calculation shows that the Laplace transform satisfies the following linear ODE:

$$\frac{1}{2} v_\lambda''(x) - \alpha \mathbb{1}_{(0,\infty)}(x) v_\lambda(x) = \lambda v_\lambda(x) - \mathbb{1}_{\mathbb{R}}(x) \quad \text{for all } x \in \mathbb{R} \setminus \{0\}. \quad (8.11)$$

This ODE is, with suitable initial conditions, uniquely solvable in $(-\infty, 0)$ and $(0, \infty)$. It will become clear from the approximation below that v_λ is of class $\mathcal{C}^1 \cap \mathcal{C}_b$ on the whole real line. Therefore we can prescribe conditions in $x = 0$ such that v_λ is the unique bounded \mathcal{C}^1 solution of (8.11). With the usual *Ansatz* for linear second-order ODEs, we get

$$v_\lambda(x) = \begin{cases} \frac{1}{\lambda} + C_1 e^{\sqrt{2\lambda}x} + C_2 e^{-\sqrt{2\lambda}x}, & x < 0, \\ \frac{1}{\alpha + \lambda} + C_3 e^{\sqrt{2(\alpha+\lambda)}x} + C_4 e^{-\sqrt{2(\alpha+\lambda)}x}, & x > 0. \end{cases}$$

Let us determine the constants C_j , $j = 1, 2, 3, 4$. The boundedness of the function v_λ gives $C_2 = C_3 = 0$. Since v_λ and v'_λ are continuous at $x = 0$, we get

$$\begin{aligned} v_\lambda(0+) &= \frac{1}{\alpha + \lambda} + C_4 = \frac{1}{\lambda} + C_1 = v_\lambda(0-), \\ v'_\lambda(0+) &= -\sqrt{2(\alpha + \lambda)} C_4 = \sqrt{2\lambda} C_1 = v'_\lambda(0-). \end{aligned}$$

Solving for C_1 gives $C_1 = -\lambda^{-1} + (\lambda(\lambda + \alpha))^{-1/2}$, and so

$$v_\lambda(0) = \int_0^\infty e^{-\lambda t} \mathbb{E}^0(e^{-\alpha G_t}) dt = \frac{1}{\sqrt{\lambda(\lambda + \alpha)}}.$$

On the other hand, by Fubini's theorem and the formula $\int_0^\infty \frac{1}{\sqrt{\pi t}} e^{-ct} dt = \frac{1}{\sqrt{c}}$:

$$\begin{aligned} \int_0^\infty e^{-\lambda t} \frac{2}{\pi} \int_0^{\pi/2} e^{-\alpha t \sin^2 \theta} d\theta dt &= \int_0^\infty e^{-\lambda t} \int_0^t \frac{e^{-\alpha r}}{\pi \sqrt{r(t-r)}} dr dt \\ &= \int_0^\infty \frac{e^{-\alpha r}}{\sqrt{\pi r}} \int_r^\infty \frac{e^{-\lambda t}}{\sqrt{\pi r(t-r)}} dt dr \\ &= \int_0^\infty \frac{e^{-(\alpha+\lambda)r}}{\sqrt{\pi r}} dr \int_0^\infty \frac{e^{-\lambda s}}{\sqrt{\pi s}} ds \\ &= \frac{1}{\sqrt{(\lambda + \alpha)\lambda}} \end{aligned}$$

Because of the uniqueness of the Laplace transform, cf. [170, Proposition 1.2], we conclude

$$g_\alpha(t) = \mathbb{E}^0 e^{-\alpha G_t} = \frac{2}{\pi} \int_0^{\pi/2} e^{-\alpha t \sin^2 \theta} d\theta = \frac{2}{\pi} \int_0^t e^{-\alpha x} d_x \arcsin \sqrt{\frac{x}{t}},$$

and, again by the uniqueness of the Laplace transform, we get (8.10).

Let us now justify the formal manipulations that led to (8.11). In order to apply Theorem 8.6, we cut and smooth the functions $f(x) = \mathbb{1}_{\mathbb{R}}(x)$ and $g(x) = \mathbb{1}_{(0,\infty)}(x)$. Pick a sequence of cut-off functions $\chi_n \in C_c(\mathbb{R})$ such that $\mathbb{1}_{[-n,n]} \leq \chi_n \leq 1$, and define $g_n(x) = 0 \vee (nx) \wedge 1$. Then $\chi_n \in \mathfrak{D}(\Delta)$, $g_n \in \mathcal{C}_b(\mathbb{R})$, $\sup_{n \geq 1} \chi_n = \mathbb{1}_{\mathbb{R}}$ and $\sup_{n \geq 0} g_n = g$. By Theorem 8.6,

$$w_n(t, x) := \mathbb{E}^x \left(\chi_n(B_t) e^{-\alpha \int_0^t g_n(B_s) ds} \right) dt, \quad t > 0, x \in \mathbb{R},$$

is, for each $n \geq 1$, the unique bounded solution of the initial value problem (8.8) with $f = f_n$ and $c = g_n$; note that the bound from Theorem 8.6 does not depend on $n \geq 1$. The Laplace transform $\lambda \mapsto v_{n,\lambda}(x)$ of $w_n(t, x)$ is the unique bounded solution of the ODE

$$\frac{1}{2} v_{n,\lambda}''(x) - \alpha \chi_n(x) v_{n,\lambda}(x) = \lambda v_{n,\lambda}(x) - g_n(x), \quad x \in \mathbb{R}. \quad (8.12)$$

Integrating two times, one can show that $\lim_{n \rightarrow \infty} v_{n,\lambda}(x) = v_\lambda(x)$, $\lim_{n \rightarrow \infty} v_{n,\lambda}'(x) = v_\lambda'(x)$ for all $x \in \mathbb{R}$ and $\lim_{n \rightarrow \infty} v_{n,\lambda}''(x) = v_\lambda''(x)$ for all $x \neq 0$; moreover, we have $v_\lambda \in \mathcal{C}^1(\mathbb{R})$, cf. Problem 8.3. Ex. 8.3

8.4 The Dirichlet problem

Up to now we have considered parabolic initial value problems. In this section we focus on the most basic elliptic boundary value problem, Dirichlet's problem for the Laplace operator with zero boundary data. Throughout this section $D \subset \mathbb{R}^d$ is a bounded open

and connected domain, and we write $\partial D = \overline{D} \setminus D$ for its boundary; $(B_t)_{t \geq 0}$ is a BM^d starting from $B_0 = x$, and $\tau = \tau_{D^c} = \inf\{t > 0 : B_t \notin D\}$ is the first exit time from the set D . We assume that the $(B_t)_{t \geq 0}$ is adapted to a right-continuous filtration $(\mathcal{F}_t)_{t \geq 0}$.³ We are interested in the following question: Find some function u on \overline{D} with

Ex. 8.5 the following properties

$$\Delta u(x) = 0 \quad \text{for } x \in D \quad (8.13a)$$

$$u(x) = f(x) \quad \text{for } x \in \partial D \quad (8.13b)$$

$$u \text{ is continuous in } \overline{D}. \quad (8.13c)$$

To get a feeling for the problem, we use similar heuristic arguments as in Sections 8.1–8.3. Assume that u is a solution of the Dirichlet problem (8.13). If ∂D is smooth, we can extend u onto \mathbb{R}^d in such a way that $u \in \mathcal{C}^2(\mathbb{R}^d)$ with $\text{supp } u$ compact. Therefore, we can apply (8.2) with $u = u(x)$ to see that

$$M_t^u = u(B_t) - u(B_0) - \int_0^t \frac{1}{2} \Delta u(B_r) dr$$

is a bounded martingale. By optional stopping, $(M_{t \wedge \tau}^u, \mathcal{F}_t)_{t \geq 0}$ is a martingale, and we see that⁴

$$M_{t \wedge \tau}^u = u(B_{t \wedge \tau}) - u(B_0) - \underbrace{\int_0^{t \wedge \tau} \frac{1}{2} \Delta u(B_r) dr}_{=0 \text{ by (8.13a)}} = u(B_{t \wedge \tau}) - u(B_0).$$

Therefore, $N_{t \wedge \tau}^u := u(B_{t \wedge \tau})$ is a bounded martingale. In particular, $\mathbb{E}^x N_0^u = \mathbb{E}^x N_{t \wedge \tau}^u$ for all $t \geq 0$. If $\tau = \tau_{D^c} < \infty$, the martingale convergence theorem yields

$$u(x) = \mathbb{E}^x N_0^u = \lim_{t \rightarrow \infty} \mathbb{E}^x N_{t \wedge \tau}^u = \mathbb{E}^x N_\tau^u = \mathbb{E}^x u(B_\tau) \stackrel{(8.13b)}{=} \mathbb{E}^x f(B_\tau).$$

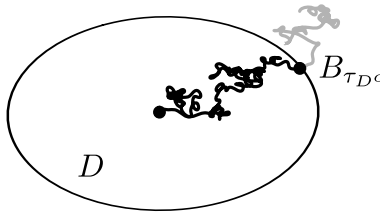


Figure 8.1. Exit time and position of a Brownian motion from the set D .

³ A condition for this can be found in Theorem 6.21.

⁴ Note that $\int_0^{t \wedge \tau} \dots dr = \int_{[0, t \wedge \tau)} \dots dr$

This shows that $\mathbb{E}^x f(B_\tau)$ is a good candidate for the solution of (8.13). But before we go on, let us quickly check that τ is indeed a.s. finite. The following result is similar to Lemma 7.24.

8.8 Lemma. *Let $(B_t)_{t \geq 0}$ be a BM^d and $D \subset \mathbb{R}^d$ be an open and bounded set. Then $\mathbb{E}^x \tau_{D^c} < \infty$ where τ_{D^c} is the first hitting time of D^c .* Ex. 8.6

Proof. Because of Lemma 5.8, $\tau = \tau_{D^c}$ is a stopping time. Since D is bounded, there is some $r > 0$ such that $D \subset \mathbb{B}(0, r)$. Define $u(x) := -\exp(-|x|^2/4r^2)$. Then $u \in \mathcal{C}_\infty^2(\mathbb{R}^d) \subset \mathfrak{D}(\Delta)$ and, for all $|x| < r$,

$$\frac{1}{2} \Delta u(x) = \frac{1}{2} \left(\frac{d}{2r^2} - \frac{|x|^2}{4r^4} \right) e^{-|x|^2/4r^2} \geq \frac{1}{2} \left(\frac{d}{2r^2} - \frac{1}{4r^2} \right) e^{-1/4} =: \kappa > 0.$$

Therefore, we can use Dynkin's formula (7.21) with $\sigma = \tau \wedge n$ to get

$$\kappa \mathbb{E}^x(\tau \wedge n) \leq \mathbb{E}^x \left(\int_0^{\tau \wedge n} \frac{1}{2} \Delta u(B_r) dr \right) = \mathbb{E}^x u(B_{\tau \wedge n}) - u(x) \leq 2\|u\|_\infty = 2.$$

By Fatou's lemma, $\mathbb{E}^x \tau \leq \lim_{n \rightarrow \infty} \mathbb{E}^x(\tau \wedge n) \leq 2/\kappa < \infty$. □ Ex. 8.7

Let us now show that $u(x) = \mathbb{E}^x f(B_\tau)$ is indeed a solution of (8.13). For the behaviour in the interior, i.e. (8.13a), the following classical notion turns out to be the key ingredient.

8.9 Definition. Let $D \subset \mathbb{R}^d$ be an open set. A function $u : D \rightarrow \mathbb{R}$ is *harmonic* (in the set D), if $u \in \mathcal{C}^2(D)$ and $\Delta u(x) = 0$.

In fact, harmonic functions are arbitrarily often differentiable. This is an interesting by-product of the proof of the next result.

8.10 Proposition. *Let $D \subset \mathbb{R}^d$ be an open and connected set and $u : D \rightarrow \mathbb{R}$ a locally bounded function. Then the following assertions are equivalent.*

- a) u is harmonic on D .
- b) u has the (spherical) mean value property: For all $x \in D$ and $r > 0$ such that $\overline{\mathbb{B}}(x, r) \subset D$ one has

$$u(x) = \int_{\partial \mathbb{B}(0, r)} u(x + z) \pi_r(dz) \quad (8.14)$$

where π_r is the normalized uniform measure on the sphere $\partial \mathbb{B}(0, r)$.

- c) u is continuous and $(u(B_{t \wedge \tau_{G^c}}), \mathcal{F}_t)_{t \geq 0}$ is a martingale w.r.t. the law \mathbb{P}^x of a Brownian motion $(B_t, \mathcal{F}_t)_{t \geq 0}$ started at any $x \in G$; here G is a bounded open set such that $\overline{G} \subset D$, and τ_{G^c} is the first exit time from the set G .

Proof. c) \Rightarrow b): Fix $x \in D$ and $r > 0$ such that $\overline{\mathbb{B}}(x, r) \subset D$, and set $\sigma := \tau_{\mathbb{B}^c(x, r)}$. Clearly, there is some bounded open set G such that $\overline{\mathbb{B}}(x, r) \subset \overline{G} \subset D$. By assumption, $(u(B_{t \wedge \tau_{G^c}}), \mathcal{F}_t)_{t \geq 0}$ is a \mathbb{P}^x -martingale.

Since $\sigma < \infty$, cf. Lemma 8.8, we find by optional stopping (Theorem A.18) that $(u(B_{t \wedge \sigma \wedge \tau_{G^c}}))_{t \geq 0}$ is a martingale. Using $\sigma \leq \tau_{G^c}$ we get

$$u(x) = \mathbb{E}^x u(B_0) = \mathbb{E}^x u(B_{t \wedge \sigma}).$$

Now $|u(B_{t \wedge \sigma})| \leq \sup_{y \in \mathbb{B}(x, r)} |u(y)| < \infty$, and by dominated convergence and the continuity of Brownian paths

$$\begin{aligned} u(x) &= \lim_{t \rightarrow \infty} \mathbb{E}^x u(B_{t \wedge \sigma}) = \mathbb{E}^x u(B_\sigma) = \int u(y) \mathbb{P}^x(B_\sigma \in dy) \\ &= \int u(z + x) \mathbb{P}^0(B_\sigma \in dz). \end{aligned}$$

Ex. 2.19 By the rotational symmetry of Brownian motion, cf. Lemma 2.14, $\mathbb{P}^0(B_\sigma \in dz)$ is a uniformly distributed probability measure on the sphere $\partial \mathbb{B}(0, r)$, and (8.14) follows.⁵

b) \Rightarrow a): Fix $x \in D$ and $\delta > 0$ such that $\overline{\mathbb{B}}(x, \delta) \subset D$. Pick $\phi_\delta \in \mathcal{C}^\infty(0, \infty)$ such that $\phi|_{[\delta^2, \infty)} \equiv 0$ and set $c_d := \int_0^\delta \phi_\delta(r^2) r^{d-1} dr$. By Fubini's theorem and changing to polar coordinates

$$\begin{aligned} u(x) &= \frac{1}{c_d} \int_0^\delta \phi_\delta(r^2) r^{d-1} u(x) dr \\ &\stackrel{(8.14)}{=} \frac{1}{c_d} \int_0^\delta \int_{\partial \mathbb{B}(0, r)} \phi_\delta(r^2) r^{d-1} u(z + x) \pi_r(dz) dr \\ &= \frac{1}{c'_d} \int_{\mathbb{B}(0, \delta)} \phi_\delta(|y|^2) u(y + x) dy \\ &= \frac{1}{c'_d} \int_{\mathbb{B}(x, \delta)} \phi_\delta(|y - x|^2) u(y) dy. \end{aligned}$$

As u is locally bounded, the differentiation lemma for parameter dependent integrals, e. g. [169, Theorem 11.5], shows that the last integral is arbitrarily often differentiable at the point x . Since $x \in D$ is arbitrary, we have $u \in \mathcal{C}^\infty(D)$.

Fix $x_0 \in D$ and assume that $\Delta u(x_0) > 0$. Because D is open and Δu continuous, there is some $\delta > 0$ such that $\overline{\mathbb{B}}(x_0, \delta) \subset D$ and $\Delta u|_{\mathbb{B}(x_0, \delta)} > 0$. Set $\sigma := \tau_{\mathbb{B}^c(x_0, \delta)}$ and take $\chi \in \mathcal{C}_c^\infty(D)$ such that $\mathbb{1}_{\mathbb{B}(x_0, \delta)} \leq \chi \leq \mathbb{1}_D$. Then we can use (8.2) for the smooth and bounded function $u_\chi := \chi u$. Using an optional stopping argument we

⁵ For a rotation $Q: \mathbb{R}^d \times \mathbb{R}^d \rightarrow \mathbb{R}^d$, $(QB_t)_{t \geq 0}$ is again a BM^d. Let $\tau = \tau_{\mathbb{B}^c(0, r)}$ and $C \in \mathcal{B}(\partial \mathbb{B}(0, r))$. Then $\mathbb{P}^0(B_\tau \in C) = \mathbb{P}^0(QB_\tau \in C) = \mathbb{P}^0(B_\tau \in Q^{-1}C)$ which shows that $\mathbb{P}^0(B_\tau \in dz)$ must be the (normalized) surface measure on the sphere $\partial \mathbb{B}(0, r)$.

conclude that

$$\begin{aligned} M_{t \wedge \sigma}^{u_\chi} &= u_\chi(B_{t \wedge \sigma}) - u_\chi(B_0) + \int_0^{t \wedge \sigma} \frac{1}{2} \Delta u_\chi(B_r) dr \\ &= u(B_{t \wedge \sigma}) - u(B_0) + \int_0^{t \wedge \sigma} \frac{1}{2} \Delta u(B_r) dr \end{aligned}$$

(use $\chi|_{\mathbb{B}(x, \delta)} \equiv 1$) is a martingale. Therefore, $\mathbb{E} M_{t \wedge \sigma}^{u_\chi} = 0$, and letting $t \rightarrow \infty$ we get

$$\mathbb{E}^{x_0} u(B_\sigma) - u(x_0) = \frac{1}{2} \mathbb{E}^{x_0} \left(\int_0^\sigma \Delta u(B_r) dr \right) > 0.$$

The mean-value property shows that the left-hand side equals 0 and we have a contradiction. Similarly one shows that $\Delta u(x_0) < 0$ gives a contradiction, thus $\Delta u(x) = 0$ for all $x \in D$.

a) \Rightarrow c): Fix any open and bounded set G such that $\overline{G} \subset D$ and pick a cut-off function $\chi \in \mathcal{C}_c^\infty$ such that $\mathbb{1}_G \leq \chi \leq \mathbb{1}_D$. Set $\sigma := \tau_{G^c}$ and $u_\chi(x) := \chi(x)u(x)$. By (8.2),

$$M_t^{u_\chi} = u_\chi(B_t) - u_\chi(B_0) - \int_0^t \frac{1}{2} \Delta u_\chi(B_r) dr$$

is a uniformly integrable martingale. Therefore, we can use optional stopping and the fact that $\chi|_G \equiv 1$ to see that

$$M_{t \wedge \sigma}^{u_\chi} = u(B_{t \wedge \sigma}) - u(B_0) - \int_0^{t \wedge \sigma} \frac{1}{2} \Delta u(B_r) dr \stackrel{\text{a)}}{=} u(B_{t \wedge \sigma}) - u(B_0)$$

is a martingale. This proves c). \square

Let us note another classical consequence of the mean-value property.

8.11 Corollary (Maximum principle). *Let $D \subset \mathbb{R}^d$ be a bounded, open and connected domain and $u \in \mathcal{C}(\overline{D}) \cap \mathcal{C}^2(D)$ be a harmonic function. Then*

$$\exists x_0 \in D : u(x_0) = \sup_{x \in D} u(x) \implies u \equiv u(x_0).$$

In other words: A non-constant harmonic function may attain its maximum only on the boundary of D .

Proof. Assume that u is a harmonic function in D and $u(x_0) = \sup_{x \in D} u(x)$ for some $x_0 \in D$. Since $u(x)$ enjoys the spherical mean-value property (8.14), we see that $u|_{\partial B_r} \equiv u(x_0)$ for all $r > 0$ such that $\overline{\mathbb{B}}(x_0, r) \subset D$. Thus, $u|_{\mathbb{B}(x_0, r)} \equiv u(x_0)$, and we see that the set $M := \{x \in D : u(x) = u(x_0)\}$ is open. Since u is continuous, M is also closed, and the connectedness of D shows that $D = M$ as M is not empty. \square

When we solve the Dirichlet problem, cf. Theorem 8.17 below, we use Proposition 8.10 to get (8.13a) and Corollary 8.11 for the uniqueness; the condition (8.13b), i. e. $\lim_{D \ni x \rightarrow x_0} \mathbb{E}^x f(B_\tau) = f(x_0)$ for $x_0 \in \partial D$, requires some kind of regularity for the hitting time $\tau = \tau_{D^c}$.

8.12 Definition. A point $x_0 \in \partial D$ is called *regular* (for D^c) if $\mathbb{P}^{x_0}(\tau_{D^c} = 0) = 1$. Every non-regular point is called *singular*.

At a regular point x_0 , Brownian motion leaves the domain D immediately; typical situations are shown in Example 8.15 below.

The point $x_0 \in \partial D$ is singular for D^c if, and only if, $\mathbb{P}^{x_0}(\tau_{D^c} > 0) = 1$. If $E \supset D$ such that $x_0 \in \partial E$, the point x_0 is also singular for E^c . This follows from $E^c \subset D^c$ and $\tau_{E^c} \geq \tau_{D^c}$.

8.13 Lemma. $x_0 \in \partial D$ is singular for D^c if, and only if, $\mathbb{P}^{x_0}(\tau_{D^c} = 0) = 0$.

Proof. By definition, any singular point x_0 satisfies $\mathbb{P}^{x_0}(\tau_{D^c} = 0) < 1$. On the other hand,

$$\{\tau_{D^c} = 0\} = \bigcap_{n=1}^{\infty} \{\tau_{D^c} \leq 1/n\} \in \bigcap_{n=1}^{\infty} \mathcal{F}_{1/n} = \mathcal{F}_{0+},$$

and by Blumenthal's 0-1-law, Corollary 6.22, $\mathbb{P}^{x_0}(\tau_{D^c} = 0) \in \{0, 1\}$. \square

The following simple regularity criterion is good enough for most situations. Recall that a *truncated cone* in \mathbb{R}^d with opening $r > 0$, vertex x_0 and direction x_1 is the set $V = \{x \in \mathbb{R}^d : x = x_0 + h \mathbb{B}(x_1, r), 0 < h < \epsilon\}$ for some $\epsilon > 0$.

8.14 Lemma (Outer cone condition. Poincaré 1890, Zaremba 1911). *Let $D \subset \mathbb{R}^d$, $d \geq 2$, be an open domain and $(B_t)_{t \geq 0}$ a BM^d . If we can touch $x_0 \in \partial D$ with the vertex of a truncated cone V such that $V \subset D^c$, then x_0 is a regular point for D^c .*

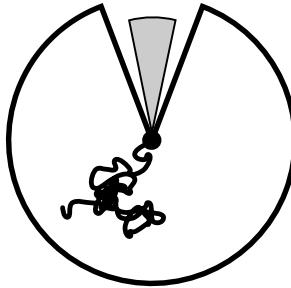


Figure 8.2. The outer cone condition.

Proof. Assume that x_0 is singular. Denote by $\tau_V := \inf\{t > 0 : B_t \in V\}$ the first hitting time of the open cone V . Since $V \subset D^c$, we have $\tau_V \geq \tau_{D^c}$, i. e.

$$\mathbb{P}^{x_0}(\tau_V > 0) \geq \mathbb{P}^{x_0}(\tau_{D^c} > 0) = 1.$$

Pick $\epsilon > 0$ so small that $\mathbb{B}(x_0, \epsilon) \setminus \{x_0\}$ is covered by finitely many copies of the cone $V = V_0$ and its rotations V_1, \dots, V_n about the vertex x_0 . Since a Brownian motion is invariant under rotations – this follows immediately from Lemma 2.14 –, we conclude that $\mathbb{P}^{x_0}(\tau_{V_j} > 0) = \mathbb{P}^{x_0}(\tau_{V_0} > 0) = 1$ for all $j = 1, \dots, n$.

Observe that $\bigcup_{j=0}^n V_j \supset \mathbb{B}(x_0, \epsilon) \setminus \{x_0\}$; therefore, $\min_{0 \leq j \leq n} \tau_{V_j} \leq \tau_{\mathbb{B}(x_0, \epsilon)}$, i. e.

$$\mathbb{P}^{x_0}(\tau_{\mathbb{B}(x_0, \epsilon)} > 0) \geq \mathbb{P}^{x_0}\left(\min_{0 \leq j \leq n} \tau_{V_j} > 0\right) = 1$$

which is impossible since $\mathbb{P}^{x_0}(B_t = x_0) = 0$ for every $t > 0$. \square

8.15 Example. Unless indicated otherwise, we assume that $d \geq 2$. The following pictures show typical examples of singular points.

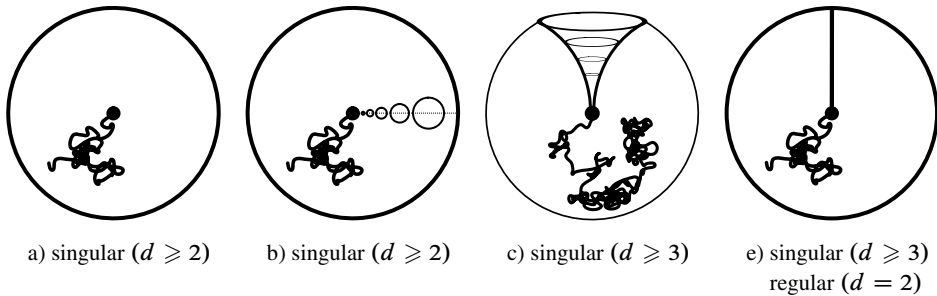


Figure 8.3. Examples of singular points.

- a) The boundary point 0 of $D = \mathbb{B}(0, 1) \setminus \{0\}$ is singular. This was the first example of a singular point and it is due to Zaremba [194]. This is easy to see since $\tau_{\{0\}} = \infty$ a. s.: By the Markov property

$$\begin{aligned} \mathbb{P}^0(\tau_{\{0\}} < \infty) &= \mathbb{P}^0(\exists t > 0 : B_t = 0) \\ &= \lim_{n \rightarrow \infty} \mathbb{P}^0(\exists t > \frac{1}{n} : B_t = 0) \\ &= \lim_{n \rightarrow \infty} \mathbb{E}^0 \left(\underbrace{\mathbb{P}^{B_{1/n}}(\exists t > 0 : B_t = 0)}_{=0, \text{ Corollary 6.16}} \right) = 0. \end{aligned}$$

Therefore, $\mathbb{P}^0(\tau_{D^c} = 0) = \mathbb{P}^0(\tau_{\{0\}} \wedge \tau_{\mathbb{B}^c(0,1)} = 0) = \mathbb{P}^0(\tau_{\mathbb{B}^c(0,1)} = 0) = 0$.

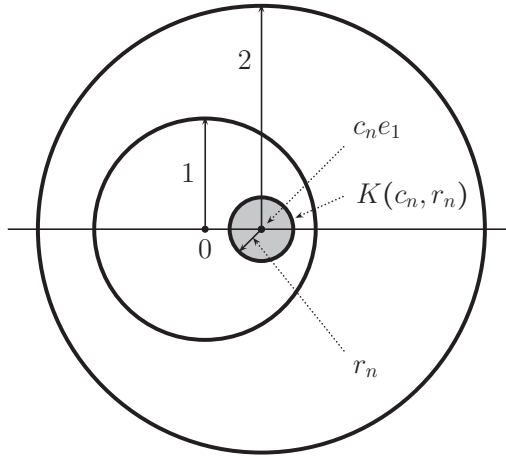


Figure 8.4. Brownian motion has to exit $\mathbb{B}(1, 0)$ before it leaves $\mathbb{B}(2, c_n e_1)$.

- b) Let $(r_n)_{n \geq 1}$ and $(c_n)_{n \geq 1}$ be sequences with $0 < r_n, c_n < 1$ which decrease to 0. We set $K(c_n, r_n) := \overline{\mathbb{B}}(c_n e_1, r_n) \subset \mathbb{R}^2$ and $e_1 = (1, 0) \in \mathbb{R}^2$. Run a Brownian motion in the set $D := \mathbb{B}(0, 1) \setminus \bigcup_{n \geq 1} K(c_n, r_n)$ obtained by removing the circles $K(c_n, r_n)$ from the unit ball. Clearly, the Brownian motion has to leave $\mathbb{B}(0, 1)$ before it leaves $\mathbb{B}(c_n e_1, 2)$, see Figure 8.4. Thus,

$$\begin{aligned} \mathbb{P}^0(\tau_{K(c_n, r_n)} < \tau_{\mathbb{B}^c(0, 1)}) &\leq \mathbb{P}^0(\tau_{K(c_n, r_n)} < \tau_{\mathbb{B}^c(c_n e_1, 2)}) \\ &= \mathbb{P}^{-c_n e_1}(\tau_{K(0, r_n)} < \tau_{\mathbb{B}^c(0, 2)}). \end{aligned}$$

Since $K(0, r_n) = \overline{\mathbb{B}}(0, r_n)$, this probability is known from Theorem 6.15:

$$\mathbb{P}^0(\tau_{K(c_n, r_n)} < \tau_{\mathbb{B}^c(0, 1)}) \leq \frac{\log 2 - \log c_n}{\log 2 - \log r_n}.$$

Take $c_n = 2^{-n}$, $r_n = 2^{-n^3}$, $n \geq 2$, and sum the respective probabilities; then

$$\sum_{n=2}^{\infty} \mathbb{P}^0(\tau_{K(c_n, r_n)} < \tau_{\mathbb{B}^c(0, 1)}) \leq \sum_{n=2}^{\infty} \frac{n+1}{n^3+1} < \infty.$$

The Borel–Cantelli lemma shows that Brownian motion visits at most finitely many of the balls $K(c_n, r_n)$ before it exits $\mathbb{B}(0, 1)$. Since $0 \notin K(c_n, r_n)$, this means that $\mathbb{P}^0(\tau_{D^c} > 0) = 1$, i. e. 0 is a singular point.

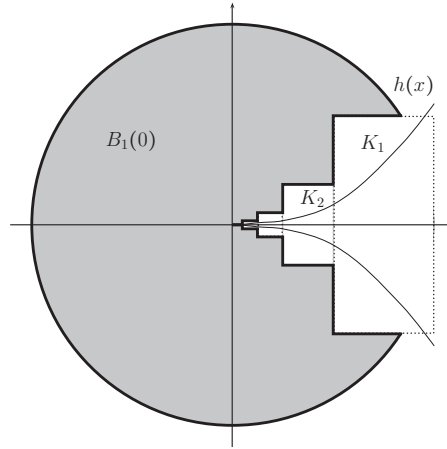


Figure 8.5. A ‘cubistic’ version of Lebesgue’s spine.

- c) (Lebesgue’s spine) In 1913 Lebesgue gave the following example of a singular point. Consider in $d = 3$ the open unit ball and remove a sharp spine, e. g. a closed cusp S which is obtained by rotating a curve $y = h(x)$ about the positive x -axis: $E := \mathbb{B}(0, 1) \setminus S$. Then 0 is a singular point for E^c .

For this, we use a similar construction as in b): We remove from $\mathbb{B}(0, 1)$ the sets K_n , $D := \mathbb{B}(0, 1) \setminus \bigcup_{n=1}^{\infty} K_n$, but the sets K_n are chosen in such a way that they cover the spine S , i. e. $D \subset E$.

Let $K_n = [2^{-n}, 2^{-n+1}] \times \overline{\mathbb{B}}(0, r_n)$ be a cylinder in \mathbb{R}^3 . The probability that a Brownian motion hits K_n can be made arbitrarily small, if we choose r_n small enough, see the technical wrap-up at the end of this paragraph. So we can arrange things in such a way that

$$\mathbb{P}^0(\tau_{K_n} < \tau_{\mathbb{B}(0,1)}) \leq p_n \quad \text{and} \quad \sum_{n=1}^{\infty} p_n < \infty.$$

As in b) we see that a Brownian motion visits at most finitely many of the cylinders K_n before exiting $\mathbb{B}(0, 1)$, i. e. $\mathbb{P}^0(\tau_{D^c} > 0) = 1$. Therefore, 0 is singular for D^c and for $E^c \subset D^c$.

Technical wrap-up. Set $K_{\epsilon}^n = [2^{-n}, 2^{-n+1}] \times \overline{\mathbb{B}}(0, \epsilon)$, $\tau_{n,\epsilon} := \tau_{K_{\epsilon}^n}$, $B_t = (w_t, b_t, \beta_t)$ and $\sigma_n = \inf\{t > 0 : w_t \geq 2^{-n}\}$.

From Corollary 6.17 we deduce that $0 < \sigma_n$, $\tau_{\mathbb{B}^c(0,1)} < \infty$ a. s. Moreover, for fixed $n \geq 1$ and any $\epsilon' \leq \epsilon$ we have $\sigma_n \leq \tau_{n,\epsilon} \leq \tau_{n,\epsilon'}$, and so

$$\{\sigma_n \leq \tau_{n,\epsilon'} \leq \tau_{\mathbb{B}^c(0,1)}\} \subset \{\sigma_n \leq \tau_{n,\epsilon} \leq \tau_{\mathbb{B}^c(0,1)}\}.$$

Since the interval $[\sigma_n, \tau_{\mathbb{B}^c(0,1)}]$ is compact we get

$$\bigcap_{\epsilon > 0} \{\sigma_n \leq \tau_{n,\epsilon} \leq \tau_{\mathbb{B}^c(0,1)}\} \subset \{\exists t \in [\sigma_n, \tau_{\mathbb{B}^c(0,1)}] : b_t = \beta_t = 0\}.$$

As

$$\mathbb{P}^0(\tau_{n,\epsilon} \leq \tau_{\mathbb{B}^c(0,1)}) = \mathbb{P}^0(\sigma_n \leq \tau_{n,\epsilon} \leq \tau_{\mathbb{B}^c(0,1)})$$

and

$$\mathbb{P}^0(\exists t > 0 : b_t = \beta_t = 0) = 0,$$

see Corollary 6.16, we conclude that $\lim_{\epsilon \rightarrow 0} \mathbb{P}^0(\tau_{n,\epsilon} \leq \tau_{\mathbb{B}^c(0,1)}) = 0$ and this is all we need. \square

Our construction shows that the spine will be exponentially sharp. Lebesgue used the function $h(x) = \exp(-1/x)$, $x > 0$, see [100, p. 285, p. 334].

- Ex. 8.8 d) Let $(b(t))_{t \geq 0}$ be a BM^1 and $\tau_{\{0\}} = \inf\{t > 0 : b(t) = 0\}$ the first return time to the origin. Then $\mathbb{P}^0(\tau_{\{0\}} = 0) = 1$. This means that in $d = 1$ a Brownian motion path oscillates infinitely often around its starting point. To see this, we begin with the remark that, because of the symmetry of a Brownian motion,

$$\mathbb{P}^0(b(t) \geq 0) = \mathbb{P}^0(b(t) \leq 0) = \frac{1}{2} \quad \text{for all } t \in [0, \epsilon], \epsilon > 0.$$

Thus, $\mathbb{P}^0(b(t) \geq 0 \forall t \in [0, \epsilon]) \leq 1/2$ and $\mathbb{P}^0(\tau_{(-\infty, 0]} \leq \epsilon) \geq 1/2$ for all $\epsilon > 0$. Letting $\epsilon \rightarrow 0$ we get $\mathbb{P}^0(\tau_{(-\infty, 0]} = 0) \geq 1/2$ and, by Blumenthal's 0-1-law, Corollary 6.22, $\mathbb{P}^0(\tau_{(-\infty, 0]} = 0) = 1$. The same argument applies to the interval $[0, \infty)$ and we get

$$\mathbb{P}^0(\tau_{\{0\}} = 0) = \mathbb{P}^0(\tau_{(-\infty, 0]} \vee \tau_{[0, \infty)} = 0) = 1.$$

- Ex. 8.9 e) (Zaremba's needle) Let $e_1 = (1, 0, \dots, 0) \in \mathbb{R}^d$, $d \geq 2$. The argument from c) shows that for the set $D = \mathbb{B}(0, 1) \setminus [0, \infty)e_1$ the point 0 is singular if $d \geq 3$. In two dimensions the situation is different: 0 is regular!



Indeed: Let $B(t) = (b(t), \beta(t))$, $t \geq 0$, be a BM^2 ; we know that the coordinate processes $b = (b(t))_{t \geq 0}$ and $\beta = (\beta(t))_{t \geq 0}$ are independent one-dimensional Brownian motions. Set $\sigma_n = \inf\{t > 1/n : b(t) = 0\}$. Since $0 \in \mathbb{R}$ is regular for $\{0\} \subset \mathbb{R}$, see d), we get that $\lim_{n \rightarrow \infty} \sigma_n = \tau_{\{0\}} = 0$ almost surely with respect to \mathbb{P}^0 . Since $\beta \perp b$, the random variable β_{σ_n} is symmetric, and so

$$\mathbb{P}^0(\beta(\sigma_n) \geq 0) = \mathbb{P}^0(\beta(\sigma_n) \leq 0) \geq \frac{1}{2}.$$

Clearly, $B(\sigma_n) = (b(\sigma_n), \beta(\sigma_n)) = (0, \beta(\sigma_n))$ and $\{\beta(\sigma_n) \geq 0\} \subset \{\tau_{D^c} \leq \sigma_n\}$, so

$$\mathbb{P}^0(\tau_{D^c} = 0) = \lim_{n \rightarrow \infty} \mathbb{P}^0(\tau_{D^c} \leq \sigma_n) \geq \lim_{n \rightarrow \infty} \mathbb{P}^0(\beta(\sigma_n) \geq 0) \geq \frac{1}{2}.$$

Now Blumenthal's 0-1-law, Corollary 6.22, applies and gives $\mathbb{P}^0(\tau_{D^c} = 0) = 1$.

8.16 Theorem. *Let $x_0 \in \partial D$ be a regular point and $\tau = \tau_{D^c}$. Then*

$$\lim_{D \ni x \rightarrow x_0} \mathbb{P}^x(\tau > h) = 0 \text{ for all } h > 0.$$

Proof. Fix $h > 0$. Since D is open, we find for every $x \in \mathbb{R}^d$

$$\begin{aligned} \mathbb{P}^x(\tau > h) &= \mathbb{P}^x(\forall s \in (0, h] : B(s) \in D) \\ &= \inf_n \mathbb{P}^x(\forall s \in (1/n, h] : B(s) \in D) \\ &= \inf_n \mathbb{E}^x \left(\mathbb{P}^{B(1/n)}(\forall s \in (0, h - 1/n] : B(s) \in D) \right) \\ &= \inf_n \mathbb{E}^x \left(\mathbb{P}^{B(1/n)}(\tau > h - 1/n) \right). \end{aligned} \tag{8.15}$$

Set $\phi(y) := \mathbb{P}^y(\tau > h - 1/n)$. Since $\phi \in \mathcal{B}_b(\mathbb{R}^d)$, the strong Feller property, Ex. 8.3 Proposition 7.3 g), shows that

$$x \mapsto \mathbb{E}^x \left(\mathbb{P}^{B(1/n)}(\tau > h - 1/n) \right) = \mathbb{E}^x \phi(B(1/n))$$

is a continuous function. Therefore,

$$\begin{aligned} \overline{\lim}_{D \ni x \rightarrow x_0} \mathbb{P}^x(\tau > h) &= \overline{\lim}_{D \ni x \rightarrow x_0} \inf_n \mathbb{E}^x \left(\mathbb{P}^{B(1/n)}(\tau > h - 1/n) \right) \\ &\leq \inf_k \overline{\lim}_{D \ni x \rightarrow x_0} \mathbb{E}^x \left(\mathbb{P}^{B(1/k)}(\tau > h - 1/k) \right) \\ &\stackrel{\text{cts.}}{=} \inf_k \mathbb{E}^{x_0} \left(\mathbb{P}^{B(1/k)}(\tau > h - 1/k) \right) \\ &\stackrel{(8.15)}{=} \mathbb{P}^{x_0}(\tau > h) = 0, \end{aligned}$$

since x_0 is regular. □

We are now ready for the main theorem of this section.

8.17 Theorem. *Let $(B_t)_{t \geq 0}$ be a BM^d , D a connected, open and bounded domain and $\tau = \tau_{D^c}$ the first hitting time of D^c . If every point in ∂D is regular for D^c and if $f : \partial D \rightarrow \mathbb{R}$ is continuous, then $u(x) = \mathbb{E}^x f(B_\tau)$ is the unique bounded solution of the Dirichlet problem (8.13).*

Proof. We know already that $u(x) = \mathbb{E}^x f(X_\tau)$ is well-defined for $x \in \overline{D}$.

1° u is harmonic in D . Let $x \in D$. Since D is open, there is some $\delta > 0$ such that $\mathbb{B}(x, \delta) \subset D$. Set $\sigma := \sigma_r := \inf\{t > 0 : B_t \notin \mathbb{B}(x, r)\}$ for $0 < r < \delta$. By the strong Markov property (6.9)

$$u(x) = \mathbb{E}^x f(B_\tau) = \mathbb{E}^x (\mathbb{E}^{B_\sigma} f(B_{\tau'})) = \mathbb{E}^x u(B_\sigma)$$

where τ' is the first exit time from D of the Brownian motion $(B_{\sigma+t} - B_{\sigma})_{t \geq 0}$. Therefore, u enjoys the spherical mean-value property (8.14) and we infer with Proposition 8.10 that u is harmonic in D , i.e. (8.13a) holds.

2° $\lim_{x \rightarrow x_0} u(x) = f(x_0)$ for all $x_0 \in \partial D$: Fix $\epsilon > 0$. Since f is continuous, there is some $\delta > 0$ such that

$$|f(x) - f(x_0)| \leq \epsilon \quad \text{for all } x \in \mathbb{B}(x_0, \delta) \cap \partial D.$$

Therefore, we find for every $h > 0$

$$\begin{aligned} |u(x) - f(x_0)| &= |\mathbb{E}^x (f(B_{\tau}) - f(x_0))| \\ &\leq \mathbb{E}^x \left(|f(B_{\tau}) - f(x_0)| \cdot \mathbb{1}_{\{\tau < h\}} \right) + 2 \|f\|_{\infty} \mathbb{P}^x(\tau \geq h) \\ &\leq \mathbb{E}^x \left(|f(B_{\tau}) - f(x_0)| \cdot \mathbb{1}_{\{\tau < h\}} \mathbb{1}_{\{\sup_{t \leq h} |B_t - x_0| < \delta\}} \right) \\ &\quad + \mathbb{E}^x \left(|f(B_{\tau}) - f(x_0)| \cdot \mathbb{1}_{\{\tau < h\}} \mathbb{1}_{\{\sup_{t \leq h} |B_t - x_0| \geq \delta\}} \right) \\ &\quad + 2 \|f\|_{\infty} \mathbb{P}^x(\tau \geq h). \end{aligned}$$

Write $I_1(x, h)$, $I_2(x, h)$ and $I_3(x, h)$ for the terms appearing in the last line.

Let $|x - x_0| \leq \delta/2$. By the triangle inequality we see that

$$\left\{ \sup_{t \leq h} |B_t + x - x_0| \geq \delta \right\} \subset \left\{ \sup_{t \leq h} |B_t| \geq \delta/2 \right\}.$$

Thus, by Doob's maximal inequality (A.13),

$$\begin{aligned} |I_2(x, h)| &\leq 2 \|f\|_{\infty} \mathbb{P}^0 \left(\sup_{t \leq h} |B_t + x - x_0| \geq \delta \right) \\ &\stackrel{|x-x_0| \leq \delta/2}{\leq} 2 \|f\|_{\infty} \mathbb{P}^0 \left(\sup_{t \leq h} |B_t| \geq \delta/2 \right) \\ &\stackrel{(A.13)}{\leq} \|f\|_{\infty} \frac{8 \mathbb{E}^0 |B_h|^2}{\delta^2} = \|f\|_{\infty} \frac{8hd}{\delta^2}. \end{aligned}$$

Because of the continuity of f we get

$$|I_1(x, h)| \leq \mathbb{E}^x \left(\epsilon \cdot \mathbb{1}_{\{\tau < h\}} \mathbb{1}_{\{\sup_{t \leq h} |B_t - x_0| < \delta\}} \right) \leq \epsilon.$$

Therefore,

$$\begin{aligned} |u(x) - f(x_0)| &\leq 2\epsilon + 2 \|f\|_{\infty} \mathbb{P}^x(\tau \geq h) \\ \text{for } |x - x_0| &\leq \frac{\delta}{2} \quad \text{and} \quad h \leq \frac{\epsilon \delta^2}{8d \|f\|_{\infty}}. \end{aligned}$$

Using Theorem 8.16, we see $\overline{\lim}_{x \rightarrow x_0} |u(x) - f(x_0)| \leq 2\epsilon$, and since $\epsilon > 0$ is arbitrary, this proves (8.13b) as well as (8.13c).

3° Uniqueness. This follows from the maximum principle, Corollary 8.11. Assume that u and w are two solutions of (8.13) with the same boundary value f . Then the functions $\pm(u - w)$ solve (8.13) with $f \equiv 0$, i. e. $\pm(u - w) \leq 0$, hence $u \equiv w$. \square

8.18 Further reading. The study of the Laplace operator and PDEs connected with the Laplacian is a central part of classical and probabilistic potential theory. A good starting point is the non-technical survey paper [138]. On the analytic side, one finds a very neat expositions in [76] on potential theory, and on general PDEs in [62]. Purely probabilistic treatments are [23], its sequel [25] and the monograph [49]. The booklet [55] has a very accessible exposition of the Dirichlet problem starting from random walks. The lecture notes [152] are a good introduction to probabilistic potential theory. Both sides, analysis and probability, are explained in the encyclopedic [42], but this is not an easy read. More general connections to analysis are explored in [7]



- [7] Bass: *Probabilistic Techniques in Analysis*.
- [23] Chung: *Lectures from Markov Processes to Brownian Motion*.
- [25] Chung, Zhao: *From Brownian Motion to Schrödinger's Equation*.
- [42] Doob: *Classical Potential Theory and Its Probabilistic Counterpart*.
- [49] Durrett: *Brownian Motion and Martingales in Analysis*.
- [55] Dynkin, Yushkevich: *Markov Processes. Theorems and Problems*.
- [62] Evans: *Partial Differential Equations*.
- [76] Helms: *Potential Theory*.
- [138] Orey: Probabilistic methods in partial differential equations.
- [152] Rao: *Brownian Motion and Classical Potential Theory*.

Problems

1. Assume that, in the setting of Lemma 8.1, the boundary function $f \in \mathcal{C}_\infty(\mathbb{R}^d)$ but not necessarily $f \in \mathcal{D}(\Delta)$. Consider (8.3) with $P_\epsilon f$ instead of f . Discuss the limit $\epsilon \rightarrow 0$.
2. Let $(B_t)_{t \geq 0}$ be a BM^d , $f : \mathbb{R}^d \rightarrow \mathbb{R}$ be a continuous function such that $\int_0^t f(B_s) ds = 0$ for all $t > 0$. Show that $f(B_s) = 0$ for all $s > 0$, and conclude that $f \equiv 0$.
3. Complete the approximation argument for Lévy's arc-sine law from Paragraph 8.7:
 - (a) Show, by a direct calculation, that $v_{n,\lambda}(x)$ converges as $n \rightarrow \infty$. Conclude from (8.12) that $v''_{n,\lambda}$ converges.

- (b) Integrate the ODE (8.12) to see that $v_{n,\lambda}(x) - v'_{n,\lambda}(0)$ has a limit. Integrating once again shows that $v'_{n,\lambda}(0)$ converges, too.
- (c) Identify the limits of $v_{n,\lambda}$, $v'_{n,\lambda}$ and $v''_{n,\lambda}$.
4. Show Theorem 8.5 with semigroup methods.
Hint: Observe that $A \int_0^t P_s g \, ds = P_t g - g = \frac{d}{dt} \int_0^t P_s g \, ds$.
5. Find the solution to the Dirichlet problem in dimension $d = 1$: $u''(x) = 0$ for all $x \in (0, 1)$, $u(0) = A$, $u(1) = b$ and u is continuous in $[0, 1]$. Compare your findings with Wald's identities, cf. Theorem 5.10 and Corollary 5.11.
6. Use Lemma 7.24 and give an alternative derivation of the result of Lemma 8.8.
7. Show that Lemma 8.8 remains true for any d -dimensional Feller process X_t with continuous paths and generator $L = \sum_{j,k=1}^d a_{jk}(x) \partial_j \partial_k + \sum_{j=1}^d b_j(x) \partial_j$ such that $a_{11}(x) \geq a_0 > 0$ and $|b_1(x)| \leq b_0 < \infty$.
Hint: Set $u(x) = e^{-x_1^2/\gamma r^2}$ and multiply it with a smooth cut-off function $\mathcal{C}_c^2(\mathbb{R}^d)$, $u|_{\mathbb{B}(0,r)} \equiv 1$, to make sure it is in $\mathfrak{D}(L)$. Then calculate $Lu(x) = L(u\chi)(x)$ for $x \in \mathbb{B}(0, r)$.
8. Use the LIL (Corollary 11.2) to give an alternative proof for the fact that a one-dimensional Brownian motion oscillates in every time interval $[0, \epsilon]$ infinitely often around its starting point.
9. (Flat cone condition; Chung 1982) Let $d \geq 2$. A *flat cone* in \mathbb{R}^d is a cone in \mathbb{R}^{d-1} . Adapt the argument of Example 8.12 d) to show the following useful regularity criterion for a BM^d : *The boundary point x_0 is regular for D^c if there is a truncated flat cone with vertex at x_0 and lying entirely in D^c .*
Hint: Example 8.15 e) for $d = 2$ and Chung [23, p. 165].
10. Let (b, β) be a BM^2 and $\sigma_n = \inf\{t > 1/n : b(t) = 0\}$. Show that the random variable $\beta(\sigma_n)$ has a probability density.

Chapter 9

The variation of Brownian paths

Recall the definition of p -variation: Let $f : [0, \infty) \rightarrow \mathbb{R}^d$ be a (non-random) function and let $\Pi = \{t_0 = 0 < t_1 < \dots < t_n = t\}$ be a finite partition of the interval $[0, t]$. By $|\Pi| := \max_{1 \leq j \leq n} (t_j - t_{j-1})$ we denote the *fineness* or *mesh* of the partition. For $p > 0$ we call

$$S_p^\Pi(f; t) := \sum_{t_{j-1}, t_j \in \Pi} |f(t_j) - f(t_{j-1})|^p = \sum_{j=1}^n |f(t_j) - f(t_{j-1})|^p \quad (9.1)$$

the p -variation sum for Π . The supremum over all finite partitions is the *strong* or *total* p -variation,

$$\text{VAR}_p(f; t) := \sup \{S_p^\Pi(f; t) : \Pi \text{ finite partition of } [0, t]\}. \quad (9.2)$$

If $\text{VAR}_1(f; t) < \infty$, the function f is said to be of *bounded* or *finite (total) variation*. Often the following notion of p -variation along a given sequence of finite partitions $(\Pi_n)_{n \geq 1}$ of $[0, t]$ with $\lim_{n \rightarrow \infty} |\Pi_n| = 0$ is used:

$$\text{var}_p(f; t) := \lim_{|\Pi_n| \rightarrow 0} S_p^{\Pi_n}(f; t). \quad (9.3)$$

Note that the strong variation $\text{VAR}_p(f; t)$ is always defined; if $\text{var}_p(f; t)$ exists, $\text{VAR}_p(f; t) \geq \text{var}_p(f; t)$. When a random function $X(s, \omega)$ is used in place of $f(s)$, we can speak of $\text{var}_p(X(\cdot, \omega); t)$ as a limit in law, in probability, in k th mean, and a. s. Note, however, that the sequence of partitions $(\Pi_n)_{n \geq 1}$ does not depend on ω .

If f is continuous, it is no restriction to require that the endpoints $0, t$ of the interval $[0, t]$ are contained in every finite partition Π . Moreover, we may restrict ourselves to partitions with rational points or points from any other dense subset (plus the endpoint t). This follows from the fact that we find for every $\epsilon > 0$ and every partition $\Pi = \{t_0 = 0 < t_1 < \dots < t_n = t\}$ some $\Pi' = \{q_0 = 0 < q_1 < \dots < q_n = t\}$ with points from the dense subset such that

$$\sum_{j=0}^n |f(t_j) - f(q_j)|^p \leq \epsilon.$$

Thus, $|S_p^\Pi(f; t) - S_p^{\Pi'}(f; t)| \leq c_{p,d} \epsilon$ with a constant depending only on p and the dimension d ; since

$$\text{VAR}_p(f; t) = \sup_n \sup \{S_p^\Pi(f; t) : \Pi \text{ partition of } [0, t] \text{ with } \#\Pi = n\},$$

we see that it is enough to calculate $\text{VAR}_p(f; t)$ along rational points.



Ex. 9.3

Ex. 9.4

9.1 The quadratic variation

Ex. 9.2 Let us determine the quadratic variation $\text{var}_2(B; t) = \text{var}_2(B(\cdot, \omega); t)$ of a Brownian motion $B(t), t \geq 0$. In view of the results of Section 2.3 it is enough to consider a one-dimensional Brownian motion throughout this section.

9.1 Theorem. *Let $(B_t)_{t \geq 0}$ be a BM¹ and $(\Pi_n)_{n \geq 1}$ be any sequence of finite partitions of $[0, t]$ satisfying $\lim_{n \rightarrow \infty} |\Pi_n| = 0$. Then the mean-square limit exists:*

$$\text{var}_2(B; t) = L^2(\mathbb{P})\text{-}\lim_{n \rightarrow \infty} S_2^{\Pi_n}(B; t) = t.$$

Usually $\text{var}_2(B; t)$ is called the *quadratic variation* of a Brownian motion.

Proof. Let $\Pi = \{t_0 = 0 < t_1 < \dots < t_n \leq t\}$ be some partition of $[0, t]$. Then we have

$$\mathbb{E} S_2^\Pi(B; t) = \sum_{j=1}^n \mathbb{E} [(B(t_j) - B(t_{j-1}))^2] = \sum_{j=1}^n (t_j - t_{j-1}) = t.$$

Therefore,

$$\mathbb{E} \left[(S_2^\Pi(B; t) - t)^2 \right] = \mathbb{V} \left[S_2^\Pi(B; t) \right] = \mathbb{V} \left[\sum_{j=1}^n (B(t_j) - B(t_{j-1}))^2 \right].$$

By (B1) the random variables $(B(t_j) - B(t_{j-1}))^2, j = 1, \dots, n$, are independent with mean zero. With Bienaymés identity we find

$$\begin{aligned} \mathbb{E} \left[(S_2^\Pi(B; t) - t)^2 \right] &\stackrel{(B1)}{=} \sum_{j=1}^n \mathbb{V} [(B(t_j) - B(t_{j-1}))^2] \\ &= \sum_{j=1}^n \mathbb{E} \left[((B(t_j) - B(t_{j-1}))^2 - (t_j - t_{j-1}))^2 \right] \\ &\stackrel{(B2)}{=} \sum_{j=1}^n \mathbb{E} \left[(B(t_j) - t_{j-1})^2 - (t_j - t_{j-1})^2 \right] \\ &\stackrel{2.12}{=} \sum_{j=1}^n (t_j - t_{j-1})^2 \underbrace{\mathbb{E} \left[(B(1)^2 - 1)^2 \right]}_{=2 \text{ cf. 2.3}} \\ &\leq 2 |\Pi| \sum_{j=1}^n (t_j - t_{j-1}) = 2 |\Pi| t \xrightarrow{|\Pi| \rightarrow 0} 0. \quad \square \end{aligned}$$

9.2 Corollary. *Almost all Brownian paths are of infinite total variation. In fact, we have $\text{VAR}_p(B; t) = \infty$ a.s. for all $p < 2$.*

Proof. Let $p = 2 - \delta$ for some $\delta > 0$. Let Π_n be any sequence of partitions of $[0, t]$ with $|\Pi_n| \rightarrow 0$. Then

$$\begin{aligned} & \sum_{t_{j-1}, t_j \in \Pi_n} (B(t_j) - B(t_{j-1}))^2 \\ & \leq \max_{t_{j-1}, t_j \in \Pi_n} |B(t_j) - B(t_{j-1})|^\delta \sum_{t_{j-1}, t_j \in \Pi_n} |B(t_j) - B(t_{j-1})|^{2-\delta} \\ & \leq \max_{t_{j-1}, t_j \in \Pi_n} |B(t_j) - B(t_{j-1})|^\delta \text{VAR}_{2-\delta}(B; t). \end{aligned}$$

The left-hand side converges, at least for a subsequence, almost surely to t . On the other hand, $\lim_{|\Pi_n| \rightarrow 0} \max_{t_{j-1}, t_j \in \Pi_n} |B(t_j) - B(t_{j-1})|^\delta = 0$ since the Brownian paths are (uniformly) continuous on $[0, t]$. This shows that $\text{VAR}_{2-\delta}(B; t) = \infty$ almost surely. \square

It is not hard to adapt the above results for intervals of the form $[a, b]$. Either we argue directly, or we observe that if $(B_t)_{t \geq 0}$ is a Brownian motion, so is the process $W(t) := B(t + a) - B(a)$, $t \geq 0$. Indeed, it is straightforward to check (B0)–(B4) for the process $W(t)$, cf. 2.9.

Recall that a function $f : [0, \infty) \rightarrow \mathbb{R}$ is called (locally) Hölder continuous of order $\alpha > 0$, if for every compact interval $[a, b] \subset [0, \infty)$ there exists a constant $c = c(f, \alpha, [a, b])$ such that

$$|f(t) - f(s)| \leq c |t - s|^\alpha \quad \text{for all } s, t \in [a, b]. \quad (9.4)$$

9.3 Corollary. *Almost all Brownian paths are nowhere (locally) Hölder continuous of order $\alpha > \frac{1}{2}$.* Ex. 9.6

Proof. Fix some interval $[a, b] \subset [0, \infty)$, $a < b$ are rational, $\alpha > \frac{1}{2}$ and assume that (9.4) holds with $f(t) = B(t, \omega)$ and some constant $c = c(\omega, \alpha, [a, b])$.

Then we find for all finite partitions Π of $[a, b]$

$$\begin{aligned} S_2^\Pi(B(\cdot, \omega), \Pi) &= \sum_{t_{j-1}, t_j \in \Pi} (B(t_j, \omega) - B(t_{j-1}, \omega))^2 \\ &\leq c(\omega)^2 \sum_{t_{j-1}, t_j \in \Pi} (t_j - t_{j-1})^{2\alpha} \leq c(\omega)^2 (b - a) |\Pi|^{2\alpha-1}. \end{aligned}$$

In view of Theorem 9.1 we know that the left-hand side converges almost surely to $b - a$ for some subsequence Π_n with $|\Pi_n| \rightarrow 0$. Since $2\alpha - 1 > 0$, the right-hand side tends to 0, and we have a contradiction.

This shows that (9.4) can only hold on a \mathbb{P} -null set $N_{a,b} \subset \Omega$. But the union of null sets $\bigcup_{0 \leq a < b, a, b \in \mathbb{Q}} N_{a,b}$ is an exceptional set which is uniform for all intervals. \square

9.2 Almost sure convergence of the variation sums

To get almost sure convergence of the variation sums is more difficult. One of the first results in this direction is due to P. Lévy for refining partitions [120, Théorème 5] and the following simple argument for dyadic and other partitions with rapidly decaying mesh-size [121, Théorème 41.1]. In both cases, however, we need an additional condition on the sequence of partitions.

9.4 Theorem (Lévy 1948). *Let $(B_t)_{t \geq 0}$ be a one-dimensional Brownian motion. Along any sequence of finite partitions Π_n of $[0, t]$ such that $\sum_{n=1}^{\infty} |\Pi_n| < \infty$ we have*

$$\text{var}_2(B; t) = \lim_{n \rightarrow \infty} S_2^{\Pi_n}(B; t) = t \quad \text{almost surely.}$$

Proof. As in the proof of Theorem 9.1 we find

$$\mathbb{E} \left[\left(S_2^{\Pi_n}(B; t) - t \right)^2 \right] = 2 \sum_{t_{j-1}, t_j \in \Pi_n} (t_j - t_{j-1})^2 \leq 2 |\Pi_n| t.$$

Now we use the Chebyshev–Markov inequality to get for every $\epsilon > 0$

$$\sum_{n=1}^{\infty} \mathbb{P} \left(\left| S_2^{\Pi_n}(B; t) - t \right| > \epsilon \right) \leq \frac{2t}{\epsilon^2} \sum_{n=1}^{\infty} |\Pi_n| < \infty.$$

By the (easy direction of the) Borel–Cantelli lemma we conclude that there is a set Ω_ϵ such that $\mathbb{P}(\Omega_\epsilon) = 1$ and for all $\omega \in \Omega_\epsilon$ there is some $N(\omega)$ such that

$$\left| S_2^{\Pi_n}(B(\cdot, \omega); t) - t \right| \leq \epsilon \quad \text{for all } n \geq N(\omega).$$

This proves that $\text{var}_2(B; t) = t$ on $\Omega_0 := \bigcap_{k \geq 1} \Omega_{1/k}$ along the sequence $(\Pi_n)_{n \geq 1}$. Since $\mathbb{P}(\Omega_0) = 1$, we are done. \square

The summability condition $\sum_{n=1}^{\infty} |\Pi_n| < \infty$ used in Theorem 9.4 is essentially an assertion on the decay of the sequence $|\Pi_n|$. If the mesh sizes decrease, we have $n|\Pi_n| \leq \sum_{j=1}^n |\Pi_j| \leq \sum_{j=1}^{\infty} |\Pi_j|$, hence $|\Pi_n| = O(1/n)$. The following result of Dudley, [44, p. 89], only requires $|\Pi_n| = o(1/\log n)$, i. e. $\lim_{n \rightarrow \infty} |\Pi_n| \log n = 0$.

9.5 Theorem (Dudley 1973). *Let $(B_t)_{t \geq 0}$ be a BM¹ and assume that Π_n is a sequence of partitions of $[0, t]$ such that $|\Pi_n| = o(1/\log n)$. Then*

$$\text{var}_2(B; t) = \lim_{n \rightarrow \infty} S_2^{\Pi_n}(B; t) = t \quad \text{almost surely.}$$

Proof. To simplify notation, we write in this proof $S_2^{\Pi_n} = S_2^{\Pi_n}(B; t)$. Then

$$\begin{aligned} S_2^{\Pi_n} - t &= \sum_{t_{j-1}, t_j \in \Pi_n} \left[(B(t_j) - B(t_{j-1}))^2 - (t_j - t_{j-1}) \right] \\ &= \sum_{t_{j-1}, t_j \in \Pi_n} (t_j - t_{j-1}) \left[\frac{(B(t_j) - B(t_{j-1}))^2}{t_j - t_{j-1}} - 1 \right] = \sum_{t_{j-1}, t_j \in \Pi_n} s_j [W_j^2 - 1] \end{aligned}$$

where

$$s_j = t_j - t_{j-1} \quad \text{and} \quad W_j = \frac{(B(t_j) - B(t_{j-1}))^2}{\sqrt{t_j - t_{j-1}}}.$$

Note that the W_j are iid $N(0, 1)$ random variables. For every $\epsilon > 0$

$$\mathbb{P}(|S_2^{\Pi_n} - t| > \epsilon) \leq \mathbb{P}(S_2^{\Pi_n} - t > \epsilon) + \mathbb{P}(t - S_2^{\Pi_n} > \epsilon);$$

we can estimate the right-hand side with the Markov inequality. For the first term we find

$$\begin{aligned} \mathbb{P}(S_2^{\Pi_n} - t > \epsilon) &\leq e^{-\epsilon\theta} \mathbb{E} e^{\theta(S_2^{\Pi_n} - t)} \\ &= e^{-\epsilon\theta} \mathbb{E} e^{\theta \sum_j s_j (W_j^2 - 1)} = e^{-\epsilon\theta} \prod_j \mathbb{E} e^{\theta s_j (W_j^2 - 1)}. \end{aligned}$$

(Below we will choose θ in such a way that the right-hand side is finite.) If we combine the elementary inequality $|e^{\lambda x} - \lambda x - 1| \leq \lambda^2 x^2 e^{|\lambda x|}$, $x, \lambda \in \mathbb{R}$, with the fact that $\mathbb{E}(W^2 - 1) = 0$, we get

$$\begin{aligned} \mathbb{P}(S_2^{\Pi_n} - t > \epsilon) &\leq e^{-\epsilon\theta} \prod_j \left[1 + \mathbb{E} \left(e^{\theta s_j (W_j^2 - 1)} - \theta s_j (W_j^2 - 1) - 1 \right) \right] \\ &\leq e^{-\epsilon\theta} \prod_j \left[1 + \mathbb{E} \left(\theta^2 s_j^2 (W_j^2 - 1)^2 e^{|\theta s_j| \cdot |W_j^2 - 1|} \right) \right]. \end{aligned}$$

A direct calculation reveals that for each $0 < \lambda_0 < 1/2$ there is a constant $C = C(\lambda_0)$ Ex. 9.7 such that

$$\mathbb{E}[(W_j^2 - 1)^2 e^{|\lambda| \cdot |W_j^2 - 1|}] \leq C < \infty \quad \text{for all } 1 \leq j \leq n, \quad |\lambda| \leq \lambda_0 < \frac{1}{2}.$$

Using the estimate $1 + y \leq e^y$ we arrive at

$$\mathbb{P}(S_2^{\Pi_n} - t > \epsilon) \leq e^{-\epsilon\theta} \prod_j [C\theta^2 s_j^2 + 1] \leq e^{-\epsilon\theta} \prod_j e^{C\theta^2 s_j^2} = e^{-\epsilon\theta + C\theta^2 s^2}$$

where $s^2 := \sum_j s_j^2$. If we choose

$$\theta = \theta_0 := \min \left\{ \frac{\epsilon}{2Cs^2}, \frac{\lambda_0}{|\Pi_n|} \right\},$$

then $|s_j \theta_0| \leq \lambda_0$ and all expectations in the calculations above are finite. Moreover,

$$\mathbb{P}(S_2^{\Pi_n} - t > \epsilon) \leq e^{-\theta_0(\epsilon - C\theta_0 s^2)} \leq e^{-\frac{1}{2}\epsilon\theta_0} \leq e^{-\frac{\epsilon\lambda_0}{2|\Pi_n|}}.$$

The same estimate holds for $\mathbb{P}(t - S_2^{\Pi_n} > \epsilon)$, and we see

$$\mathbb{P}(|S_2^{\Pi_n} - t| > \epsilon) \leq 2 \exp \left(-\frac{\epsilon\lambda_0}{2|\Pi_n|} \right).$$

Since $|\Pi_n| = o(1/\log n)$, we have $|\Pi_n| = \epsilon_n/\log n$ where $\epsilon_n \rightarrow 0$ as $n \rightarrow \infty$. Thus,

$$\mathbb{P}(|S_2^{\Pi_n} - t| > \epsilon) \leq 2n^{-\frac{\epsilon\lambda_0}{2\epsilon_n}} \leq \frac{2}{n^2} \quad \text{for all } n \geq n_0.$$

The conclusion follows now, as in the proof of Theorem 9.4, from the Borel–Cantelli lemma. \square

Yet another possibility to get almost sure convergence of the variation sums is to consider nested partitions. The proof relies on the a.s. martingale convergence theorem for backwards martingales, cf. Corollary A.8 in the appendix. We begin with an auxiliary result.

9.6 Lemma. *Let $(B_t)_{t \geq 0}$ be a BM¹ and $\mathcal{G}_s := \sigma(B_r, r \leq s; (B_u - B_v)^2, u, v \geq s)$. Then*

$$\mathbb{E}(B_t - B_s | \mathcal{G}_s) = 0 \quad \text{for all } 0 \leq s < t.$$

Proof. From Lemma 2.10 we know that the σ -algebras $\mathcal{F}_s^B = \sigma(B_r, r \leq s)$ and $\mathcal{F}_{[s, \infty)}^B := \sigma(B_t - B_s, t \geq s)$ are independent.

A typical set from a \cap -stable generator of \mathcal{G}_s is of the form

$$\Gamma = \bigcap_{j=1}^M \{B_{r_j} \in C_j\} \cap \bigcap_{k=1}^N \{(B_{v_k} - B_{u_k})^2 \in D_k\}$$

where $M, N \geq 1$, $r_j \leq s \leq u_k \leq v_k$ and $C_j, D_k \in \mathcal{B}(\mathbb{R})$. Since $\mathcal{F}_s^B \perp\!\!\!\perp \mathcal{F}_{[s, \infty)}^B$ and $\tilde{B} := -B$ is again a Brownian motion, we get

$$\begin{aligned} & \int_{\Gamma} (B_t - B_s) d\mathbb{P} \\ &= \underbrace{\int (B_t - B_s) \prod_{k=1}^N \mathbb{1}_{D_k}((B_{v_k} - B_{u_k})^2) d\mathbb{P}}_{\mathcal{F}_{[s, \infty)}^B \text{ measurable}} \cdot \underbrace{\prod_{j=1}^M \mathbb{1}_{C_j}(B_{r_j}) d\mathbb{P}}_{\mathcal{F}_s^B \text{ measurable}} \\ &= \int (B_t - B_s) \prod_{k=1}^N \mathbb{1}_{D_k}((B_{v_k} - B_{u_k})^2) d\mathbb{P} \cdot \int \prod_{j=1}^M \mathbb{1}_{C_j}(B_{r_j}) d\mathbb{P} \\ &= \int (\tilde{B}_t - \tilde{B}_s) \prod_{k=1}^N \mathbb{1}_{D_k}((\tilde{B}_{v_k} - \tilde{B}_{u_k})^2) d\mathbb{P} \cdot \int \prod_{j=1}^M \mathbb{1}_{C_j}(B_{r_j}) d\mathbb{P} \\ &= \int (B_s - B_t) \prod_{k=1}^N \mathbb{1}_{D_k}((B_{u_k} - B_{v_k})^2) d\mathbb{P} \cdot \int \prod_{j=1}^M \mathbb{1}_{C_j}(B_{r_j}) d\mathbb{P} \\ &= \cdots = - \int_{\Gamma} (B_t - B_s) d\mathbb{P}. \end{aligned}$$

This yields $\mathbb{E}(B_t - B_s | \mathcal{G}_s) = 0$. \square

9.7 Theorem (Lévy 1940). *Let $(B_t)_{t \geq 0}$ be a one-dimensional Brownian motion and let Π_n be a sequence of nested partitions of $[0, t]$, i.e. $\Pi_n \subset \Pi_{n+1} \subset \dots$, such that $\lim_{n \rightarrow \infty} |\Pi_n| = 0$. Then*

$$\text{var}_2(B; t) = \lim_{n \rightarrow \infty} S_2^{\Pi_n}(B; t) = t \quad \text{almost surely.}$$

Proof. We can assume that $\#(\Pi_{n+1} \setminus \Pi_n) = 1$, otherwise we add the missing intermediate partitions.

We show that the variation sums $S_{-n} := S_2^{\Pi_n}(B, t)$ are a martingale for the filtration $\mathcal{H}_{-n} := \sigma(S_{-n}, S_{-n-1}, S_{-n-2}, \dots)$. Fix $n \geq 1$; then $\Pi_{n+1} = \Pi_n \cup \{s\}$ and there are two consecutive points $t_j, t_{j+1} \in \Pi_n$ such that $t_j < s < t_{j+1}$. Then

$$\begin{aligned} S_{-n} - S_{-n-1} &= (B_{t_{j+1}} - B_{t_j})^2 - (B_s - B_{t_j})^2 - (B_{t_{j+1}} - B_s)^2 \\ &= 2(B_{t_{j+1}} - B_s)(B_s - B_{t_j}). \end{aligned}$$

Thus, with Lemma 9.6,

$$\mathbb{E}((B_{t_{j+1}} - B_s)(B_s - B_{t_j}) \mid \mathcal{G}_s) = (B_s - B_{t_j}) \mathbb{E}(B_{t_{j+1}} - B_s \mid \mathcal{G}_s) = 0$$

and, since $\mathcal{H}_{-n-1} \subset \mathcal{G}_s$, we can use the tower property to see

$$\mathbb{E}(S_{-n} - S_{-n-1} \mid \mathcal{H}_{-n-1}) = \mathbb{E}(\mathbb{E}(S_{-n} - S_{-n-1} \mid \mathcal{G}_s) \mid \mathcal{H}_{-n-1}) = 0.$$

Thus $(S_{-n}, \mathcal{H}_{-n})_{n \geq 1}$ is a (backwards) martingale and by the martingale convergence theorem, Corollary A.8, we see that the $\lim_{n \rightarrow \infty} S_{-n} = \lim_{n \rightarrow \infty} S_2^{\Pi_n}(B; t)$ exists almost surely. Since the L^2 limit of the variation sums is t , cf. Theorem 9.1, we get that the a. s. limit is again t . \square

9.3 Almost sure divergence of the variation sums

So far we have only considered variation sums where the underlying partition Π does not depend on ω , i.e. on the particular Brownian path. This means, in particular, that we cannot make assertions on the *strong* quadratic variation $\text{VAR}_2(B(\cdot, \omega); 1)$ of a Brownian motion. We are going to show that, in fact, the strong quadratic variation of almost all Brownian paths is infinite. This is an immediate consequence of the following result.

9.8 Theorem (Lévy 1948). *Let $(B_t)_{t \geq 0}$ be a BM¹. For almost all $\omega \in \Omega$ there is a nested sequence of random partitions $(\Pi_n(\omega))_{n \geq 1}$ such that*

$$|\Pi_n(\omega)| \xrightarrow{n \rightarrow \infty} 0 \quad \text{and} \quad \text{var}_2(B(\cdot, \omega); 1) = \lim_{n \rightarrow \infty} S_2^{\Pi_n(\omega)}(B(\cdot, \omega); 1) = +\infty.$$

The strong variation is, by definition, the supremum of the variation sums for *all* finite partitions. Thus,

9.9 Corollary (Lévy 1948). *Let $(B_t)_{t \geq 0}$ be a one-dimensional Brownian motion. Then $\text{VAR}_2(B(\cdot, \omega); 1) = \infty$ for almost all $\omega \in \Omega$.*

Proof of Theorem 9.8 (Freedman 1971). Let $w(t) := B(t, \omega)$, $t \in [0, 1]$, be a fixed Brownian trajectory, write $\Pi_n := \{t_j : t_j = j/n, j = 0, 1, \dots, n\}$ and set $\mathcal{J}_n := \{[t_{j-1}, t_j] : j = 1, \dots, n\}$. An interval $I = [a, b]$ is a k -fast interval for w , if

$$(w(b) - w(a))^2 \geq 2k(b - a).$$

1° If we can show that

$$\forall k \geq 1 \exists n = n(k) \geq k \quad \forall I \in \mathcal{J}_n : I \text{ is a } k\text{-fast interval for } w \quad (9.5)$$

then

$$S_2^{\Pi_{n(k)}}(w; 1) = \sum_{j=1}^{n(k)} |w(t_j) - w(t_{j-1})|^2 \geq 2k \sum_{j=1}^{n(k)} \frac{1}{n(k)} = 2k,$$

and so $\lim_{k \rightarrow \infty} S_2^{\Pi_{n(k)}}(w; 1) = \infty$.

On the other hand, Theorem 9.1 shows that for a Brownian motion we *cannot expect* that *all* intervals are k -fast intervals. We will, however, show that – depending on ω – there are still sufficiently many k -fast intervals. For this we have to relax the condition (9.5).

2° The following condition

$$\begin{aligned} \exists (\alpha_k)_{k \geq 1} \subset (0, 1) \quad \forall k \geq 1, m \geq 1 \exists n = n(k, m) \geq k \quad \forall I \in \mathcal{J}_m : \\ I \text{ contains at least } \lfloor \alpha_k n \rfloor + 1 \text{ many } k\text{-fast intervals from } \mathcal{J}_{mn} \end{aligned} \quad (9.6)$$

allows us to construct a nested sequence of partitions $(\Pi_n)_{n \geq 1}$ such that $|\Pi_n| \rightarrow 0$ and $S_2^{\Pi_n}(w; 1) \rightarrow \infty$.

Assume that (9.6) holds. Fix $k \geq 1, m = m_k \geq 1$ and pick the smallest integer $s_k \geq 1$ such that $(1 - \alpha_k)^{s_k} \leq \frac{1}{2}$. For $I \in \mathcal{J}_m$ we find some $n_1 = n_1(k, m)$ such that

$$\text{each } I \in \mathcal{J}_m \text{ contains at least } \lfloor \alpha_k n_1 \rfloor + 1 \text{ } k\text{-fast intervals } I' \in \mathcal{J}_{mn_1}.$$

Denote these intervals by $\mathcal{J}_{(m, n_1)}^*$. By construction, $\mathcal{J}_{mn_1} \setminus \mathcal{J}_{(m, n_1)}^*$ has no more than $(n_1 - \lfloor \alpha_k n_1 \rfloor - 1)m$ elements.

Applying (9.6) to the intervals from $\mathcal{J}_{mn_1} \setminus \mathcal{J}_{(m, n_1)}^*$ yields that there is some n_2 such that

$$\text{each } I \in \mathcal{J}_{mn_1} \setminus \mathcal{J}_{(m, n_1)}^* \text{ contains at least } \lfloor \alpha_k n_2 \rfloor + 1 \text{ } k\text{-fast intervals } I' \in \mathcal{J}_{mn_1 n_2}.$$

Denote these intervals by $\mathcal{J}_{(m,n_1,n_2)}^*$. By construction, $\mathcal{J}_{mn_1n_2} \setminus \mathcal{J}_{(m,n_1,n_2)}^*$ has at most $(n_2 - \lfloor \alpha_k n_2 \rfloor - 1)(n_1 - \lfloor \alpha_k n_1 \rfloor - 1)m$ elements.

If we repeat this procedure s_k times, there are no more than

$$m \prod_{j=1}^{s_k} (n_j - \lfloor \alpha_k n_j \rfloor - 1)$$

intervals left in $\mathcal{J}_{mn_1 \dots n_{s_k}} \setminus \mathcal{J}_{(m,n_1, \dots, n_{s_k})}^*$; their total length is at most

$$m \prod_{j=1}^{s_k} (n_j - \lfloor \alpha_k n_j \rfloor - 1) \frac{1}{m \prod_{j=1}^{s_k} n_j} \leq (1 - \alpha_k)^{s_k} \leq \frac{1}{2}.$$

The partition Π_k contains all points of the form $t = j/m$, $j = 0, \dots, m$, as well as the endpoints of the intervals from

$$\mathcal{J}_{(m,n_1)}^* \cup \mathcal{J}_{(m,n_1,n_2)}^* \cup \dots \cup \mathcal{J}_{(m,n_1,n_2,\dots,n_{s_k})}^*.$$

Since all intervals are k -fast and since their total length is at least $\frac{1}{2}$, we get

$$S_2^{\Pi_k}(w; 1) \geq 2k \frac{1}{2} = k.$$

3° In order to construct Π_{k+1} , we repeat the construction from the second step with $m = m_k \cdot n_1 \cdot n_2 \cdot \dots \cdot n_{s_k}$. This ensures that $\Pi_{k+1} \subset \Pi_k$.

4° All that remains is to show that almost all Brownian paths satisfy the condition (9.6) for a suitable sequence $(\alpha_k)_{k \geq 1} \subset (0, 1)$. Since a Brownian motion has stationary and independent increments, the random variables

$$Z_l^{(n)} = \# \left\{ 1 \leq j \leq n : \left[B\left(\frac{l}{m} + \frac{j}{mn}, \cdot\right) - B\left(\frac{l}{m} + \frac{j-1}{mn}, \cdot\right) \right]^2 \geq 2k \frac{1}{mn} \right\}$$

($l = 0, 1, \dots, m-1$) are iid binomial random variables with the parameters n and

$$p_k = \mathbb{P} \left(\left[B\left(\frac{l}{m} + \frac{j}{mn}, \cdot\right) - B\left(\frac{l}{m} + \frac{j-1}{mn}, \cdot\right) \right]^2 \geq 2k \frac{1}{mn} \right) = \mathbb{P} \left(|B(1)|^2 \geq 2k \right).$$

In the last equality we used the scaling property 2.12 of a Brownian motion. Observe that $\mathbb{E} Z_l^{(n)} = p_k n$; by the central limit theorem,

$$\mathbb{P} \left(\bigcap_{l=0}^{m-1} \{Z_l^{(n)} \geq \lfloor \tfrac{1}{2} p_k n \rfloor + 1\} \right) = \prod_{l=0}^{m-1} \mathbb{P} \left(Z_l^{(n)} \geq \lfloor \tfrac{1}{2} p_k n \rfloor + 1 \right) \xrightarrow{n \rightarrow \infty} 1,$$

and so

$$\mathbb{P} \left(\bigcup_{n=1}^{\infty} \bigcap_{l=0}^{m-1} \{Z_l^{(n)} \geq \lfloor \tfrac{1}{2} p_k n \rfloor + 1\} \right) = 1.$$

This shows that

$$\mathbb{P} \left(\bigcap_{k=1}^{\infty} \bigcap_{m=1}^{\infty} \bigcup_{n=1}^{\infty} \bigcap_{l=0}^{m-1} \{Z_l^{(n)} \geq \lfloor \tfrac{1}{2} p_k n \rfloor + 1\} \right) = 1$$

which is just (9.6) for the sequence $\alpha_k = \frac{1}{2} \mathbb{P}(|B(1)|^2 \geq 2k)$. \square

9.4 Lévy's characterization of Brownian motion

Theorem 9.1 remains valid for all continuous martingales $(X_t, \mathcal{F}_t)_{t \geq 0}$ with the property that $(X_t^2 - t, \mathcal{F}_t)_{t \geq 0}$ is a martingale. In fact, this is equivalent to saying that

$$\mathbb{E}[X_t - X_s | \mathcal{F}_s] = 0 \quad \text{for all } s \leq t, \quad (9.7)$$

and, since $\mathbb{E}[X_s X_t | \mathcal{F}_s] = X_s \mathbb{E}[X_t | \mathcal{F}_s] = X_s^2$,

$$\mathbb{E}[(X_t - X_s)^2 | \mathcal{F}_s] = t - s \quad \text{for all } s \leq t. \quad (9.8)$$

9.10 Lemma. *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a real-valued martingale with continuous sample paths such that $(X_t^2 - t, \mathcal{F}_t)_{t \geq 0}$ is a martingale. Then $\mathbb{E} X_t^4 < \infty$ and*

$$\mathbb{E}[(X_t - X_s)^4 | \mathcal{F}_s] \leq 4(t - s)^2 \quad \text{for all } s < t. \quad (9.9)$$

Ex. 9.8 *Proof.* We need the following multinomial identity for $a_0, \dots, a_n \in \mathbb{R}$ and $n \geq 1$

$$\begin{aligned} & \left(\sum_j a_j \right)^4 + 3 \sum_j a_j^4 - 4 \left(\sum_j a_j^3 \right) \left(\sum_k a_k \right) + 2 \sum_{(j,k): j < k} a_j^2 a_k^2 \\ &= 2 \sum_j a_j^2 \left(\sum_{k: k \neq j} a_k \right)^2 + 4 \left(\sum_{(j,k): j < k} a_j a_k \right)^2. \end{aligned} \quad (9.10)$$

For the (rather tedious) proof we refer to Lemma A.45 in the appendix.

Set $t_j := s + (t - s) \frac{j}{n}$, $j = 0, 1, \dots, n$, and $\delta_j := X_{t_j} - X_{t_{j-1}}$. By Fatou's lemma

$$\mathbb{E} \left(\lim_{n \rightarrow \infty} \sum_{j=1}^n \delta_j^2 \right) \leq \lim_{n \rightarrow \infty} \mathbb{E} \left(\sum_{j=1}^n \delta_j^2 \right) \stackrel{(9.8)}{=} \lim_{n \rightarrow \infty} \sum_{j=1}^n (t_j - t_{j-1}) = t - s.$$

Therefore, $\lim_{n \rightarrow \infty} \sum_{j=1}^n \delta_j^2$ is almost surely finite, and from the continuity of the sample paths of $t \mapsto X_t$ we find that

$$\lim_{n \rightarrow \infty} \sum_{j=1}^n |\delta_j|^{2+\epsilon} \leq \lim_{n \rightarrow \infty} \max_{1 \leq k \leq n} |\delta_k|^\epsilon \cdot \lim_{n \rightarrow \infty} \sum_{j=1}^n \delta_j^2 = 0.$$

Thus,

$$\begin{aligned} (X_t - X_s)^4 &= \lim_{n \rightarrow \infty} \left[\left(\sum_{j=1}^n \delta_j \right)^4 + 3 \sum_{j=1}^n \delta_j^4 - 4 \left(\sum_{j=1}^n \delta_j^3 \right) \left(\sum_{k=1}^n \delta_k \right) \right] \\ &\leq \lim_{n \rightarrow \infty} \left[\left(\sum_{j=1}^n \delta_j \right)^4 + 3 \sum_{j=1}^n \delta_j^4 - 4 \left(\sum_{j=1}^n \delta_j^3 \right) \left(\sum_{j=1}^n \delta_j \right) + 2 \sum_{k=1}^n \sum_{j < k} \delta_j^2 \delta_k^2 \right]. \end{aligned}$$

From (9.10) we see that the right-hand side is positive. Moreover, (9.10) and the conditional version of Fatou's lemma yield

$$\mathbb{E}[(X_t - X_s)^4 | \mathcal{F}_s] \leq \lim_{n \rightarrow \infty} \mathbb{E} \left[2 \sum_{j=1}^n \delta_j^2 \left(\sum_{k: k \neq j} \delta_k \right)^2 + 4 \left(\sum_{k=1}^n \sum_{j < k} \delta_j \delta_k \right)^2 \middle| \mathcal{F}_s \right].$$

If we expand the squares on the right-hand side, we see that it is of the form

$$c \sum_j \sum_{(k,l): k \neq l, k \neq j, l \neq j} \delta_j^2 \delta_k \delta_l + c' \sum_{(j,k,l,m): j < k < l < m} \delta_j \delta_k \delta_l \delta_m + 8 \sum_{(j,k): j < k} \delta_j^2 \delta_k^2.$$

Recall that $s \leq t_j < t_k < t$ and $\delta_j = X_{t_j} - X_{t_{j-1}}$, $\delta_k = X_{t_k} - X_{t_{k-1}}$. Using the tower property of conditional expectations and (9.8), we find

$$\begin{aligned} \mathbb{E}[\delta_k^2 \delta_j^2 | \mathcal{F}_s] &= \mathbb{E}[\mathbb{E}(\delta_k^2 \delta_j^2 | \mathcal{F}_{t_j}) | \mathcal{F}_s] = \mathbb{E}[\mathbb{E}(\delta_k^2 | \mathcal{F}_{t_j}) \delta_j^2 | \mathcal{F}_s] \\ &\stackrel{(9.8)}{=} \mathbb{E}[(t_k - t_{k-1}) \delta_j^2 | \mathcal{F}_s] \\ &\stackrel{(9.8)}{=} (t_k - t_{k-1})(t_j - t_{j-1}). \end{aligned}$$

If we also use (9.7), a very similar reasoning shows that

$$\mathbb{E}[\delta_j^2 \delta_k \delta_l | \mathcal{F}_s] = \mathbb{E}[\delta_j \delta_k \delta_l \delta_m | \mathcal{F}_s] = 0$$

if the indices j, k, l and m are mutually different. Finally,

$$\begin{aligned} \mathbb{E}[(X_t - X_s)^4 | \mathcal{F}_s] &\leq \lim_{n \rightarrow \infty} 8 \sum_{k=1}^n \sum_{j < k} (t_j - t_{j-1})(t_k - t_{k-1}) \\ &= \lim_{n \rightarrow \infty} 8 \frac{(t-s)^2}{n^2} \frac{n(n-1)}{2} = 4(t-s)^2, \end{aligned}$$

and the claim follows. \square

9.11 Corollary. *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a martingale with continuous sample paths such that $(X_t^2 - t, \mathcal{F}_t)_{t \geq 0}$ is a martingale. For any sequence $(\Pi_n)_{n \geq 1}$ of finite partitions of $[0, t]$ satisfying $\lim_{n \rightarrow \infty} |\Pi_n| = 0$, the mean-square limit exists:*

$$\text{var}_2(X; t) = L^2(\mathbb{P})\text{-}\lim_{n \rightarrow \infty} S_2^{\Pi_n}(X; t) = t.$$

Proof. Let $\Pi := \{0 = t_0 < t_1 < \dots < t_n = t\}$ be some partition of $[0, t]$, set $\delta_j := X_{t_j} - X_{t_{j-1}}$ and $\Delta t_j := t_j - t_{j-1}$, $j = 1, \dots, n$.

For any two increments δ_j and δ_k we find,

$$\begin{aligned} \mathbb{E}[(\delta_j^2 - \Delta t_j)(\delta_k^2 - \Delta t_k)] &= \mathbb{E}[\delta_j^2 \delta_k^2] - \Delta t_j \mathbb{E}[\delta_k^2] - \Delta t_k \mathbb{E}[\delta_j^2] + \Delta t_j \Delta t_k \\ &\stackrel{(9.8)}{=} \mathbb{E}[\delta_j^2 \delta_k^2] - \Delta t_j \Delta t_k. \end{aligned}$$

If $j \neq k$, the proof of Lemma 9.10 shows $\mathbb{E}[\delta_j^2 \delta_k^2] = \Delta t_j \Delta t_k$, i. e. the mixed terms are zero. If $j = k$ we use (9.9) to get $\mathbb{E}[\delta_j^2 \delta_k^2] = \mathbb{E}[\delta_j^4] \leq 4(\Delta t_j)^2$. Thus,

$$\begin{aligned} \mathbb{E}[(S_2^\Pi(X; t) - t)^2] &= \sum_{j=1}^n \sum_{k=1}^n \mathbb{E}[(\delta_j^2 - \Delta t_j)(\delta_k^2 - \Delta t_k)] \\ &= \sum_{j=1}^n \mathbb{E}[(\delta_j^2 - \Delta t_j)^2] \leq 3 \sum_{j=1}^n \Delta t_j^2 \leq 3|\Pi| t \xrightarrow{|\Pi| \rightarrow 0} 0, \end{aligned}$$

where $|\Pi| := \max_{1 \leq k \leq n} \Delta t_k$ is the mesh of the partition. \square

If a process satisfies the assumptions of Lemma 9.10, it is already a Brownian motion. This is Lévy's martingale characterization of Brownian motion.

9.12 Theorem (Lévy 1948). *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$, $X_0 = 0$, be a martingale with continuous sample paths such that $(X_t^2 - t, \mathcal{F}_t)_{t \geq 0}$ is a martingale. Then $(X_t)_{t \geq 0}$ is a Brownian motion.*

Proof (Gikhman, Skorokhod 1972). We have to show that

$$X_t - X_s \perp\!\!\!\perp \mathcal{F}_s \quad \text{and} \quad X_t - X_s \sim \mathbf{N}(0, t - s) \quad \text{for all } 0 \leq s \leq t.$$

This follows from

$$\mathbb{E}\left(e^{i\xi(X_t - X_s)} \middle| \mathcal{F}_s\right) = e^{-\frac{1}{2}(t-s)\xi^2} \quad \text{for all } s \leq t. \quad (9.11)$$

In fact, taking expectations in (9.11) yields

$$\mathbb{E}\left(e^{i\xi(X_t - X_s)}\right) = e^{-\frac{1}{2}(t-s)\xi^2},$$

i. e. $X_t - X_s \sim X_{t-s} \sim \mathbf{N}(0, t - s)$, and (9.11) becomes

$$\mathbb{E}\left(e^{i\xi(X_t - X_s)} \mathbb{1}_F\right) = \mathbb{E}\left(e^{i\xi(X_t - X_s)}\right) \mathbb{P}(F) \quad \text{for all } s \leq t, F \in \mathcal{F}_s;$$

this proves that $X_t - X_s \perp\!\!\!\perp F$ for all $F \in \mathcal{F}_s$.

For (9.11) we use a Taylor expansion up to order 3. Set $\Delta X = X_t - X_s$ and $h = t - s$. Then

$$\begin{aligned} e^{i\xi(X_t - X_s)} - e^{-\frac{1}{2}\xi^2(t-s)} \\ = i\xi\Delta X + \frac{\xi^2}{2}h - \frac{\xi^2}{2}(\Delta X)^2 - \frac{\xi^4}{8}h^2 - \frac{i\xi^3}{6}(\eta\Delta X)^3 + \frac{\xi^6}{48}(\theta h)^3 \end{aligned}$$

where $\eta = \eta(\omega) \in (0, 1)$ and $\theta \in (0, 1)$. If we take conditional expectations with respect to \mathcal{F}_s and observe that $\mathbb{E}(\Delta X | \mathcal{F}_s) = 0$ and $\mathbb{E}((\Delta X)^2 | \mathcal{F}_s) = h$ – this follows from (9.7) and (9.8) –, we get

$$\mathbb{E}\left(e^{i\xi(X_t - X_s)} - e^{-\frac{1}{2}\xi^2(t-s)} \middle| \mathcal{F}_s\right) = -\frac{\xi^4}{8}h^2 - \frac{i\xi^3}{6}\mathbb{E}((\eta\Delta X)^3 | \mathcal{F}_s) + \frac{\xi^6}{48}(\theta h)^3.$$

By (the conditional version of) Hölder's inequality with $p = 4/3$ and $q = 4$ we find

$$\left|\mathbb{E}((\eta\Delta X)^3 | \mathcal{F}_s)\right| \leq \mathbb{E}(|\Delta X|^3 | \mathcal{F}_s) \leq \left[\mathbb{E}((\Delta X)^4 | \mathcal{F}_s)\right]^{3/4} \stackrel{(9.9)}{\leq} 4^{3/4} h^{3/2}.$$

Therefore, we have for all $s < t$ with $h = t - s < 1$ and some constant c_ξ

$$\left|\mathbb{E}\left(e^{i\xi(X_t - X_s)} - e^{-\frac{1}{2}\xi^2(t-s)} \middle| \mathcal{F}_s\right)\right| \leq c_\xi(t-s)^{3/2}. \quad (9.12)$$

Fix $s < t$ and set $t_j := s + (t-s)j/n$, $j = 0, 1, \dots, n$, $n \geq 1$. Observe that

$$e^{i\xi(X_t - X_s)} = \prod_{j=1}^n e^{i\xi(X_{t_j} - X_{t_{j-1}})} \quad \text{and} \quad e^{-\frac{1}{2}\xi^2(t-s)} = \prod_{j=1}^n e^{-\frac{1}{2}\xi^2(t_j - t_{j-1})}.$$

Using the elementary estimate

Ex. 9.9

$$\left|\prod_{j=1}^n a_j - \prod_{j=1}^n b_j\right| \leq \sum_{j=1}^n |a_j - b_j| \quad \text{for all } a_j, b_j \in \mathbb{C}, |a_j|, |b_j| \leq 1,$$

we find

$$\begin{aligned} & \left|\mathbb{E}\left(e^{i\xi(X_t - X_s)} - e^{-\frac{1}{2}\xi^2(t-s)} \middle| \mathcal{F}_s\right)\right| \\ & \leq \mathbb{E}\left(\left|e^{i\xi(X_t - X_s)} - e^{-\frac{1}{2}\xi^2(t-s)}\right| \middle| \mathcal{F}_s\right) \\ & = \mathbb{E}\left(\left|\prod_{j=1}^n e^{i\xi(X_{t_j} - X_{t_{j-1}})} - \prod_{j=1}^n e^{-\frac{1}{2}\xi^2(t_j - t_{j-1})}\right| \middle| \mathcal{F}_s\right) \\ & \leq \sum_{j=1}^n \mathbb{E}\left(\left|e^{i\xi(X_{t_j} - X_{t_{j-1}})} - e^{-\frac{1}{2}\xi^2(t_j - t_{j-1})}\right| \middle| \mathcal{F}_s\right). \end{aligned}$$

If $n \geq 1$ is large enough, we get $h := t_j - t_{j-1} < 1$. This means that we can use (9.12) with $(s, t) = (t_{j-1}, t_j)$ and find

$$\left| \mathbb{E} \left(e^{i\xi(X_t - X_s)} - e^{-\frac{1}{2}\xi^2(t-s)} \mid \mathcal{F}_s \right) \right| \leq c_\xi \sum_{j=1}^n (t_j - t_{j-1})^{3/2} = \frac{c_\xi}{\sqrt{n}} \xrightarrow{n \rightarrow \infty} 0.$$

This proves (9.11). \square



9.13 Further reading. Almost sure convergence of the p -variation of stochastic processes is quite difficult, and no sharp conditions are known. Some of the questions raised in the landmark paper [44] are still open. Up-to-date treatments of p -variation in analysis and probability are the Aarhus lecture notes [47] and the monograph [48].

[44] Dudley: Sample functions of the Gaussian process.

[47] Dudley, Norvaiša: *An Introduction to p -variation and Young integrals*.

[48] Dudley, Norvaiša: *Differentiability of Six Operators on Nonsmooth Functions and p -Variation*.

Problems

1. Let $(\Pi_n)_{n \geq 1}$ be a sequence of refining (i. e. $\Pi_n \subset \Pi_{n+1}$) partitions of $[0, 1]$ such that $\lim_{n \rightarrow \infty} |\Pi_n| = 0$. Show that for every function $f : [0, 1] \rightarrow \mathbb{R}$ the limit

$$\lim_{n \rightarrow \infty} S_1^{\Pi_n}(f, 1) = \text{VAR}_1(f, 1) := \sup \{ S_1^\Pi(f, 1) : \Pi \text{ finite partition of } [0, 1] \}$$

exists in $[0, \infty]$ (and is, in particular, independent of the sequence $(\Pi_n)_{n \geq 1}$).

2. Let $f = (g, h) : [0, \infty) \rightarrow \mathbb{R}^2$ and $p > 0$. Show that $\text{VAR}_p(f; t) < \infty$ if, and only if, $\text{VAR}_p(g; t) + \text{VAR}_p(h; t) < \infty$.
3. Let $f \in \mathcal{C}[0, 1]$. For every partition $\Pi = \{t_0 = 0 < t_1 < \dots < t_n = 1\}$ and $\epsilon > 0$ there is some rational partition $\Pi' = \{q_0 = 0 < q_1 < \dots < q_n = 1\}$ such that

$$\sum_{j=0}^n |f(t_j) - f(q_j)|^p \leq \epsilon.$$

Deduce from this that we may calculate $\text{VAR}_p(f; 1)$ along rational points.

Hint: Show $|S_p^\Pi(f; t) - S_p^{\Pi'}(f; t)| \leq c_{p, \epsilon}$ using $(a+b)^p \leq 2^p(a^p + b^p)$, $a, b \geq 0$, $p \geq 0$.

4. Let f be continuous. Show that it does not affect the finiteness of $\text{VAR}_p(f; t)$ and the numerical value of $\text{Var}_p(f; t)$ if we restrict ourselves to partitions not containing

the endpoints 0 and t of the interval. (In this case we have to define the mesh as mesh of $\Pi \cup \{0, t\}$).

5. (Quadratic variation) Let $(B(t))_{t \geq 0}$ be a one-dimensional Brownian motion. Consider the random variables $Y_n := \sum_{k=1}^n (B(\frac{k}{n}) - B(\frac{k-1}{n}))^2$.
- Find $\mathbb{E} Y_n$ and $\mathbb{V} Y_n$.
 - Determine the probability density of Y_n .
 - Find the characteristic function $\phi_n(\xi) = \mathbb{E} e^{i\xi Y_n}$, $\xi \in \mathbb{R}$, and determine the limit $\lim_{n \rightarrow \infty} \phi_n(\xi)$.
 - Show that $\lim \mathbb{E} [(Y_n - c)^2] = 0$ and determine the constant c .
6. Show that BM^1 is almost surely not 1/2-Hölder continuous:
- For all $Z \sim \text{N}(0, 1)$ and $x > 0$ we have

$$\frac{1}{\sqrt{2\pi}} \frac{x e^{-x^2/2}}{x^2 + 1} < \mathbb{P}(Z > x) < \frac{1}{\sqrt{2\pi}} \frac{e^{-x^2/2}}{x}.$$

- Set $A_{k,n} = \{|B((k+1)2^{-n}) - B(k2^{-n})| > c\sqrt{n} 2^{-n}\}$ and show that for all $c < \sqrt{2 \log 2}$ one has $\mathbb{P}(\lim_{n \rightarrow \infty} \bigcup_{k=1}^{2^n} A_{k,n}) = 1$. For this, use the lower bound of the first part and consider the complement of the set $\lim_{n \rightarrow \infty} \bigcup_{k=1}^{2^n} A_{k,n}$.
 - Conclude from the previous part, that BM^1 is a. s. not 1/2-Hölder continuous.
7. Let $X \sim \text{N}(0, 1)$. Show that for every $\lambda_0 \in (0, \frac{1}{2})$ there is a constant $C = C(\lambda_0)$ such that $\sup_{\lambda \leq \lambda_0} \mathbb{E} \left((X^2 - 1)^2 e^{|\lambda|(X^2 - 1)} \right) \leq C < \infty$.
8. (a) Verify (9.10) for $n = 2$.
 (b) Prove (9.10) by induction.
9. Let $a_j, b_j \in \mathbb{C}$ be complex numbers with $|a_j|, |b_j| \leq 1$, $j = 1, \dots, n$. Prove

$$\left| \prod_{j=1}^n a_j - \prod_{j=1}^n b_j \right| \leq \sum_{j=1}^n |a_j - b_j|.$$

Chapter 10

Regularity of Brownian paths

We have seen in the previous chapter that Brownian paths are not Hölder continuous of any order $\alpha > 1/2$. In particular, they are not Lipschitz continuous or differentiable. In this chapter we continue the study of smoothness properties and show that the paths are actually a. s. nowhere (in t) differentiable.

10.1 Hölder continuity

We begin with a rather general criterion for the continuity of a stochastic process. Later on, in connection with stochastic differential equations, we need the following version for random fields, i. e. stochastic processes with a multi-dimensional index set.

Ex. 10.1 10.1 Theorem (Kolmogorov 1934; Slutsky 1937; Chentsov 1956). *Denote by $(\xi(x))_{x \in \mathbb{R}^n}$ a stochastic process on $(\Omega, \mathcal{A}, \mathbb{P})$ with values in \mathbb{R}^d and index set \mathbb{R}^n . If*

$$\mathbb{E} (|\xi(x) - \xi(y)|^\alpha) \leq c |x - y|^{n+\beta} \quad \text{for all } x, y \in \mathbb{R}^n \quad (10.1)$$

holds for some constants $c > 0$ and $\alpha, \beta > 0$, then $(\xi(x))_{x \in \mathbb{R}^n}$ has a modification $(\xi'(x))_{x \in \mathbb{R}^n}$ with exclusively continuous sample paths. Moreover,

$$\mathbb{E} \left[\left(\sup_{\substack{0 < |x-y| < 1 \\ x, y \in [-T, T]^n}} \frac{|\xi'(x) - \xi'(y)|}{|x - y|^\gamma} \right)^\alpha \right] < \infty, \quad (10.2)$$

for all $T > 0$ and $0 \leq \gamma < \beta/\alpha$. In particular, $x \mapsto \xi'(x)$ is a. s. locally Hölder continuous of order γ .

Proof. Let $\mathbb{N}_0 = \{0, 1, 2, \dots\}$ and $D_m := 2^{-m} \mathbb{N}_0^n \cap [0, 1)^n$; then $D = \bigcup_{m \geq 0} D_m$ are the dyadic numbers in $[0, 1)^n$. Set

$$\Delta_m := \{(x, y) \in D_m \times D_m : |x - y|_\infty \leq 2^{-m}\},$$

where $|x - y|_\infty = \max_{1 \leq k \leq n} |x^k - y^k|$. Since each $x \in D_m$ has at most $3^n - 1$ nearest neighbours in D_m , the set Δ_m contains no more than $3^n \cdot 2^{mn}$ elements.

Fix $\gamma \in (0, \beta/\alpha)$ and define $\sigma_j = \sup_{(x,y) \in \Delta_j} |\xi(x) - \xi(y)|$. Then (10.1) shows

Ex. 10.2

$$\mathbb{E}(\sigma_j^\alpha) \leq \sum_{(x,y) \in \Delta_j} \mathbb{E}(|\xi(x) - \xi(y)|^\alpha) \leq 3^n \cdot 2^{jn} \cdot c 2^{-j(n+\beta)} = c 3^n 2^{-j\beta}.$$

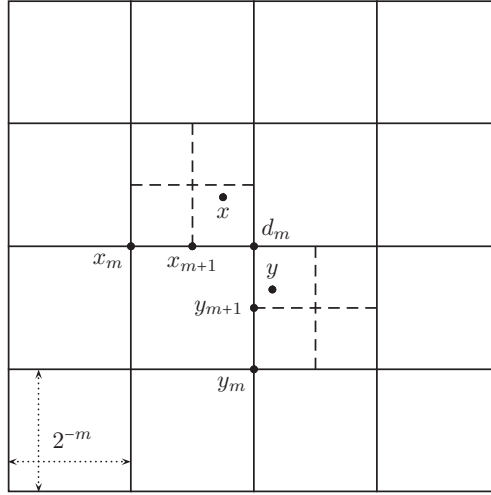


Figure 10.1. Position of the points x and y and their dyadic approximations.

Assume that $x, y \in D$ with $|x - y| < 2^{-m}$. For each $j \geq 0$ there are unique elements $x_j, y_j \in D_j$ such that x and y are in the 2^{-j} -cells $x_j + [0, 2^{-j})^n$ and $y_j + [0, 2^{-j})^n$, respectively. By construction, $\lim_{j \rightarrow \infty} x_j = x$ and $\lim_{j \rightarrow \infty} y_j = y$. For $j = m$ the condition $|x - y| < 2^{-m}$ ensures that the closed cells $x_m + [0, 2^{-m}]^n$ and $y_m + [0, 2^{-m}]^n$ have at least one element $d_m \in D_m$ in common. A possible situation is shown in the picture on the left. Thus, $(x_m, d_m), (y_m, d_m) \in \Delta_m$ and

$$|\xi(x_m) - \xi(y_m)| \leq |\xi(x_m) - \xi(d_m)| + |\xi(y_m) - \xi(d_m)| \leq 2\sigma_m.$$

Moreover, our construction shows that $(x_j, x_{j+1}), (y_j, y_{j+1}) \in \Delta_{j+1}$, $j \geq m$. Therefore,

$$\xi(x) - \xi(y) = \sum_{j \geq m} [\xi(x_{j+1}) - \xi(x_j)] + \xi(x_m) - \xi(y_m) - \sum_{j \geq m} [\xi(y_{j+1}) - \xi(y_j)]$$

and

$$\begin{aligned} |\xi(x) - \xi(y)| &\leq 2\sigma_m + \sum_{j \geq m} [|\xi(x_{j+1}) - \xi(x_j)| + |\xi(y_{j+1}) - \xi(y_j)|] \\ &\leq 2\sigma_m + 2 \sum_{j \geq m+1} \sigma_j = 2 \sum_{j \geq m} \sigma_j. \end{aligned}$$

This entails that

$$\begin{aligned}
 \sup_{x, y \in D, x \neq y} \frac{|\xi(x) - \xi(y)|}{|x - y|^\gamma} &= \sup_{m \geq 0} \sup_{\substack{x, y \in D \\ 2^{-m-1} \leq |x-y| < 2^{-m}}} \frac{|\xi(x) - \xi(y)|}{2^{-(m+1)\gamma}} \\
 &\leq \sup_{m \geq 0} \left(2 \cdot 2^{(m+1)\gamma} \sum_{j \geq m} \sigma_j \right) \\
 &= 2^{1+\gamma} \sup_{m \geq 0} \sum_{j \geq m} 2^{m\gamma} \sigma_j \\
 &\leq 2^{1+\gamma} \sum_{j=0}^{\infty} 2^{j\gamma} \sigma_j.
 \end{aligned}$$

For $\alpha \geq 1$ and $\gamma < \beta/\alpha$ we get

$$\begin{aligned}
 \left\| \sup_{x \neq y, x, y \in D} \frac{|\xi(x) - \xi(y)|}{|x - y|^\gamma} \right\|_{L^\alpha(\mathbb{P})} &\leq 2^{1+\gamma} \sum_{j=0}^{\infty} 2^{j\gamma} \|\sigma_j\|_{L^\alpha(\mathbb{P})} \\
 &\leq c^{1/\alpha} 2^{1+\gamma} \sum_{j=0}^{\infty} 2^{j\gamma} 3^{n/\alpha} 2^{-j\beta/\alpha} \\
 &= c^{1/\alpha} 2^{1+\gamma} 3^{n/\alpha} \sum_{j=0}^{\infty} 2^{j(\gamma-\beta/\alpha)} < \infty.
 \end{aligned}$$

Ex. 10.3 For $\alpha < 1$, the argument is similar, but we use $\mathbb{E}(|Z|^\alpha)$ instead of $\|Z\|_{L^\alpha(\mathbb{P})}$.

Since the expression inside the L^α -norm is almost surely finite, there exists a set $\Omega_1 \subset \Omega$ with $\mathbb{P}(\Omega_1) = 1$ such that for all $\omega \in \Omega_1$ the paths $D \ni x \mapsto \xi(x, \omega)$ are uniformly continuous:

$$|\xi(x, \omega) - \xi(y, \omega)| \leq c(\omega) |x - y|^\gamma \quad \text{for all } x, y \in D, \omega \in \Omega_1.$$

The same construction works if we replace $[0, 1]^n$ by $[-T, T]^n$, and D by the dyadic numbers $D^T \subset [-T, T]^n$. The corresponding exceptional set is now T -dependent. We define $\Omega \setminus \Omega_T$. Set $\Omega_0 = \bigcap_{T=1}^{\infty} \Omega_T$ and $D = \bigcup_{T=1}^{\infty} D^T$. Then $\mathbb{P}(\Omega_0) = 1$ and the mapping $D \ni x \mapsto \xi(x, \omega)$ is continuous if $\omega \in \Omega_0$. Since D is dense in \mathbb{R}^n , we set for all $x \in \mathbb{R}^n$

$$\xi'(x, \omega) := \begin{cases} \lim_{D \ni y \rightarrow x} \xi(y, \omega), & \omega \in \Omega_0, \\ 0, & \omega \notin \Omega_0. \end{cases} \quad (10.3)$$

Let $x_j \in D$ such that $x_j \rightarrow x$. By Fatou's lemma and the assumption (10.1)

$$\begin{aligned}
 \mathbb{E}(|\xi(x) - \xi'(x)|^\alpha) &= \mathbb{E} \left(\liminf_{j \rightarrow \infty} |\xi(x) - \xi(x_j)|^\alpha \right) \\
 &\leq \liminf_{j \rightarrow \infty} \mathbb{E}(|\xi(x) - \xi(x_j)|^\alpha) \leq c \liminf_{j \rightarrow \infty} |x - x_j|^{n+\beta} = 0.
 \end{aligned}$$

Thus, $\mathbb{P}(\xi(x) = \xi'(x)) = 1$ for all $x \in \mathbb{R}^n$. Finally,

$$\begin{aligned} \sup_{\substack{0 < |x-y| < 1 \\ x, y \in [-T, T]^n}} \frac{|\xi'(x) - \xi'(y)|}{|x - y|^\gamma} &= \sup_{\substack{0 < |x-y| < 1 \\ x, y \in [-T, T]^n}} \lim_{\substack{D^T \ni u \rightarrow x \\ D^T \ni v \rightarrow y}} \frac{|\xi(u) - \xi(v)|}{|u - v|^\gamma} \\ &\leq \sup_{\substack{0 < |u-v| < 1 \\ u, v \in D^T}} \frac{|\xi(u) - \xi(v)|}{|u - v|^\gamma} \leq c_T(\omega) \end{aligned}$$

and (10.2) follows. Since the expression under the expectation in (10.2) is almost surely finite, we get

$$|\xi'(x) - \xi'(y)| \leq c(\omega) |x - y|^\gamma \quad \text{for all } |x - y| < 1, \ x, y \in [-T, T]^n. \quad \square$$

10.2 Corollary. *A d -dimensional Brownian motion $(B_t)_{t \geq 0}$ is almost surely Hölder continuous up to order $\gamma < 1/2$.*

Proof. Since an \mathbb{R}^d -valued function is Hölder-continuous of order γ if its components are, it is enough to consider $d = 1$. Since $t \mapsto B(t)$ is continuous, there is no need to take a modification in (the proof of) Theorem 10.1. Because of (2.8) we have

$$\mathbb{E}(|B(t) - B(s)|^{2k}) = c_n |t - s|^k,$$

i. e. we can apply Theorem 10.1 with $d = n = 1$, $\alpha = 2k$ and $\beta = k - 1$. Then

$$\frac{\beta}{\alpha} = \frac{k-1}{2k} \xrightarrow{k \rightarrow \infty} \frac{1}{2}. \quad (10.4)$$

Theorem 10.1 gives for each $\gamma < (k-1)/2k$ an exceptional set $N_k = \Omega \setminus \Omega_k$ where $\mathbb{P}(N_k) = 0$. Since a γ -Hölder continuous function is also γ' -Hölder continuous for all $\gamma' < \gamma$, the assertion follows if we use the null set $N := \bigcup_{k \geq 1} N_k$. \square

10.2 Non-differentiability

We show that almost all paths of a Brownian motion are nowhere differentiable. In view of the results from Section 2.3 it is enough to consider BM^1 . The following theorem was discovered by Paley, Wiener and Zygmund. Our proof follows the classic exposition by Dvoretzky, Erdős and Kakutani.

10.3 Theorem (Paley, Wiener, Zygmund 1931). *Let $(B_t)_{t \geq 0}$ be a BM^1 . Then the path $t \mapsto B_t(\omega)$ is for almost all $\omega \in \Omega$ nowhere differentiable.* Ex. 10.4

Proof (Dvoretzky, Erdős, Kakutani 1961). Set for every $n \geq 1$

$$A_n := \{\omega \in \Omega : B(\cdot, \omega) \text{ nowhere differentiable in } [0, n)\}.$$

It is not clear if the set A_n is measurable. We will show that $\Omega \setminus A_n \subset N_n$ for a measurable null set N_n .

Assume that the function f is differentiable at $t_0 \in [0, n)$. Then

$$\exists \delta > 0 \exists L > 0 \forall t \in \mathbb{B}(t_0, \delta) : |f(t) - f(t_0)| \leq L |t - t_0|.$$

Consider for sufficiently large values of $k \geq 1$ the grid $\{\frac{j}{k} : j = 1, \dots, nk\}$. Then there exists a smallest index $j = j(k)$ such that

$$t_0 \leq \frac{j}{k} \quad \text{and} \quad \frac{j}{k}, \dots, \frac{j+3}{k} \in \mathbb{B}(t_0, \delta).$$

For $i = j+1, j+2, j+3$ we get therefore

$$\begin{aligned} |f(\frac{i}{k}) - f(\frac{i-1}{k})| &\leq |f(\frac{i}{k}) - f(t_0)| + |f(t_0) - f(\frac{i-1}{k})| \\ &\leq L(|\frac{i}{k} - t_0| + |\frac{i-1}{k} - t_0|) \\ &\leq L(\frac{4}{k} + \frac{3}{k}) = \frac{7L}{k}. \end{aligned}$$

If f is a Brownian path, this implies that for the sets

$$C_m^L := \bigcap_{k=m}^{\infty} \bigcup_{j=1}^{kn} \bigcap_{i=j+1}^{j+3} \{|B(\frac{i}{k}) - B(\frac{i-1}{k})| \leq \frac{7L}{k}\}$$

we have

$$\Omega \setminus A_n \subset \bigcup_{L=1}^{\infty} \bigcup_{m=1}^{\infty} C_m^L.$$

Our assertion follows if we can show that $\mathbb{P}(C_m^L) = 0$ for all $m, L \geq 1$. If $k \geq m$,

$$\begin{aligned} \mathbb{P}(C_m^L) &\leq \mathbb{P}\left(\bigcup_{j=1}^{kn} \bigcap_{i=j+1}^{j+3} \{|B(\frac{i}{k}) - B(\frac{i-1}{k})| \leq \frac{7L}{k}\}\right) \\ &\leq \sum_{j=1}^{kn} \mathbb{P}\left(\bigcap_{i=j+1}^{j+3} \{|B(\frac{i}{k}) - B(\frac{i-1}{k})| \leq \frac{7L}{k}\}\right) \\ &\stackrel{(B1)}{=} \sum_{j=1}^{kn} \mathbb{P}\left(\{|B(\frac{i}{k}) - B(\frac{i-1}{k})| \leq \frac{7L}{k}\}\right)^3 \\ &\stackrel{(B2)}{=} kn \mathbb{P}\left(\{|B(\frac{1}{k})| \leq \frac{7L}{k}\}\right)^3 \\ &\leq kn \left(\frac{c}{\sqrt{k}}\right)^3 \xrightarrow{k \rightarrow \infty} 0. \end{aligned}$$

For the last estimate we use $B(\frac{1}{k}) \sim k^{-1/2} B(1)$, cf. 2.12, and therefore

$$\mathbb{P}(|B(\frac{1}{k})| \leq \frac{7L}{k}) = \mathbb{P}(|B(1)| \leq \frac{7L}{\sqrt{k}}) = \frac{1}{\sqrt{2\pi}} \int_{-7L/\sqrt{k}}^{7L/\sqrt{k}} \underbrace{e^{-x^2/2}}_{\leq 1} dx \leq \frac{c}{\sqrt{k}}. \quad \square$$

10.3 Lévy's modulus of continuity

A *modulus of continuity* is a positive, continuous and strictly increasing function $\sigma : [0, \infty) \rightarrow [0, \infty)$ such that $\sigma(0) = 0$. We can use σ in order to describe (local) smoothness properties of a (random or non-random) function $f : [0, \infty) \rightarrow \mathbb{R}^d$ in terms of the following estimate:

$$|f(s) - f(t)| \leq \sigma(|s - t|) \quad \text{for all } s, t \geq 0. \quad (10.5)$$

The most important special case, $\sigma(t) = t^\alpha$, leads to Hölder ($0 < \alpha < 1$) and Lipschitz ($\alpha = 1$) continuity. For Brownian motion we have discussed this in Theorem 10.1.

If $f = (f_1, \dots, f_d)$ is d -dimensional, $|f|^2 = \sum_{j=1}^d f_j^2$, and if each coordinate function f_j satisfies (10.5) with some modulus of continuity σ_j , then f satisfies (10.5) with $\sigma^2(t) := \sum_{j=1}^d \sigma_j^2(t)$. Therefore, it is enough to consider real-valued functions.

The following theorem shows that $\sigma(t) := \sqrt{2t \log \frac{1}{t}}$ is the exact modulus of continuity for one-dimensional Brownian motion. We begin with two auxiliary results.

10.4 Lemma. *The function $\sigma(t) = \sqrt{2t \log \frac{1}{t}}$ is monotone increasing on the interval $(0, e^{-1})$. For all $\kappa \in \mathbb{R}$ and $0 < h < \frac{1}{2}$ we have*

$$\sigma(2^\kappa h) \leq \sqrt{(|\kappa| + 1)2^\kappa} \sigma(h).$$

Proof. A short calculation shows that $\frac{1}{2} \frac{d}{dt} \sigma^2(t) = \log \frac{1}{t} - 1$ is positive on $(0, e^{-1})$. Therefore $\sigma(t)$ increases on this interval. The estimate follows from

$$\frac{\sigma^2(2^\kappa h)}{\sigma^2(h)} = \frac{2 \cdot 2^\kappa h |\log 2^\kappa h|}{2h |\log h|} \leq \frac{2^\kappa (|\log 2^\kappa| + |\log h|)}{|\log h|} \leq 2^\kappa (|\kappa| + 1)$$

since $\log 2 < |\log h|$ for $h < \frac{1}{2}$. \square

10.5 Lemma. *Let $G \sim N(0, 1)$ be a standard normal random variable. Then*

$$\frac{1}{\sqrt{2\pi}} \frac{x}{x^2 + 1} e^{-x^2/2} \leq \mathbb{P}(G > x) = \frac{1}{\sqrt{2\pi}} \int_x^\infty e^{-y^2/2} dy \leq \frac{1}{\sqrt{2\pi}} \frac{1}{x} e^{-x^2/2}$$

for all $x > 0$.

Proof. The upper bound follows from

$$\int_x^\infty e^{-y^2/2} dy \leq \int_x^\infty \frac{y}{x} e^{-y^2/2} dy = \frac{1}{x} e^{-x^2/2}.$$

To see the lower bound we use integration by parts

$$\frac{1}{x^2} \int_x^\infty e^{-y^2/2} dy \geq \int_x^\infty \frac{1}{y^2} e^{-y^2/2} dy = \frac{1}{x} e^{-x^2/2} - \int_x^\infty e^{-y^2/2} dy$$

which gives

$$\int_x^\infty e^{-y^2/2} dy \geq \left(\frac{1}{x^2} + 1\right)^{-1} \frac{1}{x} e^{-x^2/2} = \frac{x}{x^2 + 1} e^{-x^2/2}. \quad \square$$

Ex. 10.5 **10.6 Theorem** (Lévy 1937). *Let $(B_t)_{t \geq 0}$ be a BM¹. Then*

$$\mathbb{P} \left(\overline{\lim}_{h \rightarrow 0} \frac{\sup_{0 \leq t \leq 1-h} |B(t+h) - B(t)|}{\sqrt{2h \log \frac{1}{h}}} = 1 \right) = 1.$$

Proof. We split the proof into two steps. We begin with the easier part and show ‘ $\lim \dots \geq 1$ ’. Since the limes superior is the largest accumulation point it is enough to consider a particular sequence $h_n = 2^{-n} \rightarrow 0$. In fact,

$$\begin{aligned} & \left\{ \overline{\lim}_{h \rightarrow 0} \frac{\sup_{0 \leq t \leq 1-h} |B(t+h) - B(t)|}{\sqrt{2h \log \frac{1}{h}}} \geq 1 \right\} \\ & \supset \left\{ \overline{\lim}_{n \rightarrow \infty} \frac{\max_{0 \leq j \leq 2^n-1} |B(\frac{j+1}{2^n}) - B(\frac{j}{2^n})|}{\sqrt{2 \cdot 2^{-n} \log 2^n}} \geq 1 \right\} \\ & = \bigcap_{r \in \mathbb{Q} \cap (0,1)} \underbrace{\left\{ \overline{\lim}_{n \rightarrow \infty} \frac{\max_{0 \leq j \leq 2^n-1} |B(\frac{j+1}{2^n}) - B(\frac{j}{2^n})|}{\sqrt{2 \cdot 2^{-n} \log 2^n}} \geq r \right\}}_{=: A_r}. \end{aligned}$$

We are going to show that $\mathbb{P}(A_r) = 1$ or, equivalently, $\mathbb{P}(\Omega \setminus A_r) = 0$. Observe that for $\omega \notin A_r$ we have necessarily $\omega \in C_n$ for infinitely many n where

$$\begin{aligned} C_n &= \left\{ \max_{0 \leq j \leq 2^n-1} |B(\frac{j+1}{2^n}) - B(\frac{j}{2^n})| \leq r \sqrt{2 \cdot 2^{-n} \log 2^n} \right\} \\ &= \bigcap_{j=0}^{2^n-1} \left\{ |B(\frac{j+1}{2^n}) - B(\frac{j}{2^n})| \leq r \sqrt{2 \cdot 2^{-n} \log 2^n} \right\}. \end{aligned}$$

By the independence and stationarity of increments (B1), (B2) we see

$$\mathbb{P}(C_n) = \left[1 - \underbrace{\frac{1}{\sqrt{2\pi 2^{-n}}} \int_r^\infty \frac{e^{-x^2/(2 \cdot 2^{-n})}}{\sqrt{2 \cdot 2^{-n} \log 2^n}} dx}_{=: I} \right]^{2^n} = (1 - I)^{2^n} \leq e^{-2^n I}.$$

The last estimate follows from the elementary convexity inequality $1 - e^{-x} \leq x$. From the lower estimate in Lemma 10.5 we get

$$\begin{aligned} 2^n I &= \frac{2^n}{\sqrt{2\pi 2^{-n}}} \int_r^\infty \frac{e^{-x^2/(2 \cdot 2^{-n})}}{\sqrt{2 \cdot 2^{-n} \log 2^n}} dx \\ &= \frac{2^n}{\sqrt{2\pi}} \int_r^\infty \frac{e^{-y^2/2}}{\sqrt{2 \log 2^n}} dy \\ &\geq \frac{1}{\sqrt{2\pi}} \frac{rn \sqrt{2 \log 2}}{1 + r^2 n 2 \log 2} \frac{2^n}{\sqrt{n}} e^{-r^2 \log 2^n} \\ &\geq c \frac{1}{\sqrt{n}} 2^{(1-r^2)n} \end{aligned}$$

for some constant $c = c_r$ and all $n \geq n_0$. Consequently,

$$\sum_{n=n_0}^\infty \mathbb{P}(C_n) \leq \sum_{n=n_0}^\infty \exp(-2^n I) \leq \sum_{n=n_0}^\infty \exp\left(-\frac{c 2^{(1-r^2)n}}{\sqrt{n}}\right) < \infty,$$

and the Borel–Cantelli lemma implies that

$$\mathbb{P}(\Omega \setminus A_r) \leq \mathbb{P}(C_n \text{ for infinitely many } n) = 0.$$

Let us now prove that ‘ $\overline{\lim} \dots \leq 1$ ’. We have to show that for almost all $\omega \in \Omega$ and every $r > 1$ there is some $h_0 = h_0(r, \omega)$ such that

$$\sup_{0 \leq t \leq 1-h} |B(t+h, \omega) - B(t, \omega)| \leq r \sqrt{2h \log \frac{1}{h}} \quad \text{for all } 0 < h \leq h_0. \quad (10.6)$$

Fix $r > 1$, write $r = \frac{1+\delta}{1-\delta}$ for some $\delta \in (0, 1)$ and denote by $\lfloor n\delta \rfloor$ the integer part of $n\delta$. Using the upper estimate from Lemma 10.5 with $x = r \sqrt{2 \log(2^n/\ell)}$, we get

for all $n \geq 1$

$$\begin{aligned}
& \mathbb{P} \left(\max_{1 \leq \ell \leq 2^{\lfloor n\delta \rfloor}} \max_{0 \leq j \leq 2^{n-\ell}} \frac{|B(\frac{j+\ell}{2^n}) - B(\frac{j}{2^n})|}{\sqrt{2 \frac{\ell}{2^n} \log \frac{2^n}{\ell}}} > r \right) \\
& \leq \sum_{\ell=1}^{2^{\lfloor n\delta \rfloor}} \sum_{j=0}^{2^{n-\ell}} \mathbb{P} \left(\frac{|B(\frac{j+\ell}{2^n}) - B(\frac{j}{2^n})|}{\sqrt{\frac{\ell}{2^n}}} > r \sqrt{2 \log \frac{2^n}{\ell}} \right) \\
& \leq \sum_{\ell=1}^{2^{\lfloor n\delta \rfloor}} 2^n \frac{2}{\sqrt{2\pi}} \frac{\exp[-r^2 \log(\frac{2^n}{\ell})]}{\sqrt{2 \log \frac{2^n}{\ell}}} \\
& \leq c \sum_{\ell=1}^{2^{\lfloor n\delta \rfloor}} \frac{2^n}{(2^{n-n\delta})^{r^2} \sqrt{2 \log 2^{n-n\delta}}} \\
& \leq c_\delta 2^{n+n\delta} 2^{-(n-n\delta)r^2} \\
& = c_\delta 2^{n(1+\delta)-nr^2(1-\delta)}. \tag{10.7}
\end{aligned}$$

By construction, $1 + \delta < r^2(1 - \delta)$, so that (10.7) is the general term of a convergent series. By the Borel–Cantelli lemma there is some $\Omega_0 \subset \Omega$ with $\mathbb{P}(\Omega_0) = 1$, such that for all $\omega \in \Omega_0$ there is some $m = m(\epsilon, \omega) \geq 1$ with

$$\max_{1 \leq \ell \leq 2^{\lfloor n\delta \rfloor}} \max_{0 \leq j \leq 2^{n-\ell}} \frac{|B(\frac{j+\ell}{2^n}, \omega) - B(\frac{j}{2^n}, \omega)|}{\sigma(\frac{\ell}{2^n})} \leq r \quad \text{for all } n \geq m. \tag{10.8}$$

Without loss of generality we can assume that $m\delta > 4$. From now on we fix $\omega \in \Omega_0$. Set $h_0 := 2^{-m(1-\delta)}$, pick $h < h_0$ and choose $n \geq m \geq 1$ in such a way that

$$2^{-(n+1)(1-\delta)} \leq h < 2^{-n(1-\delta)}. \tag{10.9}$$

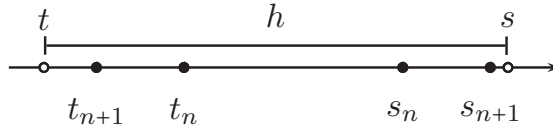


Figure 10.2. Position of the points t_n, t_{n+1} and s_n, s_{n+1} relative to t and s .

For t and $s = t + h$ there are sequences $(s_j)_{j \geq n}$ and $(t_j)_{j \geq n}$ of dyadic numbers such that $s_j, t_j \in \{k/2^j : k \geq 0\}$, $s_j \uparrow s$, $t_j \downarrow t$ and

$$|s_n - t_n| \leq h, \quad |s_j - s_{j+1}| \leq 2^{-j-1} \quad \text{and} \quad |t_j - t_{j+1}| \leq 2^{-j-1}.$$

Since $B(t)$ has continuous sample paths, we get

$$\begin{aligned} & |B(t+h) - B(t)| \\ & \leq |B(s) - B(s_n)| + |B(s_n) - B(t_n)| + |B(t_n) - B(t)| \\ & \leq \sum_{j=n}^{\infty} |B(s_{j+1}) - B(s_j)| + |B(s_n) - B(t_n)| + \sum_{j=n}^{\infty} |B(t_{j+1}) - B(t_j)| \\ & \stackrel{(10.8)}{\leq} r \sum_{j=n}^{\infty} \sigma(2^{-j-1}) + r \sigma(h) + r \sum_{j=n}^{\infty} \sigma(2^{-j-1}) \\ & = 2r \sum_{j=1}^{\infty} \sigma(2^{-j} 2^{-n}) + r \sigma(h) \\ & \stackrel{10.4}{\leq} 2r \sum_{j=1}^{\infty} \sqrt{j+1} 2^{-j/2} \sigma(2^{-n}) + r \sigma(h). \end{aligned}$$

By (10.9) and Lemma 10.4 with $\kappa = (n+1)(1-\delta) - n = 1 - \delta - n\delta$ we see that

$$\sigma(2^{-n}) \leq \sigma(2^{-n} 2^{(n+1)(1-\delta)} h) \leq \underbrace{2^{\kappa/2} \sqrt{|\kappa|+1}}_{\rightarrow 0, n \rightarrow \infty} \sigma(h) \leq (r-1)\sigma(h)$$

for sufficiently large values of $n \geq N(r)$, i.e. sufficiently small values of $h < h_0$. Thus, $2 \sum_{j=1}^{\infty} \sqrt{j+1} 2^{-j/2} \sigma(2^{-n}) \leq c(r-1)\sigma(h)$, and we find for all $h < h_0$

$$\sup_{0 \leq h \leq h_0} \sup_{0 \leq t \leq 1-h} \frac{|B(t+h) - B(t)|}{\sigma(h)} \leq cr(r-1) + r \xrightarrow[r \downarrow 1]{} 1. \quad \square$$

10.7 Corollary. *Almost all sample paths of one-dimensional Brownian motion are nowhere Hölder continuous of order $\alpha = 1/2$.*

Proof. Note that

$$\frac{\sup_{0 \leq t \leq 1-h} |B(t+h) - B(t)|}{\sqrt{h}} = \frac{\sup_{0 \leq t \leq 1-h} |B(t+h) - B(t)|}{\sqrt{2h \log \frac{1}{h}}} \sqrt{2 \log \frac{1}{h}}.$$

By Theorem 10.6 the $\overline{\lim}$ of the right-hand side tends to $+\infty$ as $h \rightarrow 0$. This means that B_t is at no point $t \in [0, 1)$ Hölder continuous.

The same argument, applied to $W(t) := B(t + k) - B(k)$ – this is again a BM^1 , see 2.9 –, completes the proof. \square



10.8 Further reading. There is a whole zoo of function spaces which can be used to measure the regularity of (random) functions. A beautiful result is the estimate [71] which can be used for many embeddings. Regularity in Besov spaces is studied in [28]. The connection of the Kolmogorov–Chentsov theorem and Sobolev embeddings is the topic of the note [168]. A related investigation for fractional Brownian motion is [63]. Moduli of continuity and (non-)differentiability are presented in [32].

[28] Ciesielski, Kerkycharian, Roynette: Quelques espaces fonctionnels associés à des processus gaussiens.

[32] Csörgő, Révész: *Strong Approximations in Probability and Statistics*.

[63] Feyel, de La Pradelle: On fractional Brownian processes.

[71] Garsia, Rodemich, Rumsey: A real variable lemma and the continuity of paths of some Gaussian processes.

[168] Schilling: Sobolev embedding for stochastic processes.

Problems

1. A *Poisson process* is a real-valued stochastic process $(N_t)_{t \geq 0}$ such that $N_0 = 0$, $N_t - N_s \sim N_{t-s}$ and for $t_0 = 0 < t_1 < \dots < t_n$ the increments $N_{t_j} - N_{t_{j-1}}$ are independent and each N_t is a Poisson random variable with parameter $t \cdot \lambda$ for some $\lambda > 0$. (In particular, $(N_t)_{t \geq 0}$ satisfies (B0), (B1), (B2)).
 - (a) Show that $(N_t)_{t \geq 0}$ does *not* satisfy the assumptions of the Kolmogorov–Slutsky–Chentsov theorem, Theorem 10.1.
 - (b) Show that (10.1) holds true for $n = 1$, $\alpha > 0$ and $\beta = 0$. Discuss the role of β for Theorem 10.1.
 - (c) Let $\lambda = 1$. Determine for the process $X_t = N_t - t$ the mean value $m(t)$ and the covariance $C(s, t) = \mathbb{E}(X_s X_t)$, $s, t \geq 0$.
2. Prove that in \mathbb{R}^n all ℓ^p -norms ($1 \leq p \leq \infty$) are equivalent:

$$\max_{1 \leq j \leq n} |x_j| \leq \left(\sum_{j=1}^n |x_j|^p \right)^{1/p} \leq n \max_{1 \leq j \leq n} |x_j| \quad \text{for all } x = (x_1, \dots, x_n) \in \mathbb{R}^n.$$

3. Show that for $\alpha \in (0, 1)$ the function $Z \mapsto \mathbb{E}(|Z|^\alpha)$ is subadditive and complete the argument in the proof of Theorem 10.1 for this case.
4. The proof of Theorem 10.3 actually shows, that almost all Brownian paths are nowhere Lipschitz continuous. Modify the argument of this proof to show that al-

most all Brownian paths are nowhere Hölder continuous for all $\alpha > 1/2$. Why does the argument break down for $\alpha = 1/2$?

Hint: It is enough to consider rational α since α -Hölder continuity implies β -Hölder continuity for all $\beta < \alpha$.

5. Use Theorem 10.6 to show that the strong p -variation $\text{VAR}_p(B;1)$ of BM^1 is for $p > 2$ finite (and, actually, 0).

Chapter 11

The growth of Brownian paths

In this chapter we study the question how fast a Brownian path grows as $t \rightarrow \infty$. Since for a BM¹ $B(t)$, $t > 0$, the process $tB(\frac{1}{t})$, $t > 0$, is again a Brownian motion, we get from the growth at infinity automatically local results for small t and vice versa. A first estimate follows from the strong law of large numbers. Write for $n \geq 1$

$$B(nt) = \sum_{j=1}^n [B(jt) - B((j-1)t)],$$

and observe that the increments of a Brownian motion are stationary and independent random variables. By the strong law of large numbers we get that $B(t, \omega) \leq \epsilon t$ for any $\epsilon > 0$ and all $t \geq t_0(\epsilon, \omega)$. Lévy's modulus of continuity, Theorem 10.6, gives a better bound. For all $\epsilon > 0$ and almost all $\omega \in \Omega$ there is some $h_0 = h_0(\epsilon, \omega)$ such

Ex. 11.1 that

$$|B(h, \omega)| \leq \sup_{0 \leq t \leq 1-h} |B(t+h, \omega) - B(t, \omega)| \leq (1+\epsilon) \sqrt{2h \log \frac{1}{h}}$$

for all $h < h_0$. Since $tB(\frac{1}{t})$ is again a Brownian motion, cf. 2.13, we see that

$$|B(t, \omega)| \leq (1+\epsilon) \sqrt{2t \log t}$$

for all $t > t_0 = h_0^{-1}$. But also this bound is not optimal.

11.1 Khintchine's Law of the Iterated Logarithm

We will show now that the function $\Lambda(t) := \sqrt{2t \log \log t}$, $t \geq 3$, describes very neatly the behaviour of the sample paths as $t \rightarrow \infty$. Still better results can be obtained from Kolmogorov's integral test which we state (without proof) at the end of this section.

The key for the proof of the next theorem is a tail estimate for the normal distribution which we know from Lemma 10.5:

$$\frac{1}{\sqrt{2\pi}} \frac{x}{x^2 + 1} e^{-x^2/2} \leq \mathbb{P}(B(1) > x) \leq \frac{1}{\sqrt{2\pi}} \frac{1}{x} e^{-x^2/2}, \quad x > 0, \quad (11.1)$$

Ex. 11.2 and the following maximal estimate which is, for example, a consequence of the re-

reflection principle, Theorem 6.9:

$$\mathbb{P}\left(\sup_{s \leq t} B(s) > x\right) \leq 2\mathbb{P}(B(t) > x), \quad x > 0, t > 0. \quad (11.2)$$

Because of the form of the numerator the next result is often called the *law of the iterated logarithm* (LIL).

11.1 Theorem (LIL. Khintchine 1933). *Let $(B(t))_{t \geq 0}$ be a BM¹. Then*

Ex. 11.3

$$\mathbb{P}\left(\overline{\lim}_{t \rightarrow \infty} \frac{B(t)}{\sqrt{2t \log \log t}} = 1\right) = 1. \quad (11.3)$$

Since $-B(t)$ and $tB(\frac{1}{t})$ are again Brownian motions, we can immediately get versions of the Khintchine LIL for the limit inferior and for $t \rightarrow 0$.

11.2 Corollary. *Let $(B_t)_{t \geq 0}$ be a BM¹. Then, almost surely,*

$$\begin{aligned} \overline{\lim}_{t \rightarrow \infty} \frac{B(t)}{\sqrt{2t \log \log t}} &= 1, & \overline{\lim}_{t \rightarrow 0} \frac{B(t)}{\sqrt{2t \log \log \frac{1}{t}}} &= 1, \\ \underline{\lim}_{t \rightarrow \infty} \frac{B(t)}{\sqrt{2t \log \log t}} &= -1, & \underline{\lim}_{t \rightarrow 0} \frac{B(t)}{\sqrt{2t \log \log \frac{1}{t}}} &= -1. \end{aligned}$$

Proof of Theorem 11.1. **1°** (Upper bound). Fix $\epsilon > 0, q > 1$ and set

$$A_n := \left\{ \sup_{0 \leq s \leq q^n} B(s) \geq (1 + \epsilon) \sqrt{2q^n \log \log q^n} \right\}.$$

From the maximal estimate (11.2), we see that

$$\begin{aligned} \mathbb{P}(A_n) &\leq 2\mathbb{P}\left(B(q^n) \geq (1 + \epsilon) \sqrt{2q^n \log \log q^n}\right) \\ &= 2\mathbb{P}\left(\frac{B(q^n)}{\sqrt{q^n}} \geq (1 + \epsilon) \sqrt{2 \log \log q^n}\right). \end{aligned}$$

Since $B(q^n)/\sqrt{q^n} \sim B(1)$, we can use the upper bound from (11.1) and find

$$\mathbb{P}(A_n) \leq \frac{1}{(1 + \epsilon)\sqrt{\pi} \sqrt{\log \log q^n}} e^{-(1 + \epsilon)^2 \log \log q^n} \leq c (n \log q)^{-(1 + \epsilon)^2}.$$

This shows that $\sum_{n=1}^{\infty} \mathbb{P}(A_n) < \infty$. Therefore, we can apply the Borel–Cantelli lemma and deduce that

$$\overline{\lim}_{n \rightarrow \infty} \frac{\sup_{0 \leq s \leq q^n} B(s)}{\sqrt{2q^n \log \log q^n}} \leq (1 + \epsilon) \quad \text{a. s.}$$

Since every $t > 1$ is in some interval of the form $[q^{n-1}, q^n]$ and since the function $\Lambda(t) = \sqrt{2t \log \log t}$ is increasing for $t > 3$, we find for all $t \geq q^{n-1} > 3$

$$\frac{B(t)}{\sqrt{2t \log \log t}} \leq \frac{\sup_{s \leq q^n} B(s)}{\sqrt{2q^n \log \log q^n}} \frac{\sqrt{2q^n \log \log q^n}}{\sqrt{2q^{n-1} \log \log q^{n-1}}}.$$

Therefore

$$\overline{\lim}_{t \rightarrow \infty} \frac{B(t)}{\sqrt{2t \log \log t}} \leq (1 + \epsilon) \sqrt{q} \quad \text{a.s.}$$

Letting $\epsilon \rightarrow 0$ and $q \rightarrow 1$ along countable sequences, we find the upper bound in (11.3).

2° If we use the estimate of step 1° for the Brownian motion $(-B(t))_{t \geq 0}$ and with $\epsilon = 1$, we see that

$$-B(q^{n-1}) \leq 2\sqrt{2q^{n-1} \log \log q^{n-1}} \leq \frac{2}{\sqrt{q}} \sqrt{2q^n \log \log q^n} \quad \text{a.s.}$$

for all sufficiently large $n \geq n_0(\epsilon, \omega)$.

3° (Lower bound). So far we have not used the independence and stationarity of the increments of a Brownian motion. The independence of the increments shows that the sets

$$C_n := \left\{ B(q^n) - B(q^{n-1}) \geq \sqrt{2(q^n - q^{n-1}) \log \log q^n} \right\}, \quad n \geq 1,$$

are independent, and the stationarity of the increments implies

$$\begin{aligned} \mathbb{P}(C_n) &= \mathbb{P} \left(\frac{B(q^n) - B(q^{n-1})}{\sqrt{q^n - q^{n-1}}} \geq \sqrt{2 \log \log q^n} \right) \\ &= \mathbb{P} \left(\frac{B(q^n - q^{n-1})}{\sqrt{q^n - q^{n-1}}} \geq \sqrt{2 \log \log q^n} \right). \end{aligned}$$

Since $B(q^n - q^{n-1})/\sqrt{q^n - q^{n-1}} \sim B(1)$ and $2x/(1+x^2) \geq x^{-1}$ for all $x > 1$, we get from the lower tail estimate of (11.1) that

$$\mathbb{P}(C_n) \geq \frac{1}{\sqrt{8\pi}} \frac{1}{\sqrt{2 \log \log q^n}} e^{-\log \log q^n} \geq \frac{c}{n \sqrt{\log n}}.$$

Therefore, $\sum_{n=1}^{\infty} \mathbb{P}(C_n) = \infty$, and since the sets C_n are independent, the Borel-Cantelli lemma applies and shows that, for infinitely many $n \geq 1$,

$$\begin{aligned} B(q^n) &\geq \sqrt{2(q^n - q^{n-1}) \log \log q^n} + B(q^{n-1}) \\ &\stackrel{2^\circ}{\geq} \sqrt{2(q^n - q^{n-1}) \log \log q^n} - \frac{2}{\sqrt{q}} \sqrt{2q^n \log \log q^n}. \end{aligned}$$

Now we divide both sides by $\sqrt{2q^n \log \log q^n}$ and get for all $q > 1$

$$\overline{\lim}_{t \rightarrow \infty} \frac{B(t)}{\sqrt{2t \log \log t}} \geq \overline{\lim}_{n \rightarrow \infty} \frac{B(q^n)}{\sqrt{2q^n \log \log q^n}} \geq \sqrt{1 - \frac{1}{q} - \frac{2}{\sqrt{q}}}$$

almost surely. Letting $q \rightarrow \infty$ along a countable sequence concludes the proof of the lower bound. \square

11.3 Remark (Fast and slow points). For every $t > 0$ $W(h) := B(t + h) - B(t)$ is again a Brownian motion. Therefore, Corollary 11.2 shows that

$$\overline{\lim}_{h \rightarrow 0} \frac{|B(t + h) - B(t)|}{\sqrt{2h \log \log \frac{1}{h}}} = 1 \quad (11.4)$$

almost surely for every fixed $t \in [0, 1]$. Using Fubini's theorem, we can choose the same exceptional set for all $t \in [0, 1]$. It is instructive to compare (11.4) with Lévy's modulus of continuity, Theorem 10.6:

$$\overline{\lim}_{h \rightarrow 0} \frac{\sup_{0 \leq t \leq 1-h} |B(t + h) - B(t)|}{\sqrt{2h \log \frac{1}{h}}} = 1. \quad (11.5)$$

By (11.4), the local modulus of continuity of a Brownian motion is $\sqrt{2h \log \log \frac{1}{h}}$; on the other hand, (11.5) means that globally (and at some necessarily random points) the modulus of continuity will be larger than $\sqrt{2h \log \frac{1}{h}}$. Therefore, a Brownian path must violate the LIL at some random points.

The random set of all times where the LIL fails,

$$E(\omega) := \left\{ t \geq 0 : \overline{\lim}_{h \rightarrow 0} \frac{|B(t + h, \omega) - B(t, \omega)|}{\sqrt{2h \log \log \frac{1}{h}}} \neq 1 \right\},$$

is almost surely uncountable and dense in $[0, \infty)$; this surprising result is due to S. J. Taylor [180], see also Orey and Taylor [139].

At the points $\tau \in E(\omega)$, Brownian motion is faster than usual. Therefore, these points are also called *fast points*. On the other hand, there are also *slow points*. A result by Kahane [93, 94] says that almost surely there exists a $t \in [0, 1]$ such that

$$\overline{\lim}_{h \rightarrow 0} \frac{|B(t + h) - B(t)|}{\sqrt{h}} < \infty.$$

11.4 Remark. Blumenthal's 0-1-law, Corollary 6.22, implies that for any deterministic function $h(t)$

$$\mathbb{P}(B(t) < h(t) \text{ for all sufficiently small } t) \in \{0, 1\}.$$

If the probability is 1, we call h an *upper function*. The following result due to Kolmogorov gives a sharp criterion whether h is an upper function.

11.5 Kolmogorov's test. *If $h(t)$ is increasing and $h(t)/\sqrt{t}$ decreasing, then h is an upper function if, and only if,*

$$\int_0^1 t^{-3/2} h(t) e^{-h^2(t)/2t} dt \text{ converges.}$$

The first proof of this result is due to Petrovsky [143], Motoo's elegant proof from 1959 can be found in Itô and McKean [85, 1.8, p. 33–36 and 4.12 pp. 161–164].

11.2 Chung's 'other' Law of the Iterated Logarithm

Let $(X_j)_{j \geq 1}$ be iid random variables such that $\mathbb{E} X_1 = 0$ and $\mathbb{E} X_1^2 < \infty$. K. L. Chung [22] proved for the maximum of the absolute value of the partial sums the following law of the iterated logarithm

$$\lim_{n \rightarrow \infty} \frac{\max_{1 \leq j \leq n} |X_1 + \cdots + X_j|}{\sqrt{n / \log \log n}} = \frac{\pi}{\sqrt{8}}.$$

A similar result holds for the running maximum of the modulus of a Brownian path. The proof resembles Khintchine's LIL for Brownian motion, but we need a more subtle estimate for the distribution of the maximum.

Using Lévy's triple law (6.19) for $a = -x$, $b = x$ and $t = 1$ we see

$$\begin{aligned} \mathbb{P} \left(\sup_{s \leq 1} |B_s| < x \right) &= \mathbb{P} \left(\inf_{s \leq 1} B_s > -x, \sup_{s \leq 1} B_s < x, B_1 \in (-x, x) \right) \\ &= \frac{1}{\sqrt{2\pi}} \int_{-x}^x \sum_{j=-\infty}^{\infty} (-1)^j e^{-(y+2jx)^2/2} dy \\ &= \frac{1}{\sqrt{2\pi}} \sum_{j=-\infty}^{\infty} (-1)^j \int_{(2j-1)x}^{(2j+1)x} e^{-y^2/2} dy \\ &= \frac{1}{\sqrt{2\pi}} \int_{\mathbb{R}} \underbrace{\sum_{j=-\infty}^{\infty} (-1)^j \mathbb{1}_{((2j-1)x, (2j+1)x)}(y)}_{=f(y)} e^{-y^2/2} dy. \end{aligned}$$

The $4x$ -periodic even function f is given by the following Fourier cosine series¹

$$f(y) = \frac{4}{\pi} \sum_{k=0}^{\infty} \frac{(-1)^k}{2k+1} \cos\left(\frac{2k+1}{2x} \pi y\right).$$

Therefore, we find

$$\begin{aligned} \mathbb{P}\left(\sup_{s \leq 1} |B_s| < x\right) &= \frac{4}{\pi} \sum_{k=0}^{\infty} \frac{1}{\sqrt{2\pi}} \frac{(-1)^k}{2k+1} \int \cos\left(\frac{2k+1}{2x} \pi y\right) e^{-y^2/2} dy \\ &\stackrel{(2.5)}{=} \frac{4}{\pi} \sum_{k=0}^{\infty} \frac{(-1)^k}{2k+1} e^{-\pi^2(2k+1)^2/(8x^2)}. \end{aligned}$$

This proves

11.6 Lemma. *Let $(B(t))_{t \geq 0}$ be a BM¹. Then*

$$\lim_{x \rightarrow 0} \mathbb{P}\left(\sup_{s \leq 1} |B(s)| < x\right) \Big/ \frac{4}{\pi} e^{-\frac{\pi^2}{8x^2}} = 1. \quad (11.6)$$

11.7 Theorem (Chung 1948). *Let $(B(t))_{t \geq 0}$ be a BM¹. Then*

$$\mathbb{P}\left(\lim_{t \rightarrow \infty} \frac{\sup_{s \leq t} |B(s)|}{\sqrt{t/\log \log t}} = \frac{\pi}{\sqrt{8}}\right) = 1. \quad (11.7)$$

Proof. **1°** (Upper bound) Set $t_n := n^n$ and

$$C_n := \left\{ \sup_{t_{n-1} \leq s \leq t_n} |B(s) - B(t_{n-1})| < \frac{\pi}{\sqrt{8}} \sqrt{\frac{t_n}{\log \log t_n}} \right\}.$$

¹ Although we have used Lévy's formula (6.19) for the joint distribution of (m_t, M_t, B_t) , it is not suitable to study the asymptotic behaviour of $\mathbb{P}(m_t < a, M_t > b, B_t \in [c, d])$ as $t \rightarrow \infty$. Our approach to use a Fourier expansion leads essentially to the following expression which is also due to Lévy [121, III.19–20, pp. 78–84]:

$$\mathbb{P}(m_t > a, M_t < b, B_t \in I) = \sum_{j=1}^{\infty} e^{-\lambda_j t} \phi_j(0) \int_I \phi_j(y) dy$$

where λ_j and ϕ_j , $j \geq 1$, are the eigenvalues and normalized eigenfunctions of the boundary value problem

$$\frac{1}{2} f''(x) = \lambda f(x) \text{ for } a < x < b \quad \text{and} \quad f(x) = 0 \text{ for } x = a, x = b.$$

It is not difficult to see that $\lambda_j = j^2 \pi^2 / 2(b-a)^2$ and

$$\phi_j(x) = \sqrt{\frac{2}{b-a}} \cos\left(\frac{j\pi}{b-a} \left[x - \frac{a+b}{2}\right]\right) \quad \text{and} \quad \phi_j(x) = \sqrt{\frac{2}{b-a}} \sin\left(\frac{j\pi}{b-a} \left[x - \frac{a+b}{2}\right]\right).$$

for odd and even indices j , respectively.

By the stationarity of the increments we see

$$\frac{\sup_{t_{n-1} \leq s \leq t_n} |B(s) - B(t_{n-1})|}{\sqrt{t_n}} \sim \frac{\sup_{s \leq t_n - t_{n-1}} |B(s)|}{\sqrt{t_n}} \leq \frac{\sup_{s \leq t_n - t_{n-1}} |B(s)|}{\sqrt{t_n - t_{n-1}}},$$

using Lemma 11.6 for $x = \pi/\sqrt{8 \log \log t_n}$, we get for sufficiently large $n \geq 1$

$$\mathbb{P}(C_n) \geq \mathbb{P}\left(\frac{\sup_{s \leq t_n - t_{n-1}} |B(s)|}{\sqrt{t_n - t_{n-1}}} < \frac{\pi}{\sqrt{8}} \frac{1}{\sqrt{\log \log t_n}}\right) \geq c \frac{4}{\pi} \frac{1}{n \log n}.$$

Therefore, $\sum_n \mathbb{P}(C_n) = \infty$; since the sets C_n are independent, the Borel–Cantelli lemma applies and shows that

$$\lim_{n \rightarrow \infty} \frac{\sup_{t_{n-1} \leq s \leq t_n} |B(s) - B(t_{n-1})|}{\sqrt{t_n / \log \log t_n}} \leq \frac{\pi}{\sqrt{8}}.$$

Ex. 11.3 On the other hand, the argument in Step 1° of the proof of Khintchine’s LIL, Theorem 11.1, yields

$$\overline{\lim}_{n \rightarrow \infty} \frac{\sup_{s \leq t_{n-1}} |B(s)|}{\sqrt{t_n / \log \log t_n}} = \overline{\lim}_{n \rightarrow \infty} \frac{\sup_{s \leq t_{n-1}} |B(s)|}{\sqrt{t_{n-1} \log \log t_{n-1}}} \frac{\sqrt{t_{n-1} \log \log t_{n-1}}}{\sqrt{t_n / \log \log t_n}} = 0.$$

Finally, using $\sup_{s \leq t_n} |B(s)| \leq 2 \sup_{s \leq t_{n-1}} |B(s)| + \sup_{t_{n-1} \leq s \leq t_n} |B(s) - B(t_{n-1})|$ we get

$$\begin{aligned} \lim_{n \rightarrow \infty} \frac{\sup_{s \leq t_n} |B(s)|}{\sqrt{t_n / \log \log t_n}} &\leq \overline{\lim}_{n \rightarrow \infty} \frac{2 \sup_{s \leq t_{n-1}} |B(s)|}{\sqrt{t_n / \log \log t_n}} + \lim_{n \rightarrow \infty} \frac{\sup_{t_{n-1} \leq s \leq t_n} |B(s) - B(t_{n-1})|}{\sqrt{t_n / \log \log t_n}} \\ &\leq \frac{\pi}{\sqrt{8}} \end{aligned}$$

which proves that $\lim_{t \rightarrow \infty} \sup_{s \leq t} |B(s)| / \sqrt{t / \log \log t} \leq \pi / \sqrt{8}$ a.s.

2° (Lower bound) Fix $\epsilon > 0$ and $q > 1$ and set

$$A_n := \left\{ \sup_{s \leq q^n} |B(s)| < (1 - \epsilon) \frac{\pi}{\sqrt{8}} \sqrt{\frac{q^n}{\log \log q^n}} \right\}.$$

Lemma 11.6 with $x = (1 - \epsilon)\pi/\sqrt{8 \log \log q^n}$ gives for large $n \geq 1$

$$\begin{aligned} \mathbb{P}(A_n) &= \mathbb{P}\left(\frac{\sup_{s \leq q^n} |B(s)|}{\sqrt{q^n}} < (1 - \epsilon) \frac{\pi}{\sqrt{8}} \frac{1}{\sqrt{\log \log q^n}}\right) \\ &\leq c \frac{4}{\pi} \left(\frac{1}{n \log q}\right)^{1/(1-\epsilon)^2}. \end{aligned}$$

Thus, $\sum_n \mathbb{P}(A_n) < \infty$, and the Borel–Cantelli lemma yields that

$$\mathbb{P} \left(\sup_{s \leq q^n} |B(s)| < (1 - \epsilon) \frac{\pi}{\sqrt{8}} \sqrt{\frac{q^n}{\log \log q^n}} \text{ for infinitely many } n \right) = 0,$$

i. e. we have almost surely that

$$\lim_{n \rightarrow \infty} \frac{\sup_{s \leq q^n} |B(s)|}{\sqrt{q^n / \log \log q^n}} \geq (1 - \epsilon) \frac{\pi}{\sqrt{8}}.$$

For $t \in [q^{n-1}, q^n]$ and large $n \geq 1$ we get

$$\begin{aligned} \frac{\sup_{s \leq t} |B(s)|}{\sqrt{t / \log \log t}} &\geq \frac{\sup_{s \leq q^{n-1}} |B(s)|}{\sqrt{q^n / \log \log q^n}} \\ &= \frac{\sup_{s \leq q^{n-1}} |B(s)|}{\sqrt{q^{n-1} / \log \log q^{n-1}}} \underbrace{\frac{\sqrt{q^{n-1} / \log \log q^{n-1}}}{\sqrt{q^n / \log \log q^n}}}_{\rightarrow 1/\sqrt{q} \text{ as } n \rightarrow \infty}. \end{aligned}$$

This proves

$$\lim_{t \rightarrow \infty} \frac{\sup_{s \leq t} |B(s)|}{\sqrt{t / \log \log t}} \geq \frac{1 - \epsilon}{\sqrt{q}} \frac{\pi}{\sqrt{8}}$$

almost surely. Letting $\epsilon \rightarrow 0$ and $q \rightarrow 1$ along countable sequences, we get the lower bound in (11.7). \square

If we replace in Theorem 11.7 $\log \log t$ by $\log |\log t|$, and if we use in the proof n^{-n} and $q < 1$ instead of n^n and $q > 1$, respectively, we get the analogue of (11.7) for $t \rightarrow 0$. Combining this with shifting $B(t + h) - B(h)$, we obtain

11.8 Corollary. *Let $(B(t))_{t \geq 0}$ be a BM^1 . Then*

$$\begin{aligned} \lim_{t \rightarrow 0} \frac{\sup_{s \leq t} |B(s)|}{\sqrt{t / \log \log \frac{1}{t}}} &= \frac{\pi}{\sqrt{8}} \quad \text{a. s.}, \\ \lim_{t \rightarrow 0} \frac{\sup_{s \leq t} |B(s + h) - B(h)|}{\sqrt{t / \log \log \frac{1}{t}}} &= \frac{\pi}{\sqrt{8}} \quad \text{a. s. for all } h \geq 0. \end{aligned}$$

Using similar methods Csörgö and Révész [32, pp. 44–47] show that

$$\lim_{h \rightarrow 0} \inf_{s \leq 1-h} \sup_{t \leq h} \frac{|B(s+t) - B(s)|}{\sqrt{h / \log \frac{1}{h}}} = \frac{\pi}{\sqrt{8}}. \quad (11.8)$$

This is the exact *modulus of non-differentiability* of a Brownian motion.



11.9 Further reading. Path properties of the LIL-type are studied in [32], [65] and [155]. Apart from the original papers which are mentioned in Remark 11.3, exceptional sets are beautifully treated in [133].

[32] Csörgő, Révész: *Strong Approximations in Probability and Statistics*.

[65] Freedman: *Brownian Motion and Diffusion*.

[133] Mörters, Peres: *Brownian Motion*.

[155] Révész: *Random Walk in Random and Non-Random Environments*.

Problems

1. Let $(B_t)_{t \geq 0}$ be a BM^1 . Use the Borel-Cantelli lemma to show that the running maximum $M_n := \sup_{0 \leq t \leq n} B_t$ cannot grow faster than $C \sqrt{n \log n}$ for any $C > 2$. Use this to show that

$$\overline{\lim}_{t \rightarrow \infty} \frac{M_t}{C \sqrt{t \log t}} \leq 1 \quad \text{a. s.}$$

Hint: Show that $M_n(\omega) \leq C \sqrt{n \log n}$ for sufficiently large $n \geq n_0(\omega)$.

2. Let $(B_t)_{t \geq 0}$ be a BM^1 . Apply Doob's maximal inequality (A.14) to the exponential martingale $M_t^\xi := e^{\xi B_t - \frac{1}{2} \xi^2 t}$ to show that

$$\mathbb{P} \left(\sup_{s \leq t} (B_s - \frac{1}{2} \xi s) > x \right) \leq e^{-x \xi}.$$

(This inequality can be used for the upper bound in the proof of Theorem 11.1, avoiding the combination of (11.2) and the upper bound in (11.1))

3. Show that the proof of Khinchine's LIL, Theorem 11.1, can be modified to give

$$\overline{\lim}_{t \rightarrow \infty} \frac{\sup_{s \leq t} |B(s)|}{\sqrt{2t \log \log t}} \leq 1.$$

Hint: Use in Step 1° of the proof $\mathbb{P}(\sup_{s \leq t} |B(s)| \geq x) \leq 4 \mathbb{P}(|B(t)| \geq x)$.

4. (Wentzell [184, p. 176]) Let $(B_t)_{t \geq 0}$ be a one-dimensional Brownian motion, $a, b > 0$ and $\tau := \inf \{t \geq 0 : B_t = b \sqrt{a + t}\}$. Show that

(a) $\mathbb{P}(\tau < \infty) = 1$;

(b) $\mathbb{E} \tau = \infty$ ($b \geq 1$);

(c) $\mathbb{E} \tau < \infty$ ($b < 1$).

Hint: Use in (b) Wald's identities. For (c) show that $\mathbb{E}(\tau \wedge n) \leq ab^2(1 - b^2)^{-1}$ for $n \geq 1$.

Chapter 12

Strassen's Functional Law of the Iterated Logarithm

From Khintchine's law of the iterated logarithm (LIL), Corollary 11.2, we know that for a one-dimensional Brownian motion

$$-1 = \liminf_{s \rightarrow \infty} \frac{B(s)}{\sqrt{2s \log \log s}} < \overline{\lim}_{s \rightarrow \infty} \frac{B(s)}{\sqrt{2s \log \log s}} = 1.$$

Since Brownian paths are a. s. continuous, it is clear that *the set of limit points of $\{B(s)/\sqrt{2s \log \log s} : s > e\}$ as $s \rightarrow \infty$ is the whole interval $[-1, 1]$.*

This observation is the starting point for an extension of Khintchine's LIL, due to Strassen [176], which characterizes the set of limit points as $s \rightarrow \infty$ of the family

$$\{Z_s(\cdot) : s > e\} \subset \mathcal{C}_{(0)}[0, 1] \quad \text{where} \quad Z_s(t) := \frac{B(st)}{\sqrt{2s \log \log s}}, \quad 0 \leq t \leq 1.$$

($\mathcal{C}_{(0)}[0, 1]$ denotes the set of all continuous functions $w : [0, 1] \rightarrow \mathbb{R}$ with $w(0) = 0$.)

If $w_0 \in \mathcal{C}_{(0)}[0, 1]$ is a limit point it is necessary that for every fixed $t \in [0, 1]$ the number $w_0(t)$ is a limit point of the family $\{Z_s(t) : s > e\} \subset \mathbb{R}$. Using Khintchine's LIL and the scaling property 2.12 of Brownian motion it is straightforward that the limit points of $\{Z_s(t) : s > e\}$ are the whole interval $[-\sqrt{t}, \sqrt{t}]$, i. e. $-\sqrt{t} \leq w_0(t) \leq \sqrt{t}$. In fact, we have

Ex. 12.1

12.1 Theorem (Strassen 1964). *Let $(B_t)_{t \geq 0}$ be a one-dimensional Brownian motion and set $Z_s(t, \omega) := B(st, \omega)/\sqrt{2s \log \log s}$. For almost all $\omega \in \Omega$,*

$$\{Z_s(\cdot, \omega) : s > e\}$$

is a relatively compact subset of $(\mathcal{C}_{(0)}[0, 1], \|\cdot\|_\infty)$, and the set of all limit points (as $s \rightarrow \infty$) is given by

Ex. 12.2

$$\mathcal{K} = \left\{ w \in \mathcal{C}_{(0)}[0, 1] : w \text{ is absolutely continuous and } \int_0^1 |w'(t)|^2 dt \leq 1 \right\}.$$

Using the Cauchy-Schwarz inequality it is not hard to see that every $w \in \mathcal{K}$ satisfies $|w(t)| \leq \sqrt{t}$ for all $0 \leq t \leq 1$.

Theorem 12.1 contains two statements: With probability one,

- a) the family $\{[0, 1] \ni t \mapsto B(st, \omega)/\sqrt{2s \log \log s} : s > e\} \subset \mathcal{C}_{(o)}[0, 1]$ contains a sequence which converges uniformly for all $t \in [0, 1]$ to a function in the set \mathcal{K} ;
- b) every function from the (non-random) set \mathcal{K} is the uniform limit of a suitable sequence $(B(s_n t, \omega)/\sqrt{2s_n \log \log s_n})_{n \geq 1}$.

For a limit point w_0 it is necessary that the probability

$$\mathbb{P}(\|Z_s(\cdot) - w_0(\cdot)\|_\infty < \delta) \quad \text{as } s \rightarrow \infty$$

is sufficiently large for any $\delta > 0$. From the scaling property of a Brownian motion we see that

$$Z_s(\cdot) - w_0(\cdot) \sim \frac{B(\cdot)}{\sqrt{2 \log \log s}} - w_0(\cdot).$$

This means that we have to calculate probabilities involving a Brownian motion with a nonlinear drift, $B(t) - \sqrt{2 \log \log s} w_0(t)$ and estimate it at an exponential scale, as we will see later. Both topics are interesting in their own right, and we will treat them separately in the following two sections. After that we return to our original problem, the proof of Strassen's Theorem 12.1.

12.1 The Cameron–Martin formula

In Chapter 4 we have introduced the Wiener space $(\mathcal{C}_{(o)}, \mathcal{B}(\mathcal{C}_{(o)}), \mu)$ as canonical probability space of a Brownian motion indexed by $[0, \infty)$. In the same way we can study a Brownian motion with index set $[0, 1]$ and the Wiener space $(\mathcal{C}_{(o)}[0, 1], \mathcal{B}(\mathcal{C}_{(o)}[0, 1]), \mu)$ where $\mathcal{C}_{(o)} = \mathcal{C}_{(o)}[0, 1]$ denotes the space of continuous functions $w : [0, 1] \rightarrow \mathbb{R}$ such that $w(0) = 0$. The Cameron–Martin theorem is about the absolute continuity of the Wiener measure under shifts in the space $\mathcal{C}_{(o)}[0, 1]$. For $w_0 \in \mathcal{C}_{(o)}[0, 1]$ we set

$$Uw(t) := w(t) + w_0(t), \quad t \in [0, 1], \quad w \in \mathcal{C}_{(o)}[0, 1].$$

Clearly, U maps $\mathcal{C}_{(o)}[0, 1]$ into itself and induces an image measure of the Wiener measure

$$\mu_U(A) = \mu(U^{-1}A) \quad \text{for all } A \in \mathcal{B}(\mathcal{C}_{(o)}[0, 1]);$$

this is equivalent to saying

$$\int_{\mathcal{C}_{(o)}} F(w + w_0) \mu(dw) = \int_{\mathcal{C}_{(o)}} F(Uw) \mu(dw) = \int_{\mathcal{C}_{(o)}} F(w) \mu_U(dw) \quad (12.1)$$

for all bounded continuous functionals $F : \mathcal{C}_{(o)}[0, 1] \rightarrow \mathbb{R}$.

We will see that μ_U is absolutely continuous with respect to μ if the shifts w_0 are absolutely continuous (a. c.) functions such that the (almost everywhere existing) deriva-

tive w'_0 is square integrable. This is the *Cameron–Martin space* which we denote by

$$\mathcal{H}^1 := \left\{ u \in \mathcal{C}_{(0)}[0, 1] : u \text{ is a. c. and } \int_0^1 |u'(t)|^2 dt < \infty \right\}. \quad (12.2)$$

12.2 Lemma. *For $u : [0, 1] \rightarrow \mathbb{R}$ the following assertions are equivalent:*

- a) $u \in \mathcal{H}^1$;
 b) $u \in \mathcal{C}_{(0)}[0, 1]$ and $\sup_{\Pi} \sum_{t_{j-1}, t_j \in \Pi} \frac{|u(t_j) - u(t_{j-1})|^2}{t_j - t_{j-1}} < \infty$ where the supremum is taken over all finite partitions $\Pi = \{0 = t_0 < t_1 < \dots < t_n = 1\}$ of $[0, 1]$.

If a) or b) is true, then $\int_0^1 |u'(t)|^2 dt = \lim_{|\Pi| \rightarrow 0} \sum_{t_{j-1}, t_j \in \Pi} |u(t_j) - u(t_{j-1})|^2 / (t_j - t_{j-1})$.

Proof. a) \Rightarrow b): Assume that $u \in \mathcal{H}^1$ and let $0 = t_0 < t_1 < \dots < t_n = 1$ be any partition of $[0, 1]$. Then

$$\sum_{j=1}^n \left[\frac{u(t_j) - u(t_{j-1})}{t_j - t_{j-1}} \right]^2 (t_j - t_{j-1}) = \int_0^1 f_n(t) dt$$

with

$$f_n(t) = \sum_{j=1}^n \left[\frac{u(t_j) - u(t_{j-1})}{t_j - t_{j-1}} \right]^2 \mathbb{1}_{[t_{j-1}, t_j)}(t).$$

By the Cauchy–Schwarz inequality

$$\begin{aligned} f_n(t) &= \sum_{j=1}^n \left[\frac{1}{t_j - t_{j-1}} \int_{t_{j-1}}^{t_j} u'(s) ds \right]^2 \mathbb{1}_{[t_{j-1}, t_j)}(t) \\ &\leq \sum_{j=1}^n \frac{1}{t_j - t_{j-1}} \int_{t_{j-1}}^{t_j} |u'(s)|^2 ds \mathbb{1}_{[t_{j-1}, t_j)}(t). \end{aligned}$$

If we integrate this inequality with respect to t , we get

$$\int_0^1 f_n(t) dt \leq \sum_{j=1}^n \int_{t_{j-1}}^{t_j} |u'(s)|^2 ds = \int_0^1 |u'(s)|^2 ds,$$

hence $M \leq \int_0^1 |u'(s)|^2 ds < \infty$, and the estimate in b) follows.

b) \Rightarrow a): Conversely, assume that b) holds and denote by M the value of the supremum in b). Let $0 \leq r_1 < s_1 \leq r_2 < s_2 \leq \dots \leq r_m < s_m \leq 1$ be the endpoints of the

intervals (r_k, s_k) , $k = 1, \dots, m$. By the Cauchy-Schwarz inequality

$$\begin{aligned} \sum_{k=1}^m |u(s_k) - u(r_k)| &\leq \left[\sum_{k=1}^m \frac{|u(s_k) - u(r_k)|^2}{s_k - r_k} \right]^{1/2} \left[\sum_{k=1}^m (s_k - r_k) \right]^{1/2} \\ &\leq M \left[\sum_{k=1}^m (s_k - r_k) \right]^{1/2}. \end{aligned}$$

This means that u is absolutely continuous and u' exists a.e. Let f_n be as in the first part of the proof and assume that the underlying partitions $\{0 = t_0 < t_1 < \dots < t_n = 1\}$ have decreasing mesh. Then the sequence $(f_n)_{n \geq 1}$ converges to $(u')^2$ at all points where u' exists. By Fatou's lemma

Ex. 12.3

$$\int_0^1 |u'(t)|^2 dt \leq \liminf_{n \rightarrow \infty} \int_0^1 f_n(t) dt = \liminf_{n \rightarrow \infty} \sum_{j=1}^n \frac{|u(t_j) - u(t_{j-1})|^2}{t_j - t_{j-1}} \leq M.$$

Finally,

$$\overline{\lim}_{n \rightarrow \infty} \int_0^1 f_n(t) dt \leq M \leq \int_0^1 |u'(t)|^2 dt \leq \liminf_{n \rightarrow \infty} \int_0^1 f_n(t) dt. \quad \square$$

As a first step we will prove the absolute continuity of μ_U for shifts from a subset \mathcal{H}_o^1 of \mathcal{H}^1 ,

$$\mathcal{H}_o^1 := \{u \in \mathcal{C}_{(0)}[0, 1] : u \in \mathcal{C}^1[0, 1] \text{ and } u' \text{ has bounded variation}\}. \quad (12.3)$$

12.3 Lemma. *Let $w_0 : [0, 1] \rightarrow \mathbb{R}$ be a function in \mathcal{H}_o^1 , $n \geq 1$ and $t_j = j/n$. Then the linear functionals*

$$G_n(w) := \sum_{j=1}^n (w(t_j) - w(t_{j-1})) \frac{w_0(t_j) - w_0(t_{j-1})}{t_j - t_{j-1}}, \quad n \geq 1, \quad w \in \mathcal{C}_{(0)}[0, 1],$$

are uniformly bounded, i. e. $|G_n(w)| \leq c \|w\|_\infty$ for all $n \geq 1$ and $w \in \mathcal{C}_{(0)}[0, 1]$; moreover, the following Riemann–Stieltjes integral exists:

$$G(w) := \lim_{n \rightarrow \infty} G_n(w) = \int_0^1 w'_0(t) dw(t).$$

Proof. The mean-value theorem shows that

$$G_n(w) = \sum_{j=1}^n (w(t_j) - w(t_{j-1})) \frac{w_0(t_j) - w_0(t_{j-1})}{t_j - t_{j-1}} = \sum_{j=1}^n (w(t_j) - w(t_{j-1})) w'_0(\theta_j)$$

with suitable intermediate points $\theta_j \in [t_{j-1}, t_j]$. It is not hard to see that

$$G_n(w) = w'_0(\theta_{n+1})w(t_n) - \sum_{j=1}^n w(t_j)(w'_0(\theta_{j+1}) - w'_0(\theta_j)) \quad (\theta_{n+1} := 1)$$

From this we conclude that

$$|G_n(w)| \leq (\|w'_0\|_\infty + \text{VAR}_1(w'_0; 1))\|w\|_\infty$$

as well as

$$\lim_{n \rightarrow \infty} G_n(w) = w'_0(1)w(1) - \int_0^1 w(t) dw'_0(t) = \int_0^1 w'_0(t) dw(t). \quad \square$$

We can now state and prove the Cameron–Martin formula for shifts in \mathcal{H}_0^1 .

12.4 Theorem (Cameron, Martin 1944). *Let $(\mathcal{C}_{(0)}[0, 1], \mathcal{B}(\mathcal{C}_{(0)}[0, 1]), \mu)$ be the Wiener space and $Uw := w + w_0$ the shift by $w_0 \in \mathcal{H}_0^1$. Then the image measure $\mu_U(A) := \mu(U^{-1}A)$, $A \in \mathcal{B}(\mathcal{C}_{(0)}[0, 1])$, is absolutely continuous with respect to μ , and for every bounded continuous functional $F : \mathcal{C}_{(0)}[0, 1] \rightarrow \mathbb{R}$ we have*

$$\begin{aligned} \int_{\mathcal{C}_{(0)}[0, 1]} F(w) \mu_U(dw) &= \int_{\mathcal{C}_{(0)}[0, 1]} F(Uw) \mu(dw) \\ &= \int_{\mathcal{C}_{(0)}[0, 1]} F(w) \frac{d\mu_U(w)}{d\mu} \mu(dw) \end{aligned} \quad (12.4)$$

with the Radon–Nikodým density

$$\frac{d\mu_U(w)}{d\mu} = \exp \left(-\frac{1}{2} \int_0^1 |w'_0(t)|^2 dt + \int_0^1 w'_0(t) dw(t) \right). \quad (12.5)$$

Proof. In order to deal with the infinite dimensional integrals appearing in (12.4) we use finite dimensional approximations. Set $t_j = j/n$, $j = 0, 1, \dots, n$, $n \geq 0$, and denote by

$$\Pi_n w(t) := \begin{cases} w(t_j), & \text{if } t = t_j, j = 0, 1, \dots, n, \\ \text{linear,} & \text{for all other } t \end{cases}$$

the piecewise linear approximation of w on the grid t_0, t_1, \dots, t_n . Obviously, we have $\Pi_n w \in \mathcal{C}_{(0)}[0, 1]$ and $\lim_{n \rightarrow \infty} \Pi_n w(t) = w(t)$ uniformly for all t .

Let $F : \mathcal{C}_{(0)}[0, 1] \rightarrow \mathbb{R}$ be a bounded continuous functional. Since $\Pi_n w$ is uniquely determined by $x = (w(t_1), \dots, w(t_n)) \in \mathbb{R}^n$, there exists a bounded continuous function $F_n : \mathbb{R}^n \rightarrow \mathbb{R}$ such that

$$F_n(x) = F(\Pi_n w) \quad \text{where} \quad x_j = w(t_j).$$

Write $x^0 = (w_0(t_1), \dots, w_0(t_n))$ and $x_0 = x_0^0 := 0$. Then

$$\begin{aligned}
 & \int_{\mathcal{C}_{(0)}[0,1]} F(\Pi_n U w) \mu(dw) \\
 &= \mathbb{E} [F_n(B(t_1) + w_0(t_1), \dots, B(t_n) + w_0(t_n))] \\
 &\stackrel{(2.10a)}{=} \frac{1}{(2\pi/n)^{n/2}} \int_{\mathbb{R}^n} F_n(x + x^0) e^{-\frac{n}{2} \sum_{j=1}^n (x_j - x_{j-1})^2} dx \\
 &= \frac{1}{(2\pi/n)^{n/2}} \int_{\mathbb{R}^n} F_n(y) e^{-\frac{n}{2} \sum_{j=1}^n (y_j - y_{j-1} - (x_j^0 - x_{j-1}^0))^2} dy \\
 &= \frac{e^{-\frac{n}{2} \sum_{j=1}^n (x_j^0 - x_{j-1}^0)^2}}{(2\pi/n)^{n/2}} \int_{\mathbb{R}^n} F_n(y) e^{n \sum_{j=1}^n (y_j - y_{j-1})(x_j^0 - x_{j-1}^0)} e^{-\frac{n}{2} \sum_{j=1}^n (y_j - y_{j-1})^2} dy.
 \end{aligned}$$

If we set

$$\begin{aligned}
 G_n(w) &:= \sum_{j=1}^n (w(t_j) - w(t_{j-1})) \frac{w_0(t_j) - w_0(t_{j-1})}{\frac{1}{n}}, \\
 \alpha_n(w_0) &:= \sum_{j=1}^n \frac{(w_0(t_j) - w_0(t_{j-1}))^2}{\frac{1}{n}},
 \end{aligned}$$

we find

$$\int_{\mathcal{C}_{(0)}[0,1]} F(\Pi_n U w) \mu(dw) = e^{-\frac{1}{2} \alpha_n(w_0)} \int_{\mathcal{C}_{(0)}[0,1]} F(\Pi_n w) e^{G_n(w)} \mu(dw). \quad (12.6)$$

Since F is bounded and continuous and $\lim_{n \rightarrow \infty} \Pi_n U w = U w$, the left-hand side converges to $\int_{\mathcal{C}_{(0)}[0,1]} F(U w) \mu(dw)$. By construction, $t_j - t_{j-1} = \frac{1}{n}$; therefore, Lemmas 12.2 and 12.3 show that

$$\lim_{n \rightarrow \infty} \alpha_n(w_0) = \int_0^1 |w'_0(t)|^2 dt \quad \text{and} \quad \lim_{n \rightarrow \infty} G_n(w) = \int_0^1 w'_0(t) dw(t).$$

Moreover, by Lemma 12.3,

$$e^{G_n(w)} \leq e^{c \|w\|_\infty} \leq e^{c \sup_{t \leq 1} w(t)} + e^{-c \inf_{t \leq 1} w(t)}.$$

By the reflection principle, Theorem 6.9, $\sup_{t \leq 1} B(t) \sim -\inf_{t \leq 1} B(t) \sim |B(1)|$. Thus

$$\int_{\mathcal{C}_{(0)}[0,1]} e^{G_n(w)} \mu(dw) \leq \mathbb{E} [e^{c \sup_{t \leq 1} B(t)}] + \mathbb{E} [e^{-c \inf_{t \leq 1} B(t)}] = 2[e^{c|B(1)|}].$$

Therefore, we can use the dominated convergence theorem to deduce that the right-

hand side of (12.6) converges to

$$e^{-\frac{1}{2} \int_0^1 |w'_0(t)|^2 dt} \int_{\mathcal{C}_{(o)}[0,1]} F(\Pi_n w) e^{\int_0^1 w'_0(t) dw(t)} \mu(dw). \quad \square$$

We want to show that Theorem 12.4 remains valid for $w_0 \in \mathcal{H}^1$. This is indeed possible, if we use a stochastic approximation technique. The drawback, however, will be that we do not have any longer an explicit formula for the limit $\lim_{n \rightarrow \infty} G_n(w)$ as a Riemann–Stieltjes integral. The trouble is that the integral

$$\int_0^1 w'_0(t) dw(t) = w(1)w_0(1) - \int_0^1 w(t) dw_0(t)$$

is for *all* integrands $w \in \mathcal{C}_{(o)}[0, 1]$ only defined in the Riemann–Stieltjes sense, if w_0 is of bounded variation, cf. Corollary A.41.

12.1 Paley–Wiener–Zygmund integrals (Paley, Wiener, Zygmund 1933). Denote by $\mathbf{BV}[0, 1]$ the set of all functions of bounded variation on $[0, 1]$. For $\phi \in \mathbf{BV}[0, 1]$ the Riemann–Stieltjes integral

$$G^\phi(w) = \int_0^1 \phi(s) dw(s) = \phi(1)w(1) - \int_0^1 w(s) d\phi(s), \quad w \in \mathcal{C}_{(o)}[0, 1],$$

is well-defined. If we write G^ϕ as the limit of a Riemann–Stieltjes sum, it is not hard to see that, on the Wiener space $(\mathcal{C}_{(o)}[0, 1], \mathcal{B}(\mathcal{C}_{(o)}[0, 1]), \mu)$,

Ex. 12.4

$$G^\phi \text{ is a normal random variable, mean 0, variance } \int_0^1 \phi^2(s) ds; \quad (12.7a)$$

$$\int_{\mathcal{C}_{(o)}[0,1]} G^\phi(w) G^\psi(w) \mu(dw) = \int_0^1 \phi(s) \psi(s) ds \text{ if } \phi, \psi \in \mathbf{BV}[0, 1]. \quad (12.7b)$$

Therefore, the map

$$L^2[0, 1] \supset \mathbf{BV}[0, 1] \ni \phi \mapsto G^\phi \in L^2(\mathcal{C}_{(o)}[0, 1], \mathcal{B}(\mathcal{C}_{(o)}[0, 1]), \mu)$$

is a linear isometry. As $\mathbf{BV}[0, 1]$ is a dense subset of $L^2[0, 1]$, we find for every function $\phi \in L^2[0, 1]$ a sequence $(\phi_n)_{n \geq 1} \subset \mathbf{BV}[0, 1]$ such that L^2 - $\lim_{n \rightarrow \infty} \phi_n = \phi$. In particular, $(\phi_n)_{n \geq 1}$ is a Cauchy sequence in $L^2[0, 1]$ and, therefore, $(G^{\phi_n})_{n \geq 1}$ is a Cauchy sequence in $L^2(\mathcal{C}_{(o)}[0, 1], \mathcal{B}(\mathcal{C}_{(o)}[0, 1]), \mu)$. Thus,

$$G^\phi(w) := \int_0^1 \phi(s) dw(t) := L^2\text{-}\lim_{n \rightarrow \infty} G^{\phi_n}(w), \quad w \in \mathcal{C}_{(o)}[0, 1]$$

defines a linear functional on $\mathcal{C}_{(o)}[0, 1]$. Note that the integral notation is *symbolical* and has *no meaning as a Riemann–Stieltjes integral* unless $\phi \in \mathbf{BV}[0, 1]$. Obviously the properties (12.7) are preserved under L^2 limits. The functional G^ϕ is often called the *Paley–Wiener–Zygmund integral*. It is a special case of the Itô integral which we will encounter in Chapter 14.

Ex. 12.5

Ex. 14.11

12.5 Corollary (Cameron, Martin 1944). *Let $(\mathcal{C}_{(o)}[0, 1], \mathcal{B}(\mathcal{C}_{(o)}[0, 1]), \mu)$ be the Wiener space and $Uw := w + w_0$ the shift by $w_0 \in \mathcal{H}^1$. Then $\mu_U(A) := \mu(U^{-1}A)$, $A \in \mathcal{B}(\mathcal{C}_{(o)}[0, 1])$ is absolutely continuous with respect to μ and for every bounded Borel measurable functional $F : \mathcal{C}_{(o)}[0, 1] \rightarrow \mathbb{R}$ we have*

$$\begin{aligned} \int_{\mathcal{C}_{(o)}[0,1]} F(w) \mu_U(dw) &= \int_{\mathcal{C}_{(o)}[0,1]} F(Uw) \mu(dw) \\ &= \int_{\mathcal{C}_{(o)}[0,1]} F(w) \frac{d\mu_U(w)}{d\mu} \mu(dw) \end{aligned} \quad (12.8)$$

with the Radon-Nikodým density

$$\frac{d\mu_U(w)}{d\mu} = \exp \left(-\frac{1}{2} \int_0^1 |w'_0(t)|^2 dt + \int_0^1 w'_0(t) dw(t) \right) \quad (12.9)$$

where the second integral is a Paley–Wiener–Zygmund integral.

Proof. Let $w_0 \in \mathcal{H}^1$ and $(\phi_n)_{n \geq 0}$ be a sequence in \mathcal{H}^1_0 with $L^2\text{-}\lim_{n \rightarrow \infty} \phi'_n = w'_0$. If F is a continuous functional, Theorem 12.4 shows for all $n \geq 1$

$$\begin{aligned} \int_{\mathcal{C}_{(o)}[0,1]} F(w + \phi_n) \mu(dw) \\ = \int_{\mathcal{C}_{(o)}[0,1]} F(w) \exp \left(-\frac{1}{2} \int_0^1 |\phi'_n(t)|^2 dt + \int_0^1 \phi'_n(t) dw(t) \right) \mu(dw). \end{aligned}$$

It is, therefore, enough to show that the terms on the right-hand side converge as $n \rightarrow \infty$. Set

$$G(w) := \int_0^1 w'_0(s) dw(s) \quad \text{and} \quad G_n(w) := \int_0^1 \phi'_n(s) dw(s).$$

Since G_n and G are normal random variables on the Wiener space, we find for all $c > 0$

$$\begin{aligned} \int_{\mathcal{C}_{(o)}[0,1]} \left| e^{G_n(w)} - e^{G(w)} \right| \mu(dw) \\ = \int_{\mathcal{C}_{(o)}[0,1]} \left| e^{G_n(w)-G(w)} - 1 \right| e^{G(w)} \mu(dw) \\ = \left(\int_{\{|G|>c\}} + \int_{\{|G|\leq c\}} \right) \left| e^{G_n(w)-G(w)} - 1 \right| e^{G(w)} \mu(dw) \\ \leq \left[\int_{\mathcal{C}_{(o)}[0,1]} \left| e^{G_n(w)-G(w)} - 1 \right|^2 \mu(dw) \right]^{1/2} \left[\int_{\{|G|>c\}} e^{2G(w)} \mu(dw) \right]^{1/2} \\ \quad + e^c \int_{\mathcal{C}_{(o)}[0,1]} \left| e^{G_n(w)-G(w)} - 1 \right| \mu(dw) \\ \leq C \left[\int_{|y|>c} e^{2y} e^{-y^2/(2a)} dy \right]^{1/2} + \frac{e^c}{\sqrt{2\pi}} \int_{\mathbb{R}} |e^{\sigma_n y} - 1| e^{-y^2/2} dy \end{aligned}$$

where $\sigma_n^2 = \int_0^1 |\phi'_n(s) - w'_0(s)|^2 ds$ and $a = \int_0^1 |w'_0(s)|^2 ds$. Thus,

$$\lim_{n \rightarrow \infty} \int_{\mathcal{C}_{(o)}[0,1]} \left| e^{G_n(w)} - e^{G(w)} \right| \mu(dw) = 0.$$

Using standard approximation arguments from measure theory, the formula (12.8) is easily extended from continuous F to all bounded measurable F . \square

12.6 Remark. It is useful to rewrite Corollary 12.5 in the language of stochastic processes. Let $(B_t)_{t \geq 0}$ be a BM¹ and define a new stochastic process by

$$X_t(\omega) := B_t(U^{-1}\omega) = B_t(\omega) - w_0(t), \quad w_0 \in \mathcal{H}^1.$$

Then X_t is, with respect to the new probability measure

$$\mathbb{Q}(d\omega) = \exp\left(-\frac{1}{2} \int_0^1 |w'_0(s)|^2 ds + \int_0^1 w'_0(s) dB_s(\omega)\right) \mathbb{P}(d\omega),$$

a Brownian motion. This is a deterministic version of the Girsanov transformation, cf. Section 17.3.

12.2 Large deviations (Schilder's theorem)

The term *large deviations* generally refers to small probabilities on an exponential scale. We are interested in lower and upper (exponential) estimates of the probability that *the whole Brownian path* is in the neighbourhood of a deterministic function $u \neq 0$. Estimates of this type have been studied by Schilder and were further developed by Borovkov, Varadhan, Freidlin and Wentzell; our presentation follows the monograph [67] by Freidlin and Wentzell.

12.7 Definition. Let $u \in \mathcal{C}_{(o)}[0, 1]$. The functional $I : \mathcal{C}_{(o)}[0, 1] \rightarrow [0, \infty]$,

$$I(u) := \begin{cases} \frac{1}{2} \int_0^1 |u'(s)|^2 ds, & \text{if } u \text{ is absolutely continuous} \\ +\infty, & \text{otherwise} \end{cases}$$

is the *action functional* of the Brownian motion $(B_t)_{t \in [0,1]}$.

The action functional I will be the exponent in the large deviation estimates for B_t . For this we need a few properties.

12.8 Lemma. *The action functional $I : \mathcal{C}_{(o)}[0, 1] \rightarrow [0, \infty]$ is lower semicontinuous, i.e. for every sequence $(w_n)_{n \geq 1} \subset \mathcal{C}_{(o)}[0, 1]$ which converges uniformly to $w \in \mathcal{C}_{(o)}[0, 1]$ we have $\liminf_{n \rightarrow \infty} I(w_n) \geq I(w)$.*

Proof. Let $(w_n)_{n \geq 1}$ and w be as in the statement of the lemma. We may assume that $\lim_{n \rightarrow \infty} I(w_n) < \infty$. For any finite partition $\Pi = \{0 = t_0 < t_1 < \dots < t_m = 1\}$ of $[0, 1]$ we have

$$\begin{aligned} \sup_{\Pi} \sum_{t_j \in \Pi} \frac{|w(t_j) - w(t_{j-1})|^2}{t_j - t_{j-1}} &= \sup_{\Pi} \lim_{n \rightarrow \infty} \sum_{t_j \in \Pi} \frac{|w_n(t_j) - w_n(t_{j-1})|^2}{t_j - t_{j-1}} \\ &\leq \lim_{n \rightarrow \infty} \sup_{\Pi} \sum_{t_j \in \Pi} \frac{|w_n(t_j) - w_n(t_{j-1})|^2}{t_j - t_{j-1}}. \end{aligned}$$

Using Lemma 12.2, we see that $I(w) \leq \lim_{n \rightarrow \infty} I(w_n)$. \square

As usual, we write

$$d(w, A) := \inf \{ \|w - u\|_{\infty} : u \in A \}, \quad w \in \mathcal{C}_{(0)}[0, 1] \quad \text{and} \quad A \subset \mathcal{C}_{(0)}[0, 1]$$

for the distance of the point w to the set A . Denote by

$$\Phi(r) := \{w \in \mathcal{C}_{(0)}[0, 1] : I(w) \leq r\}, \quad c \geq 0$$

the *sub-level sets* of the action functional I .

12.9 Lemma. *The sub-level sets $\Phi(r)$, $r \geq 0$, of the action functional I are compact subsets of $\mathcal{C}_{(0)}[0, 1]$.*

Proof. Since I is lower semicontinuous, the sub-level sets are closed. For any $r > 0$, $0 \leq s < t \leq 1$ and $w \in \Phi(r)$ we find using the Cauchy-Schwarz inequality

$$\begin{aligned} |w(t) - w(s)| &= \left| \int_s^t w'(x) dx \right| \leq \left(\int_s^t |w'(x)|^2 dx \right)^{1/2} \sqrt{t-s} \\ &\leq \sqrt{2I(w)} \sqrt{t-s} \leq \sqrt{2r} \sqrt{t-s}. \end{aligned}$$

This shows that the family $\Phi(r)$ is equi-bounded and equicontinuous. By Ascoli's theorem, $\Phi(r)$ is compact. \square

We are now ready for the upper large deviation estimate.

12.10 Theorem (Schilder 1966). *Let $(B_t)_{t \in [0, 1]}$ be a BM¹ and denote by $\Phi(r)$ the sub-level sets of the action functional I . For every $\delta > 0$, $\gamma > 0$ and $r_0 > 0$ there is some $\epsilon_0 = \epsilon_0(\delta, \gamma, r_0)$ such that*

$$\mathbb{P}(d(\epsilon B, \Phi(r_0)) \geq \delta) \leq \exp \left[-\frac{r - \gamma}{\epsilon^2} \right] \quad (12.10)$$

for all $0 < \epsilon \leq \epsilon_0$ and $0 \leq r \leq r_0$.

Proof. Set $t_j = j/n$, $j = 0, 1, \dots, n$, $n \geq 0$, and denote by

$$\Pi_n[\epsilon B](t, \omega) := \begin{cases} \epsilon B(t_j, \omega), & \text{if } t = t_j, j = 0, 1, \dots, n, \\ \text{linear}, & \text{for all other } t, \end{cases}$$

the piecewise linear approximation of the continuous function $\epsilon B_t(\omega)$ on the grid t_0, t_1, \dots, t_n . Obviously, for all $t \in [t_{k-1}, t_k]$ and $k = 1, \dots, n$,

$$\Pi_n[\epsilon B](t) = \epsilon B(t_{k-1}) + (t - t_{k-1}) \frac{\epsilon B(t_k) - \epsilon B(t_{k-1})}{t_k - t_{k-1}}.$$

Since the approximations $\Pi_n[\epsilon B](t)$ are absolutely continuous, we can calculate the action functional

$$I(\Pi_n[\epsilon B]) = \frac{1}{2} \sum_{k=1}^n n \epsilon^2 (B(t_k) - B(t_{k-1}))^2 = \frac{\epsilon^2}{2} \sum_{k=1}^n \xi_k^2 \quad (12.11)$$

where $\xi_k = \sqrt{n}(B(t_k) - B(t_{k-1})) \sim \mathbf{N}(0, 1)$ are iid standard normal random variables. Note that

$$\begin{aligned} \{d(\epsilon B, \Phi(r)) \geq \delta\} &= \{d(\epsilon B, \Phi(r)) \geq \delta, d(\Pi_n[\epsilon B], \epsilon B) < \delta\} \\ &\quad \cup \{d(\epsilon B, \Phi(r)) \geq \delta, d(\Pi_n[\epsilon B], \epsilon B) \geq \delta\} \\ &\subset \underbrace{\{\Pi_n[\epsilon B] \notin \Phi(r)\}}_{=: A_n} \cup \underbrace{\{d(\Pi_n[\epsilon B], \epsilon B) \geq \delta\}}_{=: C_n}. \end{aligned}$$

We will estimate $\mathbb{P}(A_n)$ and $\mathbb{P}(C_n)$ separately. Combining (12.11) with Chebychev's inequality gives for all $0 < \alpha < 1$

$$\begin{aligned} \mathbb{P}(A_n) &= \mathbb{P}(I(\Pi_n[\epsilon B]) > r) = \mathbb{P}\left(\frac{1}{2} \sum_{k=1}^n \xi_k^2 > \frac{r}{\epsilon^2}\right) \\ &= \mathbb{P}\left(\exp\left[\frac{1-\alpha}{2} \sum_{k=1}^n \xi_k^2\right] > \exp\left[\frac{r}{\epsilon^2} (1-\alpha)\right]\right) \\ &\leq \exp\left[-\frac{r}{\epsilon^2} (1-\alpha)\right] \underbrace{\left(\mathbb{E} \exp\left[\frac{1-\alpha}{2} \xi_1^2\right]\right)^n}_{=e^{n(1-\alpha)^2/8} \text{ by (2.6)}}. \end{aligned}$$

Therefore, we find for all $n \geq 1$, $r_0 > 0$ and $\gamma > 0$ some $\epsilon_0 > 0$ such that

$$\mathbb{P}(A_n) \leq \frac{1}{2} \exp\left[-\frac{r}{\epsilon^2} + \frac{\gamma}{\epsilon^2}\right] \quad \text{for all } 0 < \epsilon \leq \epsilon_0, 0 \leq r \leq r_0.$$

In order to estimate $\mathbb{P}(C_n)$ we use the fact that Brownian motion has stationary increments.

$$\begin{aligned}
\mathbb{P}(C_n) &= \mathbb{P}\left(\sup_{0 \leq t \leq 1} |\epsilon B(t) - \Pi_n[\epsilon B](t)| \geq \delta\right) \\
&\leq \sum_{k=1}^n \mathbb{P}\left(\sup_{t_{k-1} \leq t \leq t_k} |\epsilon B(t) - \epsilon B(t_{k-1}) - (t - t_{k-1})n\epsilon(B(t_k) - B(t_{k-1}))| \geq \delta\right) \\
&= n \mathbb{P}\left(\sup_{t \leq 1/n} |\epsilon B(t) - n\epsilon t B(1/n)| \geq \delta\right) \\
&\leq n \mathbb{P}\left(\sup_{t \leq 1/n} |B(t)| + |B(1/n)| \geq \frac{\delta}{\epsilon}\right) \\
&\leq n \mathbb{P}\left(\sup_{t \leq 1/n} |B(t)| \geq \frac{\delta}{2\epsilon}\right) \\
&\leq 2n \mathbb{P}\left(\sup_{t \leq 1/n} B(t) \geq \frac{\delta}{2\epsilon}\right).
\end{aligned}$$

We can now use the following maximal estimate which is, for example, a consequence of the reflection principle, Theorem 6.9, and the Gaussian tail estimate from Lemma 10.5

$$\mathbb{P}\left(\sup_{s \leq t} B(s) \geq x\right) \leq 2 \mathbb{P}(B(t) \geq x) \leq \frac{2}{\sqrt{2\pi t}} \frac{1}{x} e^{-x^2/(2t)} \quad x > 0, t > 0,$$

to deduce

$$\mathbb{P}(C_n) \leq 4n \mathbb{P}\left(B(1/n) \geq \frac{\delta}{2\epsilon}\right) \leq 4n \frac{2\epsilon}{\sqrt{2\pi/n} \delta} \exp\left[-\frac{\delta^2 n}{8\epsilon^2}\right].$$

Finally, pick $n \geq 1$ such that $\delta^2 n/8 > r_0$ and ϵ_0 small enough such that $\mathbb{P}(C_n)$ is less than $\frac{1}{2} \exp\left[-\frac{r}{\epsilon^2} + \frac{\gamma}{\epsilon^2}\right]$ for all $0 < \epsilon \leq \epsilon_0$ and $0 \leq r \leq r_0$. \square

12.11 Corollary. *Let $(B_t)_{t \in [0,1]}$ be a BM¹ and I the action functional. Then*

$$\overline{\lim}_{\epsilon \rightarrow 0} \epsilon^2 \log \mathbb{P}(\epsilon B \in F) \leq - \inf_{w \in F} I(w) \quad (12.12)$$

for every closed set $F \subset \mathcal{C}_0[0, 1]$.

Proof. By the definition of the sub-level set $\Phi(r)$ we have $\Phi(r) \cap F = \emptyset$ for all $r < \inf_{w \in F} I(w)$. Since $\Phi(r)$ is compact,

$$d(\Phi(r), F) = \inf_{w \in \Phi(r)} d(w, F) =: \delta_r > 0.$$

From Theorem 12.10 we know for all $0 < \epsilon \leq \epsilon_0$ that

$$\mathbb{P}[\epsilon B \in F] \leq \mathbb{P}[d(\epsilon B, \Phi(r)) > \delta_r] \leq \exp\left[-\frac{r - \gamma}{\epsilon^2}\right],$$

and so

$$\lim_{\epsilon \rightarrow 0} \epsilon^2 \log \mathbb{P}(\epsilon B \in F) \leq -r.$$

Since $r < \inf_{w \in F} I(w)$ is arbitrary, the claim follows. \square

Next we prove the lower large deviation inequality.

12.12 Theorem (Schilder 1966). *Let $(B_t)_{t \in [0,1]}$ be a BM¹ and I the action functional. For every $\delta > 0$, $\gamma > 0$ and $c > 0$ there is some $\epsilon_0 = \epsilon_0(\delta, \gamma, c)$ such that*

$$\mathbb{P}(\|\epsilon B - w_0\|_\infty < \delta) > \exp(-(I(w_0) + \gamma)/\epsilon^2) \quad (12.13)$$

for all $0 < \epsilon \leq \epsilon_0$ and $w_0 \in \mathcal{C}_{(0)}[0, 1]$ with $I(w_0) \leq c$.

Proof. By the Cameron–Martin formula (12.8), (12.9) we see that

$$\mathbb{P}(\|\epsilon B - w_0\|_\infty < \delta) = e^{-I(w_0)/\epsilon^2} \int_A e^{-\frac{1}{\epsilon} \int_0^1 w'_0(s) dw(s)} \mu(dw)$$

where $A := \{w \in \mathcal{C}_{(0)}[0, 1] : \|w\|_\infty < \delta/\epsilon\}$.

Let $C := \{w \in \mathcal{C}_{(0)}[0, 1] : \int_0^1 w'_0(s) dw(s) \leq 2\sqrt{2I(w_0)}\}$. Then

$$\begin{aligned} \int_A e^{-\frac{1}{\epsilon} \int_0^1 w'_0(s) dw(s)} \mu(dw) &\geq \int_{A \cap C} e^{-\frac{1}{\epsilon} \int_0^1 w'_0(s) dw(s)} \mu(dw) \\ &\geq e^{-\frac{2}{\epsilon} \sqrt{2I(w_0)}} \mu(A \cap C). \end{aligned}$$

Pick ϵ_0 so small that

$$\mathbb{P}(A) = \mathbb{P}\left(\sup_{t \leq 1} |B(t)| < \frac{\delta}{\epsilon}\right) \geq \frac{3}{4} \quad \text{for all } 0 < \epsilon \leq \epsilon_0.$$

Now we can use Chebyshev's inequality to find

$$\begin{aligned} \mu(\Omega \setminus C) &= \mu\left\{w \in \mathcal{C}_{(0)}[0, 1] : \int_0^1 w'_0(s) dw(s) > 2\sqrt{2I(w_0)}\right\} \\ &\leq \frac{1}{8I(w_0)} \int_{\mathcal{C}_{(0)}[0, 1]} \left[\int_0^1 w'_0(s) dw(s)\right]^2 \mu(dw) \\ &\stackrel{(12.7a)}{=} \frac{1}{8I(w_0)} \int_0^1 |w'_0(s)|^2 ds = \frac{1}{4}. \end{aligned}$$

Thus, $\mu(C) \geq \frac{3}{4}$ and, therefore, $\mu(A \cap C) = \mu(A) + \mu(C) - \mu(A \cup C) \geq 1/2$. If we make ϵ_0 even smaller, we can make sure that

$$\frac{1}{2} \exp\left[-\frac{2}{\epsilon} \sqrt{2I(w_0)}\right] \geq \exp\left[-\frac{\gamma}{\epsilon^2}\right]$$

for all $0 < \epsilon \leq \epsilon_0$ and w_0 with $I(w_0) \leq c$. \square

12.13 Corollary. Let $(B_t)_{t \in [0,1]}$ be a BM^1 and I the action functional. Then

$$\lim_{\epsilon \rightarrow 0} \epsilon^2 \log \mathbb{P}(\epsilon B \in U) \geq - \inf_{w \in U} I(w) \quad (12.14)$$

for every open set $U \subset \mathcal{C}_{(0)}[0, 1]$.

Proof. For every $w_0 \in U$ there is some $\delta_0 > 0$ such that

$$\{w \in \mathcal{C}_{(0)}[0, 1] : \|w - w_0\|_\infty < \delta_0\} \subset U.$$

Theorem 12.12 shows that

$$\mathbb{P}(\epsilon B \in U) \geq \mathbb{P}(\|\epsilon B - w_0\| < \delta_0) \geq \exp\left(-\frac{1}{\epsilon^2} (I(w_0) + \gamma)\right),$$

and so

$$\lim_{\epsilon \rightarrow 0} \epsilon^2 \log \mathbb{P}(\epsilon B \in U) \geq -I(w_0);$$

since $w_0 \in U$ is arbitrary, the claim follows. \square

12.3 The proof of Strassen's theorem

Before we start with the rigorous argument we want to add

12.14 Some heuristic considerations. Let $(B_t)_{t \geq 0}$ be a BM^1 and

$$Z_s(t, \omega) = \frac{B(st, \omega)}{\sqrt{2s \log \log s}} \sim \frac{B(t, \omega)}{\sqrt{2 \log \log s}}, \quad t \in [0, 1], s > e.$$

In order to decide whether $w_0 \in \mathcal{C}_{(0)}[0, 1]$ is a limit point of $\{Z_s : s > e\}$, we have to estimate the probability

$$\mathbb{P}(\|Z_s(\cdot) - w_0(\cdot)\|_\infty < \delta) \quad \text{as } s \rightarrow \infty$$

for all $\delta > 0$. Using the canonical version of the Brownian motion on the Wiener space $(\mathcal{C}_{(0)}[0, 1], \mathcal{B}(\mathcal{C}_{(0)}[0, 1]), \mu)$ we see for $\epsilon := 1/\sqrt{2 \log \log s}$.

$$\begin{aligned} \mathbb{P}(\|Z_s(\cdot) - w_0(\cdot)\|_\infty < \delta) &= \mathbb{P}(\|\epsilon B(\cdot) - w_0(\cdot)\|_\infty < \delta) \\ &= \mu\{w \in \mathcal{C}_{(0)}[0, 1] : \|w - \epsilon^{-1} w_0\|_\infty < \delta/\epsilon\}. \end{aligned}$$

Now assume that $w_0 \in \mathcal{H}^1$ and set $A := \{w \in \mathcal{C}_{(0)}[0, 1] : \|w\|_\infty < \delta/\epsilon\}$. The Cameron–Martin formula (12.8), (12.9) shows

$$\begin{aligned} &\mathbb{P}(\|Z_s(\cdot) - w_0(\cdot)\|_\infty < \delta) \\ &= \exp\left(-\frac{1}{2\epsilon^2} \int_0^1 |w'_0(t)|^2 dt\right) \int_A \exp\left(-\frac{1}{\epsilon} \int_0^1 w'_0(s) dw(s)\right) \mu(dw). \end{aligned}$$

Using the large deviation estimates from Corollaries 12.11 and 12.13 we see that the first term on the right-hand side is the dominating factor as $\epsilon \rightarrow 0$ (i. e. $s \rightarrow \infty$) and that

$$\mathbb{P}(\|Z_s(\cdot) - w_0(\cdot)\|_\infty < \delta) \approx (\log s)^{-\int_0^1 |w'_0(t)|^2 dt} \quad \text{as } s \rightarrow \infty. \quad (12.15)$$

Thus, the probability on the left is only larger than zero if $\int_0^1 |w'_0(t)|^2 dt < \infty$, i. e. if $w_0 \in \mathcal{H}^1$. Consider now, as in the proof of Khintchine's LIL, sequences of the form $s = s_n = q^n$, $n \geq 1$, where $q > 1$. Then (12.15) shows that

$$\sum_n \mathbb{P}(\|Z_{s_n}(\cdot) - w_0(\cdot)\|_\infty < \delta) = +\infty \quad \text{or} \quad < +\infty$$

according to $\int_0^1 |w'_0(t)|^2 dt \leq 1$ or $\int_0^1 |w'_0(t)|^2 dt > 1$, respectively. By the Borel–Cantelli lemma we infer that the family

$$\mathcal{K} = \left\{ w \in \mathcal{H}^1 : \int_0^1 |w'(t)|^2 dt \leq 1 \right\}$$

is indeed a good candidate for the set of limit points. \square

We will now give a rigorous proof of Theorem 12.1. As in Section 12.2 we write $I(u)$ for the action functional and $\Phi(r) = \{w \in \mathcal{C}_{(0)}[0, 1] : I(w) \leq r\}$ for its sub-level sets; with this notation $\mathcal{K} = \Phi(\frac{1}{2})$. Finally, set

$$\mathcal{L}(\omega) := \{w \in \mathcal{C}_{(0)}[0, 1] : w \text{ is a limit point of } \{Z_s(\cdot, \omega) : s > e\} \text{ as } s \rightarrow \infty\}.$$

We have to show that $\mathcal{L}(\omega) = \Phi(\frac{1}{2})$. We split the argument in a series of lemmas.

12.15 Lemma. $\mathcal{L}(\omega) \subset \Phi(\frac{1}{2})$ for almost all $\omega \in \Omega$.

Proof. Since $\Phi(\frac{1}{2}) = \bigcap_{\eta > 0} \Phi(\frac{1}{2} + \eta)$ it is clearly enough to show that

$$\mathcal{L}(\omega) \subset \Phi(\frac{1}{2} + \eta) \quad \text{for all } \eta > 0.$$

1° Consider the sequence $(Z_{q^n}(\cdot, \omega))_{n \geq 1}$ for some $q > 1$. From Theorem 12.10 we find for any $\delta > 0$ and $\gamma < \eta$

$$\begin{aligned} \mathbb{P}[d(Z_{q^n}, \Phi(\frac{1}{2} + \eta)) > \delta] &= \mathbb{P}[d((2 \log \log q^n)^{-1/2} B, \Phi(\frac{1}{2} + \eta)) > \delta] \\ &\leq \exp[-2(\log \log q^n)(\frac{1}{2} + \eta - \gamma)]. \end{aligned}$$

Consequently,

$$\sum_n \mathbb{P}[d(Z_{q^n}, \Phi(\frac{1}{2} + \eta)) > \delta] < \infty,$$

and the Borel–Cantelli lemma implies that, for almost all $\omega \in \Omega$,

$$d(Z_{q^n}(\cdot, \omega), \Phi(\frac{1}{2} + \eta)) \leq \delta \quad \text{for all } n \geq n_0(\omega).$$

2° Let us now fill the gaps in the sequence $(q^n)_{n \geq 1}$. Set $L(t) = \sqrt{2t \log \log t}$. Then

$$\begin{aligned} \sup_{q^{n-1} \leq t \leq q^n} \|Z_t - Z_{q^n}\|_\infty &= \sup_{q^{n-1} \leq t \leq q^n} \sup_{0 \leq r \leq 1} \left| \frac{B(rt)}{L(t)} - \frac{B(rq^n)}{L(q^n)} \right| \\ &\leq \underbrace{\sup_{q^{n-1} \leq t \leq q^n} \sup_{0 \leq r \leq 1} \frac{|B(rt) - B(rq^n)|}{L(q^n)}}_{=: X_n} + \underbrace{\sup_{q^{n-1} \leq t \leq q^n} \sup_{0 \leq r \leq 1} \frac{|B(rt)|}{L(q^n)} \left(\frac{L(q^n)}{L(t)} - 1 \right)}_{=: Y_n}. \end{aligned}$$

Since $L(t)$, $t > 3$, is increasing, we see that for sufficiently large $n \geq 1$

$$Y_n \leq \sup_{s \leq q^n} \frac{|B(s)|}{L(q^n)} \left(\frac{L(q^n)}{L(q^{n-1})} - 1 \right).$$

Khintchine's LIL, Theorem 11.1, shows that the first factor is almost surely bounded for all $n \geq n_0(\omega)$. Moreover, $\lim_{n \rightarrow \infty} L(q^n)/L(q^{n-1}) = 1$, which means that we have

$$Y_n \leq \frac{\delta}{2} \quad \text{for all } n \geq n_0(\omega, \delta, q)$$

with $\delta > 0$ as in step 1°. For the estimate of X_n we use the scaling property of a Brownian motion to get

$$\begin{aligned} \mathbb{P} \left(X_n > \frac{\delta}{2} \right) &= \mathbb{P} \left(\sup_{q^{n-1} \leq t \leq q^n} \sup_{0 \leq r \leq 1} \left| B\left(\frac{t}{q^n} r\right) - B(r) \right| > \frac{\delta}{2} \sqrt{2 \log \log q^n} \right) \\ &\leq \mathbb{P} \left(\sup_{q^{-1} \leq c \leq 1} \sup_{0 \leq r \leq 1} \left| B(cr) - B(r) \right| > \frac{\delta}{2} \sqrt{2 \log \log q^n} \right) \\ &= \mathbb{P}(\gamma B \in F) \end{aligned}$$

where $\gamma := 2\delta^{-1}(2 \log \log q^n)^{-1/2}$ and

$$F := \left\{ w \in \mathcal{C}_{(0)}[0, 1] : \sup_{q^{-1} \leq c \leq 1} \sup_{0 \leq r \leq 1} |w(cr) - w(r)| \geq 1 \right\}.$$

Ex. 12.6 Since $F \subset \mathcal{C}_{(0)}[0, 1]$ is a closed subset, we can use Corollary 12.11 and find

$$\mathbb{P} \left(X_n > \frac{\delta}{2} \right) \leq \exp \left(-\frac{1}{\gamma^2} \left(\inf_{w \in F} I(w) - h \right) \right)$$

for any $h > 0$. As $\inf_{w \in F} I(w) = q/(2q - 2)$, cf. Lemma 12.16 below, we get

$$\mathbb{P} \left(X_n > \frac{\delta}{2} \right) \leq \exp \left(-\frac{\delta^2}{4} \log \log q^n \left(\frac{q}{q-1} - 2h \right) \right)$$

and, therefore, $\sum_n \mathbb{P}[X_n > \delta/2] < \infty$ if q is close to 1. Now the Borel–Cantelli lemma implies that $X_n \leq \delta/2$ almost surely if $n \geq n_0(\omega)$ and so

$$\sup_{q^{n-1} \leq s \leq q^n} \|Z_s(\cdot, \omega) - Z_{q^n}(\cdot, \omega)\|_\infty \leq \delta \quad \text{for all } n \geq n_0(\omega, \delta, q).$$

3° Steps 1° and 2° show that for every $\delta > 0$ there is some $q > 1$ such that there is for almost all $\omega \in \Omega$ some $s_0 = s_0(\omega) = q^{n_0(\omega)+1}$ with

$$\begin{aligned} d(Z_s(\cdot, \omega), \Phi(\tfrac{1}{2} + \eta)) \\ \leq \|Z_s(\cdot, \omega) - Z_{q^n}(\cdot, \omega)\|_\infty + d(Z_{q^n}(\cdot, \omega), \Phi(\tfrac{1}{2} + \eta)) \leq 2\delta \end{aligned}$$

for all $s \geq s_0$. Since $\delta > 0$ is arbitrary and $\Phi(\frac{1}{2} + \eta)$ is closed, we get

$$L(\omega) \subset \overline{\Phi(\tfrac{1}{2} + \eta)} = \Phi(\tfrac{1}{2} + \eta). \quad \square$$

12.16 Lemma. Let $F = \left\{ w \in \mathcal{C}_{(0)}[0, 1] : \sup_{q^{-1} \leq c \leq 1} \sup_{0 \leq r \leq 1} |w(cr) - w(r)| \geq 1 \right\}$, $q > 1$, and I be the action functional of a Brownian motion. Then

$$\inf_{w \in F} I(w) = \frac{1}{2} \frac{q}{q-1}.$$

Proof. The function $w_0(t) := \frac{qt-1}{q-1} \mathbb{1}_{[1/q, 1]}(t)$ is in F , and

$$2I(w_0) = \int_{1/q}^1 \left(\frac{q}{q-1} \right)^2 dt = \frac{q}{q-1}.$$

Therefore, $2 \inf_{w \in F} I(w) \leq 2I(w_0) = q/(q-1)$. On the other hand, for every $w \in F$ there are $x \in [0, 1]$ and $c > q^{-1}$ such that

$$\begin{aligned} 1 \leq |w(x) - w(cx)|^2 &= \left| \int_{cx}^x w'(s) ds \right|^2 \leq (x - cx) \int_0^1 |w'(s)|^2 ds \\ &\leq 2(1 - q^{-1})I(w). \end{aligned} \quad \square$$

12.17 Corollary. The set $\{Z_s(\cdot, \omega) : s > e\}$ is for almost all $\omega \in \Omega$ a relatively compact subset of $\mathcal{C}_{(0)}[0, 1]$.

Proof. Let $(s_j)_{j \geq 1}$ be a sequence in $(0, \infty)$. Without loss of generality we can assume that $(s_j)_{j \geq 1}$ is unbounded. Pick a further sequence $(\delta_n)_{n \geq 1} \subset (0, \infty)$ such that $\lim_{n \rightarrow \infty} \delta_n = 0$. Because of step 3° in the proof of Lemma 12.15 there exists for every $n \geq 1$ some $w_n \in \Phi(\frac{1}{2} + \eta)$ and an index $j(n)$ such that $\|Z_{s_{j(n)}}(\cdot, \omega) - w_n\|_\infty \leq \delta_n$. Since $\Phi(\frac{1}{2} + \eta)$ is a compact set, there is a subsequence $(s'_{j(n)})_{j \geq 1} \subset (s_{j(n)})_{n \geq 1}$ such that $(Z_{s'_{j(n)}}(\cdot, \omega))_{j \geq 1}$ converges in $\mathcal{C}_{(0)}[0, 1]$. \square

12.18 Lemma. $\mathcal{L}(\omega) \supset \Phi(\frac{1}{2})$ for almost all $\omega \in \Omega$.

Proof. **1°** The sub-level sets $\Phi(r)$ are compact, cf. Lemma 12.9, and $\Phi(r) \subset \Phi(r')$ if $r \leq r'$; it follows that $\overline{\bigcup_{r < 1/2} \Phi(r)} = \Phi(\frac{1}{2})$. By the very definition, the set of limit points $\mathcal{L}(\omega)$ is a closed set. Therefore, it is enough to show that $\Phi(r) \subset \mathcal{L}(\omega)$ for all $r < 1/2$.

Being subsets of $\mathcal{C}_{(0)}[0, 1]$, the sub-level sets $\Phi(r)$ contain a countable dense subset. This means that it is enough to show that for every $w \in \Phi(r)$, every rational $r < 1/2$ and every $\epsilon > 0$ there is a sequence $(s_n)_{n \geq 1}$, $s_n \rightarrow \infty$, such that

$$\mathbb{P} \left[\|Z_{s_n}(\cdot) - w(\cdot)\|_\infty < \epsilon, \quad \text{for all } n \geq 1 \right] = 1.$$

2° As in the proof of Khintchine's LIL we set $s_n = q^n$, $n \geq 1$, for some $q > 1$, and $L(s) = \sqrt{2s \log \log s}$. Then

$$\begin{aligned} \|Z_{s_n}(\cdot) - w(\cdot)\|_\infty &= \sup_{0 \leq t \leq 1} \left| \frac{B(ts_n)}{L(s_n)} - w(t) \right| \\ &\leq \sup_{q^{-1} \leq t \leq 1} \left| \frac{B(ts_n) - B(s_{n-1})}{L(s_n)} - w(t) \right| \\ &\quad + \frac{|B(s_{n-1})|}{L(s_n)} + \sup_{t \leq q^{-1}} |w(t)| + \sup_{t \leq q^{-1}} \frac{|B(ts_n)|}{L(s_n)}. \end{aligned}$$

We will estimate the terms separately.

Since Brownian motion has independent increments, we see that the sets

$$A_n = \left\{ \sup_{q^{-1} \leq t \leq 1} \left| \frac{B(ts_n) - B(s_{n-1})}{L(s_n)} - w(t) \right| < \frac{\epsilon}{4} \right\}, \quad n \geq 1,$$

are independent. Combining the stationarity of the increments, the scaling property 2.12 of a Brownian motion and Theorem 12.12 yields for all $n \geq n_0$

$$\begin{aligned} \mathbb{P}(A_n) &= \mathbb{P} \left(\sup_{q^{-1} \leq t \leq 1} \left| \frac{B(t - q^{-1})}{\sqrt{2 \log \log s_n}} - w(t) \right| < \frac{\epsilon}{4} \right) \\ &\geq \mathbb{P} \left(\sup_{0 \leq t \leq 1 - q^{-1}} \left| \frac{B(t)}{\sqrt{2 \log \log s_n}} - v(t) \right| < \frac{\epsilon}{8} \right) \\ &\geq \mathbb{P} \left(\sup_{0 \leq t \leq 1} \left| \frac{B(t)}{\sqrt{2 \log \log s_n}} - v(t) \right| < \frac{\epsilon}{8} \right) \\ &\geq \exp(-2 \log \log s_n \cdot (I(v) + \eta)) \\ &= \left(\frac{1}{n \log q} \right)^{2(I(v) + \eta)}; \end{aligned}$$

here we choose $q > 0$ so large that $|w(q^{-1})| \leq \epsilon/8$, and use

$$v(t) = \begin{cases} w(t + \frac{1}{q}) - w(\frac{1}{q}), & \text{if } 0 \leq t \leq 1 - \frac{1}{q}, \\ w(1), & \text{if } 1 - \frac{1}{q} \leq t \leq 1. \end{cases}$$

Since $\eta > 0$ is arbitrary and

$$I(v) = \int_0^{1-1/q} |v'(t)|^2 dt = \int_{1/q}^1 |w'(s)|^2 ds \leq \int_0^1 |w'(s)|^2 ds \leq r < \frac{1}{2},$$

we find $\sum_n \mathbb{P}(A_n) = \infty$. The Borel–Cantelli lemma implies that for any $q > 1$

$$\overline{\lim}_{n \rightarrow \infty} \sup_{q^{-1} \leq t \leq 1} \left| \frac{B(ts_n) - B(s_{n-1})}{\sqrt{2 \log \log s_n}} - w(t) \right| \leq \frac{\epsilon}{4}.$$

3° Finally,

Ex. 12.7

$$\begin{aligned} \mathbb{P} \left(\frac{|B(s_{n-1})|}{L(s_n)} > \frac{\epsilon}{4} \right) &= \mathbb{P} \left(|B(1)| > \frac{\epsilon}{4} \sqrt{2q \log \log q^n} \right), \\ \sup_{t \leq q^{-1}} |w(t)| &\leq \int_0^{1/q} |w'(s)| ds \leq \frac{\sqrt{r}}{\sqrt{q}} < \frac{1}{\sqrt{q}}, \\ \mathbb{P} \left(\sup_{0 \leq t \leq q^{-1}} \frac{|B(ts_n)|}{L(s_n)} > \frac{\epsilon}{4} \right) &\leq 2 \mathbb{P} \left(|B(1)| > \frac{\epsilon}{4} \sqrt{2q \log \log q^n} \right), \end{aligned}$$

and the Borel–Cantelli lemma shows that for sufficiently large $q > 1$

$$\overline{\lim}_{n \rightarrow \infty} \left(\frac{|B(s_{n-1})|}{L(s_n)} + \sup_{t \leq q^{-1}} |w(t)| + \sup_{t \leq q^{-1}} \frac{|W(ts_n)|}{L(s_n)} \right) < \frac{3}{4} \epsilon.$$

This finishes the proof. \square

12.19 Further reading. The Cameron–Martin formula is the basis for many further developments in the analysis in Wiener and abstract Wiener spaces, see the *Further reading* section in Chapter 4; worth reading is the early account in [187]. There is an extensive body of literature on the topic of large deviations. A good starting point are the lecture notes [182] to get an idea of the topics and methods behind ‘large deviations’. Since the topic is quite technical, it is very difficult to find ‘easy’ introductions. Our favourites are [35] and [67] (which we use here), a standard reference is the classic monograph [36].



[35] Dembo, Zeitouni: *Large Deviations Techniques and Applications*.

[36] Deuschel, Stroock: *Large Deviations*.

[67] Freidlin, Wentzell: *Random Perturbations of Dynamical Systems*.

[182] Varadhan: *Large Deviations and Applications*.

[187] Wiener et al.: *Differential Space, Quantum Systems, and Prediction*.

Problems

1. Let $w \in \mathcal{C}_{(0)}[0, 1]$ and assume that for every fixed $t \in [0, 1]$ the number $w(t)$ is a limit point of the family $\{Z_s(t) : s > e\} \subset \mathbb{R}$. Show that this is not sufficient for w to be a limit point of $\mathcal{C}_{(0)}[0, 1]$.
2. Let \mathcal{K} be the set from Theorem 12.1. Show that for $w \in \mathcal{K}$ the estimate $|w(t)| \leq \sqrt{t}$, $t \in [0, 1]$ holds.
3. Let $u \in \mathcal{H}^1$ and Π_n , $n \geq 1$, be a sequence of partitions of $[0, 1]$ such that $\lim_{n \rightarrow \infty} |\Pi_n| = 0$. Show that the functions

$$f_n(t) = \sum_{t_j, t_{j-1} \in \Pi_n} \left[\frac{1}{t_j - t_{j-1}} \int_{t_{j-1}}^{t_j} u'(s) ds \right]^2 \mathbb{1}_{[t_{j-1}, t_j)}(t), \quad n \geq 1,$$

converge for $n \rightarrow \infty$ to $[u'(t)]^2$ for Lebesgue almost all t .

4. Let $\phi \in \mathbf{BV}[0, 1]$ and consider the following Riemann-Stieltjes integral

$$G^\phi(w) = \phi(1)w(1) - \int_0^1 w(s) d\phi(s), \quad w \in \mathcal{C}_{(0)}[0, 1].$$

Show that G^ϕ is a linear functional on $(\mathcal{C}_{(0)}[0, 1], \mathcal{B}(\mathcal{C}_{(0)}[0, 1]), \mu)$ satisfying

- (a) G^ϕ is a normal random variable, mean 0, variance $\int_0^1 \phi^2(s) ds$;
 - (b) $\int_{\mathcal{C}_{(0)}[0, 1]} G^\phi(w) G^\psi(w) \mu(dw) = \int_0^1 \phi(s) \psi(s) ds$ for all $\phi, \psi \in \mathbf{BV}[0, 1]$;
 - (c) If $(\phi_n)_{n \geq 1} \subset \mathbf{BV}[0, 1]$ converges in L^2 to π , then $G^\phi := \lim_{n \rightarrow \infty} G^{\phi_n}$ exists as the L^2 -limit and defines a linear functional which satisfies (a) and (b).
5. Let $w \in \mathcal{C}_{(0)}[0, 1]$. Find the densities of the following bi-variate random variables:
 - (a) $\left(\int_{1/2}^t s^2 dw(s), w(1/2) \right)$ for $1/2 \leq t \leq 1$;
 - (b) $\left(\int_{1/2}^t s^2 dw(s), w(u + 1/2) \right)$ for $0 \leq u \leq t/2, 1/2 \leq t \leq 1$;
 - (c) $\left(\int_{1/2}^t s^2 dw(s), \int_{1/2}^t s dw(s) \right)$ for $1/2 \leq t \leq 1$;
 - (d) $\left(\int_{1/2}^1 e^s dw(s), w(1) - w(1/2) \right)$.
 6. Show that $F := \left\{ w \in \mathcal{C}_{(0)}[0, 1] : \sup_{q^{-1} \leq c \leq 1} \sup_{0 \leq r \leq 1} |w(cr) - w(r)| \geq 1 \right\}$ is for every $q > 1$ a closed subset of $\mathcal{C}_{(0)}[0, 1]$.
 7. Check the (in)equalities from step 3° in the proof of Lemma 12.18.

Chapter 13

Skorokhod representation

Let $(B_t)_{t \geq 0}$ be a one-dimensional Brownian motion. Denote by $\tau = \tau_{(-a,b)^c}^\circ$, $a, b > 0$, the first entry time into the interval $(-a, b)^c$. We know from Corollary 5.11 that

$$\begin{aligned} B_\tau &\sim \frac{b}{a+b} \delta_{-a} + \frac{a}{a+b} \delta_b \quad \text{and} \\ F_{-a,b}(x) = \mathbb{P}(B_\tau \leq x) &= \frac{b}{a+b} \mathbb{1}_{[-a,b)}(x) + \mathbb{1}_{[b,\infty)}(x). \end{aligned} \tag{13.1}$$

We will now study the problem which probability distributions can be obtained in this way. More precisely, let X be a random variable with probability distribution function $F(x) = \mathbb{P}(X \leq x)$. We want to know if there is a Brownian motion $(B_t)_{t \geq 0}$ and a stopping time τ such that $X \sim B_\tau$. If this is the case, we say that X (or F) can be *embedded into a Brownian motion*. The question and its solution (for continuous F) go back to Skorokhod [172, Chapter 7].

To simplify things, we do not require that τ is a stopping time for the natural filtration $(\mathcal{F}_t^B)_{t \geq 0}$, $\mathcal{F}_t^B = \sigma(B_s : s \leq t)$, but for a larger admissible filtration $(\mathcal{F}_t)_{t \geq 0}$, cf. Definition 5.1.

If X can be embedded into $(B_t)_{t \geq 0}$, it is clear that

$$\mathbb{E}(X^k) = \mathbb{E}(B_\tau^k)$$

whenever the moments exist. Moreover, if $\mathbb{E} \tau < \infty$, Wald's identities, cf. Theorem 5.10, entail

$$\mathbb{E} B_\tau = 0 \quad \text{and} \quad \mathbb{E}(B_\tau^2) = \mathbb{E} \tau.$$

Thus, X is necessarily centred: $\mathbb{E} X = 0$. On the other hand, every centred distribution function F can be obtained as a mixture of two-point distribution functions $F_{-a,b}$. This allows us to reduce the general embedding problem to the case considered in (13.1). The elegant proof of the following lemma is due to Breiman [14].

13.1 Lemma. *Let F be a distribution function such that $\int_{\mathbb{R}} |x| dF(x) < \infty$ and $\int_{\mathbb{R}} x dF(x) = 0$. Then*

$$F(x) = \iint_{\mathbb{R}^2} F_{u,w}(x) \nu(du, dw)$$

with the mixing probability distribution

$$\nu(du, dw) = \frac{1}{\alpha} (w - u) \mathbb{1}_{\{u \leq 0 < w\}}(u, w) dF(u) dF(w)$$

and $\alpha := \frac{1}{2} \int_{\mathbb{R}} |x| dF(x)$.

Proof. Since F is centred, we see that

$$\alpha = \frac{1}{2} \int_{\mathbb{R}} |x| dF(x) = \int_{(0, \infty)} x dF(x) = - \int_{(-\infty, 0]} x dF(x).$$

Using the formula (13.1) for $F_{u,w}$ we find

$$\begin{aligned} & \alpha \iint F_{u,w}(x) \nu(du, dw) \\ &= \iint_{u \leq 0 < w} \left(\frac{w}{w-u} (w-u) \mathbb{1}_{[u,w)}(x) + (w-u) \mathbb{1}_{[w, \infty)}(x) \right) dF(u) dF(w) \\ &= \iint_{u \leq 0 < w} \left(w \mathbb{1}_{[u,w)}(x) + (w-u) \mathbb{1}_{[w, \infty)}(x) \right) dF(u) dF(w) \\ &= \iint_{u \leq 0 < w} \left(w \mathbb{1}_{[u, \infty)}(x) - u \mathbb{1}_{[w, \infty)}(x) \right) dF(u) dF(w) \\ &= \iint_{u \leq 0 < w} \left(w \mathbb{1}_{(-\infty, x]}(u) - u \mathbb{1}_{(-\infty, x]}(w) \right) dF(u) dF(w) \\ &= \underbrace{\int_{(0, \infty)} w dF(w)}_{=\alpha} \int_{(-\infty, 0]} \mathbb{1}_{(-\infty, x]}(u) dF(u) \\ &\quad - \underbrace{\int_{(-\infty, 0]} u dF(u)}_{=\alpha} \int_{(0, \infty)} \mathbb{1}_{(-\infty, x]}(w) dF(w) \\ &= \alpha \int_{\mathbb{R}} \mathbb{1}_{(-\infty, x]}(z) dF(z) = \alpha F(x). \end{aligned}$$

Finally, the monotone convergence theorem shows that ν is indeed a probability measure:

$$1 = \lim_{x \rightarrow \infty} F(x) = \iint_{\mathbb{R}^2} \lim_{x \rightarrow \infty} F_{u,w}(x) \nu(du, dw) = \nu(\mathbb{R}^2). \quad \square$$

We can now prove Skorokhod's embedding theorem.

13.2 Theorem (Skorokhod 1965). *Let X be a real random variable with distribution function F . If $\mathbb{E}|X| < \infty$ and $\mathbb{E}X = 0$, then there is a $\text{BM}^1 (B_t)_{t \geq 0}$, an admissible filtration $(\mathcal{F}_t)_{t \geq 0}$ and a stopping time τ such that $B_\tau \sim X$. Moreover, $\mathbb{E}X^2 = \mathbb{E}\tau \in [0, \infty]$.*

Proof. Let $(B_t)_{t \geq 0}$ be a BM^1 and (U, W) be a random vector with values in \mathbb{R}^2 and distribution ν as in Lemma 13.1. We can assume that $(B_t)_{t \geq 0}$ and (U, W) are independent. Therefore, the filtration $\mathcal{F}_t := \sigma(B_s : s \leq t; U; W)$ is admissible in the sense of Definition 5.1. Ex. 13.1

Consider the first exit time of B from the random interval (U, W) ,

$$\tau_{(U,W)^c}^\circ = \inf\{t \geq 0 : B_t \notin (U, W)\},$$

and write $F_{u,w}$, $u \leq 0 < w$, for the distribution function of $B_{\tau_{(u,w)^c}^\circ}$, see (13.1). Since we have

$$\{\tau_{(U,W)^c}^\circ > t\} = \left\{ \min_{0 \leq s \leq t} B_s - U > 0 \right\} \cap \left\{ \max_{0 \leq s \leq t} B_s - W < 0 \right\},$$

$\tau_{(U,W)^c}^\circ$ is an \mathcal{F}_t stopping time. Because of the independence of (U, W) and $(B_t)_{t \geq 0}$, we can use Lemma 13.1 and find

$$\begin{aligned} \mathbb{P}(B_{\tau_{U,W}} \leq x) &\stackrel{\text{tower}}{=} \mathbb{E} \left[\mathbb{E} \left(\mathbb{1}_{(-\infty, x]}(B_{\tau_{(U,W)^c}^\circ}) \mid U, W \right) \right] \\ &\stackrel{\text{indep.}}{\stackrel{A.3}{=}} \mathbb{E} \left[\mathbb{E} \left(\mathbb{1}_{(-\infty, x]}(B_{\tau_{(u,w)^c}^\circ}) \mid u=U, w=W \right) \right] \\ &= \iint \mathbb{E} \left(\mathbb{1}_{(-\infty, x]}(B_{\tau_{(u,w)^c}^\circ}) \right) \nu(du, dw) \\ &\stackrel{(13.1)}{=} \iint F_{u,w}(x) \nu(du, dw) \\ &\stackrel{13.1}{=} F(x). \end{aligned}$$

Since $B_{\tau_{U,W}} \sim F$, we get by Tonelli's theorem, Wald's identities and (5.12)

$$\begin{aligned} \int x^2 dF(x) &= \mathbb{E} \left[B_{\tau_{(U,W)^c}^\circ}^2 \right] = \iint \mathbb{E} \left[B_{\tau_{(u,w)^c}^\circ}^2 \right] \nu(du, dw) \\ &\stackrel{5.10}{=} \iint \mathbb{E} \left[\tau_{(u,w)^c}^\circ \right] \nu(du, dw) \\ &\stackrel{(5.12)}{=} \iint uw \nu(du, dw). \quad \square \end{aligned}$$

Our simple proof of Skorokhod's embedding theorem comes at the cost that we have to use an *exogenous randomization* procedure which requires the enlargement of

the natural Brownian filtration. There are several proofs which avoid this, but they are more complicated, see e. g. Sawyer [165].

Because of the strong Markov property of a Brownian motion we can use Theorem 13.2 to embed any random walk $(S_n)_{n \geq 1}$, $S_n = X_1 + \dots + X_n$, where the X_j are independent centred random variables, in a Brownian motion. For simplicity we will consider only the case where the X_j are iid.

13.3 Corollary. *Let $(X_j)_{j \geq 1}$ be a sequence of iid real random variables with common distribution F such that $\mathbb{E}|X_1| < \infty$ and $\mathbb{E} X_1 = 0$. Then there is a one-dimensional Brownian motion $(B_t)_{t \geq 0}$, an admissible filtration $(\mathcal{F}_t)_{t \geq 0}$ and a sequence $(\tau_n)_{n \geq 1}$ of \mathcal{F}_t stopping times such that the following statements hold:*

- a) *the increments $(\tau_n - \tau_{n-1})_{n \geq 1}$ are iid positive random variables ($\tau_0 := 0$);*
- b) *$\mathbb{E}(\tau_n - \tau_{n-1}) = \mathbb{E}(X_1^2)$;*
- c) *the sequences $(B_{\tau_n})_{n \geq 1}$ and $(S_n)_{n \geq 1}$, $S_n := X_1 + \dots + X_n$, have the same distribution.*

Proof. Let $(B_t)_{t \geq 0}$ be a one-dimensional Brownian motion and (U_n, W_n) , $n \geq 0$, iid random vectors such that $(U_0, W_0) \sim \nu$ with ν as in Lemma 13.1. We can assume that $(B_t)_{t \geq 0}$ and $(U_n, W_n)_{n \geq 0}$ are independent. Then

$$\begin{aligned} \mathcal{F}_t^{(n)} &:= \sigma(B_s : s \leq t; (U_0, W_0), \dots, (U_n, W_n)), \quad t \geq 0, \\ \mathcal{F}_t &:= \sigma(B_s : s \leq t; (U_n, W_n), n \geq 1), \quad t \geq 0, \end{aligned}$$

are, for every $n \geq 1$, admissible filtrations for $(B_t)_{t \geq 0}$.

Set $\tau_0 := 0$ and, recursively,

$$\tau_{n+1} := \inf \{t \geq \tau_n : B_t - B_{\tau_n} \notin (U_n, W_n)\}, \quad n = 0, 1, 2, \dots$$

As in Theorem 13.2 one can see that τ_n is an $\mathcal{F}_t^{(n)}$ stopping time and that B_{τ_n} is $\mathcal{F}_{\tau_n}^{(n)}$ measurable, cf. Lemma A.17 or Corollary 6.25.

Observe that

$$\tau_{n+1} - \tau_n = \inf \{s \geq 0 : B_{s+\tau_n} - B_{\tau_n} \notin (U_n, W_n)\}.$$

By the strong Markov property, Theorem 6.5, $(B_{s+\tau_n} - B_{\tau_n})_{s \geq 0}$ is independent of $\mathcal{F}_{\tau_n}^{(n)}$; in particular $B_{\tau_{n+1}} - B_{\tau_n} \perp\!\!\!\perp \mathcal{F}_{\tau_n}^{(n)}$. Since $(B_{s+\tau_n} - B_{\tau_n})_{s \geq 0}$ and (U_n, W_n) have the same distributions as $(B_t)_{t \geq 0}$ and (U_1, W_1) , respectively, we find that $\tau_{n+1} - \tau_n \sim \tau_1$ and

$$B_{\tau_{n+1}} - B_{\tau_n} = B_{(\tau_{n+1}-\tau_n)+\tau_n} - B_{\tau_n} \sim B_{\tau_1}.$$

By Wald's identity we get, in particular, $\mathbb{E}(\tau_{n+1} - \tau_n) = \mathbb{E} \tau_1 = \mathbb{E}(X_1^2)$.

Finally, note that τ_1, \dots, τ_n and $B_{\tau_1}, \dots, B_{\tau_n}$ are $\mathcal{F}_{\tau_n}^{(n)}$ measurable, and that τ_1 and B_{τ_1} satisfy the basic embedding Theorem 13.2. \square

13.4 Remark. We can use the embedding of random walks in Brownian motion to derive asymptotic results for random walks from corresponding results for Brownian motion, cf. [32]; these are often easier to prove. We use the law of the iterated logarithm to illustrate this point.

Let $(X_j)_{j \geq 0}$ be a sequence of iid random variables with mean zero and variance σ . By the Skorokhod embedding there are stopping times τ_n such that

$$S_n := X_1 + \dots + X_n \sim B_{\tau_n}.$$

By Corollary 13.3 and Wald's identities $(\tau_n)_{n \geq 1}$ is a random walk with mean value $\mathbb{E} \tau_n = \mathbb{E}(B_{\tau_n}^2) = \mathbb{E}(S_n^2) < \infty$. By the strong law of large numbers,

$$\lim_{n \rightarrow \infty} \frac{\tau_n}{n} = \mathbb{E} \tau_1 = \sigma^2 \quad \text{a.s.}$$

Set $\phi(s) := \sqrt{2s \log \log s}$. Then $\lim_{n \rightarrow \infty} \phi(\tau_n)/\phi(n) = \sigma$. We will show that

$$\overline{\lim}_{n \rightarrow \infty} \frac{S_n}{\sqrt{2n \log \log n}} = \overline{\lim}_{n \rightarrow \infty} \frac{S_n}{\phi(n)} = \sigma.$$

For this we claim that

$$\sup_{0 \leq r \leq \tau_n - \tau_{n-1}} |B_{r+\tau_{n-1}} - B_{\tau_n}| = O(\sqrt{n}) \quad \text{a.s. as } n \rightarrow \infty. \quad (13.2)$$

Accepting this for the time being, we find for all $s \in [\tau_{n-1}, \tau_n]$

$$\begin{aligned} \left| \frac{B_s}{\phi(s)} - \frac{B_{\tau_n}}{\phi(\tau_n)} \right| &\leq \left| \frac{B_s}{\phi(s)} - \frac{B_s}{\phi(\tau_n)} \right| + \left| \frac{B_s - B_{\tau_n}}{\phi(\tau_n)} \right| \\ &\leq \left| \frac{B_s}{\phi(s)} \right| \left| 1 - \frac{\phi(s)}{\phi(\tau_n)} \right| + \frac{1}{\phi(\tau_n)} \sup_{0 \leq r \leq \tau_n - \tau_{n-1}} |B_{r+\tau_{n-1}} - B_{\tau_n}| \\ &\leq \left| \frac{B_s}{\phi(s)} \right| \left| 1 - \frac{\phi(s)}{\phi(\tau_n)} \right| + \frac{C\sqrt{n}}{\phi(\tau_n)}. \end{aligned}$$

By the law of the iterated logarithm, Theorem 11.1, $B_s/\phi(s)$ stays bounded while the other terms tend to zero as $n \rightarrow \infty$. Thus,

$$1 = \overline{\lim}_{s \rightarrow \infty} \frac{B_s}{\phi(s)} = \overline{\lim}_{n \rightarrow \infty} \frac{B_{\tau_n}}{\phi(\tau_n)} = \overline{\lim}_{n \rightarrow \infty} \frac{S_n}{\phi(n)} \underbrace{\frac{\phi(n)}{\phi(\tau_n)}}_{\rightarrow 1/\sigma, n \rightarrow \infty} = \frac{1}{\sigma} \overline{\lim}_{n \rightarrow \infty} \frac{S_n}{\phi(n)}.$$

All that remains to be done is to check (13.2). For this we use a Borel–Cantelli argument. Observe that, by the strong Markov property,

$$\sup_{0 \leq s \leq \tau_n - \tau_{n-1}} (B_{s+\tau_{n-1}} - B_{\tau_{n-1}}) \stackrel{\text{Thm. 6.5}}{\sim} \sup_{0 \leq s \leq \tau_1} B_s.$$

Using the fact that $\sum_{n=1}^{\infty} \mu(f > n) \leq \int f d\mu$ and Doob's maximal inequality (A.14) for the stopped Brownian motion $(B_{t \wedge \tau_1})_{t \geq 0}$, we see

$$\begin{aligned} \sum_{n=1}^{\infty} \mathbb{P} \left(\sup_{0 \leq s \leq \tau_n - \tau_{n-1}} |B_{s+\tau_{n-1}} - B_{\tau_{n-1}}| > \sqrt{n} \right) &= \sum_{n=1}^{\infty} \mathbb{P} \left(\sup_{0 \leq s \leq \tau_1} B_s^2 > n \right) \\ &\leq \mathbb{E} \left(\sup_{0 \leq s \leq \tau_1} B_s^2 \right) = \mathbb{E} \left(\sup_{s \geq 0} B_{s \wedge \tau_1}^2 \right) \stackrel{\text{Doob}}{\leq} 4 \mathbb{E} (B_{\tau_1}^2) = 4 \mathbb{E} (X_1^2) < \infty. \end{aligned}$$

Now we can use the (simple direction of the) Borel–Cantelli lemma to deduce, for sufficiently large values of n ,

$$\sup_{0 \leq s \leq \tau_n - \tau_{n-1}} |B_{s+\tau_{n-1}} - B_{\tau_{n-1}}| \leq C \sqrt{n}.$$

This inequality holds in particular for $s = \tau_n - \tau_{n-1}$, and the triangle inequality gives (13.2).

We have seen in Section 3.4 that Brownian motion can be constructed as the limit of random walks. For every *finite* family $(B_{t_1}, \dots, B_{t_k})$, $t_1 < \dots < t_k$, this is just the classical central limit theorem: Let $(\epsilon_j)_{j \geq 1}$ be iid Bernoulli random variables such that $\mathbb{P}(\epsilon_1 = \pm 1) = \frac{1}{2}$ and $S_n := \epsilon_1 + \dots + \epsilon_n$, then¹

$$\frac{S_{\lfloor nt \rfloor}}{\sqrt{n}} \xrightarrow[n \rightarrow \infty]{\text{weakly}} B_t \quad \text{for all } t \in (0, 1].$$

We are interested in approximations of the whole path $[0, 1] \ni t \mapsto B_t(\omega)$ of a Brownian motion, and the above convergence will give pointwise but not uniform results. To get uniform results, it is useful to interpolate the partial sums $S_{\lfloor nt \rfloor} / \sqrt{n}$

$$S^n(t) := \frac{1}{\sqrt{n}} (S_{\lfloor nt \rfloor} - (nt - \lfloor nt \rfloor) \epsilon_{\lfloor nt \rfloor + 1}), \quad t \in [0, 1].$$

This is a piecewise linear continuous function which is equal to S_j / \sqrt{n} if $t = j/n$. We are interested in the convergence of

$$\Phi(S^n(\cdot)) \xrightarrow[n \rightarrow \infty]{} \Phi(B(\cdot))$$

¹ In fact, it is enough to assume that the ϵ_j are iid random variables with mean zero and finite variance. This universality is the reason why Theorem 13.5 is usually called *invariance principle*.

where $\Phi : (\mathcal{C}[0, 1], \mathbb{R}) \rightarrow \mathbb{R}$ is a (not necessarily linear) bounded functional which is uniformly continuous with respect to uniform convergence.

Although the following theorem implies Theorem 3.8, we cannot use it to construct Brownian motion since the proof already uses the existence of a Brownian motion. Nevertheless the result is important as it helps to understand Brownian motion intuitively and permits simple proofs of various limit theorems.

13.5 Theorem (Invariance principle. Donsker 1951). *Let $(B_t)_{t \in [0, 1]}$ be a one-dimensional Brownian motion, $(S^n(t))_{t \in [0, 1]}$, $n \geq 1$, be the sequence of processes from above, and $\Phi : \mathcal{C}([0, 1], \mathbb{R}) \rightarrow \mathbb{R}$ a uniformly continuous bounded functional. Then*

$$\lim_{n \rightarrow \infty} \mathbb{E} \Phi(S^n(\cdot)) = \mathbb{E} \Phi(B(\cdot)),$$

i. e. $S^n(\cdot)$ converges weakly to $B(\cdot)$.

Proof. Denote by $(W_t)_{t \geq 0}$ the Brownian motion into which we can embed the random walk $(S_n)_{n \geq 1}$, i. e. $W_{\tau_n} \sim S_n$ for an increasing sequence of stopping times $\tau_1 \leq \tau_2 \leq \dots$, cf. Corollary 13.3. Since our claim is a statement on the probability distributions, we can assume that $S_n = W_{\tau_n}$.

By assumption, we find for all $\epsilon > 0$ some $\eta > 0$ such that $|\Phi(f) - \Phi(g)| < \epsilon$ for all $\|f - g\|_\infty < \eta$.

The assertion follows if we can prove, for all $\epsilon > 0$ and for sufficiently large values of $n \geq N_\epsilon$, that the Brownian motions defined by $B^n(t) := n^{-1/2} W_{nt} \sim B_t$, $t \in [0, 1]$, satisfy

$$\mathbb{P}(|\Phi(S^n(\cdot)) - \Phi(B^n(\cdot))| > \epsilon) \leq \mathbb{P}\left(\sup_{0 \leq t \leq 1} |S^n(t) - B^n(t)| > \eta\right) \leq \epsilon.$$

This means that $\Phi(S^n(\cdot)) - \Phi(B^n(\cdot))$ converges in probability to 0, therefore

$$|\mathbb{E} \Phi(S^n(\cdot)) - \mathbb{E} \Phi(B(\cdot))| \stackrel{B^n \sim B}{=} |\mathbb{E} \Phi(S^n(\cdot)) - \mathbb{E} \Phi(B^n(\cdot))| \xrightarrow{n \rightarrow \infty} 0.$$

In order to see the estimate we need three simple observations. Pick $\epsilon, \eta > 0$ as above.

1° Since a Brownian motion has continuous sample paths, there is some $\delta > 0$ such that

$$\mathbb{P}\left(\exists s, t \in [0, 1], |s - t| \leq \delta : |W_t - W_s| > \eta\right) \leq \frac{\epsilon}{2}.$$

2° By Corollary 13.3 and Wald's identities, $(\tau_n)_{n \geq 1}$ is a random walk with mean $\mathbb{E} \tau_n = \mathbb{E}(B_{\tau_n}^2) = \mathbb{E}(S_n^2) < \infty$. Therefore, the strong law of large numbers tells us that

$$\lim_{n \rightarrow \infty} \frac{\tau_n}{n} = \mathbb{E} \tau_1 = 1 \quad \text{a.s.}$$

Thus, there is some $M(\eta)$ such that for all $n > m \geq M(\eta)$

$$\begin{aligned} \frac{1}{n} \max_{1 \leq k \leq n} |\tau_k - k| &\leq \frac{1}{n} \max_{1 \leq k \leq m} |\tau_k - k| + \max_{m \leq k \leq n} \left| \frac{\tau_k}{k} - 1 \right| \\ &\leq \frac{1}{n} \max_{1 \leq k \leq m} |\tau_k - k| + \eta \xrightarrow[n \rightarrow \infty]{} \eta \xrightarrow[\eta \rightarrow 0]{} 0. \end{aligned}$$

Consequently, for some $N(\epsilon, \eta)$ and all $n \geq N(\epsilon, \eta)$

$$\mathbb{P} \left(\frac{1}{n} \max_{1 \leq k \leq n} |\tau_k - k| > \frac{\delta}{3} \right) \leq \frac{\epsilon}{2}.$$

3° For every $t \in [k/n, (k+1)/n]$ there is a $u \in [\tau_k/n, \tau_{k+1}/n]$, $k = 0, 1, \dots, n-1$, such that

$$S^n(t) = n^{-1/2} W_{nu} = B^n(u).$$

Indeed: $S^n(t)$ interpolates $B^n(\tau_k)$ and $B^n(\tau_{k+1})$ linearly, i.e. $S^n(t)$ is between the minimum and the maximum of $B^n(s)$, $\tau_k \leq s \leq \tau_{k+1}$. Since $s \mapsto B^n(s)$ is continuous, there is some intermediate value u such that $S^n(t) = B^n(u)$. This situation is shown in Figure 13.1.

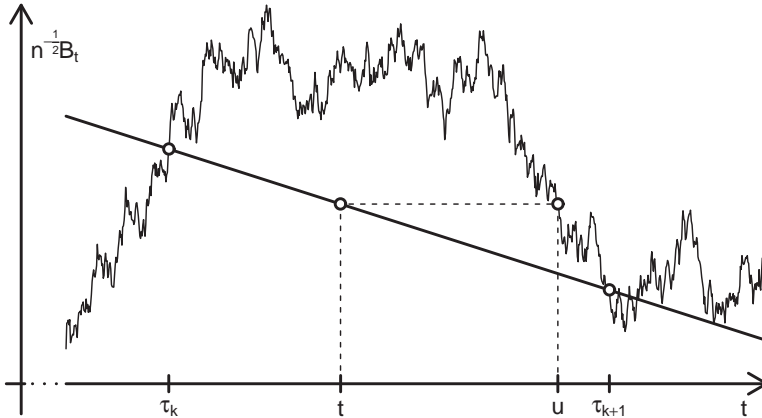


Figure 13.1. For each value $S^n(t)$ of the line segment connecting $B^n(\tau_k)$ and $B^n(\tau_{k+1})$ there is some u such that $S^n(t) = B^n(u)$.

Clearly, we have

$$\begin{aligned} & \mathbb{P} \left(\sup_{0 \leq t \leq 1} |S^n(t) - B^n(t)| > \eta \right) \\ & \leq \mathbb{P} \left(\sup_{0 \leq t \leq 1} |S^n(t) - B^n(t)| > \eta, \max_{1 \leq l \leq n} \left| \frac{\tau_l}{n} - \frac{l}{n} \right| \leq \frac{\delta}{3} \right) \\ & \quad + \mathbb{P} \left(\frac{1}{n} \max_{1 \leq l \leq n} |\tau_l - l| > \frac{\delta}{3} \right). \end{aligned}$$

If $|S^n(t) - B^n(t)| > \eta$ for some $t \in [k/n, (k+1)/n]$, then we find by 3° some $u \in [\tau_k/n, \tau_{k+1}/n]$ with $|B^n(u) - B^n(t)| > \eta$. Moreover,

$$|u - t| \leq \underbrace{\left| u - \frac{\tau_k}{n} \right|}_{\rightarrow 0, n \rightarrow \infty \text{ by } 2^\circ} + \left| \frac{\tau_k}{n} - \frac{k}{n} \right| + \left| \frac{k}{n} - t \right| \leq \frac{\delta}{3} + \frac{\delta}{3} + \frac{1}{n} \leq \delta$$

if $n \geq N(\epsilon, \eta) \vee (3/\delta)$ and $\max_{1 \leq l \leq n} \left| \frac{\tau_l}{n} - \frac{l}{n} \right| \leq \delta/3$. This, 2° and 1° give for all sufficiently large n

$$\begin{aligned} & \mathbb{P} \left(\sup_{0 \leq t \leq 1} |S^n(t) - B^n(t)| > \eta \right) \\ & \leq \mathbb{P} \left(\exists u, t \in [0, 1], |u - t| \leq \delta : |B^n(u) - B^n(t)| > \eta \right) + \frac{\epsilon}{2} \leq \epsilon. \quad \square \end{aligned}$$

Typical examples of uniformly continuous functionals are

$$\begin{aligned} \Phi_0(f) &= f(1), & \Phi_1(f) &= \sup_{0 \leq t \leq 1} |f(t) - at|, \\ \Phi_2(f) &= \int_0^1 |f(t)|^p dt, & \Phi_3(f) &= \int_0^1 |f(t) - tf(1)|^4 dt. \end{aligned}$$

With a bit of extra work one can prove Theorem 13.5 for all $\Phi : \mathcal{C}([0, 1], \mathbb{R}) \rightarrow \mathbb{R}$ which are continuous with respect to uniform convergence *at almost every Brownian path*. This version includes the following functionals

$$\begin{aligned} \Phi_4(f) &= \mathbb{1}_{\{\sup_{0 \leq t \leq 1} |f(t)| \leq 1\}}, & \Phi_5(f) &= \text{Leb}\{f > 0\} - \text{Leb}\{f < 0\}, \\ \Phi_6(f) &= \mathbb{1}_{\{\forall t : a(t) < f(t) < b(t)\}} \quad (a, b : [0, 1] \rightarrow \mathbb{R} \text{ smooth functions}). \end{aligned}$$

Using Φ_0 we could deduce the central limit theorem (without using characteristic function methods), Φ_2 gives the limit theorem

$$\lim_{n \rightarrow \infty} \mathbb{P} \left(\sum_{k=1}^n |S_k|^p \leq \lambda n^p \right) = \mathbb{P} \left(\int_0^1 |B_t|^p dt \leq \lambda \right)$$

while Φ_4 corresponds to the gambler's ruin problem and Φ_6 is Khintchine's Problem, see Sawyer [165] for a more detailed discussion.



13.6 Further reading. More applications and variants of Skorokhod's embedding theorem can be found in the monograph [65] and in the survey papers [165] and [137].

[65] Freedman: *Brownian Motion and Diffusion*.

[137] Oblój: The Skorokhod embedding problem and its offspring.

[165] Sawyer: The Skorokhod Representation.

Problems

1. In the proof of Theorem 13.2 we assumed that $(B_t)_{t \geq 0}$ and (U, W) are independent. Show that $\mathcal{F}_t := \sigma(B_s, s \leq t; U, W)$ is an admissible filtration for a Brownian motion, cf. Definition 5.1.

Chapter 14

Stochastic integrals: L^2 -Theory

A function $f : [0, T] \rightarrow \mathbb{R}$ is *Riemann integrable* if the *Riemann sums*

$$S^\Pi(f) := \sum_{j=1}^n f(\xi_j^\Pi)(s_j^\Pi - s_{j-1}^\Pi), \quad \xi_j^\Pi \in [s_{j-1}^\Pi, s_j^\Pi], \quad (14.1)$$

converge to a finite limit I . The limit is taken along all finite partitions of the interval $[0, T]$, $\Pi = \{0 = s_0^\Pi < s_1^\Pi < \dots < s_n^\Pi = T\}$, with $|\Pi| := \max_j (s_j^\Pi - s_{j-1}^\Pi) \rightarrow 0$ and it is independent of the particular sequence and the choice of the intermediate points ξ_j^Π . The number I is called the *definite Riemann integral* of f in $[0, T]$ and one writes $I = \int_0^T f(s) ds$. Riemann's definition has been modified in many ways. T. J. Stieltjes used another function g , replacing $s_j^\Pi - s_{j-1}^\Pi$ in (14.1) by $g(s_j^\Pi) - g(s_{j-1}^\Pi)$. The resulting integral $\int_0^T f(s) dg(s)$ is the *Riemann–Stieltjes integral*. The requirement that the Riemann–Stieltjes sums converge imposes some restrictions on f and g . For example, if we want that $\int_0^T f(s) dg(s)$ exists for all $f \in \mathcal{C}[0, T]$, we can show that g is necessarily of bounded variation: $\text{VAR}_1(g; t) < \infty$, see Corollary A.41. Since we know from Chapter 9 that a Brownian motion is of unbounded variation on any compact interval, it is clear that we cannot define $\int_0^T f(s) dB_s(\omega)$ as a Riemann–Stieltjes integral. Using the notion of L^2 martingales K. Itô succeeded 1942 to give a consistent and powerful meaning to $\int_0^T f(s, \omega) dB_s(\omega)$ where the class of integrand processes $f(s, \omega)$ contains all bounded processes with continuous sample paths.

14.1 Discrete stochastic integrals

Let $(\mathcal{F}_n)_{n \geq 0}$ be a filtration. An L^2 martingale $X = (X_n, \mathcal{F}_n)_{n \geq 0}$ is a martingale such that $X_n \in L^2(\mathbb{P})$ for all $n \geq 0$. We write \mathcal{M}^2 for the family of L^2 martingales. If $X \in \mathcal{M}^2$, then $X^2 = (X_n^2, \mathcal{F}_n)_{n \geq 0}$ is a submartingale and, by the Doob decomposition,

$$X_n^2 = X_0^2 + M_n + A_n \quad (14.2)$$

where $(M_n, \mathcal{F}_n)_{n \geq 0}$ is a martingale and $(A_n)_{n \geq 0}$ is a *previsible* (i. e. A_n is \mathcal{F}_{n-1} mea-

surable, $n \geq 1$) and *increasing* (i. e.: $A_n \leq A_{n+1}$, $n \geq 0$) process. This decomposition is unique. Therefore, the following definition makes sense.

14.1 Definition. Let $(X_n, \mathcal{F}_n)_{n \geq 0}$ be an L^2 martingale. The *quadratic variation* of X , $\langle X \rangle = (\langle X \rangle_n)_{n \geq 0}$, is the unique increasing and previsible process with $\langle X \rangle_0 = 0$ appearing in the Doob decomposition (14.2).

The next lemma explains why $\langle X \rangle$ is called the quadratic variation.

14.2 Lemma. Let $X = (X_n, \mathcal{F}_n)_{n \geq 0}$ be an L^2 martingale and $\langle X \rangle$ the quadratic variation. Then, for all $n > m \geq 0$,

$$\mathbb{E}[(X_n - X_m)^2 | \mathcal{F}_m] = \mathbb{E}[X_n^2 - X_m^2 | \mathcal{F}_m] = \mathbb{E}[\langle X \rangle_n - \langle X \rangle_m | \mathcal{F}_m]. \quad (14.3)$$

Proof. Note that for all $n > m \geq 0$

$$\begin{aligned} & \mathbb{E}[(X_n - X_m)^2 | \mathcal{F}_m] \\ &= \mathbb{E}[X_n^2 - 2X_n X_m + X_m^2 | \mathcal{F}_m] \\ &= \mathbb{E}[X_n^2 | \mathcal{F}_m] - 2X_m \underbrace{\mathbb{E}[X_n | \mathcal{F}_m]}_{= X_m, \text{ martingale}} + X_m^2 \\ &= \mathbb{E}[X_n^2 | \mathcal{F}_m] - X_m^2 \\ &= \mathbb{E}[X_n^2 - X_m^2 | \mathcal{F}_m] \\ &= \mathbb{E}[\underbrace{X_n^2 - \langle X \rangle_n}_{\text{martingale}} - (X_m^2 - \langle X \rangle_m) | \mathcal{F}_m] + \mathbb{E}[\langle X \rangle_n - \langle X \rangle_m | \mathcal{F}_m] \\ &= \mathbb{E}[\langle X \rangle_n - \langle X \rangle_m | \mathcal{F}_m]. \quad \square \end{aligned}$$

14.3 Definition. Let $(X_n, \mathcal{F}_n)_{n \geq 0}$ be a martingale and $(C_n, \mathcal{F}_n)_{n \geq 0}$ an adapted process. Then

$$C \bullet X_n := \sum_{j=1}^n C_{j-1}(X_j - X_{j-1}), \quad C \bullet X_0 := 0, \quad n \geq 1, \quad (14.4)$$

is called a *martingale transform*.

Let us review a few properties of the martingale transform.

14.4 Theorem. Let $(M_n, \mathcal{F}_n)_{n \geq 0}$ be an L^2 martingale and $(C_n, \mathcal{F}_n)_{n \geq 0}$ be a bounded adapted process. Then

- a) $C \bullet M$ is an L^2 martingale, i. e. $C \bullet : \mathcal{M}^2 \rightarrow \mathcal{M}^2$;
- b) $C \mapsto C \bullet M$ is linear;
- c) $\langle C \bullet M \rangle_n = C^2 \bullet \langle M \rangle_n := \sum_{j=1}^n C_{j-1}^2 (\langle M \rangle_j - \langle M \rangle_{j-1})$ for all $n \geq 1$;

- d) $\mathbb{E}[(C \bullet M_n - C \bullet M_m)^2 \mid \mathcal{F}_m] = \mathbb{E}[C^2 \bullet \langle M \rangle_n - C^2 \bullet \langle M \rangle_m \mid \mathcal{F}_m]$ for $n > m \geq 0$.
In particular,

$$\mathbb{E}[(C \bullet M_n)^2] = \mathbb{E}[C^2 \bullet \langle M \rangle_n] \quad (14.5)$$

$$= \mathbb{E}\left[\sum_{j=1}^n C_{j-1}^2 (M_j - M_{j-1})^2\right]. \quad (14.6)$$

Proof. a) By assumption, there is some constant $K > 0$ such that $|C_j| \leq K$ a.s. for all $n \geq 0$. Thus, $|C_{j-1}(M_j - M_{j-1})| \leq K(|M_{j-1}| + |M_j|) \in L^2$, which means that $C \bullet M_n \in L^2$. Let us check that $C \bullet M$ is a martingale. For all $n \geq 1$

$$\begin{aligned} \mathbb{E}(C \bullet M_n \mid \mathcal{F}_{n-1}) &= \mathbb{E}(C \bullet M_{n-1} + C_{n-1}(M_n - M_{n-1}) \mid \mathcal{F}_{n-1}) \\ &= C \bullet M_{n-1} + C_{n-1} \underbrace{\mathbb{E}((M_n - M_{n-1}) \mid \mathcal{F}_{n-1})}_{=0, \text{ martingale}} = C \bullet M_{n-1}. \end{aligned}$$

b) The linearity of $C \mapsto C \bullet M$ is obvious. In order to show

c) it is enough to prove that

- (1) $C^2 \bullet \langle M \rangle_n$ is previsible and increasing
- (2) $(C \bullet M_n)^2 - C^2 \bullet \langle M \rangle_n$ is a martingale.

By the uniqueness of the quadratic variation we see $\langle C \bullet M \rangle_n = C^2 \bullet \langle M \rangle_n$. The first point follows easily from

$$C^2 \bullet \langle M \rangle_n = \sum_{j=1}^n \overbrace{C_{j-1}^2}^{\geq 0} \underbrace{(\langle M \rangle_j - \langle M \rangle_{j-1})}_{\mathcal{F}_{j-1} \text{ mble}}.$$

In order to see the martingale property we apply Lemma 14.2 to the L^2 martingale $X = C \bullet M$ with $m = n - 1 \geq 0$ and get

$$\begin{aligned} \mathbb{E}[(C \bullet M_n)^2 - (C \bullet M_{n-1})^2 \mid \mathcal{F}_{n-1}] &\stackrel{(14.3)}{=} \mathbb{E}[(C \bullet M_n - C \bullet M_{n-1})^2 \mid \mathcal{F}_{n-1}] \\ &= \mathbb{E}[C_{n-1}^2 (M_n - M_{n-1})^2 \mid \mathcal{F}_{n-1}] \\ &= C_{n-1}^2 \mathbb{E}[(M_n - M_{n-1})^2 \mid \mathcal{F}_{n-1}] \\ &\stackrel{(14.3)}{=} C_{n-1}^2 \mathbb{E}[\langle M \rangle_n - \langle M \rangle_{n-1} \mid \mathcal{F}_{n-1}] \\ &= \mathbb{E}[C_{n-1}^2 (\langle M \rangle_n - \langle M \rangle_{n-1}) \mid \mathcal{F}_{n-1}] \\ &= \mathbb{E}[C^2 \bullet \langle M \rangle_n - C^2 \bullet \langle M \rangle_{n-1} \mid \mathcal{F}_{n-1}]. \end{aligned}$$

This shows for all $n \geq 1$

$$\begin{aligned}\mathbb{E}[(C \bullet M_n)^2 - C^2 \bullet \langle M \rangle_n \mid \mathcal{F}_{n-1}] &= \mathbb{E}[(C \bullet M_{n-1})^2 - C^2 \bullet \langle M \rangle_{n-1} \mid \mathcal{F}_{n-1}] \\ &= (C \bullet M_{n-1})^2 - C^2 \bullet \langle M \rangle_{n-1},\end{aligned}$$

i. e. $(C \bullet M)^2 - C^2 \bullet \langle M \rangle$ is a martingale.

d) using c) and Lemma 14.2 for the L^2 martingale $C \bullet M$ gives

$$\begin{aligned}\mathbb{E}[(C \bullet M_n - C \bullet M_m)^2 \mid \mathcal{F}_m] &= \mathbb{E}[(C \bullet M_n)^2 - (C \bullet M_m)^2 \mid \mathcal{F}_m] \\ &= \mathbb{E}[C^2 \bullet \langle M \rangle_n - C^2 \bullet \langle M \rangle_m \mid \mathcal{F}_m].\end{aligned}$$

If we take $m = 0$, we get (14.5). If we use Lemma 14.2 for the L^2 martingale M , we see $\mathbb{E}(\langle M \rangle_j - \langle M \rangle_{j-1} \mid \mathcal{F}_{j-1}) = \mathbb{E}((M_j - M_{j-1})^2 \mid \mathcal{F}_{j-1})$. Therefore, (14.6) follows from (14.5) by the tower property. \square

Note that $M \mapsto \langle M \rangle$ is a quadratic form. Therefore, we can use polarization to get the *quadratic covariation* which is a bilinear form:

$$\langle M, N \rangle := \frac{1}{4}(\langle M + N \rangle - \langle M - N \rangle) \quad \text{for all } M, N \in \mathcal{M}^2. \quad (14.7)$$

Ex. 14.1 By construction, $\langle M, M \rangle = \langle M \rangle$. Using polarization, we see that $M_n N_n - \langle M, N \rangle_n$ is a martingale. Moreover

14.5 Lemma. Let $(M_n, \mathcal{F}_n)_{n \geq 0}$, $(N_n, \mathcal{F}_n)_{n \geq 0}$ be L^2 martingales and assume that $(C_n, \mathcal{F}_n)_{n \geq 0}$ is a bounded adapted processes. Then

$$\langle C \bullet M, N \rangle_n = C \bullet \langle M, N \rangle_n = \sum_{j=1}^n C_{j-1} (\langle M, N \rangle_j - \langle M, N \rangle_{j-1}). \quad (14.8)$$

Moreover, if for some $I \in \mathcal{M}^2$

$$\langle I, N \rangle_n = C \bullet \langle M, N \rangle_n \quad \text{for all } N \in \mathcal{M}^2, n \geq 1, \quad (14.9)$$

then $I = C \bullet M$.

Lemma 14.5 shows that $C \bullet M$ is the only L^2 martingale satisfying (14.9).

Proof. The equality (14.8) follows from polarization and Theorem 14.4 c):

$$\langle C \bullet M \pm N \rangle = \langle C \bullet M \rangle \pm 2C \bullet \langle M, N \rangle + \langle N \rangle = C^2 \bullet \langle M \rangle \pm 2C \bullet \langle M, N \rangle + \langle N \rangle$$

and we get (14.8) by adding these two equalities. In particular, (14.9) holds for $I = C \bullet M$.

Now assume that (14.9) is true. Then

$$\langle I, N \rangle \stackrel{(14.9)}{=} C \bullet \langle M, N \rangle \stackrel{(14.8)}{=} \langle C \bullet M, N \rangle$$

or $\langle I - C \bullet M, N \rangle = 0$ for all $N \in \mathcal{M}^2$. Since $I - C \bullet M \in \mathcal{M}^2$, we can set $N = I - C \bullet M$ and find

$$\langle I - C \bullet M, I - C \bullet M \rangle = \langle I - C \bullet M \rangle = 0.$$

Thus, $\mathbb{E}[(I_n - C \bullet M_n)^2] = \mathbb{E}[\langle I - C \bullet M \rangle_n] = 0$, which shows that $I_n = C \bullet M_n$ a. s. for all $n \geq 0$. \square

14.2 Simple integrands

We can now consider stochastic integrals with respect to a Brownian motion. Throughout the rest of this chapter $(B_t)_{t \geq 0}$ is a BM¹, $(\mathcal{F}_t)_{t \geq 0}$ is an admissible filtration such that each \mathcal{F}_t contains all \mathbb{P} -null sets, e. g. the completed natural filtration $\mathcal{F}_t = \overline{\mathcal{F}_t^B}$ (cf. Theorem 6.21), and $[0, T]$, $T < \infty$, a finite interval. Although we do not explicitly state it, most results remain valid for $T = \infty$ and $[0, \infty)$. In line with the notation introduced in Section 14.1 we use

- X^τ and $X_s^\tau := X_{\tau \wedge s}$ for the stopped process $(X_t)_{t \geq 0}$;
- \mathcal{M}_T^2 for the family of \mathcal{F}_t martingales $(M_t)_{0 \leq t \leq T} \subset L^2(\mathbb{P})$;
- $\mathcal{M}_T^{2,c}$ for the L^2 martingales with (almost surely) continuous sample paths;
- $L^2(\lambda_T \otimes \mathbb{P}) = L^2([0, T] \times \Omega, \lambda \otimes \mathbb{P})$ for the set of (equivalence classes of) $\mathcal{B}[0, T] \otimes \mathcal{A}$ measurable random functions $f : [0, T] \times \Omega \rightarrow \mathbb{R}$ equipped with the norm $\|f\|_{L^2(\lambda_T \otimes \mathbb{P})}^2 = \int_0^T [\mathbb{E}(|f(s)|^2)] ds < \infty$.

14.6 Definition. A real-valued stochastic process $(f(t, \cdot))_{t \in [0, T]}$ of the form

$$f(t, \omega) = \sum_{j=1}^n \phi_{j-1}(\omega) \mathbb{1}_{[s_{j-1}, s_j)}(t) \quad (14.10)$$

where $n \geq 1$, $0 = s_0 \leq s_1 \leq \dots \leq s_n \leq T$ and $\phi_j \in L^\infty(\mathcal{F}_{s_j})$ are bounded \mathcal{F}_{s_j} measurable random variables, $j = 0, \dots, n$, is called a (*right-continuous*) *simple process*. We write \mathcal{E}_T for the family of all simple processes on $[0, T]$.

Note that we can, a bit tautologically, rewrite (14.10) as

$$f(t, \omega) = \sum_{j=1}^n f(s_{j-1}, \omega) \mathbb{1}_{[s_{j-1}, s_j)}(t). \quad (14.11)$$

14.7 Definition. Let $M \in \mathcal{M}_T^{2,c}$ be a continuous L^2 martingale and $f \in \mathcal{E}_T$. Then

$$f \bullet M_T = \sum_{j=1}^n f(s_{j-1})(M(s_j) - M(s_{j-1})) \quad (14.12)$$

is called the *stochastic integral* of $f \in \mathcal{E}_T$. Instead of $f \bullet M_T$ we also write $\int_0^T f(s) dM_s$.

Ex. 14.2 Since (14.12) is linear, it is obvious that (14.12) does not depend on the particular representation of the simple process $f \in \mathcal{E}_T$. It is not hard to see that $(f \bullet M_{s_j})_{j=0,\dots,n}$ is a martingale transform in the sense of Definition 14.3: Just take $C_j = f(s_j) = \phi_j$ and $M_j = M(s_j)$.

In order to use the results of Section 14.1 we note that for a Brownian motion

$$\begin{aligned} \langle B \rangle_{s_j} - \langle B \rangle_{s_{j-1}} &\stackrel{(14.3)}{=} \mathbb{E}[(B(s_j) - B(s_{j-1}))^2 \mid \mathcal{F}_{s_{j-1}}] \\ &\stackrel{\substack{(B1) \\ (5.1)}}{=} \mathbb{E}[(B(s_j) - B(s_{j-1}))^2] = s_j - s_{j-1}. \end{aligned} \quad (14.13)$$

Since $\langle B \rangle_0 = 0$, and since $0 = s_0 < s_1 < \dots < s_n \leq T$ is an arbitrary partition, we see that $\langle B \rangle_t = t$. We can now use Theorem 14.4 to describe the properties of the stochastic integral for simple processes.

14.8 Theorem. Let $(B_t)_{t \geq 0}$ be a BM^1 , $(M_t)_{t \leq T} \in \mathcal{M}_T^{2,c}$, $(\mathcal{F}_t)_{t \geq 0}$ be an admissible filtration and $f \in \mathcal{E}_T$. Then

- a) $f \mapsto f \bullet M_T \in L^2(\mathbb{P})$ is linear;
- b) $\langle f \bullet B \rangle_T = f^2 \bullet \langle B \rangle_T = \int_0^T |f(s)|^2 ds$;
- c) For all $f \in \mathcal{E}_T$

$$\|f \bullet B_T\|_{L^2(\mathbb{P})}^2 = \mathbb{E}[(f \bullet B_T)^2] = \mathbb{E}\left[\int_0^T |f(s)|^2 ds\right] = \|f\|_{L^2(\lambda_T \otimes \mathbb{P})}^2. \quad (14.14)$$

Proof. Let f be a simple process given by (14.10). Without loss of generality we can assume that $s_n = T$. All assertions follow immediately from the corresponding statements for the martingale transform, cf. Theorem 14.4. For b) we observe that

$$\begin{aligned} \sum_{j=1}^n |f(s_{j-1})|^2 (\langle B \rangle_{s_j} - \langle B \rangle_{s_{j-1}}) &\stackrel{(14.13)}{=} \sum_{j=1}^n |f(s_{j-1})|^2 (s_j - s_{j-1}) \\ &\stackrel{(14.11)}{=} \int_0^{s_n} |f(s)|^2 ds, \end{aligned}$$

and c) follows if we take expectations and use the fact that $((f \bullet B_{s_j})^2 - \langle f \bullet B \rangle_{s_j})_{j=0,\dots,n}$ is a martingale, i. e. for $s_n = T$ we have $\mathbb{E}(f \bullet B_T^2) = \mathbb{E}\langle f \bullet B \rangle_T$. \square

Note that for $T < \infty$ we always have $\mathcal{E}_T \subset L^2(\lambda_T \otimes \mathbb{P})$, i. e. (14.14) is always finite and $f \bullet B_T \in L^2(\mathbb{P})$ always holds true. For $T = \infty$ this remains true if we assume that $f \in L^2(\lambda_T \otimes \mathbb{P})$ and $s_n < T = \infty$. Therefore one should understand the equality in Theorem 14.8 c) as an *isometry* between $(\mathcal{E}_T, \|\cdot\|_{L^2(\lambda_T \otimes \mathbb{P})})$ and a subspace of $L^2(\mathbb{P})$.

It is natural to study $T \mapsto f \bullet M_T = \int_0^T f(s) dM_s$ as a function of time. Since for every \mathcal{F}_t stopping time τ the stopped martingale $M^\tau = (M_{s \wedge \tau})_{s \leq T}$ is again in $\mathcal{M}_T^{2,c}$, see Theorem A.18 and Corollary A.20, we can define

$$f \bullet M_\tau := f \bullet (M^\tau)_T = \sum_{j=1}^n f(s_{j-1})(M_{\tau \wedge s_j} - M_{\tau \wedge s_{j-1}}). \quad (14.15)$$

As usual, we write

$$f \bullet M_\tau \stackrel{\text{def}}{=} \int_0^T f(s) dM_s^\tau = \int_0^\tau f(s) dM_s.$$

This holds, in particular, for the constant stopping time $\tau \equiv t, 0 \leq t \leq T$.

14.9 Theorem. *Let $(B_t)_{t \geq 0}$ be a BM^1 , $(M_t)_{t \leq T} \in \mathcal{M}_T^{2,c}$, $(\mathcal{F}_t)_{t \geq 0}$ be an admissible filtration and $f \in \mathcal{E}_T$. Then*

a) $(f \bullet M_t)_{t \leq T}$ is a continuous L^2 martingale; in particular

$$\mathbb{E}[(f \bullet B_t - f \bullet B_s)^2 | \mathcal{F}_s] = \mathbb{E}\left[\int_s^t |f(r)|^2 dr | \mathcal{F}_s\right], \quad s < t \leq T.$$

b) *Localization in time.* For every \mathcal{F}_t stopping time τ with finitely many values

$$(f \bullet M)_t^\tau = f \bullet (M^\tau)_t = (f \mathbb{1}_{[0, \tau)}) \bullet M_t.$$

c) *Localization in space.* If $f, g \in \mathcal{E}_T$ are simple processes with $f(s, \omega) = g(s, \omega)$ for all $(s, \omega) \in [0, t] \times F$ for some $t \leq T$ and $F \in \mathcal{F}_\infty$, then

$$(f \bullet M_s) \mathbb{1}_F = (g \bullet M_s) \mathbb{1}_F, \quad 0 \leq s \leq t.$$

Proof. a) Let $0 \leq s < t \leq T$ and $f \in \mathcal{E}_T$. We may assume that $s, t \in \{s_0, \dots, s_n\}$ where the s_j are as in (14.10); otherwise we add s and t to the partition. From Theorem 14.4 we know that $(f \bullet M_{s_j})_j$ is a martingale, thus

$$\mathbb{E}(f \bullet M_t | \mathcal{F}_s) = f \bullet M_s.$$

Since s, t are arbitrary, we conclude that $(f \bullet M_t)_{t \leq T} \in \mathcal{M}_T^2$.

The very definition of $t \mapsto f \bullet M_t$, see (14.15), shows that $(f \bullet M_t)_{t \leq T}$ is a continuous process. Using Lemma 14.2 or Theorem 14.4 d) we get

$$\begin{aligned} \mathbb{E}[(f \bullet B_t - f \bullet B_s)^2 | \mathcal{F}_s] &= \mathbb{E}[\langle f \bullet B \rangle_t - \langle f \bullet B \rangle_s | \mathcal{F}_s] \\ &\stackrel{14.8}{=} \mathbb{E}\left[\int_0^t |f(r)|^2 dr - \int_0^s |f(r)|^2 dr | \mathcal{F}_s\right]. \end{aligned}$$

b) Write $f \in \mathcal{E}_T$ in the form (14.11). Then we have for *all* stopping times τ

$$\begin{aligned} (f \bullet M)_t^\tau &= (f \bullet M)_{t \wedge \tau} = \sum_{j=1}^n f(s_{j-1})(M_{s_j \wedge \tau \wedge t} - M_{s_{j-1} \wedge \tau \wedge t}) \\ &= \sum_{j=1}^n f(s_{j-1})(M_{s_j \wedge t}^\tau - M_{s_{j-1} \wedge t}^\tau) = f \bullet (M^\tau)_t. \end{aligned}$$

If τ takes only finitely many values, we may assume that $\tau(\omega) \in \{s_0, s_1, \dots, s_n\}$ with s_j from the representation (14.10) of f . Thus, $\tau(\omega) > s_{j-1}$ implies $\tau(\omega) \geq s_j$, and

$$\begin{aligned} f(s) \mathbb{1}_{[0, \tau)}(s) &= \sum_{j=1}^n f(s_{j-1}) \mathbb{1}_{[0, \tau)}(s) \mathbb{1}_{[s_{j-1}, s_j)}(s) \\ &= \sum_{j=1}^n \underbrace{(f(s_{j-1}) \mathbb{1}_{\{\tau > s_{j-1}\}})}_{\mathcal{F}_{s_{j-1}} \text{ mble}} \underbrace{\mathbb{1}_{[s_{j-1}, \tau \wedge s_j)}(s)}_{\tau \wedge s_j = s_j}. \end{aligned}$$

This shows that $f \mathbb{1}_{[0, \tau)} \in \mathcal{E}_T$ as well as

$$\begin{aligned} (f \mathbb{1}_{[0, \tau)}) \bullet M_t &= \sum_{j=1}^n f(s_{j-1}) \mathbb{1}_{\{\tau > s_{j-1}\}} (M_{s_j \wedge t} - M_{s_{j-1} \wedge t}) \\ &= \sum_{j=1}^n f(s_{j-1}) \underbrace{(M_{s_j \wedge \tau \wedge t} - M_{s_{j-1} \wedge \tau \wedge t})}_{=0, \text{ if } \tau \leq s_{j-1}} = f \bullet (M^\tau)_t. \end{aligned}$$

c) Because of b) we can assume that $t = T$, otherwise we replace f, g by $f \mathbb{1}_{[0, t)}$ and $g \mathbb{1}_{[0, t)}$. Refining the partitions, if necessary, $f, g \in \mathcal{E}_T$ can have the same support

points $0 = s_0 < s_1 < \dots < s_n \leq T$ in their representation (14.10). Since we know that $f(s_j)\mathbb{1}_F = g(s_j)\mathbb{1}_F$, we get

$$\begin{aligned} (f \bullet M_T)\mathbb{1}_F &= \sum_{j=1}^n f(s_{j-1})\mathbb{1}_F \cdot (M_{s_j} - M_{s_{j-1}}) \\ &= \sum_{j=1}^n g(s_{j-1})\mathbb{1}_F \cdot (M_{s_j} - M_{s_{j-1}}) = (g \bullet M_T)\mathbb{1}_F. \quad \square \end{aligned}$$

14.3 Extension of the stochastic integral to \mathcal{L}_T^2

We have seen in the previous section that the map $f \mapsto f \bullet B$ is a linear isometry from \mathcal{E}_T to $\mathcal{M}_T^{2,c}$. Therefore, we can use standard arguments to extend this mapping to the closure of \mathcal{E}_T with respect to the norm in $L^2(\lambda_T \otimes \mathbb{P})$. As before, most results remain valid for $T = \infty$ and $[0, \infty)$ but, for simplicity, we restrict ourselves to the case $T < \infty$.

14.10 Lemma. *Let $M \in \mathcal{M}_T^{2,c}$. Then $\|M\|_{\mathcal{M}_T^2} := (\mathbb{E} [\sup_{s \leq T} |M_s|^2])^{1/2}$ is a norm on $\mathcal{M}_T^{2,c}$ and $\mathcal{M}_T^{2,c}$ is closed under this norm. Moreover,*

$$\mathbb{E} \left[\sup_{s \leq T} |M_s|^2 \right] \asymp \mathbb{E} [|M_T|^2] = \sup_{s \leq T} \mathbb{E} [|M_s|^2]. \quad (14.16)$$


Proof. By Doob's maximal inequality A.10 we get that

$$\mathbb{E} [|M_T|^2] \leq \mathbb{E} \left[\sup_{s \leq T} |M_s|^2 \right] \stackrel{(A.14)}{\leq} 4 \sup_{s \leq T} \mathbb{E} [|M_s|^2] \stackrel{(A.14)}{=} 4 \mathbb{E} [|M_T|^2].$$

This proves (14.16) since $(|M_t|^2)_{t \leq T}$ is a submartingale, i.e. $t \mapsto \mathbb{E}(|M_t|^2)$ is increasing. Since $\|\cdot\|_{\mathcal{M}_T^2}$ is the $L^2(\Omega, \mathbb{P}; \mathcal{C}[0, T])$ -norm, i.e. the $L^2(\mathbb{P})$ norm (of the supremum norm) of $\mathcal{C}[0, T]$ -valued random variables, it is clear that it is again a norm. Ex. 14.3

Let $M_n \in \mathcal{M}_T^{2,c}$ with $\lim_{n \rightarrow \infty} \|M_n - M\|_{\mathcal{M}_T^2} = 0$. In particular we see that for a subsequence $n(j)$

$$\lim_{j \rightarrow \infty} \sup_{s \leq T} |M_{n(j)}(s, \omega) - M(s, \omega)| = 0 \quad \text{for almost all } \omega.$$

This shows that $s \mapsto M(s, \omega)$ is continuous. Since $L^2(\mathbb{P})$ - $\lim_{n \rightarrow \infty} M_n(s) = M(s)$ for every $s \leq T$, we find 

$$\int_F M(s) d\mathbb{P} = \lim_{n \rightarrow \infty} \int_F M_n(s) d\mathbb{P} \stackrel{\text{martingale}}{=} \lim_{n \rightarrow \infty} \int_F M_n(T) d\mathbb{P} = \int_F M(T) d\mathbb{P}$$

for all $s \leq T$ and $F \in \mathcal{F}_s$. Thus, M is a martingale, and $M \in \mathcal{M}_T^{2,c}$. □

Itô's isometry (14.14) allows us to extend the stochastic integral from the simple integrands $\mathcal{E}_T \subset L^2(\lambda_T \otimes \mathbb{P})$ to the closure $\overline{\mathcal{E}}_T$ in $L^2(\lambda_T \otimes \mathbb{P})$.

14.11 Definition. By \mathcal{L}_T^2 we denote the closure $\overline{\mathcal{E}}_T$ of the simple processes \mathcal{E}_T with respect to the norm in $L^2(\lambda_T \otimes \mathbb{P})$.

Using (14.14) we find for any sequence $(f_n)_{n \geq 1} \subset \mathcal{E}_T$ with limit $f \in \mathcal{L}_T^2$ that

$$\begin{aligned} \|f_n \bullet B_T - f_m \bullet B_T\|_{L^2(\mathbb{P})}^2 &= \mathbb{E} \left[\int_0^T |f_n(s, \cdot) - f_m(s, \cdot)|^2 ds \right] \\ &= \|f_n - f_m\|_{L^2(\lambda_T \otimes \mathbb{P})}^2 \xrightarrow{m, n \rightarrow \infty} 0. \end{aligned}$$

Therefore, for each $0 \leq t \leq T$, the limit

$$L^2(\mathbb{P})\text{-} \lim_{n \rightarrow \infty} f_n \bullet B_t \tag{14.17}$$

Ex. 14.4 exists and does not depend on the approximating sequence. By Lemma 14.10, (14.17) defines an element in $\mathcal{M}_T^{2,c}$. We will see later on (cf. Section 14.5) that \mathcal{L}_T^2 is the set of all (equivalence classes of) processes from $L^2(\lambda_T \otimes \mathbb{P})$ which have an progressively measurable representative.

14.12 Definition. Let $(B_t)_{t \geq 0}$ be a BM¹ and let $f \in \mathcal{L}_T^2$. Then the *stochastic integral* (or *Itô integral*) is defined as

$$f \bullet B_t = \int_0^t f(s) dB_s := L^2(\mathbb{P})\text{-} \lim_{n \rightarrow \infty} f_n \bullet B_t, \quad 0 \leq t \leq T, \tag{14.18}$$

where $(f_n)_{n \geq 1} \subset \mathcal{E}_T$ is any sequence approximating $f \in \mathcal{L}_T^2$ in $L^2(\lambda_T \otimes \mathbb{P})$.

The stochastic integral is constructed in such a way that all properties from Theorems 14.8 and 14.9 carry over to integrands from $f \in \mathcal{L}_T^2$.

Ex. 14.6 **14.13 Theorem.** Let $(B_t)_{t \geq 0}$ be a BM¹, $(\mathcal{F}_t)_{t \geq 0}$ be an admissible filtration and
Ex. 14.7 $f \in \mathcal{L}_T^2$. Then we have for all $t \leq T$

Ex. 14.8 a) $f \mapsto f \bullet B$ is a linear map from \mathcal{L}_T^2 into $\mathcal{M}_T^{2,c}$, i. e. $(f \bullet B_t)_{t \leq T}$ is a continuous L^2 martingale;

- b) $(f \bullet B)^2 - \langle f \bullet B \rangle$ is a martingale where $\langle f \bullet B \rangle_t := f^2 \bullet \langle B \rangle_t = \int_0^t |f(s)|^2 ds$;¹ in particular,

$$\mathbb{E}[(f \bullet B_t - f \bullet B_s)^2 | \mathcal{F}_s] = \mathbb{E}\left[\int_s^t |f(r)|^2 dr \middle| \mathcal{F}_s\right], \quad s < t \leq T. \quad (14.19)$$

- c) Itô's isometry.

$$\|f \bullet B_T\|_{L^2(\mathbb{P})}^2 = \mathbb{E}[(f \bullet B_T)^2] = \mathbb{E}\left[\int_0^T |f(s)|^2 ds\right] = \|f\|_{L^2(\lambda_T \otimes \mathbb{P})}^2. \quad (14.20)$$

- d) Maximal inequalities.

$$\mathbb{E}\left[\int_0^T |f(s)|^2 ds\right] \leq \mathbb{E}\left[\sup_{t \leq T} \left(\int_0^t f(s) dB_s\right)^2\right] \leq 4 \mathbb{E}\left[\int_0^T |f(s)|^2 ds\right]. \quad (14.21)$$

- e) Localization in time. Let τ be an \mathcal{F}_t stopping time. Then

$$(f \bullet B)_t^\tau = f \bullet (B^\tau)_t = (f \mathbb{1}_{[0, \tau)}) \bullet B_t \quad \text{for all } t \leq T.$$

- f) Localization in space. If $f, g \in \mathcal{L}_T^2$ are integrands such that $f(s, \omega) = g(s, \omega)$ for all $(s, \omega) \in [0, t] \times F$ where $t \leq T$ and $F \in \mathcal{F}_\infty$, then

$$(f \bullet B_s) \mathbb{1}_F = (g \bullet B_s) \mathbb{1}_F, \quad 0 \leq s \leq t.$$

Proof. Throughout the proof we fix $f \in \mathcal{L}_T^2$ and some sequence $f_n \in \mathcal{E}_T$, such that $L^2(\lambda_T \otimes \mathbb{P})\text{-}\lim_{n \rightarrow \infty} f_n = f$. Clearly, all statements are true for $f_n \bullet B$ and we only have to see that the L^2 limits preserve them.

- a) follows from Theorem 14.9, Lemma 14.10 and the linearity of L^2 limits.
b) note that

$$(f_n \bullet B_t)^2 - \langle f_n \bullet B \rangle_t = (f_n \bullet B_t)^2 - \int_0^t |f_n(s)|^2 ds$$

is a sequence of martingales, cf. Theorem 14.8 b). We are done if we can show that this is a Cauchy sequence in $L^1(\mathbb{P})$ since this convergence preserves the martingale

¹ Observe that we have defined the angle bracket $\langle \cdot \rangle$ only for discrete martingales. Therefore we have to define the quadratic variation here. For each fixed t this new definition is in line with 14.8 b).

property. We have

$$\begin{aligned} |\langle f_n \bullet B \rangle_t - \langle f_m \bullet B \rangle_t| &\leq \int_0^t ||f_n(s)|^2 - |f_m(s)|^2| ds \\ &= \int_0^t |f_n(s) + f_m(s)| \cdot |f_n(s) - f_m(s)| ds \\ &\leq \sqrt{\int_0^t |f_n(s) + f_m(s)|^2 ds} \sqrt{\int_0^t |f_n(s) - f_m(s)|^2 ds}. \end{aligned}$$

Taking expectations on both sides reveals that the second expression tends to zero while the first stays uniformly bounded. This follows from our assumption as $L^2(\lambda_T \otimes \mathbb{P})\text{-}\lim_{n \rightarrow \infty} f_n = f$. With essentially the same calculation we get that $L^2(\mathbb{P})\text{-}\lim_{n \rightarrow \infty} f_n \bullet B_t = f \bullet B_t$ implies $L^1(\mathbb{P})\text{-}\lim_{n \rightarrow \infty} (f_n \bullet B_t)^2 = (f \bullet B_t)^2$.

Ex. 14.9

Since both $f \bullet B$ and $(f \bullet B)^2 - f^2 \bullet \langle B \rangle$ are martingales, the technique used in the proof of Lemma 14.2 applies and yields (14.19).

c) is a consequence of the same equalities in Theorem 14.8 and the completion.

d) Apply Lemma 14.10 to the L^2 martingale $(f \bullet B_t)_{t \leq T}$ and use c):

$$\mathbb{E} \left[\sup_{t \leq T} \left(\int_0^t f(s) dB_s \right)^2 \right] \stackrel{(14.16)}{\asymp} \mathbb{E} \left[\left(\int_0^T f(s) dB_s \right)^2 \right] \stackrel{c)}{=} \mathbb{E} \left[\int_0^T |f(s)|^2 ds \right].$$

e) From the proof of Theorem 14.9 b) we know that $f_n \bullet (B^\tau)_t = (f_n \bullet B)_t^\tau$ holds for all stopping times τ and $f_n \in \mathcal{E}_T$. Using the maximal inequality d) we find

$$\begin{aligned} \mathbb{E} \left[\sup_{s \leq T} |f_n \bullet (B^\tau)_s - f_m \bullet (B^\tau)_s|^2 \right] &= \mathbb{E} \left[\sup_{s \leq T} |f_n \bullet B_{s \wedge \tau} - f_m \bullet B_{s \wedge \tau}|^2 \right] \\ &\leq \mathbb{E} \left[\sup_{s \leq T} |f_n \bullet B_s - f_m \bullet B_s|^2 \right] \\ &\stackrel{d)}{\leq} 4 \mathbb{E} \left[\int_0^T |f_n(s) - f_m(s)|^2 ds \right] \\ &\xrightarrow{m, n \rightarrow \infty} 0. \end{aligned}$$

Since $\mathcal{M}_T^{2,c}$ is complete, we conclude that

$$(f_n \bullet B)^\tau = f_n \bullet (B^\tau) \xrightarrow[n \rightarrow \infty]{\mathcal{M}_T^{2,c}} f \bullet (B^\tau).$$

On the other hand, $\lim_{n \rightarrow \infty} f_n \bullet B = f \bullet B$, hence $\lim_{n \rightarrow \infty} (f_n \bullet B)^\tau = (f \bullet B)^\tau$. This gives the first equality.

For the second equality we assume that τ is a *discrete* stopping time. Then we know from Theorem 14.9 b)

$$(f_n \mathbb{1}_{[0, \tau)}) \bullet B = (f_n \bullet B)^\tau \xrightarrow[n \rightarrow \infty]{\mathcal{M}_T^{2,c}} (f \bullet B)^\tau.$$

By the maximal inequality d) we get

$$\begin{aligned} & \mathbb{E} \left[\sup_{s \leq T} |(f_n \mathbb{1}_{[0, \tau)}) \bullet B_s - (f \mathbb{1}_{[0, \tau)}) \bullet B_s|^2 \right] \\ & \leq 4 \mathbb{E} \left[\int_0^T |f_n(s) - f(s)|^2 \underbrace{\mathbb{1}_{[0, \tau)}(s)}_{\leq 1} ds \right] \xrightarrow{n \rightarrow \infty} 0. \end{aligned}$$

This shows that $f \mathbb{1}_{[0, \tau)} \in \mathcal{L}_T^2$ and $\mathcal{M}_T^{2, c}$ - $\lim_{n \rightarrow \infty} (f_n \mathbb{1}_{[0, \tau)}) \bullet B = (f \mathbb{1}_{[0, \tau)}) \bullet B$.

Finally, let $\tau_j \downarrow \tau$ be discrete stopping times, e. g. $\tau_j = (\lfloor 2^j \tau \rfloor + 1)/2^j$ as in Lemma A.16. With calculations similar to those above we get

$$\begin{aligned} & \mathbb{E} \left[\sup_{s \leq T} |(f \mathbb{1}_{[0, \tau)}) \bullet B_s - (f \mathbb{1}_{[0, \tau_j)}) \bullet B_s|^2 \right] \\ & \leq 4 \mathbb{E} \left[\int_0^T |f(s) \mathbb{1}_{[0, \tau)}(s) - f(s) \mathbb{1}_{[0, \tau_j)}(s)|^2 ds \right] \\ & \leq 4 \mathbb{E} \left[\int_0^T |f(s) \mathbb{1}_{[\tau, \tau_j)}(s)|^2 ds \right] \end{aligned}$$

and the last expression tends to zero as $j \rightarrow \infty$ by dominated convergence. Since $f \bullet B_t^{\tau_j} \rightarrow f \bullet B^\tau$ is obvious (by continuity of the paths), the assertion follows.

f) Let $f_n, g_n \in \mathcal{E}_T$ be approximations of $f, g \in \mathcal{L}_T^2$. We will see in the proof of Theorem 14.20 below, that we can choose f_n, g_n in such a way that

$$f_n(s, \omega) = g_n(s, \omega) \quad \text{for all } n \geq 1 \quad \text{and} \quad (s, \omega) \in [0, t] \times F.$$

Since

$$\begin{aligned} \mathbb{E} \left[\sup_{s \leq t} |(f_n \bullet B_s) \mathbb{1}_F - (f \bullet B_s) \mathbb{1}_F|^2 \right] & \leq \mathbb{E} \left[\sup_{s \leq t} |f_n \bullet B_s - f \bullet B_s|^2 \right] \\ & \stackrel{\text{d)}}{\leq} \mathbb{E} \left[\int_0^t |f_n(s) - f(s)|^2 ds \right], \end{aligned}$$

we see that the corresponding assertion for step functions, Theorem 14.9 f), remains valid for the L^2 limit. \square

14.4 Evaluating Itô integrals

Using integrands from \mathcal{E}_T allows us to calculate only very few Itô integrals. For example, up to now we have the natural formula $\int_0^T \xi \cdot \mathbb{1}_{[a, b)}(s) dB_s = \xi \cdot (B_b - B_a)$ only for *bounded* \mathcal{F}_a measurable random variables ξ .

14.14 Lemma. Let $(B_t)_{t \geq 0}$ be a BM^1 , $0 \leq a < b \leq T$ and $\xi \in L^2(\mathbb{P})$ be \mathcal{F}_a measurable. Then $\xi \cdot \mathbb{1}_{[a,b)} \in \mathcal{L}_T^2$ and

$$\int_0^T \xi \cdot \mathbb{1}_{[a,b)}(s) dB_s = \xi \cdot (B_b - B_a).$$

Proof. Clearly, $[(-k) \vee \xi \wedge k] \mathbb{1}_{[a,b)} \in \mathcal{E}_T$ for every $k \geq 1$. Dominated convergence yields

$$\mathbb{E} \left[\int_0^T |(-k) \vee \xi \wedge k - \xi|^2 \mathbb{1}_{[a,b)}(s) ds \right] = (b-a) \mathbb{E} [|(-k) \vee \xi \wedge k - \xi|^2] \xrightarrow[k \rightarrow \infty]{} 0,$$

i. e. $\xi \cdot \mathbb{1}_{[a,b)} \in \mathcal{L}_T^2$. By the definition of the Itô integral we find

$$\begin{aligned} \int_0^T \xi \cdot \mathbb{1}_{[a,b)}(s) dB_s &= L^2(\mathbb{P})\text{-}\lim_{k \rightarrow \infty} \int_0^T ((-k) \vee \xi \wedge k) \mathbb{1}_{[a,b)}(s) dB_s \\ &= L^2(\mathbb{P})\text{-}\lim_{k \rightarrow \infty} ((-k) \vee \xi \wedge k)(B_b - B_a) \\ &= \xi \cdot (B_b - B_a). \end{aligned} \quad \square$$

Let us now discuss an example that shows how to calculate a stochastic integral explicitly.

Ex. 14.10 14.15 Example. Let $(B_t)_{t \geq 0}$ be a BM^1 and $T > 0$. Then

$$\int_0^T B_t dB_t = \frac{1}{2} (B_T^2 - T). \quad (14.22)$$

As a by-product of our proof we will see that $(B_t)_{t \leq T} \in \mathcal{L}_T^2$. Note that (14.22) differs from what we expect from calculus. If $u(t)$ is differentiable and $u(0) = 0$, we get

$$\int_0^T u(t) du(t) = \int_0^T u(t) u'(t) dt = \frac{1}{2} u^2(t) \Big|_{t=0}^T = \frac{1}{2} u^2(T),$$

i. e. we would expect $\frac{1}{2} B_T^2$ rather than $\frac{1}{2} (B_T^2 - T)$. On the other hand, the stochastic integral must be a martingale, see Theorem 14.13 a); and the correction ‘ $-T$ ’ turns the submartingale B_T^2 into a bone fide martingale.

Let us now verify (14.22). Pick any partition $\Pi = \{0 = s_0 < s_1 < \dots < s_n = T\}$ with mesh $|\Pi| = \max_{1 \leq j \leq n} (s_j - s_{j-1})$ and set

$$f^\Pi(t, \omega) := \sum_{j=1}^n B(s_{j-1}, \omega) \mathbb{1}_{[s_{j-1}, s_j)}(t).$$

Applying Lemma 14.14 n times, we get

$$f^\Pi \in \mathcal{L}_T^2 \quad \text{and} \quad \int_0^T f^\Pi(t) dB_t = \sum_{j=1}^n B(s_{j-1})(B(s_j) - B(s_{j-1})).$$

Observe that

$$\begin{aligned} \mathbb{E} \left[\int_0^T |f^\Pi(t) - B(t)|^2 dt \right] &= \sum_{j=1}^n \mathbb{E} \left[\int_{s_{j-1}}^{s_j} |B(s_{j-1}) - B(t)|^2 dt \right] \\ &= \sum_{j=1}^n \int_{s_{j-1}}^{s_j} \mathbb{E} [|B(s_{j-1}) - B(t)|^2] dt \\ &= \sum_{j=1}^n \int_{s_{j-1}}^{s_j} (t - s_{j-1}) dt \\ &= \frac{1}{2} \sum_{j=1}^n (s_j - s_{j-1})^2 \xrightarrow{|\Pi| \rightarrow 0} 0. \end{aligned}$$

Since \mathcal{L}_T^2 is by definition closed (w. r. t. the norm in $L^2(\lambda_T \otimes \mathbb{P})$), we conclude that $(B_t)_{t \leq T} \in \mathcal{L}_T^2$. By Itô's isometry (14.20)

$$\begin{aligned} \mathbb{E} \left[\left(\int_0^T (f^\Pi(t) - B_t) dB_t \right)^2 \right] &= \mathbb{E} \left[\int_0^T |f^\Pi(t) - B(t)|^2 dt \right] \\ &= \frac{1}{2} \sum_{j=1}^n (s_j - s_{j-1})^2 \xrightarrow{|\Pi| \rightarrow 0} 0. \end{aligned}$$

Therefore,

$$\int_0^T B_t dB_t = L^2(\mathbb{P})\text{-}\lim_{|\Pi| \rightarrow 0} \sum_{j=1}^n B(s_{j-1})(B(s_j) - B(s_{j-1})),$$

i. e. we can approximate the stochastic integral in a concrete way by Riemann–Stieltjes sums. Finally,

$$\begin{aligned} B_T^2 &= \sum_{j=1}^n (B_{s_j} - B_{s_{j-1}})^2 + 2 \sum_{j=1}^n \sum_{k=1}^{j-1} (B_{s_j} - B_{s_{j-1}})(B_{s_k} - B_{s_{k-1}}) \\ &= \underbrace{\sum_{j=1}^n (B_{s_j} - B_{s_{j-1}})^2}_{\rightarrow T, |\Pi| \rightarrow 0, \text{ Thm. 9.1}} + 2 \underbrace{\sum_{j=1}^n (B_{s_j} - B_{s_{j-1}}) B_{s_{j-1}}}_{\rightarrow \int_0^T B_t dB_t, |\Pi| \rightarrow 0} \end{aligned}$$

in $L^2(\mathbb{P})$ and, therefore, $B_T^2 - T = 2 \int_0^T B_t dB_t$.

In the calculation of Example 14.15 we used (implicitly) the following result which is interesting on its own.

Ex. 14.12 14.16 Proposition. *Let $f \in \mathcal{L}_T^2$ be a process which is, as a function of t , mean-square continuous, i. e.*

$$\lim_{[0,T] \ni s \rightarrow t} \mathbb{E} [|f(s, \cdot) - f(t, \cdot)|^2] = 0, \quad t \in [0, T].$$

Then

$$\int_0^T f(t) dB_t = L^2(\mathbb{P})\text{-}\lim_{|\Pi| \rightarrow 0} \sum_{j=1}^n f(s_{j-1})(B_{s_j} - B_{s_{j-1}})$$

holds for any sequence of partitions $\Pi = \{s_0 = 0 < s_1 < \dots < s_n = T\}$ with mesh $|\Pi| \rightarrow 0$.

Proof. As in Example 14.15 we consider the following discretization of the integrand f

$$f^\Pi(t, \omega) := \sum_{j=1}^n f(s_{j-1}, \omega) \mathbb{1}_{[s_{j-1}, s_j)}(t).$$

Since $f(t) \in L^2(\mathbb{P})$, we can use Lemma 14.14 to see $f^\Pi \in \mathcal{L}^2$. By Itô's isometry,

$$\begin{aligned} & \mathbb{E} \left[\left| \int_0^T [f(t) - f^\Pi(t)] dB_t \right|^2 \right] \\ &= \int_0^T \mathbb{E} [|f(t) - f^\Pi(t)|^2] dt \\ &= \sum_{j=1}^n \int_{s_{j-1}}^{s_j} \mathbb{E} [|f(t) - f(s_{j-1})|^2] dt \\ &\leq \sum_{j=1}^n \int_{s_{j-1}}^{s_j} \underbrace{\sup_{u, v \in [s_{j-1}, s_j]} \mathbb{E} [|f(u) - f(v)|^2]}_{\rightarrow 0, |\Pi| \rightarrow 0} dt \xrightarrow{|\Pi| \rightarrow 0} 0. \end{aligned}$$

Combining this with Lemma 14.14 finally gives

$$\sum_{j=1}^n f(s_{j-1}) \cdot (B_{s_j} - B_{s_{j-1}}) = \int_0^T f^\Pi(t) dB_t \xrightarrow[|\Pi| \rightarrow 0]{L^2(\mathbb{P})} \int_0^T f(t) dB_t. \quad \square$$

14.5 What is the closure of \mathcal{E}_T ?

Let us now see how big $\mathcal{L}_T^2 = \overline{\mathcal{E}_T}$ really is. For this we need some preparations

14.17 Definition. Let $(B_t)_{t \geq 0}$ be a BM¹ and $(\mathcal{F}_t)_{t \geq 0}$ an admissible filtration. The *progressive σ -algebra* \mathcal{P} is the family of all sets $\Gamma \subset [0, T] \times \Omega$ such that

Ex. 14.14

Ex. 14.15

$$\Gamma \cap ([0, t] \times \Omega) \in \mathcal{B}[0, t] \otimes \mathcal{F}_t \quad \text{for all } t \leq T. \quad (14.23)$$

By $L_{\mathcal{P}}^2(\lambda_T \otimes \mathbb{P})$ we denote those equivalence classes in $L^2([0, T] \times \Omega, \lambda \otimes \mathbb{P})$ which have a \mathcal{P} measurable representative. Ex. 14.16

In Lemma 6.24 we have already seen that Brownian motion itself is \mathcal{P} measurable.

14.18 Lemma. Let $T > 0$. Then $(L_{\mathcal{P}}^2(\lambda_T \otimes \mathbb{P}), \|\cdot\|_{L^2(\lambda_T \otimes \mathbb{P})})$ is a Hilbert space.

Proof. By definition, $L_{\mathcal{P}}^2(\lambda_T \otimes \mathbb{P}) \subset L^2(\lambda_T \otimes \mathbb{P})$ is a subspace and $L^2(\lambda_T \otimes \mathbb{P})$ is a Hilbert space for the norm $\|\cdot\|_{L^2(\lambda_T \otimes \mathbb{P})}$. It is, therefore, enough to show that for any sequence $(f_n)_{n \geq 0} \subset L_{\mathcal{P}}^2(\lambda_T \otimes \mathbb{P})$ such that $L^2(\lambda_T \otimes \mathbb{P})\text{-}\lim_{n \rightarrow \infty} f_n = f$, the limit f has a \mathcal{P} measurable representative. We may assume that the processes f_n are \mathcal{P} measurable. Since f_n converges in L^2 , there exists a subsequence $(f_{n(j)})_{j \geq 1}$ such that $\lim_{j \rightarrow \infty} f_{n(j)} = f$ $\lambda \otimes \mathbb{P}$ almost everywhere. Therefore, the upper limit $\phi(t, \omega) := \overline{\lim_{j \rightarrow \infty} f_{n(j)}(t, \omega)}$ is a \mathcal{P} measurable representative of f . \square

Our aim is to show that $\mathcal{L}_T^2 = L_{\mathcal{P}}^2(\lambda_T \otimes \mathbb{P})$. For this we need a simple deterministic lemma. As before we set $L^2(\lambda_T) = L^2([0, T], \mathcal{B}[0, T], \lambda)$.

14.19 Lemma. Let $\phi \in L^2(\lambda_T)$ and $T > 0$. Then

$$\Theta_n[\phi](t) := \sum_{j=1}^{n-1} \xi_{j-1} \mathbb{1}_{[t_{j-1}, t_j)}(t) \quad (14.24)$$

where

$$t_j = \frac{jT}{n}, \quad j = 0, \dots, n, \quad \xi_0 = 0 \quad \text{and} \quad \xi_{j-1} = \frac{n}{T} \int_{t_{j-2}}^{t_{j-1}} \phi(s) ds, \quad j \geq 2,$$

is a sequence of step functions with $L^2(\lambda_T)\text{-}\lim_{n \rightarrow \infty} \Theta_n[\phi] = \phi$.

Proof. If ϕ is a step function, it has only finitely many jumps, say $\sigma_1 < \sigma_2 < \dots < \sigma_m$, and set $\sigma_0 := 0$. Choose n so large that in each interval $[t_{j-1}, t_j)$ there is at most one jump σ_k . From the definition of $\Theta_n[\phi]$ it is easy to see that ϕ and $\Theta_n[\phi]$ differ exactly in the first, the last and in those intervals $[t_{j-1}, t_j)$ which contain some σ_k or whose

predecessor $[t_{j-2}, t_{j-1})$ contains some σ_k . Formally,

$$\phi(t) = \Theta_n[\phi](t) \quad \text{for all } t \notin \bigcup_{j=0}^m \left[\sigma_j - \frac{T}{n}, \sigma_j + \frac{2T}{n} \right) \cup \left[0, \frac{T}{n} \right) \cup \left[\frac{n-1}{n}T, T \right].$$

This is shown in Figure 14.1.

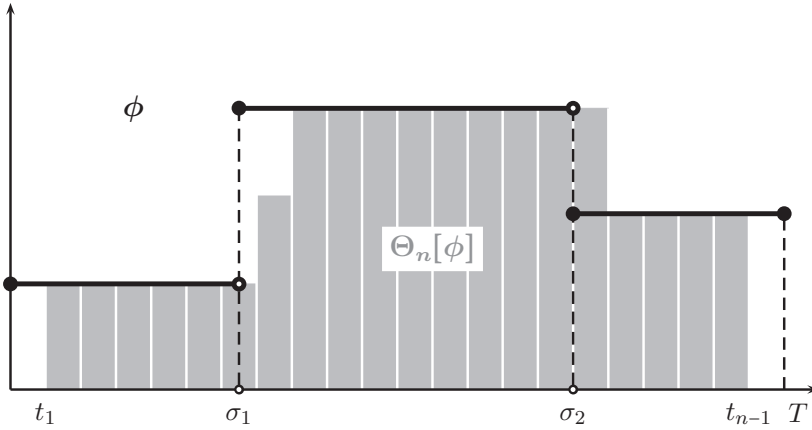


Figure 14.1. Approximation of a simple function by a left-continuous simple function.

Therefore, $\lim_{n \rightarrow \infty} \Theta_n[\phi] = \phi$ Lebesgue almost everywhere and in $L^2(\lambda_T)$.

Since the step functions are dense in $L^2(\lambda_T)$, we get for every $\phi \in L^2(\lambda_T)$ a sequence of step functions $L^2(\lambda_T)$ - $\lim_{k \rightarrow \infty} \phi^k = \phi$. Now let $\Theta_n[\phi]$ and $\Theta_n[\phi^k]$ be the step functions defined in (14.24). Then $\Theta_n[\cdot]$ is a contraction,

$$\begin{aligned} \int_0^T |\Theta_n[\phi](s)|^2 ds &= \sum_{j=1}^{n-1} \xi_{j-1}^2 (t_j - t_{j-1}) \\ &= \sum_{j=2}^{n-1} (t_j - t_{j-1}) \left[\frac{1}{t_{j-1} - t_{j-2}} \int_{t_{j-2}}^{t_{j-1}} \phi(s) ds \right]^2 \\ &\stackrel{\text{Jensen}}{\leq} \sum_{j=2}^{n-1} \underbrace{\frac{t_j - t_{j-1}}{t_{j-1} - t_{j-2}}}_{=1} \int_{t_{j-2}}^{t_{j-1}} |\phi(s)|^2 ds \\ &\leq \int_0^T |\phi(s)|^2 ds, \end{aligned}$$

and, by the triangle inequality and the contractivity of $\Theta_n[\cdot]$,

$$\begin{aligned} \|\phi - \Theta_n[\phi]\|_{L^2} &\leq \|\phi - \phi^k\|_{L^2} + \|\phi^k - \Theta_n[\phi^k]\|_{L^2} + \|\Theta_n[\phi^k] - \Theta_n[\phi]\|_{L^2} \\ &= \|\phi - \phi^k\|_{L^2} + \|\phi^k - \Theta_n[\phi^k]\|_{L^2} + \|\Theta_n[\phi^k - \phi]\|_{L^2} \\ &\leq \|\phi - \phi^k\|_{L^2} + \|\phi^k - \Theta_n[\phi^k]\|_{L^2} + \|\phi^k - \phi\|_{L^2}. \end{aligned}$$

Letting $n \rightarrow \infty$ (with k fixed) we get $\overline{\lim}_{n \rightarrow \infty} \|\phi - \Theta_n[\phi]\|_{L^2} \leq 2\|\phi^k - \phi\|_{L^2}$; now $k \rightarrow \infty$ shows that the whole expression tends to zero. \square

14.20 Theorem. *The $L^2(\lambda_T \otimes \mathbb{P})$ closure of \mathcal{E}_T is $L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$.*

Proof. Write $\mathcal{L}_T^2 = \overline{\mathcal{E}_T}$. Obviously, $\mathcal{E}_T \subset L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$. By Lemma 14.18, the set $L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$ is closed under the norm $\|\cdot\|_{\mathcal{L}_T^2}$, thus $\overline{\mathcal{E}_T} = \mathcal{L}_T^2 \subset L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$. All that remains to be shown is $L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P}) \subset \mathcal{L}_T^2$.

Let $f(t, \omega)$ be (the predictable representative of) a process in $L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$. By dominated convergence we see that the truncated process $f_C := -C \vee f \wedge C$ converges in $L^2(\lambda_T \otimes \mathbb{P})$ to f as $C \rightarrow \infty$. Therefore, we can assume that f is bounded.

Define for every ω an elementary process $\Theta_n[f](t, \omega)$, where Θ_n is as in Lemma 14.19, i.e.

$$\xi_{j-1}(\omega) := \begin{cases} \frac{n}{T} \int_{t_{j-2}}^{t_{j-1}} f(s, \omega) ds, & \text{if the integral exists} \\ 0, & \text{else.} \end{cases}$$

Note that the set of ω for which the integral does not exist is an $\mathcal{F}_{t_{j-1}}$ measurable null set, since $f(\cdot, \omega)$ is \mathcal{P} measurable and

$$\mathbb{E} \left[\int_{t_{j-2}}^{t_{j-1}} |f(s, \cdot)|^2 ds \right] \leq \mathbb{E} \left[\int_0^T |f(s, \cdot)|^2 ds \right] < \infty, \quad j \geq 2.$$

Thus, $\Theta_n[f] \in \mathcal{E}_T$. By the proof of Lemma 14.19, we have for \mathbb{P} almost all ω

$$\int_0^T |\Theta_n[f](t, \omega)|^2 dt \leq \int_0^T |f(t, \omega)|^2 dt, \quad (14.25)$$

$$\int_0^T |\Theta_n[f](t, \omega) - f(t, \omega)|^2 dt \xrightarrow{n \rightarrow \infty} 0. \quad (14.26)$$

If we take expectations in (14.26), we get by dominated convergence – the majorant is guaranteed by (14.25) – that $L^2(\lambda_T \otimes \mathbb{P})$ - $\lim_{n \rightarrow \infty} \Theta_n[f] = f$, thus $f \in \mathcal{L}_T^2$. \square

14.6 The stochastic integral for martingales

Up to now we have defined the stochastic integral with respect to a Brownian motion and for integrands in the closure of the simple processes $\mathcal{L}_T^2 = \overline{\mathcal{E}}_T$. For $f \in \mathcal{E}_T$, we can define the stochastic integral for any continuous L^2 martingale $M \in \mathcal{M}_T^{2,c}$, cf. Definition 14.6; only in the completion procedure, see Section 14.3, we used that $(B_t^2 - t, \mathcal{F}_t)_{t \leq T}$ is a martingale, that is $\langle B \rangle_t = t$. If we *assume*² that $(M_t, \mathcal{F}_t)_{t \leq T}$ is a continuous square-integrable martingale such that

$$(M_t^2 - A_t, \mathcal{F}_t)_{t \leq T} \text{ is a martingale for some} \quad (\star) \\ \text{adapted, continuous and increasing process } (A_t)_{t \leq T}$$

then we can define $f \bullet M$ for f from a suitable completion of \mathcal{E}_T .

As before, $(\mathcal{F}_t)_{t \leq T}$ is a complete filtration. Unless otherwise stated, we assume that all processes are adapted to this filtration and that all martingales are martingales w. r. t. this filtration. Let $M \in \mathcal{M}_T^{2,c}$ be such that (\star) holds and $0 = t_0 < t_1 < \dots < t_n = T$ be any partition of $[0, T]$. Then $(M_{t_j})_j$ is a martingale and (\star) shows, because of the uniqueness of the Doob decomposition, that $A_{t_j} = \langle M \rangle_{t_j}$. This allows us to use the proof of Theorem 14.8 c) and to see that the following Itô isometry holds for all simple processes $f \in \mathcal{E}_T$:

$$\|f \bullet M_T\|_{L^2(\mathbb{P})}^2 = \mathbb{E}[(f \bullet M)^2_T] = \mathbb{E}\left[\int_0^T |f(s)|^2 dA_s\right] =: \|f\|_{L^2(\mu_T \otimes \mathbb{P})}^2. \quad (14.27)$$

For fixed ω we denote by $\mu_T(\omega, ds) = dA_s(\omega)$ the measure on $[0, T]$ induced by the increasing function $s \mapsto A_s(\omega)$, i. e. $\mu_T(\omega, [s, t]) = A_t(\omega) - A_s(\omega)$ for all $s \leq t \leq T$.

This observation allows us to follow the procedure set out in Section 14.3 to define a stochastic integral for continuous square-integrable martingales.

14.21 Definition. Let $M \in \mathcal{M}_T^{2,c}$ satisfying (\star) and set $\mathcal{L}_T^2(M) := \overline{\mathcal{E}}_T$ where the closure is taken in the space $L^2(\mu_T \otimes \mathbb{P})$. Then the *stochastic integral* (or *Itô integral*) is defined as

$$f \bullet M_t = \int_0^t f(s) dM_s := L^2(\mathbb{P})\text{-}\lim_{n \rightarrow \infty} f_n \bullet M_t, \quad 0 \leq t \leq T, \quad (14.28)$$

where $(f_n)_{n \geq 1} \subset \mathcal{E}_T$ is any sequence which approximates $f \in \mathcal{L}_T^2(M)$ in the $L^2(\mu_T \otimes \mathbb{P})$ -norm.

Exactly the same arguments as in Section 14.3 can be used to show the analogue of Theorem 14.13.

² In fact, every martingale $M \in \mathcal{M}_T^{2,c}$ satisfies (\star) , see Section A.6 in the appendix.

14.22 Theorem. Let $(M_t, \mathcal{F}_t)_{t \leq T}$ be in $\mathcal{M}_T^{2,c}$ satisfying (\star) , write μ_T for the measure induced by the increasing process A_s , and let $f \in \mathcal{L}_T^2(M)$. Then we have for all $t \leq T$ Ex. 14.17

- a) $f \mapsto f \bullet M$ is a linear map from $\mathcal{L}_T^2(M)$ into $\mathcal{M}_T^{2,c}$, i. e. $(f \bullet M_t)_{t \leq T}$ is a continuous L^2 martingale;
- b) $(f \bullet M)^2 - f \bullet A$ is a martingale where $f \bullet A_t := \int_0^t |f(s)|^2 dA_s$; in particular

$$\mathbb{E}[(f \bullet M_t - f \bullet M_s)^2 | \mathcal{F}_s] = \mathbb{E}\left[\int_s^t |f(r)|^2 dA_r \mid \mathcal{F}_s\right], \quad s < t \leq T. \quad (14.29)$$

- c) Itô's isometry.

$$\|f \bullet M_T\|_{L^2(\mathbb{P})}^2 = \mathbb{E}[(f \bullet M_T)^2] = \mathbb{E}\left[\int_0^T |f(s)|^2 dA_s\right] = \|f\|_{L^2(\mu_T \otimes \mathbb{P})}^2. \quad (14.30)$$

- d) Maximal inequalities.

$$\mathbb{E}\left[\int_0^T |f(s)|^2 dA_s\right] \leq \mathbb{E}\left[\sup_{t \leq T} \left(\int_0^t f(s) dM_s\right)^2\right] \leq 4 \mathbb{E}\left[\int_0^T |f(s)|^2 dA_s\right]. \quad (14.31)$$

- e) Localization in time. Let τ be an \mathcal{F}_t stopping time. Then

$$(f \bullet M)_t^\tau = f \bullet (M^\tau)_t = (f \mathbb{1}_{[0, \tau)}) \bullet M_t \quad \text{for all } t \leq T.$$

- f) Localization in space. If $f, g \in \mathcal{L}_T^2(M)$ are such that $f(s, \omega) = g(s, \omega)$ for all $(s, \omega) \in [0, t] \times F$ for some $t \leq T$ and $F \in \mathcal{F}_\infty$, then

$$(f \bullet M_s) \mathbb{1}_F = (g \bullet M_s) \mathbb{1}_F, \quad 0 \leq s \leq t.$$

The characterization of $\overline{\mathcal{E}}_T$, see Section 14.5, relies on a deterministic approximation result, Lemma 14.19, and this remains valid in the general martingale setting.

14.23 Theorem. The $L^2(\mu_T \otimes \mathbb{P})$ closure of \mathcal{E}_T is $L_{\mathbb{P}}^2(\mu_T \otimes \mathbb{P})$, i. e. the family of all equivalence classes in $L^2(\mu_T \otimes \mathbb{P})$ which have a progressively measurable representative.

Sketch of the proof. Replace in Lemma 14.19 the measure λ_T by μ_T and the partition points t_j by the random times $\tau_j(\omega) := \inf\{s \geq 0 : A_s(\omega) > jT/n\} \wedge T$, $j \geq 1$, and $\tau_0(\omega) = 0$. The continuity of $s \mapsto A_s$ ensures that $A_{\tau_j} - A_{\tau_{j-1}} = T/n$ for all

$1 \leq j \leq \lfloor nA_T/T \rfloor$. Therefore, all arguments in the proof of Lemma 14.19 remain valid. Moreover, $\{\tau_j \leq t\} = \{A_t > jT/n\} \in \mathcal{F}_t$, i.e. the τ_j are stopping times. Since $L^2_{\mathcal{P}}(\mu_T \otimes \mathbb{P})$ is closed (Lemma 14.18), the inclusion $\bar{\mathcal{E}}_T \subset L^2_{\mathcal{P}}(\mu_T \otimes \mathbb{P})$ is clear. For the other direction we can argue as in Theorem 14.20 and see that every $f \in L^2_{\mathcal{P}}(\mu_T \otimes \mathbb{P})$ can be approximated by a sequence of random step functions of the form

$$\sum_{j=1}^{\infty} \xi_{j-1}(\omega) \mathbb{1}_{[\tau_{j-1}(\omega), \tau_j(\omega))}(t), \quad \xi_{j-1} \text{ is } \mathcal{F}_{\tau_{j-1}} \text{ mble, } |\xi_{j-1}| \leq C.$$

We are done if we can show that any process $\xi(\omega) \mathbb{1}_{[\sigma(\omega), \tau(\omega))}(t)$, where $\sigma \leq \tau$ are stopping times and ξ is a bounded and \mathcal{F}_{σ} measurable random variable, can be approximated in $L^2(\mu_T \otimes \mathbb{P})$ by a sequence from \mathcal{E}_T .

Set $\sigma^m := \sum_{k=1}^{\lfloor mT \rfloor} \frac{k}{m} \mathbb{1}_{[(k-1)/m, k/m)}(\sigma) + \frac{\lfloor mT \rfloor}{m} \mathbb{1}_{[\lfloor mT \rfloor/m, T]}(\sigma)$ and define τ^m in the same way. Then

$$\begin{aligned} \xi(\omega) \mathbb{1}_{[\sigma^m(\omega), \tau^m(\omega))}(t) &= \sum_{k=1}^{\lfloor mT \rfloor} \xi(\omega) \mathbb{1}_{\{\sigma < \frac{k-1}{m} \leq \tau\}}(\omega) \mathbb{1}_{[\frac{k-1}{m}, \frac{k}{m})}(t) \\ &\quad + \xi(\omega) \mathbb{1}_{\{\sigma < \frac{\lfloor mT \rfloor}{m} \leq \tau\}}(\omega) \mathbb{1}_{[\frac{\lfloor mT \rfloor}{m}, T]}(t). \end{aligned}$$

Because for all Borel sets $B \subset \mathbb{R}$ and $k = 1, 2, \dots, \lfloor mT \rfloor + 1$

$$\underbrace{\left\{ \xi \in B \right\} \cap \left\{ \sigma \leq (k-1)/m \right\} \cap \left\{ (k-1)/m < \tau \right\}}_{\in \mathcal{F}_{(k-1)/m}} \in \mathcal{F}_{\sigma}$$

we see that $\xi(\omega) \mathbb{1}_{[\sigma^m(\omega), \tau^m(\omega))}(t)$ is in \mathcal{E}_T . Finally,

$$\mathbb{E} \left[\int_0^T |\xi \mathbb{1}_{[\sigma^m, \tau^m)}(t) - \xi \mathbb{1}_{[\sigma, \tau)}(t)|^2 dA_t \right] \leq 2C^2 [\mathbb{E}(A_{\tau^m} - A_{\tau}) + \mathbb{E}(A_{\sigma^m} - A_{\sigma})],$$

and this tends to zero as $m \rightarrow \infty$. □



14.24 Further reading. Most books consider stochastic integrals immediately for continuous martingales. A particularly nice and quite elementary presentation is [111]. The approach of [156] is short and mathematically elegant. For discontinuous martingales standard references are [80] who use a point-process perspective, and [151] who starts out with semi-martingales as ‘good’ stochastic integrators. [109] and [110] construct the stochastic integral using random orthogonal measures. This leads to a unified treatment of Brownian motion and stationary processes.

- [80] Ikeda, Watanabe: *Stochastic Differential Equations and Diffusion Processes*.
 [109] Krylov: *Introduction to the Theory of Diffusion Processes*.
 [110] Krylov: *Introduction to the Theory of Random Processes*.
 [111] Kunita: *Stochastic Flows and Stochastic Differential Equations*.
 [151] Protter: *Stochastic Integration and Differential Equations*.
 [156] Revuz, Yor: *Continuous Martingales and Brownian Motion*.

Problems

- Let $(M_n, \mathcal{F}_n)_{n \geq 0}$ and $(N_n, \mathcal{F}_n)_{n \geq 0}$ be L^2 martingales. Show that the process $(M_n N_n - \langle M, N \rangle_n, \mathcal{F}_n)_{n \geq 0}$ is a martingale.
- Let $f \in \mathcal{E}_T$ be a simple process and $M \in \mathcal{M}_T^{2,c}$. Show that the definition of the stochastic integral $\int_0^T f(s) dM_s$ (cf. Definition 14.7) does not depend on the particular representation of the simple process (14.10).
- Show that $\|M\|_{\mathcal{M}_T^2} := (\mathbb{E}[\sup_{s \leq T} |M_s|^2])^{1/2}$ is a norm in the family \mathcal{M}_T^2 of equivalence classes of L^2 martingales (M and \widetilde{M} are equivalent if, and only if, $\sup_{0 \leq t \leq T} |M_t - \widetilde{M}_t| = 0$ a. s.).
- Show that the limit (14.17) does not depend on the approximating sequence.
- Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM^1 and τ a stopping time. Find $\langle B^\tau \rangle_t$.
- Let $(B_t)_{t \geq 0}$ be a BM^1 and $f, g \in \mathcal{L}_T^2$. Show that
 - $\mathbb{E}(f \bullet B_t \cdot g \bullet B_t | \mathcal{F}_s) = \mathbb{E}[\int_s^t f(u, \cdot) g(u, \cdot) du | \mathcal{F}_s]$ if f, g vanish on $[0, s]$.
 - $\mathbb{E}[f \bullet B_t | \mathcal{F}_s] = 0$ if f vanishes on $[0, s]$.
 - $f(t, \omega) = 0$ for all $\omega \in A \in \mathcal{F}_T$ and all $t \leq T$ implies $f \bullet B(\omega) = 0$ for all $\omega \in A$.
- Let $(B_t)_{t \geq 0}$ be a BM^1 and $(f_n)_{n \geq 1}$ a sequence in \mathcal{L}_T^2 such that $f_n \rightarrow f$ in $L^2(\lambda_T \otimes \mathbb{P})$. Then $f_n \bullet B \rightarrow f \bullet B$ in $\mathcal{M}_T^{2,c}$.
- Use the fact that any continuous, square-integrable martingale with bounded variation paths is constant (cf. Theorem A.21) to show the following:
 $\langle f \rangle \bullet B_t := \int_0^t |f(s)|^2 ds$ is the unique continuous and increasing process such that $(f \bullet B)^2 - f^2 \bullet \langle B \rangle$ is a martingale.
- Let $(X_n)_{n \geq 0}$ be a sequence of r. v. in $L^2(\mathbb{P})$. Show that $L^2\text{-}\lim_{n \rightarrow \infty} X_n = X$ implies $L^1\text{-}\lim_{n \rightarrow \infty} X_n^2 = X^2$.
- Let $(B_t)_{t \geq 0}$ be a BM^1 . Show that $\int_0^T B_s^2 dB_s = \frac{1}{3} B_t^3 - \int_0^T B_s ds, T > 0$.

11. Let $(B_t)_{t \geq 0}$ be a BM¹ and $f \in \text{BV}[0, T]$, $T < \infty$, a non-random function. Prove that $\int_0^T f(s) dB_s = f(T)B_T - \int_0^T B_s df(s)$ (the latter integral is understood as the mean-square limit of Riemann–Stieltjes sums). Conclude from this that the Itô integral extends the Paley–Wiener–Zygmund integral, cf. Paragraph 12.1.
12. Adapt the proof of Proposition 14.16 and prove that we also have

$$\lim_{|\Pi| \rightarrow 0} \mathbb{E} \left[\sup_{t \leq T} \left| \int_0^t f(s) dB_s - \sum_{j=1}^n f(s_{j-1})(B_{s_j} - B_{s_{j-1}}) \right|^2 \right] = 0.$$

13. Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM¹, $T < \infty$ and Π_n , $n \geq 1$, a sequence of partitions of $[0, T]$: $0 = t_{n,0} < t_{n,1} < \dots < t_{n,k(n)} = T$. Assume that $\lim_{n \rightarrow \infty} |\Pi_n| = 0$ and define intermediate points $\theta_{n,k}^\alpha := t_{n,k} + \alpha(t_{n,k+1} - t_{n,k})$ for some fixed $0 \leq \alpha \leq 1$. Show that the Riemann–Stieltjes sums

$$\sum_{k=1}^{k(n)-1} B(\theta_{n,k}^\alpha) (B(t_{n,k+1}) - B(t_{n,k}))$$

converge in $L^2(\mathbb{P})$. Determine the limit $L_T(\alpha)$. For which values of α is $L_t(\alpha)$, $0 \leq t \leq T$ a martingale?

14. Let $(B_t)_{t \geq 0}$ be a BM¹ and τ a stopping time. Show that $f(s, \omega) := \mathbb{1}_{[0, T \wedge \tau(\omega))}(s)$, $0 \leq s \leq T < \infty$ is progressively measurable. Use the very definition of the Itô integral to calculate $\int_0^T f(s) dB_s$ and compare the result with the localization principle of Theorem 14.13.

15. Show that

$$\mathcal{P} = \{ \Gamma : \Gamma \subset [0, T] \times \Omega, \Gamma \cap ([0, t] \times \Omega) \in \mathcal{B}[0, t] \otimes \mathcal{F}_t \text{ for all } t \leq T \}$$

is a σ -algebra.

16. Show that every right- or left-continuous adapted process is \mathcal{P} measurable.
17. Show that the process $f \bullet A_t := \int_0^t |f(s)|^2 dA_s$ appearing in Theorem 14.22 b) is adapted.

Chapter 15

Stochastic integrals: beyond \mathcal{L}_T^2

Let $(B_t)_{t \geq 0}$ be a one-dimensional Brownian motion and \mathcal{F}_t an admissible, right-continuous complete filtration, e. g. $\mathcal{F}_t = \overline{\mathcal{F}}_t^B$, cf. Theorem 6.21. We have seen in Chapter 14 that the Itô integral $\int_0^T f(s) dB_s$ exists for all integrands $f \in \mathcal{L}_T^2 = L_{\mathcal{P}}^2(\lambda_T \otimes \mathbb{P})$, i. e. for all stochastic processes f which have a representative which is \mathcal{P} measurable, cf. Definition 14.17, and satisfy

$$\mathbb{E} \left[\int_0^T |f(s, \cdot)|^2 ds \right] < \infty. \quad (15.1)$$

It is possible to weaken (15.1) and to extend the family of possible integrands.

15.1 Definition. We say that $f \in \mathcal{L}_{T, \text{loc}}^2$, if there exists a sequence $(\sigma_n)_{n \geq 0}$ of \mathcal{F}_t stopping times such that $\sigma_n \uparrow \infty$ a. s. and $f \mathbb{1}_{[0, \sigma_n)} \in \mathcal{L}_T^2$ for all $n \geq 0$. The sequence $(\sigma_n)_{n \geq 0}$ is called a *localizing sequence*.

The family \mathcal{L}_T^0 consists of all \mathcal{P} measurable processes f such that

$$\int_0^T |f(s, \cdot)|^2 ds < \infty \quad \text{a. s.} \quad (15.2)$$

15.2 Lemma. Let $f \in \mathcal{L}_T^2$. Then $\int_0^t |f(s, \cdot)|^2 ds$ is \mathcal{F}_t measurable for all $t \leq T$. Ex. 15.1

Proof. By definition, $\mathcal{L}_T^2 = \overline{\mathcal{E}}_T$, see also Theorem 14.20, i. e. there is a sequence $(f_n)_{n \geq 0} \subset \mathcal{E}_T$ which converges to f in $L^2(\lambda_T \otimes \mathbb{P})$. In particular, for every $t \in [0, T]$ there is a subsequence $(f_{n(j)})_{j \geq 1}$ such that

$$\lim_{j \rightarrow \infty} \int_0^t |f_{n(j)}(s, \cdot)|^2 ds = \int_0^t |f(s, \cdot)|^2 ds \quad \text{a. s.}$$

It is, therefore, enough to prove that $\int_0^t |f(s, \cdot)|^2 ds$ is \mathcal{F}_t measurable for $f \in \mathcal{E}_T$. This is obvious, since elementary processes are of the form (14.11): We have

$$\int_0^t |f(s, \cdot)|^2 ds = \sum_{j=1}^N |f(s_{j-1}, \cdot)|^2 (s_j - s_{j-1})$$

and all terms under the sum are \mathcal{F}_t measurable. □

15.3 Lemma. $\mathcal{L}_T^2 \subset \mathcal{L}_T^0 \subset \mathcal{L}_{T,\text{loc}}^2$.

Proof. Since a square-integrable random variable is a.s. finite, the first inclusion is clear.

To see the other inclusion, we assume that $f \in \mathcal{L}_T^0$ and define

$$\sigma_n = \inf \left\{ t \leq T : \int_0^t |f(s, \cdot)|^2 ds > n \right\}, \quad n \geq 1, \quad (15.3)$$

($\inf \emptyset = \infty$). Clearly, $\sigma_n \uparrow \infty$. Moreover,

$$\sigma_n(\omega) > t \iff \int_0^t |f(s, \omega)|^2 ds \leq n.$$

By Lemma 15.2, the integral is \mathcal{F}_t measurable; thus, $\{\sigma_n \leq t\} = \{\sigma_n > t\}^c \in \mathcal{F}_t$. This means that σ_n is a stopping time.

Ex. 14.14 Note that $f\mathbb{1}_{[0, \sigma_n)}$ is \mathcal{P} measurable since it is the product of two \mathcal{P} measurable
Ex. 15.1 random variables. From

$$\mathbb{E} \left[\int_0^T |f(s)\mathbb{1}_{[0, \sigma_n)}(s)|^2 ds \right] = \mathbb{E} \left[\int_0^{\sigma_n} |f(s)|^2 ds \right] \leq \mathbb{E} n = n,$$

we conclude that $f\mathbb{1}_{[0, \sigma_n)} \in \mathcal{L}_T^2$. □

15.4 Lemma. Let $f \in \mathcal{L}_{T,\text{loc}}^2$ and set $f_n := f\mathbb{1}_{[0, \sigma_n)}$ for a localizing sequence $(\sigma_n)_{n \geq 0}$. Then

$$(f_n \bullet B)_t^{\sigma_m} = f_m \bullet B_t \quad \text{for all } m \leq n \text{ and } t \leq \sigma_m.$$

Proof. Using Theorem 14.13 e) for $f_n, f_m \in \mathcal{L}_T^2$ we get for all $t \leq \sigma_m$

$$\begin{aligned} (f_n \bullet B)_t^{\sigma_m} &= (f_n \mathbb{1}_{[0, \sigma_m)}) \bullet B_t = (f \mathbb{1}_{[0, \sigma_n)} \mathbb{1}_{[0, \sigma_m)}) \bullet B_t = (f \mathbb{1}_{[0, \sigma_m)}) \bullet B_t \\ &= f_m \bullet B_t. \end{aligned} \quad \square$$

Lemma 15.4 allows us to define the stochastic integral for integrands in $\mathcal{L}_{T,\text{loc}}^2$.

15.5 Definition. Let $f \in \mathcal{L}_{T,\text{loc}}^2$ with localizing sequence $(\sigma_n)_{n \geq 0}$. Then

$$f \bullet B_t(\omega) := (f \mathbb{1}_{[0, \sigma_n)}) \bullet B_t(\omega) \quad \text{for all } t \leq \sigma_n(\omega) \quad (15.4)$$

is called the (extended) stochastic integral or (extended) Itô integral. We also write $\int_0^t f(s) dB_s := f \bullet B_t$.

Lemma 15.4 guarantees that this definition is consistent and independent of the localizing sequence: Indeed for any two localizing sequences $(\sigma_n)_{n \geq 0}$ and $(\tau_n)_{n \geq 0}$ we

see easily that $\rho_n := \sigma_n \wedge \tau_n$ is again such a sequence and then the telescoping argument of Lemma 15.4 applies.

We can also localize the notion of a martingale.

15.6 Definition. A *local martingale* is an adapted right-continuous process $(M_t, \mathcal{F}_t)_{t \leq T}$ for which there exists a localizing sequence $(\sigma_n)_{n \geq 0}$ such that, for every $n \geq 0$, $(M_t^{\sigma_n} \mathbb{1}_{\{\sigma_n > 0\}}, \mathcal{F}_t)_{t \leq T}$ is a martingale.

We denote by $\mathcal{M}_{T, \text{loc}}^c$ the continuous local martingales and by $\mathcal{M}_{T, \text{loc}}^{2, c}$ those local martingales where $M^{\sigma_n} \mathbb{1}_{\{\sigma_n > 0\}} \in \mathcal{M}_T^{2, c}$.

Let us briefly review the properties of the extended stochastic integral. The following theorem should be compared with its \mathcal{L}_T^2 analogue, Theorem 14.13

15.7 Theorem. Let $(B_t)_{t \geq 0}$ be a BM^1 , $(\mathcal{F}_t)_{t \geq 0}$ an admissible filtration and $f \in \mathcal{L}_{T, \text{loc}}^2$. Then we have for all $t \leq T$

- a) $f \mapsto f \bullet B$ is a linear map from $\mathcal{L}_{T, \text{loc}}^2$ to $\mathcal{M}_{T, \text{loc}}^{2, c}$, i. e. $(f \bullet B_t)_{t \leq T}$ is a continuous local L^2 martingale;
- b) $(f \bullet B)^2 - \langle f \bullet B \rangle$ is a local martingale where

$$\langle f \bullet B \rangle_t := f^2 \bullet \langle B \rangle_t = \int_0^t |f(s)|^2 ds;$$

- c) $\langle \text{no equivalent of Itô's isometry} \rangle$
- d) $\mathbb{P}(\int_0^T f(s) dB_s > \epsilon) \leq \frac{C}{\epsilon^2} + \mathbb{P}(\int_0^T |f(s)|^2 ds > C)$ for all $\epsilon > 0$, $C > 0$.
- e) *Localization in time.* Let τ be a \mathcal{F}_t stopping time. Then

$$(f \bullet B)_t^\tau = f \bullet (B^\tau)_t = (f \mathbb{1}_{[0, \tau)}) \bullet B_t \text{ for all } t \leq T.$$

- f) *Localization in space.* If $f, g \in \mathcal{L}_{T, \text{loc}}^2$ are such that $f(s, \omega) = g(s, \omega)$ for all $(s, \omega) \in [0, t] \times F$ for some $t \leq T$ and $F \in \mathcal{F}_\infty$, then

$$(f \bullet B_s) \mathbb{1}_F = (g \bullet B_s) \mathbb{1}_F, \quad 0 \leq s \leq t.$$

Proof. Throughout the proof we fix some localizing sequence $(\sigma_n)_{n \geq 1}$.

We begin with the assertion e). By the definition of the extended Itô integral and Theorem 14.13 e) we have for all $n \geq 1$

$$\begin{aligned} ((f \mathbb{1}_{[0, \sigma_n)}) \bullet B)^\tau &= \underbrace{(f \mathbb{1}_{[0, \sigma_n)} \mathbb{1}_{[0, \tau)}) \bullet B}_{=((f \mathbb{1}_{[0, \tau)}) \mathbb{1}_{[0, \sigma_n)}) \bullet B} = (f \mathbb{1}_{[0, \sigma_n)}) \bullet (B^\tau). \end{aligned}$$

Now a) follows directly from its counterpart in Theorem 14.13. To see b), we note that by Theorem 14.13 b), applied to $f \mathbb{1}_{[0, \sigma_n)}$, $(f \bullet B^{\sigma_n})^2 - \int_0^{\sigma_n \wedge \bullet} |f(s)|^2 ds$, is a martingale for each $n \geq 1$.

For d) we define $\tau := \inf\{t \leq T : \int_0^t |f(s, \cdot)|^2 ds > C\}$. As in Lemma 15.3 we see that τ is a stopping time. Therefore we get by e)

$$\begin{aligned}
 \mathbb{P}(f \bullet B_T > \epsilon) &= \mathbb{P}(f \bullet B_T > \epsilon, T < \tau) + \mathbb{P}(f \bullet B_T > \epsilon, T \geq \tau) \\
 &\stackrel{e)}{\leq} \mathbb{P}((f \mathbb{1}_{[0, \tau)}) \bullet B_T > \epsilon) + \mathbb{P}(T \geq \tau) \\
 &\leq \frac{1}{\epsilon^2} \mathbb{E} \left[|(f \mathbb{1}_{[0, \tau)}) \bullet B_T|^2 \right] + \mathbb{P} \left(\int_0^T |f(s)|^2 ds > C \right) \\
 &= \frac{1}{\epsilon^2} \mathbb{E} \left[\int_0^{T \wedge \tau} |f(s)|^2 ds \right] + \mathbb{P} \left(\int_0^T |f(s)|^2 ds > C \right) \\
 &\leq \frac{C}{\epsilon^2} + \mathbb{P} \left(\int_0^T |f(s)|^2 ds > C \right),
 \end{aligned}$$

and d) follows.

Let $f, g \in \mathcal{L}_{T, \text{loc}}^2$ with localizing sequences $(\sigma_n)_{n \geq 0}$ and $(\tau_n)_{n \geq 0}$, respectively. Then $\sigma_n \wedge \tau_n$ localizes both f and g , and by Theorem 14.13 f),

$$(f \mathbb{1}_{[0, \sigma_n \wedge \tau_n)}) \bullet B \mathbb{1}_A = (g \mathbb{1}_{[0, \sigma_n \wedge \tau_n)}) \bullet B \mathbb{1}_A \quad \text{for all } n \geq 0.$$

Because of the definition of the extended Itô integral this is just f). \square

For the extended Itô integral we also have a Riemann–Stieltjes-type approximation, cf. Proposition 14.16.

15.8 Theorem. *Let f be an \mathcal{F}_t adapted right-continuous process with left limits. Then $f \in \mathcal{L}_{T, \text{loc}}^2$.*

For every sequence $(\Pi_k)_{k \geq 1}$ of partitions of $[0, T]$ such that $\lim_{k \rightarrow \infty} |\Pi_k| = 0$, the Riemann–Stieltjes sums

$$Y_t^{\Pi_k} := \sum_{s_{j-1}, s_j \in \Pi_k} f(s_j) (B_{s_{j+1} \wedge t} - B_{s_j \wedge t})$$

converge uniformly (on compact t -sets) in probability to $f \bullet B_t$, $t \leq T$:

$$\lim_{k \rightarrow \infty} \mathbb{P} \left(\sup_{0 \leq t \leq T} \left| Y_t^{\Pi_k} - \int_0^t f(s) dB_s \right| > \epsilon \right) = 0 \quad \text{for all } \epsilon > 0. \quad (15.5)$$

Proof. Let $f^{\Pi_k}(t) := \sum_{s_{j-1}, s_j \in \Pi_k} f(s_{j-1}) \mathbb{1}_{[s_{j-1}, s_j)}(t)$. From Lemma 14.14 we infer that $f^{\Pi_k} \in \mathcal{L}_{T, \text{loc}}^2$ and $Y_t^{\Pi_k} = f^{\Pi_k} \bullet B_t$.

Since f is right-continuous with left limits, it is bounded on compact t -sets. Therefore

$$\tau_n := \inf \{s \geq 0 : |f(s)|^2 > n\} \wedge n$$

is a sequence of finite stopping times with $\tau_n \uparrow \infty$ a. s. Moreover,

$$\mathbb{E} \left[\int_0^{\tau_n \wedge T} |f(s)|^2 ds \right] \leq \mathbb{E} \left[\int_0^{\tau_n \wedge T} n ds \right] \leq \mathbb{E} \left[\int_0^T n ds \right] \leq Tn,$$

i. e. τ_n is also a localizing sequence for the stochastic integral. Therefore, by Theorem 15.7 a) and e) and Theorem 14.13 c) and d),

$$\begin{aligned} \mathbb{E} \left[\sup_{t \leq T} \left| Y_{t \wedge \tau_n}^{\Pi_k} - \int_0^{t \wedge \tau_n} f(s) dB_s \right|^2 \right] &= \mathbb{E} \left[\sup_{t \leq T} |(f^{\Pi_k} - f) \bullet B_{t \wedge \tau_n}|^2 \right] \\ &\stackrel{(14.21)}{\leq} 4 \mathbb{E} \left[|(f^{\Pi_k} - f) \bullet B_{T \wedge \tau_n}|^2 \right] \\ &\stackrel{(14.20)}{=} 4 \mathbb{E} \left[\int_0^{\tau_n \wedge T} |f^{\Pi_k}(s) - f(s)|^2 ds \right]. \end{aligned}$$

For $0 \leq s \leq \tau_n \wedge T$ the integrand is bounded by $4n$ and converges, as $k \rightarrow \infty$, to 0. By dominated convergence we see that the integral tends to 0 as $k \rightarrow \infty$. Finally, f is right-continuous with finite left-hand limits and, therefore, the set of discontinuity points $\{s \in [0, T] : f(\omega, s) \neq f(\omega, s-)\}$ is a Lebesgue null set.

Since $\lim_{n \rightarrow \infty} \tau_n = \infty$ a. s., we see that

$$\forall \delta > 0 \quad \exists N_\delta \quad \forall n \geq N_\delta : \mathbb{P}(\tau_n < T) < \delta.$$

Thus, by Chebyshev's inequality, Theorem 15.7 e) and the above calculation

$$\begin{aligned} &\mathbb{P} \left(\sup_{t \leq T} \left| Y_t^{\Pi_k} - \int_0^t f(s) dB_s \right| > \epsilon \right) \\ &\leq \mathbb{P} \left(\sup_{t \leq T} \left| Y_t^{\Pi_k} - \int_0^t f(s) dB_s \right| > \epsilon, \tau_n \geq T \right) + \mathbb{P}(\tau_n < T) \\ &\leq \mathbb{P} \left(\sup_{t \leq T} \left| Y_{t \wedge \tau_n}^{\Pi_k} - \int_0^{t \wedge \tau_n} f(s) dB_s \right| > \epsilon \right) + \mathbb{P}(\tau_n < T) \\ &\leq \frac{4}{\epsilon^2} \mathbb{E} \left[\int_0^{T \wedge \tau_n} |f^{\Pi_k}(s) - f(s)|^2 ds \right] + \delta \xrightarrow[k \rightarrow \infty]{} \delta \xrightarrow[\delta \rightarrow 0]{} 0. \quad \square \end{aligned}$$



15.9 Further reading. Localized stochastic integrals are covered by most of the books mentioned in the *Further reading* section of Chapter 14. Theorem 15.8 has variants where the driving noise is an approximation of Brownian motion. Such *Wong–Zakai* results are important for practitioners. A good source are the original papers [190] and [79].

[79] Ikeda, Nakao, Yamato: A class of approximations of Brownian motion.

[190] Wong, Zakai: Riemann–Stieltjes approximations of stochastic integrals.

Problems

1. Show that the subsequence in the proof of Lemma 15.2 can be chosen independently of $t \in [0, T]$.
2. Let τ be a stopping time. Show that $\mathbb{1}_{[0, \tau)}$ is \mathcal{P} measurable.
3. If σ_n is localizing for a local martingale, so is $\tau_n := \sigma_n \wedge n$.

Chapter 16

Itô's formula

An important consequence of the fundamental theorem of integral and differential calculus is the fact every differentiation rule has an integration counterpart. Consider, for example, the chain rule

$$(f \circ g)'(t) = f'(g(t)) \cdot g'(t) \quad \text{or} \quad f(g(t)) - f(g(0)) = \int_0^t f'(g(s)) \cdot g'(s) ds$$

which is just the substitution rule for integrals. When dealing with differential equations it is often useful to rewrite the chain rule in differential form as

$$d(f \circ g) = f' \circ g \cdot dg.$$

Let $(B_t)_{t \geq 0}$ be a BM^1 and $(\mathcal{F}_t)_{t \geq 0}$ an admissible, right-continuous complete filtration, e. g. $\mathcal{F}_t = \overline{\mathcal{F}}_t^B$, cf. Theorem 6.21. Itô's formula is the stochastic counterpart of the chain rule; it is also known as the *change-of-variable-formula* or *Itô's Lemma*. The unbounded variation of Brownian sample paths leaves, again, its traces: The chain rule has a correction term involving the quadratic variation of $(B_t)_{t \geq 0}$ and the second derivative of the function f . Here is the statement of the basic result.

16.1 Theorem (Itô 1942). *Let $(B_t)_{t \geq 0}$ be a BM^1 and let $f : \mathbb{R} \rightarrow \mathbb{R}$ be a \mathcal{C}^2 -function. Then we have for all $t \geq 0$ almost surely*

$$f(B_t) - f(B_0) = \int_0^t f'(B_s) dB_s + \frac{1}{2} \int_0^t f''(B_s) ds. \quad (16.1)$$

Before we prove Theorem 16.1, we need to discuss Itô processes and Itô differentials.

16.1 Itô processes and stochastic differentials

Let $T \in [0, \infty]$. Stochastic processes of the form

$$X_t = X_0 + \int_0^t \sigma(s) dB_s + \int_0^t b(s) ds, \quad t < T, \quad (16.2)$$

are called *Itô processes*. If the integrands satisfy $\sigma \in \mathcal{L}_{T,\text{loc}}^2$ and, almost surely, $b(\cdot, \omega) \in L^1([0, T], ds)$ (or $b(\cdot, \omega) \in L_{\text{loc}}^1([0, \infty), ds)$ if $T = \infty$) then all integrals in (16.2) make sense. Roughly speaking, an Itô process is *locally* a Brownian motion with standard deviation $\sigma(s, \omega)$ plus a drift with mean value $b(s, \omega)$. Note that these parameters change infinitesimally over time.

In dimensions other than $d = 1$, the stochastic integral is *defined* for each coordinate and we can read (16.2) as matrix and vector equality, i. e.

$$X_t^j = X_0^j + \sum_{k=1}^d \int_0^t \sigma_{jk}(s) dB_s^k + \int_0^t b_j(s) ds, \quad t < T, \quad j = 1, \dots, m, \quad (16.3)$$

where $X_t = (X_t^1, \dots, X_t^m)^\top$, $B_t = (B_t^1, \dots, B_t^d)^\top$ is a d -dimensional Brownian motion, $\sigma(t, \omega) = (\sigma_{jk}(t, \omega))_{jk} \in \mathbb{R}^{m \times d}$ is a matrix-valued stochastic process and $b(t, \omega) = (b_1(t, \omega), \dots, b_m(t, \omega))^\top \in \mathbb{R}^m$ is a vector-valued stochastic process; the coordinate processes σ_{jk} and b_j are chosen in such a way that all (stochastic) integrals in (16.3) are defined.

16.2 Definition. Let $(B_t)_{t \geq 0}$ be a BM^d , $(\sigma_t)_{t \geq 0}$ and $(b_t)_{t \geq 0}$ be two progressively measurable $\mathbb{R}^{m \times d}$ and \mathbb{R}^m -valued locally bounded processes. Then

$$X_t = X_0 + \int_0^t \sigma_s dB_s + \int_0^t b_s ds, \quad t \geq 0,$$

is an m -dimensional *Itô process*.

16.3 Remark.

- a) If we know that σ and b are locally bounded, i. e. there is some $\Omega_0 \subset \Omega$ with $\mathbb{P}(\Omega_0) = 1$ and

$$\max_{j,k} \sup_{t \leq T} |\sigma_{jk}(t, \omega)| + \max_j \sup_{t \leq T} |b_j(t, \omega)| < \infty \quad \text{for all } T > 0, \omega \in \Omega_0,$$

then $b_j(\cdot, \omega) \in L^1([0, T], dt)$, $T > 0$, almost surely; moreover, the stopping times $\tau_n(\omega) := \min_{j,k} \inf \left\{ t \geq 0 : \int_0^t |\sigma_{jk}(s, \omega)|^2 ds > n \right\}$ satisfy $\tau_n \uparrow \infty$ almost surely, i. e.

$$\mathbb{E} \left[\int_0^{\tau_n \wedge T} |\sigma_{jk}(s, \cdot)|^2 ds \right] \leq n$$

which proves that $\sigma_{jk} \in \mathcal{L}_{T,\text{loc}}^2$ for all $T > 0$.

- b) From Chapter 15 we know that $X_t = M_t + A_t$ where M_t is a (vector-valued) continuous local martingale and the coordinates of the continuous process A_t satisfy

$$|A_t^j(\omega) - A_s^j(\omega)| \leq \int_s^t |b_j(s, \omega)| ds \leq C_T(\omega)(t - s) \quad \text{for all } s, t < T,$$

i. e. A_t is locally of bounded variation.

It is convenient to rewrite (16.2) in *differential form*, i. e. as

$$dX_t(\omega) = \sigma(t, \omega) dB_t(\omega) + b(t, \omega) dt \quad (16.4)$$

where we use again, if appropriate, vector and matrix notation.

In this new notation, Example 14.15 reads

$$(B_t)^2 = 2 \int_0^t B_s dB_s + t, \quad \text{hence,} \quad d(B_t)^2 = 2B_t dB_t + dt$$

and we see that this is a special case of Itô's formula (16.1),

$$df(B_t) = f'(B_t) dB_t + \frac{1}{2} f''(B_t) dt,$$

if we set $f(x) = x^2$.

16.2 The heuristics behind Itô's formula

Let us give an intuitive argument why stochastic calculus differs from the usual calculus. Note that

$$\Delta B_t = \sqrt{\Delta t} \frac{B_{t+\Delta t} - B_t}{\sqrt{\Delta t}} = \sqrt{\Delta t} G \quad \text{and} \quad (\Delta B_t)^2 = \Delta t G^2$$

where $G \sim \mathbf{N}(0, 1)$ is a standard normal random variable. Since $\mathbb{E}|G| = \sqrt{2/\pi}$ and $\mathbb{E}G^2 = 1$, we get from $\Delta t \rightarrow 0$

$$|dB_t| \approx \sqrt{\frac{2}{\pi}} dt \quad \text{and} \quad (dB_t)^2 \approx dt.$$

This means that Itô's formula is essentially a second-order Taylor formula while in ordinary calculus we only go up to order one. Terms of higher order, for example $(dB_t)^3 \approx (dt)^{3/2}$ and $dt dB_t \approx (dt)^{3/2}$, do not give any contributions since integrals against $(dt)^{3/2}$ correspond to an integral sum with $(\Delta t)^{-1}$ many terms which are all of the order $(\Delta t)^{3/2}$ – and such a sum tends to zero as $\Delta t \rightarrow 0$. The same argument also explains why we do not have terms of second order, i. e. $(dt)^2$, in ordinary calculus.

In Section 16.5 we consider Itô's formula for a d -dimensional Brownian motion. Then we will encounter differentials of the form $dt dB_t^k$ and $dB_t^j dB_t^k$. Clearly, $dt dB_t^k \approx (dt)^{3/2}$ is negligible, while for $j \neq k$

$$(B_{t+\Delta t}^j - B_t^j)(B_{t+\Delta t}^k - B_t^k) = \Delta t \frac{(B_{t+\Delta t}^j - B_t^j)}{\sqrt{\Delta t}} \frac{(B_{t+\Delta t}^k - B_t^k)}{\sqrt{\Delta t}} \sim \Delta t G G'$$

where G and G' are two independent standard normal $N(0, 1)$ distributed random variables. Since we have $\mathbb{E}(GG') = \mathbb{E} G \mathbb{E} G' = 0$ and $(dB_t^j)^2 \approx dt$, we get, as $\Delta t \rightarrow 0$,

$$dB_t^j dB_t^k \approx \delta_{jk} dt,$$

with Kronecker's delta: $\delta_{jk} = 0$ if $j \neq k$ and $\delta_{jk} = 1$ if $j = k$. This can be nicely summed up in a multiplication table.

Table 16.1. Multiplication table for stochastic differentials.

$j \neq k$	dt	dB_t^j	dB_t^k
dt	0	0	0
dB_t^j	0	dt	0
dB_t^k	0	0	dt

16.3 Proof of Itô's formula (Theorem 16.1)

1° Let us first assume that $\text{supp } f \subset [-K, K]$ is compact. Then

$$C_f := \|f\|_\infty + \|f'\|_\infty + \|f''\|_\infty < \infty.$$

For any $\Pi := \{t_0 = 0 < t_1 < \dots < t_n = t\}$ with mesh $|\Pi| = \max_j (t_j - t_{j-1})$ we get by Taylor's theorem

$$\begin{aligned} f(B_t) - f(B_0) &= \sum_{j=1}^n (f(B_{t_j}) - f(B_{t_{j-1}})) \\ &= \sum_{j=1}^n f'(B_{t_{j-1}})(B_{t_j} - B_{t_{j-1}}) + \frac{1}{2} \sum_{j=1}^n f''(\xi_j)(B_{t_j} - B_{t_{j-1}})^2 \\ &=: J_1 + \frac{1}{2} J_2 \end{aligned}$$

with intermediate points

$$\xi_j(\omega) = B_{t_{j-1}}(\omega) + \theta_j(\omega)(B_{t_j}(\omega) - B_{t_{j-1}}(\omega))$$

for some $\theta_j(\omega) \in [0, 1]$; note that $f''(\xi_j(\omega))$ is measurable.

2° We claim that

$$J_1 = \sum_{j=1}^n f'(B_{t_{j-1}})(B_{t_j} - B_{t_{j-1}}) \xrightarrow[|\Pi| \rightarrow 0]{L^2(\mathbb{P})} \int_0^t f'(B_s) dB_s.$$

By Proposition 14.16 (or Theorem 15.8) we may use a Riemann-sum approximation of the stochastic integral provided that $s \mapsto f'(B_s)$ is continuous in $L^2(\mathbb{P})$ -sense. This, however, follows from $\|f'\|_\infty \leq C_f < \infty$ and the continuity of $x \mapsto f'(x)$ and $s \mapsto B_s$; for all sequences $(s_n)_{n \geq 1}$ with $s_n \rightarrow s$ we have

$$\lim_n \mathbb{E} \left\{ \left(f'(B_{s_n}) - f'(B_s) \right)^2 \right\} \xrightarrow[n \rightarrow \infty]{\text{dom. convergence}} 0.$$

3° For the second-order term we write

$$\begin{aligned} J_2 &:= J_{21} + J_{22} \\ &:= \sum_{j=1}^n f''(B_{t_{j-1}})(B_{t_j} - B_{t_{j-1}})^2 + \sum_{j=1}^n [f''(\xi_j) - f''(B_{t_{j-1}})](B_{t_j} - B_{t_{j-1}})^2. \end{aligned}$$

For the second term we have

$$\begin{aligned} |J_{22}| &\leq \sum_{j=1}^n |f''(\xi_j) - f''(B_{t_{j-1}})| (B_{t_j} - B_{t_{j-1}})^2 \\ &\leq \max_{1 \leq j \leq n} |f''(\xi_j) - f''(B_{t_{j-1}})| \underbrace{\sum_{j=1}^n (B_{t_j} - B_{t_{j-1}})^2}_{= S_2^\Pi(B; t) \text{ cf. (9.1)}}. \end{aligned}$$

Taking expectations and using the Cauchy-Schwarz inequality we get

$$\mathbb{E} |J_{22}| \leq \sqrt{\mathbb{E} \left(\max_{1 \leq j \leq n} |f''(\xi_j) - f''(B_{t_{j-1}})|^2 \right)} \sqrt{\mathbb{E} [(S_2^\Pi(B; t))^2]}.$$

By Theorem 9.1, $L^2(\mathbb{P})\text{-}\lim_{|\Pi| \rightarrow 0} S_2^\Pi(B; t) = t$; since $s \mapsto f''(B_s)$ is uniformly continuous on $[0, t]$ and bounded by $\|f''\|_\infty \leq C_f$, we can use dominated convergence to get

$$\lim_{|\Pi| \rightarrow 0} \mathbb{E} |J_{22}| = 0 \cdot t = 0.$$

The first term, J_{21} , converges to $\int_0^t f''(B_s) ds$. This follows from

$$\begin{aligned} &\mathbb{E} \left[\left(J_{21} - \sum_{j=1}^n f''(B_{t_{j-1}})(t_j - t_{j-1}) \right)^2 \right] \\ &= \mathbb{E} \left[\left(\sum_{j=1}^n f''(B_{t_{j-1}}) [(B_{t_j} - B_{t_{j-1}})^2 - (t_j - t_{j-1})] \right)^2 \right] \\ &= \mathbb{E} \left[\sum_{j=1}^n |f''(B_{t_{j-1}})|^2 [(B_{t_j} - B_{t_{j-1}})^2 - (t_j - t_{j-1})]^2 \right]. \end{aligned}$$

In the last equality only the pure squares survive when we multiply out the outer square. This is due to the fact that $(B_t^2 - t)_{t \geq 0}$ is a martingale: Indeed, if $j < k$, then we have $t_{j-1} < t_j \leq t_{k-1} < t_k$. For brevity we write

$$f''_{j-1} := f''(B_{t_{j-1}}), \quad \Delta_j B := B_{t_j} - B_{t_{j-1}}, \quad \Delta_j t := t_j - t_{j-1}.$$

By the tower property,

$$\begin{aligned} & \mathbb{E} \left(f''_{j-1} \{(\Delta_j B)^2 - \Delta_j t\} f''_{k-1} \{(\Delta_k B)^2 - \Delta_k t\} \right) \\ &= \mathbb{E} \left(\mathbb{E} \left[f''_{j-1} \{(\Delta_j B)^2 - \Delta_j t\} f''_{k-1} \{(\Delta_k B)^2 - \Delta_k t\} \mid \mathcal{F}_{t_{k-1}} \right] \right) \\ &= \mathbb{E} \left(\underbrace{f''_{j-1} \{(\Delta_j B)^2 - \Delta_j t\} f''_{k-1}}_{\mathcal{F}_{t_{k-1}} \text{ measurable}} \cdot \underbrace{\mathbb{E} \left[\{(\Delta_k B)^2 - \Delta_k t\} \mid \mathcal{F}_{t_{k-1}} \right]}_{=0 \text{ martingale, cf. 5.2 c}} \right) = 0, \end{aligned}$$

i. e. the mixed terms break away. From (the proof of) Theorem 9.1 we get

$$\begin{aligned} & \mathbb{E} \left[\left(J_{21} - \sum_{j=1}^n f''(B_{t_{j-1}})(t_j - t_{j-1}) \right)^2 \right] \\ &= \mathbb{E} \left[\sum_{j=1}^n |f''(B_{t_{j-1}})|^2 [(B_{t_j} - B_{t_{j-1}})^2 - (t_j - t_{j-1})]^2 \right] \\ &\leq \|f''\|_\infty^2 \mathbb{E} \left[\sum_{j=1}^n [(B_{t_j} - B_{t_{j-1}})^2 - (t_j - t_{j-1})]^2 \right] \\ &\leq 2C_f^2 |\Pi| \sum_{j=1}^n (t_j - t_{j-1}) = 2C_f^2 |\Pi| t \xrightarrow{|\Pi| \rightarrow 0} 0. \end{aligned}$$

4° So far we have shown (16.1) for functions $f \in \mathcal{C}^2$ such that $C_f < \infty$. For a general $f \in \mathcal{C}^2$ we fix $\ell \geq 0$, pick a smooth cut-off function $\chi_\ell \in \mathcal{C}_c^2$ satisfying $\mathbb{1}_{\mathbb{B}(0, \ell)} \leq \chi_\ell \leq \mathbb{1}_{\mathbb{B}(0, \ell+1)}$ and set $f_\ell(x) := f(x)\chi_\ell(x)$. Clearly, $C_{f_\ell} < \infty$, and therefore

$$f_\ell(B_t) - f_\ell(B_0) = \int_0^t f'_\ell(B_s) dB_s + \frac{1}{2} \int_0^t f''_\ell(B_s) ds.$$

Consider the stopping times

$$\tau(\ell) := \inf\{s > 0 : |B_s| \geq \ell\}, \quad \ell \geq 1,$$

and note that $\frac{d^j}{dx^j} f_\ell(B_{s \wedge \tau(\ell)}) = f^{(j)}(B_{s \wedge \tau(\ell)})$, $j = 0, 1, 2$. Thus, Theorem 15.7 shows for all $t \geq 0$

$$\begin{aligned}
 & f(B_{t \wedge \tau(\ell)}) - f(B_0) \\
 &= \int_0^{t \wedge \tau(\ell)} f'_\ell(B_s) dB_s + \frac{1}{2} \int_0^{t \wedge \tau(\ell)} f''_\ell(B_s) ds \\
 &\stackrel{15.7 \text{ e)}}{=} \int_0^t f'_\ell(B_s) \mathbb{1}_{[0, \tau(\ell))}(s) dB_s + \frac{1}{2} \int_0^t f''_\ell(B_s) \mathbb{1}_{[0, \tau(\ell))}(s) ds \\
 &= \int_0^t f'(B_s) \mathbb{1}_{[0, \tau(\ell))}(s) dB_s + \frac{1}{2} \int_0^t f''(B_s) \mathbb{1}_{[0, \tau(\ell))}(s) ds \\
 &\stackrel{15.7 \text{ e)}}{=} \int_0^{t \wedge \tau(\ell)} f'(B_s) dB_s + \frac{1}{2} \int_0^{t \wedge \tau(\ell)} f''(B_s) ds.
 \end{aligned}$$

Since $\lim_{\ell \rightarrow \infty} \tau(\ell) = \infty$ almost surely – Brownian motion does not explode in finite time – and since the (stochastic) integrals are continuous as functions of their upper boundary, the proof of Theorem 16.1 is complete. \square

16.4 Itô's formula for stochastic differentials

We will now derive Itô's formula for one-dimensional Itô processes, i. e. processes of the form

$$dX_t = \sigma(t) dB_t + b(t) dt$$

where σ, b are locally bounded (in t) and progressively measurable. We have seen in Remark 16.3 that under this assumption $b \in L^1([0, T], dt)$ and $\sigma \in \mathcal{L}^2_{T, \text{loc}}$ a. s. and for all $T > 0$. We can even show that $b \in \mathcal{L}^2_{T, \text{loc}}$. Fix some common localization sequence τ_n . Since $\overline{\mathcal{E}}_T = \mathcal{L}^2_T$ we find, for every n , sequences of simple processes $(b^\Pi)_\Pi, (\sigma^\Pi)_\Pi \subset \mathcal{E}_T$ such that

$$\sigma^\Pi \mathbb{1}_{[0, \tau_n)} \xrightarrow[|\Pi| \rightarrow 0]{L^2(\lambda_T \otimes \mathbb{P})} \sigma \mathbb{1}_{[0, \tau_n)} \quad \text{and} \quad b^\Pi \mathbb{1}_{[0, \tau_n)} \xrightarrow[|\Pi| \rightarrow 0]{L^2(\lambda_T \otimes \mathbb{P})} b \mathbb{1}_{[0, \tau_n)}.$$

Using a diagonal procedure we can achieve that the sequences $(b^\Pi)_\Pi, (\sigma^\Pi)_\Pi$ are independent of n .

16.4 Lemma. *Let $X_t^\Pi = \int_0^t \sigma^\Pi(s) dB_s + \int_0^t b^\Pi(s) ds$. Then $X_t^\Pi \rightarrow X_t$ uniformly in probability as $|\Pi| \rightarrow 0$, i. e. we have*

$$\lim_{|\Pi| \rightarrow 0} \mathbb{P} \left(\sup_{t \leq T} |X_t^\Pi - X_t| > \epsilon \right) = 0 \quad \text{for all } \epsilon > 0, T > 0. \quad (16.5)$$

Proof. The argument is similar to the one used to prove Theorem 15.7 d). Let τ_n be some localizing sequence and fix $n \geq 0$, $T > 0$ and $\epsilon > 0$. Then

$$\begin{aligned}
& \mathbb{P} \left(\sup_{t \leq T} \left| \int_0^t (\sigma^\Pi(s) - \sigma(s)) dB_s \right| > \epsilon \right) \\
& \leq \mathbb{P} \left(\sup_{t \leq T} \left| \int_0^t (\sigma^\Pi(s) - \sigma(s)) dB_s \right| > \epsilon, \tau_n > T \right) + \mathbb{P}(\tau_n \leq T) \\
& \leq \mathbb{P} \left(\sup_{t \leq T} \left| \int_0^{t \wedge \tau_n} (\sigma^\Pi(s) - \sigma(s)) dB_s \right| > \epsilon \right) + \mathbb{P}(\tau_n \leq T) \\
& \stackrel{\text{Chebyshev}}{\leq} \frac{4}{\epsilon^2} \mathbb{E} \left(\left| \int_0^{T \wedge \tau_n} (\sigma^\Pi(s) - \sigma(s)) dB_s \right|^2 \right) + \mathbb{P}(\tau_n \leq T) \\
& \stackrel{(14.20)}{\leq} \frac{4}{\epsilon^2} \mathbb{E} \left(\int_0^{T \wedge \tau_n} |\sigma^\Pi(s) - \sigma(s)|^2 ds \right) + \mathbb{P}(\tau_n \leq T) \\
& \xrightarrow[|\Pi| \rightarrow 0]{n \text{ fixed}} \mathbb{P}(\tau_n \leq T) \xrightarrow{n \rightarrow \infty} 0.
\end{aligned}$$

A similar, but simpler, calculation yields

$$\mathbb{P} \left(\sup_{t \leq T} \left| \int_0^t (b^\Pi(s) - b(s)) ds \right| > \epsilon \right) \xrightarrow{|\Pi| \rightarrow 0} 0.$$

Finally, for any two random variables $X, Y : \Omega \rightarrow [0, \infty)$,

$$\begin{aligned}
\mathbb{P}(\sup(X + Y) > \epsilon) & \leq \mathbb{P}(\sup X + \sup Y > \epsilon) \\
& \leq \mathbb{P}(\sup X > \epsilon/2) + \mathbb{P}(\sup Y > \epsilon/2)
\end{aligned}$$

from which the claim follows. \square

16.5 Theorem (Itô 1942). *Let X_t be a one-dimensional Itô process in the sense of Definition 16.2 such that $dX_t = \sigma(t) dB_t + b(t) dt$ and let $f : \mathbb{R} \rightarrow \mathbb{R}$ be a \mathcal{C}^2 -function. Then we have for all $t \geq 0$ almost surely*

$$\begin{aligned}
& f(X_t) - f(X_0) \\
& = \int_0^t f'(X_s) \sigma(s) dB_s + \int_0^t f'(X_s) b(s) ds + \frac{1}{2} \int_0^t f''(X_s) \sigma^2(s) ds \quad (16.6) \\
& =: \int_0^t f'(X_s) dX_s + \frac{1}{2} \int_0^t f''(X_s) \sigma^2(s) ds.
\end{aligned}$$

We define the stochastic integral for an Itô process $dX_t = \sigma(t) dB_t + b(t) dt$ as

$$\int_0^t f(s) dX_s := \int_0^t f(s) \sigma(s) dB_s + \int_0^t f(s) b(s) ds \quad (16.7)$$

whenever the right-hand side makes sense. It is possible to define the Itô integral for the integrator dX_t directly in the sense of Chapters 14 and 15 and to *derive* the equality (16.7) rather than to use it as a definition.

Proof of Theorem 16.5. **1°** As in the proof of Theorem 16.1 it is enough to show (16.6) locally for $t \leq \tau(\ell)$ where

$$\tau(\ell) = \inf\{s > 0 : |B_s| \geq \ell\}.$$

This means that we can assume, without loss of generality, that

$$\|f\|_\infty + \|f'\|_\infty + \|f''\|_\infty + \|\sigma\|_\infty + \|b\|_\infty \leq C < \infty.$$

2° Lemma 16.4 shows that we can approximate σ and b by elementary processes σ^Π and b^Π . Denote the corresponding Itô process by X^Π . Since $f \in \mathcal{C}_b^2$, the following limits are uniform in probability:

$$X^\Pi \xrightarrow{|\Pi| \rightarrow 0} X, \quad f(X^\Pi) \xrightarrow{|\Pi| \rightarrow 0} f(X).$$

The right-hand side of (16.6) with σ^Π , b^Π and X^Π also converges uniformly in probability by (an obvious modification of) Lemma 16.4.

This means that we have to show (16.6) only for Itô processes where σ and b are from \mathcal{E}_T for some $T > 0$.

3° Assume that $\sigma, b \in \mathcal{E}_T$ and that $\Pi = \{t_0 = 0 < t_1 < \dots < t_n = T\}$ is the (joint refinement of the) underlying partitions. Fix $t \leq T$ and assume, without loss of generality, that $t = t_k \in \Pi$. As in step 1° of the proof of Theorem 16.1 we get by Taylor's theorem

$$\begin{aligned} f(X_t) - f(X_0) &= \sum_{j=1}^k f'(X_{t_{j-1}})(X_{t_j} - X_{t_{j-1}}) + \frac{1}{2} \sum_{j=1}^k f''(\xi_j)(X_{t_j} - X_{t_{j-1}})^2 \\ &=: J_1 + \frac{1}{2} J_2 \end{aligned}$$

with

$$\xi_j(\omega) = X_{t_{j-1}}(\omega) + \theta_j(\omega)(X_{t_j}(\omega) - X_{t_{j-1}}(\omega))$$

for some $\theta_j(\omega) \in [0, 1]$.

Since σ and b are elementary processes, there are $\mathcal{F}_{t_{j-1}}$ measurable random variables σ_{j-1} and b_{j-1} such that

$$\begin{aligned} J_1 &= \sum_{j=1}^k f'(X_{t_{j-1}}) \sigma_{j-1} \cdot (B_{t_j} - B_{t_{j-1}}) + \sum_{j=1}^k f'(X_{t_{j-1}}) b_{j-1} \cdot (t_j - t_{j-1}) \\ &= J_{11} + J_{12}. \end{aligned}$$

Now $L^2(\mathbb{P})\text{-}\lim_{|\Pi| \rightarrow 0} J_{11} = \int_0^t f'(X_s) \sigma(s) dB_s$ which follows as in the proof of Theorem 16.1, step 2°. The second term J_{12} is just a Riemann sum which converges a. s.

to $\int_0^t f'(X_s)b(s) ds$. Moreover,

$$\begin{aligned}
 J_2 &= \sum_{j=1}^k f''(\xi_j)(\sigma_{j-1} \cdot (B_{t_j} - B_{t_{j-1}}) + b_{j-1} \cdot (t_j - t_{j-1}))^2 \\
 &= \sum_{j=1}^k \left[f''(\xi_j)\sigma_{j-1}^2 \cdot (B_{t_j} - B_{t_{j-1}})^2 \right. \\
 &\quad \left. + 2f''(\xi_j)\sigma_{j-1}b_{j-1} \cdot (B_{t_j} - B_{t_{j-1}})(t_j - t_{j-1}) \right. \\
 &\quad \left. + f''(\xi_j)b_{j-1}^2 \cdot (t_j - t_{j-1})^2 \right] \\
 &= J_{21} + J_{22} + J_{23}
 \end{aligned}$$

As in the Proof of 16.1, step 3°, we get $L^2(\mathbb{P})\text{-}\lim_{|\Pi| \rightarrow 0} J_{21} = \int_0^t f''(X_s)\sigma(s) ds$. Since

$$\begin{aligned}
 |J_{22}| &\leq 2 \sum_{j=1}^k \|f''\|_\infty \|\sigma\|_\infty \|b\|_\infty (t_j - t_{j-1}) \max_k |B_{t_k} - B_{t_{k-1}}| \\
 &= 2 \|f''\|_\infty \|\sigma\|_\infty \|b\|_\infty T \max_k |B_{t_k} - B_{t_{k-1}}|
 \end{aligned}$$

and since Brownian motion is almost uniformly surely continuous on $[0, T]$, we get $\lim_{|\Pi| \rightarrow 0} |J_{22}| = 0$ a. s.; a similar calculation gives $\lim_{|\Pi| \rightarrow 0} |J_{23}| = 0$. \square

16.5 Itô's formula for Brownian motion in \mathbb{R}^d

Assume that $B = (B^1, \dots, B^d)$ is a BM^d . Recall that the stochastic integral with respect to dB is defined for each coordinate, cf. Section 16.1. With this convention, the multidimensional version of Itô's formula becomes

16.6 Theorem (Itô 1942). *Let $(B_t)_{t \geq 0}$, $B_t = (B_t^1, \dots, B_t^d)$, be a d -dimensional Brownian motion, $X_t = \int_0^t \sigma(s) dB_s + \int_0^t b(s) ds$ be an m -dimensional Itô process and $f : \mathbb{R}^m \rightarrow \mathbb{R}$ be a \mathcal{C}^2 -function. Then*

$$\begin{aligned}
 f(X_t) - f(X_0) &= \sum_{k=1}^d \int_0^t \left[\sum_{j=1}^m \partial_j f(X_s) \sigma_{jk}(s) \right] dB_s^k + \sum_{k=1}^m \int_0^t \partial_k f(X_s) b_k(s) ds \\
 &\quad + \frac{1}{2} \sum_{i,j=1}^m \int_0^t \partial_i \partial_j f(X_s) \sum_{k=1}^d \sigma_{ik}(s) \sigma_{jk}(s) ds.
 \end{aligned} \tag{16.8}$$

The proof is pretty much the same as the proof in dimension $d = 1$ – just use the Taylor expansion of functions in \mathbb{R}^d rather than in \mathbb{R} .

16.7 Remark. In the literature often the following matrix and vector version of (16.8) is used:

$$\begin{aligned} f(X_t) - f(X_0) &= \int_0^t \nabla f(X_s)^\top \sigma(s) dB_s + \int_0^t \nabla f(X_s)^\top b(s) ds \\ &\quad + \frac{1}{2} \int_0^t \text{trace}(\sigma(s)^\top D^2 f(X_s) \sigma(s)) ds. \end{aligned} \quad (16.9)$$

where $\nabla = (\partial_1, \dots, \partial_d)^\top$ is the gradient, $D^2 f = (\partial_j \partial_k f)_{j,k=1}^d$ is the Hessian and dB_s is understood as a column vector from \mathbb{R}^d . In differential notation this becomes

$$\begin{aligned} df(X_t) &= \nabla f(X_t)^\top \sigma(t) dB_t + \nabla f(X_t)^\top b(t) dt \\ &\quad + \frac{1}{2} \text{trace}(\sigma(t)^\top D^2 f(X_t) \sigma(t)) dt. \end{aligned} \quad (16.10)$$

16.6 Tanaka's formula and local time

Itô's formula has been extended in various directions. Here we discuss the prototype of generalizations which relax the smoothness assumptions for the function f . The key point is that we retain some control on the (generalized) second derivative of f , e. g. if f is convex or if $f(x) = \int_0^x \phi(y) dy$ where $\phi \in \mathcal{B}_b(\mathbb{R})$ as in the Bouleau–Yor formula [151, Theorem IV.77]. In this section we consider one-dimensional Brownian motion $(B_t)_{t \geq 0}$ and the function $f(x) = |x|$.

Differentiating f (in the sense of generalized functions) yields $f'(x) = \text{sgn}(x)$ and $f''(x) = \delta_0(x)$. A formal application of Itô's formula (16.1) gives

$$|B_t| = \int_0^t \text{sgn}(B_s) dB_s + \frac{1}{2} \int_0^t \delta_0(B_s) ds. \quad (16.11)$$

In order to justify (16.11), we have to smooth out the function $f(x) = |x|$. Note that for $\epsilon > 0$

$$\begin{aligned} f_\epsilon(x) &:= \begin{cases} |x| - \frac{1}{2}\epsilon, & |x| > \epsilon, \\ \frac{1}{2\epsilon}x^2, & |x| \leq \epsilon, \end{cases} & f'_\epsilon(x) &= \begin{cases} \text{sgn}(x), & |x| > \epsilon, \\ \frac{1}{\epsilon}x, & |x| \leq \epsilon, \end{cases} \\ & & f''_\epsilon(x) &= \begin{cases} 0, & |x| > \epsilon, \\ \frac{1}{\epsilon}, & |x| < \epsilon. \end{cases} \end{aligned}$$

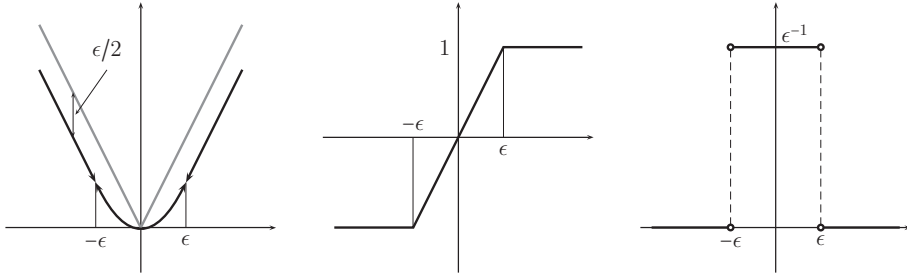


Figure 16.1. Smooth approximation of the function $x \mapsto |x|$.

The approximations f_ϵ are not yet \mathcal{C}^2 functions. Therefore, we use a Friedrichs mollifier: Pick any $\chi \in \mathcal{C}_c^\infty(-1, 1)$ with $0 \leq \chi \leq 1$, $\chi(0) = 1$ and $\int \chi(x) dx = 1$, and set for $n \geq 1$

$$\chi_n(x) := n\chi(nx) \quad \text{and} \quad f_{\epsilon,n}(x) := f_\epsilon \star \chi_n(x) = \int \chi_n(x-y)f_\epsilon(y) dy.$$

Ex. 16.9 It is not hard to see that $\lim_{n \rightarrow \infty} f_{\epsilon,n} = f_\epsilon$ and $\lim_{n \rightarrow \infty} f'_{\epsilon,n} = f'_\epsilon$ uniformly while $\lim_{n \rightarrow \infty} f''_{\epsilon,n}(x) = f''_\epsilon(x)$ for all $x \neq \pm\epsilon$.

If we apply Itô's formula (16.1) we get

$$f_{\epsilon,n}(B_t) = \int_0^t f'_{\epsilon,n}(B_s) dB_s + \frac{1}{2} \int_0^t f''_{\epsilon,n}(B_s) ds \quad \text{a. s.}$$

On the left-hand side we see $\lim_{\epsilon \rightarrow 0} \lim_{n \rightarrow \infty} f_{\epsilon,n}(B_t) = |B_t|$ a. s. and in probability.

For the expression on the right we have

$$\mathbb{E}^0 \left(\left| \int_0^t (f'_{\epsilon,n}(B_s) - \text{sgn}(B_s)) dB_s \right|^2 \right) = \mathbb{E}^0 \left(\int_0^t |f'_{\epsilon,n}(B_s) - \text{sgn}(B_s)|^2 ds \right).$$

Letting first $n \rightarrow \infty$ and then $\epsilon \rightarrow 0$, the right-hand side converges to

$$\mathbb{E}^0 \left(\int_0^t |f'_\epsilon(B_s) - \text{sgn}(B_s)|^2 ds \right) \xrightarrow[\epsilon \rightarrow 0]{\text{dom. conv.}} 0;$$

in particular, $\mathbb{P}\text{-}\lim_{n \rightarrow \infty} \int_0^t f'_\epsilon(B_s) dB_s = \int_0^t \text{sgn}(B_s) dB_s$.

Finally, set

$$\Omega_s := \{\omega : \lim_{n \rightarrow \infty} f''_{\epsilon,n}(B_s(\omega)) = f''_\epsilon(B_s(\omega))\}.$$

As $\mathbb{P}(B_s = \pm\epsilon) = 0$, we know that $\mathbb{P}(\Omega_s) = 1$ for each $s \in [0, t]$. By Fubini's theorem

$$\mathbb{P} \otimes \text{Leb} \{(\omega, s) : s \in [0, t], \omega \notin \Omega_s\} = \int_0^t \mathbb{P}(\Omega \setminus \Omega_s) ds = 0,$$

and so $\text{Leb}\{s \in [0, t] : \omega \in \Omega_s\} = t$ for \mathbb{P} almost all ω . Since $\|f''_{\epsilon,n}\|_\infty \leq (2\epsilon)^{-1}$, the dominated convergence theorem yields

$$L^2\text{-}\lim_{n \rightarrow \infty} \int_0^t f''_{\epsilon,n}(B_s) ds = \int_0^t f''_\epsilon(B_s) ds = \frac{1}{2\epsilon} \int_0^t \mathbb{1}_{(-\epsilon, \epsilon)}(B_s) ds.$$

Therefore, we have shown that

$$\mathbb{P}\text{-}\lim_{\epsilon \rightarrow 0} \frac{1}{2\epsilon} \int_0^t \mathbb{1}_{(-\epsilon, \epsilon)}(B_s) ds = |B_t| - \int_0^t \text{sgn}(B_s) dB_s$$

exists and defines a stochastic process. The same argument applies for the shifted function $f(x) = |x - a|$, $a \in \mathbb{R}$. Therefore, the following definition makes sense.

16.8 Definition (Brownian local time. Lévy 1939). Let $(B_t)_{t \geq 0}$ be a d -dimensional Brownian motion and $a \in \mathbb{R}$. The *local time at the level a up to time t* is the random process

$$L_t^a(\omega) := \mathbb{P}\text{-}\lim_{\epsilon \rightarrow 0} \frac{1}{2\epsilon} \int_0^t \mathbb{1}_{(-\epsilon, \epsilon)}(B_s - a) ds. \quad (16.12)$$

The local time L_t^a represents the total amount of time Brownian motion spends at the level a up to time t :

$$\frac{1}{2\epsilon} \int_0^t \mathbb{1}_{(-\epsilon, \epsilon)}(B_s - a) ds = \frac{1}{2\epsilon} \text{Leb}\{s \in [0, t] : B_s \in (a - \epsilon, a + \epsilon)\}.$$

We are now ready for the next result.

16.9 Theorem (Tanaka's formula. Tanaka 1963). Let $(B_t)_{t \geq 0}$ denote a one-dimensional Brownian motion. Then we have for every $a \in \mathbb{R}$

$$|B_t - a| = |a| + \int_0^t \text{sgn}(B_s - a) dB_s + L_t^a \quad (16.13)$$

where L_t^a is Brownian local time at the level a . In particular, $(t, a) \mapsto L_t^a$ (has a version which) is continuous.

Proof. The formula (16.13) follows from the preceding discussion for the shifted function $f(x) = |x - a|$. The continuity of $(t, a) \mapsto L_t^a$ is now obvious since both Brownian motion and the stochastic integral have continuous versions. \square

Tanaka's formula is the starting point for many further investigations. For example, it is possible to show that $\beta_t = \int_0^t \text{sgn}(B_s) dB_s$ is again a one-dimensional Brownian motion and $L_t^0 = \sup_{s \leq t} (-\beta_s)$. Moreover, $\mathbb{P}^0(L_\infty^a = \infty) = 1$ for any $a \in \mathbb{R}$.

Ex. 16.10



16.10 Further reading. All books mentioned in the *Further reading* section of Chapter 14 contain proofs of Itô's formula. Local times are treated in much greater detail in [99] and [156]. An interesting alternative proof of Itô's formula is given in [109].

[99] Karatzas, Shreve: *Brownian Motion and Stochastic Calculus*.

[109] Krylov: *Introduction to the Theory of Diffusion Processes*.

[156] Revuz, Yor: *Continuous Martingales and Brownian Motion*.

Problems

1. Let $(B_t)_{t \geq 0}$ be a BM¹. Use Itô's formula to obtain representations of

$$X_t = \int_0^t \exp(B_s) dB_s \quad \text{and} \quad Y_t = \int_0^t B_s \exp(B_s^2) dB_s$$

which do not contain Itô integrals.

2. (a) Use the (two-dimensional, deterministic) chain rule $d(F \circ G) = F' \circ G dG$ to deduce the formula for integration by parts for Stieltjes integrals:

$$\int_0^t f(s) dg(s) = f(t)g(t) - f(0)g(0) - \int_0^t g(s) df(s)$$

for all $f, g \in \mathcal{C}^1([0, \infty), \mathbb{R})$.

- (b) Use the Itô formula for a Brownian motion $(b_t, \beta_t)_{t \geq 0}$ in \mathbb{R}^2 to show that

$$\int_0^t b_s d\beta_s = b_t \beta_t - \int_0^t \beta_s db_s, \quad t \geq 0.$$

What happens if b and β are not independent?

3. Prove the following time-dependent version of Itô's formula: Let $(B_t)_{t \geq 0}$ be a BM¹ and $f : [0, \infty) \times \mathbb{R} \rightarrow \mathbb{R}$ be a function of class $\mathcal{C}^{1,2}$. Then

$$f(t, B_t) - f(0, 0) = \int_0^t f(s, B_s) dB_s + \int_0^t \left(\frac{\partial f}{\partial t}(s, B_s) + \frac{1}{2} \frac{\partial^2 f}{\partial x^2}(s, B_s) \right) ds.$$

Prove and state the d -dimensional counterpart.

Hint: Use the Itô formula for the $d + 1$ -dimensional Itô process (t, B_t^1, \dots, B_t^d) .

4. Prove Theorem 5.6 using Itô's formula.
5. Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM¹. Use Itô's formula to verify that the following processes are martingales:

$$X_t = e^{t/2} \cos B_t \quad \text{and} \quad Y_t = (B_t + t)e^{-B_t - t/2}.$$

6. Let $B_t = (b_t, \beta_t)$, $t \geq 0$ be a BM^2 and set $r_t := |B_t| = \sqrt{b_t^2 + \beta_t^2}$.
 - (a) Show that the stochastic integrals $\int_0^t b_s/r_s db_s$ and $\int_0^t \beta_s/r_s d\beta_s$ exist.
 - (b) Show that $W_t := \int_0^t b_s/r_s db_s + \int_0^t \beta_s/r_s d\beta_s$ is a BM^1 .
7. Let $B_t = (b_t, \beta_t)$, $t \geq 0$ be a BM^2 and $f(x+iy) = u(x, y) + iv(x, y)$, $x, y \in \mathbb{R}$, an analytic function. If $u_x^2 + v_y^2 = 1$, then $(u(b_t, \beta_t), v(b_t, \beta_t))$, $t \geq 0$, is a BM^2 .
8. Show that the d -dimensional Itô formula remains valid if we replace the real-valued function $f : \mathbb{R}^d \rightarrow \mathbb{R}$ by a complex function $f = u + iv : \mathbb{R}^d \rightarrow \mathbb{C}$.
9. (Friedrichs mollifier) Let $\chi \in \mathcal{C}_c^\infty(-1, 1)$ such that $0 \leq \chi \leq 1$, $\chi(0) = 1$ and $\int \chi(x) dx = 1$. Set $\chi_n(x) := n\chi(nx)$.
 - (a) Show that $\text{supp } \chi_n \subset [-1/n, 1/n]$ and $\int \chi_n(x) dx = 1$.
 - (b) Let $f \in \mathcal{C}(\mathbb{R})$ and $f_n := f \star \chi_n$.
Show that $|\partial^k f_n(x)| \leq n^k \sup_{y \in \mathbb{B}(x, 1/n)} |f(y)| \|\partial^k \chi\|_{L^1}$.
 - (c) Let $f \in \mathcal{C}(\mathbb{R})$ uniformly continuous. Show that $\lim_{n \rightarrow \infty} \|f \star \chi_n - f\|_\infty = 0$.
What can be said if f is continuous but not uniformly continuous?
 - (d) Let f be piecewise continuous. Show that $\lim_{n \rightarrow \infty} f \star \chi_n(x) = f(x)$ at all x where f is continuous.
10. Show that $\beta_t = \int_0^t \text{sgn}(B_s) dB_s$ is a BM^1 .
Hint: Use Lévy's characterization of a BM^1 , Theorem 9.12 or 17.5.

Chapter 17

Applications of Itô's formula

Itô's formula has many applications and we restrict ourselves to a few of them. We will use it to obtain a characterization of a Brownian motion as a martingale (Lévy's theorem 17.5) and to describe the structure of 'Brownian' martingales (Theorems 17.10, 17.12 and 17.15). In some sense, this will show that a Brownian motion is *both* a very particular martingale *and* the typical martingale. Girsanov's theorem 17.8 allows us to change the underlying probability measure which will become important if we want to solve stochastic differential equations. Finally, the Burkholder–Davis–Gundy inequalities, Theorem 17.16, provide moment estimates for stochastic integrals with respect to a Brownian motion.

Throughout this section $(B_t, \mathcal{F}_t)_{t \geq 0}$ is a BM with an admissible complete filtration, e. g. $\overline{\mathcal{F}}_t^B$, see Theorem 6.21. Recall that $\mathcal{L}_T^2 = L_{\mathcal{P}}^2(\lambda_T \otimes \mathbb{P})$.

17.1 Doléans–Dade exponentials

Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a Brownian motion on \mathbb{R} . We have seen in Example 5.2 e) that $(M_t^\xi, \mathcal{F}_t)_{t \geq 0}$ is a martingale where $M_t^\xi := e^{\xi B_t - \frac{1}{2} \xi^2 t}$, $\xi \in \mathbb{R}$. Using Itô's formula we find that M_t^ξ satisfies the following integral equation

$$M_t^\xi = 1 + \int_0^t \xi M_s^\xi dB_s \quad \text{or} \quad dM_t^\xi = \xi M_t^\xi dB_t.$$

In the usual calculus the differential equation $dy(t) = y(t) dx(t)$, $y(0) = y_0$ has the unique solution $y(t) = y_0 e^{x(t) - x(0)}$. If we compare this with the Brownian setting, we see that there appears the additional factor $-\frac{1}{2} \xi^2 t = -\frac{1}{2} \langle \xi B \rangle_t$ which is due to the second order term in Itô's formula. On the other hand, it is this factor which makes M^ξ into a martingale.

Because of this analogy it is customary to call

$$\mathcal{E}(M)_t := \exp \left(M_t - \frac{1}{2} \langle M \rangle_t \right), \quad t \geq 0, \quad (17.1)$$

for any continuous L^2 martingale $(M_t)_{t \leq T}$ the *stochastic* or *Doléans–Dade exponential*. If we use Itô's formula for the stochastic integral for martingales from Section 14.6, it is not hard to see that $\mathcal{E}(M)_t$ is itself a martingale. Here we consider only some special cases which can be directly written in terms of a Brownian motion.

17.1 Lemma. *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM¹, $f \in L^2_{\mathcal{F}}(\lambda_T \otimes \mathbb{P})$ for all $T > 0$, and assume that $|f(s, \omega)| \leq C$ for some $C > 0$ and all $s \geq 0$ and $\omega \in \Omega$. Then* Ex. 17.1

$$\exp \left(\int_0^t f(s) dB_s - \frac{1}{2} \int_0^t f^2(s) ds \right), \quad t \geq 0, \quad (17.2)$$

is a martingale for the filtration $(\mathcal{F}_t)_{t \geq 0}$.

Proof. Set $X_t = \int_0^t f(s) dB_s - \frac{1}{2} \int_0^t f^2(s) ds$. Itô's formula, Theorem 16.5, yields¹

$$\begin{aligned} e^{X_t} - 1 &= \int_0^t e^{X_s} f(s) dB_s - \frac{1}{2} \int_0^t e^{X_s} f^2(s) ds + \frac{1}{2} \int_0^t e^{X_s} f^2(s) ds \\ &= \int_0^t \exp \left(\int_0^s f(r) dB_r - \frac{1}{2} \int_0^s f^2(r) dr \right) f(s) dB_s. \end{aligned} \quad (17.3)$$

If we can show that the integrand is in $L^2_{\mathcal{F}}(\lambda_T \otimes \mathbb{P})$ for every $T > 0$, then Theorem 14.13 applies and shows that the stochastic integral, hence e^{X_t} , is a martingale.

We begin with an estimate for simple processes. Fix $T > 0$ and assume that $g \in \mathcal{E}_T$. Then $g(s, \omega) = \sum_{j=1}^n g(s_{j-1}, \omega) \mathbb{1}_{[s_{j-1}, s_j)}(s)$ for some partition $0 = s_0 < \dots < s_n = T$ of $[0, T]$ and $|g(s, \omega)| \leq C$ for some constant $C < \infty$. Since $g(s_{j-1})$ is $\mathcal{F}_{s_{j-1}}$ measurable and $B_{s_j} - B_{s_{j-1}} \perp \mathcal{F}_{s_{j-1}}$, we find

$$\begin{aligned} &\mathbb{E} \left[e^{2 \int_0^T g(r) dB_r} \right] \\ &= \mathbb{E} \left[\prod_{j=1}^n e^{2g(s_{j-1})(B_{s_j} - B_{s_{j-1}})} \right] \\ &\stackrel{\text{tower}}{=} \mathbb{E} \left[\prod_{j=1}^{n-1} e^{2g(s_{j-1})(B_{s_j} - B_{s_{j-1}})} \underbrace{\mathbb{E} \left(e^{2g(s_{n-1})(B_{s_n} - B_{s_{n-1}})} \mid \mathcal{F}_{s_{n-1}} \right)}_{\substack{(\text{A.4}) \\ 2.2} \exp \left[2g^2(s_{n-1})(s_n - s_{n-1}) \right]} \right] \\ &\leq e^{2C^2(s_n - s_{n-1})} \mathbb{E} \left[\prod_{j=1}^{n-1} e^{2g(s_{j-1})(B_{s_j} - B_{s_{j-1}})} \right]. \end{aligned}$$

Further applications of the tower property yield

$$\mathbb{E} \left[e^{2 \int_0^T g(r) dB_r} \right] \leq \prod_{j=1}^n e^{2C^2(s_j - s_{j-1})} = e^{2C^2 T}.$$

¹ Compare the following equality (17.3) with (17.1).

If $f \in L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$, there is a sequence of simple processes $(f_n)_{n \geq 1} \subset \mathcal{E}_T$ such that $\lim_{n \rightarrow \infty} f_n = f$ in $L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$ and $\lim_{n \rightarrow \infty} \int_0^T f_n(s) dB_s = \int_0^T f(s) dB_s$ in $L^2(\mathbb{P})$ and, for a subsequence, almost surely. We can assume that $|f_n(s, \omega)| \leq C$, otherwise we would consider $-C \vee f_n \wedge C$. By Fatou's lemma and the estimate for simple processes from above we get

$$\begin{aligned} \mathbb{E} \left[\left| e^{\int_0^T f(r) dB_r - \frac{1}{2} \int_0^T f^2(r) dr} f(T) \right|^2 \right] &\leq C^2 \mathbb{E} \left[e^{2 \int_0^T f(r) dB_r} \right] \\ &= C^2 \mathbb{E} \left[\lim_{n \rightarrow \infty} e^{2 \int_0^T f_n(r) dB_r} \right] \\ &\leq C^2 \varliminf_{n \rightarrow \infty} \mathbb{E} \left[e^{2 \int_0^T f_n(r) dB_r} \right] \\ &\leq C^2 e^{2C^2 T} < \infty. \end{aligned} \quad \square$$

It is clear that Lemma 17.1 also holds for BM^d and d -dimensional integrands. More interesting is the observation that we can replace the condition that the integrand is bounded by an integrability condition.

17.2 Theorem. *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$, $B_t = (B_t^1, \dots, B_t^d)$, be a d -dimensional Brownian motion and $f = (f_1, \dots, f_d)$ be a d -dimensional \mathcal{P} measurable process such that $f_j \in L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$ for all $T > 0$ and $j = 1, \dots, d$. Then*

$$M_t = \exp \left(\sum_{j=1}^d \int_0^t f_j(s) dB_s^j - \frac{1}{2} \int_0^t |f(s)|^2 ds \right) \quad (17.4)$$

is a martingale for the filtration $(\mathcal{F}_t)_{t \geq 0}$ if, and only if, $\mathbb{E} M_t = 1$.

Proof. The necessity is obvious since $M_0 = 1$. Since $f \in L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$ for all $T > 0$, we see with the d -dimensional version of Itô's formula (16.8) applied to the process $X_t = \sum_{j=1}^d \int_0^t f_j(s) dB_s^j - \frac{1}{2} \int_0^t |f(s)|^2 ds$ that

$$\begin{aligned} e^{X_t} &= 1 + \sum_{j=1}^d \int_0^t f_j(s) e^{X_s} dB_s^j \\ &= 1 + \sum_{j=1}^d \int_0^t f_j(s) \exp \left(\sum_{j=1}^d \int_0^s f_j(r) dB_r^j - \frac{1}{2} \int_0^s |f(r)|^2 dr \right) dB_s^j. \end{aligned}$$

Let $\tau_n = \inf\{s \geq 0 : |\exp(X_s)| \geq n\} \wedge n$; since $s \mapsto \exp(X_s)$ is continuous, τ_n is a stopping time and $\exp(X_{\tau_n}) \leq n$. Therefore, $f^{\tau_n} \cdot \exp(X_{\tau_n})$ is in $L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$ for all $T > 0$, and Theorem 14.13 e) shows that $M^{\tau_n} = \exp(X_{\tau_n})$ is a martingale. Clearly, $\lim_{n \rightarrow \infty} \tau_n = \infty$ which means that M_t is a local martingale, cf. Definition 15.6.

Proposition 17.3 below shows that any positive local martingale with $\mathbb{E} M_t = 1$ is already a martingale, and the claim follows. \square

The last step in the proof of Theorem 17.2 is interesting on its own:

17.3 Proposition.

- a) Every positive local martingale $(M_t, \mathcal{F}_t)_{t \geq 0}$ is a (positive) supermartingale.
 b) Let $(M_t, \mathcal{F}_t)_{t \geq 0}$ be a supermartingale with $\mathbb{E} M_t = \mathbb{E} M_0$ for all $t \geq 0$, then $(M_t)_{t \geq 0}$ is already a martingale.

Proof. a) Let τ_n be a localizing sequence for $(M_t)_{t \geq 0}$. By Fatou's lemma we see that

$$\mathbb{E} M_t = \mathbb{E} \left(\lim_{n \rightarrow \infty} M_{t \wedge \tau_n} \right) \leq \lim_{n \rightarrow \infty} \mathbb{E} M_{t \wedge \tau_n} = \mathbb{E} M_0 < \infty,$$

since M^{τ_n} is a martingale. The conditional version of Fatou's lemma shows for all $s \leq t$

$$\mathbb{E} (M_t \mid \mathcal{F}_s) \leq \lim_{n \rightarrow \infty} \mathbb{E} (M_{t \wedge \tau_n} \mid \mathcal{F}_s) = \lim_{n \rightarrow \infty} M_{s \wedge \tau_n} = M_s,$$

i. e. $(M_t)_{t \geq 0}$ is a supermartingale.

b) Let $s \leq t$ and $F \in \mathcal{F}_s$. By assumption, $\int_F M_s d\mathbb{P} \geq \int_F M_t d\mathbb{P}$. Subtracting from both sides $\mathbb{E} M_s = \mathbb{E} M_t$ gives

$$\int_F M_s d\mathbb{P} - \mathbb{E} M_s \geq \int_F M_t d\mathbb{P} - \mathbb{E} M_t.$$

Therefore,

$$\int_{F^c} M_s d\mathbb{P} \leq \int_{F^c} M_t d\mathbb{P} \quad \text{for all } s \leq t, F^c \in \mathcal{F}_s.$$

This means that $(M_t)_{t \geq 0}$ is a submartingale, hence a martingale. \square

We close this section with the Novikov condition which gives a sufficient criterion for $\mathbb{E} M_t = 1$. The following proof is essentially the proof which we used for the exponential Wald identity, cf. Theorem 5.14. We prove only the one-dimensional version, the extension to higher dimensions is obvious.

17.4 Theorem (Novikov 1972). *Let $(B_t)_{t \geq 0}$ be a BM¹ and let $f \in L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$ for all $T > 0$. If*

$$\mathbb{E} \left[\exp \left(\frac{1}{2} \int_0^\infty |f(s)|^2 ds \right) \right] < \infty, \quad (17.5)$$

then $\mathbb{E} M_t = 1$ for the stochastic exponential

$$M_t = \exp \left(\int_0^t f(s) dB_s - \frac{1}{2} \int_0^t |f(s)|^2 ds \right).$$

Proof. We have seen in the proof of Theorem 17.2 that M_t is a positive local martingale, and by Proposition 17.3 a) it is a supermartingale. In particular, $\mathbb{E} M_t \leq \mathbb{E} M_0 = 1$.

For the converse inequality write $X_t = \int_0^t f(s) dB_s$ and $\langle X \rangle_t = \int_0^t |f(s)|^2 ds$. Then $M_t = e^{X_t - \frac{1}{2} \langle X \rangle_t} = \mathcal{E}(X)_t$ and the condition (17.5) reads $\mathbb{E}[e^{\frac{1}{2} \langle X \rangle_\infty}] < \infty$.

Let $(\tau_n)_{n \geq 1}$ be a localizing sequence for the local martingale M , fix some $c \in (0, 1)$ and pick $p = p(c) > 1$ such that $p < \frac{1}{c} \wedge \frac{1}{c(2-c)}$.

Then $M_t^{\tau_n} = \exp[X_t^{\tau_n} - \frac{1}{2} \langle X \rangle_t^{\tau_n}]$ is a martingale and $\mathbb{E} M_{t \wedge \tau_n} = \mathbb{E} M_0 = 1$ for each $n \geq 1$. Using the Hölder inequality for the conjugate exponents $1/pc$ and $1/(1 - pc)$ we find

$$\begin{aligned} \mathbb{E} [\mathcal{E}(cX^{\tau_n})_t^p] &= \mathbb{E} [e^{pcX_t^{\tau_n} - \frac{1}{2} pc^2 \langle X \rangle_t^{\tau_n}}] \\ &= \mathbb{E} [e^{pc(X_t^{\tau_n} - \frac{1}{2} \langle X \rangle_t^{\tau_n})} e^{\frac{1}{2} pc(1-c) \langle X \rangle_t^{\tau_n}}] \\ &\leq \underbrace{\left[\mathbb{E} e^{X_t^{\tau_n} - \frac{1}{2} \langle X \rangle_t^{\tau_n}} \right]^{pc}}_{= [\mathbb{E} M_{t \wedge \tau_n}]^{pc} = 1} \left[\mathbb{E} e^{\frac{1}{2} \frac{pc(1-c)}{1-pc} \langle X \rangle_t^{\tau_n}} \right]^{1-pc} \\ &\leq \left(\mathbb{E} e^{\frac{1}{2} \langle X \rangle_\infty} \right)^{1-pc}. \end{aligned}$$

In the last step we used that $pc(1-c)/(1-pc) \leq 1$. This shows that the p th moment of the family $(M_t^{\tau_n})_{n \geq 1}$ is uniformly bounded, hence it is uniformly integrable.

Therefore, we can let $n \rightarrow \infty$ and find, using uniform integrability and Hölder's inequality for the exponents $1/c$ and $1/(1-c)$

$$\begin{aligned} 1 &= \mathbb{E} [e^{cX_t^{\tau_n} - \frac{1}{2} c^2 \langle X \rangle_t^{\tau_n}}] = \mathbb{E} [e^{cX_t - \frac{1}{2} c \langle X \rangle_t} e^{\frac{1}{2} c(1-c) \langle X \rangle_t}] \\ &\leq \left[\mathbb{E} e^{X_t - \frac{1}{2} \langle X \rangle_t} \right]^c \left[\mathbb{E} e^{\frac{1}{2} c \langle X \rangle_t} \right]^{1-c}. \end{aligned}$$

Since the last factor is bounded by $(\mathbb{E} e^{\langle X \rangle_\infty})^{1-c}$, we find as $c \rightarrow 1$,

$$1 \leq \mathbb{E} e^{X_t - \frac{1}{2} \langle X \rangle_t} = \mathbb{E} M_t. \quad \square$$

17.2 Lévy's characterization of Brownian motion

In the hands of Kunita and Watanabe [113] Itô's formula became a most powerful tool. We will follow their ideas and give an elegant proof of the following theorem due to P. Lévy.

17.5 Theorem (Lévy 1948). *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$, $X_0 = 0$, be a real-valued, adapted process with continuous sample paths. If both $(X_t, \mathcal{F}_t)_{t \geq 0}$ and $(X_t^2 - t, \mathcal{F}_t)_{t \geq 0}$ are martingales, then $(X_t)_{t \geq 0}$ is a Brownian motion with filtration $(\mathcal{F}_t)_{t \geq 0}$.* Ex. 17.2

Theorem 17.5 means, in particular, that BM^1 is the only continuous martingale with quadratic variation $\langle B \rangle_t \equiv t$.

The way we have stated Itô's formula, we cannot use it for $(X_t)_{t \geq 0}$ as in Theorem 17.5. Nevertheless, a close inspection of the construction of the stochastic integral in Chapters 14 and 15 shows that all results remain valid if the integrator dB_t is replaced by dX_t where $(X_t)_{t \geq 0}$ is a continuous martingale with quadratic variation $\langle X \rangle_t = t$. For step 3° of the proof of Theorem 16.1, cf. pages 237 and 238, we have to replace Theorem 9.1 by Corollary 9.11. In fact, Theorem 17.5 is an ex-post vindication for this since such martingales *already are Brownian motions*. In order not to end up with a circular conclusion, we have to go through the proofs of Chapters 14–16 again, though.



After this word of caution we can use (16.1) in the setting of Theorem 17.5 and start with the

Proof of Theorem 17.5 (Kunita, Watanabe 1967). We have to show that

$$X_t - X_s \perp \mathcal{F}_s \quad \text{and} \quad X_t - X_s \sim \mathcal{N}(0, t - s) \quad \text{for all } 0 \leq s \leq t.$$

This is equivalent to

$$\mathbb{E} \left(e^{i\xi(X_t - X_s)} \mathbb{1}_F \right) = e^{-\frac{1}{2}(t-s)\xi^2} \mathbb{P}(F) \quad \text{for all } s \leq t, F \in \mathcal{F}_s. \quad (17.6)$$

In fact, setting $F = \Omega$ yields

$$\mathbb{E} \left(e^{i\xi(X_t - X_s)} \right) = e^{-\frac{1}{2}(t-s)\xi^2},$$

i. e. $X_t - X_s \sim \mathcal{N}(0, t - s)$, and (17.6) becomes

$$\mathbb{E} \left(e^{i\xi(X_t - X_s)} \mathbb{1}_F \right) = \mathbb{E} \left(e^{i\xi(X_t - X_s)} \right) \mathbb{P}(F) \quad \text{for all } s \leq t, F \in \mathcal{F}_s;$$

this proves that $X_t - X_s \perp F$ for all $F \in \mathcal{F}_s$.

Let us now prove (17.6). Note that $f(x) := e^{i\xi x}$ is a \mathbb{C}^2 -function with

$$f(x) = e^{i\xi x}, \quad f'(x) = i\xi e^{i\xi x}, \quad f''(x) = -\xi^2 e^{i\xi x}.$$

Applying (16.1) to $f(X_t)$ and $f(X_s)$ and subtracting the respective results yields

$$e^{i\xi X_t} - e^{i\xi X_s} = i\xi \int_s^t e^{i\xi X_r} dX_r - \frac{\xi^2}{2} \int_s^t e^{i\xi X_r} dr. \quad (17.7)$$

Since $|e^{i\eta}| = 1$, the real and imaginary parts of the integrands are $L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$ -functions for any $t \leq T$ and, using Theorem 14.13 a) with X instead of B , we see that $\int_0^t e^{i\xi X_r} dX_r$ is a martingale. Hence,

$$\mathbb{E} \left(\int_s^t e^{i\xi X_r} dX_r \mid \mathcal{F}_s \right) = 0 \quad \text{a. s.}$$

Multiplying both sides of (17.7) by $e^{-i\xi X_s} \mathbb{1}_F$, where $F \in \mathcal{F}_s$, and then taking expectations gives

$$\mathbb{E} \left(e^{i\xi(X_t - X_s)} \mathbb{1}_F \right) - \mathbb{P}(F) = -\frac{\xi^2}{2} \int_s^t \mathbb{E} \left(e^{i\xi(X_r - X_s)} \mathbb{1}_F \right) dr.$$

This means that $\phi(t) := \mathbb{E} \left(e^{i\xi(X_t - X_s)} \mathbb{1}_F \right)$, $s \leq t$, is a solution of the integral equation

$$\Phi(t) = \mathbb{P}(F) + \frac{\xi^2}{2} \int_s^t \Phi(r) dr. \quad (17.8)$$

On the other hand, it is easy to see that (17.8) is also solved by

$$\psi(t) := \mathbb{P}(F) e^{-\frac{1}{2}(t-s)\xi^2}.$$

Since

$$|\phi(t) - \psi(t)| = \frac{\xi^2}{2} \left| \int_s^t (\phi(r) - \psi(r)) dr \right| \leq \frac{\xi^2}{2} \int_s^t |\phi(r) - \psi(r)| dr,$$

we can use Gronwall's lemma, Theorem A.43 in the Appendix A.8, where we set $u = |\phi - \psi|$, $a \equiv 0$ and $b \equiv 1$, to get

$$|\phi(t) - \psi(t)| \leq 0 \quad \text{for all } t \geq 0.$$

This shows that $\phi(t) = \psi(t)$, i. e. (17.8) admits *only one solution*, and (17.6) follows. \square

17.3 Girsanov's theorem

Let $(B_t)_{t \geq 0}$ be a real Brownian motion. The process $W_t := B_t - \ell t$, $\ell \in \mathbb{R}$, is called *Brownian motion with drift*. It describes a uniform motion with speed $-\ell$ which is perturbed by a Brownian motion B_t . Girsanov's theorem provides a very useful transformation of the underlying probability measure \mathbb{P} such that, under the new measure, W_t is itself a Brownian motion. This is very useful if we want to calculate, e. g. the distribution of the random variable $\sup_{s \leq t} W_s$ and other functionals of W .

17.6 Example. Let $G \sim N(m, \sigma^2)$ be on $(\Omega, \mathcal{A}, \mathbb{P})$ a real-valued normal random variable with mean m and variance σ^2 . We want to transform G into a mean zero normal random variable. Obviously, $G \rightsquigarrow G - m$ would do the trick on $(\Omega, \mathcal{A}, \mathbb{P})$ but this would also change G . Another possibility is to change the probability measure \mathbb{P} .

Observe that $\mathbb{E} e^{i\xi G} = e^{-\frac{1}{2}\sigma^2\xi^2 + im\xi}$. Since G has exponential moments, cf. Corollary 2.2, we can insert $\xi = im/\sigma^2$ and find

$$\mathbb{E} \exp\left(-\frac{m}{\sigma^2} G\right) = \exp\left(-\frac{1}{2} \frac{m^2}{\sigma^2}\right).$$

Therefore, $\mathbb{Q}(d\omega) := \exp(-\frac{m}{\sigma^2} G(\omega) + \frac{1}{2} \frac{m^2}{\sigma^2}) \mathbb{P}(d\omega)$ is a probability measure, and we find for the corresponding mathematical expectation $\mathbb{E}_{\mathbb{Q}} = \int_{\Omega} \cdots d\mathbb{Q}$

$$\begin{aligned} \mathbb{E}_{\mathbb{Q}} e^{i\xi G} &= \int e^{i\xi G} \exp\left(-\frac{m}{\sigma^2} G + \frac{1}{2} \frac{m^2}{\sigma^2}\right) d\mathbb{P} \\ &= \exp\left(\frac{1}{2} \frac{m^2}{\sigma^2}\right) \int \exp\left(i\left(\xi + i\frac{m}{\sigma^2}\right) G\right) d\mathbb{P} \\ &= \exp\left(\frac{1}{2} \frac{m^2}{\sigma^2}\right) \exp\left(-\frac{1}{2} \left(\xi + i\frac{m}{\sigma^2}\right)^2 \sigma^2 + im\left(\xi + i\frac{m}{\sigma^2}\right)\right) \\ &= \exp\left(-\frac{1}{2} \sigma^2 \xi^2\right). \end{aligned}$$

This means that, under \mathbb{Q} , we have $G \sim N(0, \sigma^2)$.

The case of a Brownian motion with drift is similar. First we need a formula which shows how conditional expectations behave under changes of measures.

Let β be a positive random variable with $\mathbb{E} \beta = 1$. Then $\mathbb{Q}(d\omega) := \beta(\omega) \mathbb{P}(d\omega)$ is a new probability measure. If \mathcal{F} is a σ -algebra in \mathcal{A} , then

$$\int_F \mathbb{E}(\beta | \mathcal{F}) d\mathbb{P} = \int_F \beta d\mathbb{P} = \mathbb{Q}(F) \quad \text{i. e.} \quad \mathbb{Q}|_{\mathcal{F}} = \mathbb{E}(\beta | \mathcal{F}) \mathbb{P}|_{\mathcal{F}}.$$

If we write \mathbb{E}_Q for the expectation w. r. t. the measure Q , we see for all $F \in \mathcal{F}$

$$\begin{aligned}\mathbb{E}_Q(\mathbb{1}_F X) &= \mathbb{E}(\mathbb{1}_F X\beta) = \mathbb{E}(\mathbb{1}_F \mathbb{E}(X\beta | \mathcal{F})) = \mathbb{E}\left(\mathbb{1}_F \frac{\overbrace{\mathbb{E}(X\beta | \mathcal{F})}^{\mathcal{F} \text{ mble}}}{\mathbb{E}(\beta | \mathcal{F})} \mathbb{E}(\beta | \mathcal{F})\right) \\ &= \mathbb{E}_Q\left(\mathbb{1}_F \frac{\mathbb{E}(X\beta | \mathcal{F})}{\mathbb{E}(\beta | \mathcal{F})}\right).\end{aligned}$$

This means that $\mathbb{E}_Q(X | \mathcal{F}) = \mathbb{E}(X\beta | \mathcal{F}) / \mathbb{E}(\beta | \mathcal{F})$.

17.7 Example. Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM¹ and set $\beta = \beta_T$ where $\beta_t = e^{\ell B_t - \ell^2 t/2}$. From Example 5.2 e) we know that $(\beta_s)_{s \leq T}$ is a martingale for the filtration $(\mathcal{F}_t)_{t \geq 0}$. Therefore, $\mathbb{E} \beta_T = 1$ and $Q := \beta_T \mathbb{P}$ is a probability measure. We want to calculate the conditional characteristic function of $W_t := B_t - \ell t$ under \mathbb{E}_Q . Let $\xi \in \mathbb{R}$ and $s \leq t \leq T$. By the tower property we get

$$\begin{aligned}\mathbb{E}_Q[e^{i\xi(W_t - W_s)} | \mathcal{F}_s] &= \mathbb{E}[e^{i\xi(W_t - W_s)} \beta_T | \mathcal{F}_s] / \mathbb{E}[\beta_T | \mathcal{F}_s] \\ &\stackrel{\text{tower}}{=} \mathbb{E}[e^{i\xi(W_t - W_s)} \mathbb{E}(\beta_T | \mathcal{F}_t) | \mathcal{F}_s] / \beta_s \\ &= \mathbb{E}[e^{i\xi(W_t - W_s)} \beta_t | \mathcal{F}_s] / \beta_s \\ &= e^{-i\ell\xi(t-s) - \frac{1}{2}\ell^2(t-s)} \mathbb{E}[e^{i\xi(B_t - B_s)} e^{\ell(B_t - B_s)} | \mathcal{F}_s] \\ &\stackrel{(B1)}{=} e^{-i\ell\xi(t-s) - \frac{1}{2}\ell^2(t-s)} \mathbb{E}[e^{i\xi(B_t - B_s)} e^{\ell(B_t - B_s)}] \\ &\stackrel{(5.1)}{=} e^{-i\ell\xi(t-s) - \frac{1}{2}\ell^2(t-s)} \mathbb{E}[e^{i\xi(B_t - B_s)} e^{\ell(B_t - B_s)}] \\ &\stackrel{(2.6)}{=} e^{-i\ell\xi(t-s) - \frac{1}{2}\ell^2(t-s)} e^{\frac{1}{2}(i\xi + \ell)^2(t-s)} \\ &= e^{-\frac{1}{2}(t-s)\xi^2}.\end{aligned}$$

Lemma 5.4 shows that $(W_t, \mathcal{F}_t)_{t \leq T}$ is a Brownian motion under Q .

We will now consider the general case.

Ex. 17.3 17.8 Theorem (Girsanov 1960). Let $(B_t, \mathcal{F}_t)_{t \geq 0}$, $B_t = (B_t^1, \dots, B_t^d)$, be a d -dimensional Brownian motion and $f = (f_1, \dots, f_d)$ be a \mathcal{P} measurable process such that $f_j \in L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$ for all $0 < T < \infty$ and $j = 1, \dots, d$. If


$$q_t = \exp\left(\sum_{j=1}^d \int_0^t f_j(s) dB_s^j - \frac{1}{2} \int_0^t |f(s)|^2 ds\right) \quad (17.9)$$

is integrable with $\mathbb{E} q_t \equiv 1$, then for every $T > 0$ the process

$$W_t := B_t - \int_0^t f(s) ds, \quad t \leq T, \quad (17.10)$$

is a BM^d under the new probability measure $\mathbb{Q} = \mathbb{Q}_T$

$$\mathbb{Q}(d\omega) := q_T(\omega) \mathbb{P}(d\omega). \quad (17.11)$$

Notice that the new probability measure \mathbb{Q} in Theorem 17.8 depends on the right endpoint T . 

Proof. Since $f_j \in L^2_{\mathcal{F}}(\lambda_T \otimes \mathbb{P})$ for all $T > 0$, the processes

$$X_t^j := \int_0^t f_j(s) dB_s^j - \frac{1}{2} \int_0^t f_j^2(s) ds, \quad j = 1, \dots, d,$$

are well-defined, and Theorem 17.2 shows that under the condition $\mathbb{E} q_t = 1$ the process $(q_t, \mathcal{F}_t)_{t \leq T}$ is a martingale. Set

$$\beta_t := \exp \left[i \langle \xi, W(t) \rangle + \frac{1}{2} t |\xi|^2 \right]$$

where $W_t = B_t - \int_0^t f(s) ds$ is as in (17.10). If we *knew* that $(\beta_t)_{t \leq T}$ is a martingale on the probability space $(\Omega, \mathcal{F}_T, \mathbb{Q} = \mathbb{Q}_T)$, then

$$\begin{aligned} \mathbb{E}_{\mathbb{Q}} \left(\exp[i \langle \xi, W(t) - W(s) \rangle] \mid \mathcal{F}_s \right) &= \mathbb{E}_{\mathbb{Q}} \left(\beta_t \exp[-i \langle \xi, W(s) \rangle] \exp[-\frac{1}{2} t |\xi|^2] \mid \mathcal{F}_s \right) \\ &= \mathbb{E}_{\mathbb{Q}} \left(\beta_t \mid \mathcal{F}_s \right) \exp[-i \langle \xi, W(s) \rangle] \exp[-\frac{1}{2} t |\xi|^2] \\ &= \beta_s \exp[-i \langle \xi, W(s) \rangle] \exp[-\frac{1}{2} t |\xi|^2] \\ &= \exp[-\frac{1}{2} (t-s) |\xi|^2]. \end{aligned}$$

Lemma 5.4 shows that $W(t)$ is a Brownian motion with respect to the probability measure \mathbb{Q} .

Let us check that $(\beta_t)_{t \leq T}$ is a martingale. We have

$$\begin{aligned} &\underbrace{\sum_{j=1}^d \int_0^t f_j(s) dB_s^j - \frac{1}{2} \int_0^t |f(s)|^2 ds + i \sum_{j=1}^d \xi_j B_t^j}_{\text{exponent of } q_t} - \underbrace{i \sum_{j=1}^d \int_0^t \xi_j f_j(s) ds + \frac{1}{2} |\xi|^2 t}_{\text{exponent of } \beta_t} \\ &= \sum_{j=1}^d \int_0^t (f_j(s) + i \xi_j) dB_s^j - \frac{1}{2} \int_0^t \langle f(s) + i \xi, f(s) + i \xi \rangle ds. \end{aligned}$$

This shows that

$$q_t \beta_t = \exp \left[\sum_{j=1}^d \int_0^t (f_j(s) + i \xi_j) dB_s^j - \frac{1}{2} \int_0^t \langle f(s) + i \xi, f(s) + i \xi \rangle ds \right],$$

and, just as for q_t , we find for the product $q_t \beta_t$ by Itô's formula that

$$q_t \beta_t = 1 + \sum_{j=1}^d \int_0^t q_s \beta_s (f_j(s) + i \xi_j) dB_s^j.$$

Therefore, $(q_t \beta_t)_{t \geq 0}$ is a local martingale. If $(\tau_k)_{k \geq 0}$ is a localizing sequence, we get for $k \geq 1$, $F \in \mathcal{F}_s$ and $s \leq t \leq T$

$$\begin{aligned} \mathbb{E}_Q(\beta_{t \wedge \tau_k} \mathbb{1}_F) &\stackrel{\text{def}}{=} \mathbb{E}(q_T \beta_{t \wedge \tau_k} \mathbb{1}_F) \\ &\stackrel{\text{tower}}{=} \mathbb{E}[\mathbb{E}(q_T \beta_{t \wedge \tau_k} \mathbb{1}_F \mid \mathcal{F}_t)] \\ &= \mathbb{E}[\underbrace{\mathbb{E}(q_T \mid \mathcal{F}_t)}_{= q_t, \text{ martingale}} \beta_{t \wedge \tau_k} \mathbb{1}_F] \\ &= \mathbb{E}[(q_t - q_{t \wedge \tau_k}) \beta_{t \wedge \tau_k} \mathbb{1}_F] + \mathbb{E}[q_{t \wedge \tau_k} \beta_{t \wedge \tau_k} \mathbb{1}_F] \\ &= \mathbb{E}[(q_t - q_{t \wedge \tau_k}) \beta_{t \wedge \tau_k} \mathbb{1}_F] + \mathbb{E}[q_{s \wedge \tau_k} \beta_{s \wedge \tau_k} \mathbb{1}_F]. \end{aligned}$$

In the last step we used the fact that $(q_t \beta_t, \mathcal{F}_t)_{t \geq 0}$ is a local martingale. Since $t \mapsto q_t$ is continuous and $q_{t \wedge \tau_k} = \mathbb{E}(q_t \mid \mathcal{F}_{t \wedge \tau_k})$, we know that $(q_{t \wedge \tau_k}, \mathcal{F}_{t \wedge \tau_k})_{k \geq 1}$ is a uniformly integrable martingale, thus, $\lim_{k \rightarrow \infty} q_{t \wedge \tau_k} = q_t$ a. s. and in $L^1(\mathbb{P})$. Moreover, we have that $|\beta_t| = e^{\frac{1}{2}|\xi|^2 t} \leq e^{\frac{1}{2}|\xi|^2 T}$. By dominated convergence,

$$\begin{aligned} \mathbb{E}_Q(\beta_t \mathbb{1}_F) &= \lim_{k \rightarrow \infty} \mathbb{E}_Q(\beta_{t \wedge \tau_k} \mathbb{1}_F) \\ &= \lim_{k \rightarrow \infty} \mathbb{E}[(q_t - q_{t \wedge \tau_k}) \beta_{t \wedge \tau_k} \mathbb{1}_F] + \lim_{k \rightarrow \infty} \mathbb{E}[q_{s \wedge \tau_k} \beta_{s \wedge \tau_k} \mathbb{1}_F] \\ &= \mathbb{E}[q_s \beta_s \mathbb{1}_F]. \end{aligned}$$

If we take $t = s$ this also shows

$$\mathbb{E}_Q(\beta_t \mathbb{1}_F) = \mathbb{E}[q_s \beta_s \mathbb{1}_F] \stackrel{t=s}{=} \mathbb{E}_Q(\beta_s \mathbb{1}_F) \quad \text{for all } F \in \mathcal{F}_s, 0 \leq s \leq T,$$

i. e. $(\beta_t, \mathcal{F}_t)_{0 \leq t \leq T}$ is a martingale under Q for every $T > 0$. □

17.4 Martingale representation – 1

Let $(B_t)_{t \geq 0}$ be a BM^1 and $(\mathcal{F}_t)_{t \geq 0}$ an admissible complete filtration. From Theorem 14.13 we know that for every $T \in [0, \infty]$ and $X \in L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$ the stochastic integral

$$M_t = x + \int_0^t X_s dB_s, \quad t \geq 0, x \in \mathbb{R}, \quad (17.12)$$

is a continuous L^2 martingale for the filtration $(\mathcal{F}_t)_{t \leq T}$ of the underlying Brownian motion. We will now show the converse: If $(M_t, \mathcal{F}_t)_{t \leq T}$ is an L^2 martingale for the complete Brownian filtration $(\mathcal{F}_t)_{t \geq 0}$, then M_t must be of the form (17.12) for some $X \in L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$.

For this we define the set

$$\mathcal{H}_T^2 := \left\{ M_T : M_T = x + \int_0^T X_s dB_s, x \in \mathbb{R}, X \in L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P}) \right\}.$$

17.9 Lemma. *The space \mathcal{H}_T^2 is a closed linear subspace of $L^2(\Omega, \mathcal{F}_T, \mathbb{P})$.*

Proof. \mathcal{H}_T^2 is obviously a linear space and, by Theorem 14.13, it is contained in $L^2(\Omega, \mathcal{F}_T, \mathbb{P})$. We show that it is a closed subspace. Let $(M_T^n)_{n \geq 1} \subset \mathcal{H}_T^2$,

$$M_T^n = x^n + \int_0^T X_s^n dB_s,$$

be an L^2 Cauchy sequence; since $L^2(\Omega, \mathcal{F}_T, \mathbb{P})$ is complete, the sequence converges to some limit $M \in L^2(\Omega, \mathcal{F}_T, \mathbb{P})$. We have to show that $M \in \mathcal{H}_T^2$. Using the Itô isometry we see for $m, n \geq 1$

$$\begin{aligned} \mathbb{E} \left[((M_T^n - x^n) - (M_T^m - x^m))^2 \right] &= \mathbb{E} \left[\left(\int_0^T (X_s^n - X_s^m) dB_s \right)^2 \right] \\ &\stackrel{(14.20)}{=} \mathbb{E} \left[\int_0^T |X_s^n - X_s^m|^2 ds \right]. \end{aligned}$$

This shows that $(X_s^n)_{s \leq T}$ is a Cauchy sequence in $L^2(\lambda_T \otimes \mathbb{P})$. Because of the completeness of $L^2(\lambda_T \otimes \mathbb{P})$ the sequence converges to some process $(X_s)_{s \leq T}$ and, for some subsequence, we get

$$\lim_{k \rightarrow \infty} X_s^{n(k)}(\omega) = X_s(\omega) \quad \text{for } \lambda_T \otimes \mathbb{P} \text{ almost all } (s, \omega) \text{ and in } L^2(\lambda_T \otimes \mathbb{P}).$$

Since pointwise limits preserve measurability, we have $(X_s)_{s \leq T} \in L^2_{\mathcal{P}}(\lambda_T \otimes \mathbb{P})$. Using again Itô's isometry, we find

$$\int_0^T X_s^n dB_s \xrightarrow[n \rightarrow \infty]{L^2(\Omega, \mathcal{F}_T, \mathbb{P})} \int_0^T X_s dB_s.$$

As L^2 -convergence implies L^1 -convergence, we get

$$\lim_{n \rightarrow \infty} x^n \stackrel{14.13 \text{ a)}}{=} \lim_{n \rightarrow \infty} \mathbb{E} \left[x^n + \int_0^T X_s^n dB_s \right] = \mathbb{E}[M],$$

and this shows that $M_T = \mathbb{E} M_T + \int_0^T X_s dB_s \in \mathcal{H}_T^2$. □

Ex. 17.6 The set \mathcal{H}_T^2 is quite large. For example, $e^{i\xi B_T}$ is for all $\xi \in \mathbb{R}$ in $\mathcal{H}_T^2 \oplus i\mathcal{H}_T^2$. This follows immediately if we apply Itô's formula to e^{M_t} with $M_t = i\xi B_t + \frac{t}{2}\xi^2$:

$$e^{i\xi B_T + \frac{T}{2}\xi^2} = 1 + \int_0^T i\xi e^{M_s} dB_s,$$

hence,

$$e^{i\xi B_T} = e^{-\frac{T}{2}\xi^2} + \int_0^T i\xi \exp\left[i\xi B_s + \frac{(s-T)}{2}\xi^2\right] dB_s,$$

and the integrand is obviously in $L_{\mathcal{P}}^2(\lambda_T \otimes \mathbb{P}) \oplus iL_{\mathcal{P}}^2(\lambda_T \otimes \mathbb{P})$.

17.10 Theorem. Let $(\mathcal{F}_t)_{t \geq 0}$ be an admissible complete filtration for the BM¹ $(B_t)_{t \geq 0}$. Every $Y \in L^2(\Omega, \mathcal{F}_T, \mathbb{P})$ is of the form $Y = y + \int_0^T X_s dB_s$ for some $y \in \mathbb{R}$ and $X \in L_{\mathcal{P}}^2(\lambda_T \otimes \mathbb{P})$.

Proof. We have to show that $\mathcal{H}_T^2 = L^2(\Omega, \mathcal{F}_T, \mathbb{P})$. Since \mathcal{H}_T^2 is a closed subspace of $L^2(\Omega, \mathcal{F}_T, \mathbb{P})$, it is enough to prove that \mathcal{H}_T^2 is dense in $L^2(\Omega, \mathcal{F}_T, \mathbb{P})$.

To do so, we take any $Y \in L^2(\Omega, \mathcal{F}_T, \mathbb{P})$ and show that

$$Y \perp \mathcal{H}_T^2 \implies Y = 0.$$

Recall that $Y \perp \mathcal{H}_T^2$ means that $\mathbb{E}(YH) = 0$ for all $H \in \mathcal{H}_T^2$. Take $H = e^{i\xi B_T}$. As $H \in \mathcal{H}_T^2 \oplus i\mathcal{H}_T^2$, we get

$$0 = \mathbb{E}[e^{i\xi B_T} Y] = \mathbb{E}[e^{i\xi B_T} \mathbb{E}(Y | B_T)] = \mathbb{E}[e^{i\xi B_T} u(B_T)].$$

Here we used the tower property and the fact that $\mathbb{E}(Y | B_T)$ can be expressed as a function u of B_T . This shows

$$\int e^{i\xi x} u(x) \mathbb{P}(B_T \in dx) = \frac{1}{\sqrt{2\pi T}} \int e^{i\xi x} u(x) e^{-\frac{x^2}{2T}} dx = 0,$$

i.e. the Fourier transform of the function $u(x)e^{-\frac{x^2}{2T}}$ is zero. This means that $u(x)e^{-\frac{x^2}{2T}} = 0$, hence $u(x) = 0$, for Lebesgue almost all x , hence $Y = 0$. \square

17.11 Corollary. Let $(M_t)_{t \leq T}$ be any L^2 martingale with respect to a complete admissible filtration $(\mathcal{F}_t)_{t \leq T}$ of a BM¹ $(B_t)_{t \geq 0}$. Then there exist some $x \in \mathbb{R}$ and $X \in L_{\mathcal{P}}^2(\lambda_T \otimes \mathbb{P})$ such that

$$M_t = x + \int_0^t X_s dB_s \text{ for all } t \leq T.$$

In particular, any martingale with respect to a complete Brownian filtration must be continuous.

Proof. By Theorem 17.10 we know that $M_T = x + \int_0^T X_s dB_s$ for some $x \in \mathbb{R}$ and $X \in L^2_{\mathcal{F}}(\lambda_T \otimes \mathbb{P})$. Define $M_t := \mathbb{E}(M_T | \mathcal{F}_t)$, $t \leq T$. By Theorem 14.13 a) we see that

$$M_t = \mathbb{E}(M_T | \mathcal{F}_t) = \mathbb{E}\left(x + \int_0^T X_s dB_s \mid \mathcal{F}_t\right) = x + \int_0^t X_s dB_s. \quad \square$$

17.5 Martingale representation – 2

In the previous section we have seen that, given a Brownian motion $(B_t, \mathcal{F}_t)_{t \geq 0}$ with a complete filtration, every L^2 martingale $(M_t, \mathcal{F}_t)_{t \leq T}$ can be represented as a stochastic integral with respect to $(B_t)_{t \geq 0}$. In particular, every L^2 martingale with respect to a complete Brownian filtration is continuous.

We will now show that for many continuous L^2 martingales $(M_t, \mathcal{F}_t)_{t \leq T}$ there is a Brownian motion $(W_t)_{t \geq 0}$ such that M_t can be written as a stochastic integral with respect to $(W_t)_{t \geq 0}$. This requires a general martingale stochastic integral as in Section 14.6. Recall that the quadratic variation of an L^2 martingale is the unique, increasing process $\langle M \rangle = (\langle M \rangle_t)_{t \leq T}$ with $\langle M \rangle_0 = 0$ such that $(M_t^2 - \langle M \rangle_t, \mathcal{F}_t)_{t \leq T}$ is a martingale. In particular,

$$\mathbb{E}[(M_t - M_s)^2 | \mathcal{F}_s] = \mathbb{E}[\langle M \rangle_t - \langle M \rangle_s | \mathcal{F}_s] \quad \text{for all } s \leq t \leq T.$$

17.12 Theorem (Doob 1953). *Let $(M_t, \mathcal{F}_t)_{t \leq T}$ be a continuous L^2 martingale such that $\langle M \rangle_t = \int_0^t m^2(s, \cdot) ds$ for some $\lambda_T \otimes \mathbb{P}$ -almost everywhere strictly positive process $m(s, \omega)$. Then there exists a Brownian motion $(W_t)_{t \leq T}$ such that the filtration $(\mathcal{F}_t)_{t \leq T}$ is admissible and*

$$M_t - M_0 = \int_0^t m(s, \cdot) dW_s.$$

Proof. Since the set $\{(s, \omega) : m(s, \omega) = 0\}$ is a $\lambda_T \otimes \mathbb{P}$ null set, we can define in all calculations below that $1/m(s, \omega) := 0$ if $m(s, \omega) = 0$. Note that

$$\mathbb{E} \int_0^t \frac{1}{m^2(s, \cdot)} d\langle M \rangle_s = \mathbb{E} \int_0^t \frac{1}{m^2(s, \cdot)} m^2(s, \cdot) ds = t.$$

Therefore, Theorem 14.22 shows that the stochastic integral

$$W_t := \int_0^t \frac{1}{m(s, \cdot)} dM_s, \quad t \leq T,$$

exists and defines a continuous L^2 martingale. Again by Theorem 14.22 the quadratic variation is

$$\langle W \rangle_t = \left\langle \int_0^t \frac{1}{m(s, \cdot)} dM_s \right\rangle_t = \int_0^t \frac{1}{m^2(s, \cdot)} d\langle M \rangle_s = \int_0^t \frac{m^2(s, \cdot)}{m^2(s, \cdot)} ds = t.$$

Lévy's martingale characterization of a Brownian motion, Theorem 17.5, shows that $(W_t, \mathcal{F}_t)_{t \leq T}$ is indeed a Brownian motion. Finally,

$$\int_0^t m(s, \cdot) dW_s = \int_0^t m(s, \cdot) \frac{1}{m(s, \cdot)} dM_s = \int_0^t dM_s = M_t - M_0. \quad \square$$

If $\{(s, \omega) : m(s, \omega) = 0\}$ is not a null set, we have to modify our argument. In this case the probability space is too small and we have to extend it by adding an independent Brownian motion $(B_t, \mathcal{G}_t)_{t \geq 0}$.² If U and V are mean zero \mathcal{F}_t resp. \mathcal{G}_t measurable random variables such that $\mathbb{E}(U | \mathcal{F}_s) = \mathbb{E}(V | \mathcal{G}_s) = 0$ for $s \leq t$, we find because of independence, that $\mathbb{E}(UV | \sigma(\mathcal{F}_s, \mathcal{G}_s)) = 0$. This follows from

$$\int_{F \cap G} UV d\mathbb{P} = \mathbb{E}(U \mathbb{1}_F V \mathbb{1}_G) = \int_F U d\mathbb{P} \int_G V d\mathbb{P} = 0$$

for all $F \in \mathcal{F}_s$, $G \in \mathcal{G}_s$, and the fact that $\{F \cap G : F \in \mathcal{F}_s, G \in \mathcal{G}_s\}$ is a \cap -stable generator of $\sigma(\mathcal{F}_s, \mathcal{G}_s)$.

17.13 Corollary (Doob 1953). *Let $(M_t, \mathcal{F}_t)_{t \leq T}$ be a continuous L^2 martingale such that $\langle M \rangle_t = \int_0^t m^2(s, \cdot) ds$. Then there is an enlargement $(\tilde{\Omega}, \tilde{\mathcal{A}}, \tilde{\mathbb{P}})$ of the underlying probability space $(\Omega, \mathcal{A}, \mathbb{P})$ and a Brownian motion $(W_t)_{t \leq T}$ on $(\tilde{\Omega}, \tilde{\mathcal{A}}, \tilde{\mathbb{P}})$ such that*

$$M_t - M_0 = \int_0^t m(s, \cdot) dW_s.$$

Proof. If necessary, we enlarge $(\Omega, \mathcal{A}, \mathbb{P})$ so that there is a further Brownian motion $(B_t, \mathcal{G}_t)_{t \leq T}$ which is independent of $(M_t, \mathcal{F}_t)_{t \leq T}$. Set $\mathcal{H}_t := \sigma(\mathcal{F}_t, \mathcal{G}_t)$. Because of the independence, both $(M_t, \mathcal{H}_t)_{t \leq T}$ and $(B_t, \mathcal{H}_t)_{t \leq T}$ are again L^2 martingales. Therefore, Theorem 14.22 shows that the stochastic integrals

Ex. 17.7

$$W_t := \int_0^t \frac{1}{m(s, \cdot)} \mathbb{1}_{\{m \neq 0\}}(s, \cdot) dM_s + \int_0^t \mathbb{1}_{\{m=0\}}(s, \cdot) dB_s, \quad s \leq T,$$

exist and define a continuous L^2 martingale for the filtration $(\mathcal{H}_t)_{t \leq T}$. From (14.29)

² This can be done by the usual product construction.

and the calculation just before the statement of Corollary 17.13 we see

$$\begin{aligned}
& \mathbb{E}[(W_t - W_s)^2 \mid \mathcal{H}_s] \\
&= \mathbb{E}\left[\left(\int_s^t \frac{1}{m(r, \cdot)} \mathbb{1}_{\{m \neq 0\}}(r, \cdot) dM_r\right)^2 \mid \mathcal{H}_s\right] \\
&\quad + \mathbb{E}\left[\left(\int_s^t \mathbb{1}_{\{m=0\}}(r, \cdot) dB_r\right)^2 \mid \mathcal{H}_s\right] \\
&\quad + 2 \mathbb{E}\left[\int_s^t \frac{1}{m(r, \cdot)} \mathbb{1}_{\{m \neq 0\}}(r, \cdot) dM_r \int_s^t \mathbb{1}_{\{m=0\}}(r, \cdot) dB_r \mid \mathcal{H}_s\right] \\
&= \mathbb{E}\left[\int_s^t \frac{1}{m^2(r, \cdot)} \mathbb{1}_{\{m \neq 0\}}(r, \cdot) d\langle M \rangle_r \mid \mathcal{H}_s\right] + \mathbb{E}\left[\int_s^t \mathbb{1}_{\{m=0\}}(r, \cdot) dr \mid \mathcal{H}_s\right] \\
&= \mathbb{E}\left[\int_s^t \frac{m^2(r, \cdot)}{m^2(r, \cdot)} \mathbb{1}_{\{m \neq 0\}}(r, \cdot) dr \mid \mathcal{H}_s\right] + \mathbb{E}\left[\int_s^t \mathbb{1}_{\{m=0\}}(r, \cdot) dr \mid \mathcal{H}_s\right] \\
&= \int_s^t (\mathbb{1}_{\{m \neq 0\}}(r, \cdot) + \mathbb{1}_{\{m=0\}}(r, \cdot)) dr = t - s.
\end{aligned}$$

By Lévy's theorem 17.5, $(W_t, \mathcal{H}_t)_{t \leq T}$ is a Brownian motion, and the rest of the proof follows along the lines of Theorem 17.12. \square

17.6 Martingales as time-changed Brownian motion

Throughout this section $(M_t, \mathcal{F}_t)_{t \geq 0}$ is a real valued, continuous martingale such that $M_t \in L^2(\mathbb{P})$. We show that there is a Brownian motion $(B_t)_{t \geq 0}$ such that $M_t = B_{\sigma(t)}$ for some random time $\sigma(t)$.

We begin with a few preparations from analysis. Let $a : [0, \infty) \rightarrow [0, \infty]$ be a right-continuous, increasing function. Then

$$\tau(s) := \inf\{t \geq 0 : a(t) > s\}, \quad \inf \emptyset = \infty, \quad (17.13)$$

is the *generalized inverse* of a .

17.14 Lemma. *Let τ be the generalized inverse of a right-continuous, increasing function $a : [0, \infty) \rightarrow [0, \infty]$.* Ex. 17.8

- a) $\tau : [0, \infty) \rightarrow [0, \infty]$ is increasing and right-continuous;
- b) $a(\tau(s)) \geq s$;
- c) $a(t) \geq s \iff \tau(s-) \leq t$;
- d) $a(t) = \inf\{s \geq 0 : \tau(s) > t\}$;

- e) $a(\tau(s)) = \sup\{u \geq s : \tau(u) = \tau(s)\};$
 f) $\tau(a(t)) = \sup\{w \geq t : a(w) = a(t)\};$
 g) if a is continuous, $a(\tau(s)) = s$.

Proof. Before we begin with the formal proof, it is helpful to draw a picture of the situation, cf. Figure 17.1. All assertions are clear if t is a point of strict increase for the function a . Moreover, if a is flat, the inverse τ makes a jump, and vice versa.

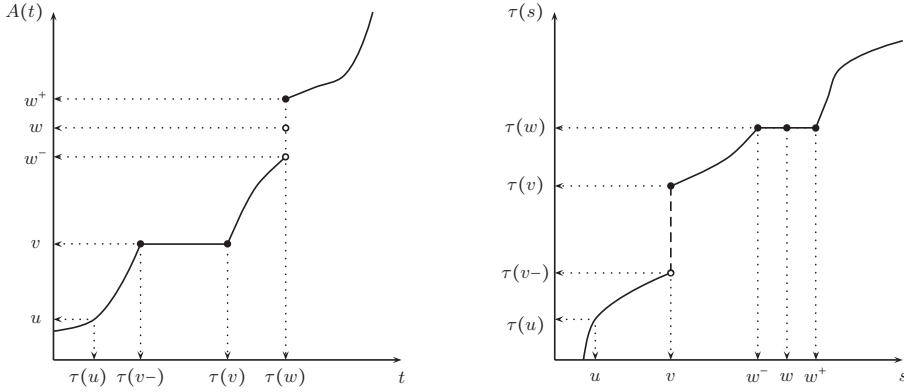


Figure 17.1. An increasing right-continuous function and its generalized right-continuous inverse.

- a) It is obvious from the definition that τ is increasing. Note that

$$\{t : a(t) > s\} = \bigcup_{\epsilon > 0} \{t : a(t) > s + \epsilon\}.$$

Therefore, $\inf\{t \geq 0 : a(t) > s\} = \inf_{\epsilon > 0} \inf\{t \geq 0 : a(t) > s + \epsilon\}$ proving right-continuity.

- b) Since $a(u) \geq s$ for all $u > \tau(s)$, we see that $a(\tau(s)) \geq s$.
 c) By the very definition of the generalized inverse τ we have

$$\begin{aligned} a(t) \geq s &\iff a(t) > s - \epsilon \quad \forall \epsilon > 0 \\ &\iff \tau(s - \epsilon) \leq t \quad \forall \epsilon > 0 \iff \tau(s-) \leq t. \end{aligned}$$

- d) From c) we see that

$$\begin{aligned} a(t) &= \sup\{s \geq 0 : a(t) \geq s\} \stackrel{c)}{=} \sup\{s \geq 0 : \tau(s-) \leq t\} \\ &= \inf\{s \geq 0 : \tau(s) > t\}. \end{aligned}$$

- e) Assume that $u \geq s$ and $\tau(u) = \tau(s)$. Then $a(\tau(s)) = a(\tau(u)) \geq u \geq s$, and so $a(\tau(s)) \geq \sup\{u > s : \tau(u) = \tau(s)\}$. For the converse, we use d) and find

$$a(\tau(s)) = \inf\{r \geq 0 : \tau(r) > \tau(s)\} \leq \sup\{u \geq s : \tau(u) = \tau(s)\}.$$

- f) Follows as e) since a and τ play symmetric roles.
- g) Because of b) it is enough to show that $a(\tau(s)) \leq s$. If $\tau(s) = 0$, this is trivial. Otherwise we find $\epsilon > 0$ such that $0 < \epsilon < \tau(s)$ and $a(\tau(s) - \epsilon) \leq s$ by the definition of $\tau(s)$. Letting $\epsilon \rightarrow 0$, we obtain $a(\tau(s)) \leq s$ by the continuity of a . \square

The following theorem was discovered by Doeblin in 1940. Doeblin never published this result but deposited it in a *pli cacheté* with the Academy of Sciences in Paris. The *pli* was finally opened in the year 2000, see the historical notes by Bru and Yor [17]. Around 25 years after Doeblin, the result was in 1965 rediscovered by Dambis and by Dubins and Schwarz in a more general context.

17.15 Theorem (Doeblin 1940; Dambis 1965; Dubins–Schwarz 1965). *Let a continuous martingale $(M_t, \mathcal{F}_t)_{t \geq 0}$, $M_t \in L^2(\mathbb{P})$, be given. Assume that the filtration is right-continuous, i. e. $\mathcal{F}_t = \mathcal{F}_{t+}$, $t \geq 0$, and that the quadratic variation satisfies $\langle M \rangle_\infty = \lim_{t \rightarrow \infty} \langle M \rangle_t = \infty$. Denote by $\tau_s = \inf\{t \geq 0 : \langle M \rangle_t > s\}$ the generalized inverse of $\langle M \rangle$. Then $(B_t, \mathcal{G}_t)_{t \geq 0} := (M_{\tau_s}, \mathcal{F}_{\tau_s})_{s \geq 0}$ is a Brownian motion and*

$$M_t = B_{\langle M \rangle_t} \text{ for all } t \geq 0.$$

Proof. Define $B_s := M_{\tau_s}$ and $\mathcal{G}_t := \mathcal{F}_{\tau_t}$. Since τ_s is the first entry time into (s, ∞) for the process $\langle M \rangle$, it is an \mathcal{F}_{t+} stopping time, cf. Lemma 5.7, which is almost surely finite because of $\langle M \rangle_\infty = \infty$.

Ex. 17.9

Let $\Omega^* = \{\omega \in \Omega : M_\cdot(\omega) \text{ and } \langle M \rangle_\cdot(\omega) \text{ are constant on the same intervals}\}$. From the definition of the quadratic variation, Theorem A.31, one can show that $\mathbb{P}(\Omega^*) = 1$, see Corollary A.32 in the appendix.

In particular, $s \mapsto M_{\tau_s}(\omega)$ is almost surely continuous: If s is a continuity point of τ_s , this is obvious. Otherwise $\tau_{s-} < \tau_s$, but for $\omega \in \Omega^*$ and all $\tau_{s-}(\omega) \leq r \leq \tau_s(\omega)$ we have $M_{\tau_{s-}}(\omega) = M_r(\omega) = M_{\tau_s}(\omega)$.

For all $k \geq 1$ and $s \geq 0$ we find

$$\mathbb{E}(M_{\tau_s \wedge k}^2) \leq \mathbb{E}(\sup_{u \leq k} M_u^2) \stackrel{\text{Doob}}{\leq} 4 \mathbb{E} M_k^2 < \infty. \quad (\text{A.14})$$

The optional stopping Theorem A.18 and Corollary A.20 show that $(M_{\tau_s \wedge k})_{s \geq 0}$ is a martingale for the filtration $(\mathcal{G}_s)_{s \geq 0}$ with $\mathcal{G}_s := \mathcal{F}_{\tau_s}$. Thus,

$$\mathbb{E}(M_{\tau_s \wedge k}^2) = \mathbb{E}(\langle M \rangle_{\tau_s \wedge k}) \leq \mathbb{E}(\langle M \rangle_{\tau_s}) = s,$$

showing that for every $s \geq 0$ the family $(M_{\tau_s \wedge k})_{k \geq 1}$ is uniformly integrable. By the continuity of M we see

$$M_{\tau_s \wedge k} \xrightarrow[k \rightarrow \infty]{L^1 \text{ and a. e.}} M_{\tau_s} =: B_s.$$

In particular, $(B_s, \mathcal{G}_s)_{s \geq 0}$ is a martingale. Fatou's lemma shows that

$$\mathbb{E} B_s^2 \leq \liminf_{k \rightarrow \infty} \mathbb{E} M_{\tau_s \wedge k}^2 \leq s,$$

i. e. B_s is also in $L^2(\mathbb{P})$.

If we apply optional stopping to the martingale $(M_t^2 - \langle M \rangle_t)_{t \geq 0}$, we find

$$\mathbb{E} [(B_t - B_s)^2 | \mathcal{G}_s] = \mathbb{E} [(M_{\tau_t} - M_{\tau_s})^2 | \mathcal{F}_{\tau_s}] = \mathbb{E} [\langle M \rangle_{\tau_t} - \langle M \rangle_{\tau_s} | \mathcal{F}_{\tau_s}] = t - s.$$

Thus $(B_s, \mathcal{G}_s)_{s \geq 0}$ is a continuous martingale whose quadratic variation process is s . Lévy's characterization of a Brownian motion, Theorem 17.5, applies and we see that $(B_s, \mathcal{G}_s)_{s \geq 0}$ is a BM¹. Finally, $B_{\langle M \rangle_t} = M_{\tau_{\langle M \rangle_t}} = M_t$ for all $t \geq 0$. \square

17.7 Burkholder–Davis–Gundy inequalities

The Burkholder–Davis–Gundy inequalities are estimates for the p th moment of a continuous martingale X in terms of the p th norm of $\sqrt{\langle X \rangle}$:

$$c_p \mathbb{E} [\langle X \rangle_T^{p/2}] \leq \mathbb{E} \left[\sup_{t \leq T} |X_t|^p \right] \leq C_p \mathbb{E} [\langle X \rangle_T^{p/2}], \quad 0 < p < \infty. \quad (17.14)$$

The constants can be made explicit and they depend only on p . In fact, (17.14) holds for continuous local martingales; we restrict ourselves to Brownian L^2 martingales, i. e. to martingales of the form

$$X_t = \int_0^t f(s) dB_s \quad \text{and} \quad \langle X \rangle_t = \int_0^t |f(s)|^2 ds$$

where $(B_t)_{t \geq 0}$ is a BM¹ and $f \in L^2_{\mathbb{P}}(\lambda_T \otimes \mathbb{P})$ for all $T > 0$, see also Sections 17.4 and 17.5.

We will first consider the case where $p \geq 2$.

17.16 Theorem (Burkholder 1966). *Let $(B_t)_{t \geq 0}$ be a BM¹ and $f \in L^2_{\mathbb{P}}(\lambda_T \otimes \mathbb{P})$ for all $T > 0$. Then, we have for all $p \geq 2$ and $q = p/(p-1)$*

$$\mathbb{E} \left[\sup_{t \leq T} \left| \int_0^t f(s) dB_s \right|^p \right] \leq C_p \mathbb{E} \left[\left(\int_0^T |f(s)|^2 ds \right)^{p/2} \right], \quad p \geq 2, \quad (17.15)$$

where $C_p = [q^p p(p-1)/2]^{p/2}$, and

$$\mathbb{E} \left[\sup_{t \leq T} \left| \int_0^t f(s) dB_s \right|^p \right] \geq c_p \mathbb{E} \left[\left(\int_0^T |f(s)|^2 ds \right)^{p/2} \right], \quad p \geq 2, \quad (17.16)$$

where $c_p = 1/(2p)^{p/2}$.

Proof. Set $X_t := \int_0^t f(s) dB_s$ and $\langle X \rangle_t = \int_0^t |f(s)|^2 ds$, cf. Theorem 14.13.

If we combine Doob's maximal inequality (A.14) and Itô's isometry (14.20) we get

$$\mathbb{E} [\langle X \rangle_T] = \mathbb{E} [|X_T|^2] \leq \mathbb{E} \left[\sup_{s \leq T} |X_s|^2 \right] \leq 4 \sup_{s \leq T} \mathbb{E} [|X_s|^2] = 4 \mathbb{E} [\langle X \rangle_T].$$

This shows that (17.15) and (17.16) hold for $p = 2$ with $C_p = 4$ and $c_p = 1/4$, respectively.

Assume that $p > 2$. By stopping with the stopping time

$$\tau_n := \inf \{ t \geq 0 : |X_t|^2 + \langle X \rangle_t \geq n \},$$

(note that $X^2 + \langle X \rangle$ is a continuous process) we can, without loss of generality, assume that both X and $\langle X \rangle$ are bounded. Since $x \mapsto |x|^p$, $p > 2$, is a \mathcal{C}^2 -function, Itô's formula, Theorem 16.6, yields

$$|X_t|^p = p \int_0^t |X_s|^{p-1} dX_s + \frac{1}{2} p(p-1) \int_0^t |X_s|^{p-2} d\langle X \rangle_s$$

where we use that $d\langle X \rangle_s = |f(s)|^2 ds$. Because of the boundedness of $|X_s|^{p-1}$, the first integral on the right-hand side is a martingale and we see with Doob's maximal inequality and the Hölder inequality for $p/(p-2)$ and $p/2$

$$\begin{aligned} \mathbb{E} \left[\sup_{s \leq T} |X_s|^p \right] &\stackrel{(A.14)}{\leq} q^p \frac{p(p-1)}{2} \mathbb{E} \left[\int_0^T |X_s|^{p-2} d\langle X \rangle_s \right] \\ &\leq q^p \frac{p(p-1)}{2} \mathbb{E} \left[\sup_{s \leq T} |X_s|^{p-2} \langle X \rangle_T \right] \\ &\leq q^p \frac{p(p-1)}{2} \left(\mathbb{E} \left[\sup_{s \leq T} |X_s|^p \right] \right)^{1-2/p} \left(\mathbb{E} [\langle X \rangle_T^{p/2}] \right)^{2/p}. \end{aligned}$$

If we divide by $(\mathbb{E} [\sup_{s \leq T} |X_s|^p])^{1-2/p}$ we find

$$\left(\mathbb{E} \left[\sup_{s \leq T} |X_s|^p \right] \right)^{2/p} \leq q^p \frac{p(p-1)}{2} (\mathbb{E} [\langle X \rangle_T^{p/2}])^{2/p}$$

and (17.15) follows with $C_p = [q^p p(p-1)/2]^{p/2}$.

For the lower estimate (17.16) we set

$$Y_t := \int_0^t \langle X \rangle_s^{(p-2)/4} dX_s.$$

Ex. 17.10 From Theorem 14.22 (or with Theorem 14.13 observing that $dX_s = f(s) dB_s$) we see that

$$\langle Y \rangle_T = \int_0^T \langle X \rangle_s^{(p-2)/2} d\langle X \rangle_s = \frac{2}{p} \langle X \rangle_T^{p/2}$$

and

$$\mathbb{E}[\langle X \rangle_T^{p/2}] = \frac{p}{2} \mathbb{E}[\langle Y \rangle_T] = \frac{p}{2} \mathbb{E}[|Y_T|^2].$$

Using Itô's formula (16.8) for the function $f(x, y) = xy$ and the Itô process $(X_t, \langle X \rangle_t^{(p-2)/4})$ we obtain

$$\begin{aligned} X_T \langle X \rangle_T^{(p-2)/4} &= \int_0^T \langle X \rangle_s^{(p-2)/4} dX_s + \int_0^T X_s d\langle X \rangle_s^{(p-2)/4} \\ &= Y_T + \int_0^T X_s d\langle X \rangle_s^{(p-2)/4}. \end{aligned}$$

If we rearrange this equality we get $|Y_T| \leq 2 \sup_{s \leq T} |X_s| \langle X \rangle_T^{(p-2)/4}$. Finally, using Hölder's inequality with $p/2$ and $p/(p-2)$,

$$\begin{aligned} \frac{2}{p} \mathbb{E}[\langle X \rangle_T^{p/2}] &= \mathbb{E}[|Y_T|^2] \leq 4 \mathbb{E}\left[\sup_{s \leq T} |X_s|^2 \langle X \rangle_T^{(p-2)/2}\right] \\ &\leq 4 \left(\mathbb{E}\left[\sup_{s \leq T} |X_s|^p\right]\right)^{2/p} (\mathbb{E}[\langle X \rangle_T^{p/2}])^{1-2/p}. \end{aligned}$$

Dividing on both sides by $(\mathbb{E}[\langle X \rangle_T^{p/2}])^{1-2/p}$ gives (17.16) with $c_p = (2p)^{-p/2}$. \square

The inequalities for $0 < p < 2$ can be proved with the following useful lemma.

17.17 Lemma. *Let X and A be positive random variables and $\rho \in (0, 1)$. If*

$$\mathbb{P}(X > x, A \leq y) \leq \frac{1}{x} \mathbb{E}(A \wedge y) \quad \text{for all } x, y > 0, \quad (17.17)$$

then

$$\mathbb{E}(X^\rho) \leq \frac{2-\rho}{1-\rho} \mathbb{E}(A^\rho).$$

Proof. Using (17.17) and Tonelli's theorem we find

$$\begin{aligned}
 \mathbb{E}(X^\rho) &= \rho \int_0^\infty \mathbb{P}(X > x) x^{\rho-1} dx \\
 &\leq \rho \int_0^\infty [\mathbb{P}(X > x, A \leq x) + \mathbb{P}(A > x)] x^{\rho-1} dx \\
 &\leq \rho \int_0^\infty \left[\frac{1}{x} \mathbb{E}(A \wedge x) + \mathbb{P}(A > x) \right] x^{\rho-1} dx \\
 &= \mathbb{E} \left[\int_0^\infty (A \wedge x) \rho x^{\rho-2} dx \right] + \mathbb{E}[A^\rho] \\
 &= \mathbb{E} \left[\int_0^A \rho x^{\rho-1} dx + \int_A^\infty A \rho x^{\rho-2} dx \right] + \mathbb{E}[A^\rho] \\
 &= \mathbb{E} \left[A^\rho - \frac{\rho A^\rho}{\rho-1} \right] + \mathbb{E}[A^\rho] \\
 &= \frac{2-\rho}{1-\rho} \mathbb{E}(A^\rho). \quad \square
 \end{aligned}$$

We want to use Lemma 17.17 for $X = \sup_{s \leq T} |X_s|^2 = \sup_{s \leq T} \left| \int_0^s f(r) dB_r \right|^2$ and $A = \langle X \rangle_T = \int_0^T |f(s)|^2 ds$. Define $\tau := \inf\{t \geq 0 : \langle X \rangle_t \geq y\}$. Then

$$\begin{aligned}
 P &:= \mathbb{P} \left(\sup_{s \leq T} |X_s|^2 > x, \langle X \rangle_T \leq y \right) = \mathbb{P} \left(\sup_{s \leq T} |X_s|^2 > x, \tau \geq T \right) \\
 &= \mathbb{P} \left(\sup_{s \leq T \wedge \tau} |X_s|^2 > x \right)
 \end{aligned}$$

and with Doob's maximal inequality (A.13) we find

$$P \leq \frac{1}{x} \mathbb{E}(|X_{T \wedge \tau}|^2) \stackrel{(*)}{=} \frac{1}{x} \mathbb{E}(\langle X \rangle_{T \wedge \tau}) = \frac{1}{x} \mathbb{E}(\langle X \rangle_T \wedge y).$$

For the equality (*) we used optional stopping and the fact that $(X_t^2 - \langle X \rangle_t)_{t \geq 0}$ is a martingale. Now we can apply Lemma 17.17 and get

$$\mathbb{E} \left[\sup_{s \leq T} |X_s|^{2\rho} \right] \leq \frac{2-\rho}{1-\rho} \mathbb{E}[\langle X \rangle_T^\rho]$$

which is the analogue of (17.15) for $0 < \rho < 2$.

In a similar way we can use Lemma 17.17 for $X = \langle X \rangle_T = \int_0^T |f(s)|^2 ds$ and $A = \sup_{s \leq T} |X_s|^2 = \sup_{s \leq T} \left| \int_0^s f(r) dB_r \right|^2$. Define $\sigma := \inf\{t > 0 : |X_t|^2 \geq y\}$. Then

$$P' := \mathbb{P} \left(\langle X \rangle_T > x, \sup_{s \leq T} |X_s|^2 \leq y \right) \leq \mathbb{P}(\langle X \rangle_T > x, \sigma > T) \leq \mathbb{P}(\langle X \rangle_{T \wedge \sigma} > x).$$

Using the Markov inequality we get

$$P' \leq \frac{1}{x} \mathbb{E}(\langle X \rangle_{T \wedge \sigma}) \leq \frac{1}{x} \mathbb{E} \left(\sup_{s \leq T \wedge \sigma} |X_s|^2 \right) \leq \frac{1}{x} \mathbb{E} \left(\sup_{s \leq T} |X_s|^2 \wedge y \right),$$

and Lemma 17.17 gives

$$\mathbb{E} \left[\sup_{s \leq T} |X_s|^{2\rho} \right] \geq \frac{1-\rho}{2-\rho} \mathbb{E} [\langle X \rangle_T^\rho]$$

which is the analogue of (17.16) for $0 < p < 2$.

Combining this with Theorem 17.16 we have shown

Ex. 17.11 17.18 Theorem (Burkholder, Davis, Gundy 1972). *Let $(B_t)_{t \geq 0}$ be a one-dimensional Brownian motion and $f \in L^2_{\mathcal{F}}(\lambda_T \otimes \mathbb{P})$ for all $T > 0$. Then, we have for $0 < p < \infty$*

$$\mathbb{E} \left[\left(\int_0^T |f(s)|^2 ds \right)^{p/2} \right] \asymp \mathbb{E} \left[\sup_{t \leq T} \left| \int_0^t f(s) dB_s \right|^p \right] \quad (17.18)$$

with finite comparison constants which depend only on $p \in (0, \infty)$.



17.19 Further reading. Further applications, e. g. to reflecting Brownian motion $|B_t|$, excursions, local times etc. are given in [80]. The monograph [109] proves Burkholder's inequalities in \mathbb{R}^d with dimension-free constants. A concise introduction to Malliavin's calculus and stochastic analysis is [171].

[80] Ikeda, Watanabe: *Stochastic Differential Equations and Diffusion Processes*.

[109] Krylov: *Introduction to the Theory of Diffusion Processes*.

[171] Shigekawa: *Stochastic Analysis*.

Problems

1. State and prove the d -dimensional version of Lemma 17.1.
2. Let $(N_t)_{t \geq 0}$ be a Poisson process with intensity $\lambda = 1$ (see Problem 1 for the definition). Show that for $\mathcal{F}_t^N := \sigma(N_r : r \leq t)$ both $X_t := N_t - t$ and $X_t^2 - t$ are martingales. Explain why this does not contradict Theorem 17.5.
3. Let $(B(t))_{t \geq 0}$ be a BM¹ and $T < \infty$. Define $\beta_T = \exp(\xi B(T) - \frac{1}{2} \xi^2 T)$, $\mathbb{Q} := \beta_T \mathbb{P}$ and $W(t) := B(t) - \xi t$, $t \in [0, T]$. Verify by a direct calculation that for all $0 < t_1 < \dots < t_n$ and $A_1, \dots, A_n \in \mathcal{B}(\mathbb{R})$

$$\mathbb{Q}(W(t_1) \in A_1, \dots, W(t_n) \in A_n) = \mathbb{P}(B(t_1) \in A_1, \dots, B(t_n) \in A_n).$$

4. Let $(B_t)_{t \geq 0}$ be a BM¹, $\Phi(y) := \mathbb{P}(B_1 \leq y)$, and set $X_t := B_t + \alpha t$ for some $\alpha \in \mathbb{R}$. Use Girsanov's theorem to show for all $0 \leq x \leq y, t > 0$

$$\mathbb{P}\left(X_t \leq x, \sup_{s \leq t} X_s \leq y\right) = \Phi\left(\frac{x - \alpha t}{\sqrt{t}}\right) - e^{2\alpha y} \Phi\left(\frac{x - 2y - \alpha t}{\sqrt{t}}\right).$$

5. Let X_t be as in Problem 4 and set $\hat{\tau}_b := \inf\{t \geq 0 : X_t = b\}$.
 (a) Use Problem 4 to find the probability density of $\hat{\tau}_b$ if $\alpha, b > 0$.
 (b) Find $\mathbb{P}(\hat{\tau}_b < \infty)$ for $\alpha < 0$ and $\alpha \geq 0$, respectively.
6. Let $(B_t)_{t \geq 0}$ be a BM¹ and denote by \mathcal{H}_T^2 the space used in Lemma 17.9. Show that $\exp[i\xi B_T]$ is contained in $\mathcal{H}_T^2 \oplus i\mathcal{H}_T^2$.
7. Assume that $(B_t, \mathcal{G}_t)_{t \geq 0}$ and $(M_t, \mathcal{F}_t)_{t \geq 0}$ are independent martingales. Show that both $(M_t, \mathcal{H}_t)_{t \leq T}$ and $(B_t, \mathcal{H}_t)_{t \leq T}$ are martingales for the enlarged filtration $\mathcal{H}_t := \sigma(\mathcal{F}_t, \mathcal{G}_t)$.
8. Show the following addition to Lemma 17.14:

$$\tau(t-) = \inf\{s \geq 0 : a(s) \geq t\} \quad \text{and} \quad a(t-) = \inf\{s \geq 0 : \tau(s) \geq t\}$$

$$\text{and } \tau(s) \geq t \iff a(t-) \leq s.$$

9. Show that, in Theorem 17.15, the quadratic variation $\langle M \rangle_t$ is a \mathcal{G}_{τ_s} stopping time.
Hint: Direct calculation, use Lemma 17.14 c) and A.15
10. Let $f \in \mathbf{BV}[0, \infty)$. Show that $\int f^{a-1} df = f^a/a$ for all $a > 0$.
11. State and prove a d -dimensional version of the Burkholder–Davis–Gundy inequalities (17.17) if $p \in [2, \infty)$.
Hint: Use the fact that all norms in \mathbb{R}^d are equivalent

Chapter 18

Stochastic differential equations

The ordinary differential equation $\frac{d}{ds}x_s = b(s, x_s)$, $x(0) = x_0$, describes the position x_t of a particle which moves with speed $b(s, x)$ depending on time and on the current position. One possibility to take into account random effects, e. g. caused by measurement errors or hidden parameters, is to add a random perturbation which may depend on the current position. This leads to an equation of the form

$$X(t + \Delta t) - X(t) = b(t, X(t))\Delta t + \sigma(t, X(t))(B(t + \Delta t) - B(t)).$$

Letting $\Delta t \rightarrow 0$ we get the following equation for *stochastic differentials*, cf. Section 16.1,

$$dX_t = b(t, X_t) dt + \sigma(t, X_t) dB_t. \quad (18.1)$$

Since (18.1) is a formal way of writing things, we need a precise definition of the solution of the stochastic differential equation (18.1).

18.1 Definition. Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM^d with admissible filtration $(\mathcal{F}_t)_{t \geq 0}$, and let $b : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^n$ and $\sigma : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^{n \times d}$ be measurable functions. A *solution* of the *stochastic differential equation (SDE)* (18.1) with initial condition $X_0 = \xi$ is a progressively measurable stochastic process $(X_t)_{t \geq 0}$, $X_t = (X_t^1, \dots, X_t^n) \in \mathbb{R}^n$, such that the following integral equation holds

$$X_t = \xi + \int_0^t b(s, X_s) ds + \int_0^t \sigma(s, X_s) dB_s, \quad t \geq 0, \quad (18.2)$$

i. e. for all $j = 1, \dots, n$,

$$X_t^j = \xi^j + \int_0^t b_j(s, X_s) ds + \sum_{k=1}^d \int_0^t \sigma_{jk}(s, X_s) dB_s^k, \quad t \geq 0. \quad (18.3)$$

(It is implicit in the definition that all stochastic integrals make sense.)

Throughout this chapter we use the standard Euclidean norms $|b|^2 = \sum_{j=1}^n b_j^2$, $|x|^2 = \sum_{k=1}^d x_k^2$ and $|\sigma|^2 = \sum_{j=1}^n \sum_{k=1}^d \sigma_{jk}^2$ for vectors $b \in \mathbb{R}^n$, $x \in \mathbb{R}^d$ and ma-

trices $\sigma \in \mathbb{R}^{n \times d}$, respectively. An easy application of the Cauchy-Schwarz inequality shows that for these norms $|\sigma x| \leq |\sigma| \cdot |x|$ holds.

The expectation \mathbb{E} and the probability measure \mathbb{P} are given by the driving Brownian motion; in particular $\mathbb{P}(B_0 = 0) = 1$.



18.1 The heuristics of SDEs

We have introduced the SDE (18.1) as a random perturbation of the ODE $\dot{x}_t = b(t, x_t)$ by a *multiplicative* noise: $\sigma(t, X_t) dB_t$. The following heuristic reasoning explains why multiplicative noise is actually a natural choice for many applications. We learned this argument from Michael Röckner.

Denote by x_t the position of a particle at time t . After Δt units of time, the particle is at $x_{t+\Delta t}$. We assume that the motion can be described by an equation of the form

$$\Delta x_t = x_{t+\Delta t} - x_t = f(t, x_t; \Delta t)$$

where f is a function which depends on the initial position x_t and the time interval $[t, t + \Delta t]$. Unknown influences on the movement, e. g. caused by hidden parameters or measurement errors, can be modelled by adding some random noise $\Gamma_{t,\Delta t}$. This leads to

$$X_{t+\Delta t} - X_t = F(t, X_t; \Delta t, \Gamma_{t,\Delta t})$$

where X_t is now a random process. Note that $F(t, X_t; 0, \Gamma_{t,0}) = 0$. In many applications it makes sense to assume that the random variable $\Gamma_{t,\Delta t}$ is independent of X_t , that it has zero mean, and that it depends only on the increment Δt . If we have no further information on the nature of the noise, we should strive for a random variable which is maximally uncertain. One possibility to measure uncertainty is entropy. We know from information theory that, among all probability densities $p(x)$ with a fixed variance σ^2 , the Gaussian density maximizes the entropy $H(p) = -\int_{-\infty}^{\infty} p(x) \log p(x) dx$, cf. Ash [2, Theorem 8.3.3]. Therefore, we assume that $\Gamma_{t,\Delta t} \sim \mathcal{N}(0, \sigma^2(\Delta t))$. By the same argument, the random variables $\Gamma_{t,\Delta t}$ and $\Gamma_{t+\Delta t,\Delta s}$ are independent and $\Gamma_{t,\Delta t+\Delta s} \sim \Gamma_{t,\Delta t} + \Gamma_{t+\Delta t,\Delta s}$. Thus, $\sigma^2(\Delta t + \Delta s) = \sigma^2(\Delta t) + \sigma^2(\Delta s)$, and if $\sigma^2(\cdot)$ is continuous, we see that $\sigma^2(\cdot)$ is a linear function. Without loss of generality we can assume that $\sigma^2(t) = t$; therefore

$$\Gamma_{t,\Delta t} \sim \Delta B_t := B_{t+\Delta t} - B_t$$

where $(B_t)_{t \geq 0}$ is a one-dimensional Brownian motion and we can write

$$X_{t+\Delta t} - X_t = F(t, X_t; \Delta t, \Delta B_t).$$

If F is sufficiently smooth, a Taylor expansion of F in the last two variables (∂_j denotes the partial derivative w. r. t. the j th variable) yields

$$\begin{aligned} X_{t+\Delta t} - X_t &= \partial_4 F(t, X_t; 0, 0) \Delta B_t + \partial_3 F(t, X_t; 0, 0) \Delta t \\ &\quad + \frac{1}{2} \partial_4^2 F(t, X_t; 0, 0) (\Delta B_t)^2 + \frac{1}{2} \partial_3^2 F(t, X_t; 0, 0) (\Delta t)^2 \\ &\quad + \partial_3 \partial_4 F(t, X_t; 0, 0) \Delta t \Delta B_t + R(\Delta t, \Delta B_t). \end{aligned}$$

The remainder term R is given by

$$R(\Delta t, \Delta B_t) = \sum_{\substack{0 \leq j, k \leq 3 \\ j+k=3}} \frac{3}{j!k!} \int_0^1 (1-\theta)^2 \partial_3^j \partial_4^k F(t, X_t; \theta \Delta t, \theta \Delta B_t) d\theta \cdot (\Delta t)^j (\Delta B_t)^k.$$

Since $\Delta B_t \sim \sqrt{\Delta t} B_1$, we get $R = o(\Delta t)$ and $\Delta t \Delta B_t = o((\Delta t)^{3/2})$. Neglecting all terms of order $o(\Delta t)$ and using $(\Delta B_t)^2 \approx \Delta t$, we get for small increments Δt

$$X_{t+\Delta t} - X_t = \underbrace{\left(\partial_3 F(t, X_t; 0, 0) + \frac{1}{2} \partial_4^2 F(t, X_t; 0, 0) \right) \Delta t}_{=: b(t, X_t)} + \underbrace{\partial_4 F(t, X_t; 0, 0) \Delta B_t}_{=: \sigma(t, X_t)}.$$

Letting $\Delta t \rightarrow 0$, we get (18.1).

18.2 Some examples

Before we study general existence and uniqueness results for the SDE (18.1), we will review a few examples where we can find the solution explicitly. Of course, we must make sure that all integrals in (18.2) exist, i. e.

$$\int_0^T |\sigma(s, X_s)|^2 ds \quad \text{and} \quad \int_0^T |b(s, X_s)| ds$$

should have an expected value or be at least almost surely finite, cf. Chapter 15. We will assume this throughout this section. In order to keep things simple we discuss only the one-dimensional situation $d = n = 1$.

Ex. 18.1 18.2 Example. Let $(B_t)_{t \geq 0}$ be a BM¹. We consider the SDE

$$X_t - x = \int_0^t b(s) ds + \int_0^t \sigma(s) dB_s \tag{18.4}$$

with deterministic coefficients $b, \sigma : [0, \infty) \rightarrow \mathbb{R}$. It is clear that X_t is a normal random variable and that $(X_t)_{t \geq 0}$ inherits the independent increments property from

the driving Brownian motion $(B_t)_{t \geq 0}$. Therefore, it is enough to calculate $\mathbb{E} X_t$ and $\mathbb{V} X_t$ in order to characterize the solution:

$$\mathbb{E}(X_t - X_0) = \int_0^t b(s) ds + \mathbb{E} \left[\int_0^t \sigma(s) dB_s \right] \stackrel{14.13 \text{ a)}}{=} \int_0^t b(s) ds$$

and

$$\begin{aligned} \mathbb{V}(X_t - X_0) &= \mathbb{E} \left[\left(X_t - X_0 - \int_0^t b(s) ds \right)^2 \right] \\ &= \mathbb{E} \left[\left(\int_0^t \sigma(s) dB_s \right)^2 \right] \stackrel{(14.20)}{=} \int_0^t \sigma^2(s) ds. \end{aligned}$$

18.3 Example (Homogeneous linear SDEs). A *homogeneous linear SDE* is an SDE of the form

$$dX_t = \beta(t)X_t dt + \delta(t)X_t dB_t$$

with non-random coefficients $\beta, \delta : [0, \infty) \rightarrow \mathbb{R}$. Let us assume that the initial condition is positive: $X_0 > 0$. Since the solution has continuous sample paths, it is clear that $X_t > 0$ for sufficiently small values of $t > 0$. Therefore it is plausible to set $Z_t := \log X_t$. By Itô's formula (16.6) we get, at least formally,

$$\begin{aligned} dZ_t &= \frac{1}{X_t} \beta(t)X_t dt + \frac{1}{X_t} \delta(t)X_t dB_t - \frac{1}{2} \frac{1}{X_t^2} \delta^2(t)X_t^2 dt \\ &= \left(\beta(t) - \frac{1}{2} \delta^2(t) \right) dt + \delta(t) dB_t. \end{aligned}$$

Thus,

$$X_t = X_0 \exp \left[\int_0^t \left(\beta(s) - \frac{1}{2} \delta^2(s) \right) ds \right] \exp \left[\int_0^t \delta(s) dB_s \right].$$

Now we can verify by a direct calculation that X_t is indeed a solution of the SDE – even if $X_0 < 0$.

If $\beta \equiv 0$, we get the formula for the stochastic exponential, see Section 17.1.

18.4 Example (Linear SDEs). A *linear SDE* is an SDE of the form

Ex. 18.3

$$dX_t = (\alpha(t) + \beta(t)X_t) dt + (\gamma(t) + \delta(t)X_t) dB_t.$$

with non-random coefficients $\alpha, \beta, \gamma, \delta : [0, \infty) \rightarrow \mathbb{R}$. Let X_t° be a random process such that $1/X_t^\circ$ is the solution of the homogeneous equation (i. e. where $\alpha = \gamma = 0$) for the initial value $X_0^\circ = 1$. From the explicit formula for $1/X_t^\circ$ of Example 18.3 it is easily seen that

$$dX_t^\circ = (-\beta(t) + \delta^2(t))X_t^\circ dt - \delta(t)X_t^\circ dB_t.$$

Set $Z_t := X_t^\circ X_t$; using Itô's formula (16.8) for the two-dimensional Itô process (X_t, X_t°) , we see

$$dZ_t = X_t dX_t^\circ + X_t^\circ dX_t - \int_0^t (\gamma(s) + \delta(s)X_s)\delta(s)X_s^\circ ds$$

and this becomes, if we insert the formulae for dX_t and dX_t° ,

$$dZ_t = (\alpha(t) - \gamma(t)\delta(t))X_t^\circ dt + \gamma(t)X_t^\circ dB_t.$$

Since we know X_t° from Example 18.3, we can thus get a formula for Z_t and then for $X_t = Z_t/X_t^\circ$.

We will now study some transformations of an SDE which help us to reduce a general SDE (18.1) to a linear SDE or an SDE with non-random coefficients.

18.5 Transformation of an SDE. Assume that $(X_t)_{t \geq 0}$ is a solution of the SDE (18.1) and denote by $f : [0, \infty) \times \mathbb{R} \rightarrow \mathbb{R}$ function such that

- $x \mapsto f(t, x)$ is monotone;
- the derivatives $f_t(t, x)$, $f_x(t, x)$ and $f_{xx}(t, x)$ exist and are continuous;
- $g(t, \cdot) := f^{-1}(t, \cdot)$ exists for each t , i.e. $f(t, g(t, x)) = g(t, f(t, x)) = x$.

Ex. 16.3 We set $Z_t := f(t, X_t)$ and use Itô's formula (16.6) to find the SDE corresponding to the process $(Z_t)_{t \geq 0}$:

$$\begin{aligned} dZ_t &= f_t(t, X_t) dt + f_x(t, X_t) dX_t + \frac{1}{2} f_{xx}(t, X_t) \sigma^2(t, X_t) dt \\ &= \underbrace{\left[f_t(t, X_t) + f_x(t, X_t) b(t, X_t) + \frac{1}{2} f_{xx}(t, X_t) \sigma^2(t, X_t) \right]}_{=: \bar{b}(t, Z_t)} dt \\ &\quad + \underbrace{f_x(t, X_t) \sigma(t, X_t)}_{=: \bar{\sigma}(t, Z_t)} dB_t \end{aligned}$$

where the new coefficients $\bar{b}(t, Z_t)$ and $\bar{\sigma}(t, Z_t)$ are given by

$$\begin{aligned} \bar{b}(t, Z_t) &= f_t(t, X_t) + f_x(t, X_t) b(t, X_t) + \frac{1}{2} f_{xx}(t, X_t) \sigma^2(t, X_t), \\ \bar{\sigma}(t, Z_t) &= f_x(t, X_t) \sigma(t, X_t) \end{aligned}$$

with $X_t = g(t, Z_t)$. We will now try to choose f and g in such a way that the transformed SDE

$$dZ_t = \bar{b}(t, Z_t) dt + \bar{\sigma}(t, Z_t) dB_t \tag{18.5}$$

has an explicit solution from which we obtain the solution $X_t = g(t, Z_t)$ of (18.1). There are two obvious choices:

Variance transform. Assume that $\bar{\sigma}(t, z) \equiv 1$. This means that

$$f_x(t, x)\sigma(t, x) \equiv 1 \iff f_x(t, x) = \frac{1}{\sigma(t, x)} \iff f(t, x) = \int_0^x \frac{dy}{\sigma(t, y)} + c.$$

This can be done if $\sigma(t, x) > 0$ and if the derivatives σ_t and σ_x exist and are continuous.

Drift transform. Assume that $\bar{b}(t, z) \equiv 0$. In this case $f(t, x)$ satisfies the partial differential equation

$$f_t(t, x) + b(t, x)f_x(t, x) + \frac{1}{2}\sigma^2(t, x)f_{xx}(t, x) = 0.$$

If $b(t, x) = b(x)$ and $\sigma(t, x) = \sigma(x)$, we have $b_t = \sigma_t = f_t = 0$, and the partial differential equation becomes an ordinary differential equation

$$b(x)f'(x) + \frac{1}{2}\sigma^2(x)f''(x) = 0$$

which has the following explicit solution

$$f(x) = c_1 + c_2 \int_0^x \exp\left(-\int_0^y \frac{2b(v)}{\sigma^2(v)} dv\right) dy.$$

18.6 Lemma. Let $(B_t)_{t \geq 0}$ be a BM¹. The SDE $dX_t = b(t, X_t)dt + \sigma(t, X_t)dB_t$ can be transformed into the form $dZ_t = \bar{b}(t)dt + \bar{\sigma}(t)dB_t$ if, and only if, the coefficients satisfy the condition

$$0 = \frac{\partial}{\partial x} \left\{ \sigma(t, x) \left(\frac{\sigma_t(t, x)}{\sigma^2(t, x)} - \frac{\partial}{\partial x} \frac{b(t, x)}{\sigma(t, x)} + \frac{1}{2} \sigma_{xx}(t, x) \right) \right\}. \quad (18.6)$$

Proof. If we use the transformation $Z_t = f(t, X_t)$ from Paragraph 18.5, it is necessary that

$$\bar{b}(t) = f_t(t, x) + b(t, x)f_x(t, x) + \frac{1}{2}f_{xx}(t, x)\sigma^2(t, x) \quad (18.7a)$$

$$\bar{\sigma}(t) = f_x(t, x)\sigma(t, x) \quad (18.7b)$$

with $x = g(t, z) = f^{-1}(t, z)$. Now we get from (18.7b)

$$f_x(t, x) = \frac{\bar{\sigma}(t)}{\sigma(t, x)}$$

and therefore

$$f_{xt}(t, x) = \frac{\bar{\sigma}_t(t)\sigma(t, x) - \bar{\sigma}(t)\sigma_t(t, x)}{\sigma^2(t, x)} \quad \text{and} \quad f_{xx}(t, x) = \frac{-\bar{\sigma}(t)\sigma_x(t, x)}{\sigma^2(t, x)}.$$

Differentiate (18.7a) in x ,

$$0 = f_{tx}(t, x) + \frac{\partial}{\partial x} \left(b(t, x) f_x(t, x) + \frac{1}{2} \sigma^2(t, x) f_{xx}(t, x) \right),$$

and plug in the formulae for f_x , f_{tx} and f_{xx} to obtain

$$\frac{\bar{\sigma}_t(t)}{\bar{\sigma}(t)} = \sigma(t, x) \left(\frac{\sigma_t(t, x)}{\sigma^2(t, x)} - \frac{\partial}{\partial x} \frac{b(t, x)}{\sigma(t, x)} + \frac{1}{2} \sigma_{xx}(t, x) \right).$$

If we differentiate this expression in x we finally get (18.6).

Since we can, with suitable integration constants, perform all calculations in reversed order, it turns out that (18.6) is a necessary and sufficient condition to transform the SDE. \square

Ex. 18.5 18.7 Example. Assume that $b(t, x) = b(x)$ and $\sigma(t, x) = \sigma(x)$ and that condition (18.6) holds, i. e.

$$0 = \frac{d}{dx} \left\{ \sigma(x) \left(-\frac{d}{dx} \frac{b(x)}{\sigma(x)} + \frac{1}{2} \sigma''(x) \right) \right\}.$$

Then the transformation

$$Z_t = f(t, X_t) = e^{ct} \int_0^{X_t} \frac{dy}{\sigma(y)}$$

converts $dX_t = b(X_t) dt + \sigma(X_t) dB_t$ into $dZ_t = c' e^{ct} dt + e^{ct} dB_t$ with suitable constants $c, c' \geq 0$; this SDE can be solved as in Example 18.2.

Ex. 18.9 18.8 Lemma. Let $(B_t)_{t \geq 0}$ be a BM¹. The SDE $dX_t = b(X_t) dt + \sigma(X_t) dB_t$ with coefficients $b, \sigma : \mathbb{R} \rightarrow \mathbb{R}$ can be transformed into a linear SDE

$$dZ_t = (\alpha + \beta Z_t) dt + (\gamma + \delta Z_t) dB_t, \quad \alpha, \beta, \gamma, \delta \in \mathbb{R},$$

if, and only if,

$$\frac{d}{dx} \left(\frac{\frac{d}{dx}(\kappa'(x)\sigma(x))}{\kappa'(x)} \right) = 0 \quad \text{where} \quad \kappa(x) = \frac{b(x)}{\sigma(x)} - \frac{1}{2} \sigma'(x). \quad (18.8)$$

Set $d(x) = \int_0^x dy/\sigma(y)$ and $\delta = - \left[\frac{d}{dx}(\kappa'(x)\sigma(x)) \right] / \kappa'(x)$. Then the transformation $Z_t = f(X_t)$ is given by

$$f(x) = \begin{cases} e^{\delta d(x)}, & \text{if } \delta \neq 0, \\ \gamma d(x), & \text{if } \delta = 0. \end{cases}$$

Proof. Since the coefficients do not depend on t , it is clear that the transformation $Z_t = f(X_t)$ should not depend on t . Using the general transformation scheme from

Paragraph 18.5 we see that the following relations are necessary:

$$b(x)f'(x) + \frac{1}{2}\sigma^2(x)f''(x) = \alpha + \beta f(x) \quad (18.9a)$$

$$\sigma(x)f'(x) = \gamma + \delta f(x) \quad (18.9b)$$

where $x = g(z)$, $z = f(x)$, i. e. $g(z) = f^{-1}(z)$.

Case 1. If $\delta \neq 0$, we set $d(x) = \int_0^x dy/\sigma(y)$; then

$$f(x) = ce^{\delta d(x)} - \frac{\gamma}{\delta}$$

is a solution to the ordinary differential equation (18.9b). Inserting this into (18.9a) yields

$$\left[\frac{\delta b(x)}{\sigma(x)} + \frac{1}{2}\sigma^2(x) \left(\frac{\delta^2}{\sigma^2(x)} - \frac{\delta\sigma'(x)}{\sigma^2(x)} \right) \right] ce^{\delta d(x)} = \beta ce^{\delta d(x)} - \frac{\beta\gamma}{\delta} + \alpha$$

which becomes

$$\left[\frac{\delta b(x)}{\sigma(x)} + \frac{1}{2}\delta^2 - \beta - \frac{\delta}{2}\sigma'(x) \right] e^{\delta d(x)} = \frac{\alpha\delta - \gamma\beta}{c\delta}.$$

Set $\kappa(x) := b(x)/\sigma(x) - \frac{1}{2}\sigma'(x)$. Then this relation reads

$$\left[\delta\kappa(x) + \frac{1}{2}\delta^2 - \beta \right] e^{\delta d(x)} = \frac{\alpha\delta - \gamma\beta}{c\delta}.$$

Now we differentiate this equality in order to get rid of the constant c :

$$\delta\kappa'(x)e^{\delta d(x)} + \left[\delta\kappa(x) + \frac{1}{2}\delta^2 - \beta \right] e^{\delta d(x)} \frac{\delta}{\sigma(x)} = 0$$

which is the same as

$$\delta\kappa(x) + \frac{1}{2}\delta^2 - \beta = -\kappa'(x)\sigma(x).$$

Differentiating this equality again yields (18.8); the formulae for δ and $f(x)$ now follow by integration.

Case 2. If $\delta = 0$ then (18.9b) becomes

$$\sigma(x)f'(x) = \gamma \quad \text{and so} \quad f(x) = \gamma \int_0^x \frac{dy}{\sigma(y)} + c.$$

Inserting this into (18.9a) shows

$$\kappa(x)\gamma = \beta\gamma d(x) + c'$$

and if we differentiate this we get $\sigma(x)\kappa'(x) \equiv \beta$, i. e. $\frac{d}{dx}(\sigma(x)\kappa'(x)) = 0$, which implies (18.8).

The sufficiency of (18.8) is now easily verified by a (lengthy) direct calculation with the given transformation. \square

18.3 Existence and uniqueness of solutions

The main question, of course, is whether an SDE has a solution and, if so, whether the solution is unique. If $\sigma \equiv 0$, the SDE becomes an ordinary differential equation, and it is well known that existence and uniqueness results for ordinary differential equations usually require a Lipschitz condition for the coefficients. Therefore it seems reasonable to require

$$\sum_{j=1}^n |b_j(t, x) - b_j(t, y)|^2 + \sum_{j=1}^n \sum_{k=1}^d |\sigma_{jk}(t, x) - \sigma_{jk}(t, y)|^2 \leq L^2 |x - y|^2 \quad (18.10)$$

for all $x, y \in \mathbb{R}^n, t \in [0, T]$ and with a (global) Lipschitz constant $L = L_T$. If we pick $y = 0$, the *global* Lipschitz condition implies the following linear growth condition

$$\sum_{j=1}^n |b_j(t, x)|^2 + \sum_{j=1}^n \sum_{k=1}^d |\sigma_{jk}(t, x)|^2 \leq M^2 (1 + |x|^2) \quad (18.11)$$

Ex. 18.10 for all $x \in \mathbb{R}^n$ and with some constant $M = M_T$ which can be estimated from below by $M^2 \geq 2L^2 + 2 \sum_{j=1}^n \sup_{t \leq T} |b_j(t, 0)|^2 + 2 \sum_{j=1}^n \sum_{k=1}^d \sup_{t \leq T} |\sigma_{jk}(t, 0)|^2$.

We begin with a stability and uniqueness result.

18.9 Theorem (Stability). *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM^d and assume that the coefficients $b : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^n$ and $\sigma : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^{n \times d}$ of the SDE (18.1) satisfy the Lipschitz condition (18.10) for all $x, y \in \mathbb{R}^n$ with the global Lipschitz constant $L = L_T$ for $T > 0$. If $(X_t)_{t \geq 0}$ and $(Y_t)_{t \geq 0}$ are any two solutions of the SDE (18.1) with \mathcal{F}_0 measurable initial conditions $X_0 = \xi \in L^2(\mathbb{P})$ and $Y_0 = \eta \in L^2(\mathbb{P})$, respectively; then*

$$\mathbb{E} \left[\sup_{t \leq T} |X_t - Y_t|^2 \right] \leq 3 e^{3L^2(T+4)T} \mathbb{E} [|\xi - \eta|^2]. \quad (18.12)$$



The expectation \mathbb{E} is given by the Brownian motion started at $B_0 = 0$.

Proof. Note that

$$X_t - Y_t = (\xi - \eta) + \int_0^t (b(s, X_s) - b(s, Y_s)) ds + \int_0^t (\sigma(s, X_s) - \sigma(s, Y_s)) dB_s.$$

The elementary estimate $(a + b + c)^2 \leq 3a^2 + 3b^2 + 3c^2$ yields

$$\begin{aligned} \mathbb{E} \left[\sup_{t \leq T} |X_t - Y_t|^2 \right] &\leq 3 \mathbb{E} [|\xi - \eta|^2] \\ &\quad + 3 \mathbb{E} \left[\sup_{t \leq T} \left| \int_0^t (b(s, X_s) - b(s, Y_s)) ds \right|^2 \right] \\ &\quad + 3 \mathbb{E} \left[\sup_{t \leq T} \left| \int_0^t (\sigma(s, X_s) - \sigma(s, Y_s)) dB_s \right|^2 \right]. \end{aligned}$$

Using the Cauchy-Schwarz inequality and (18.10), we find for the middle term

$$\begin{aligned} &\mathbb{E} \left[\sup_{t \leq T} \left| \int_0^t (b(s, X_s) - b(s, Y_s)) ds \right|^2 \right] \\ &\leq \mathbb{E} \left[\left(\int_0^T |b(s, X_s) - b(s, Y_s)| ds \right)^2 \right] \\ &\leq T \mathbb{E} \left[\int_0^T |b(s, X_s) - b(s, Y_s)|^2 ds \right] \\ &\leq L^2 T \int_0^T \mathbb{E} [|X_s - Y_s|^2] ds. \end{aligned} \tag{18.13}$$

For the third term we use the maximal inequality for stochastic integrals, (14.21) in Theorem 14.13 d), and (18.10) and get

$$\begin{aligned} &\mathbb{E} \left[\sup_{t \leq T} \left| \int_0^t (\sigma(s, X_s) - \sigma(s, Y_s)) dB_s \right|^2 \right] \\ &\leq 4 \int_0^T \mathbb{E} [|\sigma(s, X_s) - \sigma(s, Y_s)|^2] ds \\ &\leq 4L^2 \int_0^T \mathbb{E} [|X_s - Y_s|^2] ds. \end{aligned} \tag{18.14}$$

This proves

$$\begin{aligned} \mathbb{E} \left[\sup_{t \leq T} |X_t - Y_t|^2 \right] &\leq 3 \mathbb{E} [|\xi - \eta|^2] + 3L^2(T + 4) \int_0^T \mathbb{E} [|X_s - Y_s|^2] ds \\ &\leq 3 \mathbb{E} [|\xi - \eta|^2] + 3L^2(T + 4) \int_0^T \mathbb{E} \left[\sup_{r \leq s} |X_r - Y_r|^2 \right] ds. \end{aligned}$$

Now Gronwall's inequality, Theorem A.43, with $u(T) = \mathbb{E}[\sup_{t \leq T} |X_t - Y_t|^2]$, and $a(s) = 3 \mathbb{E}[|\xi - \eta|^2]$ and $b(s) = 3L^2(T + 4)$, yields

$$\mathbb{E} \left[\sup_{t \leq T} |X_t - Y_t|^2 \right] \leq 3 e^{3L^2(T+4)T} \mathbb{E} [|\xi - \eta|^2]. \quad \square$$

18.10 Corollary (Uniqueness). *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM^d and assume that the coefficients $b : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^n$ and $\sigma : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^{n \times d}$ of the SDE (18.1) satisfy*

the Lipschitz condition (18.10) for all $x, y \in \mathbb{R}^n$ with the global Lipschitz constant $L = L_T$ for $T > 0$. Any two solutions $(X_t)_{t \geq 0}$ and $(Y_t)_{t \geq 0}$ of the SDE (18.1) with the same \mathcal{F}_0 measurable initial condition $X_0 = Y_0 = \xi \in L^2(\mathbb{P})$ are indistinguishable.

Proof. The estimate (18.12) shows that $\mathbb{P}(\forall t \in [0, n] : X_t = Y_t) = 1$ for any $n \geq 1$. Therefore, $\mathbb{P}(\forall t \geq 0 : X_t = Y_t) = 1$. \square

In order to show the existence of the solution of the SDE we can use the Picard iteration scheme from the theory of ordinary differential equations.

18.11 Theorem (Existence). *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM^d and assume that the coefficients $b : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^n$ and $\sigma : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^{n \times d}$ of the SDE (18.1) satisfy the Lipschitz condition (18.10) for all $x, y \in \mathbb{R}^n$ with the global Lipschitz constant $L = L_T$ for $T > 0$ and the linear growth condition (18.11) with the constant $M = M_T$ for $T > 0$. For every \mathcal{F}_0 measurable initial condition $X_0 = \xi \in L^2(\mathbb{P})$ there exists a unique solution $(X_t)_{t \geq 0}$; this solution satisfies*

$$\mathbb{E} \left[\sup_{s \leq T} |X_s|^2 \right] \leq \kappa_T \mathbb{E} [(1 + |\xi|)^2] \quad \text{for all } T > 0. \quad (18.15)$$

Proof. The uniqueness of the solution follows from Corollary 18.10.

To prove the existence we use Picard's iteration scheme. Set

$$\begin{aligned} X_0(t) &:= \xi, \\ X_{n+1}(t) &:= \xi + \int_0^t \sigma(s, X_n(s)) dB_s + \int_0^t b(s, X_n(s)) ds. \end{aligned}$$

1° Since $(a + b)^2 \leq 2a^2 + 2b^2$, we see that

$$|X_{n+1}(t) - \xi|^2 \leq 2 \left| \int_0^t b(s, X_n(s)) ds \right|^2 + 2 \left| \int_0^t \sigma(s, X_n(s)) dB_s \right|^2.$$

With the calculations (18.13) and (18.14) – cf. the proof of Theorem 18.9 – where we set $X_s = X_n(s)$ and omit the Y -terms, and with the linear growth condition (18.11) instead of the Lipschitz condition (18.10), we get

$$\begin{aligned} \mathbb{E} \left[\sup_{t \leq T} |X_{n+1}(t) - \xi|^2 \right] &\leq 2M^2(T + 4) \int_0^T \mathbb{E} [(1 + |X_n(s)|)^2] ds \\ &\leq 2M^2(T + 4)T \mathbb{E} \left[\sup_{s \leq T} (1 + |X_n(s)|)^2 \right]. \end{aligned}$$

Starting with $n = 0$, we see recursively that the processes X_{n+1} , $n = 0, 1, \dots$, are well-defined.

2° As in step 1° we see

$$\begin{aligned} |X_{n+1}(t) - X_n(t)|^2 &\leq 2 \left| \int_0^t [b(s, X_n(s)) - b(s, X_{n-1}(s))] ds \right|^2 \\ &\quad + 2 \left| \int_0^t [\sigma(s, X_n(s)) - \sigma(s, X_{n-1}(s))] dB_s \right|^2. \end{aligned}$$

Using again (18.13) and (18.14) with $X = X_n$ and $Y = X_{n-1}$ we get

$$\mathbb{E} \left[\sup_{t \leq T} |X_{n+1}(t) - X_n(t)|^2 \right] \leq 2L^2(T+4) \int_0^T \mathbb{E} [|X_n(s) - X_{n-1}(s)|^2] ds.$$

Set $\phi_n(T) := \mathbb{E} [\sup_{t \leq T} |X_{n+1}(t) - X_n(t)|^2]$ and $c_T = 2L^2(T+4)$. Several iterations of this inequality yield

$$\begin{aligned} \phi_n(T) &\leq c_T^n \int_{t_1 \leq \dots \leq t_{n-1} \leq T} \dots \int \phi_0(t_0) dt_0 \dots dt_{n-1} \\ &\leq c_T^n \int_{t_1 \leq \dots \leq t_{n-1} \leq T} dt_0 \dots dt_{n-1} \phi_0(T) = c_T^n \frac{T^n}{n!} \phi_0(T). \end{aligned}$$

The estimate from 1° for $n = 0$ shows

$$\phi_0(T) = \mathbb{E} \left[\sup_{s \leq T} |X_1(s) - \xi|^2 \right] \leq 2M^2(T+4)T \mathbb{E} [(1 + |\xi|)^2]$$

and this means that

$$\begin{aligned} \sum_{n=0}^{\infty} \{ \mathbb{E} [\sup_{t \leq T} |X_{n+1}(t) - X_n(t)|^2] \}^{1/2} &\leq \sum_{n=0}^{\infty} \left(c_T^n \frac{T^n}{n!} \phi_0(T) \right)^{1/2} \\ &\leq C_T \sqrt{\mathbb{E} [(1 + |\xi|)^2]} \end{aligned}$$

with a constant $C_T = C(L, M, T)$.

3° For $n \geq m$ we find

$$\{ \mathbb{E} [\sup_{s \leq T} |X_n(s) - X_m(s)|^2] \}^{1/2} \leq \sum_{j=m+1}^{\infty} \{ \mathbb{E} [\sup_{s \leq T} |X_j(s) - X_{j-1}(s)|^2] \}^{1/2}.$$

Therefore, $(X_n)_{n \geq 0}$ is a Cauchy sequence in L^2 ; denote its limit by X . The above inequality shows that there is a subsequence $m(k)$ such that

$$\lim_{k \rightarrow \infty} \sup_{s \leq T} |X(s) - X_{m(k)}(s)|^2 = 0 \quad \text{a. s.}$$

Since $X_{m(k)}(\cdot)$ is continuous and adapted, we see that $X(\cdot)$ is also continuous and adapted. Moreover,

$$\begin{aligned} \left\{ \mathbb{E} \left[\sup_{s \leq T} |X(s) - \xi|^2 \right] \right\}^{1/2} &\leq \sum_{n=0}^{\infty} \left\{ \mathbb{E} \left[\sup_{t \leq T} |X_{n+1}(t) - X_n(t)|^2 \right] \right\}^{1/2} \\ &\leq C_T \sqrt{\mathbb{E} [(1 + |\xi|)^2]}, \end{aligned}$$

and (18.15) follows with $\kappa_T = C_T^2 + 1$.

4° Recall that $X_{n+1}(t) = \xi + \int_0^t \sigma(s, X_n(s)) dB_s + \int_0^t b(s, X_n(s)) ds$. Let $n = m(k)$; because of the estimate (18.15) we can use the dominated convergence theorem to get

$$X(t) = \xi + \int_0^t \sigma(s, X(s)) dB_s + \int_0^t b(s, X(s)) ds$$

which shows that X is indeed a solution of the SDE. \square

18.12 Remark. The proof of the existence of a solution of the SDE (18.1) shows, in particular, that the solution depends *measurably* on the initial condition. More precisely, if X_t^x denotes the unique solution of the SDE with initial condition $x \in \mathbb{R}^n$, then $(t, x, \omega) \mapsto X_t^x(\omega)$ is measurable with respect to $\mathcal{B}[0, t] \otimes \mathcal{B}(\mathbb{R}^n) \otimes \mathcal{F}_t$. It is enough to show this for the Picard approximations $X_n^x(t, \omega)$ since the measurability is preserved under pointwise limits. For $n = 0$ we have $X_n^x(t, \omega) = x$ and there is nothing to show. Assume we know already that $X_n^x(t, \omega)$ is $\mathcal{B}[0, t] \otimes \mathcal{B}(\mathbb{R}^n) \otimes \mathcal{F}_t$ measurable. Since

$$X_{n+1}^x(t) = x + \int_0^t b(s, X_n^x(s)) ds + \int_0^t \sigma(s, X_n^x(s)) dB_s$$

we are done, if we can show that both integral terms have the required measurability. For the Lebesgue integral this is obvious from the measurability of the integrand and the continuous dependence on the upper integral limit t . To see the measurability of the stochastic integral, it is enough to assume $d = n = 1$ and that σ is a step function. Then

$$\int_0^t \sigma(s, X_n^x(s)) dB_s = \sum_{j=1}^m \sigma(s_j, X_n^x(s_j)) (B(s_j) - B(s_{j-1}))$$

where $0 = s_0 < s_1 < \dots < s_m = t$ and this shows immediately the required measurability.

18.4 Solutions as Markov processes

Recall, cf. Remark 6.3 a), that a *Markov process* is an n -dimensional, adapted stochastic process $(X_t, \mathcal{F}_t)_{t \geq 0}$ with (right-)continuous paths such that

$$\mathbb{E}[u(X_t) | \mathcal{F}_s] = \mathbb{E}[u(X_t) | X_s] \quad \text{for all } t \geq s \geq 0, u \in \mathcal{B}_b(\mathbb{R}^n). \quad (18.16)$$

18.13 Theorem. *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a d -dimensional Brownian motion and assume that the coefficients $b(s, x) \in \mathbb{R}^n$ and $\sigma(s, x) \in \mathbb{R}^{n \times d}$ of the SDE (18.1) satisfy the global Lipschitz condition (18.10). Then the unique solution of the SDE is a Markov process.*

Proof. Consider for fixed $s \geq 0$ and $x \in \mathbb{R}^n$ the following SDE

$$X_t = x + \int_s^t b(r, X_r) dr + \int_s^t \sigma(r, X_r) dB_r, \quad t \geq s.$$

It is not hard to see that the existence and uniqueness results of Section 18.3 also hold for this SDE and we denote by $(X_t^{s,x})_{t \geq s}$ its unique solution. On the other hand, we find for every \mathcal{F}_0 measurable initial condition $\xi \in L^2(\mathbb{P})$ that

$$X_t^{0,\xi} = \xi + \int_0^t b(r, X_r^{0,\xi}) dr + \int_0^t \sigma(r, X_r^{0,\xi}) dB_r, \quad t \geq 0,$$

and so

$$X_t^{0,\xi} = X_s^{0,\xi} + \int_s^t b(r, X_r^{0,\xi}) dr + \int_s^t \sigma(r, X_r^{0,\xi}) dB_r, \quad t \geq s.$$

Because of the uniqueness of the solution, we get

$$X_t^{0,\xi} = X_t^{s, X_s^{0,\xi}} \quad \text{a. s. for all } t \geq s.$$

Moreover, $X_t^{0,\xi}(\omega)$ is of the form $\Phi(X_s^{0,\xi}(\omega), s, t, \omega)$ and the functional Φ is measurable with respect to $\mathcal{G}_s := \sigma(B_r - B_s : r \geq s)$. Since $(B_t)_{t \geq 0}$ has independent increments, we know that $\mathcal{F}_s \perp\!\!\!\perp \mathcal{G}_s$, cf. Lemma 2.10 or Theorem 6.1. Therefore, we get for all $s \leq t$ and $u \in \mathcal{B}_b(\mathbb{R}^n)$

$$\begin{aligned} \mathbb{E}[u(X_t^{0,\xi}) | \mathcal{F}_s] &= \mathbb{E}[u(\Phi(X_s^{0,\xi}, s, t, \cdot)) | \mathcal{F}_s] \\ &\stackrel{(\text{A.5})}{=} \mathbb{E}[u(\Phi(y, s, t, \cdot))] \Big|_{y=X_s^{0,\xi}} \\ &= \mathbb{E}[u(X_t^{s,y})] \Big|_{y=X_s^{0,\xi}}. \end{aligned}$$

Using the tower property for conditional expectations we see now

$$\mathbb{E}[u(X_t^{0,\xi}) | \mathcal{F}_s] = \mathbb{E}[u(X_t^{0,\xi}) | X_s^{0,\xi}]. \quad \square$$

For $u = \mathbb{1}_B$, $B \in \mathcal{B}(\mathbb{R}^n)$ the proof of Theorem 18.13 actually shows

$$p(s, X_s^{0,\xi}; t, B) := \mathbb{P}(X_t^{0,\xi} \in B \mid \mathcal{F}_s) = \mathbb{P}(X_t^{s,y} \in B) \Big|_{y=X_s^{0,\xi}}.$$

The family of measures $p(s, x; t, \cdot)$ is the *transition function* of the Markov process X . If $p(s, x; t, \cdot)$ depends only on the difference $t - s$, we say that the Markov process is *homogeneous*. In this case we write $p(t - s, x; \cdot)$.

18.14 Corollary. *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM^d and assume that the coefficients $b(x) \in \mathbb{R}^n$ and $\sigma(x) \in \mathbb{R}^{n \times d}$ of the SDE (18.1) are autonomous (i. e. they do not depend on time) and satisfy the global Lipschitz condition (18.10). Then the unique solution of the SDE is a (homogeneous) Markov process with the transition function*

$$p(t, x; B) = \mathbb{P}(X_t^x \in B), \quad t \geq 0, \quad x \in \mathbb{R}^n, \quad B \in \mathcal{B}(\mathbb{R}^n).$$

Proof. As in the proof of Theorem 18.13 we see for all $x \in \mathbb{R}^n$ and $s, t \geq 0$

$$\begin{aligned} X_{s+t}^{s,x} &= x + \int_s^{s+t} b(X_r^{s,x}) dr + \int_s^{s+t} \sigma(X_r^{s,x}) dB_r \\ &= x + \int_0^t b(X_{s+v}^{s,x}) dv + \int_0^t \sigma(X_{s+v}^{s,x}) dW_v \end{aligned}$$

Where $W_v := B_{s+v} - B_s$, $v \geq 0$, is again a Brownian motion. Thus, $(X_{s+t}^{s,x})_{t \geq 0}$ is the unique solution of the SDE

$$Y_t = x + \int_0^t b(Y_v) dv + \int_0^t \sigma(Y_v) dW_v.$$

Since W and B have the same probability distribution, we conclude that

$$\mathbb{P}(X_{s+t}^{s,x} \in B) = \mathbb{P}(X_t^{0,x} \in B) \quad \text{for all } s \geq 0, \quad B \in \mathcal{B}(\mathbb{R}^n).$$

The claim follows if we set $p(t, x; B) := \mathbb{P}(X_t^x \in B) := \mathbb{P}(X_t^{0,x} \in B)$. □

18.5 Localization procedures

We will show that the existence and uniqueness theorem for the solution of an SDE still holds under a *local* Lipschitz condition (18.10) and a *global* linear growth condition (18.11). We begin with a useful localization result for stochastic differential equations.

18.15 Lemma (Localization). *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM^d . Consider the SDEs*

$$dX_t^j = b_j(t, X_t^j) dt + \sigma_j(t, X_t^j) dB_t, \quad j = 1, 2,$$

with initial condition $X_0^1 = X_0^2 = \xi \in L^2$. Assume that ξ is \mathcal{F}_0 measurable and that the coefficients $b_j : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^n$ and $\sigma_j : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^{n \times d}$ satisfy the Lipschitz condition (18.10) for all $x, y \in \mathbb{R}^n$ with a global Lipschitz constant L . If for some $T > 0$ and $R > 0$

$$b_1|_{[0,T] \times \mathbb{B}(0,R)} = b_2|_{[0,T] \times \mathbb{B}(0,R)} \quad \text{and} \quad \sigma_1|_{[0,T] \times \mathbb{B}(0,R)} = \sigma_2|_{[0,T] \times \mathbb{B}(0,R)},$$

then we have for the stopping times $\tau_j = \inf\{t \geq 0 : |X_t^j - \xi| \geq R\} \wedge T$

$$\mathbb{P}(\tau_1 = \tau_2) = 1 \quad \text{and} \quad \mathbb{P}\left(\sup_{s \leq \tau_1} |X_s^1 - X_s^2| = 0\right) = 1. \quad (18.17)$$

Proof. Since the solutions of the SDEs are continuous processes, τ_1 and τ_2 are indeed stopping times, cf. Lemma 5.8. Observe that

$$\begin{aligned} X_{t \wedge \tau_1}^1 - X_{t \wedge \tau_1}^2 &= \int_0^{t \wedge \tau_1} [b_1(s, X_s^1) - b_2(s, X_s^2)] ds + \int_0^{t \wedge \tau_1} [\sigma_1(s, X_s^1) - \sigma_2(s, X_s^2)] dB_s \\ &= \int_0^{t \wedge \tau_1} \underbrace{[b_1(s, X_s^1) - b_2(s, X_s^1)]}_{=0} + b_2(s, X_s^1) - b_2(s, X_s^2) ds \\ &\quad + \int_0^{t \wedge \tau_1} \underbrace{[\sigma_1(s, X_s^1) - \sigma_2(s, X_s^1)]}_{=0} + \sigma_2(s, X_s^1) - \sigma_2(s, X_s^2) dB_s. \end{aligned}$$

Now we can use the elementary inequality $(a + b)^2 \leq 2a^2 + 2b^2$ and (18.13), (18.14) from the proof of the stability Theorem 18.9, with $X = X^1$ and $Y = X^2$ to deduce

$$\mathbb{E} \left[\sup_{t \leq T} |X_{t \wedge \tau_1}^1 - X_{t \wedge \tau_1}^2|^2 \right] \leq 2L^2(T + 4) \int_0^T \mathbb{E} \left[\sup_{r \leq s} |X_{r \wedge \tau_1}^1 - X_{r \wedge \tau_1}^2|^2 \right] ds.$$

Gronwall's inequality, Theorem A.43 with $a(s) = 0$ and $b(s) = 2L^2(T + 4)$ yields

$$\mathbb{E} \left[\sup_{r \leq s} |X_{r \wedge \tau_1}^1 - X_{r \wedge \tau_1}^2|^2 \right] = 0 \quad \text{for all } s \leq T.$$

Therefore, $X_{\bullet \wedge \tau_1}^1 = X_{\bullet \wedge \tau_1}^2$ almost surely and, in particular, $\tau_1 \leq \tau_2$. Swapping the roles of X^1 and X^2 we find $\tau_2 \leq \tau_1$ a. s., and the proof is finished. \square

We will use Lemma 18.15 to prove local uniqueness and existence results.

18.16 Theorem. Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a d -dimensional Brownian motion and assume that the coefficients $b : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^n$ and $\sigma : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^{n \times d}$ of the SDE (18.1) satisfy the linear growth condition (18.11) for all $x, y \in \mathbb{R}^n$ and the Lipschitz

condition (18.10) for all $x, y \in K$ for every compact set $K \subset \mathbb{R}^n$ with the local Lipschitz constants $L_{T,K}$. For every \mathcal{F}_0 measurable initial condition $X_0 = \xi \in L^2(\mathbb{P})$ there exists a unique solution $(X_t)_{t \geq 0}$, and this solution satisfies

$$\mathbb{E} \left[\sup_{s \leq T} |X_s|^2 \right] \leq \kappa_T \mathbb{E} [(1 + |\xi|)^2] \quad \text{for all } T > 0. \quad (18.18)$$

Proof. Fix $T > 0$. For $R > 0$ we find some cut-off function $\chi_R \in \mathcal{C}_c^\infty(\mathbb{R}^n)$ with $\mathbb{1}_{\mathbb{B}(0,R)} \leq \chi_R \leq 1$. Set

$$b_R(t, x) := b(t, x)\chi_R(x) \quad \text{and} \quad \sigma_R(t, x) := \sigma(t, x)\chi_R(x).$$

By Theorem 18.11 the SDE

$$dY_t = b_R(t, Y_t) dt + \sigma_R(t, Y_t) dB_t, \quad Y_0 = \chi_R(\xi)\xi$$

has a unique solution X_t^R . Set

$$\tau_R := \inf \{s \geq 0 : |X_s^R| \geq R\}.$$

For $S > R$ and all $t \geq 0$ we have

$$b_R|_{[0,t] \times \mathbb{B}(0,R)} = b_S|_{[0,t] \times \mathbb{B}(0,R)} \quad \text{and} \quad \sigma_R|_{[0,t] \times \mathbb{B}(0,R)} = \sigma_S|_{[0,t] \times \mathbb{B}(0,R)}.$$

With Lemma 18.15 we conclude that

$$\mathbb{P} \left(\sup_{t \leq T} |X_t^R - X_t^S| > 0 \right) \leq \mathbb{P}(\tau_R < T) = \mathbb{P} \left(\sup_{t \leq T} |X_t^R| \geq R \right).$$

Therefore it is enough to show that

$$\lim_{R \rightarrow \infty} \mathbb{P}(\tau_R < T) = 0. \quad (18.19)$$

Once (18.19) is established, we see that $X_t := \overline{\lim}_{R \rightarrow \infty} X_t^R$ exists uniformly for all $t \in [0, T]$ and $X_{t \wedge \tau_R} = X_{t \wedge \tau_R}^R$. Therefore, $(X_t)_{t \geq 0}$ is a continuous stochastic process. Moreover,

$$\begin{aligned} X_{t \wedge \tau_R}^R &= \chi_R(\xi)\xi + \int_0^{t \wedge \tau_R} b_R(s, X_s^R) ds + \int_0^{t \wedge \tau_R} \sigma_R(s, X_s^R) dB_s \\ &= \chi_R(\xi)\xi + \int_0^{t \wedge \tau_R} b(s, X_s) ds + \int_0^{t \wedge \tau_R} \sigma(s, X_s) dB_s. \end{aligned}$$

Letting $R \rightarrow \infty$ shows that X_t solves the SDE (18.1).

Let us now show (18.19). Since $\xi \in L^2(\mathbb{P})$, we can use Theorem 18.9 to get

$$\mathbb{E} \left[\sup_{t \leq T} |X_t^R|^2 \right] \leq \kappa_T \mathbb{E} [|\chi_R(\xi)\xi|^2] \leq \kappa_T \mathbb{E} [(1 + |\xi|)^2] < \infty,$$

and with Markov's inequality,

$$\mathbb{P}(\tau_R < T) = \mathbb{P}\left(\sup_{s \leq T} |X_s^R| \geq R\right) \leq \frac{\kappa T}{R^2} \mathbb{E}[(1 + |\xi|)^2] \xrightarrow{R \rightarrow \infty} 0.$$

To see uniqueness, assume that X and Y are any two solutions. Since X^R is unique, we conclude that

$$X_{t \wedge \tau_R \wedge T_R} = Y_{t \wedge \tau_R \wedge T_R} = X_{t \wedge \tau_R \wedge T_R}^R$$

if $\tau_R := \inf\{s \geq 0 : |X_s| \geq R\}$ and $T_R := \inf\{s \geq 0 : |Y_s| \geq R\}$. Thus, $X_t = Y_t$ for all $t \leq \tau_R \wedge T_R$. Since $\tau_R, T_R \rightarrow \infty$ as $R \rightarrow \infty$, we see that $X = Y$. \square

18.6 Dependence on the initial values

Often it is important to know how the solution $(X_t^x)_{t \geq 0}$ of the SDE (18.1) depends on the starting point $X_0 = x$. Typically, one is interested whether $x \mapsto X_t^x$ is continuous or differentiable. A first weak continuity result follows already from the stability Theorem 18.9.

18.17 Corollary (Feller continuity). *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a BM^d . Assume that the coefficients $b : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^n$ and $\sigma : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^{n \times d}$ of the SDE (18.1) satisfy the Lipschitz condition (18.10) for all $x, y \in \mathbb{R}^n$ with the global Lipschitz constant L . Denote by $(X_t^x)_{t \geq 0}$ the solution of the SDE (18.1) with initial condition $X_0 \equiv x \in \mathbb{R}^n$. Then*

$$x \mapsto \mathbb{E}[f(X_t^x)] \text{ is continuous for all } f \in \mathcal{C}_b(\mathbb{R}^n).$$

Proof. Assume that $f \in \mathcal{C}_c(\mathbb{R}^n)$ is a continuous function with compact support. Since f is uniformly continuous, we find for any $\epsilon > 0$ some $\delta > 0$ such that $|\xi - \eta| < \delta$ implies $|f(\xi) - f(\eta)| \leq \epsilon$. Now consider for $x, y \in \mathbb{R}^n$

$$\begin{aligned} & \mathbb{E}[|f(X_t^x) - f(X_t^y)|] \\ &= \mathbb{E}[|f(X_t^x) - f(X_t^y)| \mathbf{1}_{\{|X_t^x - X_t^y| \geq \delta\}}] + \mathbb{E}[|f(X_t^x) - f(X_t^y)| \mathbf{1}_{\{|X_t^x - X_t^y| < \delta\}}] \\ &\leq 2\|f\|_\infty \mathbb{E}[\mathbf{1}_{\{|X_t^x - X_t^y| \geq \delta\}}] + \epsilon \\ &\leq \frac{2\|f\|_\infty}{\delta^2} \mathbb{E}[|X_t^x - X_t^y|^2] + \epsilon \xrightarrow[|x-y| \rightarrow 0]{(18.12)} \epsilon \xrightarrow{\epsilon \rightarrow 0} 0. \end{aligned}$$

This shows that $x \mapsto \mathbb{E}[f(X_t^x)]$ is continuous for all $f \in \mathcal{C}_c(\mathbb{R}^n)$. If $f \in \mathcal{C}_b(\mathbb{R}^n)$, we choose a sequence of cut-off functions $\chi_k \in \mathcal{C}_c(\mathbb{R}^n)$ such that $\mathbf{1}_{\mathbb{B}(0,k)} \leq \chi_k \leq 1$.

Then $\chi_k f \in \mathcal{C}_c(\mathbb{R}^n)$ and we see for all $x, y \in \mathbb{R}^n$

$$\begin{aligned}
& \mathbb{E} [|f(X_t^x) - f(X_t^y)|] \\
& \leq \mathbb{E} [|\chi_k f(X_t^x) - \chi_k f(X_t^y)|] + \|f\|_\infty (\mathbb{P}(|X_t^x| \geq k) + \mathbb{P}(|X_t^y| \geq k)) \\
& \leq \mathbb{E} [|\chi_k f(X_t^x) - \chi_k f(X_t^y)|] + \frac{\|f\|_\infty}{k^2} (\mathbb{E}[|X_t^x|^2] + \mathbb{E}[|X_t^y|^2]) \\
& \leq \mathbb{E} [|\chi_k f(X_t^x) - \chi_k f(X_t^y)|] + \kappa_t \frac{\|f\|_\infty}{k^2} (|x|^2 + |y|^2) \\
& \xrightarrow[y \rightarrow x]{} 2\kappa_t \frac{\|f\|_\infty}{k^2} |x|^2 \xrightarrow[k \rightarrow \infty]{} 0. \quad \square
\end{aligned}$$

If we want to show that $x \mapsto X_t^x$ is almost surely continuous, we need the Kolmogorov–Chentsov Theorem 10.1 and Burkholder’s inequalities, Theorem 17.16.

18.18 Theorem. *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a d -dimensional Brownian motion. Assume that the coefficients $b : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^n$ and $\sigma : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^{n \times d}$ of the SDE (18.1) satisfy the Lipschitz condition (18.10) for all $x, y \in \mathbb{R}^n$ with the global Lipschitz constant L . Denote by $(X_t^x)_{t \geq 0}$ the solution of the SDE (18.1) with initial condition $X_0 \equiv x \in \mathbb{R}^n$. Then we have for all $p \geq 2$*

$$\mathbb{E} \left[\sup_{s \leq T} (1 + |X_s^x|)^p \right] \leq c_{p,T} (1 + |x|^p) \quad \text{for all } T > 0, \quad (18.20)$$

$$\mathbb{E} \left[\sup_{s \leq T} |X_s^x - X_s^y|^p \right] \leq c'_{p,T} |x - y|^p \quad \text{for all } T > 0. \quad (18.21)$$

In particular, $x \mapsto X_t^x$ (has a modification which) is continuous.

Proof. The proof of (18.20) is very similar to the proof of (18.21). Note that the global Lipschitz condition (18.10) implies the linear growth condition (18.11) which will be needed for (18.20). Therefore, we will only prove (18.21). Note that

$$X_t^x - X_t^y = (x - y) + \int_0^t (b(s, X_s^x) - b(s, X_s^y)) ds + \int_0^t (\sigma(s, X_s^x) - \sigma(s, X_s^y)) dB_s.$$

From Hölder’s inequality we see that $(a + b + c)^p \leq 3^{p-1}(a^p + b^p + c^p)$. Thus,

$$\begin{aligned}
\mathbb{E} \left[\sup_{t \leq T} |X_t^x - X_t^y|^p \right] & \leq 3^{p-1} |x - y|^p \\
& + 3^{p-1} \mathbb{E} \left[\sup_{t \leq T} \left| \int_0^t (b(s, X_s^x) - b(s, X_s^y)) ds \right|^p \right] \\
& + 3^{p-1} \mathbb{E} \left[\sup_{t \leq T} \left| \int_0^t (\sigma(s, X_s^x) - \sigma(s, X_s^y)) dB_s \right|^p \right].
\end{aligned}$$

Using the Hölder inequality and (18.10), we find for the middle term

$$\begin{aligned}
& \mathbb{E} \left[\sup_{t \leq T} \left| \int_0^t (b(s, X_s^x) - b(s, X_s^y)) ds \right|^p \right] \\
& \leq \mathbb{E} \left[\left(\int_0^T |b(s, X_s^x) - b(s, X_s^y)| ds \right)^p \right] \\
& \leq T^{p-1} \mathbb{E} \left[\int_0^T |b(s, X_s^x) - b(s, X_s^y)|^p ds \right] \\
& \stackrel{(18.10)}{\leq} L^p T^{p-1} \int_0^T \mathbb{E} [|X_s^x - X_s^y|^p] ds \\
& \leq L^p T^{p-1} \int_0^T \mathbb{E} \left[\sup_{r \leq s} |X_r^x - X_r^y|^p \right] ds.
\end{aligned}$$

For the third term we combine Burkholder's inequality (applied to the coordinate processes), the estimate (18.10) and Hölder's inequality to get

$$\begin{aligned}
& \mathbb{E} \left[\sup_{t \leq T} \left| \int_0^t (\sigma(s, X_s^x) - \sigma(s, X_s^y)) dB_s \right|^p \right] \\
& \stackrel{(17.15)}{\leq} C_p \mathbb{E} \left[\left[\int_0^T |\sigma(s, X_s^x) - \sigma(s, X_s^y)|^2 ds \right]^{p/2} \right] \\
& \stackrel{(18.10)}{\leq} C_p L^p \mathbb{E} \left[\left[\int_0^T |X_s^x - X_s^y|^2 ds \right]^{p/2} \right] \\
& \leq C_p L^p \mathbb{E} \left[\left[\int_0^T \sup_{r \leq s} |X_r^x - X_r^y|^2 ds \right]^{p/2} \right] \\
& \leq C_p L^p T^{p/2-1} \int_0^T \mathbb{E} \left[\sup_{r \leq s} |X_r^x - X_r^y|^p \right] ds.
\end{aligned}$$

This shows that

$$\mathbb{E} \left[\sup_{t \leq T} |X_t^x - Y_t^y|^p \right] \leq c |x - y|^p + c \int_0^T \mathbb{E} \left[\sup_{r \leq s} |X_r^x - Y_r^y|^p \right] ds$$

with a constant $c = c(T, p, L)$. Gronwall's inequality, Theorem A.43, yields with $u(T) = \mathbb{E}[\sup_{t \leq T} |X_t^x - X_t^y|^p]$, $a(s) = c |x - y|^p$ and $b(s) = c$,

$$\mathbb{E} \left[\sup_{t \leq T} |X_t^x - X_t^y|^p \right] \leq c e^{cT} |x - y|^p.$$

The assertion follows from the Kolmogorov–Chentsov Theorem 10.1, where we use $\alpha = \beta + n = p$. \square

18.19 Remark.

- a) A weaker form of the boundedness estimate (18.20) from Theorem 18.18 holds for all $p \in \mathbb{R}$:

$$\mathbb{E} \left[(1 + |X_t^x|)^p \right] \leq c_{p,T} (1 + |x|^2)^{p/2} \quad \text{for all } x \in \mathbb{R}^d, t \leq T, p \in \mathbb{R}. \quad (18.22)$$

Let us sketch the argument for $d = n = 1$.

Outline of the proof. Since $(1 + |x|)^p \leq (1 + 2^{p/2})(1 + |x|^2)^{p/2}$ for all $p \in \mathbb{R}$, it is enough to prove (18.22) for $\mathbb{E} \left[(1 + |X_t^x|^2)^{p/2} \right]$. Set $f(x) = (1 + x^2)^{p/2}$. Then $f \in \mathcal{C}^2(\mathbb{R})$ and we find that

$$f'(x) = px(1 + x^2)^{p/2-1} \quad \text{and} \quad f''(x) = p(1 - x^2 + px^2)(1 + x^2)^{p/2-2},$$

i. e. $|(1 + x^2)^{1/2} f'(x)| \leq C_p f(x)$ and $|(1 + x^2) f''(x)| \leq C_p f(x)$ for some absolute constant C_p . Itô's formula (16.1) yields

$$\begin{aligned} f(X_t^x) - f(x) &= \int_0^t f'(X_r^x) b(r, X_r^x) dr + \int_0^t f'(X_r^x) \sigma(r, X_r^x) dB_r \\ &\quad + \frac{1}{2} \int_0^t f''(X_r^x) \sigma^2(r, X_r^x) dr. \end{aligned}$$

The second term on the right is a local martingale, cf. Theorem 15.7 and Remark 16.3. Using a localizing sequence $(\sigma_k)_{k \geq 0}$ and the fact that $b(r, x)$ and $\sigma(r, x)$ satisfy the linear growth condition (18.11), we get for all $k \geq 1$ and $t \leq T$

$$\begin{aligned} \mathbb{E} f(X_{t \wedge \sigma_k}^x) &\leq f(x) + \mathbb{E} \left[\int_0^{t \wedge \sigma_k} \left(M(1 + |X_r^x|) |f'(X_r^x)| + M^2(1 + |X_r^x|)^2 |f''(X_r^x)| \right) dr \right] \\ &\leq f(x) + C(p, M) \int_0^t \mathbb{E} f(X_r^x) dr. \end{aligned}$$

By Fatou's lemma $\mathbb{E} f(X_t^x) \leq \liminf_{k \rightarrow \infty} \mathbb{E} f(X_{t \wedge \sigma_k}^x)$, and we can apply Gronwall's lemma, Theorem A.43. This shows for all x and $t \leq T$

$$\mathbb{E} f(X_t^x) \leq e^{C(p,M)T} f(x). \quad \square$$

- b) For the differentiability of the solution with respect to the initial conditions, we can still use the idea of the proof of Theorem 18.16. We consider the (directional) difference quotients

$$\frac{X_t^x - X_t^{x+he_j}}{h}, \quad h > 0, \quad e_j = \overbrace{(0, \dots, 0, 1, 0, \dots)}^j \in \mathbb{R}^n,$$

and show that, as $h \rightarrow 0$, the partial derivatives $\partial_j X_t^x$, $j = 1, \dots, n$, exist and satisfy the formally differentiated SDE

$$d\partial_j X_t^x = \underbrace{D_z b(t, z)}_{\substack{\in \mathbb{R}^{n \times n} \\ \in \mathbb{R}^n}} \Big|_{z=X_t^x} \underbrace{\partial_j X_t^x}_{\in \mathbb{R}^n} dt + \sum_{\ell=1}^d \underbrace{D_z \sigma_{\bullet, \ell}(z, r)}_{\substack{\in \mathbb{R}^{n \times n} \forall \ell \\ \ell \text{th column, } \in \mathbb{R}^n}} \Big|_{z=X_t^x} \underbrace{dB_t^\ell}_{\substack{\ell \text{th coordinate, } \in \mathbb{R} \\ \in \mathbb{R}^n}} \underbrace{\partial_j X_t^x}_{\in \mathbb{R}^n}.$$

As one would expect, we need now (global) Lipschitz and growth conditions also for the *derivatives* of the coefficients. We will not go into further details but refer to Kunita [111, Chapter 4.6] and [112, Chapter 3].

18.20 Corollary. *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a d -dimensional Brownian motion. Assume that the coefficients $b : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^n$ and $\sigma : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^{n \times d}$ of the SDE (18.1) satisfy the Lipschitz condition (18.10) for all $x, y \in \mathbb{R}^n$ with the global Lipschitz constant L . Denote by $(X_t^x)_{t \geq 0}$ the solution of the SDE (18.1) with initial condition $X_0 \equiv x \in \mathbb{R}^n$. Then we have for all $p \geq 2$, $s, t \in [0, T]$ and $x, y \in \mathbb{R}^n$*

$$\mathbb{E} [|X_s^x - X_t^y|^p] \leq C_{p,T} \left(|x - y|^p + (1 + |x| + |y|)^p |s - t|^{p/2} \right). \quad (18.23)$$

In particular, $(t, x) \mapsto X_t^x$ (has a modification which) is continuous.

Proof. Assume that $s \leq t$. Then the inequality $(a + b)^p \leq 2^{p-1}(a^p + b^p)$ gives

$$\mathbb{E} [|X_s^x - X_t^y|^p] \leq 2^{p-1} \mathbb{E} [|X_s^x - X_s^y|^p] + 2^{p-1} \mathbb{E} [|X_s^y - X_t^y|^p].$$

The first term can be estimated by (18.21) with $c'_{p,T}|x - y|^p$. For the second term we note that

$$|X_s^y - X_t^y|^p = \left| \int_s^t b(r, X_r^y) dr + \int_s^t \sigma(r, X_r^y) dB_r \right|^p.$$

Now we argue as in Theorem 18.18: Because of $(a + b)^p \leq 2^{p-1}(a^p + b^p)$, it is enough to estimate the integrals separately. The linear growth condition (18.11) (which follows from the global Lipschitz condition) and similar arguments as in the proof of Theorem 18.18 yield

$$\begin{aligned} \mathbb{E} \left[\left| \int_s^t b(r, X_r^y) dr \right|^p \right] &\leq |t - s|^{p-1} \mathbb{E} \left[\int_s^t |b(r, X_r^y)|^p dr \right] \\ &\stackrel{(18.11)}{\leq} |t - s|^{p-1} M^p \mathbb{E} \left[\int_s^t (1 + |X_r^y|)^p dr \right] \\ &\leq |t - s|^p M^p \mathbb{E} \left[\sup_{r \leq T} (1 + |X_r^y|)^p \right] \\ &\stackrel{(18.20)}{\leq} |t - s|^p M^p c_{p,T} (1 + |y|)^p. \end{aligned} \quad (18.24)$$

For the second integral we find with Burkholder's inequality (applied to the coordinate processes)

$$\begin{aligned} \mathbb{E} \left[\left| \int_s^t \sigma(r, X_r^y) dB_r \right|^p \right] &\stackrel{(17.15)}{\leq} C_p \mathbb{E} \left[\left(\int_s^t |\sigma(r, X_r^y)|^2 dr \right)^{p/2} \right] \\ &\leq |t-s|^{p/2} C_p M^p c_{p,T} (1+|y|)^p \end{aligned}$$

where the last estimate follows as (18.24) with $|b| \rightsquigarrow |\sigma|^2$ and $p \rightsquigarrow p/2$. Thus,

$$\mathbb{E} [|X_s^x - X_t^y|^p] \leq c_{p,T}'' \left(|x-y|^p + |t-s|^p (1+|y|)^p + |t-s|^{p/2} (1+|y|)^p \right).$$

Since $t-s \leq T$, we have $|t-s|^p \leq T^{p/2} |t-s|^{p/2}$.

If $t \leq s$, the starting points x and y change their roles in the above calculation, and we see that (18.23) is just a symmetric (in x and y) formulation of the last estimate.

The existence of a jointly continuous modification $(t, x) \mapsto X_t^x$ follows from the Kolmogorov–Chentsov Theorem 10.1 for the index set $\mathbb{R}^n \times [0, \infty) \subset \mathbb{R}^{n+1}$ and with $p = \alpha = 2(\beta + n + 1)$. \square

Ex. 18.11 18.21 Corollary. *Let $(B_t, \mathcal{F}_t)_{t \geq 0}$ be a d -dimensional Brownian motion and assume that the coefficients $b : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^n$ and $\sigma : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}^{n \times d}$ of the SDE (18.1) satisfy the Lipschitz condition (18.10) for all $x, y \in \mathbb{R}^n$ with the global Lipschitz constant L . Denote by $(X_t^x)_{t \geq 0}$ the solution of the SDE (18.1) with initial condition $X_0 \equiv x \in \mathbb{R}^n$. Then*

$$\lim_{|x| \rightarrow \infty} |X_t^x(\omega)| = \infty \text{ almost surely for every } t \geq 0. \quad (18.25)$$

Proof. For $x \in \mathbb{R}^n \setminus \{0\}$ we set $\hat{x} := x/|x|^2 \in \mathbb{R}^n$, and define for every $t \geq 0$

$$Y_t^{\hat{x}} = \begin{cases} (1 + |X_t^x|)^{-1} & \text{if } \hat{x} \neq 0, \\ 0 & \text{if } \hat{x} = 0. \end{cases}$$

Observe that $|\hat{x}| \rightarrow 0$ if, and only if, $|x| \rightarrow \infty$. If $\hat{x}, \hat{y} \neq 0$ we find for all $p \geq 2$

$$|Y_t^{\hat{x}} - Y_t^{\hat{y}}|^p = |Y_t^{\hat{x}}|^p \cdot |Y_t^{\hat{y}}|^p \cdot ||X_t^x| - |X_t^y||^p \leq |Y_t^{\hat{x}}|^p \cdot |Y_t^{\hat{y}}|^p \cdot |X_t^x - X_t^y|^p,$$

and applying Hölder's inequality two times we get

$$\mathbb{E} [|Y_t^{\hat{x}} - Y_t^{\hat{y}}|^p] \leq (\mathbb{E} [|Y_t^{\hat{x}}|^{4p}])^{1/4} (\mathbb{E} [|Y_t^{\hat{y}}|^{4p}])^{1/4} (\mathbb{E} [|X_t^x - X_t^y|^{2p}])^{1/2}.$$

From Theorem 18.18 and Remark 18.19 a) we infer

$$\mathbb{E} [|Y_t^{\hat{x}} - Y_t^{\hat{y}}|^p] \leq c_{p,T} (1 + |x|)^{-p} (1 + |y|)^{-p} |x - y|^p \leq c_{p,T} |\hat{x} - \hat{y}|^p.$$

In the last estimate we used the inequality

Ex. 18.12

$$\frac{|x - y|}{(1 + |x|)(1 + |y|)} \leq \left| \frac{x}{|x|^2} - \frac{y}{|y|^2} \right|, \quad x, y \in \mathbb{R}^n.$$

The case where $\hat{x} = 0$ is similar but easier since we use only (18.22) from Remark 18.19 a).

The Kolmogorov–Chentsov Theorem 10.1 shows that $\hat{x} \mapsto Y_t^{\hat{x}}$ is continuous at $\hat{x} = 0$, and the claim follows. \square

18.22 Further reading. The classic book by [72] is still a very good treatise on the L^2 theory of strong solutions. Weak solutions and extensions of existence and uniqueness results are studied in [99] or [80]. For degenerate SDEs you should consult [60]. The dependence on the initial value and more advanced flow properties can be found in [111] for processes with continuous paths and, for jump processes, in [112]. The viewpoint from random dynamical systems is explained in [1].



- [1] Arnold: *Random Dynamical Systems*.
- [60] Engelbert, Cherny: *Singular Stochastic Differential Equations*.
- [99] Karatzas, Shreve: *Brownian Motion and Stochastic Calculus*.
- [80] Ikeda, Watanabe: *Stochastic Differential Equations and Diffusion Processes*.
- [72] Gikhman, Skorokhod: *Stochastic Differential Equations*.
- [111] Kunita: *Stochastic Flows and Stochastic Differential Equations*.
- [112] Kunita: Stochastic differential equations based on Lévy processes and stochastic flows of diffeomorphisms.

Problems

1. Let $X = (X_t)_{t \geq 0}$ be the process from Example 18.2. Show that X is a Gaussian process with independent increments and find the covariance function $C(s, t) = \mathbb{E} X_s X_t$, $s, t \geq 0$.
Hint: Let $0 = t_0 < t_1 < \dots < t_n = t$. Find the characteristic function of the random variables $(X_{t_1} - X_{t_0}, \dots, X_{t_n} - X_{t_{n-1}})$ and $(X_{t_1}, \dots, X_{t_n})$.
2. Let $(B(t))_{t \in [0,1]}$ be a BM^1 and set $t_k = k/2^n$ for $k = 0, 1, \dots, 2^n$ and $\Delta t = 2^{-n}$. Consider the difference equation

$$\Delta X_n(t_k) = -\frac{1}{2} X_n(t_k) \Delta t + \Delta B(t_k), \quad X_n(0) = A,$$

where $\Delta X_n(t_k) = X_n(t_k + \Delta t) - X_n(t_k)$ and $\Delta B(t_k) = B(t_k + \Delta t) - B(t_k)$ for $0 \leq k \leq 2^n - 1$ and $A \sim \text{N}(0, 1)$ is independent of $(B_t)_{t \in [0,1]}$.

- (a) Find an explicit formula for $X_n(t_k)$, determine the distribution of $X_n(1/2)$ and the asymptotic distribution as $n \rightarrow \infty$.
 (b) Letting $n \rightarrow \infty$ turns the difference equation into the SDE

$$dX(t) = -\frac{1}{2} X(t) dt + dB(t), \quad X(0) = A.$$

Solve this SDE and determine the distribution of $X(1/2)$ as well as the covariance function $C(s, t) = \mathbb{E} X_s X_t$, $s, t \in [0, 1]$.

3. Show that in Example 18.4

$$X_t^\circ = \exp\left(-\int_0^t (\beta(s) - \delta^2(s)/2) ds\right) \exp\left(-\int_0^t \delta(s) dB_s\right),$$

and verify the expression given for dX_t° . Moreover, show that

$$X_t = \frac{1}{X_t^\circ} \left(X_0 + \int_0^t (\alpha(s) - \gamma(s)\delta(s)) X_s^\circ ds + \int_0^t \gamma(s) X_s^\circ dB_s \right).$$

4. Let $(B_t)_{t \geq 0}$ be a BM¹. The solution of the SDE

$$dX_t = -\beta X_t dt + \sigma dB_t, \quad X_0 = \xi$$

with $\beta > 0$ and $\sigma \in \mathbb{R}$ is an Ornstein–Uhlenbeck process.

- (a) Find X_t explicitly.
 (b) Determine the distribution of X_t . Is $(X_t)_{t \geq 0}$ a Gaussian process? Find the covariance function $C(s, t) = \mathbb{E} X_s X_t$, $s, t \geq 0$.
 (c) Determine the asymptotic distribution μ of X_t as $t \rightarrow \infty$.
 (d) Assume that the initial condition ξ is a random variable which is independent of the driving Brownian motion $(B_t)_{t \geq 0}$, write μ for its probability distribution. Find the characteristic function of $(X_{t_1}, \dots, X_{t_n})$ for any choice of $0 \leq t_1 < t_2 < \dots < t_n < \infty$.
 (e) Show that the process $(U_t)_{t \geq 0}$, where $U_t := e^{-\beta t} B(e^{2\beta t} - 1)$, has the same finite dimensional distributions as $(X_t)_{t \geq 0}$ with initial condition $X_0 = 0$.
 5. Verify the claim made in Example 18.7 using Itô's formula.
 Derive from the proof of Lemma 18.6 explicitly the form of the transformation and the coefficients in Example 18.7.
Hint: Integrate the condition (18.6) and retrace the steps of the proof of Lemma 18.6 from the end to the beginning.
 6. Let $(B_t)_{t \geq 0}$ be a BM¹. Find an SDE which has $X_t = tB_t$, $t \geq 0$, as unique solution.

7. Let $(B_t)_{t \geq 0}$ be a BM¹. Find for each of the following processes an SDE which is solved by it:

$$(a) \quad U_t = \frac{B_t}{1+t}; \quad (b) \quad V_t = \sin B_t;$$

$$(c) \quad \begin{pmatrix} X_t \\ Y_t \end{pmatrix} = \begin{pmatrix} a \cos B_t \\ b \sin B_t \end{pmatrix}, \quad a, b \neq 0.$$

8. Let $(B_t)_{t \geq 0}$ be a BM¹. Solve the following SDEs

$$(a) \quad dX_t = b \, dt + \sigma X_t \, dB_t, \quad X_0 = x_0 \text{ and } b, \sigma \in \mathbb{R};$$

$$(b) \quad dX_t = (m - X_t) \, dt + \sigma \, dB_t, \quad X_0 = x_0 \text{ and } m, \sigma > 0.$$

9. Let $(B_t)_{t \geq 0}$ be a BM¹. Use Lemma 18.8 to find the solution of the following SDE:

$$dX_t = \left(\sqrt{1 + X_t^2} + \frac{1}{2} X_t \right) dt + \sqrt{1 + X_t^2} \, dB_t.$$

10. Show that the constant M in (18.11) can be chosen in the following way:

$$M^2 \geq 2L^2 + 2 \sum_{j=1}^n \sup_{t \leq T} |b_j(t, 0)|^2 + 2 \sum_{j=1}^n \sum_{k=1}^d \sup_{t \leq T} |\sigma_{jk}(t, 0)|^2$$

11. The linear growth of the coefficients is essential for Corollary 18.21.

- (a) Consider the case where $d = n = 1$, $b(x) = -e^x$ and $\sigma(x) = 0$. Find the solution of this deterministic ODE and compare your findings for $x \rightarrow +\infty$ with Corollary 18.21

- (b) Assume that $|b(x)| + |\sigma(x)| \leq M$ for all x . Find a simpler proof of Corollary 18.21.

Hint: Find an estimate for $\mathbb{P} \left(\left| \int_0^t \sigma(X_s^x) \, dB_s \right| > r \right)$.

- (c) Assume that $|b(x)| + |\sigma(x)|$ grow, as $|x| \rightarrow \infty$, like $|x|^{1-\epsilon}$ for some $\epsilon > 0$. Show that one can still find a (relatively) simple proof of Corollary 18.21.

12. Show that for $x, y \in \mathbb{R}^n$ the inequality

$$|x - y|(1 + |x|)^{-1}(1 + |y|)^{-1} \leq ||x|^{-2}x - |y|^{-2}y|$$

holds.

Hint: Use a direct calculation after squaring both sides.

Chapter 19

On diffusions

Diffusion is a physical phenomenon which describes the tendency of two (or more) substances, e. g. gases or liquids, to reach an equilibrium: Particles of one type, say B , move or ‘diffuse’ in(to) another substance, say S , leading to a more uniform distribution of the type B particles within S . Since the particles are physical objects, it is reasonable to assume that their trajectories are continuous (in fact, differentiable) and that the movement is influenced by various temporal and spatial (in-)homogeneities which depend on the nature of the substance S . Diffusion phenomena are governed by Fick’s law. Denote by $p = p(t, x)$ the concentration of the B particles at a space-time point (t, x) , and by $J = J(t, x)$ the flow of particles. Then

$$J = -D \frac{\partial p}{\partial x} \quad \text{and} \quad \frac{\partial p}{\partial t} = -\frac{\partial J}{\partial x}$$

where D is called the diffusion constant which takes into account the geometry and the properties of the substances B and S . Einstein considered in 1905 the particle movement observed by Brown and – under the assumption that the movement is temporally and spatially homogeneous – he showed that it can be described as a diffusion phenomenon:

$$\frac{\partial p(t, x)}{\partial t} = D \frac{\partial^2 p(t, x)}{\partial x^2} \quad \text{hence} \quad p(t, x) = \frac{1}{\sqrt{4\pi D}} e^{-\frac{x^2}{4Dt}}.$$

The diffusion coefficient $D = \frac{RT}{N} \frac{1}{6\pi kP}$ depends on the absolute temperature T , the universal gas constant R , Avogadro’s number N , the friction coefficient k and the radius of the particles (i. e. atoms) P .

If the diffusion coefficient depends on time and space, $D = D(t, x)$, Fick’s law leads to a differential equation of the type

$$\frac{\partial p(t, x)}{\partial t} = D(t, x) \frac{\partial^2 p(t, x)}{\partial x^2} + \frac{\partial D(t, x)}{\partial x} \frac{\partial p(t, x)}{\partial x}.$$

In a mathematical model of a diffusion we could either use a macroscopic approach and model the particle density $p(t, x)$, or a microscopic point of view modelling the movement of the particles themselves and determine the particle density. Using probability theory we can describe the (random) position and trajectories of the particles

by a stochastic process. In view of the preceding discussion it is reasonable to require that the stochastic process

- has continuous trajectories $t \mapsto X_t(\omega)$;
- has a generator which is a second-order differential operator.

Nevertheless there is no standard definition of a diffusion process. Depending on the model, one will prefer one assumption over the other or add further requirements.

Throughout this chapter we use the following definition.

19.1 Definition. A *diffusion process* is a Feller process $(X_t)_{t \geq 0}$ with values in \mathbb{R}^d , continuous trajectories and generator $(A, \mathfrak{D}(A))$ such that $\mathcal{C}_c^\infty(\mathbb{R}^d) \subset \mathfrak{D}(A)$ and for $u \in \mathcal{C}_c^\infty(\mathbb{R}^d)$ Ex. 19.1

$$Au(x) = L(x, D)u(x) = \frac{1}{2} \sum_{i,j=1}^d a_{ij}(x) \frac{\partial^2 u(x)}{\partial x_i \partial x_j} + \sum_{j=1}^d b_j(x) \frac{\partial u(x)}{\partial x_j}. \quad (19.1)$$

The symmetric, positive semidefinite matrix $a(x) = (a_{ij}(x))_{i,j=1}^d \in \mathbb{R}^{d \times d}$ is called the *diffusion matrix*, and $b(x) = (b_1(x), \dots, b_d(x)) \in \mathbb{R}^d$ is called the *drift vector*.

Since $(X_t)_{t \geq 0}$ is a Feller process, A maps $\mathcal{C}_c^\infty(\mathbb{R}^d)$ into $\mathcal{C}_\infty(\mathbb{R}^d)$, therefore $a(\cdot)$ and $b(\cdot)$ are continuous functions.

19.2 Remark. Continuity of paths and local operators. Let $(X_t)_{t \geq 0}$ be a Feller process with generator $(A, \mathfrak{D}(A))$ such that $\mathcal{C}_c^\infty(\mathbb{R}^d) \subset \mathfrak{D}(A)$.

By Theorems 7.29 and 7.28, the continuity of the trajectories implies that A is a local operator, hence a second-order differential operator.

The converse is true if $(X_t)_{t \geq 0}$ takes values in a compact space. For this, we need the Dynkin–Kinney criterion for the continuity of the sample paths. A proof can be found in Dynkin [53, Chapter 6 § 5] or Wentzell [184, Chapter 9]. By $\mathbb{R}^d \cup \{\partial\}$ we denote the one-point compactification of \mathbb{R}^d .

19.3 Theorem (Dynkin 1952, Kinney 1953). *Let $(X_t)_{t \geq 0}$ be a strong Markov process with state space $E \subset \mathbb{R}^d \cup \{\partial\}$. If*

$$\forall \delta > 0 \quad \lim_{h \rightarrow 0} \frac{1}{h} \sup_{t \leq h} \sup_{x \in E} \mathbb{P}^x(|X_t - x| > \delta) = 0, \quad (19.2)$$

then there is a modification of $(X_t)_{t \geq 0}$ which has continuous trajectories in E .

The continuity of the trajectories implies that the generator is a local operator, cf. Theorem 7.29. If the state space E is compact, (an obvious modification of) Theorem 7.30

shows that this is the same as to say that

$$\forall \delta > 0 \quad \lim_{t \rightarrow 0} \frac{1}{t} \sup_{x \in E} \mathbb{P}^x(|X_t - x| > \delta) = 0. \quad (19.3)$$

Since $h^{-1} P(h) \leq h^{-1} \sup_{t \leq h} P(t) \leq \sup_{t \leq h} t^{-1} P(t)$ we see that (19.2) and (19.3) are equivalent conditions.

We might apply this to $(X_t)_{t \geq 0}$ with paths in $E = \mathbb{R}^d \cup \{\partial\}$, but the trouble is that continuity of $t \mapsto X_t$ in \mathbb{R}^d and $\mathbb{R}^d \cup \{\partial\}$ are quite different notions: The latter allows *explosion in finite time* as well as *sample paths coming back from ∂* , while continuity in \mathbb{R}^d means infinite life-time: $\mathbb{P}^x(\forall t \geq 0 : X_t \in \mathbb{R}^d) = 1$.

19.1 Kolmogorov's theory

Kolmogorov's seminal paper [105] *Über die analytischen Methoden der Wahrscheinlichkeitsrechnung* marks the beginning of the mathematically rigorous theory of stochastic processes in continuous time. In §§13, 14 of this paper Kolmogorov develops the foundations of the theory of diffusion processes and, we will now follow Kolmogorov's ideas.

19.4 Theorem (Kolmogorov 1933). *Let $(X_t)_{t \geq 0}$, $X_t = (X_t^1, \dots, X_t^d)$, be a Feller process with values in \mathbb{R}^d such that for all $\delta > 0$ and $i, j = 1, \dots, d$*

$$\lim_{t \rightarrow 0} \frac{1}{t} \sup_{x \in \mathbb{R}^d} \mathbb{P}^x(|X_t - x| > \delta) = 0, \quad (19.4a)$$

$$\lim_{t \rightarrow 0} \frac{1}{t} \mathbb{E}^x((X_t^j - x_j) \mathbb{1}_{\{|X_t - x| \leq \delta\}}) = b_j(x), \quad (19.4b)$$

$$\lim_{t \rightarrow 0} \frac{1}{t} \mathbb{E}^x((X_t^i - x_i)(X_t^j - x_j) \mathbb{1}_{\{|X_t - x| \leq \delta\}}) = a_{ij}(x). \quad (19.4c)$$

Then $(X_t)_{t \geq 0}$ is a diffusion process in the sense of Definition 19.1.

Proof. The Dynkin–Kinney criterion, see the discussion in Remark 19.2, shows that (19.4a) guarantees the continuity of the trajectories.

We are going to show that $\mathcal{C}_c^\infty(\mathbb{R}^d) \subset \mathcal{D}(A)$ and that $A|_{\mathcal{C}_c^\infty(\mathbb{R}^d)}$ is a differential operator of the form (19.1). The proof is similar to Example 7.9.

Let $u \in \mathcal{C}_c^\infty(\mathbb{R}^d)$. Since $\partial_j \partial_k u$ is uniformly continuous,

$$\forall \epsilon > 0 \quad \exists \delta > 0 : \max_{1 \leq i, j \leq d} \sup_{|x-y| \leq \delta} |\partial_i \partial_j u(x) - \partial_i \partial_j u(y)| \leq \epsilon.$$

Fix $\epsilon > 0$, choose $\delta > 0$ as above and set $T_t u(x) := \mathbb{E}^x u(X_t)$, $\Omega_\delta := \{|X_t - x| \leq \delta\}$. Then

$$\frac{T_t u(x) - u(x)}{t} = \underbrace{\frac{1}{t} \mathbb{E}^x ([u(X_t) - u(x)] \mathbb{1}_{\Omega_\delta})}_{=J} + \underbrace{\frac{1}{t} \mathbb{E}^x ([u(X_t) - u(x)] \mathbb{1}_{\Omega_\delta^c})}_{=J'}.$$

It is enough to show that $\lim_{t \rightarrow 0} (T_t u(x) - u(x))/t = Au(x)$ holds for all $x \in \mathbb{R}^d$, $u \in \mathcal{C}_c^\infty(\mathbb{R}^d)$ and A from (19.1), see Theorem 7.19.

The second term can be estimated by

$$|J'| \leq 2 \|u\|_\infty \frac{1}{t} \mathbb{P}^x(|X_t - x| > \delta) \xrightarrow[t \rightarrow 0]{(19.4a)} 0.$$

Now we consider J . By Taylor's formula there is some $\theta \in (0, 1)$ such that for the intermediate point $\Theta = (1 - \theta)x + \theta X_t$

$$\begin{aligned} u(X_t) - u(x) &= \sum_{j=1}^d (X_t^j - x_j) \partial_j u(x) + \frac{1}{2} \sum_{j,k=1}^d (X_t^j - x_j)(X_t^k - x_k) \partial_j \partial_k u(x) \\ &\quad + \frac{1}{2} \sum_{j,k=1}^d (X_t^j - x_j)(X_t^k - x_k) (\partial_j \partial_k u(\Theta) - \partial_j \partial_k u(x)). \end{aligned}$$

Denote the terms on the right-hand side by J_1, J_2 and J_3 , respectively. Since $\partial_j \partial_k u$ is uniformly continuous, the elementary inequality $2|ab| \leq a^2 + b^2$ yields

$$\begin{aligned} \frac{1}{t} \mathbb{E}^x (|J_3| \mathbb{1}_{\Omega_\delta}) &\leq \frac{\epsilon}{2} \sum_{i,j=1}^d \frac{1}{t} \mathbb{E}^x [|(X_t^i - x_i)(X_t^j - x_j)| \mathbb{1}_{\Omega_\delta}] \\ &\leq \frac{\epsilon d}{2} \sum_{j=1}^d \frac{1}{t} \mathbb{E}^x [(X_t^j - x_j)^2 \mathbb{1}_{\Omega_\delta}] \xrightarrow[t \rightarrow 0]{(19.4c)} \frac{\epsilon d}{2} \sum_{j=1}^d a_{jj}(x). \end{aligned}$$

Moreover,

$$\begin{aligned} \frac{1}{t} \mathbb{E}^x (J_2 \mathbb{1}_{\Omega_\delta}) &= \frac{1}{2} \sum_{j,k=1}^d \frac{1}{t} \mathbb{E}^x [(X_t^j - x_j)(X_t^k - x_k) \mathbb{1}_{\Omega_\delta}] \partial_j \partial_k u(x) \\ &\xrightarrow[t \rightarrow 0]{(19.4c)} \frac{1}{2} \sum_{i,j=1}^d a_{ij}(x) \partial_i \partial_j u(x), \end{aligned}$$

and

$$\frac{1}{t} \mathbb{E}^x (J_1 \mathbb{1}_{\Omega_\delta}) = \sum_{j=1}^d \frac{1}{t} \mathbb{E}^x [(X_t^j - x_j) \mathbb{1}_{\Omega_\delta}] \partial_j u(x) \xrightarrow[t \rightarrow 0]{(19.4b)} \sum_{j=1}^d b_j(x) \partial_j u(x).$$

Since $\epsilon > 0$ is arbitrary, the claim follows. \square

If $(X_t)_{t \geq 0}$ has a transition density, i. e. if $\mathbb{P}^x(X_t \in A) = \int_A p(t, x, y) dy$ for all $t > 0$ and $A \in \mathcal{B}(\mathbb{R}^d)$, then we can express the dynamics of the movement as a Cauchy problem. Depending on the point of view, we get Kolmogorov's *first* and *second* differential equation, cf. [105, §§ 12–14].

19.5 Proposition (Backward equation. Kolmogorov 1931). *Let $(X_t)_{t \geq 0}$ be a diffusion with generator $(A, \mathfrak{D}(A))$ such that $A|_{\mathcal{C}_c^\infty(\mathbb{R}^d)}$ is given by (19.1) with bounded drift and diffusion coefficients $b \in \mathcal{C}_b(\mathbb{R}^d, \mathbb{R}^d)$ and $a \in \mathcal{C}_b(\mathbb{R}^d, \mathbb{R}^{d \times d})$. Assume that $(X_t)_{t \geq 0}$ has a transition density $p(t, x, y)$, $t > 0$, $x, y \in \mathbb{R}^d$, satisfying*

$$\begin{aligned} p, \frac{\partial p}{\partial t}, \frac{\partial p}{\partial x_j}, \frac{\partial^2 p}{\partial x_j \partial x_k} &\in \mathcal{C}((0, \infty) \times \mathbb{R}^d \times \mathbb{R}^d), \\ p(t, \cdot, \cdot), \frac{\partial p(t, \cdot, \cdot)}{\partial t}, \frac{\partial p(t, \cdot, \cdot)}{\partial x_j}, \frac{\partial^2 p(t, \cdot, \cdot)}{\partial x_j \partial x_k} &\in \mathcal{C}_\infty(\mathbb{R}^d \times \mathbb{R}^d) \end{aligned}$$

for all $t > 0$. Then Kolmogorov's first or backward equation holds:

$$\frac{\partial p(t, x, y)}{\partial t} = L(x, D_x)p(t, x, y), \quad \text{for all } t > 0, x, y \in \mathbb{R}^d. \quad (19.5)$$

Proof. Denote by $L(x, D_x)$ the differential operator given by (19.1). By assumption, $A = L(x, D)$ on $\mathcal{C}_c^\infty(\mathbb{R}^d)$. Since $p(t, x, y)$ is the transition density, we know that the transition semigroup is given by $T_t u(x) = \int p(t, x, y)u(y) dy$.

Ex. 19.2 Using dominated convergence it is not hard to see that the conditions on $p(t, x, y)$ are such that T_t maps $\mathcal{C}_c^\infty(\mathbb{R}^d)$ to $\mathcal{C}_\infty^2(\mathbb{R}^d)$.

Since the coefficients b and a are bounded, $\|Au\|_\infty \leq \kappa \|u\|_{(2)}$, where we use the norm $\|u\|_{(2)} = \|u\|_\infty + \sum_j \|\partial_j u\|_\infty + \sum_{j,k} \|\partial_j \partial_k u\|_\infty$ in $\mathcal{C}_\infty^2(\mathbb{R}^d)$, and $\kappa = \kappa(b, a)$ is a constant. Since $\mathcal{C}_c^\infty(\mathbb{R}^d) \subset \mathfrak{D}(A)$ and $(A, \mathfrak{D}(A))$ is closed, cf. Corollary 7.11, it follows that $\mathcal{C}_\infty^2(\mathbb{R}^d) \subset \mathfrak{D}(A)$ and $A|_{\mathcal{C}_\infty^2(\mathbb{R}^d)} = L(x, D)|_{\mathcal{C}_\infty^2(\mathbb{R}^d)}$.

Ex. 19.3 Therefore, Lemma 7.10 shows that $\frac{d}{dt} T_t u = A T_t u = L(\cdot, D) T_t u$, i. e.

$$\begin{aligned} \frac{\partial}{\partial t} \int p(t, x, y)u(y) dy &= L(x, D_x) \int p(t, x, y)u(y) dy \\ &= \int L(x, D_x)p(t, x, y)u(y) dy \end{aligned}$$

for all $u \in \mathcal{C}_c^\infty(\mathbb{R}^d)$. The last equality is a consequence of the continuity and smoothness assumptions on $p(t, x, y)$ which allow us to use the differentiation lemma for parameter-dependent integrals, see e. g. [169, Theorem 11.5]. Applying it to the t -derivative, we get

$$\int \left(\frac{\partial p(t, x, y)}{\partial t} - L(x, D_x)p(t, x, y) \right) u(y) dy = 0 \quad \text{for all } u \in \mathcal{C}_c^\infty(\mathbb{R}^d)$$

which implies (19.5). □

If $L(x, D_x)$ is a second-order differential operator of the form (19.1), then the formal adjoint operator is

Ex. 19.4

$$L^*(y, D_y)u(y) = \frac{1}{2} \sum_{i,j=1}^d \frac{\partial^2}{\partial y_i \partial y_j} (a_{ij}(y)u(y)) - \sum_{j=1}^d \frac{\partial}{\partial y_j} (b_j(y)u(y))$$

provided that the coefficients a and b are sufficiently smooth.

19.6 Proposition (Forward equation. Kolmogorov 1931). *Let $(X_t)_{t \geq 0}$ denote a diffusion with generator $(A, \mathfrak{D}(A))$ such that $A|_{\mathcal{C}_c^\infty(\mathbb{R}^d)}$ is given by (19.1) with drift and diffusion coefficients $b \in \mathcal{C}^1(\mathbb{R}^d, \mathbb{R}^d)$ and $a \in \mathcal{C}^2(\mathbb{R}^d, \mathbb{R}^{d \times d})$. Assume that X_t has a probability density $p(t, x, y)$ satisfying*

Ex. 19.5

$$p, \frac{\partial p}{\partial t}, \frac{\partial p}{\partial y_j}, \frac{\partial^2 p}{\partial y_j \partial y_k} \in \mathcal{C}((0, \infty) \times \mathbb{R}^d \times \mathbb{R}^d).$$

Then Kolmogorov's second or forward equation holds:

$$\frac{\partial p(t, x, y)}{\partial t} = L^*(y, D_y)p(t, x, y), \quad \text{for all } t > 0, x, y \in \mathbb{R}^d. \quad (19.6)$$

Proof. This proof is similar to the proof of Proposition 19.5, where we use Lemma 7.10 in the form $\frac{d}{dt}T_t u = T_t A u = T_t L(\cdot, D)u$ for $u \in \mathcal{C}_c^\infty(\mathbb{R}^d)$, and then integration by parts in the resulting integral equality. The fact that we do not have to interchange the integral and the operator $L(x, D)$ as in the proof of Proposition 19.5, allows us to relax the boundedness assumptions on p , b and a , but we pay a price in the form of differentiability assumptions for the drift and diffusion coefficients. \square

If a and b do not depend on x , the transition probability $p(t, x, y)$ depends only on the difference $x - y$, and the corresponding stochastic process is spatially homogeneous, essentially a Brownian motion with a drift: $aB_t + bt$. Kolmogorov [105, §16] mentions that in this case the forward equation (19.6) was discovered (but not rigorously proved) by Bachelier [4, 5]. Often it is called *Fokker–Planck equation* since it also appears in the context of statistical physics, cf. Fokker [64] and Planck [147].

In order to understand the names ‘backward’ and ‘forward’ we have to introduce the concept of a fundamental solution of a parabolic PDE.

19.7 Definition. Let $L(s, x, D) = \sum_{i,j=1}^d a_{ij}(s, x) \partial_j \partial_k + \sum_{j=1}^d b_j(s, x) \partial_j + c(s, x)$ be a second-order differential operator on \mathbb{R}^d where $(a_{ij}(s, x)) \in \mathbb{R}^{d \times d}$ is symmetric and positive semidefinite. A *fundamental solution* of the parabolic differential operator $\frac{\partial}{\partial s} + L(s, x, D)$ on the set $[0, T] \times \mathbb{R}^d$ is a function $\Gamma(s, x; t, y)$ defined for all

$(s, x), (t, y) \in [0, T] \times \mathbb{R}^d$ with $s < t$ such that for every $f \in \mathcal{C}_c(\mathbb{R}^d)$ the function

$$u(s, x) = \int_{\mathbb{R}^d} \Gamma(s, x; t, y) f(y) dy$$

satisfies

$$-\frac{\partial u(s, x)}{\partial s} = L(s, x, D)u(s, x) \quad \text{for all } x \in \mathbb{R}^d, s < t \leq T, \quad (19.7a)$$

$$\lim_{s \uparrow t} u(s, x) = f(x) \quad \text{for all } x \in \mathbb{R}^d. \quad (19.7b)$$

Proposition 19.5 shows that $\Gamma(s, x; t, y) := p(t-s, x, y)$, $s < t$, is the fundamental solution of the parabolic differential operator $L(x, D) + \frac{\partial}{\partial s}$, whereas Proposition 19.6 means that $\Gamma(t, y; s, x) := p(s-t, y, x)$, $t < s$ is the fundamental solution of the (formal) adjoint operator $L^*(y, D) - \frac{\partial}{\partial t}$.

The backward equation has the time derivative at the left end of the time interval, and we solve the PDE in a forward direction with a condition at the right end $\lim_{s \uparrow t} p(t-s, x, y) = \delta_x(dy)$. In the forward equation the time derivative takes place at the right end of the time interval, and we solve the PDE backwards with an initial condition $\lim_{s \downarrow t} p(s-t, y, x) = \delta_y(dx)$. Of course, this does not really show in the time-homogeneous case. ... At a technical level, the proofs of Propositions 19.5 and 19.6 reveal that all depends on the identity (7.10): in the backward case we use $\frac{d}{dt}T_t = AT_t$, in the forward case $\frac{d}{dt}T_t = T_tA$.

Using probability theory we can give the following stochastic interpretation. Run the stochastic process in the time interval $[0, t]$. Using the Markov property we can perturb the movement at the beginning, i. e. on $[0, h]$ and then let it run from time h to t , or we could run it up to time $t-h$, and then perturb it at the end in $[t-h, t]$.

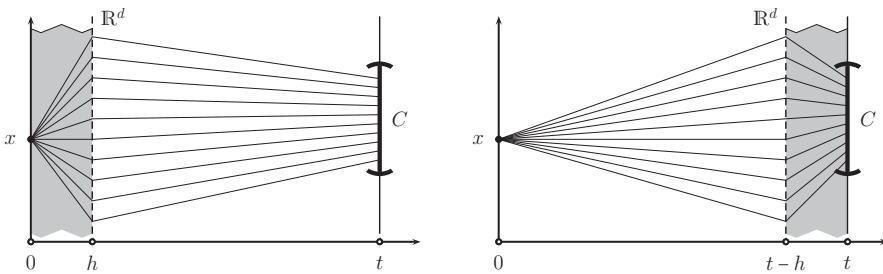


Figure 19.1. Derivation of the backward and forward Kolmogorov equations.

In the first or backward case, the Markov property gives for all $C \in \mathcal{B}(\mathbb{R}^d)$

$$\begin{aligned} \mathbb{P}^x(X_t \in C) - \mathbb{P}^x(X_{t-h} \in C) &= \mathbb{E}^x \mathbb{P}^{X_h}(X_{t-h} \in C) - \mathbb{P}^x(X_{t-h} \in C) \\ &= (T_h - \text{id}) \mathbb{P}^{(\cdot)}(X_{t-h} \in C)(x). \end{aligned}$$

If $\mathbb{P}^x(X_t \in C) = \int_C p(t, x, y) dy$ with a sufficiently smooth density, we get

$$\frac{p(t, x, y) - p(t - h, x, y)}{h} = \frac{T_h - \text{id}}{h} p(t - h, \cdot, y)(x)$$

which, as $h \rightarrow 0$, yields $\frac{\partial}{\partial t} p(t, x, y) = A_x p(t, x, y)$, i. e. the backward equation.

In the second or forward case we get

$$\begin{aligned} \mathbb{P}^x(X_t \in C) - \mathbb{P}^x(X_{t-h} \in C) &= \mathbb{E}^x \mathbb{P}^{X_{t-h}}(X_h \in C) - \mathbb{P}^x(X_{t-h} \in C) \\ &= \mathbb{E}^x ((T_h - \text{id}) \mathbb{1}_C(X_{t-h})). \end{aligned}$$

Passing to the transition densities and dividing by h , we get

$$\frac{p(t, x, y) - p(t - h, x, y)}{h} = \frac{T_h^* - \text{id}}{h} p(t - h, x, \cdot)(y)$$

and $h \rightarrow 0$ yields $\frac{\partial}{\partial t} p(t, x, y) = A_y^* p(t, x, y)$ (T_t^* , A^* are the adjoint operators).

The obvious question is, of course, which conditions on b and a ensure that there is a (unique) diffusion process such that the generator, restricted to $\mathcal{C}_c^\infty(\mathbb{R}^d)$, is a second order differential operator (19.1). If we can show that $L(x, D)$ generates a Feller semi-group or has a nice fundamental solution for the corresponding forward or backward equation for $L(x, D)$, then we could use Kolmogorov's construction, cf. Section 4.2 or Remark 7.6, to construct the corresponding process.

For example, we could use the Hille–Yosida theorem (Theorem 7.16 and Lemma 7.18); the trouble is then to verify condition c) from Theorem 7.16 which amounts to finding a solution to an elliptic partial differential equation of the type $\lambda u(x) - L(x, D)u(x) = f(x)$ for $f \in \mathcal{C}_\infty(\mathbb{R}^d)$. In a non-Hilbert space setting, this is not an easy task.

Another possibility would be to construct a fundamental solution for the backward equation. Here is a standard result in this direction. For a proof we refer to Friedman [68, Chapter I.6].

19.8 Theorem. *Let $L(x, D) = \sum_{i,j=1}^d a_{ij}(x) \partial_i \partial_j + \sum_{j=1}^d b_j(x) \partial_j + c(x)$ be a differential operator on \mathbb{R}^d such that*

$$\sum_{i,j=1}^d a_{ij}(x) \xi_i \xi_j \geq \kappa |\xi|^2 \text{ for some constant } \kappa \text{ and all } x, \xi \in \mathbb{R}^d,$$

$a_{ij}(\cdot), b_j(\cdot), c(\cdot) \in \mathcal{C}_b(\mathbb{R}^d)$ are locally Hölder continuous.

Then there exists a fundamental solution $\Gamma(s, x; t, y)$ such that

- $\Gamma, \frac{\partial \Gamma}{\partial s}, \frac{\partial \Gamma}{\partial x_j}$ and $\frac{\partial^2 \Gamma}{\partial x_j \partial x_k}$ are continuous in $(s, x; t, y)$;

- $\Gamma(s, x; t, y) \leq c(t-s)^{-d/2} e^{-C|x-y|^2/(t-s)},$
 $\frac{\partial \Gamma(s, x; t, y)}{\partial x_j} \leq c(t-s)^{-(d+1)/2} e^{-C|x-y|^2/(t-s)};$
- $-\frac{\partial \Gamma(s, x; t, y)}{\partial s} = L(x, D_x)\Gamma(s, x; t, y);$
- $\lim_{s \uparrow t} \int \Gamma(s, x; t, y) f(x) dx = f(y)$ for all $f \in \mathcal{C}_b(\mathbb{R}^d).$

19.2 Itô's theory

It is possible to extend Theorem 19.8 to increasing and degenerate coefficients, and we refer the interested reader to the exposition in Stroock and Varadhan [177, Chapter 3]. Here we will follow Itô's idea to use stochastic differential equations (SDEs) and construct a diffusion directly. In Chapter 18 we have discussed the uniqueness, existence and regularity of SDEs of the form

$$dX_t = b(X_t) dt + \sigma(X_t) dB_t, \quad X_0 = x, \quad (19.8)$$

for a $\text{BM}^d (B_t)_{t \geq 0}$ and coefficients $b : \mathbb{R}^d \rightarrow \mathbb{R}^d$, $\sigma : \mathbb{R}^d \rightarrow \mathbb{R}^{d \times d}$. The key ingredient was the following global Lipschitz condition

$$\sum_{j=1}^d |b_j(x) - b_j(y)|^2 + \sum_{j,k=1}^d |\sigma_{jk}(x) - \sigma_{jk}(y)|^2 \leq C^2 |x - y|^2 \quad \text{for all } x, y \in \mathbb{R}^d. \quad (19.9)$$

If we take $y = 0$, (19.9) entails the linear growth condition.

19.9 Theorem. *Let $(B_t)_{t \geq 0}$ be a BM^d and $(X_t^x)_{t \geq 0}$ be the solution of the SDE (19.8). Assume that the coefficients b and σ satisfy the global Lipschitz condition (19.9). Then $(X_t^x)_{t \geq 0}$ is a diffusion process in the sense of Definition 19.1 with infinitesimal generator*

$$L(x, D)u(x) = \frac{1}{2} \sum_{i,j=1}^d a_{ij}(x) \frac{\partial^2 u(x)}{\partial x_i \partial x_j} + \sum_{j=1}^d b_j(x) \frac{\partial u(x)}{\partial x_j}$$

where $a(x) = \sigma(x)\sigma^\top(x)$.

Proof. The uniqueness, existence, and Markov property of $(X_t^x)_{t \geq 0}$ follow from Corollary 18.10, Theorem 18.11 and Corollary 18.14, respectively. The (stochastic) integral is, as a function of its upper limit, continuous, cf. Theorem 14.13 or The-

orem 15.7, and so $t \mapsto X_t^x(\omega)$ is continuous. From Corollary 18.21 we know that $\lim_{|x| \rightarrow \infty} |X_t^x| = \infty$ for every $t > 0$. Therefore, we conclude with the dominated convergence theorem that

Ex. 18.11



$$x \mapsto T_t u(x) := \mathbb{E} u(X_t^x) \text{ is in } \mathcal{C}_\infty(\mathbb{R}^d) \text{ for all } t \geq 0 \text{ and } u \in \mathcal{C}_\infty(\mathbb{R}^d).$$

The operators T_t are clearly sub-Markovian and contractive. Let us show that $(T_t)_{t \geq 0}$ is strongly continuous, hence a Feller semigroup. Take $u \in \mathcal{C}_c^2(\mathbb{R}^d)$ and apply Itô's formula, Theorem 16.6. Then

$$\begin{aligned} u(X_t^x) - u(x) &= \sum_{j,k=1}^d \int_0^t \sigma_{jk}(X_s^x) \partial_j u(X_s^x) dB_s^k + \sum_{j=1}^d \int_0^t b_j(X_s^x) \partial_j u(X_s^x) ds \\ &\quad + \frac{1}{2} \sum_{i,j=1}^d \int_0^t a_{ij}(X_s^x) \partial_i \partial_j u(X_s^x) ds. \end{aligned}$$

Since the first term on the right is a martingale, its expected value vanishes, and we get

$$\begin{aligned} |T_t u(x) - u(x)| &= |\mathbb{E} u(X_t^x) - u(x)| \\ &\leq \sum_{j=1}^d \mathbb{E} \left| \int_0^t b_j(X_s^x) \partial_j u(X_s^x) ds \right| + \frac{1}{2} \sum_{i,j=1}^d \mathbb{E} \left| \int_0^t a_{ij}(X_s^x) \partial_i \partial_j u(X_s^x) ds \right| \\ &\leq t \sum_{j=1}^d \|b_j \partial_j u\|_\infty + t \sum_{i,j=1}^d \|a_{ij} \partial_i \partial_j u\|_\infty \xrightarrow{t \rightarrow 0} 0 \end{aligned}$$

Approximate $u \in \mathcal{C}_\infty(\mathbb{R}^d)$ by a sequence $(u_n)_{n \geq 0} \subset \mathcal{C}_c^2(\mathbb{R}^d)$. Then

$$\begin{aligned} \|T_t u - u\|_\infty &\leq \|T_t(u - u_n)\|_\infty + \|T_t u_n - u_n\|_\infty + \|u - u_n\|_\infty \\ &\leq 2\|u - u_n\|_\infty + \|T_t u_n - u_n\|_\infty. \end{aligned}$$

Letting first $t \rightarrow 0$ and then $n \rightarrow \infty$ proves strong continuity on $\mathcal{C}_\infty(\mathbb{R}^d)$. We can now determine the generator of the semigroup. Because of Theorem 7.19 we only need to calculate the limit $\lim_{t \rightarrow 0} (T_t u(x) - u(x))/t$ for each $x \in \mathbb{R}^d$. Using Itô's formula as before, we see that

$$\begin{aligned} &\frac{u(X_t^x) - u(x)}{t} \\ &= \sum_{j=1}^d \mathbb{E} \left(\frac{1}{t} \int_0^t b_j(X_s^x) \partial_j u(X_s^x) ds \right) + \frac{1}{2} \sum_{i,j=1}^d \mathbb{E} \left(\frac{1}{t} \int_0^t a_{ij}(X_s^x) \partial_i \partial_j u(X_s^x) ds \right) \\ &\xrightarrow{t \rightarrow 0} \sum_{j=1}^d b_j(x) \partial_j u(x) + \frac{1}{2} \sum_{i,j=1}^d a_{ij}(x) \partial_i \partial_j u(x). \end{aligned}$$

For the limit in the last line we use dominated convergence and the fact that $t^{-1} \int_0^t \phi(s) ds \rightarrow \phi(0+)$ as $t \rightarrow 0$. \square

Theorem 19.9 shows what we have to do, if we want to construct a diffusion with a given generator $L(x, D)$ of the form (19.1): the drift coefficient b must be globally Lipschitz, and the positive definite diffusion matrix $a = (a_{ij}) \in \mathbb{R}^{d \times d}$ should be of the form $a = \sigma \sigma^\top$ where $\sigma = (\sigma_{jk}) \in \mathbb{R}^{d \times d}$ is globally Lipschitz continuous.

Let $M \in \mathbb{R}^{d \times d}$ and set $\|M\|^2 := \sup_{|\xi|=1} \langle M^\top M \xi, \xi \rangle$. Then $\|M\|$ is a submultiplicative matrix norm which is compatible with the Euclidean norm on \mathbb{R}^d , i.e. $\|MN\| \leq \|M\| \cdot \|N\|$ and $|Mx| \leq \|M\| \cdot |x|$ for all $x \in \mathbb{R}^d$. For a symmetric matrix M we write $\lambda_{\min}(M)$ and $\lambda_{\max}(M)$ for the smallest and largest eigenvalue, respectively; we have $\|M\|^2 = \lambda_{\max}(MM^\top)$ and, if M is semidefinite, $\|M\| = \lambda_{\max}(M)$.¹

19.10 Lemma (van Hemmen–Ando 1980). *Assume that $B, C \in \mathbb{R}^{d \times d}$ are two symmetric, positive semidefinite matrices such that $\lambda_{\min}(B) + \lambda_{\min}(C) > 0$. Then their unique positive semidefinite square roots satisfy*

$$\|\sqrt{C} - \sqrt{B}\| \leq \frac{\|C - B\|}{\max\{\lambda_{\min}^{1/2}(B), \lambda_{\min}^{1/2}(C)\}}.$$

Proof. A simple change of variables shows that

$$\frac{1}{\sqrt{\lambda}} = \frac{2}{\pi} \int_0^\infty \frac{dr}{r^2 + \lambda}, \quad \text{hence} \quad \sqrt{\lambda} = \frac{2}{\pi} \int_0^\infty \frac{\lambda dr}{r^2 + \lambda}.$$

Since B, C are positive semidefinite, we can use this formula to determine the square roots, e.g. for B ,

$$\sqrt{B} = \frac{2}{\pi} \int_0^\infty B(s^2 \text{id} + B)^{-1} ds$$

Set $R_s^B := (s^2 \text{id} + B)^{-1}$ and observe that $BR_s^B - CR_s^C = s^2 R_s^B (B - C) R_s^C$. Thus,

$$\sqrt{C} - \sqrt{B} = \frac{2}{\pi} \int_0^\infty s^2 R_s^C (C - B) R_s^B ds.$$

Since the matrix norm is submultiplicative and compatible with the Euclidean norm, we find for all $\xi \in \mathbb{R}^d$ with $|\xi| = 1$ using the Cauchy-Schwarz inequality

$$\langle R_s^C (C - B) R_s^B \xi, \xi \rangle \leq |R_s^C (C - B) R_s^B \xi| \leq \|R_s^B\| \cdot \|R_s^C\| \cdot \|C - B\|.$$

¹ In fact, if $M = (m_{jk})$, then $\|M\|^2 = \sum_{jk} m_{jk}^2$.

Without loss of generality we may assume that $\lambda_{\min}(C) \geq \lambda_{\min}(B)$. Observe that $\|R_s^B\| = \lambda_{\max}(R_s^B) = (s^2 + \lambda_{\min}(B))^{-1}$. Thus,

$$\begin{aligned} \|\sqrt{C} - \sqrt{B}\| &\leq \frac{2}{\pi} \int_0^\infty \|R_s^B\| \|R_s^C\| s^2 ds \cdot \|C - B\| \\ &\leq \frac{2}{\pi} \int_0^\infty \frac{s^2}{s^2 + \lambda_{\min}(B)} \frac{1}{s^2 + \lambda_{\min}(C)} ds \cdot \|C - B\| \\ &\leq \frac{2}{\pi} \int_0^\infty \frac{1}{s^2 + \lambda_{\min}(C)} ds \cdot \|C - B\| = \frac{\|C - B\|}{\sqrt{\lambda_{\min}(C)}}. \quad \square \end{aligned}$$

19.11 Corollary. *Let $L(x, D)$ be the differential operator (19.1) and assume that the diffusion matrix $a(x)$ is globally Lipschitz continuous (Lipschitz constant C) and uniformly positive definite, i. e. $\inf_x \lambda_{\min}(a(x)) = \kappa > 0$. Moreover, assume that the drift vector is globally Lipschitz continuous. Then there exists a diffusion process in the sense of Definition 19.1 such that its generator $(A, \mathfrak{D}(A))$ coincides on the test functions $\mathcal{C}_c^\infty(\mathbb{R}^d)$ with $L(x, D)$.*

Proof. Denote by $\sigma(x)$ the unique positive definite square root of $a(x)$. Using Lemma 19.10 we find

$$\|\sigma(x) - \sigma(y)\| \leq \frac{1}{\kappa} \|a(x) - a(y)\| \leq \frac{C}{\kappa} |x - y|, \quad \text{for all } x, y \in \mathbb{R}^d.$$

On a finite dimensional vector space all norms are comparable, i. e. (19.9) holds for σ (possibly with a different Lipschitz constant). Therefore, we can apply Theorem 19.9 and obtain the diffusion process $(X_t)_{t \geq 0}$ as the unique solution to the SDE (19.8). \square

It is possible to replace the uniform positive definiteness of the matrix $a(x)$ by the weaker assumption that $a_{ij}(x) \in \mathcal{C}_b^2(\mathbb{R}^d, \mathbb{R})$ for all $i, j = 1, \dots, d$, cf. [80, Proposition IV.6.2] or [177, Theorem 5.2.3].

The square root σ is only unique if we require that it is also positive semidefinite. Let us briefly discuss what happens if $\sigma\sigma^\top = a = \rho\rho^\top$ where ρ is some $d \times d$ matrix which need not be positive semidefinite. Let $(\sigma, b, (B_t)_{t \geq 0})$ be as in Theorem 19.9 and denote by $(X_t^x)_{t \geq 0}$ the unique solution of the SDE (19.8). Assume that $(W_t)_{t \geq 0}$ is a further Brownian motion which is independent of $(B_t)_{t \geq 0}$ and that $(Y_t^x)_{t \geq 0}$ is the unique solution of the SDE

$$dY_t = b(Y_t) dt + \rho(Y_t) dW_t \quad Y_0 = x \in \mathbb{R}^d.$$

The proof of Theorem 19.9 works literally for $(Y_t^x)_{t \geq 0}$ since we only need the existence and uniqueness of the solution of the SDE, and the equality $\rho\rho^\top = a$. This means that $(X_t^x)_{t \geq 0}$ and $(Y_t^x)_{t \geq 0}$ are two Feller processes whose generators, $(A, \mathfrak{D}(A))$ and $(C, \mathfrak{D}(C))$, say, coincide on the set $\mathcal{C}_c^2(\mathbb{R}^d)$. It is tempting to claim that the processes coincide, but for this we need to know that $A|_{\mathcal{C}_c^2} = C|_{\mathcal{C}_c^2}$ implies $(A, \mathfrak{D}(A)) =$

$(C, \mathfrak{D}(C))$. A sufficient criterion would be that \mathcal{C}_c^2 is an operator core for A and C , i. e. A and C are the closures of $A|_{\mathcal{C}_c^2}$ and $C|_{\mathcal{C}_c^2}$, respectively. This is far from obvious.

Using Stroock-and-Varadhan's martingale problem technique, one can show that – under even weaker conditions than those of Corollary 19.11 – the diffusion process $(X_t^x)_{t \geq 0}$ is unique in the following sense: Every process whose generator coincides on $\mathcal{C}_c^2(\mathbb{R}^d)$ with $L(x, D)$, has the same finite dimensional distributions as $(X_t^x)_{t \geq 0}$. On the other hand, this is the best we can hope for. For example, we have

$$a(x) = \sigma(x)\sigma^\top(x) = \rho(x)\rho^\top(x) \quad \text{if} \quad \rho(x) = \sigma(x)U$$

for any orthogonal matrix $U \in \mathbb{R}^{d \times d}$. If U is not a function of x , we see that

$$\begin{aligned} dY_t &= b(Y_t)dt + \rho(Y_t)dB_t = b(Y_t)dt + \sigma(Y_t)UdB_t \\ &= b(Y_t)dt + \sigma(Y_t)d(UB)_t. \end{aligned}$$

Of course, $(UB_t)_{t \geq 0}$ is again a Brownian motion, but it is different from the original Brownian motion $(B_t)_{t \geq 0}$. This shows that the solution $(Y_t^x)_{t \geq 0}$ will, even in this simple case, not coincide with $(X_t^x)_{t \geq 0}$ in a pathwise sense. Nevertheless, $(X_t^x)_{t \geq 0}$ and $(Y_t^x)_{t \geq 0}$ have the same finite dimensional distributions. In other words: We cannot expect that there is a strong one-to-one relation between L and the stochastic process.



19.12 Further reading. The interplay of PDEs and SDEs is nicely described in [8] and [66]. The contributions of Stroock and Varadhan, in particular, the characterization of stochastic processes via the martingale problem, opens up a whole new world. Good introductions are [61] and, with the focus on diffusion processes, [177]. The survey [9] gives a brief introduction to the ideas of Malliavin's calculus, see also the *Further reading* recommendations in Chapter 4 and 17.

[8] Bass: *Diffusions and Elliptic Operators*.

[9] Bell: Stochastic differential equations and hypoelliptic operators.

[61] Ethier, Kurtz: *Markov Processes: Characterization and Convergence*.

[66] Freidlin: *Functional Integration and Partial Differential Equations*.

[177] Stroock, Varadhan: *Multidimensional Diffusion Processes*.

Problems

1. Let $(A, \mathfrak{D}(A))$ be the generator of a diffusion process in the sense of Definition 19.1 and denote by a, b the diffusion and drift coefficients. Show that $a \in \mathcal{C}(\mathbb{R}^d, \mathbb{R}^{d \times d})$ and $b \in \mathcal{C}(\mathbb{R}^d, \mathbb{R}^d)$.

2. Show that under the assumptions of Proposition 19.5 we can interchange integration and differentiation: $\frac{\partial^2}{\partial x_j \partial x_k} \int p(t, x, y) u(y) dy = \int \frac{\partial^2}{\partial x_j \partial x_k} p(t, x, y) u(y) dy$ and that the resulting function is in $\mathcal{C}_\infty(\mathbb{R}^d)$.
3. Denote by $\|u\|_{(2)} = \|u\|_\infty + \sum_j \|\partial_j u\|_\infty + \sum_{j,k} \|\partial_j \partial_k u\|_\infty$. Assume that $A : \mathcal{C}_c^\infty(\mathbb{R}^d) \rightarrow \mathcal{C}_\infty(\mathbb{R}^d)$ is a closed operator such that $\|Au\|_\infty \leq C \|u\|_{(2)}$. Show that $\mathcal{C}_c^2(\mathbb{R}^d) = \overline{\mathcal{C}_c^\infty(\mathbb{R}^d)}^{\|\cdot\|_{(2)}} \subset \mathfrak{D}(A)$.
4. Let $L = L(x, D_x) = \sum_{ij} a_{ij}(x) \partial_i \partial_j + \sum_j b_j(x) \partial_j + c(x)$ be a second order differential operator. Determine its formal adjoint $L^*(y, D_y)$, i. e. the linear operator satisfying $\langle Lu, \phi \rangle_{L^2} = \langle u, L^* \phi \rangle_{L^2}$ for all $u, \phi \in \mathcal{C}_c^\infty(\mathbb{R}^d)$. What is the formal adjoint of the operator $L(x, D) + \frac{\partial}{\partial t}$?
5. Complete the proof of Proposition 19.6 (Kolmogorov's forward equation).
6. Let $(B_t)_{t \geq 0}$ be a BM¹. Show that $X_t := (B_t, \int_0^t B_s ds)$ is a diffusion process and determine its generator.
7. (Carré-du-champ operator) Let $(X_t)_{t \geq 0}$ be a diffusion process with generator $L = L(x, D) = A|_{\mathcal{C}_c^\infty}$ as in (19.1). Write $M_t^u = u(X_t) - u(X_0) - \int_0^t Lf(X_r) dr$.
 - (a) Show that $M^u = (M_t^u)_{t \geq 0}$ is an L^2 -martingale for all $u \in \mathcal{C}_c^\infty(\mathbb{R}^d)$.
 - (b) Show that $M^f = (M_t^f)_{t \geq 0}$ is local martingale for all $f \in \mathcal{C}^2(\mathbb{R}^d)$.
 - (c) Denote by $\langle M^u, M^\phi \rangle_t$, $u, \phi \in \mathcal{C}_c^\infty(\mathbb{R}^d)$, the quadratic covariation of the L^2 martingales M^u and M^ϕ , cf. (14.7) and Theorem 14.13 b). Show that

$$\begin{aligned} \langle M^u, M^\phi \rangle_t &= \int_0^t (L(u\phi) - uL\phi - \phi Lu)(X_s) ds \\ &= \int_0^t \nabla u(X_s) \cdot a(X_s) \nabla \phi(X_s) ds. \end{aligned}$$

($x \cdot y$ denotes the Euclidean scalar product and $\nabla = (\partial_1, \dots, \partial_d)^\top$.)

Remark: The operator $\Gamma(u, \phi) := L(u\phi) - uL\phi - \phi Lu$ is called *carré-du-champ* or *mean-field operator*.

Chapter 20

Simulation of Brownian motion

by Björn Böttcher

This chapter presents a concise introduction to simulation, in particular focusing on the normal distribution and Brownian motion.

20.1 Introduction

Recall that for iid random variables X_1, \dots, X_n with a common distribution F for any $\omega \in \Omega$ the tuple $(X_1(\omega), \dots, X_n(\omega))$ is called a (*true*) *sample* of F . In simulations the following notion of a sample is commonly used.

20.1 Definition. A tuple of values (x_1, \dots, x_n) which is indistinguishable from a true sample of F by statistical inference is called a *sample* of F . An element of such a tuple is called a *sample value* of F , and it is denoted by

$$x_j \stackrel{s}{\sim} F.$$

Note that the statistical inference has to ensure both the distributional properties and the independence. A naive idea to verify the distributional properties would be, for instance, the use of a Kolmogorov–Smirnov test on the distribution. But this is not a good test for the given problem, since a true sample would fail such a test with confidence level α with probability α . Instead one should do at least repeated Kolmogorov–Smirnov tests and then test the p -values for uniformity. In fact, there is a wide range of possible tests, for example the NIST (National Institute of Standards and Technology)¹ has defined a set of tests which should be applied for reliable results.

In theory, letting n tend to ∞ and allowing all possible statistical tests, only a true sample would be a sample. But in practice it is impossible to get a true sample of a continuous distribution, since a physical experiment as sample mechanism clearly

¹ http://csrc.nist.gov/groups/ST/toolkit/rng/stats_tests.html

yields measurement and discretization errors. Moreover for the generation of samples with the help of a computer we only have a discrete and finite set of possible values due to the internal number representation. Thus one considers n large, but not arbitrarily large.

Nevertheless, in simulations one treats a sample as if it were a true sample, i.e. any distributional relation is translated to samples. For example, given a distribution function F and the uniform distribution $\mathbf{U}_{(0,1)}$ we have

$$X \sim \mathbf{U}_{(0,1)} \implies F^{-1}(X) \sim F$$

since

$$\mathbb{P}(F^{-1}(X) \leq x) = \mathbb{P}(X \leq F(x)) = F(x)$$

and thus

$$x \stackrel{s}{\sim} \mathbf{U}_{(0,1)} \implies F^{-1}(x) \stackrel{s}{\sim} F \quad (20.1)$$

holds. This yields the following algorithm.

20.2 Algorithm (Inverse transform method). Let F be a distribution function.

1. Generate $u \stackrel{s}{\sim} \mathbf{U}_{(0,1)}$

2. Set $x := F^{-1}(u)$

Then $x \stackrel{s}{\sim} F$.

By this algorithm a sample of $\mathbf{U}_{(0,1)}$ can be transformed to a sample of any other distribution. Therefore the generation of samples from the uniform distribution is the basis for all simulations. For this various algorithms are known and they are implemented in most of the common programming languages. Readers interested in the details of the generation of samples of the uniform distribution are referred to [117, 118].

From now on we will assume that we can generate $x \stackrel{s}{\sim} \mathbf{U}_{(0,1)}$. Thus we can apply Algorithm 20.2 to get samples of a given distribution as the following examples illustrate.

20.3 Example (Scaled and shifted Uniform distribution $\mathbf{U}_{(a,b)}$). For $a < b$ the distribution function of $\mathbf{U}_{(a,b)}$ is given by

$$F(x) = \frac{x-a}{b-a} \quad \text{for } x \in [a, b]$$

and thus

$$u \stackrel{s}{\sim} \mathbf{U}_{(0,1)} \implies (b-a)u + a \stackrel{s}{\sim} \mathbf{U}_{(a,b)}.$$

20.4 Example (Bernoulli distribution β_p and $\mathbf{U}_{\{0,1\}}$, $\mathbf{U}_{\{-1,1\}}$). Let $p \in [0, 1]$. Given $U \sim \mathbf{U}_{(0,1)}$, then $\mathbb{1}_{[0,p]}(U) \sim \beta_p$. In particular, we get

$$u \stackrel{s}{\sim} \mathbf{U}_{(0,1)} \implies \mathbb{1}_{[0,\frac{1}{2}]}(u) \stackrel{s}{\sim} \mathbf{U}_{\{0,1\}} \quad \text{and} \quad 2 \cdot \mathbb{1}_{[0,\frac{1}{2}]}(u) - 1 \stackrel{s}{\sim} \mathbf{U}_{\{-1,1\}}.$$

20.5 Example (Exponential distribution Exp_λ). The distribution function of Exp_λ with parameter $\lambda > 0$ is $F(x) = 1 - e^{-\lambda x}$, $x \geq 0$, and thus

$$F^{-1}(y) = -\frac{1}{\lambda} \ln(1 - y).$$

Now using

$$U \sim \mathbf{U}_{(0,1)} \implies 1 - U \sim \mathbf{U}_{(0,1)}$$

yields with the inverse transform method

$$u \stackrel{s}{\sim} \mathbf{U}_{(0,1)} \implies -\frac{1}{\lambda} \ln(u) \stackrel{s}{\sim} \text{Exp}_\lambda.$$

Before focusing on the normal distribution, let us introduce two further algorithms which can be used to transform a sample of a distribution G to a sample of a distribution F .

20.6 Algorithm (Acceptance–rejection method). Let F and G be probability distribution functions with densities f and g , such that there is a constant $c > 0$ with $f(x) \leq cg(x)$ for all x .

1. Generate $y \stackrel{s}{\sim} G$
2. Generate $u \stackrel{s}{\sim} \mathbf{U}_{(0,1)}$
3. If $ucg(y) \leq f(y)$, then $x := y$
otherwise restart with 1.

Then $x \stackrel{s}{\sim} F$.

Proof. If $Y \sim G$ and $U \sim \mathbf{U}_{(0,1)}$ are independent, then the algorithm generates a sample value of the distribution for $A \in \mathcal{B}$ by

$$\mathbb{P}(Y \in A \mid Uc g(Y) \leq f(Y)) = \frac{\mathbb{P}(Y \in A, Uc g(Y) \leq f(Y))}{\mathbb{P}(Uc g(Y) \leq f(Y))} = \int_A f(y) dy.$$

This follows from the fact that $g(x) = 0$ implies $f(x) = 0$, and

$$\begin{aligned}
 \mathbb{P}(Y \in A, Ucg(Y) \leq f(Y)) &= \int_0^1 \int_A \mathbb{1}_{\{(v,z): vcg(z) \leq f(z)\}}(u, y) g(y) dy du \\
 &= \int_{A \setminus \{g=0\}} \int_0^{\frac{f(y)}{cg(y)}} du g(y) dy \\
 &= \int_{A \setminus \{g=0\}} \frac{f(y)}{cg(y)} g(y) dy \\
 &= \frac{1}{c} \int_{A \setminus \{g=0\}} f(y) dy \\
 &= \frac{1}{c} \int_A f(y) dy.
 \end{aligned}$$

In particular, we see for $A = \mathbb{R}$

$$\mathbb{P}(Ucg(Y) \leq f(Y)) = \mathbb{P}(Y \in \mathbb{R}, Ucg(Y) \leq f(Y)) = \frac{1}{c}. \quad \square$$

20.7 Example (One-sided normal distribution). We can apply the acceptance–rejection method to the one-sided normal distribution

$$f(x) = 2 \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} \mathbb{1}_{[0,\infty)}(x),$$

and the Exp_1 distribution $g(x) = e^{-x} \mathbb{1}_{[0,\infty)}(x)$ (for sampling see Example 20.5) with the smallest possible c (such that $f(x) \leq cg(x)$ for all x) given by

$$c = \sup_x \frac{f(x)}{g(x)} = \sup_x \sqrt{\frac{2}{\pi}} e^{-\frac{x^2}{2}+x} \mathbb{1}_{[0,\infty)}(x) = \sqrt{\frac{2e}{\pi}} \approx 1.315.$$

The name of Algorithm 20.6 is due to the fact that in the first step a sample is proposed and in the third step this is either accepted or rejected. In the case of rejection a new sample is proposed until acceptance. The algorithm can be visualized as throwing darts onto the set $B = \{(x, y) \in \mathbb{R}^2 : 0 \leq y \leq cg(x)\}$ uniformly. Then the x -coordinate of the dart is taken as a sample (accepted) if, and only if, the dart is within $C = \{(x, y) \in \mathbb{R}^2 : 0 \leq y \leq f(x)\}$; note that $C \subset B$ is satisfied by our assumptions. This also shows that, on average, the algorithm proposes $\lambda^2(B)/\lambda^2(C) = c$ times a sample until acceptance.

In contrast the following algorithm does not produce *unused* samples.

20.8 Algorithm (Composition method). Let F be a distribution with density f of the form

$$f(x) = \int g_y(x) \mu(dy)$$

where μ is a probability measure and g_y is for each y a density.

1. Generate $y \stackrel{s}{\sim} \mu$
2. Generate $x \stackrel{s}{\sim} g_y$

Then $x \stackrel{s}{\sim} F$.

Proof. Let $Y \sim \mu$ and, conditioned on $\{Y = y\}$, let $X \sim g_y$, i. e.

$$P(X \leq x) = \int \mathbb{P}(X \leq x \mid Y = y) \mu(dy) = \iint_{-\infty}^x g_y(z) dz \mu(dy) = \int_{-\infty}^x f(z) dz.$$

Thus we have $X \sim F$. □

20.9 Example (Normal distribution). We can write the density of the standard normal distribution as

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} = \frac{1}{2} \underbrace{\frac{2}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} \mathbb{1}_{(-\infty, 0]}(x)}_{= g_{-1}(x)} + \frac{1}{2} \underbrace{\frac{2}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} \mathbb{1}_{(0, \infty)}(x)}_{= g_1(x)}$$

i. e. $f(x) = \int_{-\infty}^{\infty} g_y(x) \mu(dy)$ with $\mu(dy) = \frac{1}{2}\delta_{-1}(dy) + \frac{1}{2}\delta_1(dy)$.

Note that $X \sim g_1$ can be generated by Example 20.7 and that

$$X \sim g_1 \implies -X \sim g_{-1}$$

holds. Thus by Example 20.3 and the composition method we obtain the following algorithm.

20.10 Algorithm (Composition of one-sided normal samples).

1. Generate $y \stackrel{s}{\sim} \mathbf{U}_{\{-1, 1\}}$, $z \stackrel{s}{\sim} f$ (where the function f denotes the density of the one-sided normal distribution, see Example 20.7)
2. Set $x := yz$

Then $x \stackrel{s}{\sim} \mathbf{N}(0, 1)$.

In the next section we will see further algorithms to generate samples of the normal distribution.

20.2 Normal distribution

We start with the one-dimensional normal distribution and note that for $\mu \in \mathbb{R}$ and $\sigma > 0$

$$X \sim \mathcal{N}(0, 1) \implies \sigma X + \mu \sim \mathcal{N}(\mu, \sigma^2).$$

Thus we have the following algorithm to transform samples of the standard normal distribution to samples of a general normal distribution.

20.11 Algorithm (Change of mean and variance). Let $\mu \in \mathbb{R}$ and $\sigma > 0$.

1. Generate $y \stackrel{s}{\sim} \mathcal{N}(0, 1)$
2. Set $x := \sigma y + \mu$

Then $x \stackrel{s}{\sim} \mathcal{N}(\mu, \sigma^2)$.

Based on the above algorithm we can focus in the following on the generation of samples of $\mathcal{N}(0, 1)$.

The standard approach is the inverse transform method, Algorithm 20.2, with a numerical inversion of the distribution function of the standard normal distribution.

20.12 Algorithm (Numerical inversion).

1. Generate $u \stackrel{s}{\sim} \mathcal{U}_{(0,1)}$
2. Set $x := F^{-1}(u)$ (where F is the $\mathcal{N}(0, 1)$ -distribution function)

Then $x \stackrel{s}{\sim} \mathcal{N}(0, 1)$.

A different approach which generates two samples at once is the following

20.13 Algorithm (Box–Muller-method).

1. Generate $u_1, u_2 \stackrel{s}{\sim} U_{(0,1)}$
2. Set $r := \sqrt{-2 \ln(u_1)}$, $v := 2\pi u_2$
3. Set $x_1 := r \cos(v)$, $x_2 := r \sin(v)$

Then $x_j \stackrel{s}{\sim} N(0, 1)$.

Proof. Let $X_1, X_2 \sim N(0, 1)$ be independent. Thus they have the joint density $f_{X_1, X_2}(x_1, x_2) = \frac{1}{2\pi} \exp\left(-\frac{x_1^2 + x_2^2}{2}\right)$. Now we express the vector (X_1, X_2) in polar coordinates, i. e.

$$\begin{pmatrix} X_1 \\ X_2 \end{pmatrix} = \begin{pmatrix} R \cos V \\ R \sin V \end{pmatrix} \quad (20.2)$$

and the probability density becomes, in polar coordinates, $r > 0, v \in [0, 2\pi)$

$$f_{R,V}(r, v) = f_{X_1, X_2}(r \cos v, r \sin v) \det \begin{pmatrix} \cos v & -r \sin v \\ \sin v & r \cos v \end{pmatrix} = e^{-\frac{1}{2}r^2} \frac{r}{2\pi}.$$

Thus R and V are independent, $V \sim U_{[0, 2\pi)}$ and the distribution of R^2 is

$$\mathbb{P}(R^2 \leq x) = \int_0^{\sqrt{x}} e^{-\frac{1}{2}r^2} r \, dr = \int_0^x \frac{1}{2} e^{-\frac{1}{2}s} \, ds = 1 - e^{-\frac{1}{2}x},$$

i. e. $R^2 \sim \text{Exp}_{1/2}$.

The second step of the algorithm is just an application of the inverse transform method, Examples 20.3 and 20.5, to generate samples of R and V , in the final step (20.2) is applied. \square

A simple application of the central limit theorem can be used to get (approximately) standard normal distributed samples.

20.14 Algorithm (Central limit method).

1. Generate $u_j \stackrel{s}{\sim} U_{(0,1)}$, $j = 1, \dots, 12$
2. Set $x := \sum_{j=1}^{12} u_j - 6$

Then $x \stackrel{s}{\sim} N(0, 1)$ approximately.

Proof. Let $U_j \sim U_{(0,1)}$, $j = 1, 2, \dots, 12$, be independent, then $\mathbb{E}(U_j) = \frac{1}{2}$ and

$$\mathbb{V} U_j = \mathbb{E} U_j^2 - (\mathbb{E} U_j)^2 = \frac{x^3}{3} \Big|_0^1 - \frac{1}{4} = \frac{4}{12} - \frac{3}{12} = \frac{1}{12}.$$

By the central limit theorem

$$\frac{\sum_{j=1}^{12} U_j - 12 \mathbb{E} U_1}{\sqrt{12 \mathbb{V} U_1}} = \sum_{j=1}^{12} U_j - 6 \sim N(0, 1) \quad (\text{approximately}). \quad \square$$

Clearly one could replace the constant 12 in the first step of the above algorithm by any other integer n , but then the corresponding sum in step 2 would have to be rescaled by $(n/12)^{-\frac{1}{2}}$ and this would require a further calculation. In the above setting it can be shown that the pointwise difference of the standard normal distribution function and the distribution function of $\left(\sum_{j=1}^{12} U_j - 12 \mathbb{E} U_1\right) / \sqrt{12 \mathbb{V} U_1}$ for iid $U_j \sim U_{(0,1)}$ is always less than 0.0023, see Section 3.1 in [160].

Now we have seen several ways to generate samples of the normal distribution, i. e. we know how to simulate increments of Brownian motion.

20.3 Brownian motion

A simulation of a continuous time process is always based on a finite sample, since we can only generate finitely many values in finite time.

Brownian motion is usually simulated on a discrete time grid $0 = t_0 < t_1 < \dots < t_n$, and it is common practice for visualizations to interpolate the simulated values linearly. This is justified by the fact that Brownian Motion has continuous paths.

Throughout this section $(B_t)_{t \geq 0}$ denotes a one-dimensional standard Brownian motion. The simplest method to generate a Brownian path uses the property of stationary and independent increments.

20.15 Algorithm (Independent increments). Let $0 = t_0 < t_1 < \dots < t_n$ be a time grid.

Initialize $b_0 := 0$

For $j = 1$ to n

1. Generate $y \stackrel{s}{\sim} N(0, t_j - t_{j-1})$

2. Set $b_{t_j} := b_{t_{j-1}} + y$

Then $(b_0, b_{t_1}, \dots, b_{t_n}) \stackrel{s}{\sim} (B_0, B_{t_1}, \dots, B_{t_n})$.

This method works particularly well if the step size $\delta = t_j - t_{j-1}$ is constant for all j , in this case one only has to generate $y \stackrel{s}{\sim} N(0, \delta)$ repeatedly. A drawback of the method is, that any refinement of the discretization yields a new simulation and, thus, a different path.

To get a refinement of an already simulated path one can use the idea of Lévy's original argument, see Section 3.2.

20.16 Algorithm (Interpolation. Lévy's argument). Let $0 \leq s_0 < s < s_1$ be fixed times and (b_{s_0}, b_{s_1}) be a sample value of (B_{s_0}, B_{s_1}) .

1. Generate $b_s \stackrel{s}{\sim} N\left(\frac{(s_1-s)b_{s_0} + (s-s_0)b_{s_1}}{s_1-s_0}, \frac{(s-s_0)(s_1-s)}{s_1-s_0}\right)$

Then $(b_{s_0}, b_s, b_{s_1}) \stackrel{s}{\sim} (B_{s_0}, B_s, B_{s_1})$ given that $B_{s_0} = b_{s_0}$ and $B_{s_1} = b_{s_1}$.

Proof. In Section 3.2 we have seen that in the above setting

$$\mathbb{P}(B_s \in \cdot \mid B_{s_0} = b_{s_0}, B_{s_1} = b_{s_1}) = N(m_s, \sigma_s^2)$$

with

$$m_s = \frac{(s_1-s)b_{s_0} + (s-s_0)b_{s_1}}{s_1-s_0} \quad \text{and} \quad \sigma_s^2 = \frac{(s-s_0)(s_1-s)}{s_1-s_0}. \quad \square$$

Thus, for $0 = t_0 < t_1 < \dots < t_n$ and $(b_0, b_{t_1}, \dots, b_{t_n}) \stackrel{s}{\sim} (B_0, B_{t_1}, \dots, B_{t_n})$ we get

$$(b_0, b_{t_1}, \dots, b_{t_j}, b_r, b_{t_{j+1}}, \dots, b_{t_n}) \stackrel{s}{\sim} (B_0, B_{t_1}, \dots, B_{t_j}, B_r, B_{t_{j+1}}, \dots, B_{t_n})$$

for some $r \in (t_j, t_{j+1})$ by generating the intermediate point b_r given $(b_{t_j}, b_{t_{j+1}})$. Repeated applications yield an arbitrary refinement of a given discrete simulation of a Brownian path. Taking only the dyadic refinement of the unit interval we get the Lévy–Ciesielski representation, cf. Section 3.1.

20.17 Algorithm (Lévy–Ciesielski). Let $J \geq 1$ be the order of refinement.

Initialize $b_0 := 0$

Generate $b_1 \stackrel{s}{\sim} N(0, 1)$

For $j = 0$ to $J - 1$

For $l = 0$ to $2^j - 1$

1. Generate $y \stackrel{s}{\sim} N(0, 1)$

2. Set $b_{(2l+1)/2^{j+1}} = \frac{1}{2}(b_{l/2^j} + b_{(l+1)/2^j}) + 2^{-(\frac{j}{2}+1)}y$

Then $(b_0, b_{1/2^J}, b_{2/2^J}, \dots, b_1) \stackrel{s}{\sim} (B_0, B_{1/2^J}, B_{2/2^J}, \dots, B_1)$.

Proof. As noted at the end of Section 3.2, the expansion of B_t at dyadic times has only finitely many terms and

$$B_{(2l+1)/2^{j+1}} = \frac{1}{2} (B_{l/2^j} + B_{(l+1)/2^j}) + 2^{-(\frac{j}{2}+1)} Y$$

where $Y \sim N(0, 1)$. □

Analogous to the central limit method for the normal distribution we can also use Donsker's invariance principle, Theorem 3.8, to get an approximation to a Brownian path.

20.18 Algorithm (Donsker's invariance principle). Let n be the number of steps.

Initialize $b_0 := 0$

1. Generate $u_j \stackrel{s}{\sim} U_{\{-1,1\}}$, $j = 1, \dots, n$

2. For $k = 1$ to n

Set $b_{k/n} = \frac{1}{\sqrt{n}} u_k + b_{(k-1)/n}$

Then $(b_0, b_{1/n}, \dots, b_{2/n}, \dots, b_1) \stackrel{s}{\sim} (B_0, B_{1/n}, B_{2/n}, \dots, B_1)$ approximately.

Clearly one could also replace the Bernoulli random walk in the above algorithm by a different symmetric random walk with finite variance.

20.4 Multivariate Brownian motion

Since a d -dimensional Brownian motion has independent one-dimensional Brownian motions as components, cf. Corollary 2.15, we can use the algorithms from the last section to generate the components separately.

In order to generate a Q -Brownian motion we need to simulate samples of a multivariate normal distribution with correlated components.

20.19 Algorithm (Two-dimensional Normal distribution). Let an expectation vector $\mu \in \mathbb{R}^2$, variances $\sigma_1^2, \sigma_2^2 > 0$ and correlation $\rho \in [-1, 1]$ be given.

1. Generate $y_1, y_2, y_3 \stackrel{s}{\sim} \mathbf{N}(0, 1)$
2. Set $x_1 := \sigma_1 \left(\sqrt{1 - |\rho|} y_1 + \sqrt{|\rho|} y_3 \right) + \mu_1$
 $x_2 := \sigma_2 \left(\sqrt{1 - |\rho|} y_2 + \text{sgn}(\rho) \sqrt{|\rho|} y_3 \right) + \mu_2$

Then $(x_1, x_2)^\top \stackrel{s}{\sim} \mathbf{N} \left(\mu, \begin{pmatrix} \sigma_1^2 & \rho \sigma_1 \sigma_2 \\ \rho \sigma_1 \sigma_2 & \sigma_2^2 \end{pmatrix} \right)$.

Proof. Let $Y_1, Y_2, Y_3 \sim \mathbf{N}(0, 1)$ be independent random variables. Then

$$\sigma_1 \sqrt{1 - |\rho|} \begin{pmatrix} Y_1 \\ 0 \end{pmatrix} + \sigma_2 \sqrt{1 - |\rho|} \begin{pmatrix} 0 \\ Y_2 \end{pmatrix} + \sqrt{|\rho|} \begin{pmatrix} \sigma_1 Y_3 \\ \text{sgn}(\rho) \sigma_2 Y_3 \end{pmatrix}$$

is a normal random variable since it is a sum of independent $\mathbf{N}(0, \cdot)$ distributed random vectors. Observe that the last vector is bivariate normal with completely dependent components. Calculating the covariance matrix yields the statement. \square

This algorithm shows nicely how the correlation of the components is introduced into the sample. In the next algorithm this is less obvious, but the algorithm works in any dimension.

20.20 Algorithm (Multivariate Normal distribution). Let a positive semidefinite covariance matrix $Q \in \mathbb{R}^{d \times d}$ and an expectation vector $\mu \in \mathbb{R}^d$ be given. Calculate the Cholesky decomposition $\Sigma \Sigma^\top = Q$

1. Generate $y_1, \dots, y_d \stackrel{s}{\sim} \mathbf{N}(0, 1)$ and let y be the vector with components y_1, \dots, y_d
2. Set $x := \mu + \Sigma y$

Then $x \stackrel{s}{\sim} \mathbf{N}(\mu, Q)$.

Proof. Compare with the discussion following Definition 2.17. \square

Now we can simulate Q -Brownian motion using the independence and stationarity of the increments.

20.21 Algorithm (Q -Brownian motion). Let $0 = t_0 < t_1 < \dots < t_n$ be a time grid and $Q \in \mathbb{R}^{d \times d}$ be a positive semidefinite covariance matrix.

Initialize $b_k^0 := 0, k = 1, \dots, d$

For $j = 1$ to n

1. Generate $y \stackrel{s}{\sim} N(0, (t_j - t_{j-1})Q)$
2. Set $b^{t_j} := b^{t_{j-1}} + y$

Then $(b^0, b^{t_1}, \dots, b^{t_n}) \stackrel{s}{\sim} (B_0, B_{t_1}, \dots, B_{t_n})$, where $(B_t)_{t \geq 0}$ is a Q -Brownian motion.

20.22 Example. Figure 20.1 shows the sample path of a Q -Brownian motion for $t \in [0, 10]$ with step width 0.0001 and correlation $\rho = 0.8$, i.e. $Q = \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix}$. Note that it seems as if the correlation were visible, but – especially since it is just a single sample and the scaling of the axes is not given – one should be careful when drawing conclusions from samples.

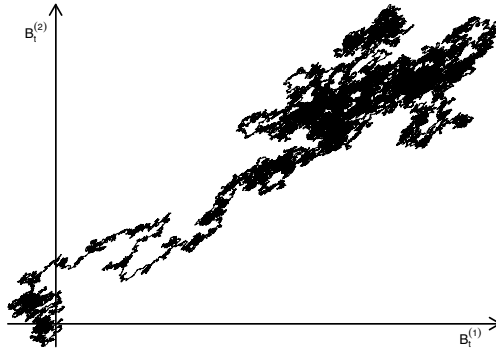


Figure 20.1. Simulation of a 2-dimensional Q -Brownian motion.

20.5 Stochastic differential equations

Based on Chapter 18 we consider in this section a process $(X_t)_{t \geq 0}$ defined by the stochastic differential equation

$$X_0 = x_0 \quad \text{and} \quad dX_t = b(t, X_t) dt + \sigma(t, X_t) dB_t. \quad (20.3)$$

Corollary 18.10 and Theorem 18.11 provide conditions which ensure that (20.3) has a unique solution. We are going to present two algorithms that allow us to simulate the solution.

20.23 Algorithm (Euler scheme). Let b and σ in (20.3) satisfy for all $x, y \in \mathbb{R}$, $0 \leq s < t \leq T$ and some $K > 0$

$$|b(t, x) - b(t, y)| + |\sigma(t, x) - \sigma(t, y)| \leq K |x - y|, \quad (20.4)$$

$$|b(t, x)| + |\sigma(t, x)| \leq K(1 + |x|), \quad (20.5)$$

$$|b(t, x) - b(s, x)| + |\sigma(t, x) - \sigma(s, x)| \leq K(1 + |x|)\sqrt{t - s}. \quad (20.6)$$

Let $n \geq 1$, $0 = t_0 < t_1 < \dots < t_n = T$ be a time grid, and x_0 be the initial position.

For $j = 1$ to n

1. Generate $y \stackrel{s}{\sim} N(0, t_j - t_{j-1})$
2. Set $x_{t_j} := x_{t_{j-1}} + b(t_{j-1}, x_{t_{j-1}})(t_j - t_{j-1}) + \sigma(t_{j-1}, x_{t_{j-1}}) y$

Then $(x_0, x_{t_1}, \dots, x_{t_n}) \stackrel{s}{\sim} (X_0, X_{t_1}, \dots, X_{t_n})$ approximately, where $(X_t)_{t \geq 0}$ is the solution to (20.3).

Note that the conditions (20.4) and (20.5) ensure the existence of a unique solution.

The approximation error can be measured in various ways. For this denote, in the setting of the algorithm, the maximal step width by $\delta := \max_{j=1, \dots, n} (t_j - t_{j-1})$ and the sequence of random variables corresponding to step 2 by $(X_{t_j}^\delta)_{j=0, \dots, n}$. Given a Brownian motion $(B_t)_{t \geq 0}$, we can define $X_{t_j}^\delta$, by $X_{t_0}^\delta := x_0$, and for $j = 1, \dots, n$ by

$$X_{t_j}^\delta := X_{t_{j-1}}^\delta + b(t_{j-1}, X_{t_{j-1}}^\delta)(t_j - t_{j-1}) + \sigma(t_{j-1}, X_{t_{j-1}}^\delta)(B_{t_j} - B_{t_{j-1}}). \quad (20.7)$$

The process $(X_{t_j}^\delta)_{j=0, \dots, n}$ is called the *Euler scheme* approximation of the solution $(X_t)_{t \geq 0}$ to the equation (20.3).

We say that an approximation scheme has *strong order* of convergence β if

$$\mathbb{E} |X_T^\delta - X_T| \leq C \delta^\beta.$$

The following result shows that the algorithm provides an approximate solution to (20.3).

20.24 Theorem. *In the setting of Algorithm 20.23 the strong order of convergence of the Euler scheme is $1/2$. Moreover, the convergence is locally uniform, i. e.*

$$\sup_{t \leq T} \mathbb{E} |X_t^\delta - X_t| \leq c_T \sqrt{\delta},$$

where $(X_t^\delta)_{t \leq T}$ is the piecewise constant extension of the Euler scheme from the discrete time set t_0, \dots, t_n to $[0, T]$ defined by

$$X_t^\delta := X_{t_{n(t)}}^\delta$$

with $n(t) := \max\{n : t_n \leq t\}$.

In the same setting and with some additional technical conditions, cf. [3] and [104], one can show that the *weak order* of convergence of the Euler scheme is 1, i. e.

$$|\mathbb{E} g(X_T^\delta) - \mathbb{E} g(X_T)| \leq c_T \delta$$

for any $g \in \mathcal{C}^4$ whose first four derivatives grow at most polynomially.

Proof of Theorem 20.24. Let b, σ, T, n , be as in Algorithm 20.23. Furthermore let $(X_t)_{t \geq 0}$ be the solution to the corresponding stochastic differential equation (20.3), $(X_{t_j}^\delta)_{j=0, \dots, n}$ be the Euler scheme defined by (20.7) and $(X_t^\delta)_{t \leq T}$ be its piecewise constant extension onto $[0, T]$.

Note that this allows us to rewrite X_t^δ as

$$X_t^\delta = x_0 + \int_0^{t_{n(t)}} b(t_{n(s)}, X_{t_{n(s)}}^\delta) ds + \int_0^{t_{n(t)}} \sigma(t_{n(s)}, X_{t_{n(s)}}^\delta) dB_s.$$

Thus, we can calculate the error by

$$\begin{aligned} Z(T) &:= \sup_{t \leq T} \mathbb{E} [|X_t^\delta - X_t|^2] \leq \mathbb{E} \left[\sup_{t \leq T} |X_t^\delta - X_t|^2 \right] \\ &\leq 3Z_1(T) + 3Z_2(T) + 3Z_3(T), \end{aligned} \tag{20.8}$$

where Z_1, Z_2 and Z_3 will be given explicitly below. We begin with the first term and estimate it analogously to the proof of Theorem 18.9, see (18.13), (18.14), and apply

the estimate (20.4).

$$\begin{aligned}
 Z_1(T) &:= \mathbb{E} \left[\sup_{t \leq T} \left| \int_0^{t_{n(t)}} \left(b(t_{n(s)}, X_{t_{n(s)}}^\delta) - b(t_{n(s)}, X_s) \right) ds \right. \right. \\
 &\quad \left. \left. + \int_0^{t_{n(t)}} \left(\sigma(t_{n(s)}, X_{t_{n(s)}}^\delta) - \sigma(t_{n(s)}, X_s) \right) dB_s \right|^2 \right] \\
 &\leq 2T \int_0^T \mathbb{E} [|b(t_{n(s)}, X_{t_{n(s)}}^\delta) - b(t_{n(s)}, X_s)|^2] ds \\
 &\quad + 8 \int_0^T \mathbb{E} [|\sigma(t_{n(s)}, X_{t_{n(s)}}^\delta) - \sigma(t_{n(s)}, X_s)|^2] ds \\
 &\leq (2TK^2 + 8K^2) \int_0^T \mathbb{E} [|X_{t_{n(s)}}^\delta - X_s|^2] ds \\
 &\leq C \int_0^T Z(s) ds.
 \end{aligned}$$

The second term in (20.8) can be estimated in a similar way.

$$\begin{aligned}
 Z_2(T) &:= \mathbb{E} \left[\sup_{t \leq T} \left| \int_0^{t_{n(t)}} \left(b(t_{n(s)}, X_s) - b(s, X_s) \right) ds \right. \right. \\
 &\quad \left. \left. + \int_0^{t_{n(t)}} \left(\sigma(t_{n(s)}, X_s) - \sigma(s, X_s) \right) dB_s \right|^2 \right] \\
 &\leq 2T \int_0^T \mathbb{E} [|b(t_{n(s)}, X_s) - b(s, X_s)|^2] ds \\
 &\quad + 8 \int_0^T \mathbb{E} [|\sigma(t_{n(s)}, X_s) - \sigma(s, X_s)|^2] ds \\
 &\leq (2TK^2 + 8K^2) \int_0^T (1 + \mathbb{E} [|X_s|])^2 (\sqrt{s - t_{n(s)}})^2 ds \leq C\delta,
 \end{aligned}$$

where we used (20.6) for the second inequality, and $|s - t_{n(s)}| < \delta$ and (18.15) for the last inequality. The terms Z_1 and Z_2 only consider the integral up to time $t_{n(t)}$, thus the third term takes care of the remainder up to time t . It can be estimated using (20.5), $|s - t_{n(s)}| < \delta$ and (18.15).

$$\begin{aligned}
 Z_3(T) &:= \mathbb{E} \left[\sup_{t \leq T} \left| \int_{t_{n(t)}}^t b(s, X_s) ds + \int_{t_{n(t)}}^t \sigma(s, X_s) dB_s \right|^2 \right] \\
 &\leq 2\delta \int_{t_{n(t)}}^t \mathbb{E} [|b(s, X_s)|^2] ds + 8 \int_{t_{n(t)}}^t \mathbb{E} [|\sigma(s, X_s)|^2] ds \leq C\delta.
 \end{aligned}$$

Note that all constants appearing in the estimates above may depend on T , but do not depend on δ . Thus we have shown

$$Z(T) \leq C_T \delta + C_T \int_0^T Z(s) ds,$$

where C_T is independent of δ . By Gronwall's lemma, Theorem A.43, we obtain that $Z(T) \leq C_T e^{TC_T} \delta$ and, finally,

$$\sup_{t \leq T} \mathbb{E} [|X_t^\delta - X_t|] \leq \sqrt{Z(T)} \leq c_T \sqrt{\delta}. \quad \square$$

In the theory of ordinary differential equations one obtains higher order approximation schemes with the help of Taylor's formula. Similarly for stochastic differential equations an application of a stochastic Taylor–Itô formula yields the following scheme, which converges with strong order 1 to the solution of (20.3), for a proof see [104]. There it is also shown that the weak order is 1, thus it is the same as for the Euler scheme.

20.25 Algorithm (Milstein scheme). Let b and σ satisfy the conditions of Algorithm 20.23 and $\sigma(t, \cdot) \in \mathcal{C}^2$. Let $T > 0$, $n \geq 1$, $0 = t_0 < t_1 < \dots < t_n = T$ be a time grid and x_0 be the starting position.

For $j = 1$ to n

1. Generate $y \stackrel{s}{\sim} \mathbf{N}(0, t_j - t_{j-1})$
2. Set $x_{t_j} := x_{t_{j-1}} + b(t_{j-1}, x_{t_{j-1}})(t_j - t_{j-1}) + \sigma(t_{j-1}, x_{t_{j-1}}) y$
 $\quad + \frac{1}{2} \sigma(t_{j-1}, x_{t_{j-1}}) \sigma^{(0,1)}(t_{j-1}, x_{t_{j-1}}) (y^2 - (t_j - t_{j-1}))$
 (where $\sigma^{(0,1)}(t, x) := \frac{\partial}{\partial x} \sigma(t, x)$)

Then $(x_0, x_{t_1}, \dots, x_{t_n}) \stackrel{s}{\sim} (X_0, X_{t_1}, \dots, X_{t_n})$ approximately.

In the literature one finds a variety of further algorithms for the numerical evaluation of SDEs. Some of them have higher orders of convergence than the Euler or Milstein scheme, see for example [104] and the references given therein. Depending on the actual simulation task one might be more interested in the weak or strong order of convergence. The strong order is important for pathwise approximations whereas the weak order is important if one is only interested in distributional quantities.

We close this chapter by one of the most important applications of samples: The Monte Carlo method.

20.6 Monte Carlo method

The Monte Carlo method is nothing but an application of the strong law of large numbers. Recall that for iid random variables $(Y_j)_{j \geq 1}$ with finite expectation

$$\frac{1}{n} \sum_{j=1}^n Y_j \xrightarrow{n \rightarrow \infty} \mathbb{E} Y_1 \quad \text{almost surely.}$$

Thus, if we are interested in the expectation of $g(X)$ where X is a random variable and g a measurable function, we can approximate $\mathbb{E} g(X)$ by the following algorithm.

20.26 Algorithm (Monte Carlo method). Let X be a random variable with distribution F and g be a function such that $\mathbb{E} |g(X)| < \infty$. Let $n \geq 1$.

1. Generate $x_j \stackrel{s}{\sim} F, j = 1, \dots, n$
2. Set $y := \frac{1}{n} \sum_{j=1}^n g(x_j)$

Then y is an approximation of $\mathbb{E} g(X)$.

The approximation error is, by the central limit theorem and the theorem of Berry–Esseen, proportional to $1/\sqrt{n}$.

20.27 Example. In applications one often defines a model by an SDE of the form (20.3) and one is interested in the value of $\mathbb{E} g(X_t)$, where g is some function, and $(X_t)_{t \geq 0}$ is the solution of the SDE. We know how to simulate the normal distributed increments of the underlying Brownian motion, e. g. by Algorithm 20.13. Hence we can simulate samples of X_t by the Euler scheme, Algorithm 20.23. Doing this repeatedly corresponds to step 1 in the Monte Carlo method. Therefore, we get by step 2 an approximation of $\mathbb{E} g(X_t)$.



20.28 Further reading. General monographs on simulation are [3] which focuses, in particular, on steady states, rare events, Gaussian processes and SDEs, and [162] which focuses in particular on Monte Carlo simulation, variance reduction and Markov chain Monte Carlo. More on the numerical treatment of SDEs can be found in [104]. A good starting point for the extension to Lévy processes and applications in finance is [30].

[3] Asmussen, Glynn: *Stochastic Simulation: Algorithms and Analysis*.

[30] Cont, Tankov: *Financial modelling with jump processes*.

[104] Kloeden, Platen: *Numerical Solution of Stochastic Differential Equations*.

[162] Rubinstein, Kroese: *Simulation and the Monte Carlo Method* Wiley.

Appendix

A.1 Kolmogorov's existence theorem

In this appendix we give a proof of Kolmogorov's theorem, Theorem 4.8, on the existence of stochastic processes with prescribed finite dimensional distributions. In addition to the notation introduced in Chapter 4, we use $\pi_J^L : (\mathbb{R}^d)^L \rightarrow (\mathbb{R}^d)^J$ for the projection onto the coordinates from $J \subset L \subset I$. Moreover, $\pi_J := \pi_J^I$ and we write \mathcal{H} for the family of all *finite* subsets of I . Finally

$$\mathcal{Z} := \left\{ \pi_J^{-1}(B) : J \in \mathcal{H}, B \in \mathcal{B}((\mathbb{R}^d)^J) \right\}$$

denotes the family of *cylinder sets* and we call $Z = \pi_J(B)$ a *J-cylinder* with *basis* B . If $J \in \mathcal{H}$ and $\#J = j$, we can identify $(\mathbb{R}^d)^J$ with \mathbb{R}^{jd} .

For the proof of Theorem 4.8 we need the following auxiliary result.

A.1 Lemma (Regularity). *Every probability measure μ on $(\mathbb{R}^n, \mathcal{B}(\mathbb{R}^n))$ is outer regular,*

$$\mu(B) = \inf \{ \mu(U) : U \supset B, U \subset \mathbb{R}^n \text{ open} \} \quad (\text{A.1})$$

and inner (compact) regular,

$$\begin{aligned} \mu(B) &= \sup \{ \mu(F) : F \subset B, F \text{ closed} \} \\ &= \sup \{ \mu(K) : K \subset B, K \text{ compact} \}. \end{aligned} \quad (\text{A.2})$$

Proof. Consider the family

$$\Sigma := \left\{ A \subset \mathbb{R}^n : \forall \epsilon > 0 \exists U \text{ open}, F \text{ closed}, F \subset A \subset U, \mu(U \setminus F) < \epsilon \right\}.$$

It is easy to see that Σ is a σ -algebra¹ which contains all closed sets. Indeed, if $F \subset \mathbb{R}^n$ is closed, the sets

$$U_n := F + \mathbb{B}(0, 1/n) = \{x + y : x \in F, y \in \mathbb{B}(0, 1/n)\}$$

¹ *Indeed:* $\emptyset \in \Sigma$ is obvious. Pick $A \in \Sigma$. For every $\epsilon > 0$ there are closed and open sets $F_\epsilon \subset A \subset U_\epsilon$ such that $\mu(U_\epsilon \setminus F_\epsilon) < \epsilon$. Observe that $U_\epsilon^c \subset A^c \subset F_\epsilon^c$, that U_ϵ^c is closed, that F_ϵ^c is open and that

$$F_\epsilon^c \setminus U_\epsilon^c = F_\epsilon^c \cap (U_\epsilon^c)^c = F_\epsilon^c \cap U_\epsilon = U_\epsilon \setminus F_\epsilon.$$

Thus, $\mu(F_\epsilon^c \setminus U_\epsilon^c) = \mu(U_\epsilon \setminus F_\epsilon) < \epsilon$, and we find that $A^c \in \Sigma$.

are open and $\bigcap_{n \geq 1} U_n = F$. By the continuity of measures, $\lim_{n \rightarrow \infty} \mu(U_n) = \mu(F)$, which shows that $F \subset U_n$ and $\mu(U_n \setminus F) < \epsilon$ for sufficiently large values of $n > N(\epsilon)$. This proves $F \in \Sigma$ and we conclude that

$$\mathcal{B}(\mathbb{R}^n) = \sigma(\text{closed subsets}) \subset \sigma(\Sigma) = \Sigma.$$

In particular, we can construct for every $B \in \mathcal{B}(\mathbb{R}^n)$ sequences of open $(U_n)_n$ and closed $(F_n)_n$ sets such that $F_n \subset B \subset U_n$ and $\lim_{n \rightarrow \infty} \mu(U_n \setminus F_n) = 0$. Therefore,

$$\mu(B \setminus F_n) + \mu(U_n \setminus B) \leq 2\mu(U_n \setminus F_n) \xrightarrow{n \rightarrow \infty} 0$$

which shows that the infimum in (A.1) and the supremum in the first equality of (A.2) are attained.

Finally, replace in the last step the closed sets F_n with the compact sets

$$K_n := F_n \cap \overline{\mathbb{B}}(0, R), \quad R = R(n).$$

Using the continuity of measures, we get $\lim_{R \rightarrow \infty} \mu(F_n \setminus F_n \cap \overline{\mathbb{B}}(0, R)) = 0$, and for sufficiently large $R = R(n)$ with $R(n) \rightarrow \infty$,

$$\mu(B \setminus K_n) \leq \mu(B \setminus F_n) + \mu(F_n \setminus K_n) \xrightarrow{n \rightarrow \infty} 0.$$

This proves the second equality in (A.2). \square

A.2 Theorem (Theorem 4.8. Kolmogorov 1933). *Let $I \subset [0, \infty)$ and $(p_J)_{J \in \mathcal{H}}$ be a projective family of probability measures p_J defined on $((\mathbb{R}^d)^J, \mathcal{B}((\mathbb{R}^d)^J))$, $J \in \mathcal{H}$. Then there exists a unique probability measure μ on the space $((\mathbb{R}^d)^I, \mathcal{B}^I)$ such that $p_J = \mu \circ \pi_J^{-1}$ where π_J is the canonical projection of $(\mathbb{R}^d)^I$ onto $(\mathbb{R}^d)^J$.*

The measure μ is often called the *projective limit* of the family $(p_J)_{J \in \mathcal{H}}$.

Let $A_n \in \Sigma$, $n \geq 1$. For $\epsilon > 0$ there exist sets $F_n \subset A_n \subset U_n$ satisfying $\mu(U_n \setminus F_n) < \epsilon/2^n$. Therefore,

$$\Phi_N := \overbrace{F_1 \cup \dots \cup F_N}^{\text{closed}} \subset A_1 \cup \dots \cup A_N \subset U := \overbrace{\bigcup_{j \geq 1} U_j}^{\text{open}}.$$

Since

$$\bigcap_{N \geq 1} U \setminus \Phi_N = \bigcup_{j \geq 1} U_j \setminus \bigcup_{k \geq 1} F_k = \bigcup_{j \geq 1} \left(U_j \setminus \bigcup_{k \geq 1} F_k \right) \subset \bigcup_{j \geq 1} (U_j \setminus F_j)$$

we find by the continuity and the σ -subadditivity of the (finite) measure μ

$$\lim_{N \rightarrow \infty} \mu(U \setminus \Phi_N) = \mu\left(\bigcup_{j \geq 1} U_j \setminus \bigcup_{k \geq 1} F_k\right) \leq \sum_{j \geq 1} \mu(U_j \setminus F_j) \leq \sum_{j \geq 1} \frac{\epsilon}{2^j} = \epsilon.$$

Thus, $\mu(U \setminus \Phi_N) \leq 2\epsilon$ for all sufficiently large $N > N(\epsilon)$. This proves $\bigcup_{n \geq 1} A_n \in \Sigma$.

Proof. We use the family $(p_J)_{J \in \mathcal{H}}$ to define a (pre-)measure on the cylinder sets \mathcal{Z} ,

$$\mu \circ \underbrace{\pi_J^{-1}}_{\in \mathcal{Z}}(B) := p_J(B) \quad \forall B \in \mathcal{B}^J. \quad (\text{A.3})$$

Since \mathcal{Z} is an algebra of sets, i. e. $\emptyset \in \mathcal{Z}$ and \mathcal{Z} is stable under finite unions, intersections and complements, we can use Carathéodory's extension theorem to extend μ onto $\mathcal{B}^I = \sigma(\mathcal{Z})$.

Let us, first of all, check that μ is well-defined and that it does not depend on the particular representation of the cylinder set Z . Assume that $Z = \pi_J^{-1}(B) = \pi_K^{-1}(B')$ for $J, K \in \mathcal{H}$ and $B \in \mathcal{B}^J$, $B' \in \mathcal{B}^K$. We have to show that

$$p_J(B) = p_K(B').$$

Set $L := J \cup K \in \mathcal{H}$. Then

$$\begin{aligned} Z &= \pi_J^{-1}(B) = (\pi_J^L \circ \pi_L)^{-1}(B) = \pi_L^{-1} \circ (\pi_J^L)^{-1}(B), \\ Z &= \pi_K^{-1}(B') = (\pi_K^L \circ \pi_L)^{-1}(B') = \pi_L^{-1} \circ (\pi_K^L)^{-1}(B'). \end{aligned}$$

Thus, $(\pi_J^L)^{-1}(B) = (\pi_K^L)^{-1}(B')$ and the projectivity (4.6) of the family $(p_J)_{J \in \mathcal{H}}$ gives

$$p_J(B) = \pi_J^L(p_L)(B) = p_L((\pi_J^L)^{-1}(B)) = p_L((\pi_K^L)^{-1}(B')) = p_K(B').$$

Next we verify that μ is an additive set function: $\mu(\emptyset) = 0$ is obvious. Let $Z, Z' \in \mathcal{Z}$ such that $Z = \pi_J^{-1}(B)$ and $Z' = \pi_K^{-1}(B')$. Setting $L := J \cup K$ we can find suitable basis sets $A, A' \in \mathcal{B}^L$ such that

$$Z = \pi_L^{-1}(A) \quad \text{and} \quad Z' = \pi_L^{-1}(A').$$

If $Z \cap Z' = \emptyset$ we see that $A \cap A' = \emptyset$, and μ inherits its additivity from p_L :

$$\begin{aligned} \mu(Z \cup Z') &= \mu(\pi_L^{-1}(A) \cup \pi_L^{-1}(A')) = \mu(\pi_L^{-1}(A \cup A')) \\ &= p_L(A \cup A') = p_L(A) + p_L(A') \\ &= \dots = \mu(Z) + \mu(Z'). \end{aligned}$$

Finally, we prove that μ is σ -additive on \mathcal{Z} . For this we show that μ is continuous at the empty set, i. e. for every decreasing sequence of cylinder sets $(Z_n)_{n \geq 1} \subset \mathcal{Z}$, $\bigcap_{n \geq 1} Z_n = \emptyset$ we have $\lim_{n \rightarrow \infty} \mu(Z_n) = 0$. Equivalently,

$$Z_n \in \mathcal{Z}, \quad Z_n \supset Z_{n+1} \supset \dots, \quad \mu(Z_n) \geq \alpha > 0 \implies \bigcap_{n \geq 1} Z_n \neq \emptyset.$$

Write for such a sequence of cylinder sets $Z_n = \pi_{J_n}^{-1}(B_n)$ with suitable $J_n \in \mathcal{H}$ and $B_n \in \mathcal{B}^{J_n}$. Since for every $J \subset K$ any J -cylinder is also a K -cylinder, we may assume without loss of generality that $J_n \subset J_{n+1}$.

Note that $(\mathbb{R}^d)^{J_n} \cong (\mathbb{R}^d)^{\#J_n}$; therefore we can use Lemma A.1 to see that p_{J_n} is inner compact regular. Thus,

$$\text{there exist compact sets } K_n \subset B_n : \quad p_{J_n}(B_n \setminus K_n) \leq 2^{-n}\alpha.$$

Write $Z'_n := \pi_{J_n}^{-1}(K_n)$ and observe that

$$\mu(Z_n \setminus Z'_n) = \mu(\pi_{J_n}^{-1}(B_n \setminus K_n)) = p_{J_n}(B_n \setminus K_n) \leq 2^{-n}\alpha.$$

Although the Z_n are decreasing, this might not be the case for the Z'_n ; therefore we consider $\widehat{Z}_n := Z'_1 \cap \dots \cap Z'_n \subset Z_n$ and find

$$\begin{aligned} \mu(\widehat{Z}_n) &= \mu(Z_n) - \mu(Z_n \setminus \widehat{Z}_n) \\ &= \mu(Z_n) - \mu\left(\bigcup_1^n \underbrace{(Z_n \setminus Z'_k)}_{\subset Z_k \setminus Z'_k}\right) \\ &\geq \mu(Z_n) - \sum_1^n \mu(Z_k \setminus Z'_k) \\ &\geq \alpha - \sum_1^n 2^{-k}\alpha > 0. \end{aligned}$$

We are done if we can show that $\bigcap_{n \geq 1} Z'_n = \bigcap_{n \geq 1} \widehat{Z}_n \neq \emptyset$. For this we construct some element $z \in \bigcap_{n \geq 1} \widehat{Z}_n$.

1° For every $m \geq 1$ we pick some $f_m \in \widehat{Z}_m$. Since the sequence $(\widehat{Z}_n)_{n \geq 1}$ is decreasing,

$$f_m \in \widehat{Z}_n \subset Z'_n, \text{ hence } \pi_{J_n}(f_m) \in K_n \text{ for all } m \geq n \text{ and } n \geq 1.$$

Since projections are continuous, and since continuous maps preserve compactness,

$$\pi_t(f_m) = \pi_t^{J_n} \circ \pi_{J_n}(f_m) \in \underbrace{\pi_t^{J_n}(K_n)}_{\text{compact}} \text{ for all } m \geq n, n \geq 1 \text{ and } t \in J_n.$$

2° The index set $H := \bigcup_{n \geq 1} J_n$ is countable; denote by $H = (t_k)_{k \geq 1}$ some enumeration. For each $t_k \in H$ there is some $n(k) \geq 1$ such that $t_k \in J_{n(k)}$. Moreover, by 1°,

$$(\pi_{t_k}(f_m))_{m \geq n(k)} \subset \pi_{t_k}^{J_{n(k)}}(K_{n(k)}).$$

Since the sets $\pi_{t_k}^{J_{n(k)}}(K_{n(k)})$ are compact, we can take repeatedly subsequences,

$$\begin{aligned} (\pi_{t_1}(f_m))_{m \geq n(1)} \subset \text{compact set} &\implies \exists (f_m^1)_m \subset (f_m)_m : \exists \lim_{m \rightarrow \infty} \pi_{t_1}(f_m^1) \\ (\pi_{t_2}(f_m^1))_{m \geq n(2)} \subset \text{compact set} &\implies \exists (f_m^2)_m \subset (f_m^1)_m : \exists \lim_{m \rightarrow \infty} \pi_{t_2}(f_m^2) \\ &\vdots \\ (\pi_{t_k}(f_m^{k-1}))_{m \geq n(k)} \subset \text{compact set} &\implies \exists (f_m^k)_m \subset (f_m^{k-1})_m : \exists \lim_{m \rightarrow \infty} \pi_{t_k}(f_m^k). \end{aligned}$$

The diagonal sequence $(f_m^m)_{m \geq 1}$ satisfies

$$\exists \lim_{m \rightarrow \infty} \pi_t(f_m^m) = z_t \text{ for each } t \in H.$$

3° By construction, $\pi_t(f_m) \in \pi_t^{J_n}(K_n)$ for all $t \in J_n$ and $m \geq n$. Therefore,

$$\pi_t(f_m^m) \in \pi_t^{J_n}(K_n) \text{ for all } m \geq n \text{ and } t \in J_n.$$

Since K_n is compact, $\pi_t^{J_n}(K_n)$ is compact, hence closed, and we see that

$$z_t = \lim_m \pi_t(f_m^m) \in \pi_t^{J_n}(K_n) \text{ for every } t \in J_n.$$

Using the fact that $n \geq 1$ is arbitrary, find

$$z(t) := \begin{cases} z_t & \text{if } t \in H; \\ * & \text{if } t \in I \setminus H \end{cases}$$

(* stands for any element from E) defines an element $z = (z(t))_{t \in I} \in (\mathbb{R}^d)^I$ with the property that for all $n \geq 1$

$$\pi_{J_n}(z) \in K_n \text{ if, and only if, } z \in \pi_{J_n}^{-1}(K_n) = Z'_n \supset \widehat{Z}_n.$$

This proves $z \in \bigcap_{n \geq 1} Z'_n = \bigcap_{n \geq 1} \widehat{Z}_n$ and $\bigcap_{n \geq 1} \widehat{Z}_n \neq \emptyset$. □

A.2 A property of conditional expectations

We assume that you have some basic knowledge of conditioning and conditional expectations with respect to a σ -algebra. Several times we will use the following result which is not always contained in elementary expositions. Throughout this section let $(\Omega, \mathcal{A}, \mathbb{P})$ be a probability space.

A.3 Lemma. *Let $X : (\Omega, \mathcal{A}) \rightarrow (D, \mathcal{D})$ and $Y : (\Omega, \mathcal{A}) \rightarrow (E, \mathcal{E})$ be two random variables. Assume that $\mathcal{X}, \mathcal{Y} \subset \mathcal{A}$ are σ -algebras such that X is \mathcal{X}/\mathcal{D} measurable, Y*

is \mathcal{Y}/\mathcal{E} measurable and $\mathcal{X} \perp\!\!\!\perp \mathcal{Y}$. Then

$$\mathbb{E}[\Phi(X, Y) \mid \mathcal{X}] = [\mathbb{E} \Phi(x, Y)] \Big|_{x=X} = \mathbb{E}[\Phi(X, Y) \mid X] \quad (\text{A.4})$$

holds for all bounded $\mathcal{D} \times \mathcal{E}/\mathcal{B}(\mathbb{R})$ measurable functions $\Phi : D \times E \rightarrow \mathbb{R}$.

If $\Psi : D \times \Omega \rightarrow \mathbb{R}$ is bounded and $\mathcal{D} \otimes \mathcal{Y}/\mathcal{B}(\mathbb{R})$ measurable, then

$$\mathbb{E}[\Psi(X(\cdot), \cdot) \mid \mathcal{X}] = [\mathbb{E} \Psi(x, \cdot)] \Big|_{x=X} = \mathbb{E}[\Psi(X(\cdot), \cdot) \mid X]. \quad (\text{A.5})$$

Proof. Assume first that $\Phi(x, y)$ is of the form $\phi(x)\psi(y)$. Then

$$\begin{aligned} \mathbb{E}[\underbrace{\phi(X)\psi(Y)}_{=\Phi(X,Y)} \mid \mathcal{X}] &= \phi(X) \mathbb{E}[\psi(Y) \mid \mathcal{X}] \stackrel{Y \perp\!\!\!\perp X}{=} \phi(X) \mathbb{E}[\psi(Y)] \\ &= \mathbb{E}[\underbrace{\phi(x)\psi(Y)}_{=\Phi(x,Y)}] \Big|_{x=X}. \end{aligned}$$

Now fix some $F \in \mathcal{X}$ and pick $\phi(x) = \mathbb{1}_A(x)$ and $\psi(y) = \mathbb{1}_B(y)$, $A \in \mathcal{D}$, $B \in \mathcal{E}$. Our argument shows that

$$\int_F \mathbb{1}_{A \times B}(X(\omega), Y(\omega)) \mathbb{P}(d\omega) = \int_F \mathbb{E} \mathbb{1}_{A \times B}(x, Y) \Big|_{x=X(\omega)} \mathbb{P}(d\omega), \quad F \in \mathcal{X}.$$

Both sides of this equality (can be extended from $\mathcal{D} \times \mathcal{E}$ to) define measures on the product σ -algebra $\mathcal{D} \otimes \mathcal{E}$. By linearity, this becomes

$$\int_F \Phi(X(\omega), Y(\omega)) \mathbb{P}(d\omega) = \int_F \mathbb{E} \Phi(x, Y) \Big|_{x=X(\omega)} \mathbb{P}(d\omega), \quad F \in \mathcal{X},$$

for $\mathcal{D} \otimes \mathcal{E}$ measurable positive step functions Φ ; using standard arguments from the theory of measure and integration, we get this equality for positive measurable functions and then for all bounded measurable functions. The latter is, however, equivalent to the first equality in (A.4).

The second equality follows if we take conditional expectations $\mathbb{E}(\cdots \mid X)$ on both sides of the first equality and use the tower property of conditional expectation.

The formulae (A.5) follow in a similar way. \square

A simple consequence of Lemma A.3 are the following formulae.

A.4 Corollary. Let $X : (\Omega, \mathcal{A}) \rightarrow (D, \mathcal{D})$ and $Y : (\Omega, \mathcal{A}) \rightarrow (E, \mathcal{E})$ be two random variables. Assume that $\mathcal{X}, \mathcal{Y} \subset \mathcal{A}$ are σ -algebras such that X is \mathcal{X}/\mathcal{D} measurable, Y is \mathcal{Y}/\mathcal{E} measurable and $\mathcal{X} \perp\!\!\!\perp \mathcal{Y}$. Then

$$\mathbb{E} \Phi(X, Y) = \int_D \mathbb{E} \Phi(x, Y) \mathbb{P}(X \in dx) = \mathbb{E} \left[\int_D \Phi(x, Y) \mathbb{P}(X \in dx) \right] \quad (\text{A.6})$$

holds for all bounded $\mathcal{D} \times \mathcal{E}/\mathcal{B}(\mathbb{R})$ measurable functions $\Phi : D \times E \rightarrow \mathbb{R}$.

If $\Psi : D \times \Omega \rightarrow \mathbb{R}$ is bounded and $\mathcal{D} \otimes \mathcal{Y}/\mathcal{B}(\mathbb{R})$ measurable, then

$$\mathbb{E} \Psi(X(\cdot), \cdot) = \int_D \mathbb{E} \Psi(x, \cdot) \mathbb{P}(X \in dx) = \mathbb{E} \left[\int_D \Psi(x, \cdot) \mathbb{P}(X \in dx) \right]. \quad (\text{A.7})$$

A.3 From discrete to continuous time martingales

We assume that you have some basic knowledge of discrete time martingales, e. g. as in [169]. The key to the continuous time setting is the following simple observation: If $(X_t, \mathcal{F}_t)_{t \geq 0}$ is a martingale, then

$$(X_{t_j}, \mathcal{F}_{t_j})_{j \geq 1} \text{ is a martingale for any sequence } 0 \leq t_1 < t_2 < t_3 < \dots$$

We are mainly interested in estimates and convergence results for martingales. Our strategy is to transfer the corresponding results from the discrete time setting to continuous time. The key result is *Doob's upcrossing estimate*. Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a real-valued martingale, $I \subset [0, \infty)$ be a finite index set with $t_{\min} = \min I$, $t_{\max} = \max I$, and $a < b$. Then

$$U(I; [a, b]) = \max \left\{ m \mid \begin{array}{l} \exists \sigma_1 < \tau_1 < \sigma_2 < \tau_2 < \dots < \sigma_m < \tau_m, \\ \sigma_j, \tau_j \in I, X(\sigma_j) < a < b < X(\tau_j) \end{array} \right\}$$

is the number of *upcrossings* of $(X_s)_{s \in I}$ across the strip $[a, b]$. Figure A.1 shows two upcrossings.

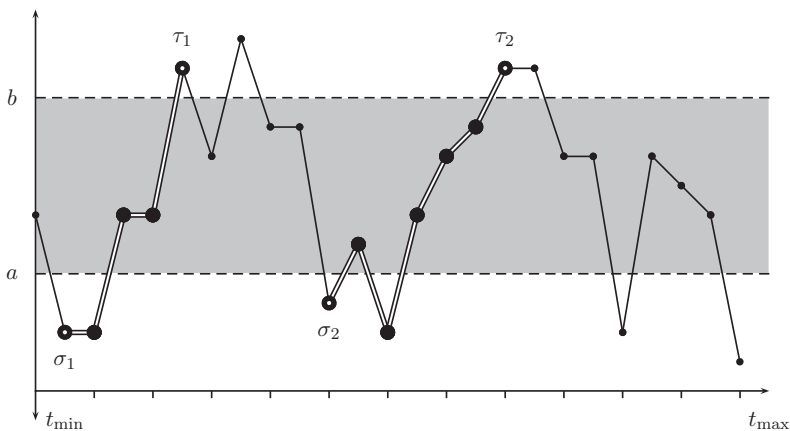


Figure A.1. Two upcrossings over the interval $[a, b]$.

Since $(X_s, \mathcal{F}_s)_{s \in I}$ is a discrete time martingale, we know

$$\mathbb{E}[U(I; [a, b])] \leq \frac{1}{b-a} \mathbb{E}[(X(t_{\max}) - a)^+]. \quad (\text{A.8})$$

A.5 Lemma (Upcrossing estimate). *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a real-valued submartingale and $D \subset [0, \infty)$ a countable subset. Then*

$$\mathbb{E}[U(D \cap [0, t]; [a, b])] \leq \frac{1}{b-a} \mathbb{E}[(X(t) - a)^+] \quad (\text{A.9})$$

for all $t > 0$ and $-\infty < a < b < \infty$.

Proof. Let $D_N \subset D$ be finite index sets such that $D_N \uparrow D$. If $t_{\max} := \max D_N \cap [0, t]$, we get from (A.8) with $I = D_N$

$$\mathbb{E}[U(D_N \cap [0, t]; [a, b])] \leq \frac{1}{b-a} \mathbb{E}[(X(t_{\max}) - a)^+] \leq \frac{1}{b-a} \mathbb{E}[(X(t) - a)^+].$$

For the last estimate observe that $(X_t - a)^+$ is a submartingale since $x \mapsto x^+$ is convex and increasing. Since $\sup_N U(D_N \cap [0, t]; [a, b]) = U(D \cap [0, t]; [a, b])$, (A.9) follows from monotone convergence. \square

A.6 Theorem (Martingale convergence theorem. Doob 1953). *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a martingale in \mathbb{R}^d or a submartingale in \mathbb{R} with continuous sample paths. If $\sup_{t \geq 0} \mathbb{E}|X_t| < \infty$, then the limit $\lim_{t \rightarrow \infty} X_t$ exists almost surely and defines a finite \mathcal{F}_∞ measurable random variable X_∞ .*

If $(X_t, \mathcal{F}_t)_{t \geq 0}$ is a martingale of the form $X_t = \mathbb{E}(Z | \mathcal{F}_t)$ for some $Z \in L^1(\mathbb{P})$, then we have $X_\infty = \mathbb{E}(Z | \mathcal{F}_\infty)$ and $\lim_{t \rightarrow \infty} X_t = X_\infty$ almost surely and in $L^1(\mathbb{P})$.

Proof. Since a d -dimensional function converges if, and only if, all coordinate functions converge, we may assume that X_t is real valued. By Doob's upcrossing estimate

$$\mathbb{E}[U(Q^+; [a, b])] \leq \frac{\sup_{t \geq 0} \mathbb{E}[(X_t - a)^+]}{b-a} < \infty$$

for all intervals $[a, b]$ and $Q^+ := Q \cap [0, \infty)$. Thus,

$$U(Q^+; [a, b])(\omega) < \infty \quad \text{for all } \omega \notin N_{a,b}$$

where $N_{a,b} \subset \Omega$ with $\mathbb{P}(N_{a,b}) = 0$. Set $N := \bigcup_{a < b, a, b \in \mathbb{Q}^+} N_{a,b}$ and $\Omega_0 := \Omega \setminus N$. Then $\mathbb{P}(\Omega_0) = 1$.

We claim that for $\omega \in \Omega_0$ the limit $\lim_{t \rightarrow \infty} X_t(\omega)$ exists. Otherwise we could find $\alpha, \beta \in \mathbb{Q}$ such that

$$\liminf_{s \rightarrow \infty} X_s(\omega) < \alpha < \beta < \overline{\lim}_{s \rightarrow \infty} X_s(\omega).$$

By the continuity of the sample paths,

$$\lim_{s \rightarrow \infty} X_s(\omega) = \lim_{Q^+ \ni r \rightarrow \infty} X_r(\omega) \quad \text{and} \quad \overline{\lim}_{s \rightarrow \infty} X_s(\omega) = \overline{\lim}_{Q^+ \ni r \rightarrow \infty} X_r(\omega),$$

which means that there must be infinitely many upcrossings of $(X_t)_{t \in Q^+}$ over the interval $[\alpha, \beta]$; this is impossible since $U(Q^+; [\alpha, \beta])(\omega) < \infty$.

Set $X_\infty := \lim_{t \rightarrow \infty} X_t \mathbb{1}_{\Omega_0}$. By Fatou's Lemma,

$$\mathbb{E} |X_\infty| = \mathbb{E} \left[\lim_{t \rightarrow \infty} |X_t| \right] \leq \lim_{t \rightarrow \infty} \mathbb{E} |X_t| \leq \sup_{t \geq 0} \mathbb{E} |X_t| < \infty$$

which shows that $X_\infty \in L^1$ and that X_∞ is a.s. finite.

Assume now that $X_t = \mathbb{E}(Z | \mathcal{F}_t)$. From

$$|X_t| = |\mathbb{E}(Z | \mathcal{F}_t)| \leq \mathbb{E}(|Z| | \mathcal{F}_t)$$

we conclude that $\sup_{t \geq 0} \mathbb{E} |X_t| \leq \mathbb{E} |Z| < \infty$ and that $(X_t)_{t \geq 0}$ is uniformly integrable. By the L^1 -convergence theorem for discrete uniformly integrable martingales we find that for any sequence $t_j \uparrow \infty$ the limit $\lim_{j \rightarrow \infty} X_{t_j} = X'_\infty$ exists in L^1 and almost surely. By the uniqueness of a.s. limits we conclude that $X_\infty = X'_\infty$ almost surely.

Fix $t > 0$. For all $t_j \geq t$ and $F \in \mathcal{F}_t \subset \mathcal{F}_{t_j}$ we find, because of L^1 -convergence,

$$\int_F Z d\mathbb{P} \stackrel{X_{t_j} = \mathbb{E}(Z | \mathcal{F}_{t_j})}{=} \int_F X_{t_j} d\mathbb{P} \xrightarrow{j \rightarrow \infty} \int_F X_\infty d\mathbb{P},$$

i. e.

$$\int_F Z d\mathbb{P} = \int_F X_\infty d\mathbb{P} \quad \text{for all } F \in \bigcup_{t \geq 0} \mathcal{F}_t.$$

Since $\bigcup_{t \geq 0} \mathcal{F}_t$ is a \cap -stable generator of \mathcal{F}_∞ , we see $\mathbb{E}(Z | \mathcal{F}_\infty) = X_\infty$. \square

The next theorem allows us to identify all martingales of the form $X_t = \mathbb{E}(Z | \mathcal{F}_t)$.

A.7 Theorem (Closure of a martingale. Doob 1953). *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a martingale in \mathbb{R}^d with continuous paths. Then the following assertions are equivalent*

- a) $(X_t)_{t \geq 0}$ is uniformly integrable;
- b) $X_\infty = \lim_{t \rightarrow \infty} X_t$ exists a.s. and in L^1 and $(X_t, \mathcal{F}_t)_{t \in [0, \infty]}$ is a martingale.

Proof. a) \Rightarrow b): Since $(X_t)_{t \geq 0}$ is uniformly integrable, we have $\sup_{t \geq 0} \mathbb{E} |X_t| < \infty$. Therefore $\lim_{t \rightarrow \infty} X_t = X_\infty$ almost surely, and by Vitali's convergence theorem we find that $\lim_{j \rightarrow \infty} X_{t_j} = X_\infty$ in L^1 for all increasing sequences $t_j \uparrow \infty$. Thus, the limit $\lim_{t \rightarrow \infty} X_t = X_\infty$ exists in L^1 , and we find for all $F \in \mathcal{F}_t$ and $T \geq t$

$$\int_F X_\infty d\mathbb{P} \stackrel{L^1\text{-conv.}}{=} \lim_{T \rightarrow \infty} \int_F X_T d\mathbb{P} \stackrel{F \in \mathcal{F}_t}{=} \int_F X_t d\mathbb{P};$$

hence, $\mathbb{E}(X_\infty | \mathcal{F}_t) = X_t$ and $(X_t)_{t \in [0, \infty]}$ is a martingale.

b) \Rightarrow a): Since $X_t = \mathbb{E}(X_\infty | \mathcal{F}_t)$ for $X_\infty \in L^1(\mathbb{P})$, we get $|X_t| \leq \mathbb{E}(|X_\infty| | \mathcal{F}_t)$ and, therefore, $\mathbb{E} |X_t| \leq \mathbb{E} |X_\infty|$. Hence, we find for all $t \geq 0$, $R > 0$ and $K > 0$

$$\begin{aligned}
 \int_{\{|X_t| > R\}} |X_t| d\mathbb{P} &\leq \int_{\{|X_t| > R\}} \mathbb{E}(|X_\infty| | \mathcal{F}_t) d\mathbb{P} = \int_{\{|X_t| > R\}} |X_\infty| d\mathbb{P} \\
 &= \int_{\{|X_t| > R\} \cap \{|X_\infty| > K\}} |X_\infty| d\mathbb{P} + \int_{\{|X_t| > R\} \cap \{|X_\infty| \leq K\}} |X_\infty| d\mathbb{P} \\
 &\leq \int_{\{|X_\infty| > K\}} |X_\infty| d\mathbb{P} + K \mathbb{P}(|X_t| > R) \\
 &\leq \int_{\{|X_\infty| > K\}} |X_\infty| d\mathbb{P} + \frac{K}{R} \mathbb{E} |X_t| \\
 &\leq \int_{\{|X_\infty| > K\}} |X_\infty| d\mathbb{P} + \frac{K}{R} \mathbb{E} |X_\infty| \\
 &\xrightarrow{R \rightarrow \infty} \int_{\{|X_\infty| > K\}} |X_\infty| d\mathbb{P} \xrightarrow[K \rightarrow \infty]{\text{dom. conv.}} 0. \quad \square
 \end{aligned}$$

For backwards submartingales, i. e. for submartingales with a decreasing index set, the assumptions of the convergence theorems (A.6) and (A.7) are always satisfied.

A.8 Corollary (Backwards convergence theorem). *Let (X_t, \mathcal{F}_t) be a submartingale and let $t_1 \geq t_2 \geq t_3 \geq \dots$, $t_j \downarrow t_\infty$, be a decreasing sequence. Then $(X_{t_j}, \mathcal{F}_{t_j})_{j \geq 1}$ is a (backwards) submartingale which is uniformly integrable. In particular, the limit $\lim_{j \rightarrow \infty} X_{t_j}$ exists in L^1 and almost surely.*

Proof. That $(X_{t_j})_{j \geq 1}$ is a (backwards directed) submartingale is obvious. Now we use that $\mathbb{E} X_0 \leq \mathbb{E} X_t$ and that $(X_t^+)_{t \geq 0}$ is a submartingale. Therefore, we get for $t \leq t_0$

$$\mathbb{E} |X_t| = 2 \mathbb{E} X_t^+ - \mathbb{E} X_t \leq 2 \mathbb{E} X_t^+ - \mathbb{E} X_0 \leq 2 \mathbb{E} X_{t_1}^+ - \mathbb{E} X_0$$

which shows that $\sup_{j \geq 1} \mathbb{E} |X_{t_j}| < \infty$. By the martingale convergence theorem we get

$$X_{t_j} \xrightarrow{j \rightarrow \infty} Z$$

almost surely for some \mathcal{F}_{t_∞} measurable random variable $Z \in L^1(\mathbb{P})$. By the submartingale property, $\mathbb{E} X_{t_j}$ is decreasing, i. e. the limit $\lim_{j \rightarrow \infty} \mathbb{E} X_{t_j} = \inf_{j \geq 1} \mathbb{E} X_{t_j}$ exists (and is finite). Thus,

$$\forall \epsilon > 0 \exists N = N_\epsilon \forall j \geq N : |\mathbb{E} X_{t_j} - \mathbb{E} X_{t_N}| < \epsilon.$$

This means that we get for all $j \geq N$, i. e. $t_j \leq t_N$:

$$\begin{aligned} \int_{\{|X_{t_j}| \geq R\}} |X_{t_j}| d\mathbb{P} &= 2 \int_{\{|X_{t_j}| \geq R\}} X_{t_j}^+ d\mathbb{P} - \int_{\{|X_{t_j}| \geq R\}} X_{t_j} d\mathbb{P} \\ &= 2 \int_{\{|X_{t_j}| \geq R\}} X_{t_j}^+ d\mathbb{P} - \mathbb{E} X_{t_j} + \int_{\{|X_{t_j}| < R\}} X_{t_j} d\mathbb{P} \end{aligned}$$

and since X_t and X_t^+ are submartingales

$$\begin{aligned} &\leq 2 \int_{\{|X_{t_j}| \geq R\}} X_{t_N}^+ d\mathbb{P} - \mathbb{E} X_{t_j} + \int_{\{|X_{t_j}| < R\}} X_{t_N} d\mathbb{P} \\ &\leq 2 \int_{\{|X_{t_j}| \geq R\}} X_{t_N}^+ d\mathbb{P} + (\epsilon - \mathbb{E} X_{t_N}) + \int_{\{|X_{t_j}| < R\}} X_{t_N} d\mathbb{P} \\ &= 2 \int_{\{|X_{t_j}| \geq R\}} X_{t_N}^+ d\mathbb{P} + \epsilon - \int_{\{|X_{t_j}| \geq R\}} X_{t_N} d\mathbb{P} \\ &= \int_{\{|X_{t_j}| \geq R\}} |X_{t_N}| d\mathbb{P} + \epsilon. \end{aligned}$$

Thus,

$$\begin{aligned} &\int_{\{|X_{t_j}| \geq R\}} |X_{t_j}| d\mathbb{P} \\ &\leq \int_{\{|X_{t_j}| \geq R\}} |X_{t_N}| d\mathbb{P} + \epsilon \\ &\leq \underbrace{\int_{\{|X_{t_j}| \geq R\} \cap \{|X_{t_N}| \leq R/2\}} |X_{t_N}| d\mathbb{P}}_{\text{note: } |X_{t_N}| \leq R/2 \leq |X_{t_j}|/2} + \int_{\{|X_{t_j}| \geq R\} \cap \{|X_{t_N}| \geq R/2\}} |X_{t_N}| d\mathbb{P} + \epsilon \\ &\leq \frac{1}{2} \int_{\{|X_{t_j}| \geq R\} \cap \{|X_{t_N}| \leq R/2\}} |X_{t_j}| d\mathbb{P} + \int_{\{|X_{t_N}| \geq R/2\}} |X_{t_N}| d\mathbb{P} + \epsilon \\ &\leq \frac{1}{2} \int_{\{|X_{t_j}| \geq R\}} |X_{t_j}| d\mathbb{P} + \int_{\{|X_{t_N}| \geq R/2\}} |X_{t_N}| d\mathbb{P} + \epsilon. \end{aligned}$$

This shows that

$$\int_{\{|X_{t_j}| \geq R\}} |X_{t_j}| d\mathbb{P} \leq 2 \int_{\{|X_{t_N}| \geq R/2\}} |X_{t_N}| d\mathbb{P} + 2\epsilon \xrightarrow[R \rightarrow \infty]{\text{dom. conv.}} 2\epsilon \xrightarrow[\epsilon \rightarrow 0]{} 0$$

holds uniformly for all $j \geq N$. Therefore, $(X_{t_j})_{j \geq 1}$ is uniformly integrable. By Vitali's convergence theorem we conclude that $L^1\text{-}\lim_{j \rightarrow \infty} X_{t_j}$ exists. \square

We will also need some maximal inequalities for martingales. Recall that for a real-valued discrete time submartingale $(X(i), \mathcal{F}_i)_{i \in I}$ where $t_{\min} = \min I$, $t_{\max} = \max I$ and $r > 0$ the following estimates hold

$$\mathbb{P} \left(\max_{s \in I} X(s) \geq r \right) \leq \frac{1}{r} \mathbb{E} [X^+(t_{\max})], \quad (\text{A.10})$$

$$\mathbb{P} \left(\min_{s \in I} X(s) \leq -r \right) \leq \frac{1}{r} \left(\mathbb{E} [X^+(t_{\max})] - \mathbb{E} [X(t_{\min})] \right). \quad (\text{A.11})$$

Since for every d -dimensional martingale or positive submartingale $(X(t), \mathcal{F}_t)_{t \geq 0}$ such that $X(t) \in L^p(\mathbb{P})$, $p \geq 1$, the process $(|X(t)|^p, \mathcal{F}_t)_{t \geq 0}$ is a submartingale, we get

$$\mathbb{P} \left(\max_{s \in I} |X(s)| \geq r \right) \leq \frac{1}{r^p} \mathbb{E} [|X(t_{\max})|^p] \quad \text{for all } r > 0. \quad (\text{A.12})$$

A.9 Lemma (Doob 1953). *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a positive submartingale or a d -dimensional martingale, $X_t \in L^p(\mathbb{P})$ for some $p \geq 1$ and $D \subset [0, \infty)$ a countable subset. Then*

$$\mathbb{P} \left(\sup_{s \in D \cap [0, t]} |X_s| \geq r \right) \leq \frac{1}{r^p} \mathbb{E} [|X_t|^p] \quad \text{for all } r > 0, t \geq 0. \quad (\text{A.13})$$

Proof. Let $D_N \subset D$ be finite sets such that $D_N \uparrow D$. We use (A.12) with the index set $I = D_N \cap [0, t]$, $t_{\max} = \max D_N \cap [0, t]$ and observe that for the submartingale $(|X_t|^p, \mathcal{F}_t)_{t \geq 0}$ the estimate

$$\mathbb{E} [|X(t_{\max})|^p] \leq \mathbb{E} [|X(t)|^p]$$

holds. Since

$$\left\{ \max_{s \in D_N \cap [0, t]} |X(s)| \geq r \right\} \uparrow \left\{ \sup_{s \in D \cap [0, t]} |X(s)| \geq r \right\}$$

we obtain (A.13) by using (A.12) and the continuity of measures. \square

Finally we can deduce Doob's maximal inequality.

A.10 Theorem (L^p maximal inequality. Doob 1953). *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a positive submartingale or a d -dimensional martingale, $X_t \in L^p(\mathbb{P})$ for some $p > 1$ and $D \subset [0, \infty)$ a countable subset. Then for $p^{-1} + q^{-1} = 1$ and all $t > 0$*

$$\mathbb{E} \left[\sup_{s \in D \cap [0, t]} |X_s|^p \right] \leq q^p \sup_{s \in [0, t]} \mathbb{E} [|X_s|^p] \leq q^p \mathbb{E} [|X_t|^p]. \quad (\text{A.14})$$

If $(X_t)_{t \geq 0}$ is continuous, (A.14) holds for $D = [0, \infty)$.

Proof. Let $D_N \subset D$ be finite index sets with $D_N \uparrow D$. The first estimate follows immediately from the discrete version of Doob's inequality,

$$\mathbb{E} \left[\sup_{s \in D_N \cap [0, t]} |X_s|^p \right] \leq q^p \sup_{s \in D_N \cap [0, t]} \mathbb{E} [|X_s|^p] \leq q^p \sup_{s \in [0, t]} \mathbb{E} [|X_s|^p]$$

and a monotone convergence argument. For the second estimate observe that we have $\mathbb{E}[|X_s|^p] \leq \mathbb{E}[|X_t|^p]$ for all $s \leq t$.

Finally, if $t \mapsto X_t$ is continuous and D a dense subset of $[0, \infty)$, we know that

$$\sup_{s \in D \cap [0, t]} |X_s|^p = \sup_{s \in [0, t]} |X_s|^p. \quad \square$$

A.4 Stopping and sampling

Recall that a stochastic process $(X_t)_{t \geq 0}$ is said to be *adapted* to the filtration $(\mathcal{F}_t)_{t \geq 0}$ if X_t is for each $t \geq 0$ an \mathcal{F}_t measurable random variable.

A.4.1 Stopping times

A.11 Definition. Let $(\mathcal{F}_t)_{t \geq 0}$ be a filtration. A random time $\tau : \Omega \rightarrow [0, \infty]$ is called an (\mathcal{F}_t) -*stopping time* if

$$\{\tau \leq t\} \in \mathcal{F}_t \quad \text{for all } t \geq 0. \quad (\text{A.15})$$

Here are a few rules for stopping times which should be familiar from the discrete time setting.

A.12 Lemma. Let $\sigma, \tau, \tau_j, j \geq 1$ be stopping times with respect to a filtration $(\mathcal{F}_t)_{t \geq 0}$. Then

- a) $\{\tau < t\} \in \mathcal{F}_t$;
- b) $\tau + \sigma, \tau \wedge \sigma, \tau \vee \sigma, \sup_j \tau_j$ are \mathcal{F}_t stopping times;
- c) $\inf_j \tau_j, \varliminf_j \tau_j, \overline{\lim}_j \tau_j$ are \mathcal{F}_{t+} stopping times.

Proof. The proofs are similar to the discrete time setting. Let $t \geq 0$.

- a) $\{\tau < t\} = \bigcup_{k \geq 1} \{\tau \leq t - \frac{1}{k}\} \in \bigcup_{k \geq 1} \mathcal{F}_{t - \frac{1}{k}} \subset \mathcal{F}_t$.

b) We have for all $t \geq 0$

$$\begin{aligned}
 \{\sigma + \tau > t\} &= \{\sigma = 0, \tau > t\} \cup \{\sigma > t, \tau = 0\} \cup \{\sigma > 0, \tau \geq t\} \\
 &\quad \cup \{0 < \tau < t, \tau + \sigma > t\} \\
 &= \underbrace{\{\sigma = 0, \tau > t\}}_{=\{\sigma \leq 0\} \cap \{\tau \leq t\}^c \in \mathcal{F}_t} \cup \underbrace{\{\sigma > t, \tau = 0\}}_{=\{\sigma \leq t\}^c \cap \{\tau \leq 0\} \in \mathcal{F}_t} \cup \underbrace{\{\sigma > 0, \tau \geq t\}}_{=\{\sigma \leq 0\}^c \cap \{\tau < t\}^c \in \mathcal{F}_t} \\
 &\quad \cup \bigcup_{r \in (0, t) \cap \mathbb{Q}} \underbrace{\{r < \tau < t, \sigma > t - r\}}_{=\{\tau \leq r\}^c \cap \{\tau < t\} \cap \{\sigma \leq t - r\}^c \in \mathcal{F}_t}.
 \end{aligned}$$

This shows that $\{\sigma + \tau \leq t\} = \{\sigma + \tau > t\}^c \in \mathcal{F}_t$ for all $t \geq 0$.

- $\{\sigma \vee \tau \leq t\} = \{\sigma \leq t\} \cap \{\tau \leq t\} \in \mathcal{F}_t$;
- $\{\sigma \wedge \tau \leq t\} = \{\sigma \leq t\} \cup \{\tau \leq t\} \in \mathcal{F}_t$;
- $\{\sup_j \tau_j \leq t\} = \bigcap_j \{\tau_j \leq t\} \in \mathcal{F}_t$;
- $\{\inf_j \tau_j < t\} = \bigcup_j \{\tau_j < t\} = \bigcup_j \bigcup_k \{\tau_j \leq t - \frac{1}{k}\} \in \mathcal{F}_t$; thus,

$$\{\inf_j \tau_j \leq t\} = \bigcap_{n \geq 1} \{\inf_j \tau_j \leq t + \frac{1}{n}\} \in \bigcap_{n \geq 1} \mathcal{F}_{t + \frac{1}{n}} = \mathcal{F}_{t+}.$$

c) Since \lim_j and $\overline{\lim}_j$ are combinations of inf and sup, the claim follows from the fact that $\inf_j \tau_j$ and $\sup_j \tau_j$ are \mathcal{F}_{t+} stopping times. \square



A.13 Example. The infimum of \mathcal{F}_t stopping times need not be an \mathcal{F}_t stopping time. To see this, let τ be an \mathcal{F}_{t+} stopping time (but not an \mathcal{F}_t stopping time). Then $\tau_j := \tau + \frac{1}{j}$ are \mathcal{F}_t stopping times:

$$\{\tau_j \leq t\} = \{\tau \leq t - \frac{1}{j}\} \in \mathcal{F}_{(t - \frac{1}{j})+} \subset \mathcal{F}_t.$$

But $\tau = \inf_j \tau_j$ is, by construction, not an \mathcal{F}_t stopping time.

We will now associate a σ -algebra with a stopping time.

A.14 Definition. Let τ be an \mathcal{F}_t stopping time. Then

$$\begin{aligned}
 \mathcal{F}_\tau &:= \left\{ A \in \mathcal{F}_\infty := \sigma\left(\bigcup_{t \geq 0} \mathcal{F}_t\right) : A \cap \{\tau \leq t\} \in \mathcal{F}_t \ \forall t \geq 0 \right\}, \\
 \mathcal{F}_{\tau+} &:= \left\{ A \in \mathcal{F}_\infty := \sigma\left(\bigcup_{t \geq 0} \mathcal{F}_t\right) : A \cap \{\tau \leq t\} \in \mathcal{F}_{t+} \ \forall t \geq 0 \right\}.
 \end{aligned}$$

It is not hard to see that \mathcal{F}_τ and $\mathcal{F}_{\tau+}$ are σ -algebras and that for $\tau \equiv t$ we have $\mathcal{F}_\tau = \mathcal{F}_t$. Note that, despite the slightly misleading notation, \mathcal{F}_τ is a family of sets which does not depend on ω .



A.15 Lemma. *Let $\sigma, \tau, \tau_j, j \geq 1$ be \mathcal{F}_t stopping times. Then*

- a) τ is \mathcal{F}_τ measurable;
- b) if $\sigma \leq \tau$ then $\mathcal{F}_\sigma \subset \mathcal{F}_\tau$;
- c) if $\tau_1 \geq \tau_2 \geq \dots \geq \tau_n \downarrow \tau$, then $\mathcal{F}_{\tau+} = \bigcap_{j \geq 1} \mathcal{F}_{\tau_j+}$. If $(\mathcal{F}_t)_{t \geq 0}$ is right-continuous,

$$\text{then } \mathcal{F}_\tau = \bigcap_{j \geq 1} \mathcal{F}_{\tau_j}.$$

Proof. a) We have to show that $\{\tau \leq s\} \in \mathcal{F}_\tau$ for all $s \geq 0$. But this follows from

$$\{\tau \leq s\} \cap \{\tau \leq t\} = \{\tau \leq s \wedge t\} \in \mathcal{F}_{s \wedge t} \subset \mathcal{F}_t \quad \forall t \geq 0.$$

b) Let $F \in \mathcal{F}_\sigma$. Then we have for all $t \geq 0$

$$F \cap \{\tau \leq t\} = F \cap \{\tau \leq t\} \cap \underbrace{\{\sigma \leq \tau\}}_{=\Omega} = \underbrace{F \cap \{\sigma \leq t\}}_{\in \mathcal{F}_t \text{ as } F \in \mathcal{F}_\sigma} \cap \underbrace{\{\tau \leq t\}}_{\in \mathcal{F}_t} \in \mathcal{F}_t.$$

Thus, $F \in \mathcal{F}_\tau$ and $\mathcal{F}_\sigma \subset \mathcal{F}_\tau$.

c) We know from Lemma A.12 that $\tau := \inf_j \tau_j$ is an \mathcal{F}_{t+} stopping time. Applying b) to the filtration $(\mathcal{F}_{t+})_{t \geq 0}$ we get

$$\tau_j \geq \tau \quad \forall j \geq 1 \xRightarrow{\text{b)}} \bigcap_{j \geq 1} \mathcal{F}_{\tau_j+} \supset \mathcal{F}_{\tau+}.$$

On the other hand, if $F \in \mathcal{F}_{\tau_j+}$ for all $j \geq 1$, then

$$F \cap \{\tau < t\} = F \cap \bigcup_{j \geq 1} \{\tau_j < t\} = \bigcup_{j \geq 1} F \cap \{\tau_j < t\} \in \mathcal{F}_t$$

where we use that

$$F \cap \{\tau_j < t\} = \bigcup_{k \geq 1} F \cap \{\tau_j \leq t - \frac{1}{k}\} \in \bigcup_{k \geq 1} \mathcal{F}_{t - \frac{1}{k}} \subset \mathcal{F}_t.$$

As in (5.9) we see now that $F \cap \{\tau \leq t\} \in \mathcal{F}_{t+}$.

If $(\mathcal{F}_t)_{t \geq 0}$ is right-continuous, we have $\mathcal{F}_t = \mathcal{F}_{t+}$ and $\mathcal{F}_\tau = \mathcal{F}_{\tau+}$. □

We close this section with a result which allows us to approximate stopping times from above by a decreasing sequence of stopping times with countably many values. This helps us to reduce many assertions to the discrete time setting.

A.16 Lemma. *Let τ be an \mathcal{F}_t stopping time. Then there exists a decreasing sequence of \mathcal{F}_t stopping times with*

$$\tau_1 \geq \tau_2 \geq \dots \geq \tau_n \downarrow \tau = \inf_j \tau_j$$

and each τ_j takes only countably many values:

$$\tau_j(\omega) = \begin{cases} \tau(\omega), & \text{if } \tau(\omega) = \infty; \\ k2^{-j}, & \text{if } (k-1)2^{-j} \leq \tau(\omega) < k2^{-j}, \quad k \geq 1. \end{cases}$$

Moreover, $F \cap \{\tau_j = k2^{-j}\} \in \mathcal{F}_{k/2^j}$ for all $F \in \mathcal{F}_{\tau+}$.

Proof. Fix $t \geq 0$ and $j \geq 1$ and pick $k = k(t, j)$ such that $(k-1)2^{-j} \leq t < k2^{-j}$. Then

$$\{\tau_j \leq t\} = \{\tau_j \leq (k-1)2^{-j}\} = \{\tau < (k-1)2^{-j}\} \in \mathcal{F}_{(k-1)/2^j} \subset \mathcal{F}_t,$$

i. e. the τ_j 's are stopping times.

From the definition it is obvious that $\tau_j \downarrow \tau$: Note that $\tau_j(\omega) - \tau(\omega) \leq 2^{-j}$ if $\tau(\omega) < \infty$. Finally, if $F \in \mathcal{F}_{\tau+}$, we have

$$\begin{aligned} F \cap \{\tau_j = k2^{-j}\} &= F \cap \{(k-1)2^{-j} \leq \tau < k2^{-j}\} \\ &= F \cap \{\tau < k2^{-j}\} \cap \{\tau < (k-1)2^{-j}\}^c \\ &= \bigcup_{n \geq 1} \underbrace{\left(F \cap \left\{\tau \leq k2^{-j} - \frac{1}{n}\right\}\right)}_{\in \mathcal{F}_{(k/2^j - 1/n)+} \subset \mathcal{F}_{k/2^j}} \underbrace{\cap \{\tau < (k-1)2^{-j}\}^c}_{\in \mathcal{F}_{k/2^j}} \in \mathcal{F}_{k/2^j}. \quad \square \end{aligned}$$

A.4.2 Optional sampling

If we evaluate a continuous martingale at an increasing sequence of stopping times, we still have a martingale. The following technical lemma has exactly the same proof as Lemma 6.24 and its Corollary 6.25 – all we have to do is to replace in these proofs Brownian motion B by the continuous process X .

A.17 Lemma. *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a d -dimensional adapted stochastic process with continuous sample paths and let τ be an \mathcal{F}_t stopping time. Then $X(\tau)\mathbb{1}_{\{\tau < \infty\}}$ is an \mathcal{F}_τ measurable random variable.*

Recall the optional sampling theorem for stopping times $\alpha, \beta : \Omega \rightarrow \{t_j\}_{j \geq 1}$ with values in a discrete set:

$$(X(t_j), \mathcal{F}_{t_j})_{j \geq 1} \text{ submartingale, } \alpha \leq \beta \leq C \implies X(\alpha) \leq \mathbb{E}(X(\beta) | \mathcal{F}_\alpha).$$

A.18 Theorem (Optional stopping, Doob 1953). *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a submartingale with continuous sample paths and let $\sigma \leq \tau$ be \mathcal{F}_t stopping times. Then we have for all $k \geq 0$ that $X_{\sigma \wedge k}, X_{\tau \wedge k} \in L^1(\mathbb{P})$ and*

$$X(\sigma \wedge k) \leq \mathbb{E}(X(\tau \wedge k) \mid \mathcal{F}_{\sigma \wedge k}). \quad (\text{A.16})$$

If $\sigma \leq \tau$ are bounded stopping times, i. e. if $\tau \leq K$ a. s. for some constant K , or if $\mathbb{P}(\tau < \infty) = 1$, $X(\sigma), X(\tau) \in L^1(\mathbb{P})$ and $\lim_{k \rightarrow \infty} \mathbb{E}[|X(k)| \mathbb{1}_{\{\tau > k\}}] = 0$, then

$$X(\sigma) \leq \mathbb{E}(X(\tau) \mid \mathcal{F}_\sigma). \quad (\text{A.17})$$

Proof. Using Lemma A.16 we can approximate $\sigma \leq \tau$ by discrete-valued stopping times $\sigma_j \downarrow \sigma$ and $\tau_j \downarrow \tau$. The construction in Lemma A.16 shows that we still have $\sigma_j \leq \tau_j$.

Using the discrete version of Doob's optional sampling theorem with $\alpha = \sigma_j \wedge k$ and $\beta = \tau_j \wedge k$ reveals that $X(\sigma_j \wedge k), X(\tau_j \wedge k) \in L^1(\mathbb{P})$ and

$$X(\sigma_j \wedge k) \leq \mathbb{E}[X(\tau_j \wedge k) \mid \mathcal{F}_{\sigma_j \wedge k}].$$

Since $\mathcal{F}_{\sigma \wedge k} \subset \mathcal{F}_{\sigma_j \wedge k}$, the tower property gives

$$\mathbb{E}[X(\sigma_j \wedge k) \mid \mathcal{F}_{\sigma \wedge k}] \leq \mathbb{E}[X(\tau_j \wedge k) \mid \mathcal{F}_{\sigma \wedge k}]. \quad (\text{A.18})$$

By discrete optional sampling, $(X(\sigma_j \wedge k), \mathcal{F}_{\sigma_j \wedge k})_{j \geq 1}$ and $(X(\tau_j \wedge k), \mathcal{F}_{\tau_j \wedge k})_{j \geq 1}$ are backwards submartingales and as such, cf. Corollary A.8, uniformly integrable. Therefore,

$$X(\sigma_j \wedge k) \xrightarrow[j \rightarrow \infty]{\text{a. s., } L^1} X(\sigma \wedge k) \quad \text{and} \quad X(\tau_j \wedge k) \xrightarrow[j \rightarrow \infty]{\text{a. s., } L^1} X(\tau \wedge k).$$

This shows that we have for all $F \in \mathcal{F}_{\sigma \wedge k}$

$$\begin{aligned} \int_F X(\sigma \wedge k) d\mathbb{P} &= \lim_{j \rightarrow \infty} \int_F X(\sigma_j \wedge k) d\mathbb{P} \stackrel{(\text{A.18})}{\leq} \lim_{j \rightarrow \infty} \int_F X(\tau_j \wedge k) d\mathbb{P} \\ &= \int_F X(\tau \wedge k) d\mathbb{P} \end{aligned}$$

which proves (A.17).

Assume now that $X(\sigma), X(\tau) \in L^1(\mathbb{P})$. If $F \in \mathcal{F}_\sigma$, then

$$F \cap \{\sigma \leq k\} \cap \{\sigma \wedge k \leq t\} = F \cap \{\sigma \leq k\} \cap \{\sigma \leq t\} = F \cap \{\sigma \leq k \wedge t\} \in \mathcal{F}_{k \wedge t} \subset \mathcal{F}_t,$$

which means that $F \cap \{\sigma \leq k\} \in \mathcal{F}_{\sigma \wedge k}$. Since $\sigma \leq \tau < \infty$ we have by assumption

$$\int_{F \cap \{\sigma > k\}} |X(k)| d\mathbb{P} \leq \int_{\{\sigma > k\}} |X(k)| d\mathbb{P} \leq \int_{\{\tau > k\}} |X(k)| d\mathbb{P} \xrightarrow[k \rightarrow \infty]{} 0.$$

Using (A.16) we get by dominated convergence

$$\begin{aligned}
 \int_F X(\sigma \wedge k) d\mathbb{P} &= \overbrace{\int_{F \cap \{\sigma \leq k\}} X(\sigma \wedge k) d\mathbb{P}}^{\rightarrow \int_F X(\sigma) d\mathbb{P}, k \rightarrow \infty} + \overbrace{\int_{F \cap \{\sigma > k\}} X(k) d\mathbb{P}}^{\rightarrow 0, k \rightarrow \infty} \\
 &\leq \int_{F \cap \{\sigma \leq k\}} X(\tau \wedge k) d\mathbb{P} + \int_{F \cap \{\sigma > k\}} X(k) d\mathbb{P} \\
 &= \overbrace{\int_{F \cap \{\tau \leq k\}} X(\tau \wedge k) d\mathbb{P}}^{\rightarrow \int_F X(\tau) d\mathbb{P}, k \rightarrow \infty} + \overbrace{\int_{F \cap \{\tau > k\}} X(k) d\mathbb{P}}^{\rightarrow 0, k \rightarrow \infty}
 \end{aligned}$$

which is exactly (A.17). \square

A.19 Corollary. Let $(X_t, \mathcal{F}_t)_{t \geq 0}$, $X_t \in L^1(\mathbb{P})$, be an adapted process with continuous sample paths. The process $(X_t, \mathcal{F}_t)_{t \geq 0}$ is a submartingale if, and only if,

$$\mathbb{E} X_\sigma \leq \mathbb{E} X_\tau \text{ for all bounded } \mathcal{F}_t \text{ stopping times } \sigma \leq \tau. \quad (\text{A.19})$$

Proof. If $(X_t, \mathcal{F}_t)_{t \geq 0}$ is a submartingale, then (A.19) follows at once from Doob's optional stopping theorem. Conversely, assume that (A.19) holds true. Let $s < t$, pick any $F \in \mathcal{F}_s$ and set $\rho := t \mathbb{1}_{F^c} + s \mathbb{1}_F$. Obviously, ρ is a bounded stopping time and if we use (A.19) with $\sigma := \rho$ and $\tau := t$ we see

$$\mathbb{E} (X_s \mathbb{1}_F) + \mathbb{E} (X_t \mathbb{1}_{F^c}) = \mathbb{E} X_\rho \leq \mathbb{E} X_t.$$

Therefore,

$$\mathbb{E} (X_s \mathbb{1}_F) \leq \mathbb{E} (X_t \mathbb{1}_F) \text{ for all } s < t \text{ and } F \in \mathcal{F}_s,$$

i. e. $X_s \leq \mathbb{E}(X_t | \mathcal{F}_s)$. \square

A.20 Corollary. Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a submartingale which is uniformly integrable and has continuous paths. If $\sigma \leq \tau$ are stopping times with $\mathbb{P}(\tau < \infty) = 1$, then all assumptions of Theorem A.18 are satisfied.

Proof. Since $(X_t)_{t \geq 0}$ is uniformly integrable, we get

$$\sup_{t \geq 0} \mathbb{E} |X(t)| \leq c < \infty.$$

Moreover, since X_t^+ is a submartingale, we have for all $k \geq 0$

$$\begin{aligned}
 \mathbb{E} |X(\sigma \wedge k)| &= 2 \mathbb{E} X^+(\sigma \wedge k) - \mathbb{E} X(\sigma \wedge k) \stackrel{(\text{A.16})}{\leq} 2 \mathbb{E} X^+(k) - \mathbb{E} X(0) \\
 &\leq 2 \mathbb{E} |X(k)| + \mathbb{E} |X(0)| \leq 3c.
 \end{aligned}$$

By Fatou's lemma we get

$$\mathbb{E} |X(\sigma)| \leq \liminf_{k \rightarrow \infty} \mathbb{E} |X(\sigma \wedge k)| \leq 3c,$$

and, similarly, $\mathbb{E} |X(\tau)| \leq 3c$. Finally,

$$\begin{aligned} \mathbb{E} [|X(k)| \mathbb{1}_{\{\tau > k\}}] &= \mathbb{E} [|X(k)| (\mathbb{1}_{\{\tau > k\} \cap \{|X(k)| > R\}} + \mathbb{1}_{\{\tau > k\} \cap \{|X(k)| \leq R\}})] \\ &\leq \sup_n \mathbb{E} [|X(n)| \mathbb{1}_{\{|X(n)| > R\}}] + R \mathbb{P}(\tau > k) \\ &\xrightarrow[k \rightarrow \infty]{} \sup_n \mathbb{E} [|X(n)| \mathbb{1}_{\{|X(n)| > R\}}] \xrightarrow[R \rightarrow \infty]{\text{unif. integrable}} 0. \quad \square \end{aligned}$$

A.21 Remark.

- a) If $(X_t, \mathcal{F}_t)_{t \geq 0}$ is a martingale, we have '=' in (A.16), (A.17) and (A.19).
 b) Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a submartingale and ρ an \mathcal{F}_t stopping time. Then $(X_t^\rho, \mathcal{F}_t)$, $X_t^\rho := X_{\rho \wedge t}$, is a submartingale.

Indeed: X^ρ is obviously adapted and continuous. For any two bounded stopping times $\sigma \leq \tau$ it is clear that $\rho \wedge \sigma \leq \rho \wedge \tau$ are bounded stopping times, too. Therefore,

$$\mathbb{E} X_\sigma^\rho \stackrel{\text{def}}{=} \mathbb{E} X_{\rho \wedge \sigma} \stackrel{\text{A.19}}{\leq} \mathbb{E} X_{\rho \wedge \tau} \stackrel{\text{def}}{=} \mathbb{E} X_\tau^\rho,$$

and yet another application of Corollary A.19 shows that X^ρ is a submartingale for $(\mathcal{F}_t)_{t \geq 0}$.

- c) Note that $(X_t = X_{t+}, \mathcal{F}_{t+})$ is also a submartingale. If $\sigma \leq \tau$ are \mathcal{F}_{t+} stopping times, then (A.16) and (A.17) remain valid if we use $\mathcal{F}_{(\sigma \wedge k)+}, \mathcal{F}_{\sigma+}$ etc.

We close this section with a uniqueness result for compensators.

A.22 Proposition. *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a martingale with paths which are continuous and of bounded variation on compact intervals. Then $X_t = X_0$ a.s. for all $t \geq 0$.*

Proof. Without loss of generality we may assume that $X_0 = 0$. Denote by

$$V_t := \text{VAR}_1(X; [0, t]) \quad \text{and} \quad \tau_n = \inf\{t \geq 0 : V_t \geq n\}$$

the strong variation of $(X_t)_{t \geq 0}$ and the stopping time when V_t exceeds n for the first time. Since $t \mapsto X_t$ is continuous, so is $t \mapsto V_t$, cf. Paragraph A.34 f).

By optional stopping, Theorem A.18, we know that $(X_t^{\tau_n}, \mathcal{F}_{t \wedge \tau_n})_{t \geq 0}$ is, for every $n \geq 1$, a martingale. An application of Doob's maximal inequality yields

$$\begin{aligned} \mathbb{E} \left[\sup_{s \leq t} (X_s^{\tau_n})^2 \right] &\stackrel{(A.14)}{\leq} 4 \mathbb{E} [(X_t^{\tau_n})^2] = \mathbb{E} \left[\sum_{j=1}^N ((X_{\frac{j}{N}t}^{\tau_n})^2 - (X_{\frac{j-1}{N}t}^{\tau_n})^2) \right] \\ &= \mathbb{E} \left[\sum_{j=1}^N (X_{\frac{j}{N}t}^{\tau_n} - X_{\frac{j-1}{N}t}^{\tau_n})^2 \right] \\ &\leq \mathbb{E} \left[V_t \sup_{1 \leq j \leq N} |X_{\frac{j}{N}t}^{\tau_n} - X_{\frac{j-1}{N}t}^{\tau_n}| \right] \\ &\leq n \mathbb{E} \left[\sup_{1 \leq j \leq N} |X_{\frac{j}{N}t}^{\tau_n} - X_{\frac{j-1}{N}t}^{\tau_n}| \right]; \end{aligned}$$

in particular, the expectation on the left-hand side is finite. Since the paths are uniformly continuous on compact intervals, we find

$$\sup_{1 \leq j \leq N} |X_{\frac{j}{N}t}^{\tau_n} - X_{\frac{j-1}{N}t}^{\tau_n}| \xrightarrow[N \rightarrow \infty]{} 0 \quad \text{almost surely.}$$

On the other hand, $\sup_{1 \leq j \leq N} |X_{\frac{j}{N}t}^{\tau_n} - X_{\frac{j-1}{N}t}^{\tau_n}| \leq \text{VAR}(X^{\tau_n}; [0, t]) \leq V_{t \wedge \tau_n} \leq n$, and dominated convergence gives

$$\mathbb{E} \left[\sup_{s \leq t} (X_s^{\tau_n})^2 \right] \leq n \lim_{N \rightarrow \infty} \mathbb{E} \left[\sup_{1 \leq j \leq N} |X_{\frac{j}{N}t}^{\tau_n} - X_{\frac{j-1}{N}t}^{\tau_n}| \right] = 0.$$

Therefore, $X_s = 0$ a. s. on $[0, k \wedge \tau_n]$ for all $k \geq 1$ and $n \geq 1$. The claim follows since $\tau_n \uparrow \infty$ almost surely as $n \rightarrow \infty$. \square

A.23 Corollary. *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be an L^2 -martingale such that both $(X_t^2 - A_t, \mathcal{F}_t)_{t \geq 0}$ and $(X_t^2 - A'_t, \mathcal{F}_t)_{t \geq 0}$ are martingales for some adapted processes $(A_t)_{t \geq 0}$ and $(A'_t)_{t \geq 0}$ having continuous paths of bounded variation on compact sets and $A_0 = A'_0 = 0$. Then $\mathbb{P}(A_t = A'_t \forall t \geq 0) = 1$.*

Proof. By assumption $A_t - A'_t = (X_t^2 - A'_t) - (X_t^2 - A_t)$, i. e. $A - A'$ is a martingale with continuous paths which are of bounded variation on compact sets. The claim follows from Proposition A.22. \square

A.5 Remarks on Feller processes

A Markov process $(X_t, \mathcal{F}_t)_{t \geq 0}$ is said to be a *Feller process* if the transition semigroup (7.1) has the Feller property, i. e. P_t is a strongly continuous Markov semigroup on the space $\mathcal{C}_\infty(\mathbb{R}^d) = \{u \in \mathcal{C}(\mathbb{R}^d) : \lim_{|x| \rightarrow \infty} u(x) = 0\}$, cf. Remark 7.4. In

Remark 7.6 we have seen how to construct a Feller process from any given Feller semigroup. Throughout this section we assume that the filtration is right-continuous, i. e. $\mathcal{F}_t = \mathcal{F}_{t+}$ for all $t \geq 0$, and complete. Let us briefly discuss some properties of Feller processes.

A.24 Theorem. *Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a Feller process. Then there exists a Feller process $(\tilde{X}_t)_{t \geq 0}$ whose sample paths are right-continuous with finite left limits and which has the same finite dimensional distributions as $(X_t)_{t \geq 0}$.*

Proof. See Blumenthal and Gettoor [13, Theorem I.9.4, p. 46] or Ethier and Kurtz [61, Theorem 4.2.7 p. 169]. \square

From now on we will assume that every Feller process has right-continuous paths.

A.25 Theorem. *Every Feller process $(X_t, \mathcal{F}_t)_{t \geq 0}$ has the strong Markov property (6.9).*

Proof. The argument is similar to the proof of Theorem 6.5. Let σ be a stopping time and approximate σ by the stopping times $\sigma_j := (\lfloor 2^j \sigma \rfloor + 1)/2^j$, cf. Lemma A.16. Let $t > 0$, $x \in \mathbb{R}^d$, $u \in \mathcal{C}_\infty(\mathbb{R}^d)$ and $F \in \mathcal{F}_{\sigma+}$. Since $\{\sigma < \infty\} = \{\sigma_n < \infty\}$ we find by dominated convergence and the right-continuity of the sample paths

$$\begin{aligned} \int_{F \cap \{\sigma < \infty\}} u(X_{t+\sigma}) d\mathbb{P}^x &= \lim_{j \rightarrow \infty} \int_{F \cap \{\sigma_j < \infty\}} u(X_{t+\sigma_j}) d\mathbb{P}^x \\ &= \lim_{j \rightarrow \infty} \sum_{k=0}^{\infty} \mathbb{E}^x [u(X_{t+\sigma_j}) \mathbb{1}_{F \cap \{\sigma_j = k/2^j\}}] \\ &= \lim_{j \rightarrow \infty} \sum_{k=0}^{\infty} \mathbb{E}^x [u(X_{t+k/2^j}) \mathbb{1}_{F \cap \{\sigma_j = k/2^j\}}]. \end{aligned}$$

By Lemma A.16, $F \cap \{\sigma_j = k/2^j\} \in \mathcal{F}_{\sigma+}$; using the tower and the Markov property, we get

$$\begin{aligned} &\mathbb{E}^x [\mathbb{E} (u(X_{t+k/2^j}) \mathbb{1}_{F \cap \{\sigma_j = k/2^j\}} \mid \mathcal{F}_{k/2^j})] \\ &= \mathbb{E}^x [\mathbb{1}_{F \cap \{\sigma_j = k/2^j\}} \mathbb{E}^x (u(X_{t+k/2^j}) \mid \mathcal{F}_{k/2^j})] \\ &= \mathbb{E}^x [\mathbb{1}_{F \cap \{\sigma_j = k/2^j\}} \mathbb{E}^{X_{k/2^j}} u(X_t)] \\ &= \mathbb{E}^x [\mathbb{1}_{F \cap \{\sigma_j = k/2^j\}} \mathbb{E}^{X_{\sigma_j}} u(X_t)] \end{aligned}$$

and, again by dominated convergence and the right-continuity

$$\begin{aligned}
 \int_{F \cap \{\sigma < \infty\}} u(X_{t+\sigma}) d\mathbb{P}^x &= \lim_{j \rightarrow \infty} \sum_{k=0}^{\infty} \mathbb{E}^x [\mathbb{1}_{F \cap \{\sigma_j = k/2^j\}} \mathbb{E}^{X_{\sigma_j}} u(X_t)] \\
 &= \lim_{j \rightarrow \infty} \mathbb{E}^x [\mathbb{1}_{F \cap \{\sigma < \infty\}} \mathbb{E}^{X_{\sigma_j}} u(X_t)] \\
 &= \mathbb{E}^x [\mathbb{1}_{F \cap \{\sigma < \infty\}} \mathbb{E}^{X_{\sigma}} u(X_t)]. \quad \square
 \end{aligned}$$

A.26 Theorem. Let $(X_t, \mathcal{F}_t)_{t \geq 0}$ be a Feller process and $\tau_x := \inf\{t > 0 : X_t \neq x\}$ the first hitting time of the set $\mathbb{R}^d \setminus \{x\}$. Then there exists some $\lambda(x) \in [0, \infty]$ such that $\mathbb{P}^x(\tau_x \geq t) = e^{-\lambda(x)t}$.

Proof. Under \mathbb{P}^x we have, a. s. $\tau_x > t$ if, and only if, $\sup_{s \leq t} |X_s - x| = 0$. Therefore,

$$\mathbb{P}^x(\tau_x > t + s) = \mathbb{P}^x(\tau_x > s, \sup_{r \leq t} |X_{r+s} - X_s| = 0)$$

By the Markov property we find

$$\begin{aligned}
 \mathbb{P}^x(\tau_x > s, \sup_{r \leq t} |X_{r+s} - X_s| = 0) &= \mathbb{E}^x(\mathbb{1}_{\{\tau_x > s\}} \mathbb{P}^{X_s}[\sup_{r \leq t} |X_r - X_0| = 0]) \\
 &\stackrel{X_s=x}{=} \mathbb{E}^x(\mathbb{1}_{\{\tau_x > s\}}) \mathbb{P}^x[\sup_{r \leq t} |X_r - x| = 0].
 \end{aligned}$$

Thus, $\mathbb{P}^x(\tau_x > t + s) = \mathbb{P}^x(\tau_x > s) \mathbb{P}^x(\tau_x > t)$. As

$$\lim_{n \rightarrow \infty} \mathbb{P}^x(\tau_x > t - 1/n) = \mathbb{P}^x(\tau_x \geq t),$$

we find for all $s, t > 0$ that $\mathbb{P}^x(\tau_x \geq t + s) = \mathbb{P}^x(\tau_x \geq s) \mathbb{P}^x(\tau_x \geq t)$. Since $s \mapsto \mathbb{P}^x(\tau_x \geq s)$ is left-continuous and $\mathbb{P}^x(\tau_x \geq 0) = 1$, this functional equation has a unique solution of the form $e^{-\lambda(x)t}$ with some $\lambda(x) \in [0, \infty]$. \square

A.6 The Doob–Meyer decomposition

We will now show that every martingale $M \in \mathcal{M}_T^{2,c}$ satisfies

$$\begin{aligned}
 (M_t^2 - A_t, \mathcal{F}_t)_{t \leq T} \quad \text{is a martingale for some} \\
 \text{adapted, continuous and increasing process } (A_t)_{t \leq T}. \quad (\star)
 \end{aligned}$$

We begin with the uniqueness of the process A . This follows directly from Corollary A.23.

A.27 Lemma. Assume that $M \in \mathcal{M}_T^{2,c}$ and that there are two continuous and increasing processes $(A_t)_{t \leq T}$ and $(A'_t)_{t \leq T}$ such that both $(M_t^2 - A_t)_{t \leq T}$ and $(M_t^2 - A'_t)_{t \leq T}$ are martingales. Then $A_t - A_0 = A'_t - A'_0$ a. s. with the same null set for all $t \leq T$.

Let us now turn to the existence of the process A from (★). The following elementary proof is due to Kunita [111, Chapter 2.2]. We consider first bounded martingales $M \in \mathcal{M}_T^{2,c}$.

A.28 Theorem. *Let $M \in \mathcal{M}_T^{2,c}$ such that $\sup_{t \leq T} |M_t| \leq \kappa$ for some constant $\kappa < \infty$. Let $(\Pi_n)_{n \geq 1}$ be any sequence of finite partitions of $[0, T]$ satisfying $\lim_{n \rightarrow \infty} |\Pi_n| = 0$. Then the mean-square limit*

$$L^2(\mathbb{P})\text{-}\lim_{n \rightarrow \infty} \sum_{t_{j-1}, t_j \in \Pi_n} (M_{t \wedge t_j} - M_{t \wedge t_{j-1}})^2$$

exists uniformly for all $t \leq T$ and defines an increasing and continuous process $(\langle M \rangle_t)_{t \geq 0}$ such that $M^2 - \langle M \rangle$ is a martingale.

The proof is based on two lemmas. Let $\Pi = \{0 = t_0 < t_1 < \dots < t_n = T\}$ be a partition of $[0, T]$ and denote by $M^{\Pi, t} = (M_{t_j}^t, \mathcal{F}_{t_j \wedge t})_{t_{j-1}, t_j \in \Pi, 0 \leq t \leq T}$, where $M_{t_j}^t = M_{t \wedge t_j}$. Consider the martingale transform

$$N_T^{\Pi, t} := M^{\Pi, t} \bullet M_T^{\Pi, t} = \sum_{k=1}^n M_{t \wedge t_{j-1}} (M_{t \wedge t_j} - M_{t \wedge t_{j-1}}).$$

A.29 Lemma. $(N_T^{\Pi, t}, \mathcal{F}_t)_{t \leq T}$ is a continuous martingale and $\mathbb{E}[(N_T^{\Pi, t})^2] \leq \kappa^4$. Moreover,

$$M_t^2 - M_0^2 = \sum_{j=1}^n (M_{t \wedge t_j} - M_{t \wedge t_{j-1}})^2 + 2N_T^{\Pi, t}, \quad (\text{A.20})$$

and $\mathbb{E}[(\sum_{j=1}^n (M_{t \wedge t_j} - M_{t \wedge t_{j-1}})^2)^2] \leq 16\kappa^4$.

Proof. The martingale property is just Theorem 14.9 a). For the first estimate we use Theorem 14.4

$$\begin{aligned} \mathbb{E}[(N_T^{\Pi, t})^2] &= \mathbb{E}[(M^{\Pi, t} \bullet M_T^{\Pi, t})^2] \stackrel{(14.5)}{=} \mathbb{E}[(M^{\Pi, t})^2 \bullet \langle M^{\Pi, t} \rangle_T] \\ &\leq \kappa^2 \mathbb{E}[\langle M^{\Pi, t} \rangle_T] = \kappa^2 \mathbb{E}[M_T^2 - M_0^2] \leq \kappa^4. \end{aligned}$$

The identity (A.20) follows from

$$\begin{aligned}
 M_t^2 - M_0^2 &= \sum_{j=1}^n (M_{t \wedge t_j}^2 - M_{t \wedge t_{j-1}}^2) \\
 &= \sum_{j=1}^n (M_{t \wedge t_j} + M_{t \wedge t_{j-1}})(M_{t \wedge t_j} - M_{t \wedge t_{j-1}}) \\
 &= \sum_{j=1}^n [(M_{t \wedge t_j} - M_{t \wedge t_{j-1}})^2 + 2M_{t \wedge t_{j-1}}(M_{t \wedge t_j} - M_{t \wedge t_{j-1}})] \\
 &= \sum_{j=1}^n (M_{t \wedge t_j} - M_{t \wedge t_{j-1}})^2 + 2N_T^{\Pi, t}.
 \end{aligned}$$

From this and the first estimate we get

$$\begin{aligned}
 &\mathbb{E} \left[\left(\sum_{j=1}^n (M_{t \wedge t_j} - M_{t \wedge t_{j-1}})^2 \right)^2 \right] \\
 &= \mathbb{E} \left[(2N_T^{\Pi, t} + \underbrace{M_t^2 - M_0^2}_{\leq 2\kappa^2})^2 \right] \leq 8 \mathbb{E} [(N_T^{\Pi, t})^2] + 8\kappa^4 \leq 16\kappa^4.
 \end{aligned}$$

In the penultimate inequality we used the fact that $(a + b)^2 \leq 2a^2 + 2b^2$. \square

A.30 Lemma. *Let $(\Pi_n)_{n \geq 1}$ be a sequence of partitions of $[0, T]$ such that the mesh size $|\Pi_n| = \max_{t_{j-1}, t_j \in \Pi_n} (t_j - t_{j-1})$ tends to 0 as $n \rightarrow \infty$. Then*

$$\lim_{m, n \rightarrow \infty} \mathbb{E} \left[\sup_{0 \leq t \leq T} |N_T^{\Pi_n, t} - N_T^{\Pi_m, t}|^2 \right] = 0.$$

Proof. For any $\Pi = \{0 = t_0 < t_1 < \dots < t_n = T\}$ we define $f_u^{\Pi, t} := M_{t \wedge t_{j-1}}$ if $t_{j-1} \leq u < t_j$. Note that for any two partitions Π, Π' and all $u \leq T$ we have

$$|f_u^{\Pi, t} - f_u^{\Pi', t}| \leq \sup_{|r-s| \leq |\Pi| \wedge |\Pi'|} |M_{t \wedge s} - M_{t \wedge r}|.$$

By the continuity of the sample paths, this quantity tends to zero as $|\Pi|, |\Pi'| \rightarrow 0$.

Fix two partitions Π, Π' of the sequence $(\Pi_n)_{n \geq 1}$ and denote the elements of Π by t_j and those of $\Pi \cup \Pi'$ by u_k . The process $f_u^{\Pi, t}$ is constructed in such a way that

$$\begin{aligned}
 N_T^{\Pi, t} &= M^{\Pi, t} \bullet M_T^{\Pi, t} = \sum_{t_{j-1}, t_j \in \Pi} M_{t \wedge t_{j-1}} (M_{t \wedge t_j} - M_{t \wedge t_{j-1}}) \\
 &= \sum_{u_{k-1}, u_k \in \Pi \cup \Pi'} f_{u_{k-1}}^{\Pi, t} (M_{t \wedge u_k} - M_{t \wedge u_{k-1}}) = f^{\Pi, t} \bullet M_T^{\Pi \cup \Pi', t}.
 \end{aligned}$$

Thus,

$$N_T^{\Pi,t} - N_T^{\Pi',t} = (f^{\Pi,t} - f^{\Pi',t}) \bullet M_T^{\Pi \cup \Pi',t},$$

and we find by Doob's maximal inequality A.10,

$$\begin{aligned} & \mathbb{E} \left[\sup_{s \leq t} |N_T^{\Pi,s} - N_T^{\Pi',s}|^2 \right] \\ & \leq 4 \mathbb{E} [|N_T^{\Pi,t} - N_T^{\Pi',t}|^2] \\ & = 4 \mathbb{E} [|(f^{\Pi,t} - f^{\Pi',t}) \bullet M_T^{\Pi \cup \Pi',t}|^2] \\ & \stackrel{(4.6)}{=} 4 \mathbb{E} \left[\sum_{u_{k-1}, u_k \in \Pi \cup \Pi'} (f_{u_{k-1}}^{\Pi,t} - f_{u_{k-1}}^{\Pi',t})^2 (M_{t \wedge u_k} - M_{t \wedge u_{k-1}})^2 \right] \\ & \leq 4 \mathbb{E} \left[\sup_{u \leq T} |f_u^{\Pi,t} - f_u^{\Pi',t}|^2 \sum_{u_{k-1}, u_k \in \Pi \cup \Pi'} (M_{t \wedge u_k} - M_{t \wedge u_{k-1}})^2 \right] \\ & \leq 4 \left(\mathbb{E} \left[\sup_{u \leq T} |f_u^{\Pi,t} - f_u^{\Pi',t}|^4 \right] \right)^{1/2} \\ & \quad \cdot \left(\mathbb{E} \left[\left(\sum_{u_{k-1}, u_k \in \Pi \cup \Pi'} (M_{t \wedge u_k} - M_{t \wedge u_{k-1}})^2 \right) \right] \right)^{1/2}. \end{aligned}$$

The first term tends to 0 as $|\Pi|, |\Pi'| \rightarrow 0$ while the second factor is bounded by $4\kappa^2$, cf. Lemma A.29. \square

Proof of Theorem A.28. Let $(\Pi_n)_{n \geq 0}$ be a sequence of refining partitions such that $|\Pi_n| \rightarrow 0$ and set $\Pi := \bigcup_{n \geq 0} \Pi_n$.

By Lemma A.30, the limit L^2 - $\lim_{n \rightarrow \infty} N_T^{\Pi_n,t}$ exists uniformly for all $t \in [0, T]$ and the same is true, because of (A.20), for L^2 - $\lim_{n \rightarrow \infty} \sum_{t_{j-1}, t_j \in \Pi_n} (M_{t \wedge t_j} - M_{t \wedge t_{j-1}})^2$. Therefore, there is a subsequence such that

$$\langle M \rangle_t := \lim_{k \rightarrow \infty} \sum_{t_{j-1}, t_j \in \Pi_n(k)} (M_{t \wedge t_j} - M_{t \wedge t_{j-1}})^2$$

exists a. s. uniformly for $t \in [0, t]$; due to the uniform convergence, $\langle M \rangle_t$ is a continuous process.

For $s, t \in \Pi$ with $s < t$ there is some $k \geq 1$ such that $s, t \in \Pi_n(k)$. Obviously,

$$\sum_{t_{j-1}, t_j \in \Pi_n(k)} (M_{s \wedge t_j} - M_{s \wedge t_{j-1}})^2 \leq \sum_{t_{j-1}, t_j \in \Pi_n(k)} (M_{t \wedge t_j} - M_{t \wedge t_{j-1}})^2$$

and, as $k \rightarrow \infty$, this inequality also holds for the limiting process $\langle M \rangle$. Since Π is dense and $t \mapsto \langle M \rangle_t$ continuous, we get $\langle M \rangle_s \leq \langle M \rangle_t$ for all $s \leq t$.

From (A.20) we see that $M_t^2 - \sum_{t_{j-1}, t_j \in \Pi_{n(k)}} (M_{t \wedge t_j} - M_{t \wedge t_{j-1}})^2$ is a martingale. From Lemma 14.10 we know that the martingale property is preserved under L^2 limits, and so $M^2 - \langle M \rangle$ is a martingale. \square

Let us, finally, remove the assumption that M is bounded.

A.31 Theorem (Doob–Meyer decomposition. Meyer 1962, 1963). *For all square integrable martingales $M \in \mathcal{M}_T^{2,c}$ there exists a unique adapted, continuous and increasing process $(\langle M \rangle_t)_{t \leq T}$ such that $M^2 - \langle M \rangle$ is a martingale and $\langle M \rangle_0 = 0$.*

If $(\Pi_n)_{n \geq 1}$ is any refining sequence of partitions of $[0, T]$ with $|\Pi_n| \rightarrow 0$, then $\langle M \rangle$ is given by

$$\langle M \rangle_t = \lim_{n \rightarrow \infty} \sum_{t_{j-1}, t_j \in \Pi_n} (M_{t \wedge t_j} - M_{t \wedge t_{j-1}})^2$$

where the limit is uniform (on compact t -sets) in probability.

Proof. Define a sequence of stopping times $\tau_n := \inf \{t \leq T : |M_t| \geq n\}$, $n \geq 1$, ($\inf \emptyset = \infty$). Clearly, $\lim_{n \rightarrow \infty} \tau_n = \infty$, $M^{\tau_n} \in \mathcal{M}_T^{2,c}$ and $|M_t^{\tau_n}| \leq n$ for all $t \in [0, T]$. Since for any partition Π of $[0, T]$ and all $m \leq n$

$$\sum_{t_{j-1}, t_j \in \Pi} \left(M_{t \wedge \tau_m \wedge t_j}^{\tau_n} - M_{t \wedge \tau_m \wedge t_{j-1}}^{\tau_n} \right)^2 = \sum_{t_{j-1}, t_j \in \Pi} \left(M_{t \wedge t_j}^{\tau_m} - M_{t \wedge t_{j-1}}^{\tau_m} \right)^2,$$

the following process is well-defined

$$\langle M \rangle_t(\omega) := \langle M^{\tau_n} \rangle_t(\omega) \quad \text{for all } (\omega, t) : t \leq \tau_n(\omega) \wedge T$$

and we find that $\langle M \rangle_t$ is continuous and $\langle M \rangle_t^{\tau_n} = \langle M^{\tau_n} \rangle_t$.

Using Fatou's lemma, the fact that $(M^{\tau_n})^2 - \langle M^{\tau_n} \rangle$ is a martingale, and Doob's maximal inequality A.10, we have

$$\begin{aligned} \mathbb{E}(\langle M \rangle_T) &= \mathbb{E} \left(\lim_{n \rightarrow \infty} \langle M \rangle_T^{\tau_n} \right) \leq \lim_{n \rightarrow \infty} \mathbb{E}(\langle M^{\tau_n} \rangle_T) = \lim_{n \rightarrow \infty} \mathbb{E}((M_T^{\tau_n})^2 - M_0^2) \\ &\leq \mathbb{E} \left(\sup_{t \leq T} M_t^2 \right) \\ &\leq 4 \mathbb{E}(M_T^2). \end{aligned}$$

Therefore, $\langle M \rangle_T \in L^1(\mathbb{P})$. Moreover,

$$\mathbb{E} \left(\sup_{t \leq T} |\langle M \rangle_t - \langle M^{\tau_n} \rangle_t| \right) = \mathbb{E} \left(\sup_{t \leq T} |\langle M \rangle_t - \langle M \rangle_{\tau_n}| \mathbb{1}_{\{t > \tau_n\}} \right) \leq \mathbb{E}(\langle M \rangle_T \mathbb{1}_{\{T > \tau_n\}})$$

which shows, by the dominated convergence theorem, that $L^1\text{-}\lim_{n \rightarrow \infty} \langle M \rangle_t^{\tau_n} = \langle M \rangle_t$ uniformly for $t \in [0, T]$. Similarly,

$$\mathbb{E} \left(\sup_{t \leq T} (M_t - M_t^{\tau_n})^2 \right) = \mathbb{E} \left(\sup_{t \leq T} (M_t - M_{\tau_n})^2 \mathbb{1}_{\{t > \tau_n\}} \right) \leq 4 \mathbb{E} \left(\sup_{t \leq T} M_t^2 \mathbb{1}_{\{T > \tau_n\}} \right).$$

Since, by Doob's maximal inequality, $\sup_{t \leq T} M_t^2 \in L^1$, we can use dominated convergence to see that $L^1\text{-}\lim_{n \rightarrow \infty} (M_t^{\tau_n})^2 = M_t^2$ uniformly for all $t \in [0, T]$.

Since L^1 -convergence preserves the martingale property, $M^2 - \langle M \rangle$ is a martingale.

Using that $\langle M \rangle_t = \langle M \rangle_t^{\tau_n} = \langle M^{\tau_n} \rangle_t$ for $t \leq \tau_n$, we find for every $\epsilon > 0$

$$\begin{aligned} & \mathbb{P} \left(\sup_{t \leq T} \left| \langle M \rangle_t - \sum_{t_{j-1}, t_j \in \Pi} (M_{t \wedge t_j} - M_{t \wedge t_{j-1}})^2 \right| > \epsilon \right) \\ & \leq \mathbb{P} \left(\sup_{t \leq T} \left| \langle M^{\tau_n} \rangle_t - \sum_{t_{j-1}, t_j \in \Pi} (M_{t \wedge t_j}^{\tau_n} - M_{t \wedge t_{j-1}}^{\tau_n})^2 \right| > \epsilon, T \leq \tau_n \right) + \mathbb{P}(T > \tau_n) \\ & \leq \frac{1}{\epsilon^2} \mathbb{E} \left(\sup_{t \leq T} \left| \langle M^{\tau_n} \rangle_t - \sum_{t_{j-1}, t_j \in \Pi} (M_{t \wedge t_j}^{\tau_n} - M_{t \wedge t_{j-1}}^{\tau_n})^2 \right|^2 \right) + \mathbb{P}(T > \tau_n) \\ & \xrightarrow[|\Pi| \rightarrow 0]{\text{Thm. A.28}} \mathbb{P}(T > \tau_n) \xrightarrow{n \rightarrow \infty} 0. \end{aligned} \quad \square$$

A.32 Corollary. Let $(M_t, \mathcal{F}_t)_{t \geq 0} \subset L^2(\mathbb{P})$ be a continuous martingale with respect to a right-continuous filtration. Then M and the quadratic variation $\langle M \rangle$ are, with probability one, constant on the same t -intervals.

Proof. Fix $T > 0$. We have to show that there is a set $\Omega^* \subset \Omega$ such that $\mathbb{P}(\Omega^*) = 1$ and for all $[a, b] \subset [0, T]$ and $\omega \in \Omega^*$

$$M_t(\omega) = M_a(\omega) \text{ for all } t \in [a, b] \iff \langle M \rangle_t(\omega) = \langle M \rangle_a(\omega) \text{ for all } t \in [a, b].$$

“ \Rightarrow ”: By Theorem A.31 there is a set $\Omega' \subset \Omega$ of full \mathbb{P} -measure and a sequence of partitions $(\Pi_n)_{n \geq 1}$ of $[0, T]$ such that $\lim_{n \rightarrow \infty} |\Pi_n| = 0$ and

$$\langle M \rangle_t(\omega) = \lim_{n \rightarrow \infty} \sum_{t_{j-1}, t_j \in \Pi_n} (M_{t_j \wedge t}(\omega) - M_{t_{j-1} \wedge t}(\omega))^2, \quad \omega \in \Omega'.$$

If M is constant on $[a, b] \subset [0, T]$, we see that for all $t \in [a, b]$

$$\langle M \rangle_t(\omega) - \langle M \rangle_a(\omega) = \lim_{n \rightarrow \infty} \sum_{t_{j-1}, t_j \in \Pi_n \cap [a, b]} (M_{t_j \wedge t}(\omega) - M_{t_{j-1} \wedge t}(\omega))^2 = 0 \quad \omega \in \Omega'.$$

“ \Leftarrow ”: Let Ω' be as above. The idea is to show that for every rational $q \in [0, T]$ there is a set $\Omega_q \subset \Omega'$ with $\mathbb{P}(\Omega_q) = 1$ and such that for every $\omega \in \Omega_q$ a $\langle M \rangle$ -(ω)-constancy

interval of the form $[q, b]$ is also an interval of constancy for $M_*(\omega)$. Then the claim follows for the set $\Omega^* = \bigcap_{q \in \mathbb{Q} \cap [0, T]} \Omega_q$.

The shifted process $N_t := M_{t+q} - M_q$, $0 \leq t \leq T - q$ is a continuous L^2 martingale for the filtration $(\mathcal{F}_{t+q})_{q \in [0, T-q]}$; moreover, $\langle N \rangle_t = \langle M \rangle_{t+q} - \langle M \rangle_q$.

Set $\tau = \inf\{s > 0 : \langle N \rangle_s > 0\}$, note that τ is a stopping time (since $\langle N \rangle$ has continuous paths), and consider the stopped processes $\langle N^\tau \rangle_t = \langle N \rangle_{t \wedge \tau}$ and $N_t^\tau = N_{t \wedge \tau}$.

By definition, $\langle N^\tau \rangle_t \equiv 0$ a.s., hence $\langle M^\tau \rangle_t \equiv \langle M \rangle_q$ almost surely. Since $\mathbb{E}(N_{\tau \wedge t}^2) = \mathbb{E}\langle N^\tau \rangle_t = 0$, we conclude that $M_{t \wedge \tau + q} - M_q = N_t^\tau = 0$ a.s., i.e. there is a set Ω_q of full measure such that for every $\omega \in \Omega_q$ every interval of constancy $[q, b]$ of $\langle M \rangle$, is also an interval of constancy for M_* . On the other hand, the direction “ \Rightarrow ” shows that a constancy interval of M_* cannot be longer than that of $\langle M \rangle$, hence they coincide. \square

A.7 BV functions and Riemann–Stieltjes integrals

Here we collect some material on functions of bounded variation, BV-functions for short, and Riemann–Stieltjes integrals. Our main references are Hewitt & Stromberg [77, Chapters V.8, V.9, V.17], Kestelman [101, Chapter XI], Rudin [163, Chapter 6] and Riesz & Sz.-Nagy [159, Chapters I.4–9, III.49–60].

A.7.1 Functions of bounded variation

Let $f : [a, b] \rightarrow \mathbb{R}$ be a function and let $\Pi = \{t_0 = a < t_1 < \dots < t_n = b\}$ be a finite partition of the interval $[a, b] \subset \mathbb{R}$. By $|\Pi| := \max_{1 \leq j \leq n} (t_j - t_{j-1})$ we denote the *fineness* or *mesh* of Π . We call

$$S_1^\Pi(f; [a, b]) := \sum_{t_j, t_{j-1} \in \Pi} |f(t_j) - f(t_{j-1})| = \sum_{j=1}^n |f(t_j) - f(t_{j-1})| \quad (\text{A.21})$$

the *variation sum*. The supremum of the variation sums over all finite partitions

$$\text{VAR}_1(f; [a, b]) := \sup \{S_p^\Pi(f; t) : \Pi \text{ finite partition of } [a, b]\} \in [0, \infty] \quad (\text{A.22})$$

is the *strong variation* or *total variation*.

A.33 Definition. A function $f : [a, b] \rightarrow \mathbb{R}$ is said to be of *bounded variation* (also: *finite variation*) if $\text{VAR}_1(f; [a, b]) < \infty$. We denote by $\text{BV}[a, b]$ the set of all functions $f : [a, b] \rightarrow \mathbb{R}$ of bounded variation; if the domain is clear, we just write **BV**.

Let us collect a few properties of functions of bounded variation.

A.34 Properties of BV-functions. Let $f, g : [a, b] \rightarrow \mathbb{R}$ be real functions.

- a) If f is monotone, then $f \in \mathbf{BV}$ and $\text{VAR}_1(f; [a, b]) = |f(b) - f(a)|$;
- b) If $f \in \mathbf{BV}$, then $\sup_{t \in [a, b]} |f(t)| < \infty$;
- c) If $f, g \in \mathbf{BV}$, then $f \pm g$, $f \cdot g$ and $f/(|g| + \epsilon)$, $\epsilon > 0$, are of bounded variation;
- d) Let $a < c < b$. Then $\text{VAR}_1(f; [a, b]) = \text{VAR}_1(f; [a, c]) + \text{VAR}_1(f; [c, b])$ in $[0, \infty]$;
- e) $f \in \mathbf{BV}$ if, and only if, there exist two increasing functions $\phi, \psi : [a, b] \rightarrow \mathbb{R}$ such that $f = \phi - \psi$.
In particular, $f \in \mathbf{BV}[a, b]$ has for every $t \in (a, b)$ finite left- and right-hand limits $f(t \pm)$, and the number of discontinuities is at most countable.
- f) If $f \in \mathbf{BV}$ is continuous at t_0 , then $t \mapsto \text{VAR}_1(f; [a, t])$ is continuous at $t = t_0$;
- g) Every $f \in \mathbf{BV}$ has a decomposition $f(t) = f_c(t) + f_s(t)$ where $f_c \in \mathbf{BV} \cap \mathcal{C}$ and $f_s(t) = \sum_{a \leq s \leq t} [f(s+) - f(s-)]$ are uniquely determined;
- h) Every $f \in \mathbf{BV}$ has Lebesgue almost everywhere a finite derivative.

A.7.2 The Riemann–Stieltjes Integral

The Riemann–Stieltjes integral is a generalization of the classical Riemann integral. It was introduced by T. J. Stieltjes in his memoir [174, in particular: pp. 69–76] on continued fractions and the general moment problem. While Stieltjes considers only positive integrators (‘mass distributions’), it is F. Riesz [157] who remarked that one can consider integrators from \mathbf{BV} .

Let $f, \alpha : [a, b] \rightarrow \mathbb{R}$ and let $\Pi = \{t_0 = a < t_1 < \dots < t_n = b\}$ be a partition of the interval $[a, b] \subset \mathbb{R}$ with mesh $|\Pi|$; moreover, ξ_j^Π denotes an arbitrary point from the partition interval $[t_j, t_{j+1}]$, $j = 0, \dots, n-1$.

A.35 Definition. Let $\alpha : [a, b] \rightarrow \mathbb{R}$. A function $f : [a, b] \rightarrow \mathbb{R}$ is called (*Riemann–Stieltjes*) *integrable* if the *Stieltjes sums*

$$S^\Pi(f; \alpha) := \sum_{t_j, t_{j-1} \in \Pi} f(\xi_{j-1}^\Pi)(\alpha(t_j) - \alpha(t_{j-1})) \quad (\text{A.23})$$

converge, independently of the choice of the points $\xi_j^\Pi \in [t_j, t_{j+1}]$, $t_j, t_{j+1} \in \Pi$, to a finite limit as $|\Pi| \rightarrow 0$. The common value is called the *Riemann–Stieltjes integral* of f , and we write

$$\int_a^b f(t) d\alpha(t) := \lim_{|\Pi| \rightarrow 0} \sum_{t_j, t_{j-1} \in \Pi} f(\xi_{j-1}^\Pi)(\alpha(t_j) - \alpha(t_{j-1})). \quad (\text{A.24})$$

With $\alpha(t) = t$ we get the Riemann integral as a special case. Let us briefly collect the main features of the Riemann–Stieltjes integral.

A.36 Properties of the Riemann–Stieltjes integral. Let $f, \alpha : [a, b] \rightarrow \mathbb{R}$ be real functions and $a \leq c \leq b$.

- a) $(f, \alpha) \mapsto \int_a^b f(t) d\alpha(t)$ is bilinear;
- b) If $\int_a^b f(t) d\alpha(t)$ exists, so does $\int_c^d f(t) d\alpha(t)$ for all $a \leq c < d \leq b$;
- c) $\int_a^b f(t) d\alpha(t) = \int_a^c f(t) d\alpha(t) + \int_c^b f(t) d\alpha(t)$;
- d) $|\int_a^b f(t) d\alpha(t)| \leq \|f\|_\infty \cdot \text{VAR}_1(\alpha; [a, b])$;
- e) If f is piecewise continuous, $\alpha \in \text{BV}[a, b]$ and α and f have no common discontinuities, then f is Riemann–Stieltjes integrable;
- f) If f is Riemann–Stieltjes integrable, then f and α cannot have common discontinuities;
- g) If f is Riemann integrable and if α is absolutely continuous, then f is Riemann–Stieltjes integrable and $\int_a^b f(t) d\alpha(t) = \int_a^b f(t) \alpha'(t) dt$;
- h) If f is Riemann–Stieltjes integrable, then $x \mapsto \int_a^x f(t) d\alpha(t)$ is continuous at continuity points of α ;
- i) If α is right-continuous and increasing, then $\mu_\alpha(s, t] := \alpha(t) - \alpha(s)$ defines a positive measure on $\mathcal{B}[a, b]$ and $\int_a^b f(t) d\alpha(t) = \int_{(a, b]} f(t) \mu_\alpha(dt)$;

The property A.36 i) implies that we can embed the Riemann–Stieltjes integral into the Lebesgue theory if $\alpha \in \text{BV}$ is right-continuous. Observe that in this case we can make sure that $\alpha = \phi - \psi$ for *right-continuous* increasing functions ϕ, ψ . Then, $\mu_\alpha := \mu_\phi - \mu_\psi$ is a signed measure on $\mathcal{B}[a, b]$ with μ_ϕ, μ_ψ as in A.36 i). The integral $\int_{(a, b]} f(s) \mu_\alpha(ds)$ is sometimes called the *Lebesgue–Stieltjes integral*.

The integration by parts formula and the chain rule for Riemann–Stieltjes integrals are particularly simple. The chain rule is the deterministic counterpart of Itô’s lemma.

A.37 Theorem (Integration by parts). Let $\alpha, f : [a, b] \rightarrow \mathbb{R}$. If $\int_a^b f(t) d\alpha(t)$ exists, so does $\int_a^b \alpha(t) df(t)$; in this case

$$f(b)\alpha(b) - f(a)\alpha(a) = \int_a^b f(t) d\alpha(t) + \int_a^b \alpha(t) df(t).$$

A.38 Theorem (Chain rule). Let $\alpha \in \mathcal{C}[0, T]$ be increasing and $f \in \mathcal{C}^1[\alpha(0), \alpha(T)]$. Then

$$f(\alpha(t)) - f(\alpha(0)) = \int_0^t f'(\alpha(s)) d\alpha(s) = \int_{\alpha(0)}^{\alpha(t)} f'(x) dx \quad \text{for all } t \leq T.$$

Alternatively, if β is strictly increasing and $g \in \mathcal{C}[a, b]$, then

$$\int_a^b g(s) d\beta(s) = \int_{\beta(a)}^{\beta(b)} g(\beta^{-1}(x)) dx.$$

The property A.36 d) implies that $A[f] := \int_a^b f(t) d\alpha(t)$, $\alpha \in \mathbf{BV}$, is a continuous linear form on the Banach space $(\mathcal{C}[a, b], \|\cdot\|_\infty)$. In the already mentioned note [157] F. Riesz showed that *all* continuous linear forms are of this type.

A.39 Theorem (Riesz representation theorem). *Let $A : \mathcal{C}[a, b] \rightarrow \mathbb{R}$ be a continuous linear form, i. e. a linear map satisfying $|A[f]| \leq c_A \|f\|_\infty$. Then there is a function $\alpha \in \mathbf{BV}[a, b]$, depending only on A , such that $A[f] = \int_a^b f(t) d\alpha(t)$.*

In fact, it is not even necessary to assume that the linear form is *continuous*. As soon as $f \mapsto \int_a^b f(t) d\alpha(t)$ is finite for every $f \in \mathcal{C}[a, b]$, then α is necessarily of bounded variation. For this we need

A.40 Theorem (Banach–Steinhaus theorem). *Let $(\mathcal{X}, \|\cdot\|_\mathcal{X})$, $(\mathcal{Y}, \|\cdot\|_\mathcal{Y})$ be two Banach spaces and $S^\Pi : \mathcal{X} \rightarrow \mathcal{Y}$ be a family of linear maps indexed by Π . Then*

$$\sup_{\Pi} \|S^\Pi[x]\|_\mathcal{Y} < \infty, \quad \forall x \in \mathcal{X}, \quad \text{implies that} \quad \sup_{\Pi} \underbrace{\sup_{\|x\|_\mathcal{X} \leq 1} \|S^\Pi[x]\|_\mathcal{Y}}_{\text{operator norm } \|S^\Pi\|_{\mathcal{X} \rightarrow \mathcal{Y}}} < \infty.$$

A.41 Corollary. *Let $\alpha : [a, b] \rightarrow \mathbb{R}$. If $A[f] := \int_a^b f(t) d\alpha(t)$ exists for all continuous functions $f \in \mathcal{C}[a, b]$, then $\alpha \in \mathbf{BV}[a, b]$.*

Proof. We apply Theorem A.40 to

$$(\mathcal{X}, \|\cdot\|_\mathcal{X}) = (\mathcal{C}[a, b], \|\cdot\|_\infty) \quad \text{and} \quad (\mathcal{Y}, \|\cdot\|_\mathcal{Y}) = (\mathbb{R}, |\cdot|)$$

and $S^\Pi[f] := \sum_{t_j, t_{j-1} \in \Pi} f(t_{j-1})(\alpha(t_j) - \alpha(t_{j-1}))$ where

$$\Pi = \{t_0 = a < t_1 < \cdots < t_n = b\}$$

is a partition of $[a, b]$.

For the saw-tooth function

$$f_\Pi(s) := \begin{cases} \text{sgn}(\alpha(t_j) - \alpha(t_{j-1})), & \text{if } s = t_{j-1}, j = 1, \dots, n, \\ \text{piecewise linear,} & \text{at all other points,} \end{cases}$$

we find

$$S^\Pi[f_\Pi] = \sum_{t_j, t_{j-1} \in \Pi} |\alpha(t_j) - \alpha(t_{j-1})| \leq \sup_{\|f\|_\infty \leq 1} |S^\Pi[f]|.$$

By assumption $\lim_{|\Pi| \rightarrow 0} S^\Pi[f] = \int_a^b f(t) d\alpha(t) < \infty$ for all $f \in \mathcal{C}[a, b]$. Thus, $\sup_\Pi |S^\Pi[f]| < \infty$ and, by the Banach–Steinhaus theorem,

$$\text{VAR}_1(f; [a, b]) = \sup_\Pi S^\Pi[f_\Pi] \leq \sup_\Pi \sup_{\|f\|_\infty \leq 1} |S^\Pi[f]| < \infty. \quad \square$$

A.42 Remark. Corollary A.41 shows that the stochastic integral $\int_0^t f(s) dB_s(\omega)$ cannot be defined for every $\omega \in \Omega$ as a Riemann–Stieltjes integral, if the class of admissible integrands contains all (non-random!) continuous functions f . Since for continuous integrands the Riemann–Stieltjes and the Lebesgue–Stieltjes integrals coincide, it is also not possible to use a Lebesgue–Stieltjes approach.

A.8 Some tools from analysis

A.8.1 Gronwall’s lemma

Gronwall’s lemma is a well-known result from ordinary differential equations. It is often used to get uniqueness of the solutions of certain ODEs.

A.43 Theorem (Gronwall’s lemma). *Let $u, a, b : [0, \infty) \rightarrow [0, \infty)$ be positive measurable functions satisfying the inequality*

$$u(t) \leq a(t) + \int_0^t b(s)u(s) ds \quad \text{for all } t \geq 0. \quad (\text{A.25})$$

Then

$$u(t) \leq a(t) + \int_0^t a(s)b(s) \exp\left(\int_s^t b(r) dr\right) ds \quad \text{for all } t \geq 0. \quad (\text{A.26})$$

If $a(t) = a$ and $b(t) = b$ are constants, the estimate (A.26) simplifies and reads

$$u(t) \leq ae^{bt} \quad \text{for all } t \geq 0. \quad (\text{A.26}')$$

Proof. Set $y(t) := \int_0^t b(s)u(s) ds$. Since this is the primitive of $b(t)u(t)$ we can rewrite (A.25) for Lebesgue almost all $t \geq 0$ as

$$y'(t) - b(t)y(t) \leq a(t)b(t). \quad (\text{A.27})$$

If we set $z(t) := y(t) \exp\left(-\int_0^t b(s) ds\right)$ and substitute this into (A.27), we arrive at

$$z'(t) \leq a(t)b(t) \exp\left(-\int_0^t b(s) ds\right)$$

Lebesgue almost everywhere. On the other hand, $z(0) = y(0) = 0$, hence we get by

integration

$$z(t) \leq \int_0^t a(s)b(s) \exp\left(-\int_0^s b(r) dr\right) ds$$

or

$$y(t) \leq \int_0^t a(s)b(s) \exp\left(\int_s^t b(r) dr\right) ds$$

for all $t > 0$. This implies (A.26) since $u(t) \leq a(t) + y(t)$. \square

A.8.2 Completeness of the Haar functions

Denote by $(H_n)_{n \geq 0}$ the Haar functions introduced in Paragraph 3.2. We will show here that they are a complete orthonormal system in $L^2([0, 1], ds)$, hence a orthonormal basis. For some background on Hilbert spaces and orthonormal bases see, e. g. [169, Chapters 21, 22].

A.44 Theorem. *The Haar functions are a complete orthonormal basis of $L^2([0, 1], ds)$.*

Proof. We show that any $f \in L^2$ with $\langle f, H_n \rangle_{L^2} = \int_0^1 f(s)H_n(s) ds = 0$ for all $n \geq 0$ is Lebesgue almost everywhere zero. Indeed, if $\langle f, H_n \rangle_{L^2} = 0$ for all $n \geq 0$ we find by induction that

$$\int_{k/2^j}^{(k+1)/2^j} f(s) ds = 0 \quad \text{for all } k = 0, 1, \dots, 2^j - 1, j \geq 0. \quad (\text{A.28})$$

Let us quickly verify the induction step. Assume that (A.28) is true for some j ; fix $k \in \{0, 1, \dots, 2^j - 1\}$. The interval $I = [\frac{k}{2^j}, \frac{k+1}{2^j})$ supports the Haar function H_{2^j+k} , and on the left half of the interval, $I_1 = [\frac{k}{2^j}, \frac{2k+1}{2^{j+1}})$, it has the value $2^{j/2}$, on the right half, $I_2 = [\frac{2k+1}{2^{j+1}}, \frac{k+1}{2^j})$, it is $-2^{j/2}$. By assumption

$$\langle f, H_{2^j+k} \rangle_{L^2} = 0, \quad \text{thus,} \quad \int_{I_1} f(s) ds = \int_{I_2} f(s) ds.$$

From the induction assumption (A.28) we know, however, that

$$\int_{I_1 \cup I_2} f(s) ds = 0, \quad \text{thus,} \quad \int_{I_1} f(s) ds = - \int_{I_2} f(s) ds.$$

This is only possible if $\int_{I_1} f(s) ds = \int_{I_2} f(s) ds = 0$. Since I_1, I_2 are generic intervals of the next refinement, the induction is complete.

Because of (A.28) the primitive $F(t) := \int_0^t f(s) ds$ is zero for all $t = k2^{-j}$, $j \geq 0$, $k = 0, 1, \dots, 2^j$; since the dyadic numbers are dense in $[0, 1]$ and since $F(t)$ is continuous, we see that $F(t) \equiv 0$ and, therefore, $f \equiv 0$ almost everywhere. \square

A.8.3 A multinomial identity

The following non-standard binomial identity is useful in Lévy's characterization of Brownian motion, cf. Lemma 9.10 and Theorem 9.12.

A.45 Lemma. *Let $n \geq 1$ and $a_1, \dots, a_n \in \mathbb{R}$. Then*

$$\begin{aligned} & \left(\sum_j a_j \right)^4 + 3 \sum_j a_j^4 - 4 \left(\sum_j a_j^3 \right) \left(\sum_k a_k \right) + 2 \sum_{(j,k): j < k} a_j^2 a_k^2 \\ &= 2 \sum_j a_j^2 \left(\sum_{k: k \neq j} a_k \right)^2 + 4 \left(\sum_{(j,k): j < k} a_j a_k \right)^2 \end{aligned}$$

Proof. Expanding the term on the left we find

$$\begin{aligned} & \left(\sum_{(j,k): j < k} a_j a_k \right)^2 \\ &= \underbrace{\sum_j a_j^2 \left(\sum_{k: k \neq j} a_k \right)^2 - \sum_{(j,k): j < k} a_j^2 a_k^2}_{\mathbf{A}} + 6 \underbrace{\sum_{(j,k,l,m): j < k < l < m} a_j a_k a_l a_m}_{\mathbf{B}} \end{aligned}$$

Note that **A** contains all terms of the form a^2bc, ab^2c, abc^2 which appear *twice* and a^2b^2 which appear *only once*. (Here and in the sequel the letters a, b, c, d stand for mutually distinct elements.) Since the first term of **A** contains $a^2b^2 + b^2a^2$, i. e. with double multiplicity, we have to *subtract* one set of a^2b^2 terms. In **B** we collect all terms of the form $abcd$ arise in $\binom{4}{2,2} = 6$ ways as product of two pairs. This explains the multiplicity 6.

Again by expanding we get

$$\left(\sum_j a_j \right)^4 = -3 \sum_j a_j^4 + 4 \left(\sum_j a_j^3 \right) \left(\sum_k a_k \right) \quad (\text{A.29})$$

$$+ 6 \sum_j \sum_{(k,l): k \neq l, k \neq j, l \neq j} a_j^2 a_k a_l + 3 \sum_j \sum_{k: k \neq j} a_j^2 a_k^2 \quad (\text{A.30})$$

$$+ 24 \sum_{(j,k,l,m): j < k < l < m} a_j a_k a_l a_m. \quad (\text{A.31})$$

In the first line we collect all terms of the form a^4 and a^3b which occur $\binom{4}{4}$ and $\binom{4}{3,1}$ times, respectively. The second line contains all terms of the form a^2bc, ab^2c, abc^2 which have multiplicity $\binom{4}{2,1,1} = 12$ and a^2b^2 which have multiplicity $\binom{4}{2,2} = 6$. Since the sums do not distinguish between ab and ba , we get the coefficients 6 and 3.

Finally, line three contains all terms of the form $abcd$, each of which arises in exactly $\binom{4}{1,1,1,1} = 24$ ways.

Since the sum in (A.31) is just $4\mathbf{B}$, with \mathbf{B} from the beginning of the proof, we see

$$\begin{aligned}
 \left(\sum_j a_j\right)^4 &= -3 \sum_j a_j^4 + 4 \left(\sum_j a_j^3\right) \left(\sum_k a_k\right) \\
 &\quad + 6 \left\{ \sum_j a_j^2 \left(\sum_{k:k \neq j} a_k\right)^2 - \sum_j \sum_{k:k \neq j} a_j^2 a_k^2 \right\} + 3 \sum_j \sum_{k:k \neq j} a_j^2 a_k^2 \\
 &\quad + 4 \left\{ \left(\sum_{(j,k):j < k} a_j a_k\right)^2 - \sum_j a_j^2 \left(\sum_{k:k \neq j} a_k\right)^2 + \sum_{(j,k):j < k} a_j^2 a_k^2 \right\} \\
 &= -3 \sum_j a_j^4 + 4 \left(\sum_j a_j^3\right) \left(\sum_k a_k\right) - 2 \sum_{(j,k):j < k} a_j^2 a_k^2 \\
 &\quad + 2 \sum_j a_j^2 \left(\sum_{k:k \neq j} a_k\right)^2 + 4 \left(\sum_{(j,k):j < k} a_j a_k\right)^2. \quad \square
 \end{aligned}$$

Bibliography

- [1] Arnold, L.: *Random Dynamical Systems*. Springer, Berlin 2009.
- [2] Ash, R. B.: *Information Theory*. Interscience Publishers, New York 1965. Reprinted by: Dover, New York 1990.
- [3] Asmussen, S., Glynn, P. W.: *Stochastic Simulation: Algorithms and Analysis*. Springer, Berlin 2007.
- [4] Bachelier, L.: Théorie de la spéculation. *Ann. Sci. Ec. Norm. Super.* **17** (1900) 21–86. English translations in [31] pp. 17–78 and [6].
- [5] Bachelier, L.: *Calcul des Probabilités*. Gauthier–Villars, Paris 1912. Reprinted by Éditions Jacques Gabay, Paris 1992.
- [6] Bachelier, L., Davis, M. (ed.), Etheridge, A. (ed.): *Louis Bachelier's Theory of Speculation: The Origins of Modern Finance*. Princeton University Press, Princeton (NJ) 2006.
- [7] Bass, R.: *Probabilistic Techniques in Analysis*. Springer, New York 1995.
- [8] Bass, R.: *Diffusions and Elliptic Operators*. Springer, New York 1998.
- [9] Bell, D. R.: Stochastic differential equations and hypoelliptic operators. In: Rao, M. M. (ed.): *Real and Stochastic Analysis. New Perspectives*. Birkhäuser, Boston 2004, 9–42.
- [10] Billingsley, P.: *Convergence of Probability Measures*. Wiley, New York 1968. 2nd edn. Wiley, New York 1999.
- [11] Billingsley, P.: *Probability and Measure (3rd edn)*. Wiley, New York 1995.
- [12] Blumenthal, R. M.: An extended Markov property. *Trans. Am. Math. Soc.* (1957) **82** 52–72.
- [13] Blumenthal, R. M., Gettoor, R. K.: *Markov Processes and Potential Theory*. Academic Press, New York 1968. Reprinted by Dover, Mineola (NY) 2007.
- [14] Breiman, L.: On the tail behavior of sums of independent random variables. *Z. Wahrscheinlichkeitstheorie verw. Geb.* **9** (1967) 20–25.
- [15] Brown, R.: On the existence of active molecules in organic and inorganic bodies. *The Philosophical Magazine and Annals of Philosophy (new Series)* **4** (1828) 161–173.
- [16] Brown, R.: Additional remarks on active molecules. *The Philosophical Magazine and Annals of Philosophy (new Series)* **6** (1828) 161–166.
- [17] Bru, B., Yor, M.: Comments on the life and mathematical legacy of Wolfgang Doeblin. *Finance and Stochastics* **6** (2002) 3–47.
- [18] Burkholder, D. L.: Martingale Transforms. *Ann. Math. Statist.* (1966) **37** 1494–1504. Reprinted in [20].
- [19] Burkholder, D. L., Davis, B. J., Gundy, R. F.: Integral inequalities for convex functions of operators on martingales. In: Le Cam, L. *et al.* (eds.): *Proceedings of the 6th Berkeley Symposium on Mathematical Statistics and Probability*, Vol. II, pp. 223–240, University of California Press, Berkely (CA) 1972. Reprinted in [20].
- [20] Burkholder, D. L., Davis, B. (ed.), Song, R. (ed.): *Selected Works of Donald L. Burkholder*. Springer, New York 2011.

- [21] Chentsov, N. N.: Weak convergence of stochastic processes whose trajectories have no discontinuities of the second kind and the “heuristic” approach to the Kolmogorov-Smirnov tests. *Theor. Probab. Appl.* **1** (1956), 140–144.
- [22] Chung, K. L.: On the maximum partial sums of sequences of independent random variables. *Trans. Am. Math. Soc.* **61** (1948) 205–233.
- [23] Chung, K. L.: *Lectures from Markov Processes to Brownian Motion*. Springer, New York 1982. Enlarged and reprinted as [24].
- [24] Chung, K. L., Walsh, J. B.: *Markov Processes, Brownian Motion, and Time Symmetry (Second Edition)*. Springer, Berlin 2005. This is the 2nd edn. of [23].
- [25] Chung, K. L., Zhao, Z.: *From Brownian Motion to Schrödinger’s Equation*. Springer, Berlin 1995.
- [26] Ciesielski, Z.: On Haar functions and on the Schauder basis of the space $C_{(0,1)}$. *Bull. Acad. Polon. Sci. Sér. Math. Astronom. Phys.* **7** (1959) 227–232.
- [27] Ciesielski, Z.: *Lectures on Brownian Motion, Heat Conduction and Potential Theory*. Aarhus Matematisk Institut, Lecture Notes Ser. **3**, Aarhus 1966.
- [28] Ciesielski, Z., Kerkycharian, G., Roynette, B.: Quelques espaces fonctionnels associés à des processus gaussiens. *Studia Math.* **107** (1993) 171–204.
- [29] Cohen, L.: The history of noise. *IEEE Signal Processing Magazine* (November 2005) 20–45.
- [30] Cont, R., Tankov, P.: *Financial Modelling With Jump Processes*. Chapman & Hall/CRC, Boca Raton 2004.
- [31] Cootner, P. (ed.): *The Random Character of Stock Market Prices*. The MIT Press, Cambridge (MA) 1964.
- [32] Csörgő, M., Révész, P.: *Strong Approximations in Probability and Statistics*. Academic Press, New York 1981.
- [33] Dambis, K. E.: On the decomposition of continuous submartingales. *Theory Probab. Appl.* **10** (1965) 401–410.
- [34] DaPrato, G., Zabczyk, J.: *Stochastic Equations in Infinite Dimensions*. Cambridge University Press, Cambridge 1992.
- [35] Dembo, A., Zeitouni, O.: *Large Deviations Techniques and Applications (2nd ed.)*. Springer, New York 1993.
- [36] Deuschel, J.-D., Stroock, D. W.: *Large Deviations*. Academic Press 1989.
- [37] Doeblin, W.: Sur l’équation de Kolmogoroff. *C.R. Acad. Sci. Paris* **331** (2000) 1059–1102. (First printing of the *pli cacheté* deposited with the *Académie des Sciences* in 1940)
- [38] Doléans-Dade, C.: On the existence and unicity of solutions of stochastic integral equations. *Z. Wahrscheinlichkeitstheorie verw. Geb.* **36** (1976) 93–101.
- [39] Donsker, M.: An invariance principle for certain probability limit theorems. *Mem. Am. Math. Soc.* **6** (1951) 12 pp. (paper no. 4)
- [40] Doob, J. L.: *Stochastic Processes*. Interscience Publishers, New York 1953. Reprinted by Wiley, New York 1990.
- [41] Doob, J. L.: A probability approach to the heat equation. *Trans. Am. Math. Soc.* **80** (1955) 216–280.
- [42] Doob, J. L.: *Classical Potential Theory and Its Probabilistic Counterpart*. Springer, Berlin 1984.
- [43] Dubins, L. E., Schwarz, G.: On continuous martingales. *Proc. Natl. Acad. Sci. USA* **53** (1965) 913–916.

- [44] Dudley, R. M.: Sample functions of the Gaussian process. *Ann. Probab.* **1** (1973) 66–103. Reprinted in [46].
- [45] Dudley, R. M.: *Uniform Central Limit Theorems*. Cambridge University Press, Cambridge 1999.
- [46] Dudley, R. M., Giné, E. (ed.) et al.: *Selected Works of R. M. Dudley*. Springer, New York 2010.
- [47] Dudley, R. M., Norvaiša, R.: *An Introduction to p -variation and Young integrals*. MaPhySto Lecture Notes **1**, Aarhus 1998.
- [48] Dudley, R. M., Norvaiša, R.: *Differentiability of Six Operators on Nonsmooth Functions and p -Variation*. Springer, Lecture Notes in Mathematics **1703**, Berlin 1999.
- [49] Durrett, R.: *Brownian Motion and Martingales in Analysis*. Wadsworth Advanced Books & Software, Belmont (CA) 1984.
- [50] Durrett, R.: *Probability: Theory and Examples (4th edn)*. Cambridge University Press, Cambridge 2010.
- [51] Dynkin, E. B.: Criteria of continuity and of absence of discontinuities of the second kind for trajectories of a Markov stochastic process. *Izv. Akad. Nauk SSSR, Ser. Matem.* **16** (6) (1952) 563–572. (In Russian)
- [52] Dynkin, E. B.: Infinitesimal operators of Markov processes. *Theor. Probab. Appl.* **1** (1956) 34–54. Reprinted in [56]
- [53] Dynkin, E. B.: *Die Grundlagen der Theorie der Markoffschen Prozesse*. Springer, Berlin 1961. English translation: *Theory of Markov Processes*, Dover, Mineola (NY) 2006.
- [54] Dynkin, E. B.: *Markov Processes (2 vols.)*. Springer, Berlin 1965.
- [55] Dynkin, E. B., Yushkevich, A. A.: *Markov Processes. Theorems and Problems*. Plenum Press, New York 1969. Russian Original, German Translation: *Sätze und Aufgaben über Markoffsche Prozesse*, Springer 1969.
- [56] Dynkin, E. B., Yushkevich, A. A. (ed.) et al.: *Selected Papers of E. B. Dynkin with Commentary*. American Mathematical Society, Providence (RI) 2000.
- [57] Einstein, A.: Über die von der molekularkinetischen Theorie der Wärme geforderte Bewegung von in ruhenden Flüssigkeiten suspendierten Teilchen. *Ann. Phys.* **17** (1905) 549–560. Reprinted in [69], English translation in [58].
- [58] Einstein, A., Fürth, R. (ed.): *Investigations on the Theory of the Brownian movement*. Dover, New York 1956.
- [59] Eliot, G.: *Middlemarch*. William Blackwood and Sons, Edinburgh 1871.
- [60] Engelbert, H.-J., Cherny, A. S.: *Singular Stochastic Differential Equations*. Springer, Lecture Notes in Mathematics **1858**, Berlin 2004.
- [61] Ethier, S. N., Kurtz, T. G.: *Markov Processes. Characterization and Convergence*. Wiley, New York 1986.
- [62] Evans, L. C.: *Partial Differential Equations*. American Mathematical Society, Providence (RI) 1998.
- [63] Feyel, D., de La Pradelle, A.: On fractional Brownian processes. *Potential Anal.* **10** (1999) 273–288.
- [64] Fokker, A. D.: *Die mittlere Energie rotierender elektrischer Dipole im Strahlungsfeld*. *Ann. d. Phys.* **43** (1914) 810–820.
- [65] Freedman, D.: *Brownian Motion and Diffusion*. Holden-Day, San Francisco (CA) 1971. Reprinted by Springer, New York 1982.
- [66] Freidlin, M.: *Functional Integration and Partial Differential Equations*. Princeton University Press, Ann. Math. Studies **109**, Princeton (NJ) 1985.

- [67] Freidlin, M. I., Wentzell, A. D.: *Random Perturbations of Dynamical Systems (Second Edition)*. Springer, New York 1998.
- [68] Friedman, A.: *Partial Differential Equations of Parabolic Type*. Prentice-Hall, Englewood Cliffs (NJ) 1964. Reprinted by Dover, Mineola (NY) 2008.
- [69] Fürth, R. (ed.), Einstein, A., Smoluchowski, M. von: *Untersuchungen über die Theorie der Brownschen Bewegung – Abhandlung über die Brownsche Bewegung und verwandte Erscheinungen*. Harri Deutsch, Ostwalds Klassiker Bd. **199** (reprint of the separate volumes 199 and 207), Frankfurt am Main 1997. English translation [58].
- [70] Fukushima, M., Osima, Y., Takeda, M.: *Dirichlet Forms and Symmetric Markov Processes (2nd edn)*. De Gruyter, Berlin 2011.
- [71] Garsia, A. M., Rodemich, E., Rumsey, H.: A real variable lemma and the continuity of paths of some Gaussian processes. *Indiana Univ. Math. J.* **20** (1970) 565–578.
- [72] Gikhman, I. I., Skorokhod, A. W.: *Stochastic Differential Equations*. Springer, Berlin 1972.
- [73] Girsanov, I. V.: On transforming a certain class of stochastic processes by absolutely continuous substitution of measures. *Theor. Probab. Appl.* **5** (1960) 285–301.
- [74] Gradshteyn, I. S., Ryzhik, I. M.: *Table of Integrals, Series, and Products (4th corrected and enlarged edn)*. Academic Press, San Diego 1980.
- [75] Helms, L. L.: *Introduction to Potential Theory*. Wiley–Interscience, New York 1969.
- [76] Helms, L. L.: *Potential Theory*. Springer, London 2009. Extended version of [75].
- [77] Hewitt, E., Stromberg, K.: *Real and Abstract Analysis*. Springer, New York 1965.
- [78] Hille, E.: *Functional Analysis and Semi-groups*. Am. Math. Society, New York 1948. (2nd ed. 1957 jointly with R. S. Phillips)
- [79] Ikeda, N., Nakao, S., Yamato, Y.: A class of approximations of Brownian motion. *Publ. RIMS Kyoto Univ.* **13** (1977) 285–300.
- [80] Ikeda, N., Watanabe, S.: *Stochastic Differential Equations and Diffusion Processes (Second Edn)*. North-Holland, Amsterdam 1989.
- [81] Itô, K.: Differential equations determining a Markoff process (in Japanese). *J. Pan-Japan Math. Coll.* **1077** (1942). Translation reprinted in [87].
- [82] Itô, K.: Semigroups in Probability Theory. In: *Functional Analysis and Related Topics, 1991. Proceedings of the International Conference in Memory of Professor Kôzaku Yosida held at RIMS, Kyoto University, Japan, July 29 - Aug. 2, 1991*. Springer, Lecture Notes in Mathematics **1540**, Berlin 1993, 69–83.
- [83] Itô, K.: *Stochastic Processes. Lectures Given at Aarhus University*. Springer, Berlin 2004. Reprint with commentaries of Itô's 1969 Aarhus Lecture Notes Series **16** volume.
- [84] Itô, K.: *Essentials of Stochastic Processes*. Am. Math. Society, Providence (RI) 2006.
- [85] Itô, K., McKean, H. P.: *Diffusion Processes and Their Sample Paths*, Springer, Berlin 1965.
- [86] Itô, K., Nisio, M.: On the convergence of sums of independent Banach space valued random variables. *Osaka J. Math.* **5** (1968) 35–48.
- [87] Itô, K., Stroock, D. W. (ed.), Varadhan, S. R. S. (ed.): *Kiyosi Itô. Selected Papers*. Springer, New York 1987.
- [88] Jacod, J., Protter, P.: *Probability Essentials*. Springer, Berlin 2000.
- [89] John, F.: *Partial Differential Equations (4th edn)*. Springer, Berlin 1982.
- [90] Kac, M.: On distributions of certain Wiener functionals. *Trans. Am. Math. Soc.* (1949) **65** 1–13. Reprinted in [92].

- [91] Kac, M.: On some connections between probability theory and differential and integral equations. In: Neyman, J. (ed.): *Proceedings of the Second Berkely Symposium on Mathematical Statistics and Probability, August 1950*, University of California Press, Berkeley 1951, 189–225. Reprinted in [92].
- [92] Kac, M., Baclawski, K. (ed.), Donsker, M. D. (ed.): *Mark Kac: Probability, Number Theory, and Statistical Physics*. The MIT Press, Cambridge (MA) 1979.
- [93] Kahane, J.-P.: Sur l'irrégularité locale du mouvement brownien. *C. R. Acad. Sci, Paris Sér. A* **278** (1974) 331–333.
- [94] Kahane, J.-P.: Sur les zéros et les instants de ralentissement du mouvement brownien. *C. R. Acad. Sci, Paris Sér. A* **282** (1976) 431–433.
- [95] Kahane, J.-P.: *Some random series on functions (2nd edn)*. Cambridge University Press, Cambridge 1985.
- [96] Kakutani, S.: On Brownian motions in n -space. *Proc. Imperial Acad. Japan* (1944) **20** 648–652. Reprinted in [98].
- [97] Kakutani, S.: Two-dimensional Brownian motion and harmonic functions. *Proc. Imperial Acad. Japan* (1944) **20** 706–714. Reprinted in [98].
- [98] Kakutani, S., Kallman, R. R. (ed.): *Shizuo Kakutani – Selected Papers (2 vols.)*. Birkhäuser, Boston 1986.
- [99] Karatzas, I., Shreve, S. E.: *Brownian Motion and Stochastic Calculus*. Springer, New York 1988.
- [100] Kellogg, O. D.: *Foundations of Modern Potential Theory*. Springer, Berlin 1929. Reprinted by Dover, New York 1956.
- [101] Kestelman, H.: *Modern Theories of Integration*. Dover, New York 1960 (2nd edn).
- [102] Khintchine, A.: *Asymptotische Gesetze der Wahrscheinlichkeitsrechnung*. Springer, *Ergeb. Math. Grenzgeb.* Bd. **2**, Heft 4, Berlin 1933.
- [103] Kinney, J. R.: Continuity criteria of sample functions of Markov processes. *Trans. Am. Math. Soc.* **74** (1953) 280–302.
- [104] Kloeden, P. E., Platen, E.: *Numerical Solution of Stochastic Differential Equations*. Springer, Berlin 1992.
- [105] Kolmogoroff, A. N., Über die analytischen Methoden in der Wahrscheinlichkeitsrechnung. *Math. Ann.* **104** (1931) 415–458. English translation in [107].
- [106] Kolmogoroff, A. N., *Grundbegriffe der Wahrscheinlichkeitsrechnung*. Springer, *Ergeb. Math. Grenzgeb.* Bd. **2**, Heft 3, Berlin 1933.
- [107] Kolmogorov, A. N., Shiryaev, A. N. (ed.): *Selected Works of A. N. Kolmogorov (3 vols.)*. Kluwer, Dordrecht 1993.
- [108] Kowalski, E.: Bernstein polynomials and Brownian Motion. *Am. Math. Monthly* (2006) **113** 865–886.
- [109] Krylov, N. V.: *Introduction to the Theory of Diffusion Processes*. American Mathematical Society, Providence (RI) 1995.
- [110] Krylov, N. V.: *Introduction to the Theory of Random Processes*. American Mathematical Society, Providence (RI) 2002.
- [111] Kunita, H.: *Stochastic Flows and Stochastic Differential Equations*. Cambridge University Press, Cambridge 1990.
- [112] Kunita, H.: Stochastic differential equations based on Lévy processes and stochastic flows of diffeomorphisms. In: Rao, M. M. (ed.): *Real and Stochastic Analysis. New Perspectives*. Birkhäuser, Boston 2004, 305–373.

- [113] Kunita, H., Watanabe, S.: On square integrable martingales. *Nagoya Math. J.* **30** (1967) 209–245.
- [114] Kuo, H.-H.: *Gaussian Measures in Banach Spaces*. Springer, Lecture Notes in Mathematics **463**, Berlin 1975.
- [115] Lebesgue, H.: Sur des cas d'impossibilité du problème du Dirichlet ordinaire. *C. R. Séances Soc. Math. France* **41** (1913). Reprinted in [116, Vol. IV, pp. 125–126]
- [116] Lebesgue, H.: *Œuvres Scientifiques En Cinq Volumes (5 vols.)*. L'enseignement mathématique & Institut de Mathématiques Université de Genève, Geneva 1973.
- [117] L'Ecuyer, P.: Uniform random number generation. *Ann. Oper. Res.* (1994) **53**, 77–120.
- [118] L'Ecuyer, P., Panneton, F.: Fast random number generators based on linear recurrences modulo 2: overview and comparison. In: Kuhl, M. E., Steiger, N. M., Armstrong, F. B., Joines, J. A. (eds.): *Proceedings of the 37th Winter Simulation Conference, Orlando, FL, 2005*. ACM 2005. IEEE, New York 2006, 110–119.
- [119] Lévy, P.: Sur certains processus stochastiques homogènes. *Compositio Math.* **7** (1939) 283–339. Reprinted in [122, vol. 4].
- [120] Lévy, P.: Le mouvement brownien plan. *Am. J. Math.* **62** (1940) 487–550. Reprinted in [122, vol. 5].
- [121] Lévy, P.: *Processus Stochastiques et Mouvement Brownien*. Gauthier-Villars, Paris 1948.
- [122] Lévy, P., Dugué, D. (ed.): *Œuvres de Paul Lévy (6 vols.)*. Gauthier-Villars, Paris 1973–1980.
- [123] Lifshits, M. A.: *Gaussian Random Functions*. Kluwer, Dordrecht 1995.
- [124] Lumer, G., Phillips, R. S.: Dissipative operators in a Banach space. *Pacific J. Math.* **11** (1961) 679–698.
- [125] MacDonald, D. K. C.: *Noise and Fluctuations*. Wiley, New York 1962. Reprinted by Dover, Mineola (NY) 2006.
- [126] Malliavin, P.: Stochastic calculus of variation and hypoelliptic operators. In: Itô, K. (ed.): *Proceedings Intl. Symposium on SDEs, Kyoto 1976*, Wiley, New York 1978, 195–263.
- [127] Malliavin, P.: *Integration and Probability*. Springer, Berlin 1995.
- [128] Malliavin, P.: *Stochastic Analysis*. Springer, Berlin 1997.
- [129] Marcus, M., Rosen, J.: *Markov Processes, Gaussian Processes, and Local Times*. Cambridge University Press, Cambridge 2006.
- [130] Mazo, R. M.: *Brownian Motion. Fluctuations, Dynamics and Applications*. Oxford University Press, Oxford 2002.
- [131] Meyer, P. A.: A decomposition theorem for supermartingales. *Illinois J. Math.* **6** (1962) 193–205.
- [132] Meyer, P. A.: Decompositions of supermartingales. The uniqueness theorem. *Illinois J. Math.* **7** (1963) 1–17.
- [133] Mörters, P., Peres, Y.: *Brownian Motion*. Cambridge University Press, Cambridge 2010.
- [134] Motoo, M.: Proof of the law of iterated logarithm through diffusion equation. *Ann. Inst. Math. Statist.* **10** (1958) 21–28.
- [135] Nelson, E.: *Dynamical Theories of Brownian Motion*. Princeton University Press, Princeton (NJ) 1967.
- [136] Novikov, A. A.: On an identity for stochastic integrals. *Theory Probab. Appl.* **16** (1972) 717–720.

- [137] Oblój, J.: The Skorokhod embedding problem and its offspring. *Probab. Surveys* **1** (2004) 321–390.
- [138] Orey, S.: Probabilistic methods in partial differential equations. In: Littman, W. (ed.): *Studies in Partial Differential Equations*. The Mathematical Association of America, 1982, 143–205.
- [139] Orey, S., Taylor, S. J.: How often on a Brownian path does the law of iterated logarithm fail? *Proc. London Math. Soc.* **28** (1974) 174–192.
- [140] Paley, R. E. A. C., Wiener, N., Zygmund, A.: Notes on random functions. *Math. Z.* **37** (1933) 647–668.
- [141] Partzsch, L.: *Vorlesungen zum eindimensionalen Wienerischen Prozeß*. Teubner, Texte zur Mathematik Bd. **66**, Leipzig 1984.
- [142] Pazy, A.: *Semigroups of Linear Operators and Applications to Partial Differential Equations*. Springer, New York 1983.
- [143] Petrowsky, I.: Zur ersten Randwertaufgabe der Wärmeleitungsgleichung. *Compositio Math.* **1** (1935) 383–419.
- [144] Perrin, J.: Mouvement brownien et réalité moléculaire. *Ann. Chimie Physique (8ème sér.)* **18** (1909) 5–114.
- [145] Perrin, J.: *Les Atomes*. Librairie Félix Alcan, Paris 1913.
- [146] Pinsky, M. A.: *Introduction to Fourier Analysis and Wavelets*. Brooks/Cole, Pacific Grove (CA) 2002.
- [147] Planck, M.: Über einen Satz der statistischen Dynamik und seine Erweiterung in der Quantentheorie. *Sitzungsber. Königl. Preuß. Akad. Wiss. phys.-math. Kl.* **10.5** (1917) 324–341.
- [148] Poincaré, H.: Sur les équations aux dérivées partielles de la physique mathématique. *Am. J. Math.* **12** (1890) 211–294.
- [149] Port, S., Stone, S.: *Brownian Motion and Classical Potential Theory*. Academic Press, New York 1978.
- [150] Prevôt, C., Röckner, M.: *A Concise Course on Stochastic Partial Differential Equations*. Springer, Lecture Notes in Mathematics **1905**, Berlin 2007.
- [151] Protter, P. E.: *Stochastic Integration and Differential Equations (Second Edn.)*. Springer, Berlin 2004.
- [152] Rao, M.: *Brownian Motion and Classical Potential Theory*. Aarhus Universitet Matematisk Institut, Lecture Notes Series **47**, Aarhus 1977.
- [153] Rényi, A.: *Probability Theory*. North Holland, Amsterdam 1970. Reprinted by Dover, Mineola (NY) 2007.
- [154] Reuter, G. E. H.: Denumerable Markov processes and the associated contraction semi-groups on l . *Acta Math.* **97** (1957) 1–46.
- [155] Révész, P.: *Random Walk in Random and Non-Random Environments (2nd ed.)*. World Scientific, New Jersey 2005.
- [156] Revuz, D., Yor, M.: *Continuous Martingales and Brownian Motion*. Springer, Berlin 2005 (3rd edn.).
- [157] Riesz, F.: Sur les opérations fonctionnelles linéaires. *C. R. Acad. Sci. Paris* **149** (1909) 974–977. Reprinted in [158, pp. 400–403].
- [158] Riesz, F., Császár, Á. (ed.): *Frédéric Riesz – Œuvres complètes*. Gauthier-Villars, Paris 1960.
- [159] Riesz, F., Sz. Nagy, B.: *Leçons d'Analyse Fonctionnelle*. Akadémiai Kiadó, Budapest 1952. English translation: *Functional Analysis*. Dover, Mineola (NY) 1990.

- [160] Ripley, B. D.: *Stochastic Simulation*. Wiley, Series in Probab. Stat., Hoboken (NJ) 1987.
- [161] Rogers, L. C. G., Williams, D.: *Diffusions, Markov Processes, and Martingales (2nd edn, 2 vols.)*. Wiley, Chichester 1987 & 1994. Reprinted by Cambridge University Press, Cambridge 2000.
- [162] Rubinstein, R. Y., Kroese, D. P.: *Simulation and the Monte Carlo Method*. Wiley, Hoboken (NJ) 2008.
- [163] Rudin, W.: *Principles of Mathematical Analysis*. McGraw-Hill, Auckland 1976 (3rd edn).
- [164] Rudin, W.: *Real and Complex Analysis*. McGraw-Hill, New York 1986 (3rd edn).
- [165] Sawyer, S.: The Skorokhod Representation. *Rocky Mountain J. Math.* **4** (1974) 579–596.
- [166] Schilder, M.: Some asymptotic formulae for Wiener integrals. *Trans. Am. Math. Soc.* **125** (1966) 63–85.
- [167] Slutsky, E. E.: Qualche proposizione relativa alla teoria delle funzioni aleatorie. *Giorn. Ist. Ital. Attuari* **8** (1937) 183–199.
- [168] Schilling, R. L.: Sobolev embedding for stochastic processes. *Expo. Math.* **18** (2000) 239–242.
- [169] Schilling, R. L.: *Measures, Integrals and Martingales*. Cambridge University Press, Cambridge 2005.
- [170] Schilling, R. L., Song, R., Vondraček, Z.: *Bernstein Functions. Theory and Applications*. De Gruyter, Berlin 2010.
- [171] Shigekawa, I.: *Stochastic Analysis*. American Mathematical Society, Providence (RI) 2004.
- [172] Skorokhod, A. V.: *Studies In The Theory Of Random Processes*. Addison-Wesley, Reading (MA) 1965.
- [173] von Smoluchowski, M.: Zur kinetischen Theorie der Brownschen Molekularbewegung und der Suspensionen. *Annalen der Physik (4. Folge)* **21** (1906) 756–780. Reprinted in [69], English translation in [58].
- [174] Stieltjes, T. J.: Recherches sur les fractions continues. *Ann. Fac. Sci. Toulouse* **8** (1894) 1–122 and **9** (1895) 1–47. Translated and reprinted in [175] vol. 2.
- [175] Stieltjes, T. J., van Dijk, G. (ed.): *Thomas Jan Stieltjes - Œuvres Complètes – Collected Papers (2 vols.)*. Springer, Berlin 1993.
- [176] Strassen, V.: An invariance principle for the law of the iterated logarithm. *Z. Wahrscheinlichkeitstheorie verw. Geb.* **3** (1964) 211–246.
- [177] Stroock, D. W., Varadhan, S. R. S.: *Multidimensional Diffusion Processes*. Springer, Berlin 1979.
- [178] Tanaka, H.: Note on continuous additive functionals of the 1-dimensional Brownian path. *Z. Wahrscheinlichkeitstheor. verw. Geb.* **1** (1963) 251–257. Reprinted in [179]
- [179] Tanaka, H., Maejima, M. (ed.), Shiga, T. (ed.): *Stochastic Processes. Selected Papers of Hiroshi Tanaka*. World Scientific, New Jersey 2002.
- [180] Taylor, S. J.: Regularity of irregularities on a Brownian path. *Ann. Inst. Fourier* **24.2** (1974) 195–203.
- [181] van Hemmen, J. L., Ando, T.: An inequality for trace ideals. *Commun. Math. Phys.* **76** (1980) 143–148.
- [182] Varadhan, S. R. S.: *Large Deviations and Applications*. SIAM, CBMS-NSF Lecture Notes **46**, Philadelphia (PA) 1984.
- [183] Wald, A.: On cumulative sums of random variables. *Ann. Math. Statist.* (1944) **15** 283–296.

- [184] Wentzell, A. D.: *Theorie zufälliger Prozesse*. Birkhäuser, Basel 1979. Russian original. English translation: *Course in the Theory of Stochastic Processes*, McGraw-Hill 1981.
- [185] Wiener, N.: Differential-space. *J. Math. & Phys.* (1923) **58**, 131–174. Reprinted in [188] vol. 1.
- [186] Wiener, N.: Un problème de probabilités dénombrables. *Bull. Soc. Math. France* (1924) **11**, 569–578. Reprinted in [188] vol. 1.
- [187] Wiener, N., Siegel, A., Rankin, B., Martin, W. T.: *Differential Space, Quantum Systems, and Prediction*. The MIT Press, Cambridge (MA) 1966.
- [188] Wiener, N., Masani, P. (ed.): *Collected Works With Commentaries (4 vols)*. The MIT Press, Cambridge (MA) 1976–86.
- [189] Williams, D.: *Probability with Martingales*. Cambridge University Press, Cambridge 1991.
- [190] Wong, E., Zakai, M.: Riemann–Stieltjes Approximations of Stochastic Integrals, *Z. Wahrscheinlichkeitstheor. verw. Geb.* **12** (1969) 87–97.
- [191] Yamasaki, Y.: *Measures on Infinite Dimensional Spaces*. World Scientific, Singapore 1985.
- [192] Yosida, K.: On the differentiability and the representation of one-parameter semigroup of linear operators. *J. Math. Soc. Japan* **1** (1948) 244–253. Reprinted in [193].
- [193] Yosida, K., Itô, K. (ed.): *Kôzaku Yosida Collected Papers*. Springer, Tokyo 1992.
- [194] Zaremba, S.: Sur le principe de Dirichlet. *Acta Math.* (1911) **34**, 293–316.

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